Parameterization of copula functions for bivariate survival data in the `surrosurv` package. Modelling and simulation

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Let define the joint survival function of S and T via a copula function:

$$S(s,t) = P(S > s, T > t) = C(u,v)|_{u=S_S(s), v=S_T(t)}\,, \tag{1}$$

where $S_S(\cdot) = P(S > s)$ and $S_T(\cdot) = P(T > t)$ are the marginal survival functions of S and T.

Modelling

In the case of possibly right-censored data, the individual contribution to the likelihood is

- $S(s,t) = C(u,v)|_{S_S(s),S_T(t)}$ if S is censored at time s and T is censored at time t,
- $-\frac{\partial}{\partial t} S(s,t) = \frac{\partial}{\partial v} C(u,v)\big|_{S_S(s),S_T(t)} f_T(t)$ if S is censored at time s and $T = t$,
- $-\frac{\partial}{\partial s} S(s,t) = \frac{\partial}{\partial v}\ C(u,v)|_{S_S(s),S_T(t)} f_S(s)$ if $S = s$ and T is censored at time t,
- $\frac{\partial^2}{\partial s \partial t} S(s,t) = \frac{\partial^2}{\partial u \partial v} C(u,v)\Big|_{S_S(s),S_T(t)} f_S(s) f_t(t)$ if $S = s$ and $T = t$.

Clayton copula

The bivariate Clayton [1978] copula is defined as

$$C(u,v) = \left(u^{-\theta} + v^{-\theta} - 1\right)^{-1/\theta}, \qquad \theta > 0. \tag{2}$$

The first derivative with respect to u is

$$\begin{aligned} \frac{\partial}{\partial u} C(u,v) &= \left(u^{-\theta} + v^{-\theta} - 1\right)^{-\frac{1+\theta}{\theta}} u^{-(1+\theta)} \\ &= \left[\frac{C(u,v)}{u}\right]^{1+\theta}. \end{aligned} \tag{3}$$

The second derivative with respect to u and v is

$$\frac{\partial^2}{\partial u \partial v} C(u,v) = (1+\theta) \frac{C(u,v)^{1+2\theta}}{(uv)^{1+\theta}}. \tag{4}$$

The Kendall [1938]'s tau for the Clayton copula is

$$\tau = \frac{\theta}{\theta + 2}. \tag{5}$$

Plackett copula

The bivariate Plackett [1965] copula is defined as

$$C(u,v) = \frac{\left[Q - R^{1/2}\right]}{2(\theta - 1)}, \qquad \theta > 0, \tag{6}$$

with

$$\begin{aligned} Q &= 1 + (\theta - 1)(u + v), \\ R &= Q^2 - 4\theta(\theta - 1)uv. \end{aligned} \tag{7}$$

Given that

$$\frac{\partial}{\partial u} Q = \theta - 1, \tag{8}$$

$$\frac{\partial}{\partial u} R = 2(\theta - 1)\left(1 - (\theta + 1)v + (\theta - 1)u\right), \tag{9}$$

the first derivative of $C(u,v)$ with respect to u is

$$\begin{aligned} \frac{\partial}{\partial u} C(u,v) &= \frac{1}{2}\left[1 - \frac{1 - (\theta + 1)v + (\theta - 1)u}{R^{1/2}}\right] \\ &= \frac{1}{2}\left[1 - \frac{Q - 2\theta v}{R^{1/2}}\right]. \end{aligned} \tag{10}$$

By defining

$$f = 1 - (\theta + 1)v + (\theta - 1)u, \tag{11}$$

$$g = R^{1/2} \tag{12}$$

and given that

$$f' = \frac{\partial}{\partial u} f = -(\theta + 1), \tag{13}$$

$$g' = \frac{\partial}{\partial u} g = \frac{\theta - 1}{R^{1/2}}\Big(1 - (\theta + 1)u + (\theta - 1)v\Big), \tag{14}$$

then, the second derivative with respect to u and v is

$$\begin{aligned} \frac{\partial^2}{\partial u \partial v} C(u,v) &= -\frac{f'g - fg'}{2g^2} \\ &= -\frac{\theta}{R^{3/2}}\Big[1 + (\theta - 1)(u + v - 2uv)\Big] \\ &= -\frac{\theta}{R^{3/2}}\Big[Q - 2(\theta - 1)uv\Big]. \end{aligned} \tag{15}$$

The Kendall's tau for the Plackett copula cannot be computed analyticaly and is obtained numerically.

