## The Impact of Baseline Incidence Rates on Burden of Disease Assessment of Air Pollution and Onset Childhood Asthma: Analysis of Data from the Contiguous United States

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**Abbreviations**

**AC:** Attributable number of cases

**ACBS:** Asthma Call Back Survey

**AF:** Attributable fraction of cases

**BRFSS:** Behavioral Risk Factor Surveillance System

**CDC:** Center for Disease Control and Prevention

**CRF**: Concentration-Response Function

**D.C.:** District of Columbia

**EPA:** United States Environmental Protection Agency

**U.S.:** United States

**LUR:** Land-use regression

**NHGIS:** National Historical Geographic Information System

**PAF:** Population attributable fraction

**IR:** Incidence rate

**PR:** Prevalence rate

**ppb**: Parts per billion

**TRAP:** Traffic-related air pollution

**Abstract**

**Background**

Burden of disease (BoD) assessments typically rely on national-level incidence rates of the health outcomes of interest. The impact of using a flat national-level incidence rate, versus a more granular spatially varying rate, remains unknown and understudied in the literature. At the same time, there have been more publications estimating burden of onset childhood asthma attributable to air pollution exposures, as emerging evidence demonstrates that traffic-related air pollution (TRAP) leads to the onset of the disease. In this paper, we estimated the number of incident asthma cases among children attributable to Nitrogen Dioxide (NO2), a good marker of TRAP, in the contiguous U.S. We used both a national-level and newly generated state-specific asthma incidence rates and compared results from the two approaches.

**Methods**

We estimated the number and percentage of incident childhood asthma cases potentially attributable to NO2 using standard BoD assessment methods. We combined children (<18 years) counts and 2010 NO2 exposures at populated U.S. census blocks with a flat national-level and newly generated state-specific asthma incidence rates. We calculated incidence rates using raw data from 2006-2010 Behavioral Risk Factor Surveillance System (BRFSS) and the nested Asthma Call Back (ACS) surveys collected by the Center for Disease Control and Prevention (CDC). We sourced concentration response functions (CRF) from the latest meta-analysis on TRAP and risk of onset childhood asthma. NO2 concentrations were obtained from a previously validated land-use regression (LUR) model. We stratified BoD estimates by urban or rural status and median household income, exploring trends across 48 states and the District of Columbia (D.C.).

**Results**

The mean (min-max) NO2 concentration(s) was 13.2 (1.5-58.3) ug/m3 and was highest in census-designated urbanized areas. We were able to estimate childhood asthma incidence rates in 32 states which ranged from 4.3 (Montana) to 17.7 (D.C.) per 1,000 at-risk children. The 17 states that did not have data to calculate an incidence rate were assigned the national aggregate asthma incidence rate of 11.6 per 1,000 at-risk children. Using the state-specific incidence rates resulted in a 7.2% reduction (absolute number=10,192) in the number of incident asthma cases attributable to NO2, which amounted to a 1.7% reduction in the attributable fraction at the U.S. level. The change in the attributable number of cases across the states was more prominent and ranged from -50% (Iowa, Louisiana, Montana) to +33.6% (Texas). The State of California had the largest absolute decrease, with 6,2000 fewer attributable cases, while the State of Texas had the largest increase, with 3,600 additional cases, followed by New York (1,700). Stratifying our analyses by socioeconomic status and urban versus rural status produced new trends compared to what we observed in past analyses employing a national-level incidence rate.

**Conclusion**

Using state-specific versus national-level asthma incidence rates resulted in a small change in the attributable BoD at the U.S. level, but had a more prominent impact on the BoD estimates at the state level. To our knowledge, this is the first study to analyse and document the impact of using a flat versus a spatially varying asthma incidence rate in the context of air pollution BoD assessment.

**Introduction**

Burden of disease (BoD) assessment is a powerful and relatively practical method to estimate the number/percentage of premature mortality and morbidity cases that may be attributable to environmental exposures. Such estimates can indicate how many cases of premature deaths and/or disease may be prevented by eliminating or reducing the exposure of interest. In the context of air pollution exposure, BoD methods have become increasingly popular and predominantly used to assess the burden of premature mortality attributable to air pollution at the global, national, regional and local scales (Cohen et al., 2017, Cohen et al., 2005, Ostro and Organization, 2004, Lelieveld et al., 2015, Bhalla et al., 2014, Tainio, 2015, Mueller et al., 2017). The prior focus on premature mortality may reflect the availability of well-established epidemiological data that associate air pollution with premature death. It could also reflect the level of advancement in BoD, which is a relatively new practice still concerned with the most extreme outcome (death). Yet, to truly map, grasp and communicate the public health impact of air pollution exposures, extending BoD beyond premature mortality is required, especially to chronic disease. Chronic diseases are important as they have significant impacts on the quality of life of individuals and families, affect productivity at work and school, and can result in death and imply significant health care costs which may be preventable. Further, given the ubiquity of air pollution exposure, especially in urban areas, the relatively modest-sized risk estimates from epidemiology translate into a large, yet modifiable, BoD.

One chronic disease that recently received more attention in the context of air pollution is the onset of childhood asthma. Asthma is a burdensome disease that is often cited as the most impactful chronic illness of childhood (Gasana et al., 2012, National Survey of Children's Health, 2012). It is the third leading cause of hospitalization in children aged 15 and under and the leading cause of school absenteeism due to a chronic disease (American Lung Association, 2019, Hsu et al., 2016). In the U.S. alone, 6 million children had ongoing asthma in 2016 (Zahran et al., 2018). The economic burden of asthma in the U.S., including costs incurred by absenteeism and mortality, was $81.9 billion in 2013 (Nurmagambetov et al., 2018). The U.S. Center for Disease Control and Prevention (CDC) estimates the number of missed school days in a single year, 2008, reached 10.4 million for children with asthma (CDC, 2010).

There is emerging evidence that exposure to air pollution, primarily when traffic-related, is associated with the onset of children’s asthma (Khreis et al., 2017), and more recent studies confirm these associations (Rancière et al., 2016, Rice et al., 2018, Lee et al., 2018, Pennington et al., 2018). However, a limited number of studies have investigated the burden of incident childhood asthma attributable to traffic-related air pollution (TRAP). All previous BoD studies investigating this topic (Achakulwisut et al., 2019, Khreis et al., 2018c, Khreis et al., 2018b, Alotaibi et al., 2019, Anenberg et al., 2018) highlighted several data gaps that may impact final BoD estimates and potentially result in uncertainty and error. The accuracy of the BoD estimate is dependent on the accuracy of input data, namely: 1) the air pollution exposure levels and distribution, 2) the concentration-response functions, and 3) the baseline asthma incidence rates. Some studies investigated the impacts of different input datasets on final BoD estimates and found that different exposure assessment methods (e.g. dispersion versus land-use regression modeling) can result in up to a 3% difference in BoD estimates (Khreis et al., 2018a). Alotaibi et al. (2019) explored the impact of uncertainty in concentration-response functions (CRF) on final BoD estimates and found that using the most conservative end of the CRF (the lower 95% CI) can reduce the estimated burden by up to 19% when compared to the central estimate. On the other hand, using the most extreme end of the CRF (the upper 95% CI) can increase the estimated burden by up to 14% when compared to the central estimate. Finally, one study (Khreis et al., 2018a) and their follow-up work in Khreis et al. (2019) showed that using a national versus a local baseline asthma incidence rate can result in up to an 11% difference in BoD estimates; however, this analysis was limited to one medium-sized city in England.

The impact asthma baseline incidence rates have on BoD estimates has not been thoroughly studied. Childhood asthma can be challenging to diagnose and ascertain, and national-level incidence rates are likely to vary substantially, in particular among urban and rural populations. Previous literature relied on national-level asthma incidence rates, which is in line with practice by prominent institutions and studies such as the Global Burden of Disease analyses. In this paper, we explore the potential impact of using state-specific varying asthma incidence rates on the final burden of childhood asthma due to NO2 and compare the change in BoD estimates to those produced by Alotaibi et al. (2019) who used a national-level asthma incidence rate, as typically practiced (Achakulwisut et al., 2019, Khreis et al., 2018c, Perez et al., 2009, Perez et al., 2013, Khreis et al., 2018b, Anenberg et al., 2018). Using these more granular state-specific asthma incidence rates, we also explored trends in the BoD estimates by socioeconomic status and urban versus rural status to confirm (or otherwise) previous trends observed in past analysis. We selected NO2 as the exposure of interest, as more studies underline this pollutant’s CRF and it has been the most commonly used pollutant in previous epidemiological and BoD assessments (Khreis et al., 2017). this supports our aim of establishing the difference in BoD estimates from previous analyses, by *only* altering the baseline childhood asthma incidence rate and not the exposure.

**Methods**

*Study area and time point*

We analyzed data for the 48 contiguous U.S. States and the District of Columbia (D.C.) for 2010 at the census block level, the smallest geographical unit available. Population counts, urban or rural living location status and annual NO2 concentrations were all available at the census block level. However, median household income was only available at the census block group level, which is one level higher than the census block (US Census Bureau, 2010). Childhood asthma incidence rates were calculated at the state level. NO2 concentrations were not available for states or territories outside the contiguous U.S. (Alaska, Hawaii and Puerto Rico), and hence these were excluded from the analysis.

*Census data*

We included populated census blocks of the contiguous U.S. for the year 2010, as obtained from the National Historical Geographic Information System (NHGIS) website (Manson et al., 2018, US Census Bureau, 2010). Each block included information on the total population of children <18 years old, and whether the census block was designated as an urban or a rural block. Census-designated urban areas were defined by the census bureau using multiple criteria including total population thresholds, density, nonresidential urban land use (e.g. paved areas and airports), and distance to other urban developed areas (US Census Bureau, 2016). Census blocks are the basic geographical units of urban areas. Further, census-designated urban areas are classified into two subtypes; urban clusters (≥2,500 to <50,000 people) or urbanized areas (≥50,000 people). The median household income in the past 12 months using 2010 inflation adjusted dollars was divided into five categories consistent with two previous relevant publications: <$20,000, $20,000 to <$35,000, $35,000 to <$50,000, $50,000 to <$75,000 and ≥$75,000 (Clark et al., 2017, Alotaibi et al., 2019). Census blocks were assigned the same median household income of the census block group they resided within. There were 2,686 (0.04%) census blocks with missing median household income data in 2010. These census blocks were assigned a “Not defined” status in the analysis of median household income. Table 1 summarizes the geographical and demographic data across all census blocks included in this analysis.

*Table 1: Census data description, year 2010*

|  |  |  |
| --- | --- | --- |
| **Geographic characteristics** | **Total populated census blocks** | 6,182,882 |
| **Total census-designated urban areas** | 3,590,278 (58%) |
| **Demographic characteristics** | **Total population** | 306,675,006 |
| **Total population of children (birth – 18)** | 73,690,271 (24%) |
| **Mean (range) number of children in census blocks** | 12 (0-2214) |
| **Population of children by living location** | **Rural** | 13,763,183 (19%) |
| **Urban clusters (≥2,500 and <50,000 people)** | 6,994,464 (9%) |
| **Urbanized area (≥50,000 people)** | 52,932,624 (72%) |
| **Population of children by median household income** | **<$20,000** | 2,614,804 (4%) |
| **$20,000 to <$35,000** | 12,770,843 (17%) |
| **$35,000 to <$50,000** | 18,573,954 (25%) |
| **$50,000 to <$75,000** | 21,953,876 (30%) |
| **≥$75,000** | 17,763,239 (24%) |

*NO2 exposure assessment*

Annual average NO2 concentrations for each populated census block were available at the centroid location for the year 2010. Concentrations were derived from a land-use regression model (LUR) developed by Bechle et al. (2015) that incorporates spatial and temporal air pollutant data. The spatial data is derived from the U.S. Environmental Protection Agency (EPA) air quality monitoring data, satellite data and several GIS covariates including impervious surfaces, elevation, major, minor and residential roads, and distance to coast. The temporal data of the LUR model is incorporated by scaling the spatial data with the average monthly monitor readings for 11 consecutive years. The model achieves a relatively high predictive power as demonstrated using hold-out cross validation when compared to similar NO2 LUR models (Vienneau et al., 2013, Beelen et al., 2009, Hystad et al., 2011, Novotny et al., 2011) with an R2 reaching 82%. The LUR model has been used in multiple studies including Clark et al. (2017) and Alotaibi et al. (2019). A detailed description of the model can be found at Bechle et al. (2015). NO2 concentrations were converted from ppb to ug/m3through multiplying by 1.88 (WHO, 2005). Exposure data was matched with census blocks using a unique identifier for each census block as provided in the NHGIS dataset.

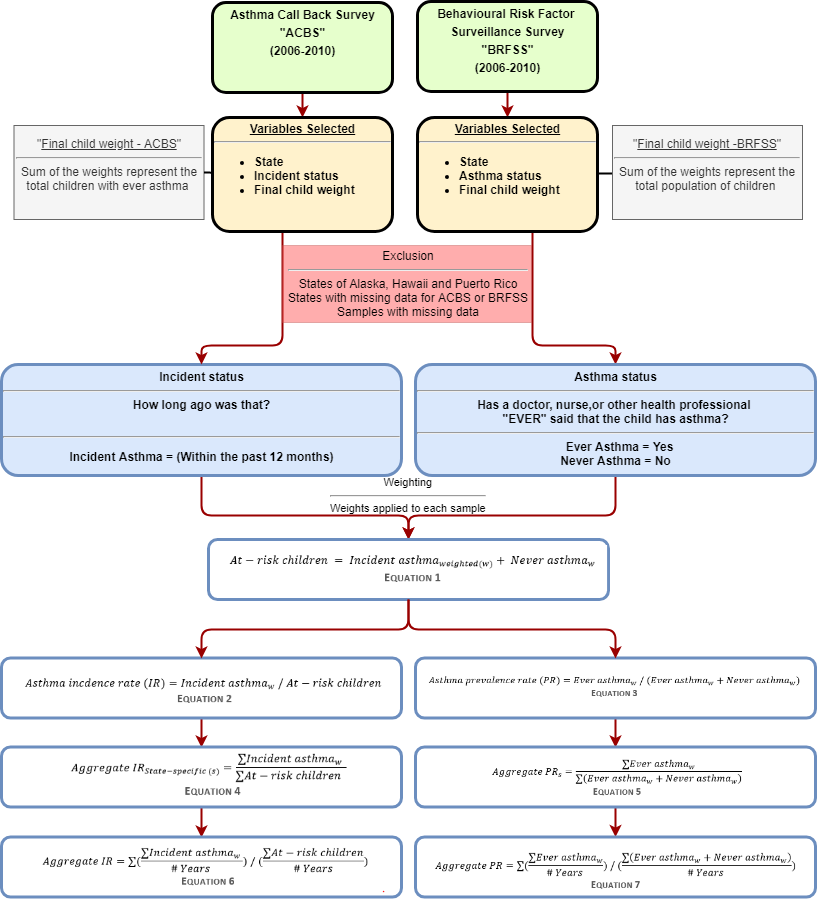
*Concentration-response functions*

We used an incident asthma CRF of 1.05 (95% CI = 1.02-1.07) per 4 ug/m3 of NO2. The CRF was obtained from a meta-analysis of 20 studies examining the association between exposure to TRAP and the risk of developing asthma among children from birth to 18 years of age (Khreis et al., 2017). These CRF represent data from the most recent analyses on TRAP and the onset of childhood asthma, and have been used in several published peer-reviewed BoD assessments (Khreis et al., 2018c, Khreis et al., 2018b, Achakulwisut et al., 2019, Alotaibi et al., 2019, Khreis, In press).

