514.05. Statistical Analysis of Astronomical Measurement Errors

Richard A. Bilonick, PhD

R. A. Bilonick Statistics Consultancy LLC

rabilonick@gmail.com

237th Meeting of the American Astronomical Society

January 15, 2021

4D> 4B> 4B> B 990

514.05 5051-05-01

514.05. Statistical Analysis of Astronomical Measurement Errors

Richard A. Bilonick, PhD

R. A. Bilonick Statistics Consultancy LLC

rabilonick@genail.com 237th Meeting of the American Astronomical Society

January 15, 2021

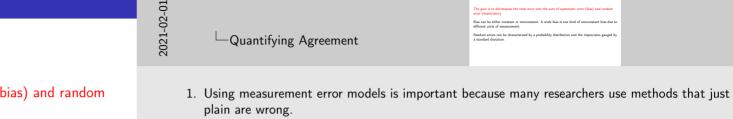
R. A. Bilonick 514.05 January 15, 2021 1 / 17

Quantifying Agreement

The goal is to decompose the total error into the sum of systematic error (bias) and random error (imprecision).

Bias can be either constant or nonconstant. A scale bias is one kind of nonconstant bias due to different units of measurement.

Random errors can be characterized by a probability distribution and the imprecision gauged by a standard deviation.



514.05

2. If you get the calibration curve wrong, you will have a false picture of how methods compare.

Quantifying Agreement

The goal is to decompose the total error into the sum of systematic error (bias) and randon Riss can be either constant or nonconstant. Δ scale bias is one kind of nonconstant bias due to

- 3. Ordinary regression rarely will produce a curve is even approximately close to the correct calibration curve.
- 4. And there are TWO different regression lines to boot! They both can't be correct!

R. A. Bilonick 514.05 January 15, 2021

DEEP2 Spectroscopic & Photometric Data

Galaxies 1432

z_{spec} Spectroscopy

Zfink S. Finkelstein

z_{font} A. Fontana

z_{pfor} J. Pfor

z_{salv} M. Salvato

Zwikl T. Wiklind

 z_{wuvt} S. Wuyts

R. A. Bilonick

Funding for the DEEP2 Galaxy Redshift Survey has been provided by NSF grants AST-95-09298, AST-0071048, AST-0507428, and AST-0507483 as well as NASA LTSA grant NNG04GC89G. Jeffrey Newman and Ditran Kodra kindly provided the redshift data.

514.05



January 15, 2021

514.05

Calaina 1422
Aqua. Spectroscopy
Ana S. Friedmann
Aqua. Spe

- 1. There were 1432 galaxies which had their redshift measured by all seven methods.
- 2. A one-factor measurement error model would not work for all seven methods because the six photometric methods all depend on the same photometric information.
- 3. Even a two-factor model will not work here because the spectroscopic measurements were not repeated making it impossible to separate the true redshift variance from the random error.

Ordinary Regression Model

In an ordinary regression model, one variable must be a response (dependent) variable and the other must be an explanatory (independent) variable which is measured WITHOUT random error.

The decision to use ordinary regression limits you to two measurement methods at a time.

The decision on dependent/independent variables is consequential if the independent variable is contaminated by random error.

Depending on which of the 2 methods is chosen as the independent variable, you will get a different regression line, i.e., there are ALWAYS 2 regression lines (so long at the correlation does not equal 1 or -1).

It is well known that regressing against an independent variable with random measurement error biases the slope estimate toward zero. The greater the random error, the greater the bias toward zero.



2021-02-01

Ordinary Regression Model

In an ordinary regression model, one variable must be a response (dependent) variable and the other must be an explanately (independent) variable which is measured WITHOUT random over.

In diction to use ordinary regression limits you to two measurement methods at a time.

The decision to use ordinary regression limits you to two measurement methods at a time.

The decision to use ordinary regression function of the composition variable in December or which is a Composition of which of the decision to use ordinary regression (see in the consistent of the confidence or which is a format an independent variable in the consistent of the confidence or which the confidence of the confidence of the confidence of the confidence of the confidence or which the confidence of t

- 1. At minimum, pairwise regression would be inefficient and would not use all available information
- 2. You should not use ordinary regression when both methods are contaminated by random error.
- 3. All methods are contaminated by random error to one extent or another. It's always best to account for it.

