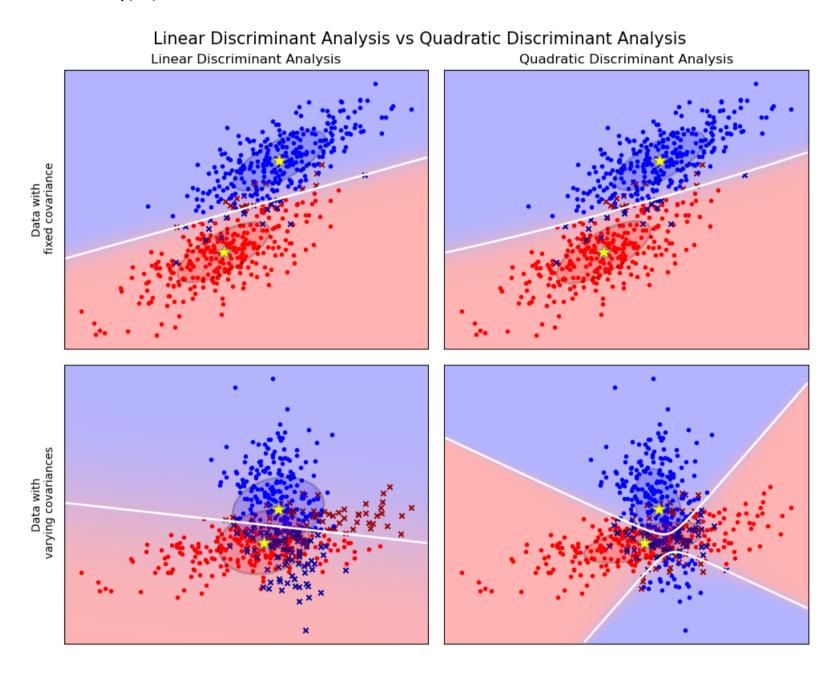
1.2. Linear and Quadratic Discriminant Analysis

Linear Discriminant Analysis (<u>LinearDiscriminantAnalysis</u>) and Quadratic Discriminant Analysis (<u>QuadraticDiscriminantAnalysis</u>) are two classic classifiers, with, as their names suggest, a linear and a quadratic decision surface, respectively.

These classifiers are attractive because they have closed-form solutions that can be easily computed, are inherently multiclass, have proven to work well in practice, and have no hyperparameters to tune.



The plot shows decision boundaries for Linear Discriminant Analysis and Quadratic Discriminant Analysis. The bottom row demonstrates that Linear Discriminant Analysis can only learn linear boundaries, while Quadratic Discriminant Analysis can learn quadratic boundaries and is therefore more flexible.

Examples:

Linear and Quadratic Discriminant Analysis with covariance ellipsoid: Comparison of LDA and QDA on synthetic data.

1.2.1. Dimensionality reduction using Linear Discriminant Analysis

<u>LinearDiscriminantAnalysis</u> can be used to perform supervised dimensionality reduction, by projecting the input data to a linear subspace consisting of the directions which maximize the separation between classes (in a precise sense discussed in the mathematics section below). The dimension of the output is necessarily less than the number of classes, so this is in general a rather strong dimensionality reduction, and only makes sense in a multiclass setting.

This is implemented in the transform method. The desired dimensionality can be set using the n_components parameter. This parameter has no influence on the fit and predict methods.

Examples:

Comparison of LDA and PCA 2D projection of Iris dataset: Comparison of LDA and PCA for dimensionality reduction of the Iris dataset

1.2.2. Mathematical formulation of the LDA and QDA classifiers

Both LDA and QDA can be derived from simple probabilistic models which model the class conditional distribution of the data P(X|y=k) for each class k. Predictions can then be obtained by using Bayes' rule, for each training sample $x \in \mathcal{R}^d$:

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https://scikit-learn.org/stable/modules/lda_qda.html

$$P(y=k|x) = rac{P(x|y=k)P(y=k)}{P(x)} = rac{P(x|y=k)P(y=k)}{\sum_l P(x|y=l) \cdot P(y=l)}$$

and we select the class k which maximizes this posterior probability.