*Asthma incidence and prevalence rates*

An incidence rate (IR) is defined as the number of new cases of a disease within a specified time period among an at-risk population (Mausner and Kramer, 1985). To estimate the childhood asthma IR aggregated for the years 2006 through 2010 among U.S. states, we obtained the Behavioral Risk Factor Surveillance System (BRFSS) and Asthma Call Back Survey (ACBS) child data sets (CDC, 2011, CDC, 2009), which can be found at the CDC website <https://www.cdc.gov/brfss/>. We followed methods described by Winer et al. (2012) to estimate the asthma incidence rates, and present our steps in Figure 1. The ACBS and BRFSS define children as birth to 18 years of age people, which is in line with the meta-analysis from where we sourced the CRF (Khreis et al., 2017). The following variables were extracted from the surveys: state, asthma status question (from the BRFSS), incident status question (from the ACBS), and children sample weights from both surveys. All analysis was conducted using R statistical software (R Core Team, 2018).

*Figure 1: Childhood asthma incidence rate estimation flow chart*

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To determine the “asthma status” of children, respondents to the BRFSS were asked “Has a doctor, nurse, or other health professional EVER said that the child has asthma?”, If the answer was “Yes”, the respondent was designated as “Ever asthma”. If the answer was “No”, the respondent was designated as “Never asthma”. Respondents with children designated as “Ever asthma” were requested to participate in the ACBS follow up. To determine the “incident status” of children, respondents to the ACBS were asked: “How old was the [name of child] when a doctor or other health professional first said [he/she] had asthma? How long ago was that?” If the answer to the latter part of this question was “within the past 12 months”, the respondent was designated as an “Incident asthma”, while other responses were not relevant to the analysis described next.

Each respondent (sample) from the BRFSS and ACBS was assigned a weight to adjust for the disproportionate population sample selection as compared to the state’s overall population distribution, the variation in probability of selection, the actual response of each respondent, or nonresponse (Garbe et al., 2011, Korn and Graubard, 2011). To simplify this, the weight of each sample represents the number of children within each state, with similar characteristics (age, sex and race) to the sample. Weights were used to convert samples to population estimates of children. For example, if respondent (X) had a weight of 150, her/his response to survey questions represented answers of 150 children within their state, with similar age, sex and race characteristics. The sum of childhood weights for the BRFSS represent the total population of children within each state, while the sum of weights for the ACBS represent the total population of children with “Ever asthma” within each state.

“At-risk children” were then estimated by taking the weighted sum of respondents designated as “Incident asthma” and “Never asthma”, as shown in Equation 1.

*Equation 1*

The asthma incidence rate (IR) was the weighted “Incident asthma” divided by “At-risk children”, as shown in Equation 2.

*Equation 2*

The asthma prevalence rate (PR) was the weighted “Ever asthma” divided by the sum of weighted “Ever asthma” and weighted “Never asthma”, as shown in Equation 3.

*Equation 3*

To estimate the state-specific aggregate asthma IR and PR, we summed numerators and denominators across available years (2006-2010) for each state separately, as shown in Equation 4 and 5.

*Equation 4*

*Equation 5*

For states with no available data regarding state-specific incidence rates(n = 17) and/or state-specific prevalence rates(n = 8), we assigned them the overall aggregate IR (11.6 per 1,000 at-risk children) and aggregated PR (13.1 per 100 children) across all the available data. In order to aggregate across all the states we re-weighted the data to adjust for the number of available years across each state. For example, the state of Arizona had two years of available data (2006 and 2007). To aggregate the IR for Arizona across all available years, we summed the weighted “Incident asthma” across all the years and divided it by two (the number of years with available data for the state of Arizona). We then divided the results by the sum of “At-risk children” across all the two years, again divided by two, as shown in Equation 6.

*Equation 6*

The aggregate asthma PR across all available years was estimated as shown in Equation 7.

*Equation 7*

*Burden of disease estimation*

To estimate the BoD of incident childhood asthma attributable to NO2 exposure, we followed standard methods described in Alotaibi et al. (2019) with the following steps:

The total number of at-risk children residing in a census block was estimated for each state. This was done by subtracting the total number of children within the census block multiplied by the state-specific aggregate PR (from Equation 5 or 7) from the total number of children within the same census block, as shown in Equation 8.

*Equation 8*

We then estimated the number of childhood asthma incident cases within each census block by multiplying the state-specific aggregate asthma IR (from Equation 4 or 6) by the at-risk children in each census block (from Equation 8), as shown in Equation 9.

*Equation 9*

We then calculated the relative risk (RRdiff) for asthma onset due to the exposure difference between the estimated exposure levels from the LUR model (NO2 concentration at the centroid of each census block) and no exposure (zero concentration for NO2) at each census block, as shown in Equation 10.

*Equation 10*

Where RR is the CRF and RRunit is the exposure unit (4 ug/m3) for the CRF as extracted from Khreis et al. (2017). The population attributable fraction (PAF) was also estimated at each census block, as shown in Equation 11.

*Equation 11*

The attributable number of asthma incident cases (AC) was estimated by multiplying the PAF with the total number of asthma incident cases at each census block (from Equation 9), as shown in Equation 12.

*Equation 12*

The attributable number of asthma incident cases for each census block was then summed across the state to obtain state total AC estimates, and the entire country to obtain the national AC estimates, as shown in Equation 13.

*Equation 13*

**Results**

*NO2 concentrations and trends*

The mean (min-max) NO2 concentrations were 13.2 (1.5-58.3) ug/m3 (Table 2) with the highest mean NO2 concentrations in urbanized areas (18.4 ug/m3) (Figure S1) and among the highest median household income group of ≥$75,000 (16.5 ug/m3) followed by the lowest median household income group of <$20,000 (16.1 ug/m3) (Figure S2). When stratifying NO2 concentrations by median household income groups for urban and rural areas, rural areas had an increasing average concentration as income increased, while urban clusters has a decreasing average concentration as income increased, showing a U-shaped trend (Figure S3 and Figure S4). South Dakota had the lowest mean NO2 concentration (5.2 ug/m3), while D.C. had the highest (26.3 ug/m3) (Table S1 and Figure S5). Figure S6 and Figure S7 demonstrate NO2 concentrations across median household income and urban/rural location, separately for each state.

*Table 2: NO2 concentrations (ug/m3) by strata*

|  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- |
|  |  | **Mean** | **Min** | **25%** | **Median** | **75%** | **Max** |
| **Total** |  | 13.2 | 1.5 | 7.9 | 11.4 | 16.6 | 58.3 |
| **By living location** | **Rural** | 8.0 | 1.5 | 6.0 | 7.8 | 9.8 | 37.7 |
| **Urban cluster** | 12.0 | 1.6 | 9.6 | 11.9 | 14.2 | 35.6 |
| **Urbanized area** | 18.4 | 2.6 | 13.0 | 17.0 | 22.1 | 58.3 |
| **By median household income** | **<$20,000** | 16.1 | 2.0 | 10.4 | 14.9 | 20.1 | 56.8 |
| **$20,000 to <$35,000** | 13.2 | 1.6 | 8.1 | 11.7 | 16.7 | 58.3 |
| **$35,000 to <$50,000** | 11.8 | 1.5 | 7.0 | 10.0 | 14.5 | 58.0 |
| **$50,000 to <$75,000** | 12.8 | 1.6 | 7.6 | 10.8 | 15.7 | 55.7 |
| **≥$75,000** | 16.5 | 2.1 | 10.9 | 14.9 | 20.6 | 55.5 |

*ACBS and BRFSS results*

Overall, there were 32 states where we were able to extract childhood asthma incidence rates and 41 states with childhood asthma prevalence rates (Table S2-S4). The total childhood samples included for the period 2006-2010 were 293,464 samples from the BRFSS and 16,156 samples from the ACBS (Table 3). The BRFSS samples ranged between 55,094 samples (2006) and 61,862 (2008). The ACBS samples ranged between 2,017 samples (2006) and 4,095 (2009). The weighted estimates represent the childhood population counts of available states from the BRFSS and the ACBS for the years when the survey was conducted.

Across all available states, the overall aggregate asthma incidence rate for the years 2006-2010 was 11.6 per 1,000 at-risk children (Table 3). The state of Montana had the lowest aggregate childhood asthma incidence rate (IR = 4.3 per 1,000 at-risk children), while D.C. had the highest aggregate childhood asthma incidence rate (IR = 17.7 per 1,000 at-risk children) (Table S2). States that did not have an incidence rate available (n = 19 states) were assigned the overall aggregate asthma incidence rate of 11.6 per 1,000 at-risk children (Table S2-S4).

The overall aggregate asthma prevalence rate for the years 2006-2010 was 13.1 per 100 children (Table 3). The state of Iowa had the lowest aggregate childhood asthma prevalence rate (PR = 8.4 per 100 children), while D.C. had the highest aggregate childhood asthma prevalence rate (PR = 19.9 per 100 children) (Table S2). States that did not have a prevalence rate available (n = 8 states) were assigned the overall aggregate asthma prevalence rate of 13.1 per 100 children.

*Table 3: Childhood asthma survey summaries*

|  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- |
|  | **2006** | **2007** | **2008** | **2009** | **2010** | **Total** |
| **BRFSS sample (weighted)** | 55,094 (50,674,742) | 59,487 (43,661,381) | 61,862 (53,327,550) | 59,821 (47,747,373) | 57,200 (39,975,264) | 293,464 |
| **Ever asthma sample (weighted)** | 7,168 (6,493,224) | 7,971 (5,763,409) | 8,255 (7,218,400) | 8,126 (6,279,938) | 7,483 (5,158,455) | 39,003 |
| **ACBS Sample (weighted)** | 2,017 (4,580,870) | 2,797 (5,459,638) | 3,924 (4,343,245) | 4,095 (4,154,076) | 2,196 (3,116,669) | 16,156 |
| **Incident case sample (weighted)** | 154 (404,276) | 173 (312,917) | 169 (385,818) | 153 (297,546) | 160 (319,743) | 809 |
| **At-risk sample (weighted)** | 48,080 (30,825,589) | 51,689 (36,050,557) | 53,776 (26,491,259) | 51,848 (25,942,087) | 49,877 (22,900,850) | 255,270 |
| **Incidence rate** | 13.1 | 8.7 | 14.6 | 11.5 | 14.0 | 11.6\* |
| **Prevalence rate** | 12.8 | 13.2 | 13.5 | 13.2 | 12.9 | 13.1\*\* |
| **Number of states included** | 18 | 26 | 20 | 17 | 17 | 32\*\*\* |

*\*Aggregate asthma incidence rate per 1,000 at-risk children*

*\*\*Aggregate asthma prevalence rate per 100 children*

*\*\*Total number of states included in the aggregate asthma incidence rate estimation*

*Asthma incident cases*

Using state-specific asthma incidence rates, the estimated number of incident cases of childhood asthma in 2010 were 747,437 (Table 4). By living location, 19% lived in a rural area, while 9% and 72% lived in an urban cluster and urbanized area, respectively. The largest percentage of childhood asthma cases (30%) lived in an income block group of $50,000 to <$75,000, while the lowest percentage (4%) lived in the lowest income block group of <$20,000. The state with the lowest number of estimated incident cases of childhood asthma was Montana with 866 cases, while the state with the largest number was Texas with 99,084 cases (Table S5).