R. A. Bilonick 514.05 January 15, 2021 4 / 17

Measurement Error Model

In a measurement error model, all measurement methods are treated as responses (dependent) variables.

They are dependent on one or more latent (unobserved) variables.

The number of latent variables and their relationships with the measurements depends on how the data were collected.

The measurement error model can be implemented using a structural equation model (SEM). Relationships among the latent variables represents the structure.



514.05

To Wassurement Error Model

Measurement error model, all measurement methods are treated as response (dependent) varieties.

They are dependent on one or more latent (underwed) varieties.

The measurement method their relationships with the measurements depend on loss the data varieties and their relationships with the measurements depend on loss the data varieties and their relationships with the measurements depend on loss the data varieties and their relationships among the latest varieties are present that structure.

- 1. The latent variable for the true but unknown redshift represents the assumption that all methods are trying to measure the same quantity.
- 2. The structural equation model is essentially a set of simultaneous equations with distributional assumptions for the random error and the latent variable(s)
- 3. When determining the calibration equations, the latent variable vanishes.

R. A. Bilonick 514.05 January 15, 2021 5 / 17

Simple One-Factor Measurement Error Model – Path Diagram

The path diagram shows the relationship between the true but unknown redshift ζ and the observed measurements Z_1 , Z_2 , and Z_3 .

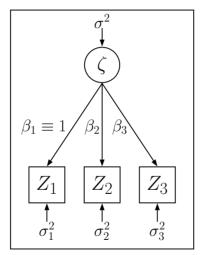
This type of one-factor model would be reasonable for either independent repeats using the same method or three different photometric redshift methods.

$$\zeta \sim N(\bar{\zeta}, \sigma^2)$$

$$Z_1 = \zeta + \epsilon_1, \quad \epsilon_1 \sim N(0, \sigma_1^2)$$

$$Z_2 = \alpha_2 + \beta_2 \zeta + \epsilon_2, \quad \epsilon_2 \sim N(0, \sigma_2^2)$$

$$Z_3 = \alpha_3 + \beta_3 \zeta + \epsilon_3, \quad \epsilon_3 \sim N(0, \sigma_2^2)$$





514.05

Simple One-Factor Measurement Error Model - Path Diagram

The path diagram about the indicationly between the interaction of the diagram of the indicational paths and the interaction of the diagram of the indication of

- 1. Path diagrams are a great way to illustrate the measurement error model or any SEM.
- 2. However, when the model is very large, it can be difficult to represent the entire model by a path diagram.
- 3. If you can draw the path diagram it will help guide you in writing the SEM code.

R. A. Bilonick 514.05 January 15, 2021 6 /

Calibration Equations

Using some simple algebra and taking expectations, the calibration curve between any two types of photometric measurements Z_i and $Z_{i'}$ can be easily derived as:

$$E[Z_i] = \left[\alpha_i - \frac{\beta_i}{\beta_{i'}}\alpha_{i'}\right] + \left(\frac{\beta_i}{\beta_{i'}}\right)E[Z_{i'}]$$

The estimation of the parameters clearly requires a nonlinear methodology.



- 1. While there are always two DIFFERENT regreassion equations for any pair of methods, there is only ONE calibration curve relating the methods.
- 2. You must have at least three measurements for each object in order to estimate all the model parameters.
- These three or more measurements per object could be three or more different methods or fewer methods if replaced by independent repeats of some of the other methods.