More specifically, for linear and quadratic discriminant analysis, P(x|y) is modeled as a multivariate Gaussian distribution with density:

$$P(x|y=k) = rac{1}{(2\pi)^{d/2}|\Sigma_k|^{1/2}} \mathrm{exp}\left(-rac{1}{2}(x-\mu_k)^t \Sigma_k^{-1}(x-\mu_k)
ight)$$

where d is the number of features.

1.2.2.1. QDA

According to the model above, the log of the posterior is:

$$egin{split} \log P(y=k|x) &= \log P(x|y=k) + \log P(y=k) + Cst \ &= -rac{1}{2} \log |\Sigma_k| - rac{1}{2} (x-\mu_k)^t \Sigma_k^{-1} (x-\mu_k) + \log P(y=k) + Cst, \end{split}$$

where the constant term Cst corresponds to the denominator P(x), in addition to other constant terms from the Gaussian. The predicted class is the one that maximises this log-posterior.

Note: Relation with Gaussian Naive Bayes

If in the QDA model one assumes that the covariance matrices are diagonal, then the inputs are assumed to be conditionally independent in each class, and the resulting classifier is equivalent to the Gaussian Naive Bayes classifier naive_bayes.GaussianNB.

1.2.2.2. LDA

LDA is a special case of QDA, where the Gaussians for each class are assumed to share the same covariance matrix: $\Sigma_k = \Sigma$ for all k. This reduces the log posterior to:

$$\log P(y=k|x) = -rac{1}{2}(x-\mu_k)^t\Sigma^{-1}(x-\mu_k) + \log P(y=k) + Cst.$$

The term $(x - \mu_k)^t \Sigma^{-1}(x - \mu_k)$ corresponds to the <u>Mahalanobis Distance</u> between the sample x and the mean μ_k . The Mahalanobis distance tells how close x is from μ_k , while also accounting for the variance of each feature. We can thus interpret LDA as assigning x to the class whose mean is the closest in terms of Mahalanobis distance, while also accounting for the class prior probabilities.

The log-posterior of LDA can also be written [3] as:

$$\log P(y=k|x) = \omega_k^t x + \omega_{k0} + Cst.$$

where $\omega_k = \Sigma^{-1}\mu_k$ and $\omega_{k0} = -\frac{1}{2}\mu_k^t\Sigma^{-1}\mu_k + \log P(y=k)$. These quantities correspond to the coef_ and intercept_ attributes, respectively.

From the above formula, it is clear that LDA has a linear decision surface. In the case of QDA, there are no assumptions on the covariance matrices Σ_k of the Gaussians, leading to quadratic decision surfaces. See [1] for more details.

1.2.3. Mathematical formulation of LDA dimensionality reduction

First note that the K means μ_k are vectors in \mathbb{R}^d , and they lie in an affine subspace H of dimension at most K-1 (2 points lie on a line, 3 points lie on a plane, etc).

As mentioned above, we can interpret LDA as assigning x to the class whose mean μ_k is the closest in terms of Mahalanobis distance, while also accounting for the class prior probabilities. Alternatively, LDA is equivalent to first *sphering* the data so that the covariance matrix is the identity, and then assigning x to the closest mean in terms of Euclidean distance (still accounting for the class priors).

Computing Euclidean distances in this d-dimensional space is equivalent to first projecting the data points into H, and computing the distances there (since the other dimensions will contribute equally to each class in terms of distance). In other words, if x is closest to μ_k in the original space, it will also be the case in H. This shows that, implicit in the LDA classifier, there is a dimensionality reduction by linear projection onto a K-1 dimensional space.

We can reduce the dimension even more, to a chosen L, by projecting onto the linear subspace H_L which maximizes the variance of the μ_k^* after projection (in effect, we are doing a form of PCA for the transformed class means μ_k^*). This L corresponds to the n_components parameter used in the <u>transform</u> method. See [1] for more details.