*Attributable number of cases and fraction*

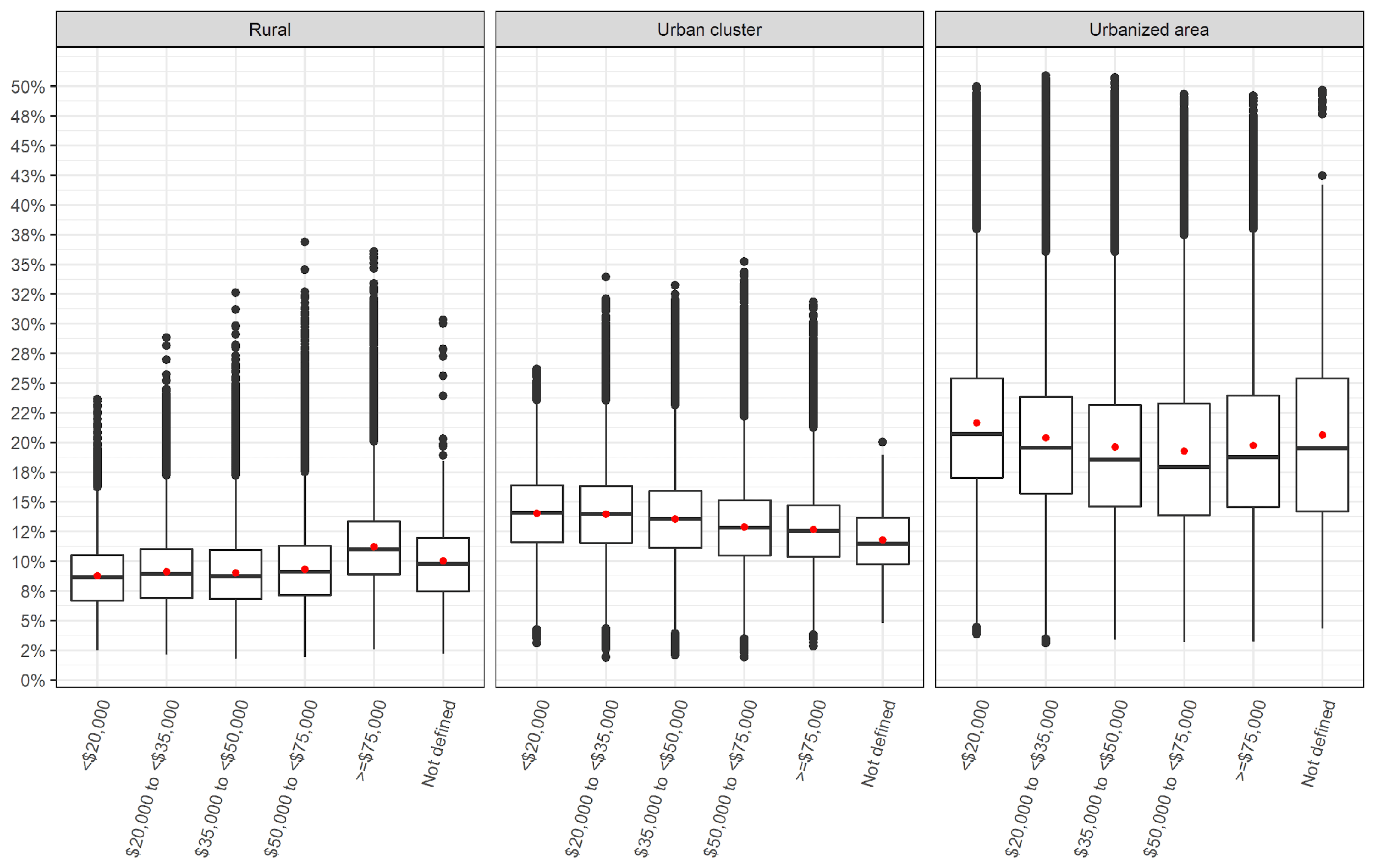
We estimated a total of 131,739 childhood asthma incident cases attributable to NO2 exposure, accounting for 17.6% of all childhood asthma incident cases (Table 4). By living location, urbanized areas had the largest number of attributable cases totaling 108,745 cases and the highest percentage of all asthma incident cases at 20.3%. Rural areas had a total of 13,788 cases and accounted for the lowest percentage of all asthma cases at 9.8%, while urban clusters had only 9,206 cases representing 13.1% of all asthma incident cases (Table 4). By median household income, census blocks with incomes of $50,000 to <$75,000 had the largest number of cases attributable to NO2; 37,253 cases accounting for 16.8% of all asthma incident cases. However, the income group with the largest proportion of asthma cases attributable to NO2 exposure was the lowest income group <$20,000, accounting for 20.8% of all asthma incident cases (Table 4).

The distribution of attributable fractions at the census block level shows that the mean value was higher in Urbanized areas compared to rural areas (Figure S8), and followed a U shape distribution by income group (Figure S9). When examining the distribution of attributable fraction across income groups stratified by living location we observe that the mean value increased by increasing income group in rural areas, decreased by increasing income group in urban clusters and presented as a U shape in urbanized areas (Figure 2).

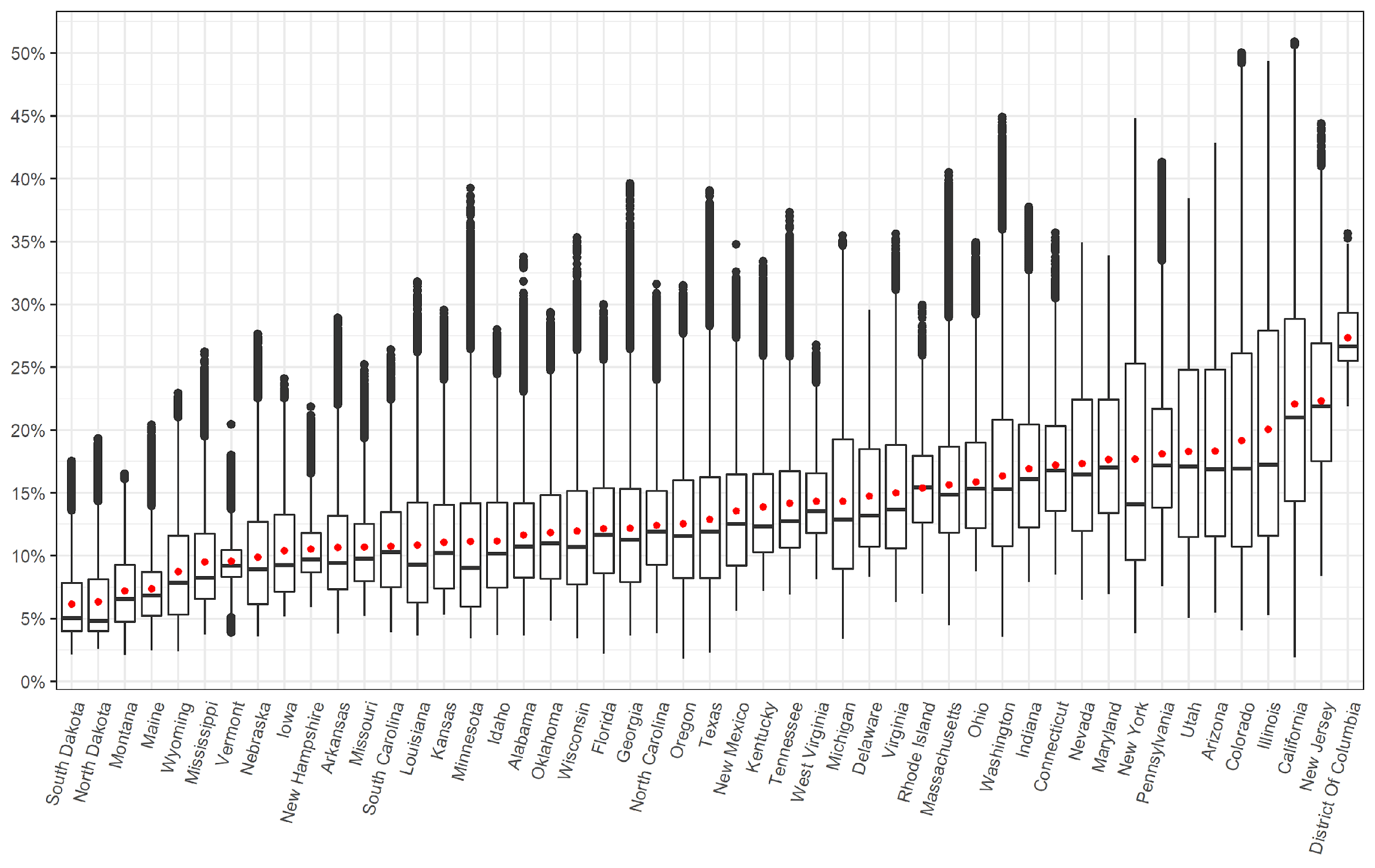
The state with the lowest number of estimated attributable cases was Montana with 69 cases, while the state with the largest number of estimated attributable cases was California with 19,205 cases (Table S4 ). The state with the lowest attributable fraction was South Dakota (7.5%), while the state with the highest attributable fraction was D.C. (26.9%) (Table S5). When examining the distribution of attributable fractions across all census blocks we observe that the state with the lowest average value was South Dakota while the state with the largest average value was District of Columbia (Figure 3).

*Table 4: Comparing the results of the burden of disease using state-specific estimates vs original estimates*

|  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  |  | **Results using flat national-level IR** | | | **Results using state-specific IR** | | | **Difference** | | | **Difference (%)** | | |
|  |  | **Incident cases** | **AC** | **AF** | **Incident cases** | **AC** | **AF** | **Incident cases** | **AC** | **AF** | **Incident cases** | **AC** | **AF** |
|  | **Total** | 794,934 | 141,931 | 17.9% | 747,437 | 131,739 | 17.6% | -47,497 | -10,192 | -0.3% | -6.0% | -7.2% | -1.7% |
| **By living location (% of Total)** | **Rural** | 148,470 (19%) | 14,466 (10%) | 9.7% | 140,799 (19%) | 13,788 (10%) | 9.8% | -7,671 | -678 | 0.1% | -5.2% | -4.7% | 1.0% |
| **Urban cluster** | 75,453 (9%) | 9,844 (7%) | 13.0% | 70,524 (9%) | 9,206 (7%) | 13.1% | -4,929 | -638 | 0.1% | -6.5% | -6.5% | 0.8% |
| **Urbanized area** | 571,011 (72%) | 117,621 (83%) | 20.6% | 536,113 (72%) | 108,745 (83%) | 20.3% | -34,898 | -8,876 | -0.3% | -6.1% | -7.5% | -1.5% |
| **By median household income (% of Total)** | **<$20,000** | 28,207 (4%) | 5,892 (4%) | 20.9% | 27,770 (4%) | 5,786 (4%) | 20.8% | -437 | -106 | -0.1% | -1.5% | -1.8% | -0.5% |
| **$20,000 to <$35,000** | 137,765 (17%) | 25,794 (18%) | 18.7% | 132,843 (18%) | 24,699 (19%) | 18.6% | -4,922 | -1,095 | -0.1% | -3.6% | -4.2% | -0.5% |
| **$35,000 to <$50,000** | 200,367 (25%) | 34,549 (24%) | 17.2% | 188,466 (25%) | 32,088 (24%) | 17.0% | -11,901 | -2,461 | -0.2% | -5.9% | -7.1% | -1.2% |
| **$50,000 to <$75,000** | 236,827 (30%) | 40,540 (29%) | 17.1% | 221,334 (30%) | 37,253 (28%) | 16.8% | -15,493 | -3,287 | -0.3% | -6.5% | -8.1% | -1.8% |
| **≥$75,000** | 191,621 (24%) | 35,128 (25%) | 18.3% | 176,880 (24%) | 31,885 (24%) | 18.0% | -14,741 | -3,243 | -0.3% | -7.7% | -9.2% | -1.6% |

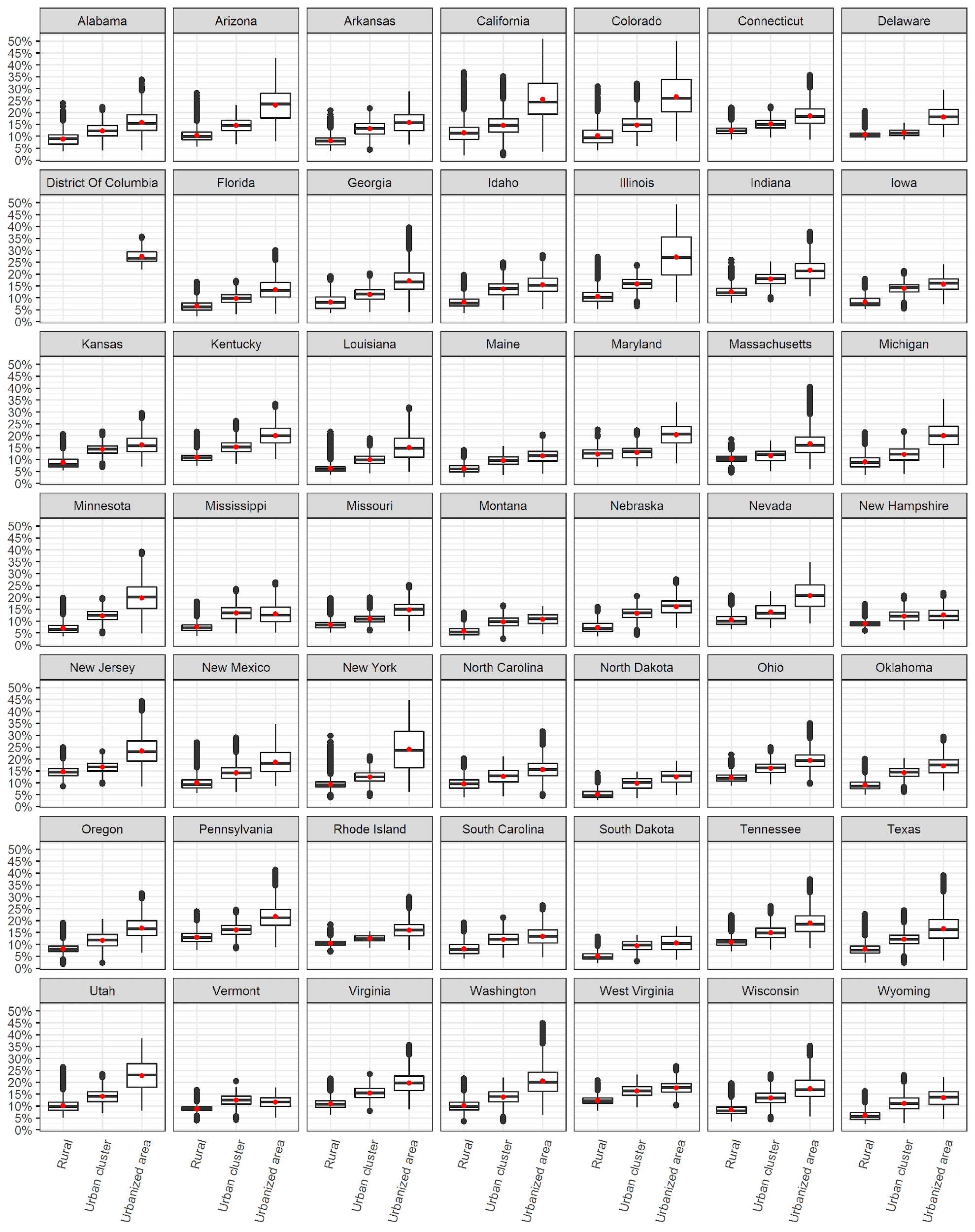
*Figure 2: Distribution of attributable fractions by median household income group stratified into living location*

*\*Red dot represents the mean value while the midline represents the median value of attributable fractions across all the census blocks*

