

# Blocking in multi-stage experiments

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## Abstract

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Is it better to align the Stage 2 blocks with the Stage 1 blocks as  
far as possible  
or to make them as orthogonal to each other as possible?

In either case, how should treatments be assigned?

## Example 1: the problem

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units		treatments		
source	df	source	df	EMS
Mean	1	Mean	1	$6\sigma_B^2 + 6\sigma_L^2 + \sigma^2 + q_0$
Blocks	5			$6\sigma_B^2 + 6\sigma_L^2 + \sigma^2$
Units[B]	30	$F$	1	$\sigma^2 + q(F)$
		$G$	2	$\sigma^2 + q(G)$
		$F\#G$	2	$\sigma^2 + q(FG)$
		residual	25	$\sigma^2$

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We have lost 5 residual degrees of freedom, and gained nothing.

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This extends to three or more stages.

Brien, Harch, Correll and Bailey (2011) call this  
“confounding big with big”.

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**Design 2a** Align batches with lots;  
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Units[B]	30	$G$	2	$\sigma^2 + q(G)$
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		residual	26	$\sigma^2$

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The variance for the contrast between levels of  $F$  involves  $\sigma_L^2$  as well as  $\sigma_B^2$ , so it is larger than it needs to be.

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There are 4 residual degrees of freedom for testing the main effect of  $F$ , and this cannot be increased.

## Example 2: design 2b

**Design 2b** Cross batches with lots to form a square array;  
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and then randomize rows and columns.

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Batches	5	$F$	1	$6\sigma_B^2 + \sigma^2 + q(F)$
		residual	4	$6\sigma_B^2 + \sigma^2$
Lots	5			$6\sigma_L^2 + \sigma^2$
B#L	25	$G$	2	$\sigma^2 + q(G)$
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Main effect of  $F$  has smaller variance than before, and same df.

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Main effect of  $F$  has smaller variance than before, and same df.  
Other treatment effects have same (small) variance,  
and df reduced from 26 to 21.

## Some principles

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*If a treatment factor has to be applied to large units such as blocks in one stage,  
then it will have relatively few residual degrees of freedom.  
In order not to reduce these further, try to confound  
the whole of this block term with the same term in other stages.*

## Example 2: design 2c

**Design 2c** Make 3 squares by crossing pairs of batches (shown as rows) with pairs of lots (shown as columns).

$F1$	$G1, G2, G3$	$G1, G2, G3$
$F2$	$G1, G2, G3$	$G1, G2, G3$

$F1$	$G1, G2, G3$	$G1, G2, G3$
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Randomize levels of  $F$  to rows within each square;  
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Randomize levels of  $F$  to rows within each square;  
randomize levels of  $G$  within each corner of each square.

**Remark** Squares is the **supremum** of Batches and Lots:  
 $\text{Squares} = \text{Batches} \vee \text{Lots}$ .

## Example 2: skeleton anova for design 2c

units		treatments		
source	df	source	df	EMS
Mean	1	Mean	1	$6\sigma_B^2 + 6\sigma_L^2 + \sigma_1^2 + q_0$
Squares	2			$6\sigma_B^2 + 6\sigma_L^2 + \sigma_1^2$
Batches[S]	3	F	1	$6\sigma_B^2 + \sigma_1^2 + q(F)$
		residual	2	$6\sigma_B^2 + \sigma_1^2$
Lots[S]	3			$6\sigma_L^2 + \sigma_1^2$
B#L[S]	3			$\sigma_1^2$
Units[B,L,S]	24	G	2	$\sigma^2 + q(G)$
		F#G	2	$\sigma^2 + q(FG)$
		residual	20	$\sigma^2$

The randomization argument suggests that  $\sigma_1^2 \neq \sigma^2$ .

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Both residual df have decreased, and nothing has been gained.

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Any further blocking of either batches or lots reduces the already-small residual df for main effects.

## More general numbers

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Stage 1 groups them into  $b$  batches of size  $s$ ;

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we need  $m$  arrays of  $b/m$  batches crossed with  $c/m$  lots,  
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we need  $m$  arrays of  $b/m$  batches crossed with  $c/m$  lots,  
with each intersection containing  $Nm/bc$  experimental units,  
where  $m$  divides  $b$ ,  $m$  divides  $c$  and  $bc$  divides  $Nm$ ,  
so  $m$  divides  $\gcd\{b, c\}$ , and  $\gcd\{b, c\}$  divides  $Nm/\text{lcm}\{b, c\}$ .