R. A. Bilonick 514.05 January 15, 2021 7 /

Method of Moments Estimators for 3 Methods

The method of moments equates the observed covariance matrix S to the theoretical covariance matrix Σ resulting in six equations (setting $\beta_1 = 1$) in six unknowns:

$$S_{1}^{2} = \sigma^{2} + \sigma_{1}^{2}$$

$$S_{2}^{2} = \beta_{2}^{2}\sigma^{2} + \sigma_{2}^{2}$$

$$S_{3}^{2} = \beta_{3}^{2}\sigma^{2} + \sigma_{3}^{2}$$

$$S_{12} = \beta_{2}\sigma^{2}$$

$$S_{13} = \beta_{3}\sigma^{2}$$

$$S_{23} = \beta_{2}\beta_{3}\sigma^{2}$$



2021-02-01

R. A. Bilonick 514.05 January 15, 2021 8 / 1

Method of Moments Estimators for 3 Methods

The method of moments against the dispersion state it. So to the Moments contained and moments against the dispersion contained and moments against the dispersion contained and moments are stated as a subject of the s

- 1. Here we have six egations in six unknowns for the scale parameters and variance components.
- 2. Because the number of eqations equals the number of parameters we get a unique solution.
- 3. In this case, we can not get a gauge of how well the model fits because there will be zero degree of freedom. This is not a big problem but it's always better to have more than three methods an or repeats - the more the better.

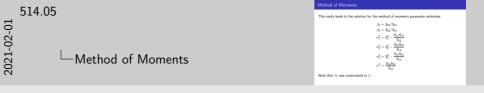
Method of Moments

This easily leads to the solution for the method of moments parameter estimates:

$$eta_2 = S_{23}/S_{13}$$
 $eta_3 = S_{23}/S_{12}$
 $\sigma_1^2 = S_1^2 - rac{S_{12}S_{13}}{S_{23}}$
 $\sigma_2^2 = S_2^2 - rac{S_{12}S_{23}}{S_{13}}$
 $\sigma_3^2 = S_3^2 - rac{S_{13}S_{23}}{S_{12}}$
 $\sigma^2 = rac{S_{12}S_{13}}{S_{23}}$

Note that β_1 was constrained to 1.





- 1. Here the solution is relatively easy.
- 2. In practice, we would prefer to use maximum likelihood to estimate the parameters.

R. A. Bilonick 514.05 January 15, 2021 9 / 17

Method of Moments – Solutions for lphas and $ar{\zeta}$

Setting the sample means (\bar{Z}_i) equal to their expected values given the model:

$$ar{Z}_1 = ar{\zeta}$$
 $ar{Z}_2 = lpha_2 + eta_2 ar{\zeta}$
 $ar{Z}_3 = lpha_3 + eta_3 ar{\zeta}$

which yielded the following solution:

$$\bar{\zeta} = \bar{Z}_1$$
 $\alpha_2 = \bar{Z}_2 - \frac{S_{23}}{S_{13}} \bar{Z}_1$
 $\alpha_3 = \bar{Z}_3 - \frac{S_{23}}{S_{12}} \bar{Z}_1$

Note that α_1 was constrained to 0.



 $\frac{1000}{5}$ Setting the sample mean (2) equal to their expected values given the model: $\frac{\lambda_1 - \zeta_2}{\lambda_2} = \alpha_1 + \beta_1 \zeta_2$ which yielded the following solutions: $\frac{\zeta}{\zeta} = \frac{\lambda_1}{\alpha_2} = \frac{\zeta_2}{\delta_{12}} \frac{\lambda_1}{\lambda_2}$ Wheth hold of Moments — Solutions for α s and $\frac{\zeta}{\zeta} = \frac{\lambda_1}{\alpha_2} = \frac{\zeta_2}{\delta_{12}} \frac{\lambda_1}{\lambda_2}$ Note that α_1 was constrained to 0.

- 1. We solve for the intercepts separately.
- 2. We need the scale parameter estimates to complete the intercept estimates.
- 3. We need both the intercepts and the scale parameter estimates to determine the calibration curve

R. A. Bilonick 514.05 January 15, 2021 10

Method of Moments Estimates – Finkelstein (Z_1) , Fontana (Z_2) , & Pfor (Z_3)

$$\begin{bmatrix} S_1^2 = 0.165 & S_{12} = 0.157 & S_{13} = 0.15 \\ S_2^2 = 0.185 & S_{23} = 0.154 \\ S_3^2 = 0.173 \end{bmatrix} \qquad \qquad \bar{Z}_1 = \bar{\zeta} = 0.7689 \\ \bar{Z}_2 = \alpha_2 + \beta_2 \bar{\zeta} = 0.7852$$

