LECTURE 5.3: PERMUTATION TESTS

STAT 1361/2360: STATISTICAL LEARNING AND DATA SCIENCE

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Overview

- Last time we talked about the bootstrap and used confidence intervals to motivate the idea
- Today we'll introduce **permutation tests**
 - ► Form of hypothesis testing that requires far fewer assumptions than the parametric versions you're used to
 - Accordingly, we now focus on hypothesis testing as the motivation for permutation tests
- First, a quick review of the general inference framework we discussed last time:



Inference Overview

The fundamental steps in statistical inference:

- 1. We have some (population distribution) F with some associated parameter $\boldsymbol{\theta}$
- 2. We take a sample from that population and calculate an estimate $\hat{\theta}$
- 3. We want to use $\hat{\theta}$ to *infer* something about θ

But ... In order to make any kind of inference, we need to know something about the distribution of $\hat{\theta}$



Inference Overview

That final point – needing to know something about the distribution of $\hat{\theta}$ – should be readily apparent in the context of hypothesis testing



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- 2. Define a test statistic t
- 3. Determine what the distribution of t would be if the null hypothesis were true (i.e. under H_0)
- Calculate the test statistic and determine whether it's "reasonable" to think it could have come from that distribution.
 - ▶ If not, we reject H_0 . If so, we fail to reject H_0 .



- Remember, we think of H_0 as being true "by default" we only reject H_0 if we find sufficient enough evidence to do so.
- The logic behind arguing that we have sufficient evidence to reject *H*⁰ is:
 - 1. If H_0 were true, then t would have to come from this distribution



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4. Therefore, H_0 must not be (or, is likely not) true

Hypothesis Testing Overview

 Thus, in order to do any kind of hypothesis test, we need to know what the distribution of our test statistic t would look like if the null hypothesis were true

** Sidenote: The terms "statistic" and "estimator" are referring to the same kind of fundamental object – both are just functions of the data. Thus, whether we talk about needing the distribution of an estimator $\hat{\theta}$ or a test statistic t, we're talking about the same kind of thing.



Let's go through an example:

Suppose we have two groups of observations and we want to test whether there is a significant difference in the group means:

$$H_0: \mu_1 = \mu_2 \iff \mu_1 - \mu_2 = 0$$

 $H_1: \mu_1 \neq \mu_2 \iff \mu_1 - \mu_2 \neq 0$

Data (unpaired) of the form:

Group 1: $X_1, ..., X_{n_1}$ Group 2: $Y_1, ..., Y_{n_2}$



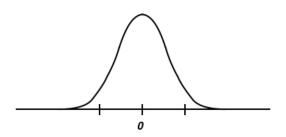
The natural test statistic to use here would be $t = \hat{\mu}_1 - \hat{\mu}_2$

How would we know the distribution of *t*?

- 1. If we assume the observations from each group come from a normal distribution:
 - ▶ Then we know $\hat{\mu}_1$ is approximately normal and $\hat{\mu}_2$ is approximately normal, so $\hat{\mu}_1 \hat{\mu}_2$ is approximately normal
- 2. If n_1 and n_2 are both reasonably large
 - ► Then we can appeal to the central limit theorem to know that both group means and hence the difference in group means are approximately normal



Now, if the null hypothesis were true and either of those two situations apply, then we know $t = \hat{\mu}_1 - \hat{\mu}_2$ would come from a normal distribution with mean o, so we can calculate our value of t and compare





BUT ... what if neither of those situations apply?

This is actually a very common situation in many applied fields:

- We have two groups of observations (and often times there looks to be a clear difference in group means)
- But we only have a few observations in each group and no reason to think that observations in each group came from a normal distribution

So what do we do in this situation?



Let's say we were willing to accept that everything about the distributions from the two groups was the same except (possibly) the means

Then under the null hypothesis, the means (and thus the entire distributions) would be identical

Group 1	Group 2
x_1	y_1
x_2	y_2
•	•
•	•
x_{n_1}	y_{n_2}



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How can we exploit this fact?



Data Format

Let's think about the data in a different format:

Observation	Group Label
x_1	1
x_2	1
•	•
	•
x_{n_1}	1
y_1	2
y_2	2
•	•
•	•
y_{n_2}	2



Group Label Randomness

Under the null hypothesis, the "Group Label" is random

Key Point #2: If this were true and we randomly shuffled the groups (i.e. permuted the Group Label) and recalculated $t_1^* = \hat{\mu}_1^* - \hat{\mu}_2^*$ according to these new groups, I should not expect t_1^* to be all that different from the test statistic I calculated on the original (correct) Group Labels, t



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But how different is too different to think H_0 isn't true? What if we did this twice to compute t_1^* and t_2^* ?



Permutation Test Idea

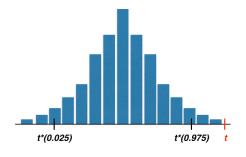
Idea: Repeat this a bunch of times, each time permuting the Group Labels and recomputing t_i^* . Then look at the distribution of these "permuted" test statistics and see where your original statistic t falls in that distribution.



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Idea: Repeat this a bunch of times, each time permuting the Group Labels and recomputing t_i^* . Then look at the distribution of these "permuted" test statistics and see where your original statistic t falls in that distribution.

 \implies If t lies at the extremes of this distribution, then we say it's likely it came from a different distribution, which means we reject H_0





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- 3. Repeat step (2) a large number of times (we'll say B times) to get $t_1^*, ..., t_B^*$
- 4. Look at distribution of the permuted test statistics (t_i^*) if the original statistic t lies within the $\alpha/2$ and $1 \alpha/2$ quantiles, we fail to reject H_0 . Otherwise we reject.