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For orthogonality, the only possibility is three  $4 \times 3$  arrays.

•	•	•
•	•	•
•	•	•
•	•	•

•	•	•
•	•	•
•	•	•
•	•	•

•	•	•
•	•	•
•	•	•
•	•	•

## *m* Arrays

The between-Arrays stratum variance is

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so these  $m - 1$  degrees of freedom are typically not used for inference or estimation.

It is desirable to keep  $m$  small.

## Powers of the same small prime

Suppose that the number of batches and the number of lots are both powers of  $p$ , where  $p = 2$  or  $p = 3$ , that several  $p$ -level treatment factors  $F_1, F_2, \dots$  must be applied to whole batches in Stage 1, and several  $p$ -level treatment factors  $G_1, G_2, \dots$  must be applied to whole lots in Stage 2.

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If each stage is a single replicate of the relevant treatments, we may be able to take  $m = p$  and sacrifice information only on high-order interactions.

## Powers of the same small prime

Suppose that the number of batches and the number of lots are both powers of  $p$ , where  $p = 2$  or  $p = 3$ , that several  $p$ -level treatment factors  $F_1, F_2, \dots$  must be applied to whole batches in Stage 1, and several  $p$ -level treatment factors  $G_1, G_2, \dots$  must be applied to whole lots in Stage 2.

If each stage is a single replicate of the relevant treatments, we may be able to take  $m = p$  and sacrifice information only on high-order interactions.

For example, if  $p = 2$  and  $F_1, F_2, F_3$  and  $F_4$  are applied to whole batches in Stage 1 while  $G_1, G_2$  and  $G_3$  are applied to whole lots in Stage 2, we can use two  $8 \times 4$  arrays, and confound  $F_1F_2F_3F_4$  and  $G_1G_2G_3$  with each other and with arrays.

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This technique is called **post-fractionation** by Bisgaard (1997) and Vivacqua and Bisgaard (2009).

## Three stages

Suppose that there are three stages, and that treatment factors  $F$ ,  $G$  and  $H$  are applied in Stages 1, 2, 3 respectively.

Suppose that each treatment factor must be applied to whole blocks in its stage.

We already know that we should try to make the blocks from each stage as orthogonal as possible to blocks from every other stage.

An example was investigated by Mee and Bates (1998).

## Example 5: the problem

There are 16 experimental units.

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In Stage 3, the units must be processed in 4 pods of size 4.

Treatment factor  $H$  has 2 levels, which are applied in Stage 3, and these must be applied to whole pods.

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In Stage 3, the units must be processed in 4 pods of size 4.

Treatment factor  $H$  has 2 levels, which are applied in Stage 3, and these must be applied to whole pods.

How should we design the experiment?

## Example 5: design 5a

**Design 5a** Form the experimental units into 2 arrays of size  $2 \times 2 \times 2$ . The first coordinate indicates the batch, the second coordinate indicates the lot, and the third coordinate indicates the pod. Within each array, randomize levels of  $F$  to batches, levels of  $G$  to lots, and levels of  $H$  to pods.

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The single df between Arrays is wasted.

All 7 treatment df have different variances,  
each with just one residual df.

## Example 5: design 5b

**Design 5b** Let  $\Lambda$  be a  $4 \times 4$  Latin square.

Stage 1: identify the batches with the rows of  $\Lambda$ .

Stage 2: identify the lots with the columns of  $\Lambda$ .

Stage 3: identify the pods with the letters  $\Lambda$ .

## Example 5: design 5b

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Can we apply levels of  $F, G, H$  to rows, columns, letters respectively in such a way that all treatment interactions are orthogonal to rows, columns and letters?

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Can we apply levels of  $F, G, H$  to rows, columns, letters respectively in such a way that all treatment interactions are orthogonal to rows, columns and letters?

Mee and Bates found a cunning way of doing this, using the non-cyclic Latin square of order 4.