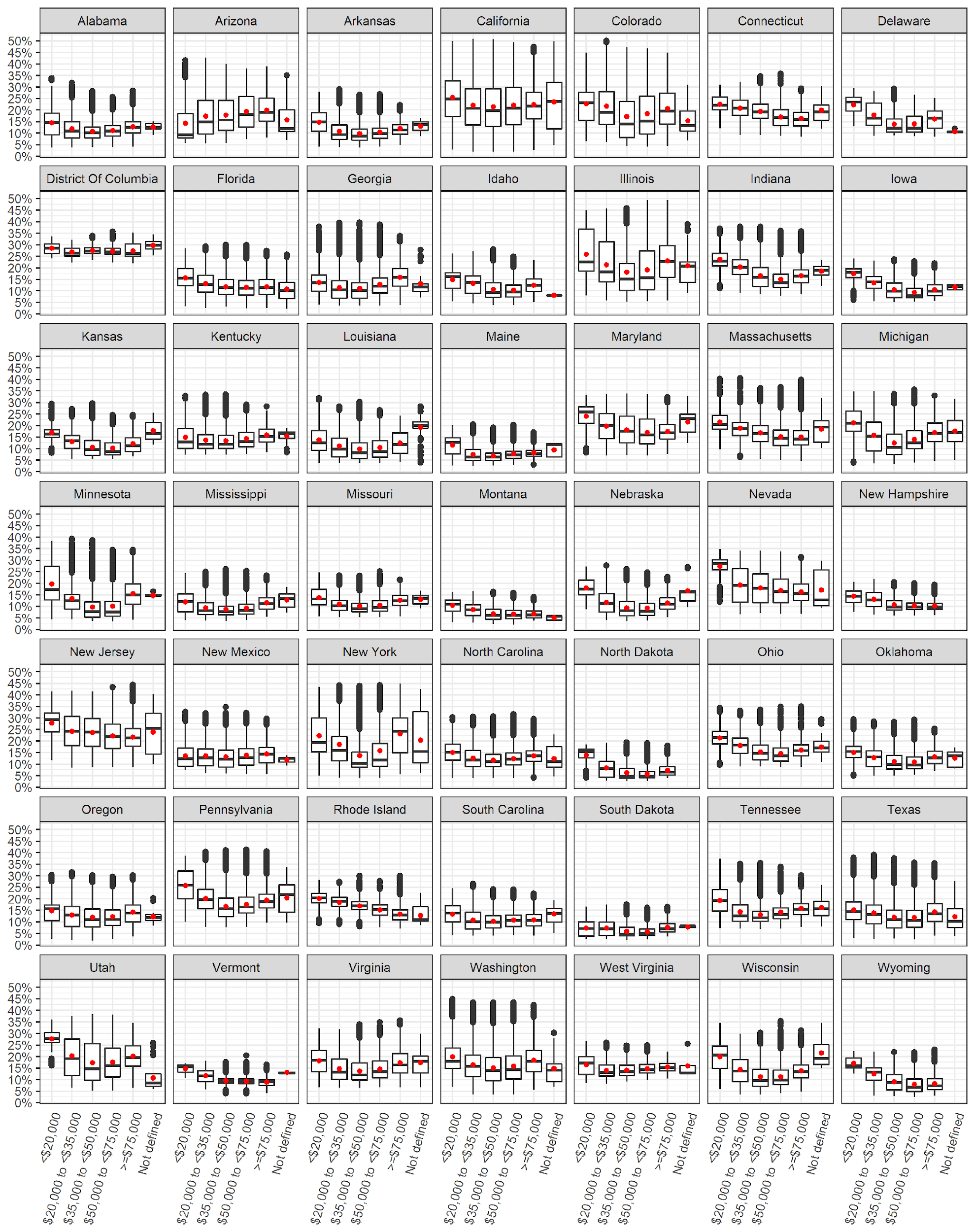
*Figure 3: Distribution of attributable fractions by state*

Figures 4 and 5 present the distribution of attributable fractions across all census blocks by living location and median household income group for each state. The majority of states broadly follow a distribution similar to the national level as shown in Figure S8 and Figure S9, with a few exceptions (By living location see; Delaware, Maryland, Mississippi, Vermont. By median household income see; Arizona, Connecticut, D.C., Florida, Maine, Massachusetts, Montana, Nevada, New Hampshire, New Jersey, New Mexico, Vermont, West Virginia, Rhode Island and Wyoming).

*Figure 4: Distribution of attributable fraction by state and living location*

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*Figure 5: Distribution of attributable fraction by state and median household income group*

****

*Comparing total asthma incident cases using state-specific versus a flat national-level incidence rate*

Using state-specific asthma incidence rates, the overall number of incident asthma cases was reduced by 47,497 (6%) compared to estimates that used a flat national-level asthma incidence rate of 12.5 per 1,000 at-risk children (Table 4) (Alotaibi et al., 2019). By living location, the largest relative change was among urban clusters, with a decrease of 4,929 (6.5%) cases, followed by urbanized areas with a reduction of 34,898 (6.1%) cases. By income group, the largest relative change in the number of cases was among the highest income groups with a decrease of 14,741 (7.7%) cases, while the smallest relative change was among the lowest income group, with a decrease of 437 (1.5%) cases (Table 4). California had the largest decrease in numbers of total childhood asthma incident cases (24,441 cases), while Texas had the largest increase in numbers of total childhood asthma incident cases (25,019 cases) (Table S5). Montana had the largest relative reduction in total childhood asthma incident cases (64.1%), while Texas had the largest relative increase (33.8%) (Table S5).

*Comparing attributable asthma incident cases due to NO2 using state-specific versus a flat national-level incidence rate*

The total number of attributable cases was reduced by 10,192 (7.2%) when compared to the flat national-level asthma incidence rate (Table 4). By living location, urbanized areas had the largest relative change with a decrease of 8,876 (7.5%) cases, while rural areas had the least relative change with a decrease of 678 (4.7%) cases attributable to NO2 exposure. By income group, the highest income group had the largest relative change with a decrease in attributable cases by 3,243 (9.2%) while the lowest income group had the least relative change with a decrease of 106 (1.8%) cases. California had the largest decrease in attributable cases (6,190 cases), while Texas had the largest increase (3,615 cases) (Table S5).

*Comparing attributable asthma incident fractions due to NO2 using state-specific versus a flat national-level incidence rate*

The overall attributable fraction was reduced by 0.3% (a 1.7% reduction). In terms of living location, urbanized areas had the largest relative reduction by 1.5%, while rural areas had a relative increase by 1%. In terms of income group, the largest relative reduction was 1.8% for the $50,000 to <$75,000 income strata (Table 4). The attributable fraction across states did not differ when using state-specific asthma incidence rates, the difference observed across some states in (*Table S5*) are due to rounding errors.

**Discussion**

*Summary and key findings*

Using validated NO2 exposure data and U.S. Census data from the year 2010, we investigated the impact baseline childhood asthma incidence rates have on BoD estimates across 48 U.S. states and D.C. to document differences in BoD estimates when using state-specific versus a flat national-level asthma incidence rate. We utilized data from 2006-2010 surveys from the CDC to estimate the number of at-risk children, state-specific incidence and prevalence rates, and applied these estimates to population data at the census block level, in the most comprehensive exercise to date. Previous literature relied on national-level asthma incidence rates for BoD assessments and the impact of this simplification was unknown. Using the state-specific asthma incidence rates, we also explored trends in BoD estimates stratifying our analyses by socioeconomic status and urban versus rural status to confirm (or otherwise) trends we observed in past analysis which employed a flat national-level incidence rate (Alotaibi et al., 2019).

At the national level, the difference in the estimated BoD using state-specific versus a national-level incidence rate was relatively small. Using the state-specific incidence rates resulted in a 7.2% reduction of 10,192 attributable number of cases, which amounted to a 1.7% reduction in the attributable fraction. While a 1.7% reduction in the attributable fraction is relatively small and needs to be carefully weighed against the effort required to produce state-specific estimates, its implications might be more significant depending on if/how BoD estimates are used in policy making. For example, according to Perry et al. (2019), the average annual costs of asthma per child in the U.S. ranged from $3,076 to $13,612. In simplistic terms, using the state-specific versus the national-level incidence rate could result in a reduction ranging from $31,350,592 to $138,733,504 in estimated burden costs, per year. This is significant. More interestingly, however, is the variation across the states which we expect to be more accurately captured when using the state-specific incidence rates. Using the state-specific incidence rates, which we expected to more accurately capture between-state variation, resulted in a change in the attributable number of cases ranging from -64.1% (Montana) to 33.6% (Texas). California had the largest absolute decrease in the number of attributable cases by 6,190 cases while the state of Texas had the largest increase by 3,615 cases, followed by New York (1,750).

Further, stratifying our analyses by socioeconomic status and urban versus rural living produced new trends than what we observed in past analysis which employed a flat national-level incidence rate (Alotaibi et al., 2019). For example, Most of the change in the attributable number of cases (-7.5%) occurred in urbanized areas among the highest median household income group of ≥$75,000 (-9.2%). The distribution of the attributable fractions across all census blocks by living location and median household income group was investigated separately for each state, rather than the national aggregation which we followed in previous analysis (Alotaibi et al., 2019). The majority of states broadly followed the U-shaped distribution observed at the national level; where the lowest and the highest median household income groups had the highest burden. There were many exceptions, however, including Arizona, Connecticut, D.C., Florida, Maine, Massachusetts, Montana, Nevada, New Hampshire, New Jersey, New Mexico, Vermont, West Virginia, Rhode Island and Wyoming. On the other hand, there were states where this U-shaped distribution was more prominent including Illinois and New York. This may reflect a trend of the lowest income populations living in the most polluted census blocks due to financial restrain, a trend that is well established in the environmental justice literature. On the other hand, it is possible, although counter-intuitive, that the highest income populations live in highly polluted census blocks close to busy downtowns and central business districts. If this was the case, it is clear that this trend does not apply to all states and may even differ by rural, urban cluster and urbanized area status (Figure S4, S6 and S7). Previous work suggested that metropolitan areas in particular exhibit considerable heterogeneity when it comes to socioeconomic status and exposure to air pollution. For example, in cities like New York, wealthy neighborhoods have been associated with higher concentrations of pollution (Hajat et al., 2013). These trends warrant further investigation, ideally using more refined air pollution and asthma incidence estimates at an even more granular spatial scale.

*Comparison with previous studies*

All previous BoD studies investigating this topic (Achakulwisut et al., 2019, Khreis et al., 2018c, Khreis et al., 2018b, Alotaibi et al., 2019, Anenberg et al., 2018) have identified several data gaps which might impact the final BoD estimates and potentially introduce uncertainty and error. In fact, these gaps are applicable to BoD studies of air pollution and any health outcome. Previous studies stressed that the accuracy of the BoD estimate is dependent on the accuracy of input data, namely: 1) the air pollution exposure levels and distribution, 2) the concentration-response functions, and 3) the baseline asthma incidence rates. Some of the studies cited above investigated the impacts of different input datasets on final BoD estimates and found that different exposure assessment methods (e.g. dispersion versus land use regression modeling) can result in up to 3% difference in BoD estimates (Khreis et al., 2018a). Alotaibi et al. (2019) explored the impact of uncertainty in concentration-response functions (CRF) on final BoD estimates and found that using the most conservative end of the CRF (the lower 95% CI) can reduce the estimated burden by up to 19% when compared to the central estimate. On the other hand, using the most extreme end of the CRF (the upper 95% CI) can increase the estimated burden by up to 14% when compared to the central estimate. Finally, one English study (Khreis et al., 2018a) and their follow-up work in Khreis et al. (2019) showed that using a national versus a local baseline asthma incidence rate can result in up to 11% difference in BoD estimates, but this analysis was limited to one medium-sized city North of England.

The impact that the asthma baseline incidence rates have on BoD estimates has not yet been thoroughly studied. Childhood asthma can be challenging to diagnose, and national-level incidence rates might vary substantially at smaller spatial resolutions and between urban and rural populations. Previous literature relied on national-level asthma incidence rates, which is also in line with practice by prominent institutions and studies such as the Global Burden of Disease analyses. In this paper, we aimed to explore the potential impact of using state-specific varying asthma incidence rates on the final burden of childhood asthma due to NO2 and compare the change in BoD estimates to those produced by Alotaibi et al. (2019) who used a national-level asthma incidence rate, as is typically practiced (Achakulwisut et al., 2019, Khreis et al., 2018c, Perez et al., 2009, Perez et al., 2013, Khreis et al., 2018b, Anenberg et al., 2018). Using this more granular state-specific asthma incidence rates, we also explored trends in the BoD estimates by socioeconomic status and urban versus rural status to confirm (or otherwise) previous trends we observed in past analysis. We selected NO2 as the exposure of interest, as more studies underline this pollutant’s CRF and as it has been the most commonly used pollutant in previous similar analyses. Further, the selection of the pollutant in this instance is less relevant as our aim was to establish the difference in BoD estimates from previous analyses, only by altering the baseline childhood asthma incidence rate.

*Strengths and limitations*

We selected NO2 as the exposure of interest, as more studies underline this pollutant’s CRF and as it has been the most commonly used pollutant in previous similar analyses. The results in this paper re-affirm our finding in the main paper, more specifically the sensitivity analysis. In our main paper we argued that changing or testing ranges of asthma IR in the analysis did not alter, in magnitude, the end results as compared to a change in the CRF function (Check the sensitivity matrix).

