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Statistical Methods in AI (CSE/ECE 471)

Lecture-17: MLE, MAP and Bayesian Estimation

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Ravi Kiran

Center for Visual Information Technology (CVIT), IIIT Hyderabad



$$\text{PROBABILITY} = \frac{\text{EVENT}}{\text{OUTCOMES}}$$



Data – a probability-based perspective

- The basis for Statistical Learning Theory

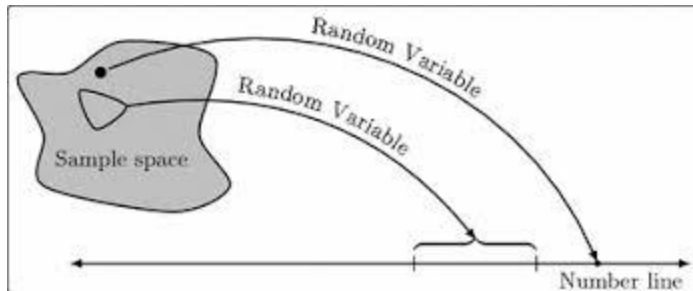
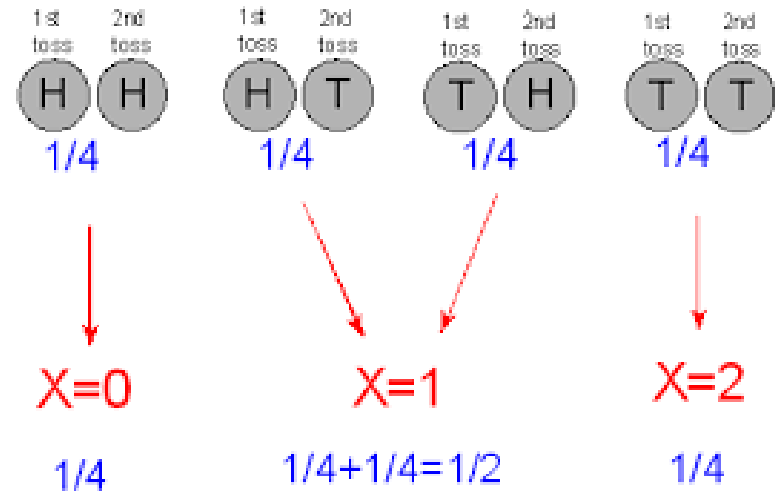
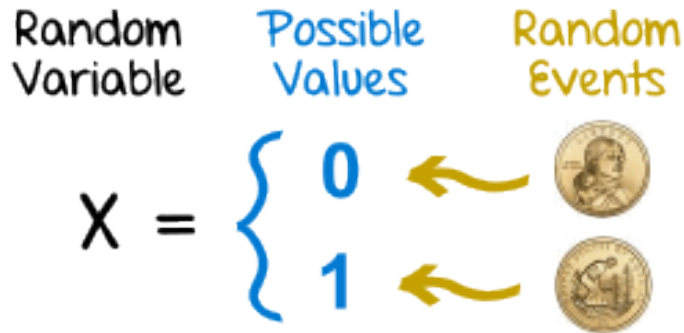


Then we observe candies drawn from some bag: ● ● ● ● ● ● ● ● ● ●

- Domain described by random variables (r.v.)
 - $X = \{\text{apple, grape}\}$
 - $b_i \in [1,5]$
- **Data = Instantiation of some or all r.v.'s in the domain**

Random Variables

R.V. = A numerical value from a random experiment

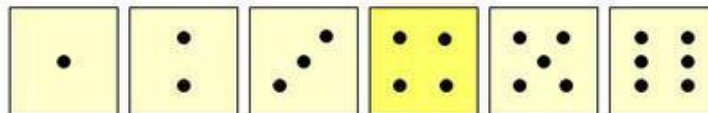




Discrete Random Variables

- Can only take on a countable number of values

Examples:



- Roll a die twice**

Let X be the number of times 4 comes up
(then X could be 0, 1, or 2 times)

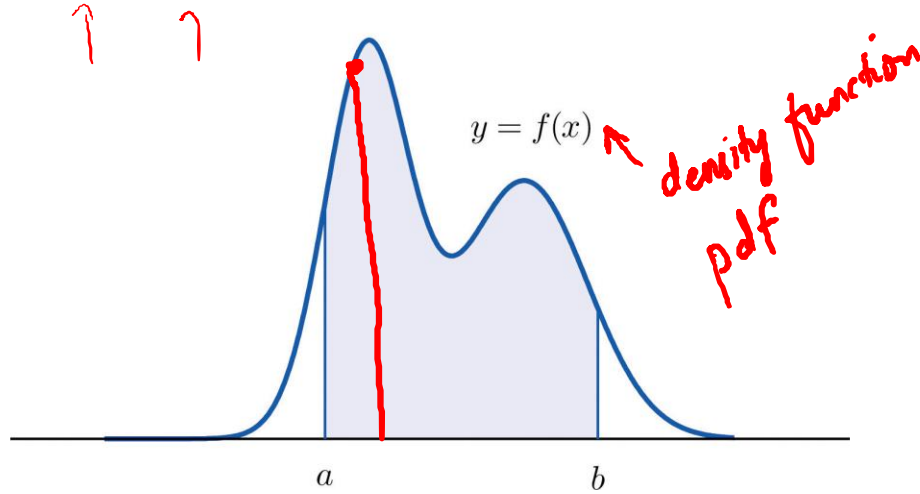
- Toss a coin 5 times.**

Let X be the number of heads
(then $X = 0, 1, 2, 3, 4, \text{ or } 5$)

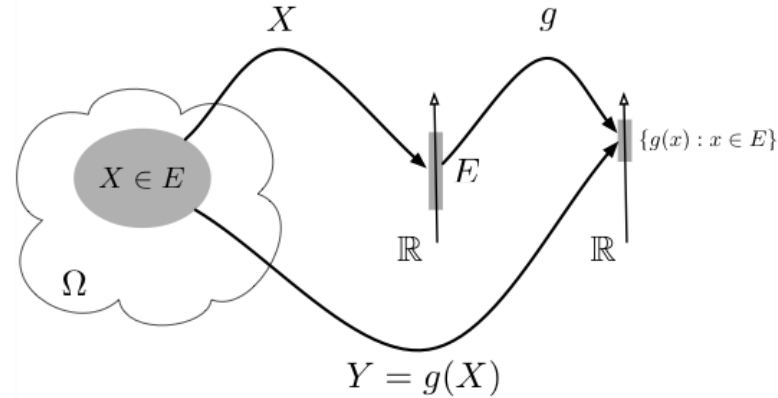


Continuous random variable

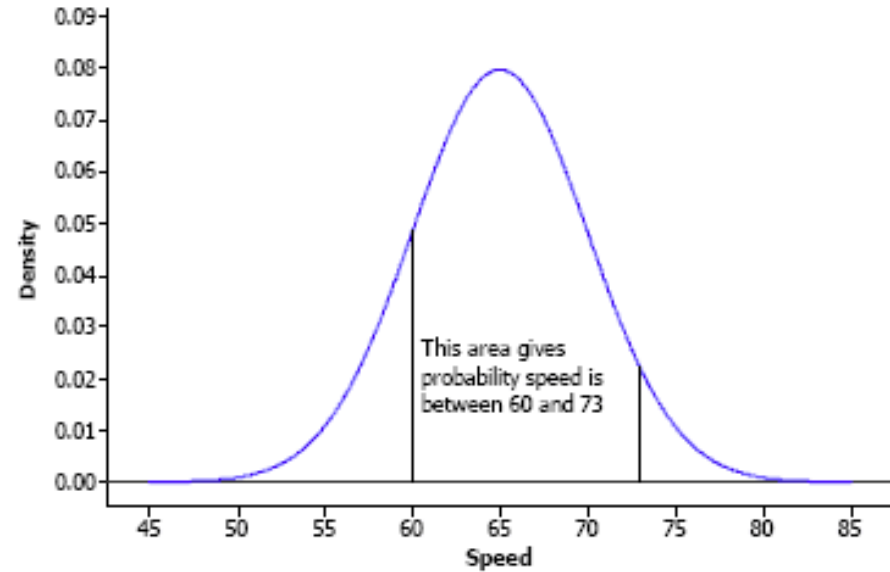
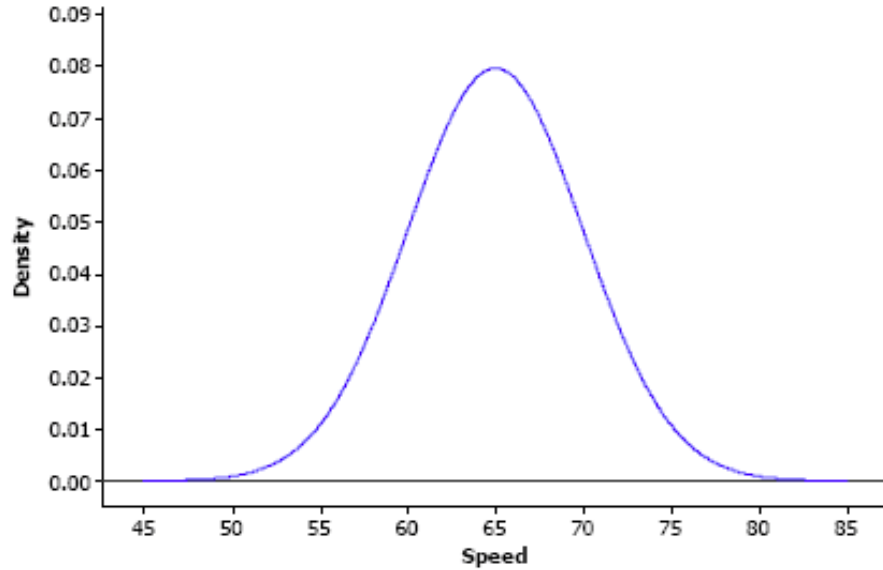
$P(a < X < b) = \text{area of shaded region}$



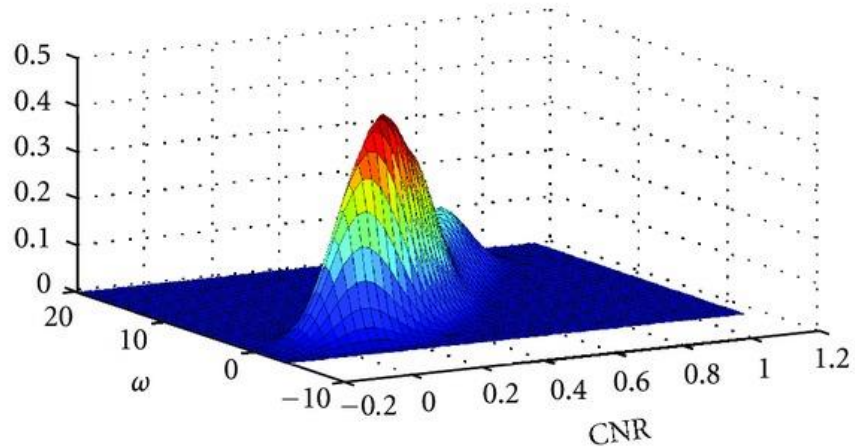
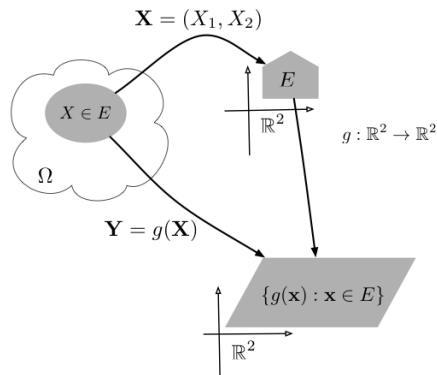
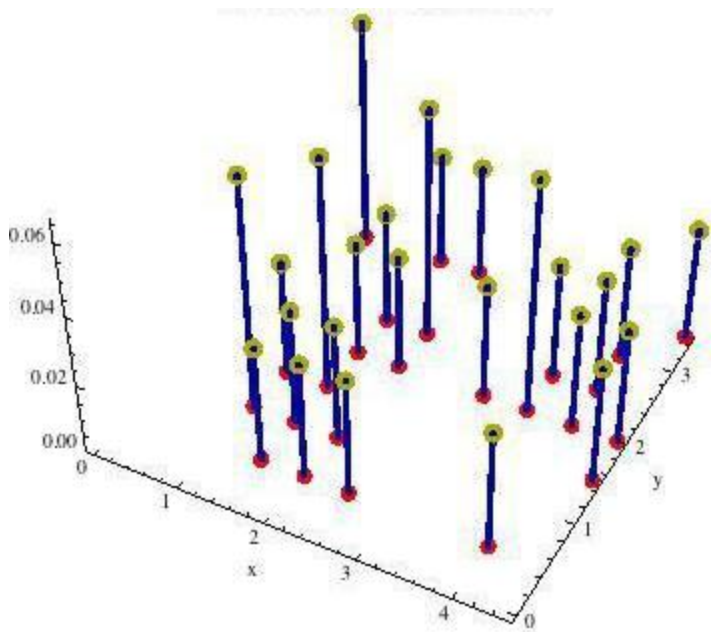
$$P(Y = 0.2) = 0$$



Continuous random variable



Random vectors



Data \rightarrow r.v.