$$\beta_2 = S_{23}/S_{13} = 1.029$$

$$\beta_3 = S_{23}/S_{12} = 0.983$$

$$\sigma_1^2 = S_1^2 - \frac{S_{12}S_{13}}{S_{23}} = 0.01196, \ \sigma_1 = 0.109$$

$$\sigma_2^2 = S_2^2 - \frac{S_{12}S_{23}}{S_{13}} = 0.023, \ \sigma_2/\beta_2 = 0.148$$

$$\sigma_3^2 = S_3^2 - \frac{S_{13}S_{23}}{S_{12}} = 0.02542, \ \sigma_3/\beta_3 = 0.162$$

$$\sigma^2 = \frac{S_{12}S_{13}}{S_{12}} = 0.153, \ \sigma = 0.391$$

$$ar{Z}_1 = ar{\zeta} = 0.7689$$

$$\bar{Z}_2 = \alpha_2 + \beta_2 \bar{\zeta} = 0.7852$$

$$\bar{Z}_3 = \alpha_3 + \beta_3 \bar{\zeta} = 0.7849$$

$$\bar{\zeta} = \bar{Z}_1 = 0.769$$
 $c_2 = \bar{Z}_2 = \frac{S_{23}}{2} \bar{Z}_1 = -0.006$

$$\alpha_2 = \bar{Z}_2 - \frac{S_{23}}{S_{13}}\bar{Z}_1 = -0.006$$

$$\alpha_3 = \bar{Z}_3 - \frac{S_{23}}{S_{12}}\bar{Z}_1 = 0.029$$



514.05 $\tilde{Z}_1 = \zeta = 0.7689$ $\tilde{Z}_2 = \alpha_2 + \beta_2 \tilde{\zeta} = 0.7852$ $\tilde{Z}_3 = \alpha_3 + \beta_3 \tilde{\zeta} = 0.7849$ $\sigma_1^2 = S_1^2 - \frac{S_{12}S_{13}}{2} = 0.01196, \ \sigma_1 = 0.109$ $\sigma_2^2 = S_2^2 - \frac{S_{12}S_{23}}{C} = 0.023, \ \sigma_2/\beta_2 = 0.148$ Method of Moments Estimates – $\alpha_2 = \tilde{Z}_2 - \frac{S_{23}}{c}\tilde{Z}_1 = -0.005$ $\sigma_3^2 = S_3^2 - \frac{S_{13}S_{23}}{S_{13}} = 0.02542, \ \sigma_3/\beta_3 = 0.162$ $\alpha_3 = \tilde{Z}_3 - \frac{S_{23}}{a}\tilde{Z}_1 = 0.029$ Finkelstein (Z_1) , Fontana (Z_2) , & Pfor $(Z_3)^{\frac{(2.5)}{11}-0.151,\,\sigma-0.391}$

1. The solutions are easy enough to do by hand but usually we would use a computer program.

R. A. Bilonick 514.05 January 15, 2021

Results Using merror R Package – Converted to $eta_1=1$ Constraint

	п	σ_i	σ_{σ_i}	α_i	β_i	LB	UB	$\frac{\sigma_i}{\beta_i}$
Z_{fink}	1424	0.109	0.031	-0.008	0.996	0.101	0.118	0.110
Z_{font}	1424	0.152	0.035	-0.014	1.025	0.145	0.161	0.149
Z_{pfor}	1424	0.159	0.036	0.021	0.979	0.152	0.168	0.163
True Redshift	1424	0.392	0.077			0.378	0.408	

The merror package uses the constraint that the geometric average of the β s equals 1. To convert this so that instead, $\beta_1 = 1$, we divide each β by $\beta_1 = 0.996$:

Method	σ_{i}	$\beta_1\beta_2\beta_3=1$	$eta_1=1$	σ_i/β_i
Finkelstein	0.109	0.996	1.000	0.109
Fontana	0.152	1.025	1.029	0.148
Pfor	0.159	0.979	0.983	0.162

514.05

R. A. Bilonick



- 1. There are many ways to impose the constraints.
- 2. Which ever way you impose the constraints, the conclusions in relative terms are always the same
- 3. The calibration curves are invariant to the constraints.