## Example 5: skeleton anova for design 5b

units		treatments		
source	df	source	df	EMS
Mean	1	Mean	1	$4\sigma_R^2 + 4\sigma_C^2 + 4\sigma_L^2 + \sigma^2 + q_0$
Rows	3	$F$	1	$4\sigma_R^2 + \sigma^2 + q(F)$
			2	$4\sigma_R^2 + \sigma^2$
Columns	3	$G$	1	$4\sigma_C^2 + \sigma^2 + q(G)$
			2	$4\sigma_C^2 + \sigma^2$
Letters	3	$H$	1	$4\sigma_L^2 + \sigma^2 + q(H)$
			2	$4\sigma_L^2 + \sigma^2$
Units[R,C,L]	6	$F\#G$	1	$\sigma^2 + q(FG)$
		$F\#H$	1	$\sigma^2 + q(FH)$
		$G\#H$	1	$\sigma^2 + q(GH)$
		$F\#G\#H$	1	$\sigma^2 + q(FGH)$
		residual	2	$\sigma^2$

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source	df	source	df	EMS
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Rows	3	$F$	1	$4\sigma_R^2 + \sigma^2 + q(F)$
		residual	2	$4\sigma_R^2 + \sigma^2$
Columns	3	$G$	1	$4\sigma_C^2 + \sigma^2 + q(G)$
		residual	2	$4\sigma_C^2 + \sigma^2$
Letters	3	$H$	1	$4\sigma_L^2 + \sigma^2 + q(H)$
		residual	2	$4\sigma_L^2 + \sigma^2$
Units[R,C,L]	6	$F\#G$	1	$\sigma^2 + q(FG)$
		$F\#H$	1	$\sigma^2 + q(FH)$
		$G\#H$	1	$\sigma^2 + q(GH)$
		$F\#G\#H$	1	$\sigma^2 + q(FGH)$
		residual	2	$\sigma^2$

For this Latin square, the decomposition into strata can be justified by randomization (Bailey, 1982).

## Example 5: design 5b by design key

The design key introduced by Patterson (1965) gives a clean construction of design 5b. All factors and pseudofactors have two levels, and arithmetic is modulo 2.

Rows:  $R_1, R_2, R_1 + R_2$ .

Columns:  $C_1, C_2, C_1 + C_2$ .

Letters:  $L_1 = R_1 + C_1, L_2 = R_2 + C_2, L_1 + L_2 = R_1 + R_2 + C_1 + C_2$ .

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$F = R_1$ .

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$F = R_1$ .

$G = C_2$ .

$H = R_1 + R_2 + C_1 + C_2$ .

$F + G = R_1 + C_2$ .

$F + H = R_2 + C_1 + C_2$ .

$G + H = R_1 + R_2 + C_1$ .

$F + G + H = R_2 + C_1$ .

## Non-orthogonality

So far, we have assumed that treatments applied in Stage i  
**either** must be applied to whole blocks in Stage i  
**or** can be orthogonal to blocks in Stage i.

Suppose that neither of these conditions holds?

## Example 6: the problem

There are 36 experimental units.

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Stage 1 has 12 blocks of size 3, in 4 superblocks of 3 blocks.

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Treatment factor  $G$  has 3 levels, which are applied in Stage 2, in such a way that each combination of  $F$  and  $G$  occurs 4 times.

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Forced:  $F$  applied to whole blocks.

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Good idea: Apply each level of  $F$  to one block per superblock.

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Good idea: Align superblocks and rows.

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Good idea: Make the design in columns as efficient as possible.

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Good idea: Apply each level of  $G$  to one exp. unit per block.

Good idea: Align superblocks and rows.

Good idea: Form the 12 blocks and 9 columns into three arrays of size  $4 \times 3$ .

Good idea: Make the design in columns as efficient as possible.

Can we achieve all of this?

## Example 6: align superblocks with rows, and make arrays

Row 1 = Superblock 1

Row 2 = Superblock 2

Row 3 = Superblock 3

Row 4 = Superblock 4

●	●	●
●	●	●
●	●	●
●	●	●

Row 1 = Superblock 1

Row 2 = Superblock 2

Row 3 = Superblock 3

Row 4 = Superblock 4

●	●	●
●	●	●
●	●	●
●	●	●

Row 1 = Superblock 1

Row 2 = Superblock 2

Row 3 = Superblock 3

Row 4 = Superblock 4

●	●	●
●	●	●
●	●	●
●	●	●

## Example 6: design

	Array 1			Array 2			Array 3		
Row 1	$F1G1$	$F1G2$	$F1G3$	$F2G1$	$F2G2$	$F2G3$	$F3G1$	$F3G2$	$F3G3$