* We can derive from this finding, that using national level incidence rates to estimate asthma burden due to TRAP is acceptable even if asthma incidence rates is not available at a finer level.
* The state-specific attributable fractions did not change. The reason is that the incident rate is applied uniformly across the state (spatially), thus the total asthma cases and total attributable cases will change with equal proportion when applying the new asthma incidence rate. The attributable fraction is a function of CRF and exposure estimate regardless of the IR. Had we applied an incidence rate based on other factors like age, gender, race, income group, then the attributable fraction across the state would differ since the change in incidence rate won’t be uniform within the state.
* Third, we relied on national paediatric asthma incidence rates in our analysis, but subnational variations most likely exist, especially between urban and rural populations

**Conclusions**

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**Supplementary Material**

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*Table S1: NO2 concentration (ug/m3) by state*

|  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- |
| **State** | **Mean** | **Min** | **25%** | **Median** | **75%** | **Max** |
| Alabama | 10.3 | 3.0 | 7.1 | 9.3 | 12.5 | 33.8 |
| Arizona | 17.0 | 4.6 | 10.1 | 15.1 | 23.4 | 45.9 |
| Arkansas | 9.3 | 3.2 | 6.2 | 8.1 | 11.6 | 28.0 |
| California | 21.1 | 1.6 | 12.7 | 19.3 | 27.9 | 58.3 |
| Colorado | 18.1 | 3.4 | 9.3 | 15.2 | 24.8 | 56.9 |
| Connecticut | 15.6 | 7.3 | 11.9 | 15.0 | 18.6 | 36.2 |
| Delaware | 13.2 | 7.1 | 9.3 | 11.6 | 16.7 | 28.7 |
| D.C. | 26.3 | 20.2 | 24.2 | 25.4 | 28.5 | 36.1 |
| Florida | 10.7 | 1.8 | 7.4 | 10.2 | 13.7 | 29.2 |
| Georgia | 10.8 | 3.0 | 6.8 | 9.8 | 13.6 | 41.4 |
| Idaho | 9.8 | 3.1 | 6.4 | 8.8 | 12.6 | 26.9 |
| Illinois | 19.0 | 4.4 | 10.1 | 15.5 | 26.9 | 55.7 |
| Indiana | 15.4 | 6.7 | 10.7 | 14.4 | 18.7 | 38.9 |
| Iowa | 9.1 | 4.3 | 6.1 | 8.0 | 11.7 | 22.6 |
| Kansas | 9.7 | 4.5 | 6.3 | 8.8 | 12.4 | 28.7 |
| Kentucky | 12.4 | 6.1 | 8.9 | 10.8 | 14.8 | 33.3 |
| Louisiana | 9.6 | 3.0 | 5.3 | 8.0 | 12.6 | 31.4 |
| Maine | 6.3 | 2.0 | 4.4 | 5.8 | 7.5 | 18.7 |
| Maryland | 16.1 | 5.9 | 11.8 | 15.3 | 20.8 | 34.0 |
| Massachusetts | 14.1 | 3.7 | 10.3 | 13.2 | 17.0 | 42.5 |
| Michigan | 12.9 | 2.8 | 7.7 | 11.3 | 17.5 | 35.9 |
| Minnesota | 9.9 | 2.9 | 5.0 | 7.8 | 12.5 | 40.8 |
| Mississippi | 8.3 | 3.1 | 5.6 | 7.0 | 10.2 | 24.9 |
| Missouri | 9.3 | 4.4 | 6.8 | 8.4 | 11.0 | 23.8 |
| Montana | 6.2 | 1.7 | 4.0 | 5.5 | 8.0 | 14.8 |
| Nebraska | 8.6 | 3.0 | 5.2 | 7.7 | 11.1 | 26.5 |
| Nevada | 15.9 | 5.5 | 10.5 | 14.7 | 20.8 | 35.2 |
| New Hampshire | 9.1 | 5.0 | 7.4 | 8.4 | 10.3 | 20.2 |
| New Jersey | 21.0 | 7.1 | 15.8 | 20.2 | 25.7 | 48.1 |
| New Mexico | 12.1 | 4.7 | 7.9 | 11.0 | 14.8 | 35.0 |
| New York | 16.6 | 3.2 | 8.3 | 12.4 | 23.9 | 48.7 |
| North Carolina | 11.0 | 3.2 | 8.0 | 10.4 | 13.5 | 31.1 |
| North Dakota | 5.4 | 2.1 | 3.3 | 4.0 | 6.9 | 17.6 |
| Ohio | 14.3 | 7.5 | 10.7 | 13.6 | 17.3 | 35.2 |
| Oklahoma | 10.4 | 4.1 | 7.0 | 9.5 | 13.1 | 28.5 |
| Oregon | 11.1 | 1.5 | 7.0 | 10.1 | 14.3 | 31.0 |
| Pennsylvania | 16.6 | 6.4 | 12.2 | 15.5 | 20.1 | 43.7 |
| Rhode Island | 13.8 | 5.9 | 11.1 | 13.7 | 16.2 | 29.2 |
| South Carolina | 9.4 | 3.3 | 6.4 | 8.9 | 11.9 | 25.1 |
| South Dakota | 5.2 | 1.8 | 3.3 | 4.2 | 6.7 | 15.8 |
| Tennessee | 12.7 | 5.9 | 9.2 | 11.2 | 15.0 | 38.3 |
| Texas | 11.5 | 1.9 | 7.0 | 10.4 | 14.5 | 40.6 |
| Utah | 17.0 | 4.3 | 10.0 | 15.4 | 23.4 | 39.8 |
| Vermont | 8.3 | 3.3 | 7.1 | 7.9 | 9.1 | 18.7 |
| Virginia | 13.5 | 5.3 | 9.2 | 12.0 | 17.1 | 36.1 |
| Washington | 14.9 | 2.9 | 9.3 | 13.6 | 19.1 | 48.9 |
| West Virginia | 12.7 | 6.9 | 10.3 | 11.9 | 14.9 | 25.5 |
| Wisconsin | 10.6 | 2.8 | 6.6 | 9.3 | 13.5 | 35.7 |
| Wyoming | 7.6 | 2.0 | 4.5 | 6.7 | 10.1 | 21.4 |

*Table S2: Available childhood asthma incidence rates by state and year*

|  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- |
| **State** | **2006\*** | **2007\*** | **2008\*** | **2009\*** | **2010\*** | **Aggregate IR\*** | **Aggregate PR\*\*** |
| Alabama |  |  |  |  |  | 11.6 | 14.4 |
| Arizona | 23.7 | 6.8 |  |  |  | 15.2 | 13.1 |
| Arkansas |  |  |  |  |  | 11.6 | 13.1 |
| California | 12.1 | 6.5 |  |  |  | 9.3 | 12.2 |
| Colorado |  |  |  |  |  | 11.6 | 13.1 |
| Connecticut |  | 9.9 | 14.1 | 10.8 | 13.5 | 12 | 16 |
| Delaware |  |  |  |  |  | 11.6 | 18.2 |
| D.C. | 5.3 | 28.8 |  |  |  | 17.7 | 19.9 |
| Florida |  |  |  |  |  | 11.6 | 13.1 |
| Georgia | 6.4 | 5.8 | 9.1 | 16.6 | 6.9 | 9.1 | 15.1 |
| Idaho |  |  |  |  |  | 11.6 | 9 |
| Illinois |  | 4.2 |  | 9.2 |  | 6.7 | 12.4 |
| Indiana | 25.4 | 9.3 | 13.4 | 9.9 | 17.6 | 15.2 | 12.8 |
| Iowa | 5 | 4 | 9.9 |  |  | 6.3 | 8.4 |
| Kansas | 7.8 | 9.9 | 9.9 | 8.3 | 9 | 9 | 11.6 |
| Kentucky |  |  |  |  |  | 11.6 | 14 |
| Louisiana |  |  |  | 5.8 |  | 5.8 | 13 |
| Maine | 13 | 8.7 | 5.8 |  |  | 9.2 | 13.2 |
| Maryland | 16.2 | 8.6 | 11 | 17.3 | 2.3 | 11.2 | 14.8 |
| Massachusetts |  |  |  |  |  | 11.6 | 13.1 |
| Michigan | 5.3 | 7.7 | 5.2 | 13.4 | 29.3 | 12 | 13.6 |
| Minnesota |  |  |  |  |  | 11.6 | 9.5 |
| Mississippi |  | 10.8 |  |  | 17.2 | 14 | 14.2 |
| Missouri | 21.2 | 10.3 | 7.2 |  |  | 12.9 | 13.9 |
| Montana | 2.8 | 2 |  | 3.7 | 8.5 | 4.3 | 9.7 |
| Nebraska | 11.9 | 8.3 | 8.9 | 3.3 | 12.9 | 9.1 | 9.3 |
| Nevada |  |  |  |  |  | 11.6 | 10.9 |
| New Hampshire | 11.5 | 13.8 | 10.4 |  |  | 12 | 12.1 |
| New Jersey |  |  | 6.3 | 12.5 | 10.5 | 9.8 | 14.3 |
| New Mexico |  | 3.2 | 9.5 |  | 7.2 | 6.7 | 12 |
| New York | 12.9 | 6.1 | 28.4 | 11.2 |  | 14.7 | 15.8 |
| North Carolina |  |  |  |  |  | 11.6 | 13.1 |
| North Dakota |  |  |  |  |  | 11.6 | 8.9 |
| Ohio |  | 13.1 | 17 |  |  | 15.1 | 12.3 |
| Oklahoma |  | 9.2 | 10.1 |  | 12.9 | 10.8 | 14 |
| Oregon |  | 11.1 |  |  |  | 11.1 | 11.1 |
| Pennsylvania |  | 21.8 |  |  | 4.3 | 13.2 | 13.9 |
| Rhode Island |  |  | 15.3 | 13.2 |  | 14.3 | 16.1 |
| South Carolina |  |  |  |  |  | 11.6 | 13.1 |
| South Dakota |  |  |  |  |  | 11.6 | 13.1 |
| Tennessee |  |  |  |  |  | 11.6 | 13.1 |
| Texas | 14.4 |  | 18.2 | 12.5 | 21 | 16.6 | 13.1 |
| Utah |  | 15.4 | 11.9 | 5.6 | 9.3 | 10.4 | 10.2 |
| Vermont | 13.5 | 4.4 | 8.5 | 21.2 | 10.4 | 11.5 | 13.8 |
| Virginia |  |  |  |  |  | 11.6 | 13.6 |
| Washington |  |  |  | 7.9 | 5.6 | 6.8 | 10.8 |
| West Virginia |  | 11.8 |  |  |  | 11.8 | 12.7 |
| Wisconsin | 12.3 |  |  |  |  | 12.3 | 10.6 |
| Wyoming |  |  |  |  |  | 11.6 | 9.5 |

*\*Incidence rate per 1,000 at-risk children*

*\*\* Prevalence rate per 100 children*

*[Note: The grey highlight indicates states/years with no available data]*

*Table S3: Childhood asthma survey summary by state (Total of 2006-2010)*

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
| **State** | **Total ACBS sample** | **Total BRFSS sample** | **Total ever asthma** | **Total incident cases** |
| Arizona | 103 | 5,535 | 699 | 10 |
| California | 172 | 11,801 | 1,543 | 13 |
| Connecticut | 549 | 7,112 | 1,132 | 47 |
| D.C. | 69 | 4,101 | 685 | 6 |
| Georgia | 545 | 9,433 | 1,455 | 26 |
| Illinois | 122 | 6,187 | 778 | 6 |
| Indiana | 500 | 9,824 | 1,361 | 41 |
| Iowa | 245 | 8,084 | 724 | 19 |
| Kansas | 827 | 14,699 | 1,839 | 50 |
| Louisiana | 88 | 8,829 | 1,214 | 4 |
| Maine | 376 | 4,523 | 644 | 23 |
| Maryland | 624 | 13,093 | 1,897 | 44 |
| Michigan | 680 | 10,762 | 1,524 | 43 |
| Mississippi | 208 | 10,816 | 1,527 | 14 |
| Missouri | 262 | 5,646 | 814 | 20 |
| Montana | 286 | 8,609 | 909 | 17 |
| Nebraska | 717 | 17,883 | 1,644 | 53 |
| New Hampshire | 232 | 5,285 | 664 | 19 |
| New Jersey | 458 | 15,410 | 2,230 | 32 |
| New Mexico | 287 | 5,554 | 765 | 17 |
| New York | 404 | 7,083 | 1,079 | 28 |
| Ohio | 351 | 7,989 | 1,138 | 32 |
| Oklahoma | 299 | 8,611 | 1,291 | 21 |
| Oregon | 165 | 4,793 | 579 | 13 |
| Pennsylvania | 209 | 14,760 | 2,090 | 12 |
| Rhode Island | 169 | 7,127 | 1,209 | 11 |
| Texas | 780 | 16,749 | 2,293 | 55 |
| Utah | 573 | 14,417 | 1,617 | 45 |
| Vermont | 597 | 8,784 | 1,220 | 40 |
| Washington | 594 | 9,706 | 1,165 | 33 |
| West Virginia | 85 | 5,089 | 663 | 5 |
| Wisconsin | 140 | 5,170 | 611 | 10 |

\**Incidence rate* per 1,000 at-risk children

\**Prevalence rate per 100 children*

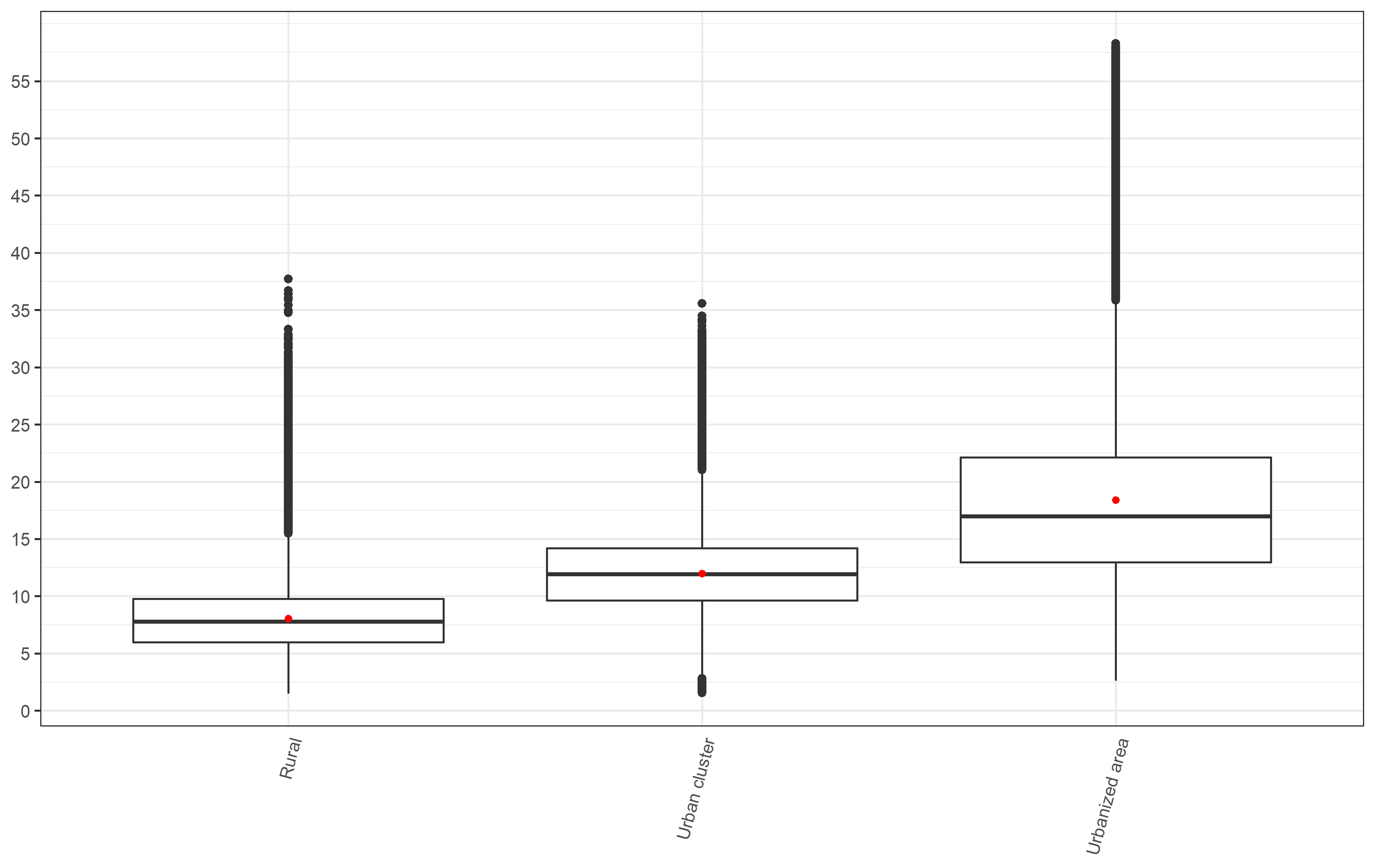
*Table S4: Aggregated childhood asthma weighted survey summary (2006-2010)*