Results Using OpenMx R Package

Parameter estimates for a single-factor model for only photometric redshifts for Finkelstein, Fontana, and Pfor. These estimates are the same as previously estimated by the method of moments. Bootstrap sampling (bias corrected) was used to approximate the 95% intervals.

	Lower	Estimate	Upper			
		Scale Bias				
eta_{fink}		1.000				
$eta_{ extsf{font}}$	0.980	1.029	1.090			
$eta_{ extit{pfor}}$	0.941	0.983	1.021			
Bias Intercept						
α_{fink}		0.000				
$lpha_{\mathit{font}}$	-0.049	-0.006	0.028			
α_{nfor}	0.004	0.029	0.057			

	Lower	Estimate	Upper	
	SD for	True Redshif	t Values	
σ_p	0.359	0.391	0.416	
	Imprecision SDs			
σ_{fink}	0.064	0.109	0.132	
$\sigma_{ extit{font}}/eta_{ extit{font}}$	0.082	0.148	0.185	
$\sigma_{ m pfor}/eta_{ m pfor}$	0.091	0.162	0.192	

ļ	0.109	0.132	
<u> </u>	0.148	0.185	
	0.162	0.192	
↓ □	→ ←	← 豊 → - 豊	99 0

2021-02-01

R. A. Bilonick 514.05 January 15, 2021 514.05 Eastern and Blog. There estimates are the same as associately estimated by the mothed of Scale Rias SD for True Redshift Values 0.980 1.029 1.09 a- 0.350 0.301 0.416 0.941 0.983 1.02 Results Using OpenMx R Package σ_{Enk} 0.054 0.109 0.132

1. The OpenMx model yields the same results as the merror package – so long as the same constraints are used in both.

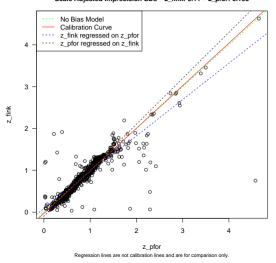
region -0.049 -0.006 0.028

 $\sigma_{fant}/\beta_{fant}$ 0.082 0.148 0.185 $\sigma_{plar}/\beta_{plar}$ 0.091 0.162 0.192

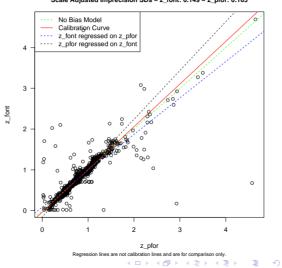
- 2. Bootstrapping should provide more realistic confidence intervals.
- 3. OpenMx can provide confidence intervals for any function of the parameters.

Calibration Plots - Invariant to the Constraints

Calibration Curve: z_fink = -0.03 + 1.017 z_pfor and z_pfor = 0.029 + 0.983 z_fink Scale Adjusted Imprecision SDs - z_fink: 0.11 - z_pfor: 0.163



Calibration Curve: z_font = -0.036 + 1.047 z_pfor and z_pfor = 0.035 + 0.955 z_font
Scale Adjusted Imprecision SDs - z_font: 0.149 - z_pfor: 0.163



R. A. Bilonick 514.05 January 15, 2021 14 / 17



- 1. The calibration curves show how wrong ordinary regression can be.
- 2. You won't know exactly how wrong the ordinary regression curves will be until you fit the correct measurement error model.
- 3. In my experience, researchers often use the regression curve the inferior of the two curves.

Conclusions

R A Bilonick

- Ordinary regression causes the estimate of the scale bias to be biased toward 0. This bias was large for the photometric methods. (It also occurs with spectroscopic redshift, but is smaller in magnitude).
- Before you can compare method imprecision, the imprecision standard deviations must be divided by their corresponding scale biases (β s).
- At minimum, you need 3 methods compared on the same objects, or if only two methods, at least one must have independent repeats. (Otherwise, you would be forced to make unverifiable assumptions in order to reduce the number of parameters needed to be estimated.)
- Although 3 methods is the minimum, the more methods compared on the same objects the better.
- It is not necessary to have independent repeats so long as you have at least 3 different methods, but having repeats can be useful.
- A one-factor model is not reasonable if you want to include both spectroscopic and photometric methods in the same model. (See 235.06 for an example of including both spectroscopic and photometric methods in the same comparison more than one latent factor is needed.)