Relative frequency

Relative frequency is the same as **experimental probability**.
We use relative frequency to predict probabilities from experimental data.

The experiment

This spinner was spun 40 times and the results recorded in this table:



Colour	Frequency
Blue	20
Yellow	10
Red	5
Green	5

Relative frequency

$$\frac{\text{frequency of event}}{\text{total number of trials}}$$

Event means **one possible outcome**;
here, one colour on the spinner.

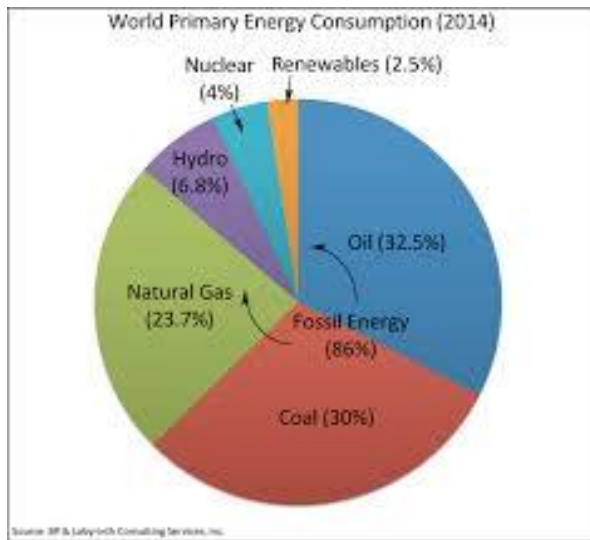
There were 20 blues recorded...

$$P(\text{blue}) = \frac{20}{40}$$

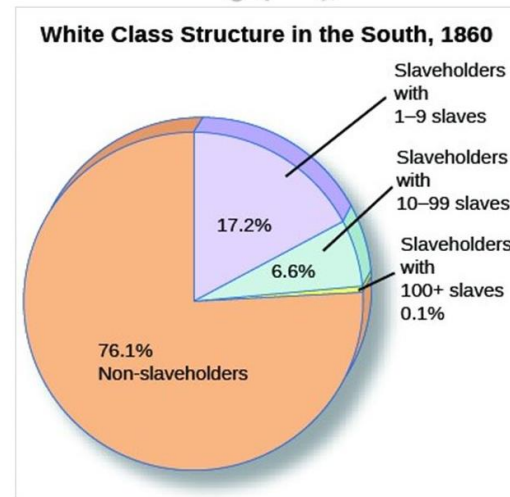
...out of 40 spins.

$$\text{Simplify: } P(\text{blue}) = \frac{20}{40} = \frac{2}{4} = \frac{1}{2}$$

Discrete Prior distributions

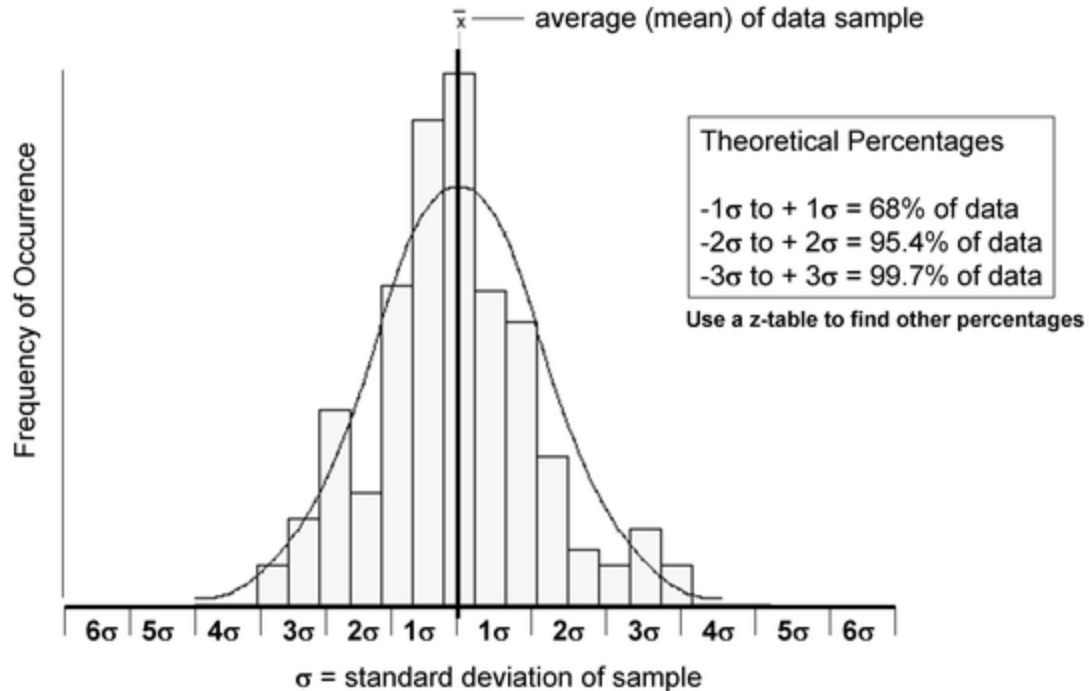


Slave-Owning Population (1860)



Data → r.v.

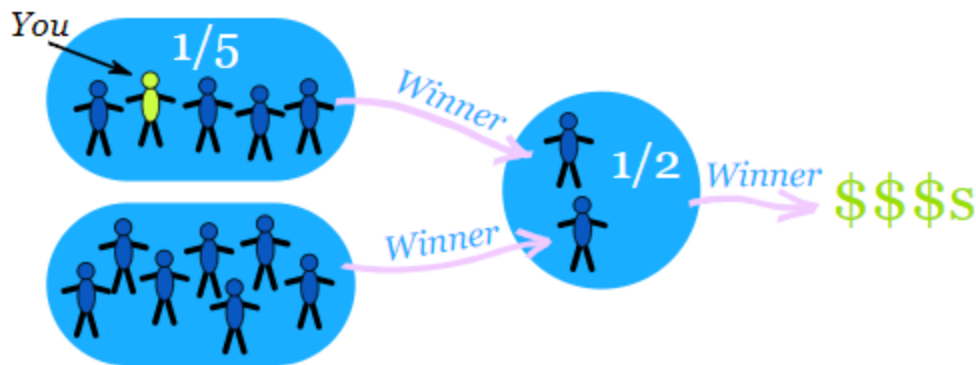
Normal Distribution Curve, Fit to a Histogram



Independent Events

Imagine there are two groups:

- A member of each group gets randomly chosen for the winners circle,
- **then** one of those gets randomly chosen to get the big money prize:



What is your chance of winning the big prize?

Independent vs. Dependent Events



Using the bag of marbles on the left, what is the probability of pulling a black marble two times in a row? $P(\text{black}, \text{black})$

When you put 1st marble back in
(*Independent Events*)

$$\frac{2}{10} * \frac{2}{10}$$

$$\frac{1}{5} * \frac{1}{5} = \frac{1}{25}$$

When you KEEP 1st marble
(*Dependent Events*)

$$\frac{2}{10} * \frac{1}{9}$$

$$\frac{1}{5} * \frac{1}{9}$$

Independent Events

The outcome of one event **does not** affect the outcome of the other.

If A and B are independent events then the probability of both occurring is


$$P(A \text{ and } B) = P(A) \times P(B)$$

Dependent Events

The outcome of one event affects the outcome of the other.

If A and B are dependent events then the probability of both occurring is

$$P(A \text{ and } B) = P(A) \times P(B|A)$$



Probability of B given A

Independent vs. Dependent Events



Using the bag of marbles on the left, what is the probability of pulling a black marble two times in a row? $P(\text{black, black})$

When you put 1st marble back in
(*Independent Events*)

$$\frac{2}{10} * \frac{2}{10}$$

$$\frac{1}{5} * \frac{1}{5} = \frac{1}{25}$$

When you KEEP 1st marble
(*Dependent Events*)

$$\frac{2}{10} * \frac{1}{9}$$



$$\frac{1}{5} * \frac{1}{9}$$

$$P(A \text{ and } B) = P(A) \times P(B)$$


$$P(A \text{ and } B) = P(A) \times P(B | A)$$

Probability of B given A

Marginal Probabilities

	$\begin{cases} x = 1 & (\text{Rains}) \\ x = 0 & (\text{Doesn't rain}) \end{cases}$	$\begin{cases} \Pr(x = 1) = 0.6 \\ \Pr(x = 0) = 0.4 \end{cases}$
	 $\begin{cases} y = 1 & (\text{Have umbrella}) \\ y = 0 & (\text{Don't have umbrella}) \end{cases}$	$\begin{cases} \Pr(y = 1) = 0.3 \\ \Pr(y = 0) = 0.7 \end{cases}$

Joint Probability

	$\begin{cases} x = 1 & (\text{Rains}) \\ x = 0 & (\text{Doesn't rain}) \end{cases}$	$\begin{cases} \Pr(x = 1) = 0.6 \\ \Pr(x = 0) = 0.4 \end{cases}$
	$\begin{cases} y = 1 & (\text{Have umbrella}) \\ y = 0 & (\text{Don't have umbrella}) \end{cases}$	$\begin{cases} \Pr(y = 1) = 0.3 \\ \Pr(y = 0) = 0.7 \end{cases}$

$$\Pr(x = 0) = \sum_{y=0}^1 \Pr(x=0, y)$$

$$= \Pr(x=0, y=0) + \Pr(x=0, y=1)$$

$$= 0.28 + 0.12 = 0.4$$

Case 1: Rains but you have an umbrella

$$\begin{aligned} \Pr(x = 1, y = 1) &= \Pr(x = 1) \times \Pr(y = 1) \\ &= 0.6 \times 0.3 \\ &= 0.18 \end{aligned}$$

Case 2: Rains but you DON'T have an umbrella

$$\begin{aligned} \Pr(x = 1, y = 0) &= \Pr(x = 1) \times \Pr(y = 0) \\ &= 0.6 \times 0.7 \\ &= 0.42 \end{aligned}$$

Conditional Probability



Given

$x = 1$ (Rains)



What's the Probability of
 $y = 1$ (Bring umbrella)



$\begin{cases} x = 1 & \text{(Rains)} \\ x = 0 & \text{(Doesn't rain)} \end{cases}$

$\begin{cases} y = 1 & \text{(Have umbrella)} \\ y = 0 & \text{(Don't have umbrella)} \end{cases}$