Gumbel–Hougaard copula

The bivariate Gumbel [1960]–Hougaard [1986] copula is defined as

$$C(u,v) = \exp\Big(-Q^{\theta}\Big), \qquad \theta \in (0,1), \tag{16}$$

with

$$Q = (-\ln u)^{1/\theta} + (-\ln v)^{1/\theta}. \tag{17}$$

Given that

$$\frac{\partial}{\partial u} Q = -\frac{(-\ln u)^{1/\theta - 1}}{\theta u}, \tag{18}$$

then, the first derivative with respect to u is

$$\frac{\partial}{\partial u} C(u,v) = \frac{(-\ln u)^{1/\theta-1}}{u} C(u,v) Q^{\theta-1} \tag{19}$$

and the second derivative with respect to u and v is

$$\frac{\partial^2}{\partial u \partial v} C(u,v) = \frac{\Big[(-\ln u)(-\ln v)\Big]^{1/\theta-1}}{uv} C(u,v) Q^{\theta-2} \left[\frac{1}{\theta} - 1 + Q^{\theta}\right]. \tag{20}$$

The Kendall's tau for the Gumbel–Hougaard copula is

$$\tau = 1 - \theta. \tag{21}$$

Simulation

The function `simData.cc()` generates data from a Clayton copula model. First, the time value for the surrogate endpoint S is generated from its (exponential) marginal survival function:

$$S = -\log(U_S/\lambda_S), \qquad \text{with } U_S := S_S(S) \sim U(0,1). \tag{22}$$

Then, the time value for the true endpoint T is generated conditionally on the value s of S. The conditional survival function of $T \mid S$ is

$$S_{T|S}(t \mid s) = \frac{-\frac{\partial}{\partial s} S(s,t)}{-\frac{\partial}{\partial s} S(s,0)} = \frac{\frac{\partial}{\partial u} C(u,v)}{\frac{\partial}{\partial u} C(u,1)} \tag{23}$$

As the Clayton copula is used, we get (see Equtaion 3)

$$\begin{aligned} S_{T|S}(t \mid s) &= \left[\frac{C(S_S(s), S_T(t))}{C(S_S(s),1)}\right]^{1+\theta} = \left[\frac{U_S^{-\theta} + S_T(t)^{-\theta} - 1}{U_S^{-\theta}}\right]^{\frac{1+\theta}{\theta}} \\ &= \Big[1 + U_S^{\theta}(S_T(t)^{-\theta} - 1)\Big]^{\frac{1+\theta}{\theta}} \end{aligned} \tag{24}$$

By generating uniform random values for $U_T := S_{T|S}(T \mid s) \sim U(0,1)$, the values for $T \mid S$ are obtained as follows:

$$\begin{aligned} U_T &= \Big[1 + U_S^{\theta}(S_T(T)^{-\theta} - 1)\Big]^{\frac{1+\theta}{\theta}} \\ S_T(T) &= \left[\left(U_T^{-\frac{\theta}{1+\theta}} - 1\right) U_S^{-\theta} + 1\right]^{-1/\theta} \\ T &= -\log(U_T'/\lambda_T), \qquad \text{with } U_T' = \left[\left(U_T^{-\frac{\theta}{1+\theta}} - 1\right) U_S^{-\theta} + 1\right]^{-1/\theta}. \end{aligned} \tag{25}$$

References

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