Getting started with the glmmADMB package

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1 Introduction/quick start

glmmADMB is a package, built on the open source AD Model Builder nonlinear fitting engine, for fitting generalized linear mixed models and extensions.

- response distributions: Poisson, binomial, negative binomial (NB1 and NB2 parameterizations), Gamma, Beta, truncated Poisson and negative binomial; Gaussian coming soon
- link functions: log, logit, probit, cloglog, inverse, identity
- zero-inflation (models with a constant zero-inflation value only); hurdle models via truncated Poisson/NB
- single or multiple (nested or crossed) random effects
- offsets
- post-fit MCMC chain for characterizing uncertainty

As of version 0.6.5, the package has been greatly revised to allow a wider range of response and link functions and to allow models with multiple random effects. For now, the resulting package is slower than the old (single-random-effect version), but we hope to increase its speed in the future.

In order to use glmmADMB effectively you should already be reasonably familiar with generalized linear mixed models (GLMMs), which in turn requires familiarity with (i) generalized linear models (e.g. the special cases of logistic, binomial, and Poisson regression) and (ii) 'modern' mixed models (those working via maximization of the marginal likelihood rather than by manipulating sums of squares).

In order to fit a model in glmmADMB you need to:

• specify a model for the fixed effects, in the standard R (Wilkinson-Rogers) formula notation (see ?formula or Section 11.1 of the Introduction to R. Formulae can also include offsets.

• specify a model for the random effects, in the notation that is common to the nlme and lme4 packages. Random effects are specified as e|g, where e is an effect and g is a grouping factor (which must be a factor variable, or a nesting of/interaction among factor variables). For example, the formula would be 1|block for a random-intercept model or time|block for a model with random variation in slopes through time across groups specified by block. A model of nested random effects (block within site) would be 1|site/block; a model of crossed random effects (block and year) would be (1|block)+(1|year).

Random effects can be specified either in a separate random argument (as in nlme) or as part of the model formula (as in lme4).

- choose the error distribution by specifying the family (as a string: e.g. "poisson" or "binomial")
- specify a link function (as a string: e.g. "logit" or "log".
- optionally specify that zero-inflation is present zeroInflation=TRUE. In the current version, zero-inflation can only be specified as a single constant term across the entire model i.e. it cannot vary across groups or with covariates.

2 Owls data

These data, taken from [3] and ultimately from [2], quantify the number of negotiations among owlets (owl chicks) in different nests *prior* to the arrival of a provisioning parent as a function of food treatment (deprived or satiated), the sex of the parent, and arrival time. The total number of calls from the nest is recorded, along with the total brood size, which is used as an offset to allow the use of a Poisson response.

Since the same nests are measured repeatedly, the nest is used as a random effect. The model can be expressed as a zero-inflated generalized linear mixed model (ZIGLMM).

First we draw some pictures (Figures 1, 2).

Load the glmmADMB package to get access to the Owls data set; load the ggplot2 graphics package.

> library(glmmADMB)
> library(ggplot2)

Various small manipulations of the data set: (1) reorder nests by mean negotiations per chick, for plotting purposes; (2) add log brood size variable (for offset); (3) rename response variable.

```
> Owls <- transform(Owls,
    Nest=reorder(Nest,NegPerChick),
    logBroodSize=log(BroodSize),
    NCalls=SiblingNegotiation)
```

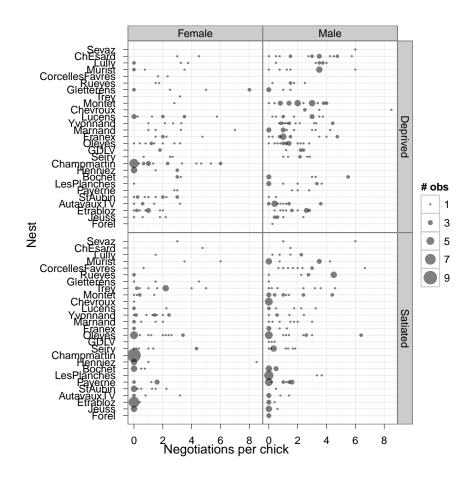


Figure 1: Basic view of owl data (arrival time not shown).