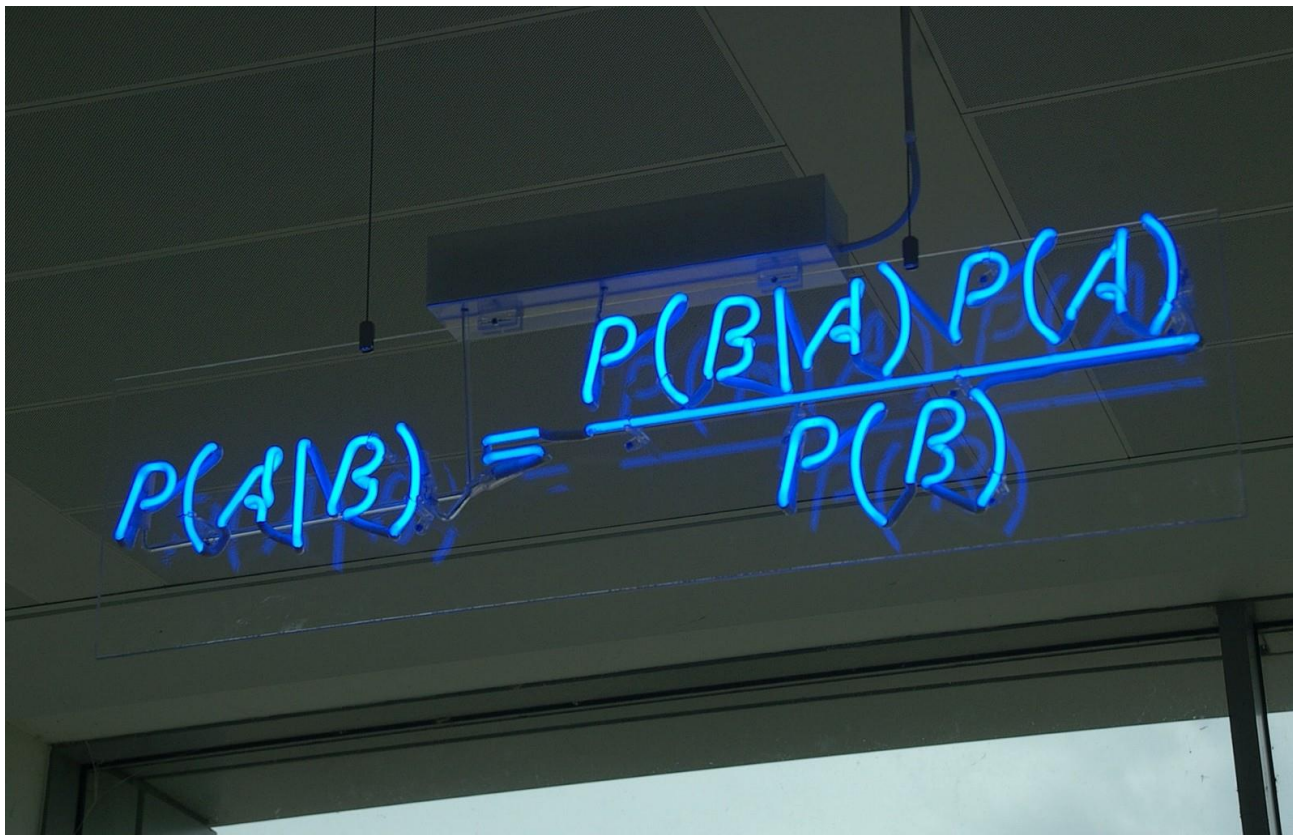
$$\Pr(x = 1) = 0.6$$

$$\Pr(x = 0) = 0.4$$

$$\Pr(y = 1) = 0.3$$

$$\Pr(y = 0) = 0.7$$

Bayes' Rule



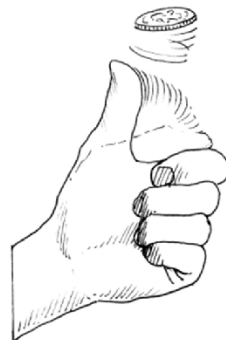
A photograph of a blue neon sign mounted on a ceiling, displaying the formula for Bayes' Rule. The sign is illuminated with a bright blue light, and the background is dark. The formula is written in a stylized, handwritten font. The sign is slightly tilted and has some visible wear and tear, including a small hole in the neon tube.

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

Our first machine learning problem:

Parameter estimation: MLE, MAP

Estimating Probabilities



Flipping a Coin

I have a coin, if I flip it, what's the probability that it will fall with the head up?

Let us flip it a few times to estimate the probability:



The estimated probability is: $\frac{3}{5}$ "Frequency of heads"

Flipping a Coin



The estimated probability is: $3/5$ "Frequency of heads"

Questions:

- (1) Why frequency of heads???
- (2) How good is this estimation???
- (3) Why is this a machine learning problem???

We are going to answer these questions

Question (1)

Why frequency of heads???

- Frequency of heads is exactly the *maximum likelihood estimator* for this problem
- MLE has nice properties
(interpretation, statistical guarantees, simple)

Maximum Likelihood Estimation

MLE for Bernoulli distribution

Data, $D =$



$$D = \{X_i\}_{i=1}^n, X_i \in \{H, T\}$$

$$P(\text{Heads}) = \theta, P(\text{Tails}) = 1 - \theta$$

Flips are **i.i.d.**:

- **Independent** events
- **Identically distributed** according to Bernoulli distribution

MLE: Choose θ that maximizes the probability of observed data

Maximum Likelihood Estimation

MLE: Choose θ that maximizes the probability of observed data

$$\begin{aligned}\hat{\theta}_{MLE} &= \arg \max_{\theta} P(D | \theta) \leftarrow p(x_1, x_2, \dots, x_n | \theta) \\ &= \arg \max_{\theta} \prod_{i=1}^n P(X_i | \theta) \quad \text{Independent draws} \\ &= \arg \max_{\theta} \prod_{i: X_i = H} \theta \prod_{i: X_i = T} (1 - \theta) \quad \text{Identically distributed} \\ &= \arg \max_{\theta} \underbrace{\theta^{\alpha_H} (1 - \theta)^{\alpha_T}}_{J(\theta)} \quad \begin{matrix} \alpha_H \\ \alpha_T \end{matrix}\end{aligned}$$

Maximum Likelihood Estimation

MLE: Choose θ that maximizes the probability of observed data

$$\hat{\theta}_{MLE} = \arg \max_{\theta} P(D|\theta)$$

Question (2)

How good is this MLE estimation???

$$E_{\eta}[\alpha_H] = np$$

$$\hat{\theta}_{MLE} = \frac{\alpha_H}{\alpha_H + \alpha_T}$$

$$\checkmark = \underbrace{\frac{\alpha_H}{n}}$$

$$E\left[\frac{\alpha_H}{n}\right] = \frac{np}{n} = p$$

$$E_{\mathcal{D}} \left[(g^{(\mathcal{D})}(\mathbf{x}) - f(\mathbf{x}))^2 \right] = \underbrace{E_{\mathcal{D}} \left[(g^{(\mathcal{D})}(\mathbf{x}) - \bar{g}(\mathbf{x}))^2 \right]}_{\text{var}(\mathbf{x})} + \underbrace{(\bar{g}(\mathbf{x}) - f(\mathbf{x}))^2}_{\text{bias}(\mathbf{x})}$$

$$\left(E_{\eta}[\text{estimator}] - \theta_p \right)^2$$

How many flips do I need?

I flipped the coins 5 times: 3 heads, 2 tails

$$\hat{\theta}_{MLE} = \frac{3}{5}$$

What if I flipped 30 heads and 20 tails?

$$\hat{\theta}_{MLE} = \frac{30}{50}$$

- Which estimator should we trust more?
- The more the merrier???

Simple bound

Let θ^* be the true parameter.

For $n = \alpha_H + \alpha_T$, and $\hat{\theta}_{MLE} = \frac{\alpha_H}{\alpha_H + \alpha_T}$

For any $\epsilon > 0$:

Hoeffding's inequality:

$$P(|\hat{\theta} - \theta^*| \geq \epsilon) \leq 2e^{-2n\epsilon^2}$$



Probably Approximate Correct (PAC) Learning

I want to know the coin parameter θ , within $\epsilon = 0.1$ error with probability at least $1 - \delta = 0.95$.

How many flips do I need?

$$P(|\hat{\theta} - \theta^*| \geq \epsilon) \leq \underline{2e^{-2n\epsilon^2}} \leq \underline{\delta} \quad \sigma.95$$
$$e^{-2n\epsilon^2} \leq \frac{\delta}{2}$$
$$-2n\epsilon^2 \leq \ln\left(\frac{\delta}{2}\right) \quad \Rightarrow \quad n \geq \frac{1}{2\epsilon^2} \ln\left(\frac{2}{\delta}\right)$$
$$\ln\left(\frac{2}{\delta}\right) \leq 2n\epsilon^2$$

Sample complexity:

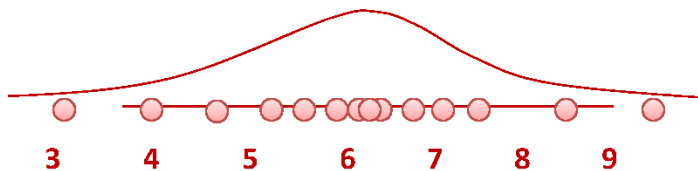
$$n \geq \frac{\ln(2/\delta)}{2\epsilon^2}$$

Question (3)

Why is this a machine learning problem???

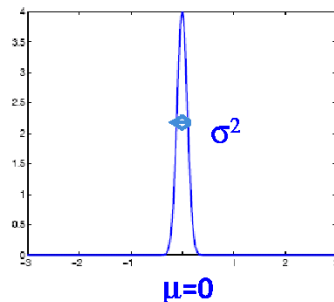
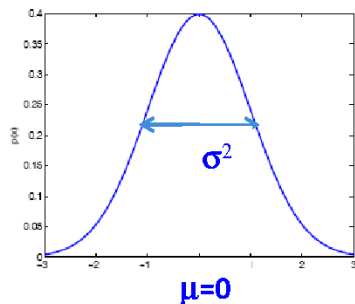
- improve their **performance** (accuracy of the predicted prob.)
- at some **task** (predicting the probability of heads)
- with **experience** (the more coins we flip the better we are)

What about continuous features?

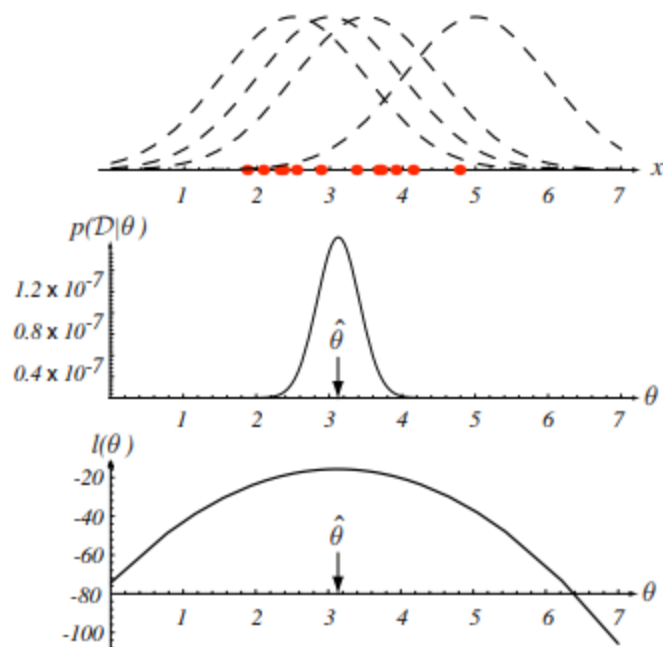


Let us try Gaussians...

$$p(x | \mu, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right) = \mathcal{N}_x(\mu, \sigma)$$



Example: Maximum Likelihood Estimate of the Mean



MLE for Gaussian mean and variance

Choose $\theta = (\mu, \sigma^2)$ that maximizes the probability of observed data

$$\begin{aligned}\hat{\theta}_{MLE} &= \arg \max_{\theta} P(D | \theta) \\ &= \arg \max_{\theta} \prod_{i=1}^n P(X_i | \theta) && \text{Independent draws} \\ &= \arg \max_{\theta} \prod_{i=1}^n \frac{1}{2\sigma^2} e^{-(X_i - \mu)^2 / 2\sigma^2} && \text{Identically distributed} \\ &= \arg \max_{\theta = (\mu, \sigma^2)} \underbrace{\frac{1}{2\sigma^2} e^{-\sum_{i=1}^n (X_i - \mu)^2 / 2\sigma^2}}_{J(\theta)}\end{aligned}$$

MLE for Gaussian mean and variance

$$\hat{\mu}_{MLE} = \frac{1}{n} \sum_{i=1}^n x_i$$
$$\hat{\sigma}_{MLE}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mu})^2$$

Note: MLE for the variance of a Gaussian is **biased**

[Expected result of estimation is **not** the true parameter!]