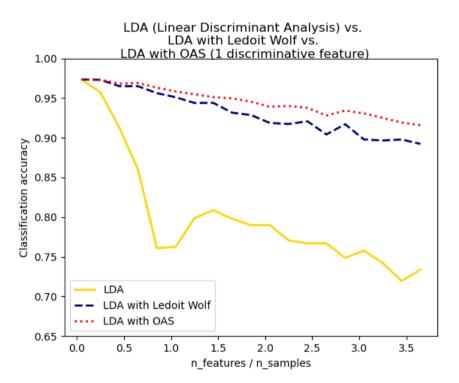
1.2.4. Shrinkage and Covariance Estimator

Shrinkage is a form of regularization used to improve the estimation of covariance matrices in situations where the number of training samples is small compared to the number of features. In this scenario, the empirical sample covariance is a poor estimator, and shrinkage helps improving the generalization performance of the classifier. Shrinkage LDA can be used by setting the shrinkage parameter of the LinearDiscriminantAnalysis class to 'auto'. This automatically determines the optimal shrinkage parameter in an analytic way following the lemma introduced by Ledoit and Wolf [2]. Note that currently shrinkage only works when setting the solver parameter to 'lsqr' or 'eigen'.

The shrinkage parameter can also be manually set between 0 and 1. In particular, a value of 0 corresponds to no shrinkage (which means the empirical covariance matrix will be used) and a value of 1 corresponds to complete shrinkage (which means that the diagonal matrix of variances will be used as an estimate for the covariance matrix). Setting this parameter to a value between these two extrema will estimate a shrunk version of the covariance matrix.

The shrunk Ledoit and Wolf estimator of covariance may not always be the best choice. For example if the distribution of the data is normally distributed, the Oracle Shrinkage Approximating estimator sklearn.covariance.OAS yields a smaller Mean Squared Error than the one given by Ledoit and Wolf's formula used with shrinkage="auto". In LDA, the data are assumed to be gaussian conditionally to the class. If these assumptions hold, using LDA with the OAS estimator of covariance will yield a better classification accuracy than if Ledoit and Wolf or the empirical covariance estimator is used.

The covariance estimator can be chosen using with the covariance_estimator parameter of the discriminant_analysis.LinearDiscriminantAnalysis class. A covariance estimator should have a fit method and a covariance_ attribute like all covariance estimators in the sklearn.covariance module.



Examples:

Normal, Ledoit-Wolf and OAS Linear Discriminant Analysis for classification: Comparison of LDA classifiers with Empirical, Ledoit Wolf and OAS covariance estimator.

1.2.5. Estimation algorithms

Using LDA and QDA requires computing the log-posterior which depends on the class priors P(y=k), the class means μ_k , and the covariance matrices.

The 'svd' solver is the default solver used for <u>LinearDiscriminantAnalysis</u>, and it is the only available solver for <u>QuadraticDiscriminantAnalysis</u>. It can perform both classification and transform (for LDA). As it does not rely on the calculation of the covariance matrix, the 'svd' solver may be preferable in situations where the number of features is large. The 'svd' solver cannot be used with shrinkage. For QDA, the use of the SVD solver relies on the fact that the covariance matrix Σ_k is, by definition, equal to $\frac{1}{n-1}X_k^tX_k=\frac{1}{n-1}VS^2V^t$ where V comes from the SVD of the (centered) matrix: $X_k=USV^t$. It turns out that we can compute the log-posterior above without having to explicitly compute Σ : computing S and V via the SVD of X is enough. For LDA, two SVDs are computed: the SVD of the centered input matrix X and the SVD of the class-wise mean vectors.

The 'lsqr' solver is an efficient algorithm that only works for classification. It needs to explicitly compute the covariance matrix Σ , and supports shrinkage and custom covariance estimators. This solver computes the coefficients $\omega_k = \Sigma^{-1}\mu_k$ by solving for $\Sigma\omega = \mu_k$, thus avoiding the explicit computation of the inverse Σ^{-1} .

The 'eigen' solver is based on the optimization of the between class scatter to within class scatter ratio. It can be used for both classification and transform, and it supports shrinkage. However, the 'eigen' solver needs to compute the covariance matrix, so it might not be suitable for situations with a high number of features.

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[1] (<u>1,2</u>)

"The Elements of Statistical Learning", Hastie T., Tibshirani R., Friedman J., Section 4.3, p.106-119, 2008.

[<u>2</u>]

Ledoit O, Wolf M. Honey, I Shrunk the Sample Covariance Matrix. The Journal of Portfolio Management 30(4), 110-119, 2004.

[<u>3</u>]

R. O. Duda, P. E. Hart, D. G. Stork. Pattern Classification (Second Edition), section 2.6.2.

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