## Example 6: design

	Array 1			Array 2			Array 3		
Row 1	$F1G1$	$F1G2$	$F1G3$	$F2G1$	$F2G2$	$F2G3$	$F3G1$	$F3G2$	$F3G3$
Row 2	$F2G3$	$F2G1$	$F2G2$	$F3G3$	$F3G1$	$F3G2$	$F1G3$	$F1G1$	$F1G2$
Row 3	$F3G3$	$F3G1$	$F3G2$	$F1G3$	$F1G1$	$F1G2$	$F2G3$	$F2G1$	$F2G2$
Row 4	$F1G2$	$F1G3$	$F1G1$	$F2G2$	$F2G3$	$F2G1$	$F3G2$	$F3G3$	$F3G1$

## Example 6: design

	Array 1			Array 2			Array 3		
Row 1	$F1G1$	$F1G2$	$F1G3$	$F2G1$	$F2G2$	$F2G3$	$F3G1$	$F3G2$	$F3G3$
Row 2	$F2G3$	$F2G1$	$F2G2$	$F3G3$	$F3G1$	$F3G2$	$F1G3$	$F1G1$	$F1G2$
Row 3	$F3G3$	$F3G1$	$F3G2$	$F1G3$	$F1G1$	$F1G2$	$F2G3$	$F2G1$	$F2G2$
Row 4	$F1G2$	$F1G3$	$F1G1$	$F2G2$	$F2G3$	$F2G1$	$F3G2$	$F3G3$	$F3G1$

$F1$	$F1$	$F1$	$F2$	$F2$	$F2$	$F3$	$F3$	$F3$
$G3$	$G1$	$G2$	$G3$	$G1$	$G2$	$G3$	$G1$	$G2$

## Example 6: design

	Array 1			Array 2			Array 3		
Row 1	$F1G1$	$F1G2$	$F1G3$	$F2G1$	$F2G2$	$F2G3$	$F3G1$	$F3G2$	$F3G3$
Row 2	$F2G3$	$F2G1$	$F2G2$	$F3G3$	$F3G1$	$F3G2$	$F1G3$	$F1G1$	$F1G2$
Row 3	$F3G3$	$F3G1$	$F3G2$	$F1G3$	$F1G1$	$F1G2$	$F2G3$	$F2G1$	$F2G2$
Row 4	$F1G2$	$F1G3$	$F1G1$	$F2G2$	$F2G3$	$F2G1$	$F3G2$	$F3G3$	$F3G1$

$F1$	$F1$	$F1$	$F2$	$F2$	$F2$	$F3$	$F3$	$F3$
$G3$	$G1$	$G2$	$G3$	$G1$	$G2$	$G3$	$G1$	$G2$

## Example 6: design

	Array 1			Array 2			Array 3		
Row 1	$F1G1$	$F1G2$	$F1G3$	$F2G1$	$F2G2$	$F2G3$	$F3G1$	$F3G2$	$F3G3$
Row 2	$F2G3$	$F2G1$	$F2G2$	$F3G3$	$F3G1$	$F3G2$	$F1G3$	$F1G1$	$F1G2$
Row 3	$F3G3$	$F3G1$	$F3G2$	$F1G3$	$F1G1$	$F1G2$	$F2G3$	$F2G1$	$F2G2$
Row 4	$F1G2$	$F1G3$	$F1G1$	$F2G2$	$F2G3$	$F2G1$	$F3G2$	$F3G3$	$F3G1$

$F1$	$F1$	$F1$	$F2$	$F2$	$F2$	$F3$	$F3$	$F3$
$G3$	$G1$	$G2$	$G3$	$G1$	$G2$	$G3$	$G1$	$G2$

The design in columns is factorially balanced,  
with canonical efficiency factors  $15/16$  for both main effects  
and  $3/4$  for the interaction.

## Example 6: skeleton anova

units		treatments			
source	df	cef	source	df	EMS
Mean	1	1	Mean	1	$9\sigma_R^2 + 9\sigma_S^2 + 4\sigma_C^2 + 3\sigma_B^2 + \sigma^2 + q_0$
Rows	3				$9\sigma_R^2 + 9\sigma_S^2 + 3\sigma_B^2 + \sigma^2$
Arrays	2	$\frac{1}{16}$	F	2	$4\sigma_C^2 + 3\sigma_B^2 + \sigma^2 + \frac{1}{16}q(F)$
Blocks[R,A]	6	$\frac{15}{16}$	F	2	$3\sigma_B^2 + \sigma^2 + \frac{15}{16}q(F)$
			residual	4	$3\sigma_B^2 + \sigma^2$
Columns[A]	6	$\frac{1}{16}$	G	2	$4\sigma_C^2 + \sigma^2 + \frac{1}{16}q(G)$
		$\frac{1}{4}$	F#G	4	$4\sigma_C^2 + \sigma^2 + \frac{1}{4}q(FG)$
Units[B,C,R,A]	18	$\frac{15}{16}$	G	2	$\sigma^2 + \frac{15}{16}q(G)$
		$\frac{3}{4}$	F#G	4	$\sigma^2 + \frac{3}{4}q(FG)$
			residual	12	$\sigma^2$

## A strategy for non-orthogonality

Suppose that, in Stage 1, treatment factor  $F$  has canonical efficiency factors

$p$  in Batches  
 $q$  in Units[Batches],

where  $p + q = 1$ .