Unbiased variance estimator: $\hat{\sigma}_{unbiased}^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \hat{\mu})^2$

$$E[\hat{\sigma}_{MLE}^2] \neq \sigma^2 \quad E[\hat{\sigma}_{unbiased}^2] = \sigma^2$$

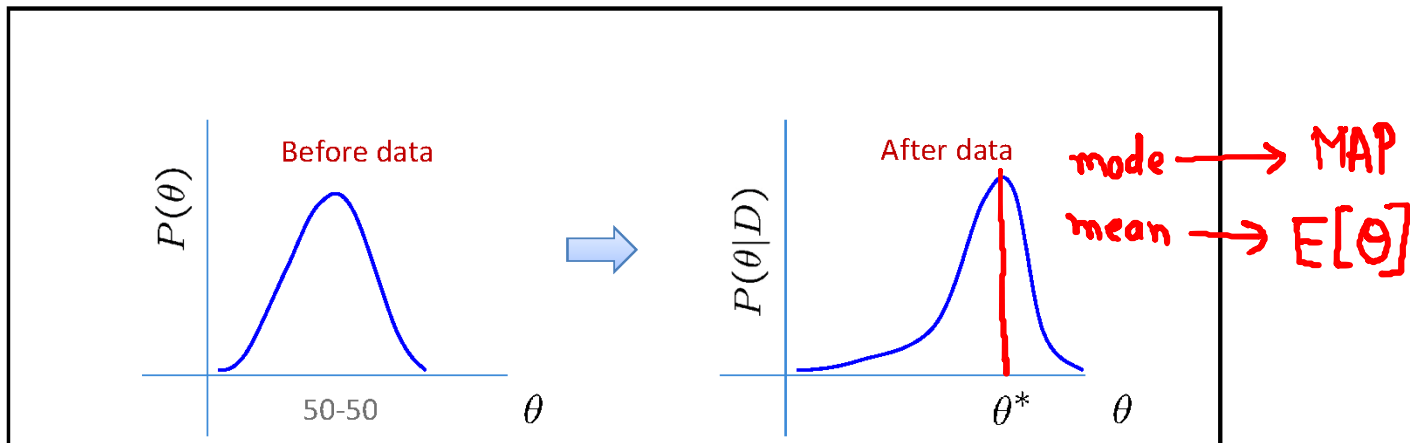
What about prior knowledge?
(MAP Estimation)

What about prior knowledge?

We know the coin is “close” to 50-50. What can we do now?

The Bayesian way...

Rather than estimating a single θ , we obtain a distribution over possible values of θ



Prior distribution

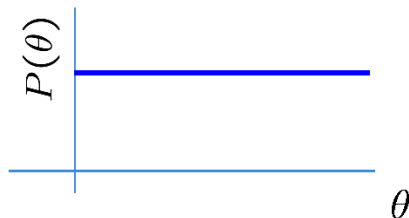
What prior? What distribution do we want for a prior?

- Represents expert knowledge (philosophical approach)
- Simple posterior form (engineer's approach)

$$p(\theta|D) = \frac{\overbrace{p(D|\theta)p(\theta)}}{p(D)}$$

Uninformative priors:

- Uniform distribution



Conjugate priors:

- Closed-form representation of posterior
- $P(\theta)$ and $P(\theta|D)$ have the same form

In order to proceed we will need:

Bayes Rule



Bayes, Thomas (1763) An essay towards solving a problem in the doctrine of chances. *Philosophical Transactions of the Royal Society of London*, 53:370-418

Chain Rule & Bayes Rule

Chain rule:

$$P(X, Y) = P(X|Y)P(Y) = P(Y|X)P(X)$$

Bayes rule:

$$P(X|Y) = \frac{P(Y|X)P(X)}{P(Y)}$$

Bayes rule is important for reverse conditioning.

Bayesian Learning

- Use Bayes rule:

$$P(\theta \mid \mathcal{D}) = \frac{P(\mathcal{D} \mid \theta)P(\theta)}{P(\mathcal{D})}$$

- Or equivalently:

$$P(\theta \mid \mathcal{D}) \propto P(\mathcal{D} \mid \theta)P(\theta)$$

posterior likelihood prior

MLE vs. MAP

- Maximum Likelihood estimation (MLE)

Choose value that maximizes the probability of observed data

$$\hat{\theta}_{MLE} = \arg \max_{\theta} P(D|\theta)$$



- Maximum *a posteriori* (MAP) estimation

Choose value that is most probable given observed data and prior belief

$$\begin{aligned}\hat{\theta}_{MAP} &= \arg \max_{\theta} P(\theta|D) \\ &= \arg \max_{\theta} P(D|\theta)P(\theta)\end{aligned}$$

When is MAP same as MLE?

MAP estimation for Binomial distribution

Coin flip problem: Likelihood is Binomial



$$P(\mathcal{D} \mid \theta) = \binom{n}{\alpha_H} \theta^{\alpha_H} (1 - \theta)^{\alpha_T}$$

If the prior is Beta distribution,

$$P(\theta) = \frac{\theta^{\beta_H-1} (1 - \theta)^{\beta_T-1}}{B(\beta_H, \beta_T)} \sim \text{Beta}(\beta_H, \beta_T)$$

\Rightarrow posterior is Beta distribution

Beta function: $B(x, y) = \int_0^1 t^{x-1} (1 - t)^{y-1} dt$

MAP estimation for Binomial distribution

Likelihood is Binomial: $P(\mathcal{D} | \theta) = \binom{n}{\alpha_H} \theta^{\alpha_H} (1 - \theta)^{\alpha_T}$

Prior is Beta distribution: $P(\theta) = \frac{\theta^{\beta_H-1} (1 - \theta)^{\beta_T-1}}{B(\beta_H, \beta_T)} \sim \text{Beta}(\beta_H, \beta_T)$

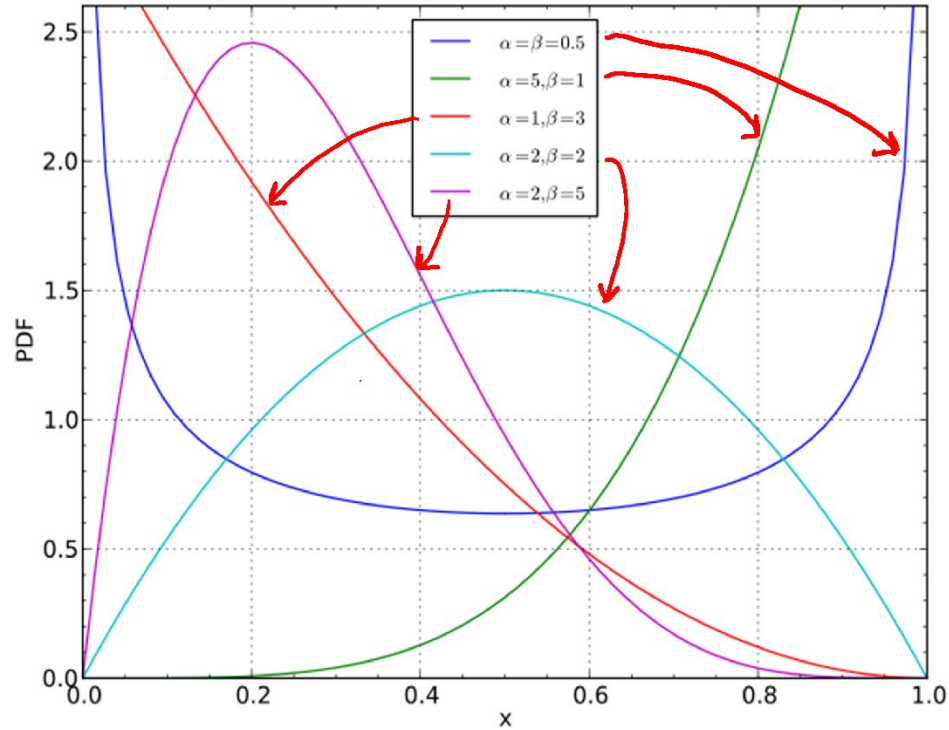
⇒ posterior is Beta distribution

$$P(\theta | D) \sim \text{Beta}(\beta_H + \alpha_H, \beta_T + \alpha_T)$$

$P(\theta)$ and $P(\theta | D)$ have the same form! [Conjugate prior]

$$\begin{aligned} \hat{\theta}_{MAP} &= \arg \max_{\theta} P(\theta | D) = \arg \max_{\theta} P(D | \theta) P(\theta) \\ &= \frac{\alpha_H + \beta_H - 1}{\alpha_H + \beta_H + \alpha_T + \beta_T - 2} \end{aligned}$$

Beta distribution

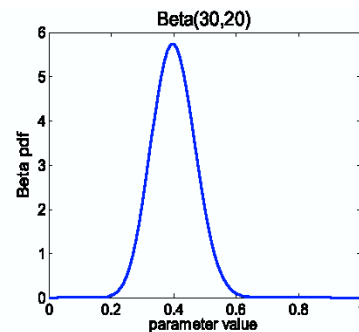
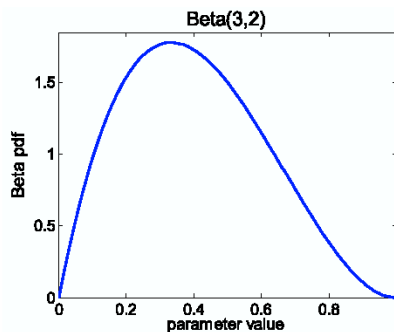
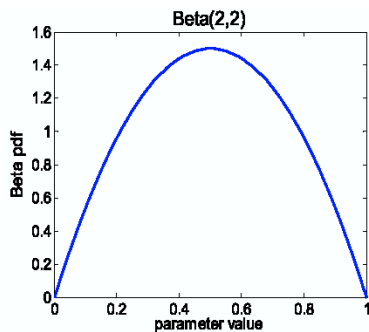


More concentrated as values of α, β increase

Beta conjugate prior

$$P(\theta) \sim \text{Beta}(\beta_H, \beta_T)$$

$$P(\theta|D) \sim \text{Beta}(\beta_H + \alpha_H, \beta_T + \alpha_T)$$



As $n = \alpha_H + \alpha_T$
increases

As we get more samples, effect of prior is “washed out”

From Binomial to Multinomial

Example: Dice roll problem (6 outcomes instead of 2)

Likelihood is $\sim \text{Multinomial}(\theta = \{\theta_1, \theta_2, \dots, \theta_k\})$

$$P(\mathcal{D} | \theta) = \theta_1^{\alpha_1} \theta_2^{\alpha_2} \dots \theta_k^{\alpha_k}$$



If prior is Dirichlet distribution,

$$P(\theta) = \frac{\prod_{i=1}^k \theta_i^{\beta_i - 1}}{B(\beta_1, \dots, \beta_k)} \sim \text{Dirichlet}(\beta_1, \dots, \beta_k)$$

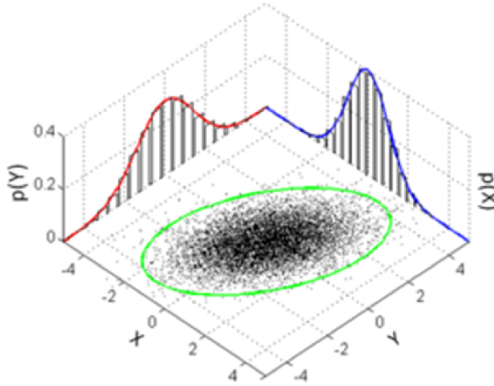
Then posterior is Dirichlet distribution

$$P(\theta | D) \sim \text{Dirichlet}(\beta_1 + \alpha_1, \dots, \beta_k + \alpha_k)$$