|  |  |  |  |
| --- | --- | --- | --- |
| **State** | **Aggregated weighted**  **incident cases** | **Aggregated at-risk**  **children** | **Years of available**  **data** |
| Arizona | 42,622 | 2,802,422 | 2 |
| California | 156,599 | 16,850,453 | 2 |
| Connecticut | 32,939 | 2,734,478 | 4 |
| D.C. | 3,184 | 179,493 | 2 |
| Georgia | 94,786 | 10,458,074 | 5 |
| Illinois | 37,799 | 5,673,571 | 2 |
| Indiana | 105,219 | 6,936,762 | 5 |
| Iowa | 11,510 | 1,829,734 | 3 |
| Kansas | 27,509 | 3,059,760 | 5 |
| Louisiana | 5,379 | 931,966 | 1 |
| Maine | 6,662 | 722,763 | 3 |
| Maryland | 64,871 | 5,816,584 | 5 |
| Michigan | 126,102 | 10,491,065 | 5 |
| Mississippi | 18,264 | 1,300,917 | 2 |
| Missouri | 46,410 | 3,600,272 | 3 |
| Montana | 3,296 | 768,012 | 4 |
| Nebraska | 18,262 | 2,014,605 | 5 |
| New Hampshire | 9,423 | 788,302 | 3 |
| New Jersey | 51,472 | 5,274,310 | 3 |
| New Mexico | 8,857 | 1,327,496 | 3 |
| New York | 221,226 | 15,027,481 | 4 |
| Ohio | 71,568 | 4,755,245 | 2 |
| Oklahoma | 24,628 | 2,285,659 | 3 |
| Oregon | 8,328 | 752,768 | 1 |
| Pennsylvania | 62,292 | 4,733,925 | 2 |
| Rhode Island | 5,476 | 384,117 | 2 |
| Texas | 381,999 | 22,992,023 | 4 |
| Utah | 30,221 | 2,902,955 | 4 |
| Vermont | 6,498 | 563,280 | 5 |
| Washington | 18,647 | 2,752,373 | 2 |
| West Virginia | 3,847 | 325,031 | 1 |
| Wisconsin | 14,404 | 1,174,447 | 1 |

*Table S5: State-specific results and comparison*

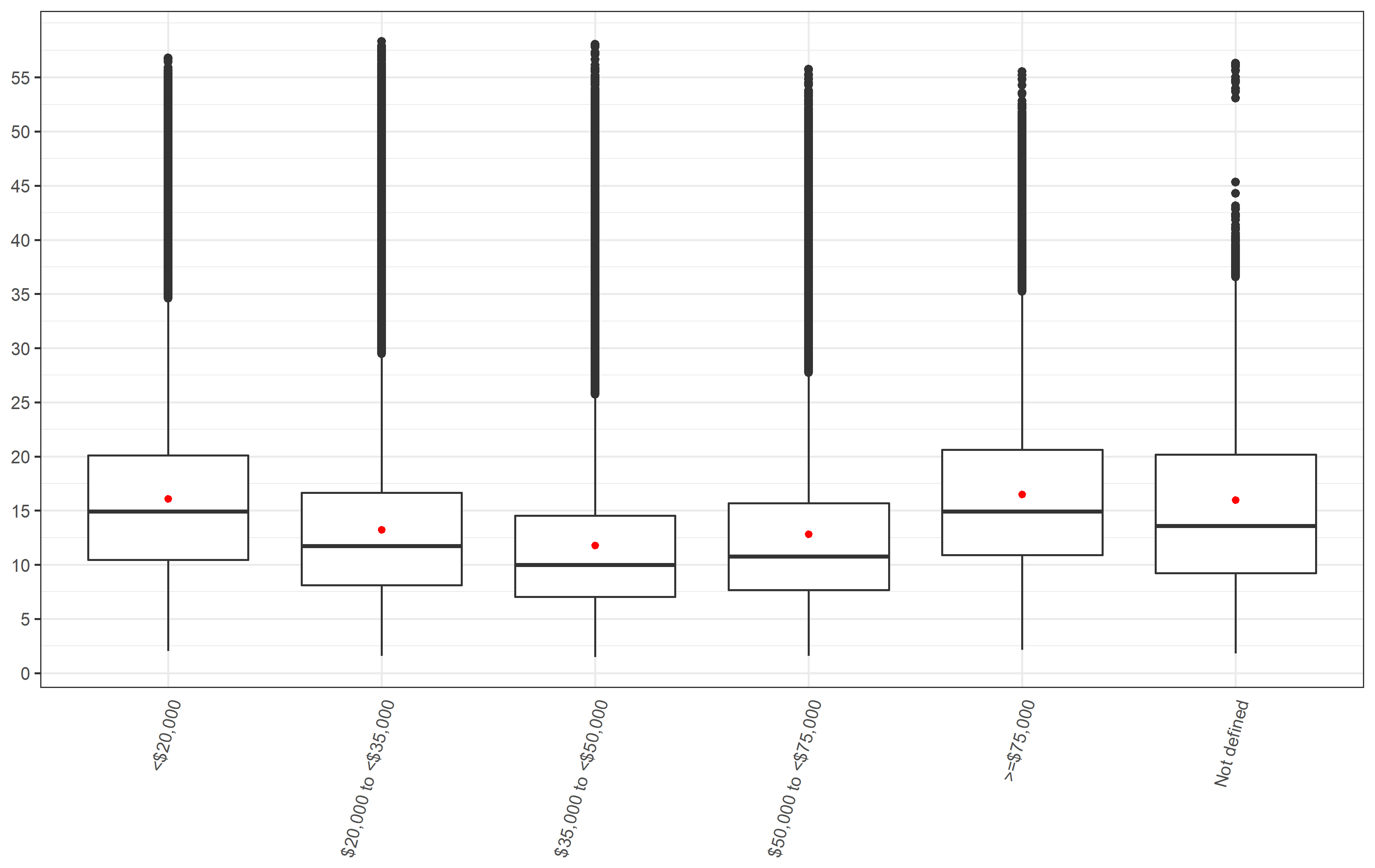
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  | **Results using flat national-level IR** | | | **Results using state-specific IR** | | | **Difference** | | | **Difference (%)** | | |
| **State** | **Incident cases** | **AC** | **AF** | **Incident cases** | **AC** | **AF** | **Incident cases** | **AC** | **AF** | **Incident cases** | **AC** | **AF** |
| **Alabama** | 12,200 | 1,400 | 11.5% | 11,300 | 1,300 | 11.5% | -900 | -100 | 0.0% | -7.4% | -7.1% | 0.3% |
| **Arizona** | 17,600 | 3,800 | 21.6% | 21,500 | 4,600 | 21.4% | 3,900 | 800 | -0.2% | 22.2% | 21.1% | -0.9% |
| **Arkansas** | 7,700 | 900 | 11.7% | 7,200 | 800 | 11.1% | -500 | -100 | -0.6% | -6.5% | -11.1% | -4.9% |
| **California** | 100,300 | 25,400 | 25.3% | 75,800 | 19,200 | 25.3% | -24,500 | -6,200 | 0.0% | -24.4% | -24.4% | 0.0% |
| **Colorado** | 13,200 | 3,100 | 23.5% | 12,400 | 2,900 | 23.4% | -800 | -200 | -0.1% | -6.1% | -6.5% | -0.4% |
| **Connecticut** | 8,800 | 1,600 | 18.2% | 8,300 | 1,500 | 18.1% | -500 | -100 | -0.1% | -5.7% | -6.3% | -0.6% |
| **Delaware** | 2,200 | 400 | 18.2% | 2,000 | 300 | 15.0% | -200 | -100 | -3.2% | -9.1% | -25.0% | -17.5% |
| **D.C.** | 1,100 | 300 | 27.3% | 1,400 | 400 | 28.6% | 300 | 100 | 1.3% | 27.3% | 33.3% | 4.8% |
| **Florida** | 43,200 | 5,500 | 12.7% | 40,500 | 5,200 | 12.8% | -2,700 | -300 | 0.1% | -6.3% | -5.5% | 0.8% |
| **Georgia** | 26,900 | 3,900 | 14.5% | 19,200 | 2,800 | 14.6% | -7,700 | -1,100 | 0.1% | -28.6% | -28.2% | 0.6% |
| **Idaho** | 4,600 | 600 | 13.0% | 4,500 | 600 | 13.3% | -100 | 0 | 0.3% | -2.2% | 0.0% | 2.2% |
| **Illinois** | 33,800 | 8,300 | 24.6% | 18,300 | 4,500 | 24.6% | -15,500 | -3,800 | 0.0% | -45.9% | -45.8% | 0.1% |
| **Indiana** | 17,300 | 3,100 | 17.9% | 21,300 | 3,900 | 18.3% | 4,000 | 800 | 0.4% | 23.1% | 25.8% | 2.2% |
| **Iowa** | 7,900 | 1,000 | 12.7% | 4,200 | 500 | 11.9% | -3,700 | -500 | -0.8% | -46.8% | -50.0% | -6.0% |
| **Kansas** | 7,800 | 1,100 | 14.1% | 5,800 | 800 | 13.8% | -2,000 | -300 | -0.3% | -25.6% | -27.3% | -2.2% |
| **Kentucky** | 11,000 | 1,600 | 14.5% | 10,300 | 1,500 | 14.6% | -700 | -100 | 0.0% | -6.4% | -6.3% | 0.1% |
| **Louisiana** | 12,100 | 1,400 | 11.6% | 5,600 | 700 | 12.5% | -6,500 | -700 | 0.9% | -53.7% | -50.0% | 8.0% |
| **Maine** | 3,000 | 200 | 6.7% | 2,200 | 200 | 9.1% | -800 | 0 | 2.4% | -26.7% | 0.0% | 36.4% |
| **Maryland** | 14,600 | 2,800 | 19.2% | 12,800 | 2,500 | 19.5% | -1,800 | -300 | 0.4% | -12.3% | -10.7% | 1.8% |
| **Massachusetts** | 15,300 | 2,500 | 16.3% | 14,400 | 2,400 | 16.7% | -900 | -100 | 0.3% | -5.9% | -4.0% | 2.0% |
| **Michigan** | 25,300 | 4,200 | 16.6% | 24,400 | 4,100 | 16.8% | -900 | -100 | 0.2% | -3.6% | -2.4% | 1.2% |
| **Minnesota** | 13,900 | 2,100 | 15.1% | 13,500 | 2,100 | 15.6% | -400 | 0 | 0.4% | -2.9% | 0.0% | 3.0% |
| **Mississippi** | 8,200 | 800 | 9.8% | 9,100 | 900 | 9.9% | 900 | 100 | 0.1% | 11.0% | 12.5% | 1.4% |
| **Missouri** | 15,400 | 1,800 | 11.7% | 15,800 | 1,900 | 12.0% | 400 | 100 | 0.3% | 2.6% | 5.6% | 2.9% |
| **Montana** | 2,400 | 200 | 8.3% | 900 | 100 | 11.1% | -1,500 | -100 | 2.8% | -62.5% | -50.0% | 33.3% |
| **Nebraska** | 5,000 | 600 | 12.0% | 3,800 | 500 | 13.2% | -1,200 | -100 | 1.2% | -24.0% | -16.7% | 9.6% |
| **Nevada** | 7,200 | 1,400 | 19.4% | 6,900 | 1,400 | 20.3% | -300 | 0 | 0.8% | -4.2% | 0.0% | 4.3% |
| **New Hampshire** | 3,100 | 300 | 9.7% | 3,000 | 300 | 10.0% | -100 | 0 | 0.3% | -3.2% | 0.0% | 3.3% |
| **New Jersey** | 22,300 | 5,400 | 24.2% | 17,300 | 4,200 | 24.3% | -5,000 | -1,200 | 0.1% | -22.4% | -22.2% | 0.3% |
| **New Mexico** | 5,600 | 900 | 16.1% | 3,000 | 500 | 16.7% | -2,600 | -400 | 0.6% | -46.4% | -44.4% | 3.7% |
| **New York** | 46,700 | 11,800 | 25.3% | 53,600 | 13,500 | 25.2% | 6,900 | 1,700 | -0.1% | 14.8% | 14.4% | -0.3% |
| **North Carolina** | 24,600 | 3,200 | 13.0% | 23,100 | 3,000 | 13.0% | -1,500 | -200 | 0.0% | -6.1% | -6.3% | -0.2% |
| **North Dakota** | 1,600 | 100 | 6.3% | 1,600 | 100 | 6.3% | 0 | 0 | 0.0% | 0.0% | 0.0% | 0.0% |
| **Ohio** | 29,500 | 5,000 | 16.9% | 36,100 | 6,200 | 17.2% | 6,600 | 1,200 | 0.2% | 22.4% | 24.0% | 1.3% |
| **Oklahoma** | 10,000 | 1,300 | 13.0% | 8,600 | 1,200 | 14.0% | -1,400 | -100 | 1.0% | -14.0% | -7.7% | 7.3% |
| **Oregon** | 9,300 | 1,300 | 14.0% | 8,500 | 1,200 | 14.1% | -800 | -100 | 0.1% | -8.6% | -7.7% | 1.0% |
| **Pennsylvania** | 30,100 | 6,000 | 19.9% | 31,600 | 6,300 | 19.9% | 1,500 | 300 | 0.0% | 5.0% | 5.0% | 0.0% |
| **Rhode Island** | 2,400 | 400 | 16.7% | 2,700 | 400 | 14.8% | 300 | 0 | -1.9% | 12.5% | 0.0% | -11.1% |
| **South Carolina** | 11,700 | 1,300 | 11.1% | 10,900 | 1,200 | 11.0% | -800 | -100 | -0.1% | -6.8% | -7.7% | -0.9% |
| **South Dakota** | 2,200 | 200 | 9.1% | 2,100 | 200 | 9.5% | -100 | 0 | 0.4% | -4.5% | 0.0% | 4.8% |
| **Tennessee** | 16,100 | 2,500 | 15.5% | 15,100 | 2,400 | 15.9% | -1,000 | -100 | 0.4% | -6.2% | -4.0% | 2.4% |
| **Texas** | 74,100 | 10,700 | 14.4% | 99,100 | 14,300 | 14.4% | 25,000 | 3,600 | 0.0% | 33.7% | 33.6% | -0.1% |
| **Utah** | 9,400 | 1,900 | 20.2% | 8,100 | 1,700 | 21.0% | -1,300 | -200 | 0.8% | -13.8% | -10.5% | 3.8% |
| **Vermont** | 1,400 | 100 | 7.1% | 1,300 | 100 | 7.7% | -100 | 0 | 0.5% | -7.1% | 0.0% | 7.7% |
| **Virginia** | 20,000 | 3,400 | 17.0% | 18,700 | 3,200 | 17.1% | -1,300 | -200 | 0.1% | -6.5% | -5.9% | 0.7% |
| **Washington** | 17,100 | 3,000 | 17.5% | 9,600 | 1,700 | 17.7% | -7,500 | -1,300 | 0.2% | -43.9% | -43.3% | 0.9% |
| **West Virginia** | 4,200 | 600 | 14.3% | 4,000 | 600 | 15.0% | -200 | 0 | 0.7% | -4.8% | 0.0% | 5.0% |
| **Wisconsin** | 14,400 | 2,100 | 14.6% | 14,700 | 2,200 | 15.0% | 300 | 100 | 0.4% | 2.1% | 4.8% | 2.6% |
| **Wyoming** | 1,500 | 100 | 6.7% | 1,400 | 100 | 7.1% | -100 | 0 | 0.5% | -6.7% | 0.0% | 7.1% |

