

Integration

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This portfolio covers the intuition behind MCMC methods.

Properties of Markov Chains

Definitions

We define a discrete Markov chain $\mathbf{X} := (X_n)_{n \geq 0}$ on a measurable space (x, \mathcal{X}) , so that for $A \in \mathcal{X}$ we have

$$P(X_n \in A | X_0 = x, \dots, X_{n-1} = x_{n-1}) = P(X_n \in A | X_{n-1} = x_{n-1}) \quad (1)$$

LLN and CLT

We now define some theorems for the case of Markov Chain

Theorem: LLN For a time-invariant, positive Harris Markov chain $\mathbf{X} = (X_n)_{n \geq 0}$ with stationary distribution π and any $f \in L_1(X, \pi) = \{f : \pi(|f|) < \infty\}$

$$\lim_{n \rightarrow \infty} \frac{1}{n} S_n(f) = \pi(f) \quad (2)$$

almost surely for any initial distribution.

Theorem: CLT For a time-invariant, positive Harris and geometrically ergodic Markov chain $\mathbf{X} = (X_n)_{n \geq 0}$ with stationary distribution π and $\pi(|f|^{2+\delta}) < \infty$ for some $\delta > 0$,

$$n^{\frac{1}{2}} \{n^{-1} S_n(f) - \pi(f)\} \xrightarrow{L} \mathcal{N}(0, \sigma^2(f)) \quad (3)$$

as $n \rightarrow \infty$, where $\bar{f} = f - \pi(f)$ and

$$\sigma^2(f) = \mathbb{E}_X [\bar{f}(X_0)^2] + 2 \sum_{k=1}^{\infty} \mathbb{E}_{\pi} [\bar{f}(X_0) \bar{f}(X_k)] < \infty \quad (4)$$

Metropolis-Hastings

The most common method of constructing Markov chains is the Metropolis-Hastings algorithm.

Algorithm

We assume π has a density w.r.t a measure λ .

We then specify a proposal transition kernel Q , with density q with respect to λ :

$$Q(x, dz) = q(x, z)\lambda(dz) \quad (5)$$

Aiming to obtain $P_{MH}(x, \cdot)$, we perform two steps

1. Simulate $x^* \sim Q(x, \cdot)$
2. Take $x = x^*$ with probability

$$\alpha_{MH}(x, x^*) := \min \left\{ 1, \frac{\pi(x^*)q(x^*, x)}{\pi(x)q(x, x^*)} \right\} \quad (6)$$

Thus we only need π up to a constant and to simulate from $Q(x, \cdot)$ to obtain sample from P_{MH} .

The code for this is now given below, where the input **Q** is a list containing a **density** function and a **sample** function:

```
make.MH_kernel <- function(pi,Q){  
  q <- Q$density  
  P <- function(x) {  
    x_new <- Q$sample(x)  
    alpha <- min(1, pi(x_new)*q(x_new,x)/pi(x)/q(x,x_new))  
    ifelse(runif(1) < alpha, x_new, x)  
  }  
  return(P)  
}
```

Normal Distribution Example

We first need to construct a function which produces the **Q** as described above for a normal distribution. For simplicity, we will only consider the univariate proposal

```
make.normal_Q <- function(sigma) {  
  Q <- list()  
  Q$sample <- function(x) {  
    x + sigma*rnorm(1)  
  }  
  Q$density <- function(x,y) {  
    dnorm(y-x, sd=sigma)  
  }  
  return(Q)  
}
```

Can now use these two functions to obtain a **P** we can simulate a Markov chain from

```
Q_normal <- make.normal_Q(1)  
P <- make.MH_kernel(dnorm, Q_normal)
```

Now we can obtain a Markov chain:

```
n <- 100000

xs<- rep(0,n)
x<- 1

for (i in 1:n) {
  x <- P(x)
  xs[i] <- x
}
```