For Multinomial, conjugate prior is Dirichlet distribution.

http://en.wikipedia.org/wiki/Dirichlet_distribution

Conjugate prior for Gaussian?



$$(2\pi)^{-\frac{k}{2}} |\Sigma|^{-\frac{1}{2}} e^{-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})' \Sigma^{-1}(\mathbf{x}-\boldsymbol{\mu})},$$

Conjugate prior on mean: **Gaussian**

Conjugate prior on covariance matrix: **Inverse Wishart**

$$\frac{|\Psi|^{\frac{\nu}{2}}}{2^{\frac{\nu p}{2}} \Gamma_p(\frac{\nu}{2})} |\mathbf{X}|^{-\frac{\nu+p+1}{2}} e^{-\frac{1}{2} \text{tr}(\Psi \mathbf{X}^{-1})}$$

Bayesians vs. Frequentists

You are no good when sample is small

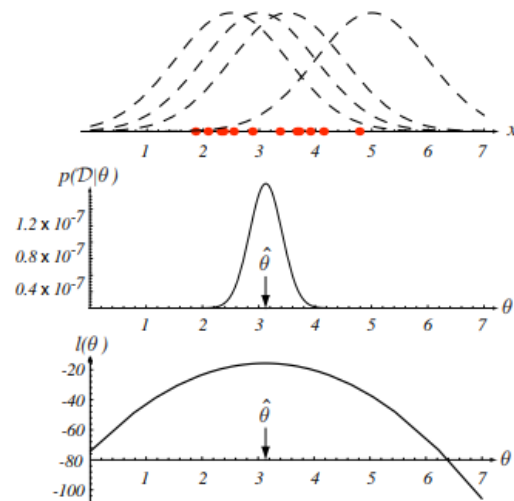
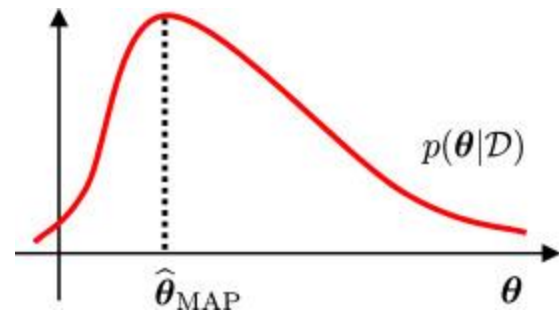


You give a different answer for different priors

$$\begin{aligned}
\hat{\Theta}_{MAP} &= \operatorname{argmax}_{\Theta} \operatorname{prob}(\Theta|\mathcal{X}) \\
&= \operatorname{argmax}_{\Theta} \frac{\operatorname{prob}(\mathcal{X}|\Theta) \cdot \operatorname{prob}(\Theta)}{\operatorname{prob}(\mathcal{X})} \\
&= \operatorname{argmax}_{\Theta} \operatorname{prob}(\mathcal{X}|\Theta) \cdot \operatorname{prob}(\Theta) \\
&= \operatorname{argmax}_{\Theta} \prod_{\mathbf{x}_i \in \mathcal{X}} \operatorname{prob}(\mathbf{x}_i|\Theta) \cdot \operatorname{prob}(\Theta)
\end{aligned}$$

$$\begin{aligned}
\hat{\Theta}_{MAP} &= \operatorname{argmax}_{\Theta} \operatorname{prob}(\Theta|\mathcal{X}) \\
&= \operatorname{argmax}_{\Theta} \frac{\operatorname{prob}(\mathcal{X}|\Theta) \cdot \operatorname{prob}(\Theta)}{\operatorname{prob}(\mathcal{X})} \\
&= \operatorname{argmax}_{\Theta} \operatorname{prob}(\mathcal{X}|\Theta) \cdot \operatorname{prob}(\Theta) \\
&= \operatorname{argmax}_{\Theta} \prod_{\mathbf{x}_i \in \mathcal{X}} \operatorname{prob}(\mathbf{x}_i|\Theta) \cdot \operatorname{prob}(\Theta)
\end{aligned}$$

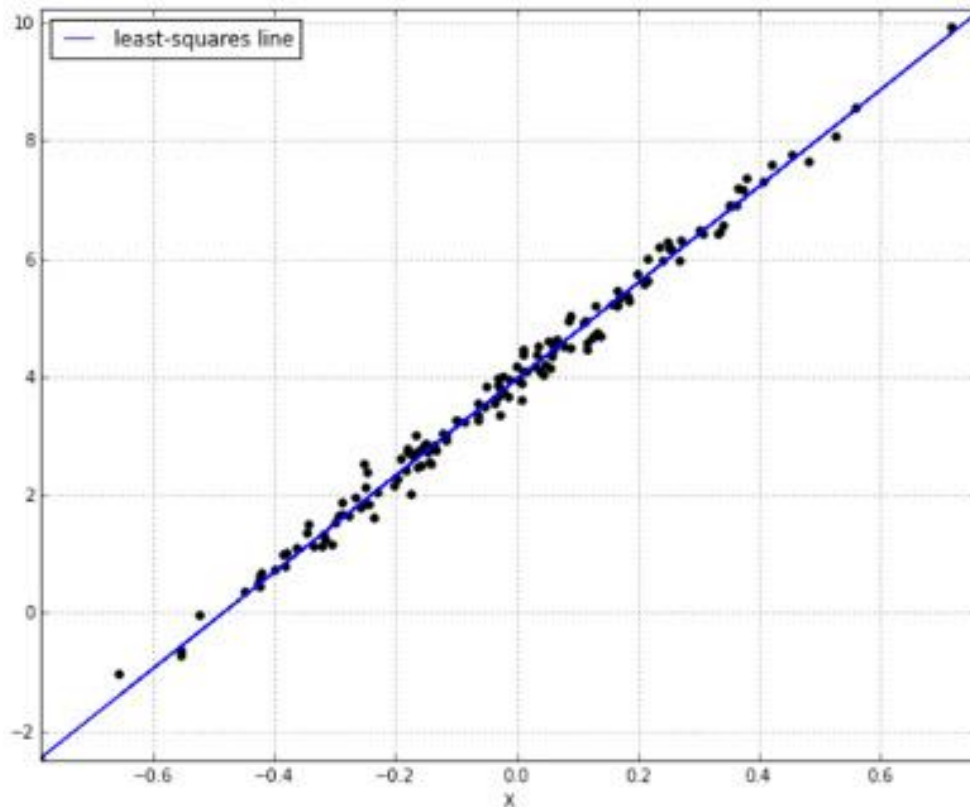
$$\hat{\Theta}_{MAP} = \operatorname{argmax}_{\Theta} \left(\sum_{\mathbf{x}_i \in \mathcal{X}} \log \operatorname{prob}(\mathbf{x}_i|\Theta) + \log \operatorname{prob}(\Theta) \right)$$



$$y_i = \theta^T x_i + \epsilon_i$$

\hat{y}_i

$$\arg \min_{\theta} \sum_{i=1}^n (y_i - \hat{y}_i)^2$$



$$y = 4 + 8x + \text{noise}$$

Linear Regression Param Estimation - A Probabilistic Perspective

$$y_i = \theta^T x_i + \epsilon_i \qquad \epsilon_i \sim \mathcal{N}(0, \sigma^2)$$

$$\begin{matrix} X = \mu + Y \\ Y \sim \mathcal{N}(0, \sigma^2) \end{matrix} \quad \Rightarrow \quad X \sim \mathcal{N}(\mu, \sigma^2) \quad \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

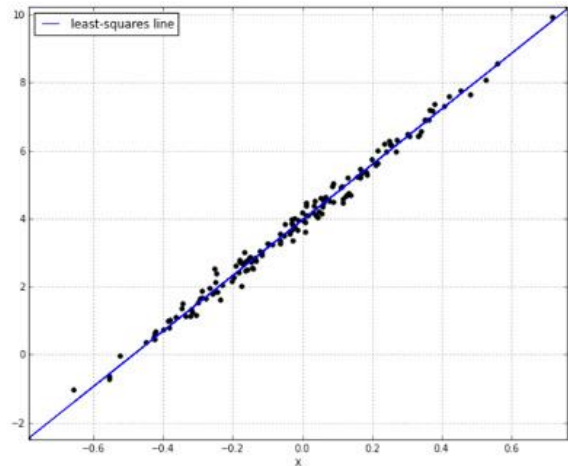
$$y_i \sim \mathcal{N}(\theta^T x_i, \sigma^2)$$

$p(y_i = y_i | x_i, \theta)$

likelihood = $p(\underbrace{y_1}_{c_1} | x_1, \theta) * p(y_2 | x_2, \theta) \dots * p(y_n | x_n, \theta)$

$$\prod_{i=1}^n \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(y_i - \theta^T x_i)^2}{2\sigma^2}}$$

$$\sigma^2 2\pi^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum_1^n (y_i - \theta^T x_i)^2}$$