*Figure S1: Distribution of NO2 concentrations (ug/m3) by living location*

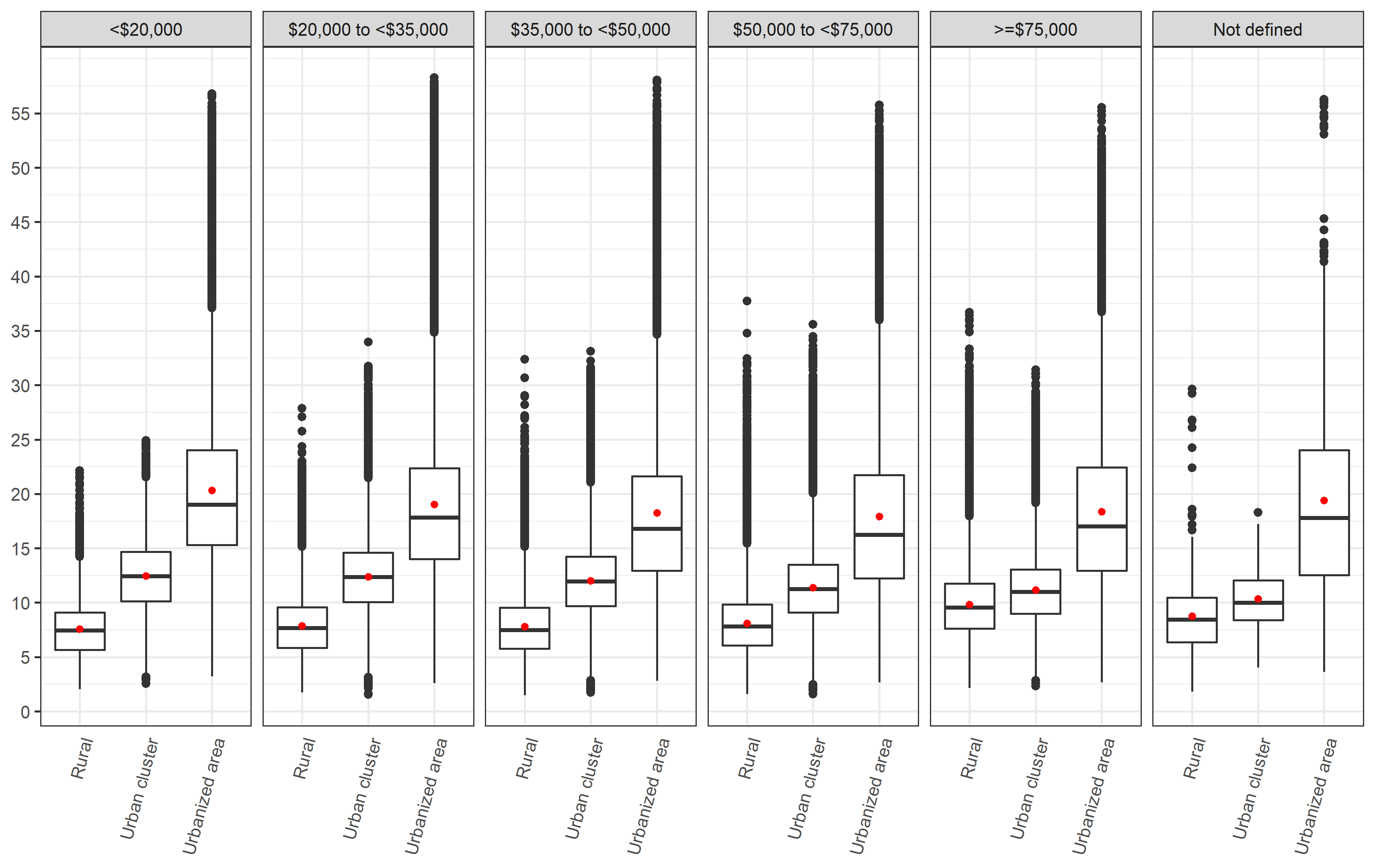


*\*Red dot represents the mean value while the midline represents the median value across all census blocks*

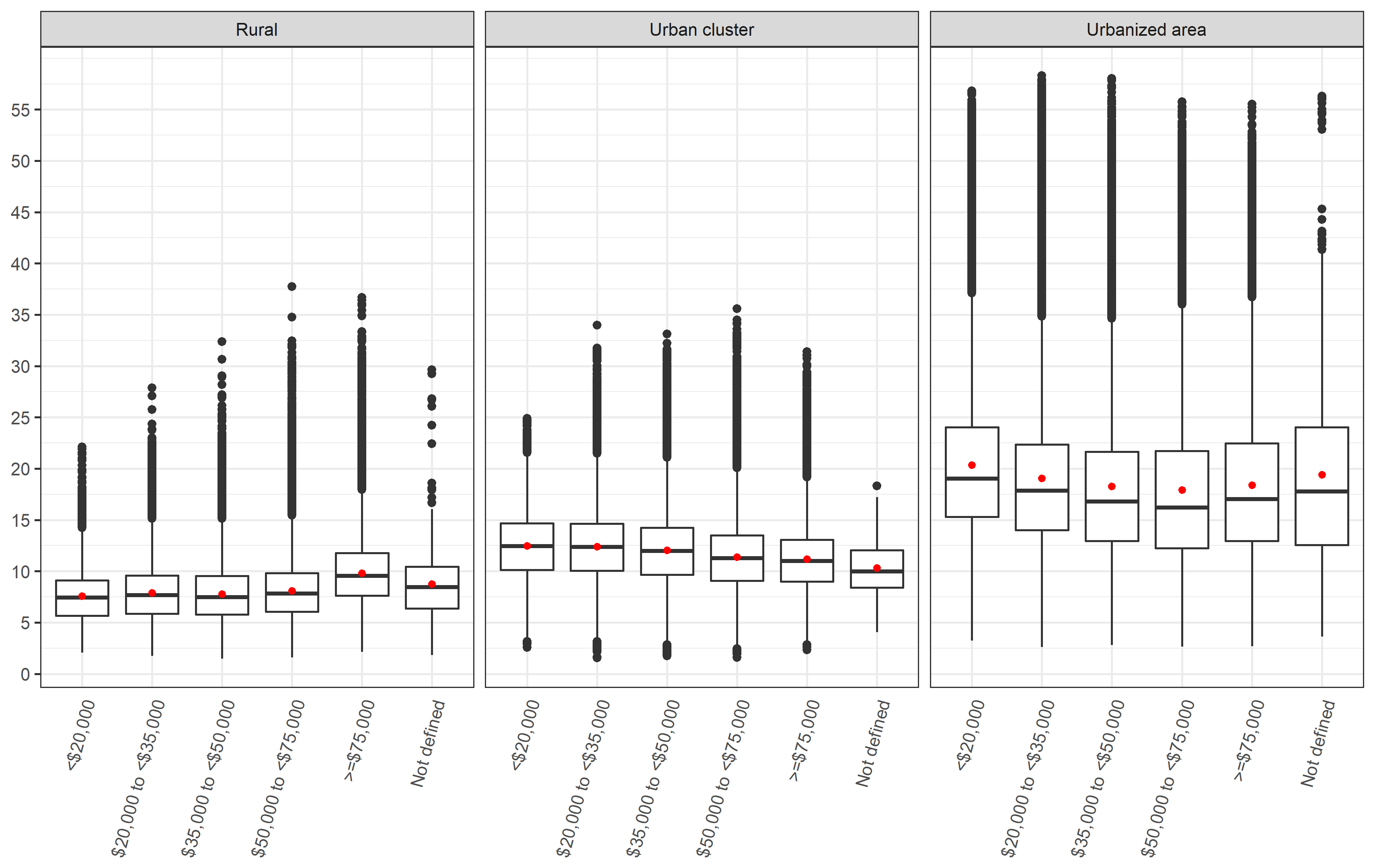
*Figure S2: Distribution of NO2 concentrations (ug/m3) by median household income group*



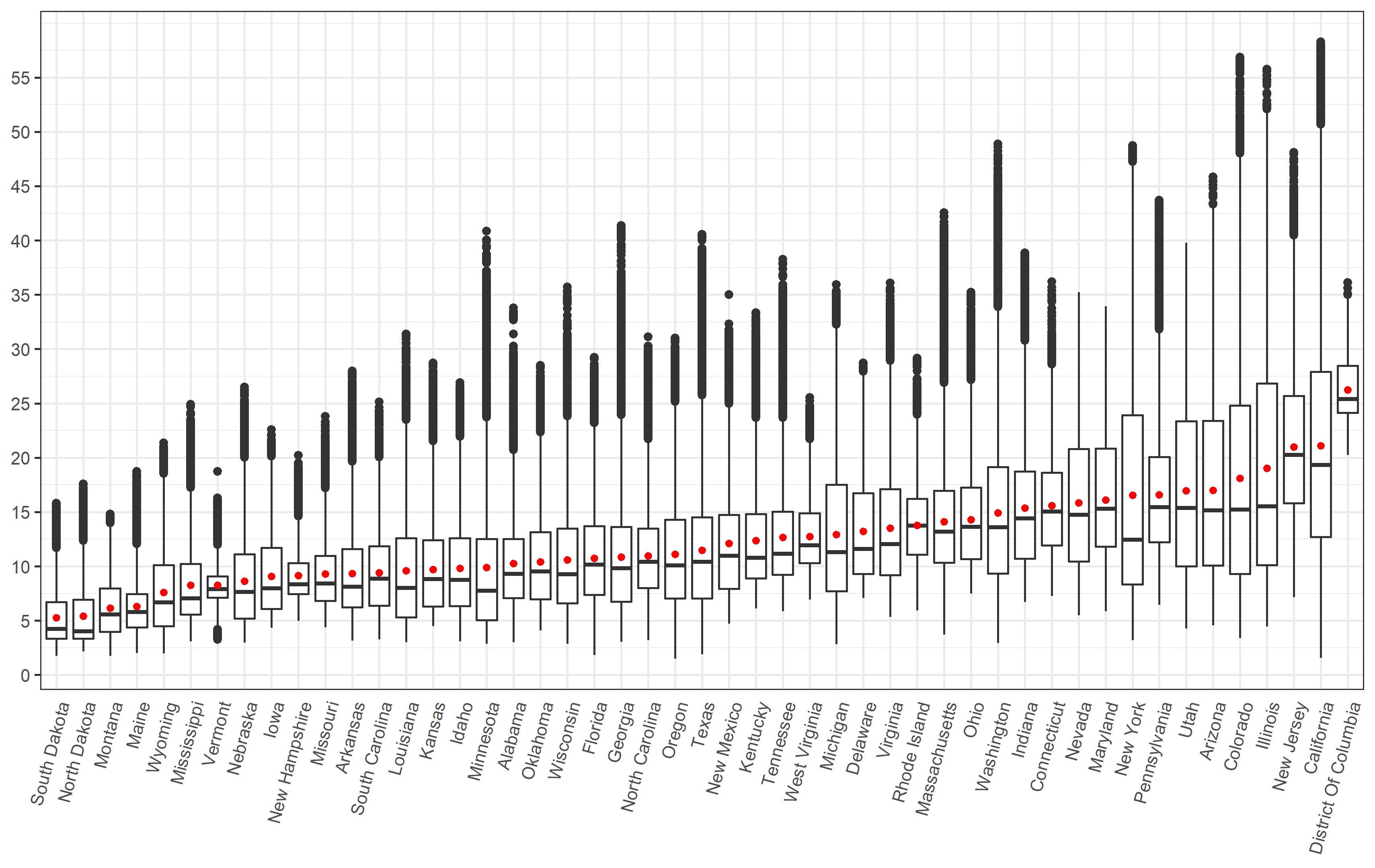
*Figure S3: Distribution of NO2 concentrations (ug/m3) by living location stratified into median household income group*