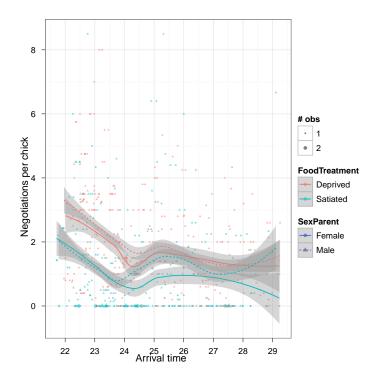
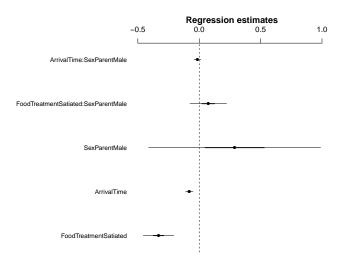


Figure 2: Basic view of owl data, #2 (nest identity not shown)

```
Now fit some models:
  The basic glmmadmb fit — a zero-inflated Poisson model.
> fit_zipoiss <- glmmadmb(NCalls~(FoodTreatment+ArrivalTime)*SexParent+
                                      offset(logBroodSize)+(1|Nest),
                                      data=Owls,
                                      zeroInflation=TRUE,
                                      family="poisson")
> summary(fit_zipoiss)
Call:
glmmadmb(formula = NCalls ~ (FoodTreatment + ArrivalTime) * SexParent +
    offset(logBroodSize) + (1 | Nest), data = Owls, family = "poisson",
    zeroInflation = TRUE)
Coefficients:
                                    Estimate Std. Error z value Pr(>|z|)
(Intercept)
                                      2.8562 0.3871 7.38 1.6e-13 ***
FoodTreatmentSatiated
                                                 0.0635 -5.22 1.8e-07 ***
                                     -0.3314
ArrivalTime
                                     -0.0807
                                                 0.0156
                                                          -5.18 2.3e-07 ***
SexParentMale
                                                0.3575
                                                          0.81
                                                                    0.42
                                      0.2882
FoodTreatmentSatiated:SexParentMale
                                      0.0740
                                                 0.0761
                                                          0.97
                                                                    0.33
                                                                    0.29
ArrivalTime:SexParentMale
                                     -0.0150
                                                 0.0143
                                                         -1.05
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
Number of observations: total=599, Nest=27
Random effect variance(s):
$Nest
            (Intercept)
(Intercept)
               0.14001
Zero-inflation: 0.25833 (std. err.: 0.018107)
Log-likelihood: -1985.3
  The coefplot2 package knows about glmmadmb fits:
```

> library(coefplot2)
> coefplot2(fit_zipoiss)



We can also try a standard zero-inflated negative binomial model; the default is the "NB2" parameterization (variance = $\mu(1 + \mu/k)$).

Alternatively, use an "NB1" fit (variance = $\phi \mu$).

Relax the assumption that total number of calls is strictly proportional to brood size (i.e. using log(brood size) as an offset):

Every change we have made so far improves the fit — changing distributions improves it enormously, while changing the role of brood size makes only a modest (-1 AIC unit) difference:

```
> library(bbmle)
> AICtab(fit_zipoiss,fit_zinbinom,fit_zinbinom1,fit_zinbinom1_bs)
```

```
      dAIC
      df

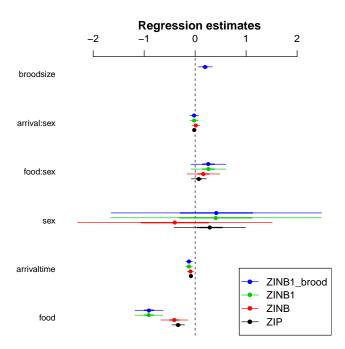
      fit_zinbinom1_bs
      0.0
      10

      fit_zinbinom1
      1.2
      9

      fit_zinbinom
      68.7
      9

      fit_zipoiss
      637.0
      8
```

Compare the parameter estimates:



2.1 Hurdle models

In contrast to zero-inflated models, hurdle models treat zero-count and non-zero outcomes as two completely separate categories, rather than treating the zero-count outcomes as a mixture of structural and sampling zeros.

As of version 0.6.7.1, glmmADMB includes truncated Poisson and negative binomial familes and hence can fit hurdle models. The two parts of the model

have to be fitted separately, however. First we fit a truncated distribution to the non-zero outcomes:

Then we fit a model to the binary part of the data (zero vs. non-zero). In this case, I started by fitting a simple (intercept-only) model with intercept-level random effects only. This comes a bit closer to matching the previous (zero-inflation) models, which treated zero-inflation as a single constant level across the entire data set (in fact, leaving out the random effects and just using glmmADMB(nz~1,data=0wls,family="binomial"), or glm(nz~1,data=0wls,family="binomial"), would be an even closer match). I then fitted a more complex binary model—this is all a matter of judgment about how complex a model it's worth trying to fit to a given data set—but it does look as though the zero-inflation varies with arrival time and satiation.