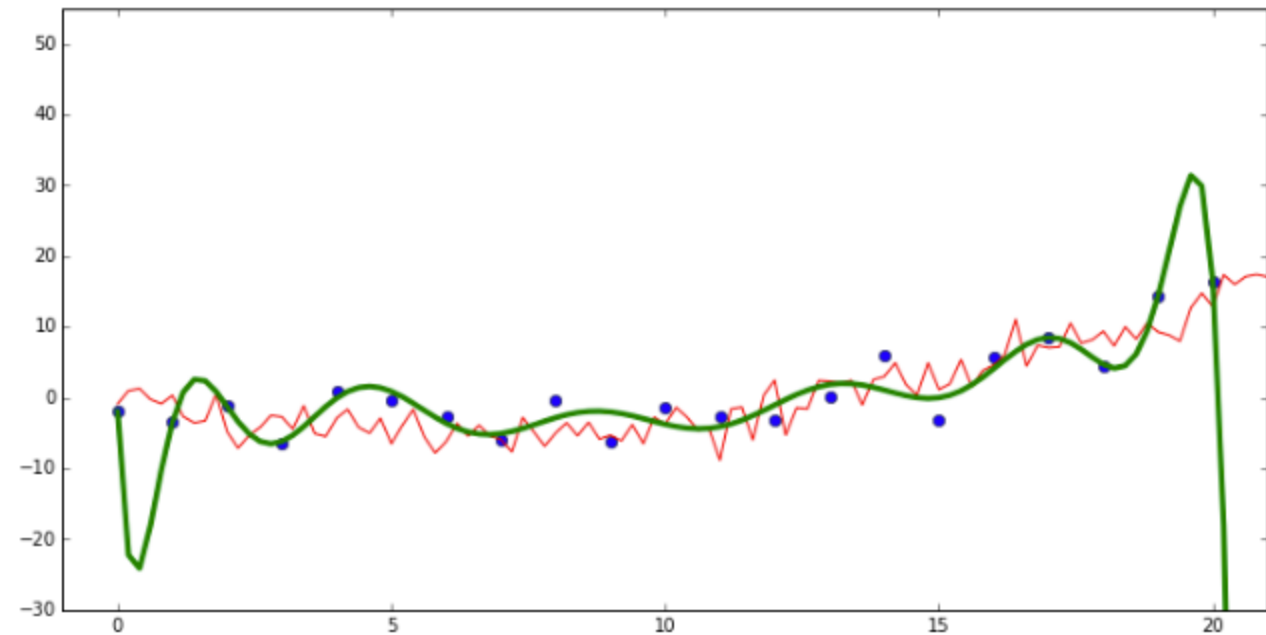


$(x_1, c_1) (x_2, c_2) \dots (x_n, c_n)$
 $p(c = a | x)$

$$\arg \min_{\theta} \sum_{i=1}^n (y_i - \hat{y}_i)^2$$

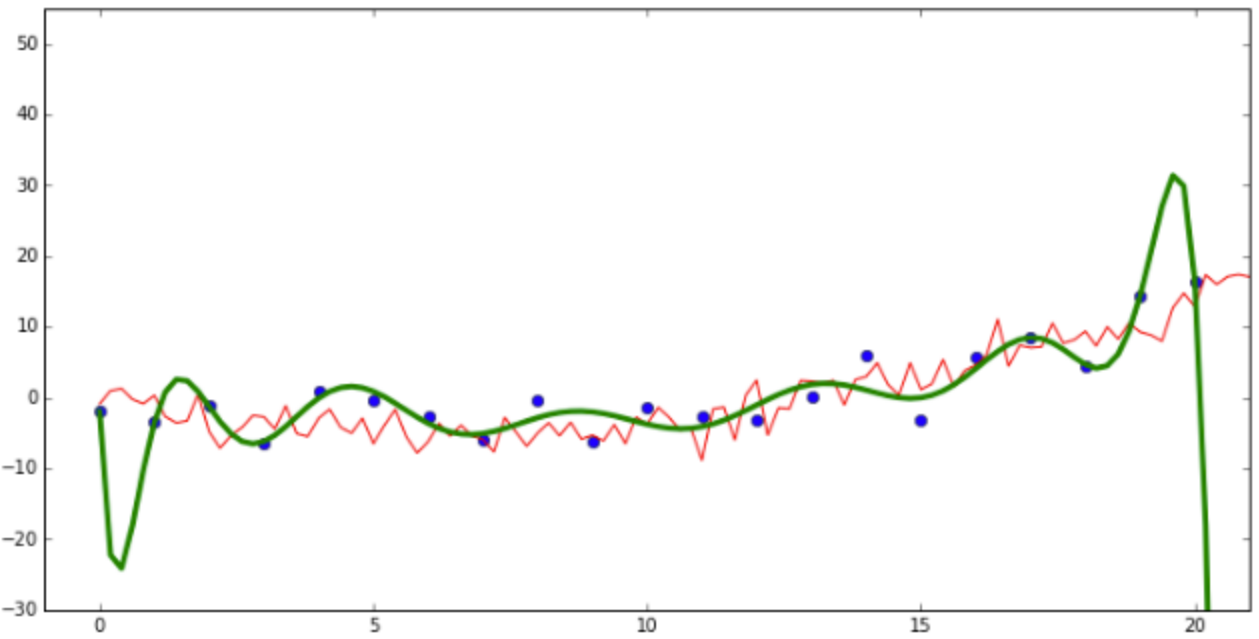
Minimizing the residual sum of squares is equivalent to maximizing the likelihood of the data

$y = ax + bx^2$ Actual



14 degree poly
(hypothesis set)

$$y = \underline{ax} + \underline{bx^2} \quad \text{Actual}$$



14 degree poly
(hypothesis set)

$$y = ax + bx^2 + \underline{\epsilon}$$

$$\epsilon \sim N(0, \tau)$$

$$X \sim N(0, \sigma^2)$$

$$Y = X + c$$

$$Y \sim N(c, \sigma^2)$$

$$P(y_1 | \mu, \sigma^2) \cdot P(y_2 | \mu, \sigma^2) \cdots$$

$$\arg \min_{\theta} \sum_1^n (y_i - \hat{y}_i)^2 + \lambda \sum_1^p \theta^2$$

Ridge (L2) Regression

$$\text{posterior} = \frac{\text{likelihood} \times \text{prior}}{\text{marginal likelihood}}$$

$$\arg \min_{\theta} \sum_1^n (y_i - \hat{y}_i)^2 + \lambda \sum_1^p \theta^2$$

$$\text{posterior} \propto \text{likelihood} \times \text{prior}$$

$$\theta \sim \mathcal{N}(0, \tau^2)$$

$$\sigma^2 2\pi^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum_1^n (y_i - \hat{y}_i)^2} \times \tau^2 2\pi^{-\frac{p}{2}} e^{-\frac{1}{2\tau^2} \sum_1^p \theta^2}$$

$$e^{-\frac{1}{2\sigma^2} \sum_1^n (y_i - \hat{y}_i)^2 - \frac{1}{2\tau^2} \sum_1^p \theta^2} \times \sigma^2 2\pi^{-\frac{n}{2}} \times \tau^2 2\pi^{-\frac{p}{2}}$$

$$\arg \max_{\theta} e^{-\frac{1}{2\sigma^2} \sum_1^n (y_i - \hat{y}_i)^2 - \frac{1}{2\tau^2} \sum_1^p \theta^2}$$

$$\arg \max_{\theta} -\frac{1}{2\sigma^2} \sum_1^n (y_i - \hat{y}_i)^2 - \frac{1}{2\tau^2} \sum_1^p \theta^2$$

$$\arg \max_{\theta} -1(\sum_1^n (y_i - \hat{y}_i)^2 + \frac{\sigma^2}{\tau^2} \sum_1^p \theta^2)$$

$$\arg \min_{\theta} \sum_1^n (y_i - \hat{y}_i)^2 + \frac{\sigma^2}{\tau^2} \sum_1^p \theta^2$$

$$\text{posterior} = \frac{\text{likelihood} \times \text{prior}}{\text{marginal likelihood}}$$

$$\arg \min_{\theta} \sum_1^n (y_i - \hat{y}_i)^2 + \lambda \sum_1^p \theta^2$$

$$\text{posterior} \propto \text{likelihood} \times \text{prior}$$

$$\sigma^2 2\pi^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum_1^n (y_i - \hat{y}_i)^2} \times \tau^2 2\pi^{-\frac{p}{2}} e^{-\frac{1}{2\tau^2} \sum_1^p \theta^2}$$

$$e^{-\frac{1}{2\sigma^2} \sum_1^n (y_i - \hat{y}_i)^2 - \frac{1}{2\tau^2} \sum_1^p \theta^2} \times \sigma^2 2\pi^{-\frac{n}{2}} \times \tau^2 2\pi^{-\frac{p}{2}}$$

$$\arg \max_{\theta} e^{-\frac{1}{2\sigma^2} \sum_1^n (y_i - \hat{y}_i)^2 - \frac{1}{2\tau^2} \sum_1^p \theta^2}$$

$$\arg \max_{\theta} -\frac{1}{2\sigma^2} \sum_1^n (y_i - \hat{y}_i)^2 - \frac{1}{2\tau^2} \sum_1^p \theta^2$$

$$\arg \max_{\theta} -1(\sum_1^n (y_i - \hat{y}_i)^2 + \frac{\sigma^2}{\tau^2} \sum_1^p \theta^2)$$

$$\arg \min_{\theta} \sum_1^n (y_i - \hat{y}_i)^2 + \frac{\sigma^2}{\tau^2} \sum_1^p \theta^2$$

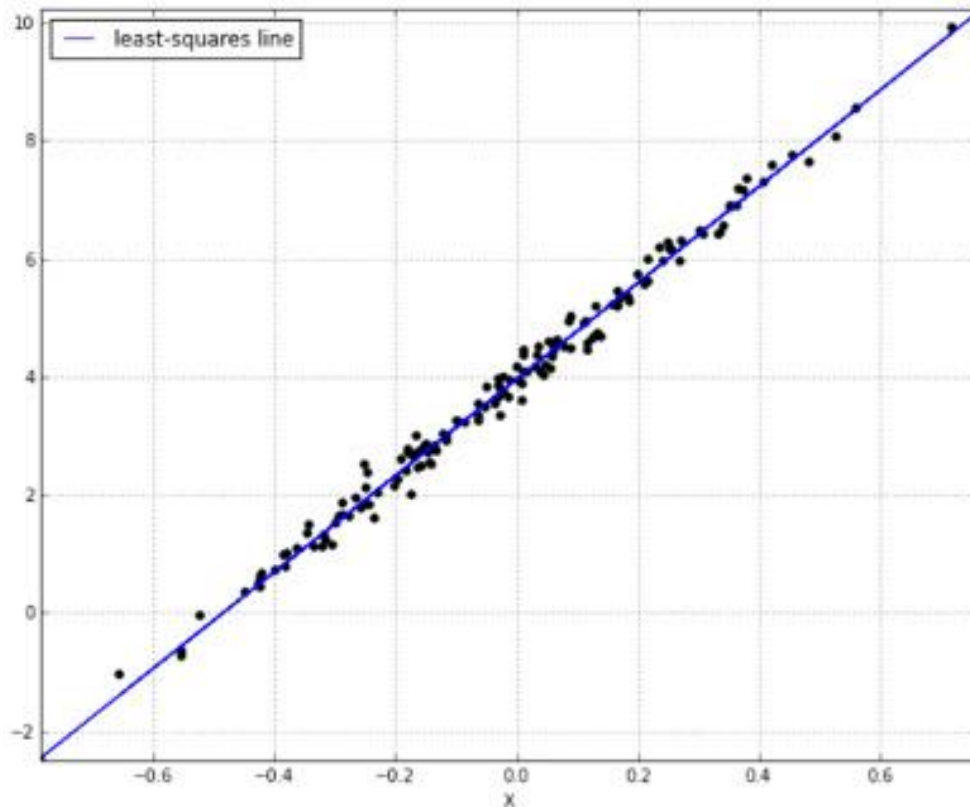
Ridge regression = MAP estimate with a zero-mean Gaussian prior, with λ proportional to τ^2 .

- Lower variance on the prior → Higher λ value in the ridge regression solution.

$$y_i = \theta^T x_i + \epsilon_i$$

\hat{y}_i

$$\arg \min_{\theta} \sum_{i=1}^n (y_i - \hat{y}_i)^2$$



$$y = 4 + 8x + \text{noise}$$

Linear Regression Param Estimation - A Probabilistic Perspective

$$y_i = \theta^T x_i + \epsilon_i \qquad \epsilon_i \sim \mathcal{N}(0, \sigma^2)$$

$$\begin{matrix} X = \mu + Y \\ Y \sim \mathcal{N}(0, \sigma^2) \end{matrix} \quad \Rightarrow \quad X \sim \mathcal{N}(\mu, \sigma^2) \quad \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

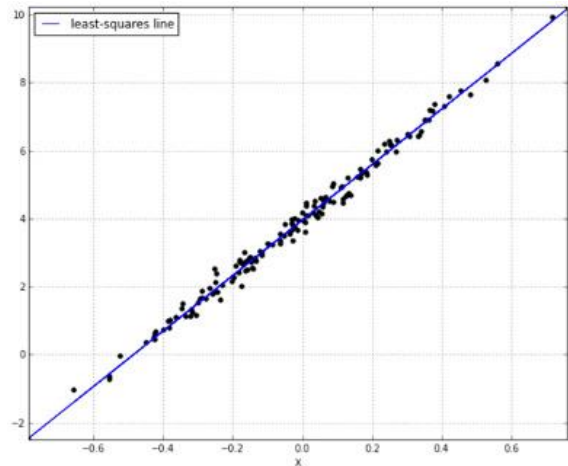
$$y_i \sim \mathcal{N}(\theta^T x_i, \sigma^2)$$

$p(y_i = y_i | x_i, \theta)$

likelihood = $p(\underbrace{y_1}_{c_1} | x_1, \theta) * p(y_2 | x_2, \theta) \dots * p(y_n | x_n, \theta)$

$$\prod_{i=1}^n \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(y_i - \theta^T x_i)^2}{2\sigma^2}}$$

$$\sigma^2 2\pi^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum_1^n (y_i - \theta^T x_i)^2}$$