514.05

January 15. 2021 15 / 17

514.05

2021-02-01

 $lue{}$ Conclusions

Conclusions

- Ordinary regression causes the estimate of the scale bias to be biased toward 0. This bias was large for the photometric methods. (It also occurs with spectroscopic redshift, but is smaller in magnitude)
- Before you can compare method imprecision, the imprecision standard deviations must be divided by their corresponding scale biases (8s).
- one must have independent repeats. (Otherwise, you would be forced to make unverifiable assumptions in order to reduce the number of parameters needed to be estimated.)
- Although 3 methods is the minimum, the more methods compared on the same objects the better.
 It is not necessary to have independent repeats so long as you have at least 3 different methods,
- A one-factor model is not reasonable if you want to include both spectroscopic and photometric methods in the same model. (See 235.06 for an example of including both spectroscopic and

References I

Richard A. Bilonick, *merror: Accuracy and Precision of Measurements*, 2015, R package version 2.0.2.

Steven M. Boker, Michael C. Neale, Hermine H. Maes, Michael J. Wilde, Michael Spiegel, Timothy R. Brick, Ryne Estabrook, Timothy C. Bates, Paras Mehta, Timo von Oertzen, Ross J. Gore, Michael D. Hunter, Daniel C. Hackett, Julian Karch, Andreas M. Brandmaier, Joshua N. Pritikin, Mahsa Zahery, Robert M. Kirkpatrick, Yang Wang, Charles Driver, Massachusetts Institute of Technology, S. G. Johnson, Association for Computing Machinery, Dieter Kraft, Stefan Wilhelm, Sarah Medland, Carl F. Falk, Matt Keller, Manjunath B G, and The Regents of the University of California., *Openmx 2.9.9 user guide*, 2018.

Michael D. Hunter, State Space Modeling in an Open Source, Modular, Structural Equation Modeling Environment, Structural Equation Modeling (in press), 1–18.



514.05

erences I

- chard A. Bilonick, merror: Accuracy and Precision of Measurements, 2015, R package
- Strown M. Boher, Michael C. Mash, Hormins H. Mass, Michael J. Wilds, Michael Spiegel, Truntely P. Botk, Spie Estabuch, Timotty C. Batts, Para Michael, Timo von Ostetton, 1800a. J. Gorn, Michael D. Huster, Daniel C. Halckett, Jahlus Karch, Andreas M. Brandmaier, Joshua N. Poritkin, Mahasa Zahery, Robert M. Kriepatrick, Yang Wong, Charles Driver, Manaschaustis Institute of Technology, S. G. Johnson, Association for Computing Machinery, Deter Furth, Study Whelm, Sarial Mediland, Carl F. Talk, Matt. Heller, Magnaged Sel. Gard The Regulate of the University of Littlerina, Journes 202 1.

user guste, 2018.

Michael D. Hunter, State Space Modeling in an Open Source, Modular, Structural Equation Modeling Emvironment, Structural Equation Modeling (in press), 1–18.

R. A. Bilonick 514.05 January 15, 2021 16 / 17

References II

- JL Jaech, Statistical Analysis of Measurement Errors, John Wiley & Sons, Inc., New York, NY, USA (1985).
- Michael C. Neale, Michael D. Hunter, Joshua N. Pritikin, Mahsa Zahery, Timothy R. Brick, Robert M. Kirkpatrick, Ryne Estabrook, Timothy C. Bates, Hermine H. Maes, and Steven M. Boker, *OpenMx 2.0: Extended structural equation and statistical modeling*, Psychometrika **81** (2016), no. 2, 535–549.
 - Joshua N. Pritikin, Michael D. Hunter, and Steven M. Boker, *Modular Open-source Software for Item Factor Analysis*, Educational and Psychological Measurement **75** (2015), no. 3, 458–474.
- R Core Team, R: A language and environment for statistical computing, R Foundation for Statistical Computing, Vienna, Austria, 2018.



2021-02-01

R. A. Bilonick 514.05 January 15, 2021 17 / 1