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Suppose that, in Stage 1, treatment factor  $F$  has canonical efficiency factors

$$\begin{aligned} p &\text{ in Batches} \\ q &\text{ in Units[Batches],} \end{aligned}$$

where  $p + q = 1$ .

Possible strategies for Stage 2:

- ( $F \perp L$ ) make  $F$  orthogonal to Lots;
- (BigwithBig) confound (the  $F$ -part of) Batches with Lots and the  $F$ -part of  $U[B]$  with  $U[L]$ ;
- (Compensate) confound the  $F$ -part of Batches with  $U[L]$  and the  $F$ -part of  $U[B]$  with Lots.

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- (Compensate) confound the  $F$ -part of Batches with  $U[L]$  and the  $F$ -part of  $U[B]$  with Lots.

If  $p = 1$  then  $F$  is confounded with Batches, so ( $F \perp L$ ) is the same as (Compensate) and this is best.

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If  $q = 1$  then  $F$  is orthogonal to Batches, so ( $F \perp L$ ) is the same as (BigwithBig) and this is best.

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If  $q = 1$  then  $F$  is orthogonal to Batches, so ( $F \perp L$ ) is the same as (BigwithBig) and this is best.

For all values of  $p$ , ( $F \perp L$ ) is best if it is possible.

## A strategy for non-orthogonality, continued

### Theorem

If Stage 1 has  $b$  batches of size  $s$  and Stage 2 has  $c$  lots of size  $k$ , then (BigwithBig) is better than (Compensate) in the sense of giving smaller variances for the estimators of contrasts between levels of  $F$  if and only if

$$\frac{q}{p} > \frac{\sigma^2}{(s\sigma_B^2 + \sigma^2)} \frac{(k\sigma_L^2 + \sigma^2)}{(s\sigma_B^2 + k\sigma_L^2 + \sigma^2)}.$$

## A strategy for non-orthogonality, continued

### Theorem

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If  $q > p$  then (BigwithBig) is better; otherwise, prior estimates of the relatives magnitudes of  $\sigma^2$ ,  $\sigma_B^2$  and  $\sigma_L^2$  are required to make the decision.

## Another strategy for non-orthogonality

Stage 1 has  $b$  batches of size  $s$  and Stage 2 has  $c$  lots of size  $k$ , so  $N = bs = ck$ .

If  $s$  divides  $k$  then we could nest batches within lots.

If  $sk$  divides  $N$  then we could cross batches with lots in  $N/sk$  arrays.

If  $k = ts$  and  $N = msk$  then we could do either.

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Which is better?

## Another strategy for non-orthogonality, continued

Stage 1 has  $mts$  batches of size  $s$  and

Stage 2 has  $ms$  lots of size  $ts$ .

Let  $\Delta$  be a design for  $F$  in  $mts$  batches of size  $s$ .

Let  $\Gamma$  be a design for  $F$  in  $ms$  lots of size  $ts$ .

Let  $\text{Nest}(\Delta, \Gamma)$  be a design where each lot contains  $t$  batches,  
the design in batches is  $\Delta$  and the design in lots is  $\Gamma$ .

Let  $\text{Cross}(\Delta, \Gamma)$  be a design with  $m$  arrays of  $ts$  batches crossed  
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### Theorem

*The variances for the estimators of contrasts between levels of  $F$  are no bigger for  $\text{Nest}(\Delta, \Gamma)$  than for  $\text{Cross}(\Delta, \Gamma)$ .*

This is not quite the whole story,

because it may be possible to construct a  $\text{Cross}(\Delta, \Gamma)$  for better  
block designs  $\Delta$  and  $\Gamma$  than a  $\text{Nest}(\Delta, \Gamma)$ .

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Many multi-phase experiments are similar.

This all leads to interesting questions.  
Thank you for listening.