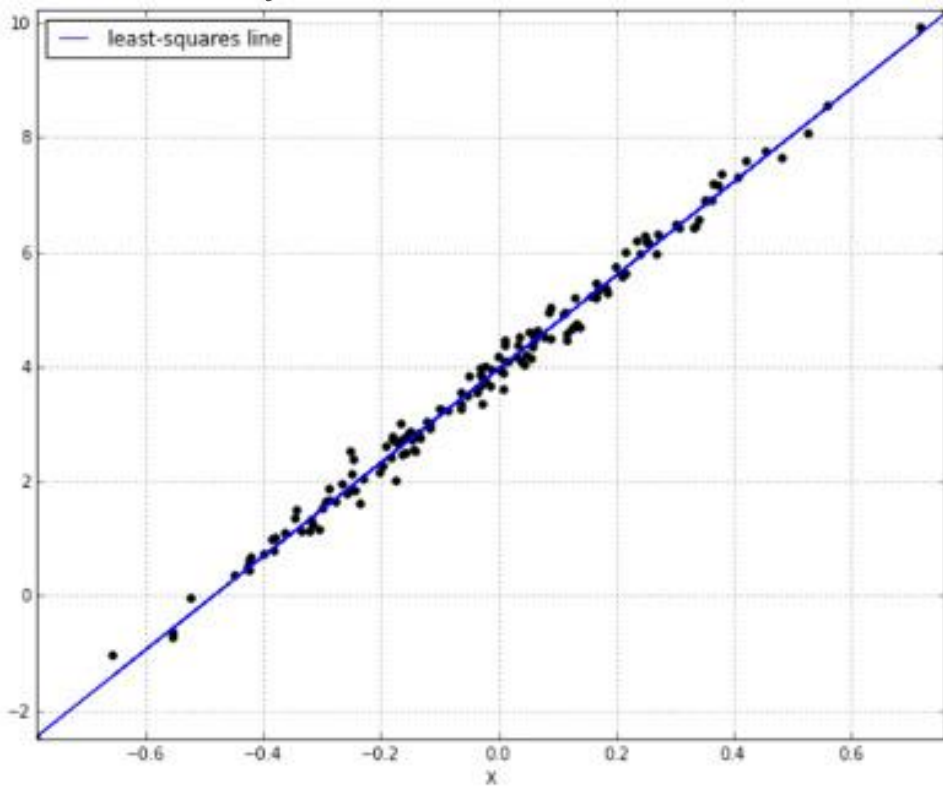
$(x_1, c_1) (x_1, c_2) \dots (x_n, c_n)$

$p(c = a | x)$

$$\arg \min_{\theta} \sum_{i=1}^n (y_i - \hat{y}_i)^2$$

Minimizing the residual sum of squares is equivalent to maximizing the likelihood of the data

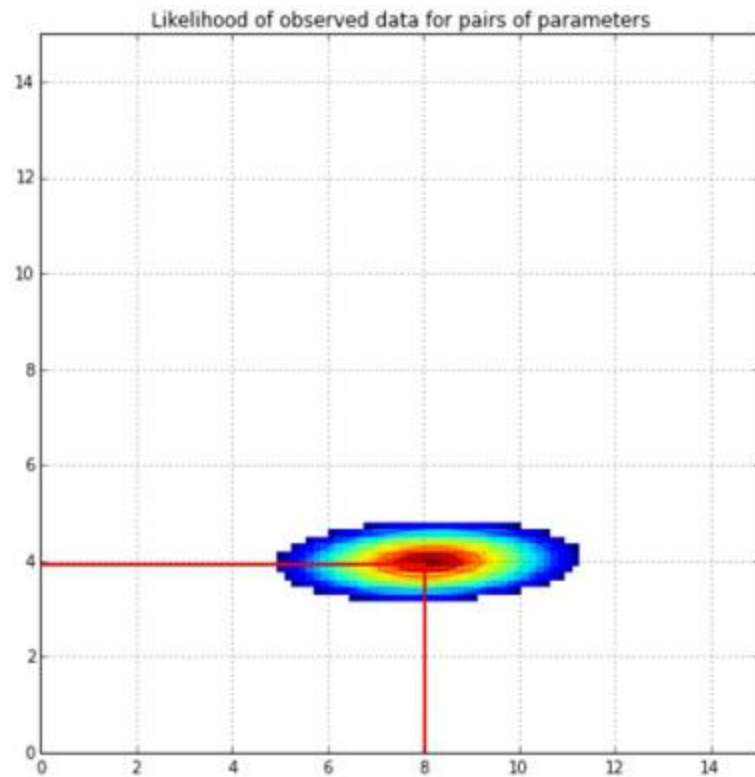
$$y = 4 + 8x + \text{noise}$$



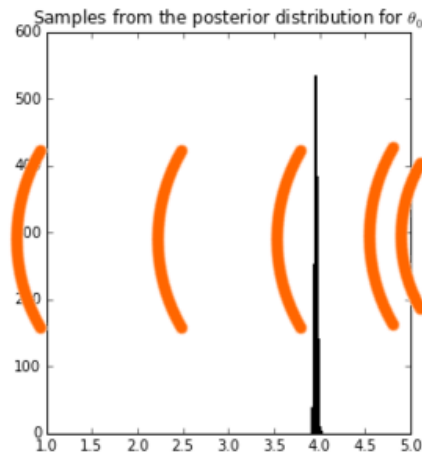
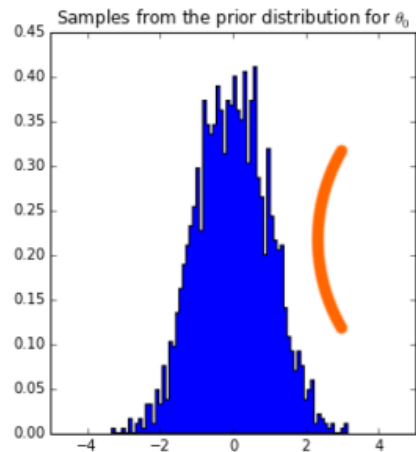
$x_i, y_i, \sigma^2, \mu, \theta$
 $\sim \mathcal{N}(y_i, \sigma^2)$
 θ_i

$$\mathcal{L}(\theta) = \sigma^2 2\pi^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum_1^n (y_i - \theta^T x_i)^2}$$

θ_2



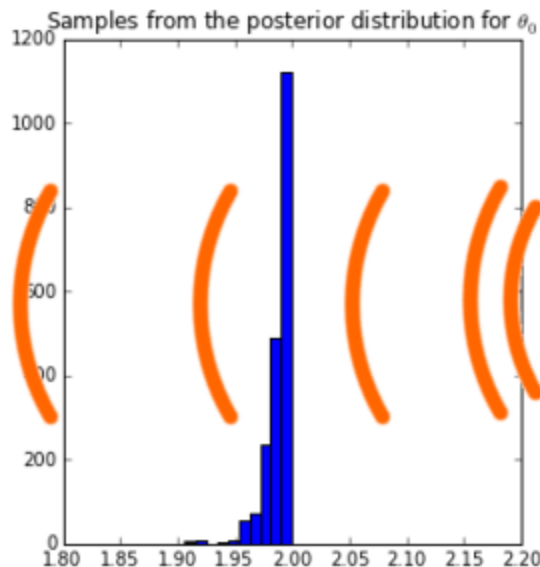
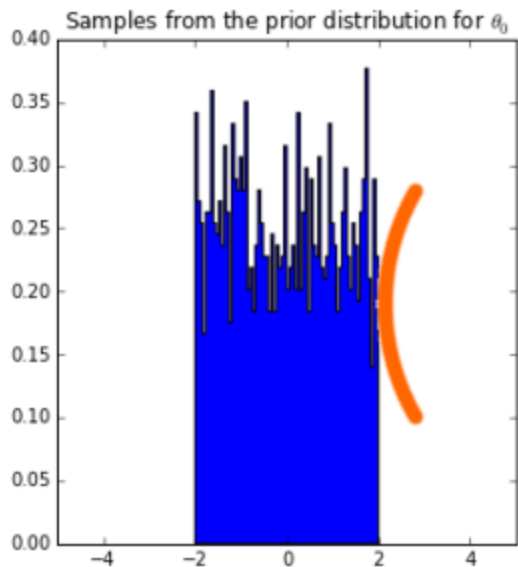
Going Bayesian – Introduce parameter priors



data is like a magnet that
attracts probability mass

$$\begin{aligned}\hat{\Theta}_{MAP} &= \operatorname{argmax}_{\Theta} \operatorname{prob}(\Theta | \mathcal{X}) \\ &= \operatorname{argmax}_{\Theta} \frac{\operatorname{prob}(\mathcal{X} | \Theta) \cdot \operatorname{prob}(\Theta)}{\operatorname{prob}(\mathcal{X})} \\ &= \operatorname{argmax}_{\Theta} \operatorname{prob}(\mathcal{X} | \Theta) \cdot \operatorname{prob}(\Theta) \\ &= \operatorname{argmax}_{\Theta} \prod_{\mathbf{x}_i \in \mathcal{X}} \operatorname{prob}(\mathbf{x}_i | \Theta) \cdot \operatorname{prob}(\Theta)\end{aligned}$$

Choosing a sensible prior is important !

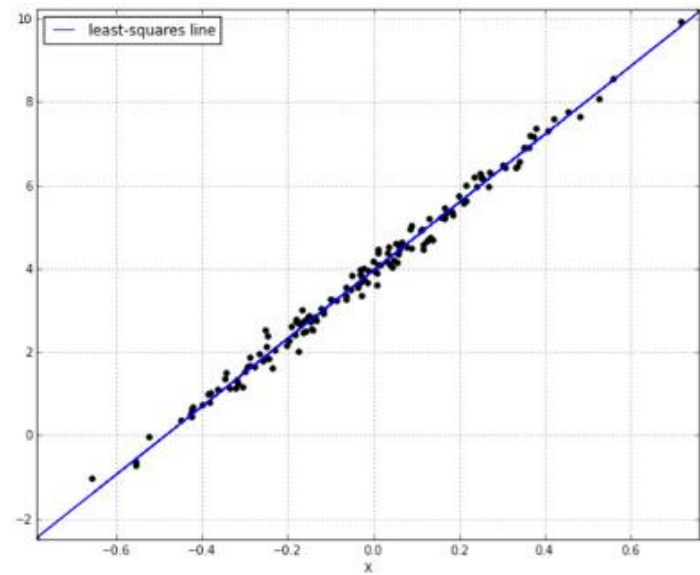


$$\begin{aligned}
 \hat{\Theta}_{MAP} &= \operatorname{argmax}_{\Theta} \operatorname{prob}(\Theta | \mathcal{X}) \\
 &= \operatorname{argmax}_{\Theta} \frac{\operatorname{prob}(\mathcal{X} | \Theta) \cdot \operatorname{prob}(\Theta)}{\operatorname{prob}(\mathcal{X})} \\
 &= \operatorname{argmax}_{\Theta} \operatorname{prob}(\mathcal{X} | \Theta) \cdot \operatorname{prob}(\Theta) \\
 &= \operatorname{argmax}_{\Theta} \prod_{\mathbf{x}_i \in \mathcal{X}} \operatorname{prob}(\mathbf{x}_i | \Theta) \cdot \operatorname{prob}(\Theta)
 \end{aligned}$$

$$\hat{\mathbf{y}}^* = [\text{MAP estimate of } \theta_0] + [\text{MAP estimate of } \theta_1] \mathbf{x}^*$$

Plug in MAP estimates ?