2.2 MCMC fitting

AD Model Builder has the capability to run a *post hoc* Markov chain to assess variability — that is, it uses the MLE as a starting point and the estimated sampling distribution (variance-covariance matrix) of the parameters as a candidate distribution, and "jumps around" the parameter space in a consistent way (Metropolis-Hastings?) to generate a series of samples from a posterior distribution of the parameter distribution (assuming flat priors: please see the ADMB documentation, or [1], for more details).

This is very convenient, but tends to be a bit slow. In the example below, I ran a chain of 50,000 MCMC iterations — on examination, the default chain of 1000 iterations was much too short — which took about two hours on a modern (2011) laptop.

family="nbinom1",
mcmc=TRUE,
mcmc.opts=mcmcControl(mcmc=50000))

Convert the MCMC chain to an mcmc object which the coda package can handle:

- > library(coda)
- > m <- as.mcmc(fit_zinbinom1_bs_mcmc\$mcmc)</pre>

Look at the trace plots. (Unfortunately, the MCMC variables are not properly labeled yet!)

- > library(scapeMCMC)
- > plotTrace(m)

V1	V2	V3	V4	V5	V6	V7
				opportuni		
V8	V9	V10	V11	V12	V13	V14
V15	V16	V17	V18	V19	V20	V21
V22	V23	V24	V25	V26	V27	V28
V29	V30	V31	V32	V33	V34	V35
	head head apprince of the			earn weeter fall-out his dragatill.	The effect of the property of the state of t	
V36	V37		•			•
Litandia) aldaminia sh Isang resista presiden						

- > gg <- geweke.diag(m)</pre>
- > range(effectiveSize(m))
- [1] 316.7461 702.1153
- > HPDinterval(m)

```
-8.21234181 -4.30098848
۷5
    -1.04757046 3.24003606
۷6
     0.88345874 8.90789301
۷7
    -0.31651160 3.56829008
8V
    -2.22957611 1.55683873
۷9
    -1.36784463 -0.45857212
    1.59636460 1.95655313
V10
V11
    -1.49081983 -1.02366435
V12 -1.56828489 -0.72608096
V13 -1.61582871 -0.50460814
V14 -1.90121523 -0.76022085
V15
    -1.06732858 -0.20816054
V16
    -0.20797827 0.66810830
V17 -1.34831422 -0.43570542
V18 -0.70588824 0.07192203
V19
    -1.43300082 -0.41339674
V20 -1.35590230 -0.22022910
V21 -0.51441482 0.78808287
V22 -0.51891395 0.25602298
V23
    -0.23575700 0.73582321
V24 -0.54402216 0.26956438
V25
     0.03551954 0.79823805
V26 -0.80082990 1.12111573
V27
    -0.09841576 0.70041477
V28 -0.62273846 0.21209616
V29
     0.16680329 1.21003569
     0.96304138 2.07080017
V30
V31 -0.47786300 0.52939624
V32
     0.14796009 0.92470871
V33
     0.94182995 2.02695564
V34
     0.46302084
                 1.29059423
V35
     0.59005525
                 1.41125115
V36
     0.78509092 1.80918334
V37
     0.32249976 1.23303743
attr(,"Probability")
[1] 0.9503386
> plotDens(m)
> ## very slow and not particularly useful for this large a set of parameters \dots
> plotSplom(m,pch=".")
```

lower

0.04086837 0.13503793

39.65328510 48.75123447

-10.48568580 -6.10491571

V1

٧2

V3

upper

3 Other information

The standard set of accessors is available:

coef extract (fixed-effect) coefficients

fixef a synonym for coef, for consistency with nlme/lme4

ranef extract random effect coefficients ("BLUPs" or "conditional modes")

residuals extract (Pearson) residuals

fitted fitted values

predict predicted values (based only on fixed effects, not on random effects), possibly with standard errors (based only on uncertainty of fixed effects), possibly for new data

logLik extract log-likelihood

AIC extract AIC

summary print summary

stdEr extract standard errors of coefficients

vcov extract estimated variance-covariance matrix of coefficients

VarCorr extract variance-covariance matrices of random effects

confint extract confidence intervals of fixed-effect coefficients

Missing: specifying starting values; MCMC; general troubleshooting (extra arguments, running outside R)

References

- [1] Benjamin M. Bolker. *Ecological Models and Data in R.* Princeton University Press, Princeton, NJ, 2008.
- [2] A. Roulin and L. Bersier. Nestling barn owls beg more intensely in the presence of their mother than in the presence of their father. *Animal Behaviour*, 74:1099–1106, 2007.
- [3] Alain F. Zuur, Elena N. Ieno, Neil J. Walker, Anatoly A. Saveliev, and Graham M. Smith. *Mixed Effects Models and Extensions in Ecology with R.* Springer, 1 edition, March 2009.