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Note: The q quantile is simply the value T such that $100 \times q\%$ of the observed values of the t_i^* are less than or equal to T.



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Suppose I have two variables X_1 and X_2 and I want to test if they're correlated:

$$H_0 = \operatorname{corr}(X_1, X_2) = r = 0$$

$$H_1 = corr(X_1, X_2) = r > 0$$

How could I do this – what would I be permuting?



Permutation Test for Correlation

Permutation Test for testing correlation:

- 1. Compute original test statistic $r = \widehat{\text{corr}}(X_1, X_2)$
- 2. Fix the order of X_1 , shuffle the order of X_2 , and (re)compute $r_1^* = \widehat{\text{corr}}(X_1, X_2^*)$
- 3. Repeat step (2) a large number of times (say B times) to get $r_1^*, ..., r_B^*$
- 4. Look at distribution of the permuted test statistics $(r_i^*'s)$ if the original statistic r lies below the 1α quantile, we fail to reject H_0 . Otherwise, we reject.



Permutation Test for Variable Importance

More generally, suppose I build a model with p predictors and I want to know if some predictor – say X_1 – is important. How could we (informally) use this permutation idea to test this?



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- 1. Build model with original data and compute the error Err₀
- 2. Fix X_2 , ..., X_p , permute the values of X_1 , rebuild the model and (re)compute the error Err_1^*



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- 1. Build model with original data and compute the error Err₀
- 2. Fix $X_2, ..., X_p$, permute the values of X_1 , rebuild the model and (re)compute the error Err_1^*
- 3. Repeat step (2) a large number of times (say B times) to get $\operatorname{Err}_{1}^{*},...,\operatorname{Err}_{B}^{*}$
- 4. Look at distribution of model errors when X_1 is permuted. If Err_0 lies above the α quantile of the permuted errors, then the original model isn't doing significantly better with the true values of X_1 , so we could conclude X_1 isn't that important.



- Determining which quantiles of the permutation distribution to compare the original test statistic with is going to depend on what kind of test you're doing – see last three examples (group means - middle quantiles; correlation - upper quantile; variable importance - lower quantile)
- We can actually get p-values for these tests. Recall a p-value is just the probability of seeing a test statistic *more* extreme than what you observed if H_0 is true
 - ► We can use the permutation distribution to estimate this ⇒ percentage of permuted test statistics more extreme than that calculated on original data



P-value Example

Example: Suppose we're doing a permutation test for correlation. Our original test statistic is r = 0.31. We do 1000 permutations, each time calculating the (permuted) test statistics to get $r_1^*, ..., r_{1000}^*$.

Say that of those 1000 permuted test statistics, 6 are greater than 0.31. Then our p-value would be

$$p = \frac{\text{# More Extreme Test Statistics}}{\text{Total # Permuted Test Statistics}} = \frac{6}{1000} = 0.006$$



- The theory behind permutation tests is actually quite involved, so we aren't getting into that here
- The formal development of permutation tests usually requires that the hypotheses be something like

$$H_0 = \mathcal{F}_1 = \mathcal{F}_2$$
$$H_1 = \mathcal{F}_1 \neq \mathcal{F}_2$$

in the case where we're testing equality between two groups, so \mathcal{F}_1 would be the distribution of the data from Group 1 and \mathcal{F}_2 the distribution of the data from Group 2

► The *power* of the test would then depend on the choice of test statistic



- Recall that the *power* of a hypothesis test is defined as the probability of (correctly) rejecting the null hypothesis H_0 when it is false.
- The idea behind permutation tests is actually quite simple

 you might ask why we even bother with the
 parametric-style hypothesis tests you learned in intro stats
 - ► Permutation tests suffer from lower power than their parametric counterparts
 - ► Keep in mind though that many times there is no nice parametric counterpart – back to the idea of needing to know the distribution of a statistic/estimator



Permutation tests are very often used to test group differences in small sample size settings. What can go wrong here?

- 1. Even if H_0 is true, might be an outlier in one of the groups
 - ► Could cause you to falsely reject *H*₀ parametric counterparts (usually *t*-tests) would also suffer
- 2. Even if H_0 is true, all samples in one group might be smaller than all samples in another
 - ► Likely would cause you to falsely reject *H*₀; *t*-tests would also suffer, but may be able to overcome it if differences were small enough



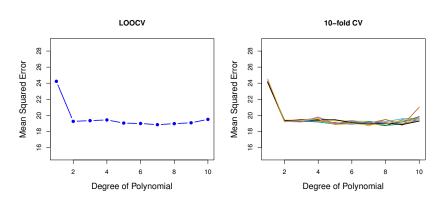


Cross-validation

- k-fold Cross-validation gives us a "better" (more robust)
 way of estimating generalizability error of a particular
 model (problems with training error and test error based
 on one split)
- Very useful in terms of deciding which models are outperforming others
 - ► Can be used to compare across different kinds of models (e.g. Logistic Regression vs. LDA v.s. QDA) or choosing the "best" from a class of models (e.g. choosing the value of a tuning parameter like the *k* in *k*NN or the number of terms in linear models)



Cross-validation



ISLR Fig. 5.4: Cross-validation error (MSE) from a linear model when different degrees of polynomial for one covariate (feature) are included in the model. Left: LOOCV. Right: Several different runs of 10-fold CV.



Bootstrapping and Permutation Tests

- In order to do any kind of inference (CIs or hypothesis tests), we need to know something about the distribution of a statistic/estimator
- In many cases, there are no mathematical results that tell us what the distribution is, or the data do not satisfy the assumptions necessary in order to assume a particular distribution
- In such instances, the bootstrap and permutation tests can be used in order to perform familiar kinds of statistical inference (CIs = bootstrap; Hypothesis Tests = bootstrap or permutation tests)