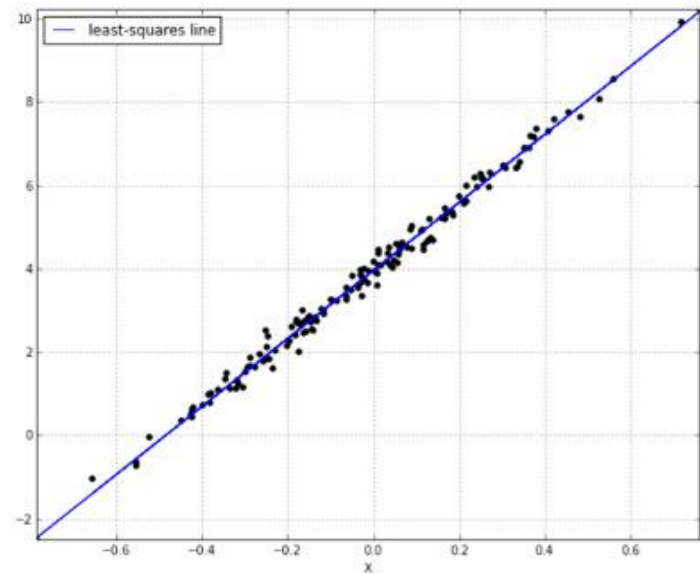
$$\begin{aligned}\hat{\Theta}_{MAP} &= \operatorname{argmax}_{\Theta} \operatorname{prob}(\Theta|\mathcal{X}) \\ &= \operatorname{argmax}_{\Theta} \frac{\operatorname{prob}(\mathcal{X}|\Theta) \cdot \operatorname{prob}(\Theta)}{\operatorname{prob}(\mathcal{X})} \\ &= \operatorname{argmax}_{\Theta} \operatorname{prob}(\mathcal{X}|\Theta) \cdot \operatorname{prob}(\Theta) \\ &= \operatorname{argmax}_{\Theta} \prod_{\mathbf{x}_i \in \mathcal{X}} \operatorname{prob}(\mathbf{x}_i|\Theta) \cdot \operatorname{prob}(\Theta)\end{aligned}$$



$$\hat{y}^* = [\text{MAP estimate of } \theta_0] + [\text{MAP estimate of } \theta_1]x^*$$

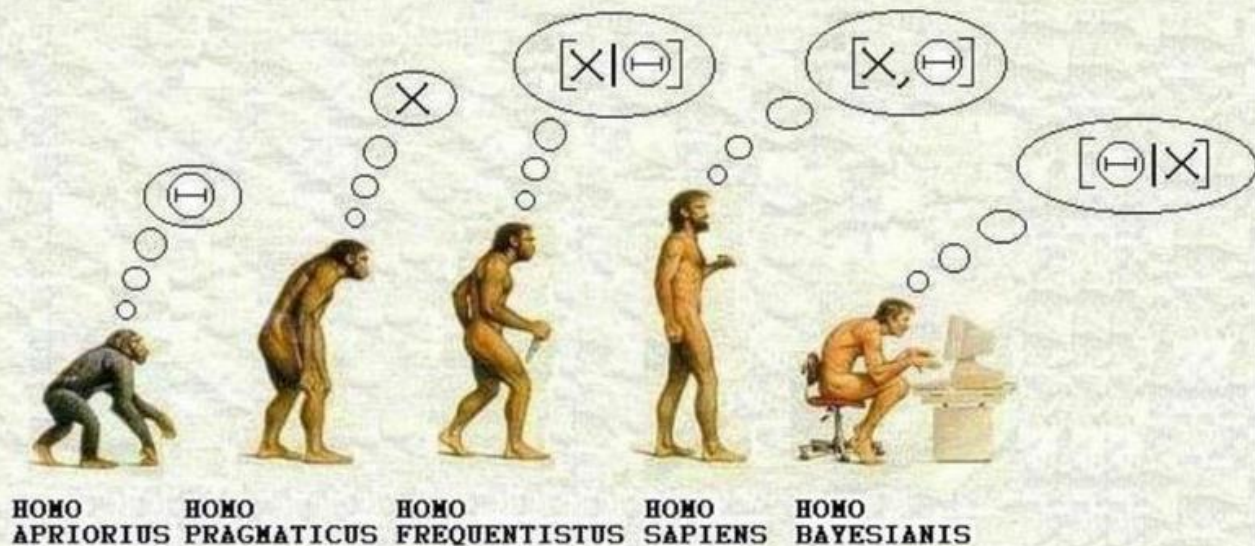
Plug in MAP estimates ?

$$\begin{aligned}\hat{\Theta}_{MAP} &= \operatorname{argmax}_{\Theta} \operatorname{prob}(\Theta|\mathcal{X}) \\ &= \operatorname{argmax}_{\Theta} \frac{\operatorname{prob}(\mathcal{X}|\Theta) \cdot \operatorname{prob}(\Theta)}{\operatorname{prob}(\mathcal{X})} \\ &= \operatorname{argmax}_{\Theta} \operatorname{prob}(\mathcal{X}|\Theta) \cdot \operatorname{prob}(\Theta) \\ &= \operatorname{argmax}_{\Theta} \prod_{\mathbf{x}_i \in \mathcal{X}} \operatorname{prob}(\mathbf{x}_i|\Theta) \cdot \operatorname{prob}(\Theta)\end{aligned}$$



$$\hat{y}^* = [\text{MAP estimate of } \theta_0] + [\text{MAP estimate of } \theta_1]x^*$$

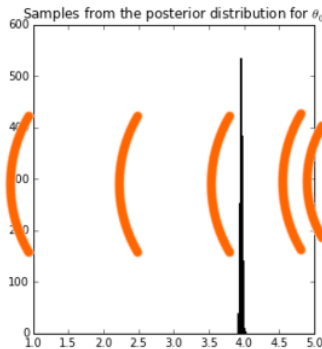
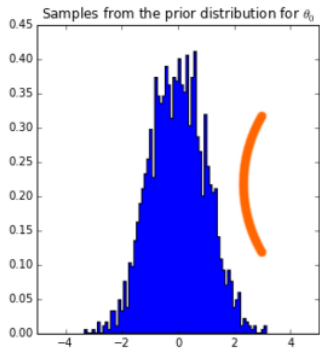
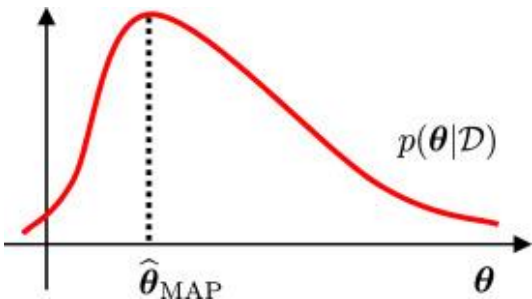
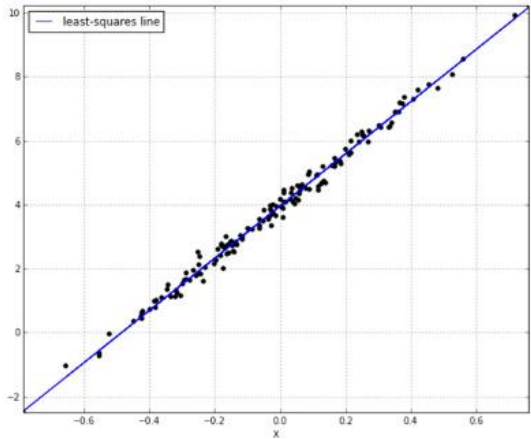
(YET ANOTHER) HISTORY OF LIFE AS WE KNOW IT...



Credit:unknown

Go Full Bayesian !

$$y_i = \theta^T x_i + \epsilon_i$$

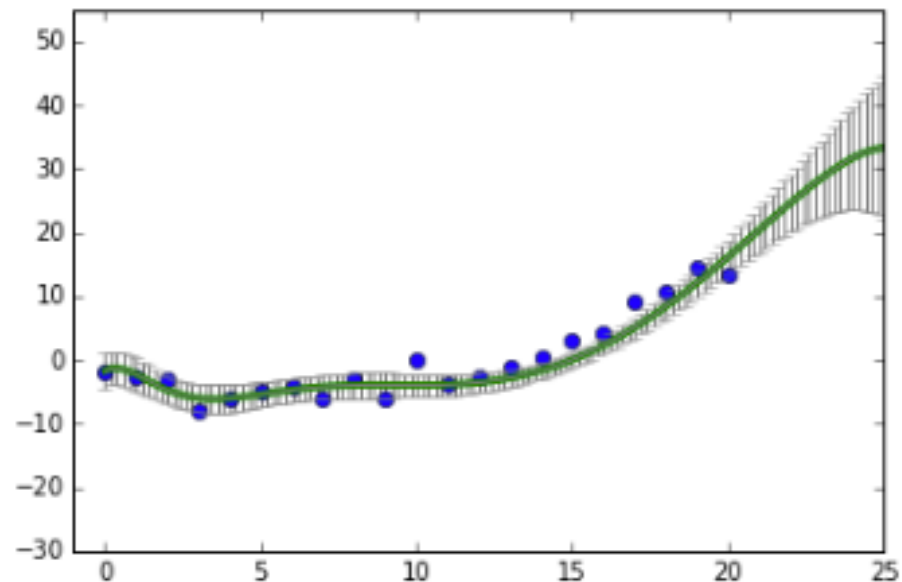
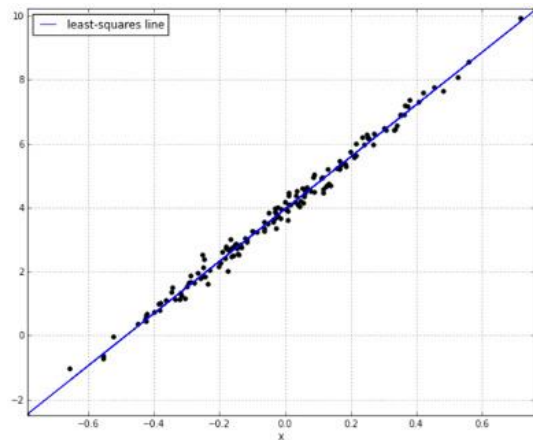


training data

$$p(y^*|x^*, X, y) = \int_{\theta} p(y^*|x^*, \theta) \underbrace{p(\theta|X, y)}_{\text{training data}} d\theta$$

We get a *probability distribution* for the outcome y^* at each x^*

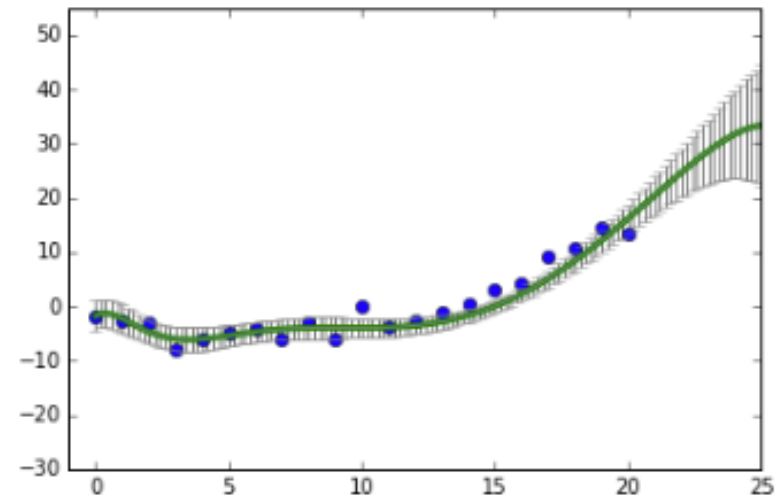
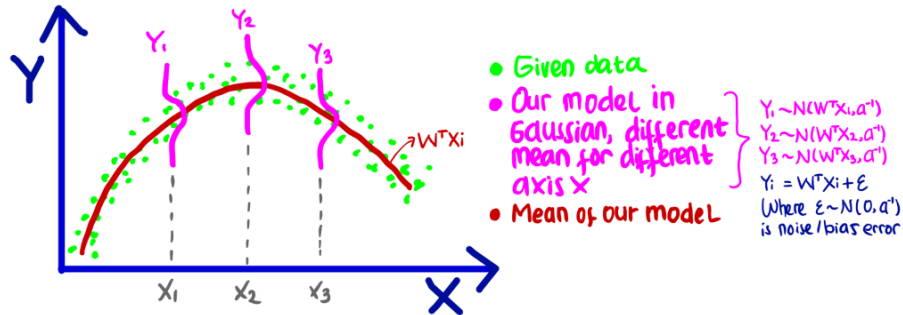
Go Full Bayesian !



$$p(y^*|x^*, X, y) = \int_{\theta} p(y^*|x^*, \theta) p(\theta|X, y) d\theta$$

Retains information about the level of uncertainty around each prediction

Bayesian Linear Regression !



$$p(y^*|x^*, X, y) = \int_{\theta} p(y^*|x^*, \theta) p(\theta|X, y) d\theta$$

Retains information about the level of uncertainty around each prediction

Regression Comparison