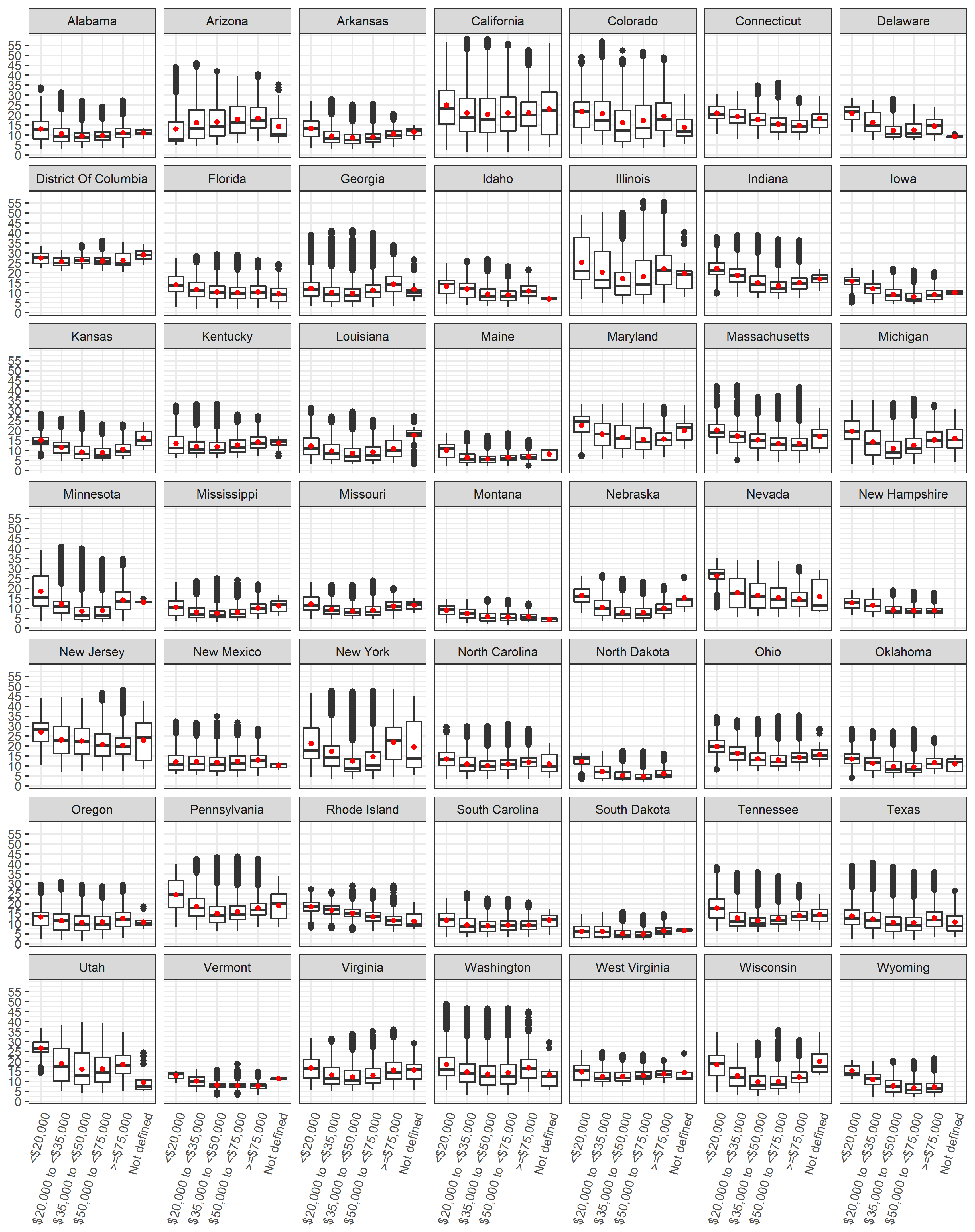
*Figure S4: Distribution of NO2 concentrations (ug/m3) by median household income group stratified into living location*



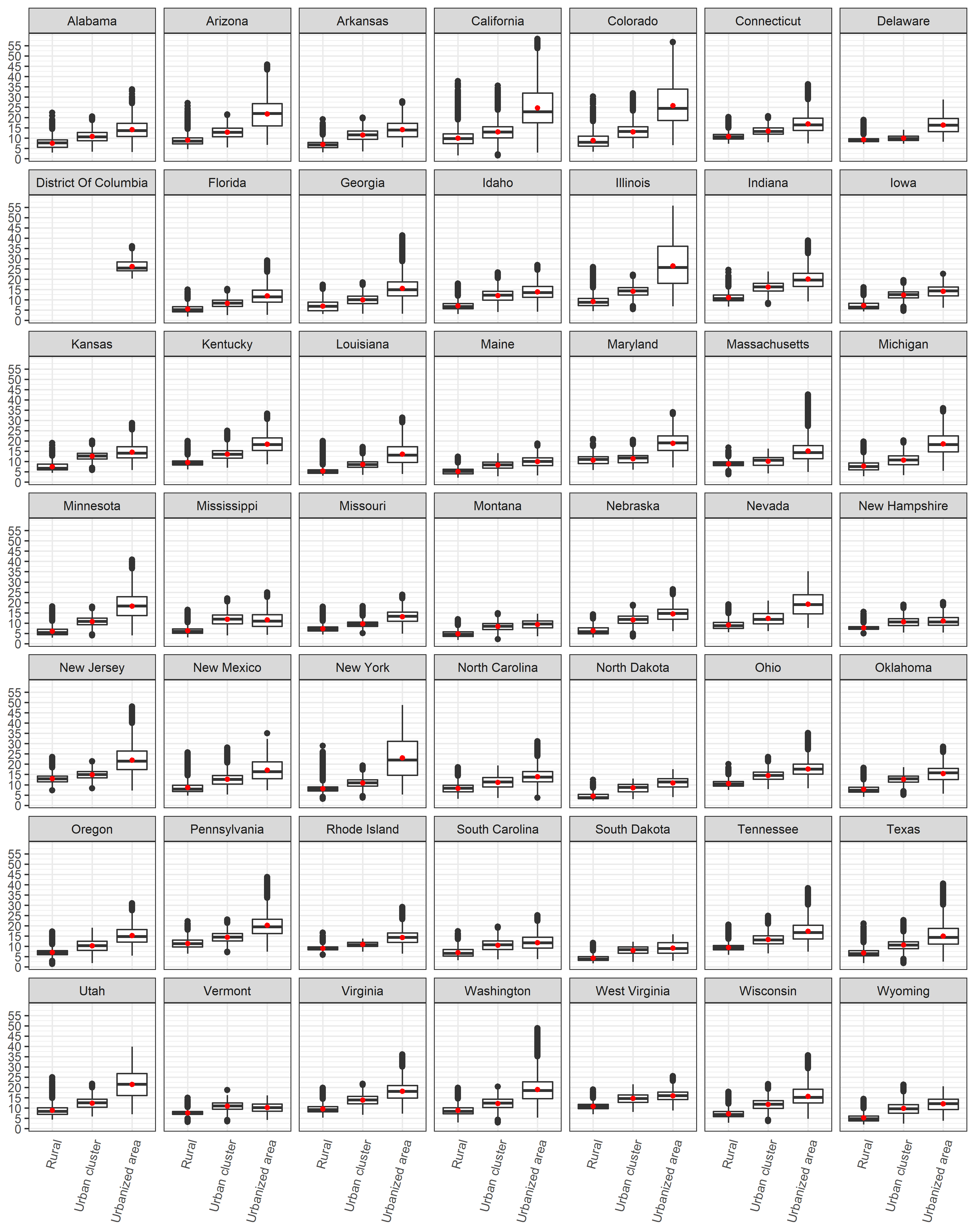
*Figure S5: Distribution of NO2 concentrations (ug/m3) by state*



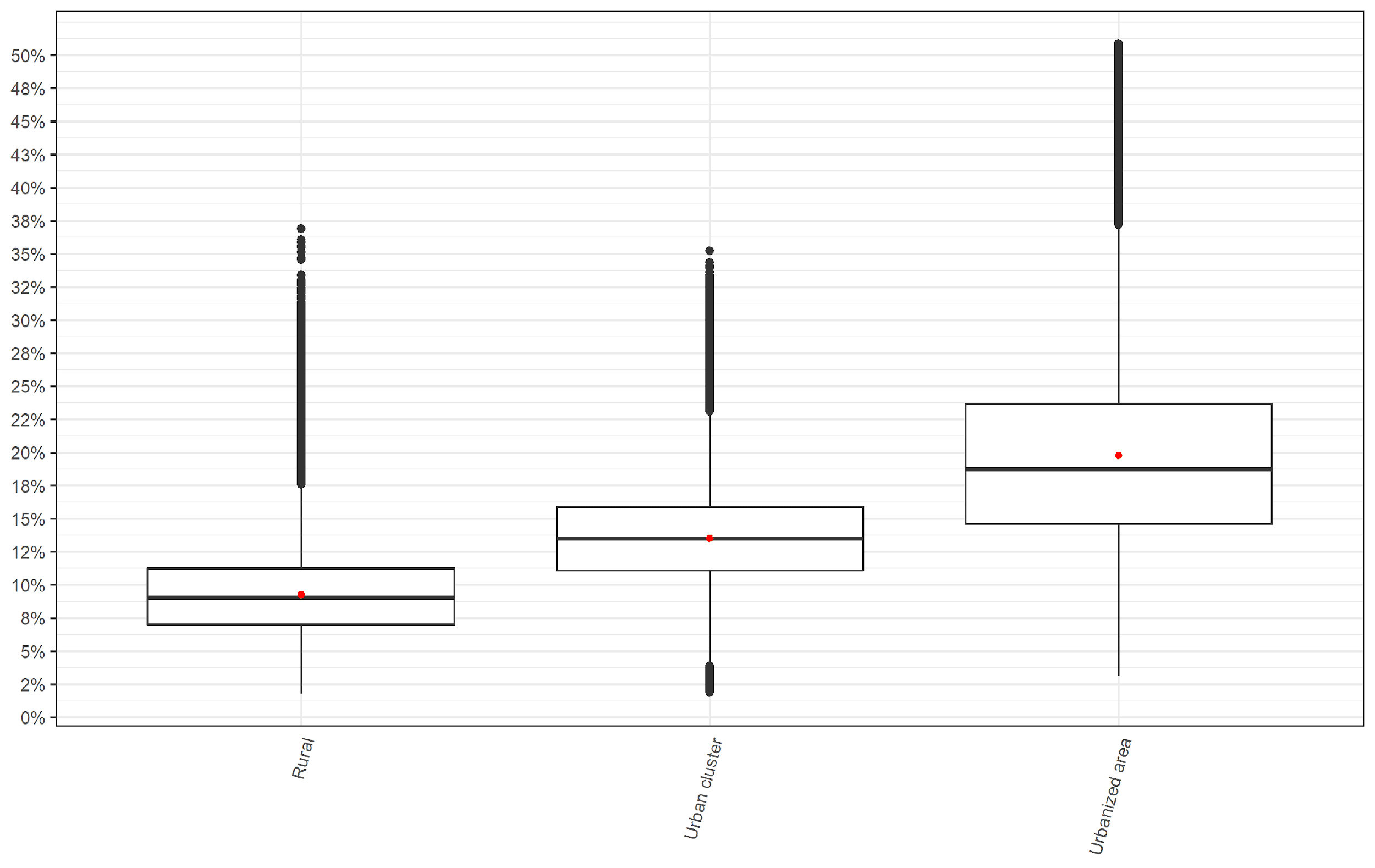
*Figure S6: Distribution of NO2 concentrations (ug/m3) by state and median household income group*



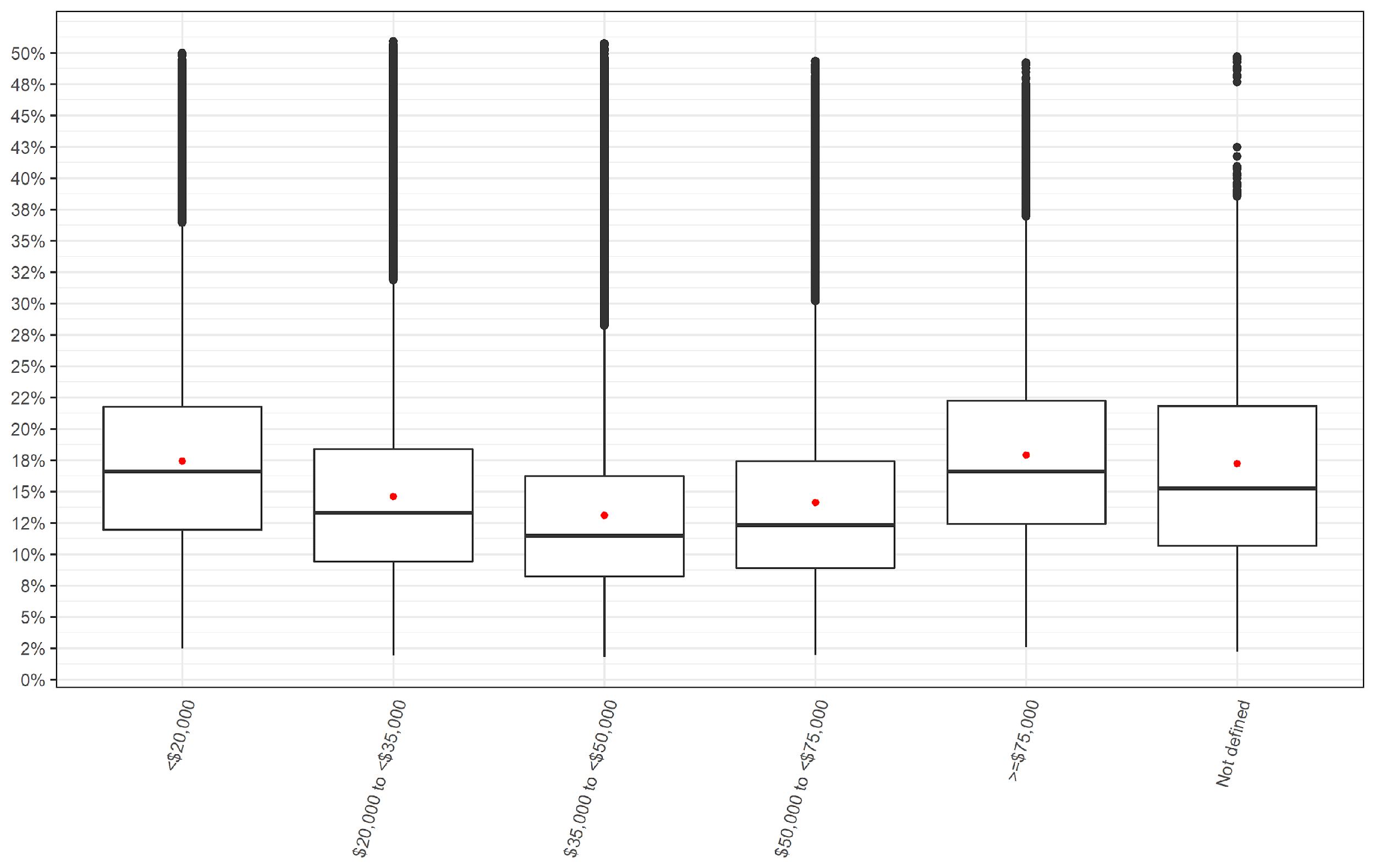
*Figure S7: Distribution of NO2 concentrations (ug/m3) by state and living location*



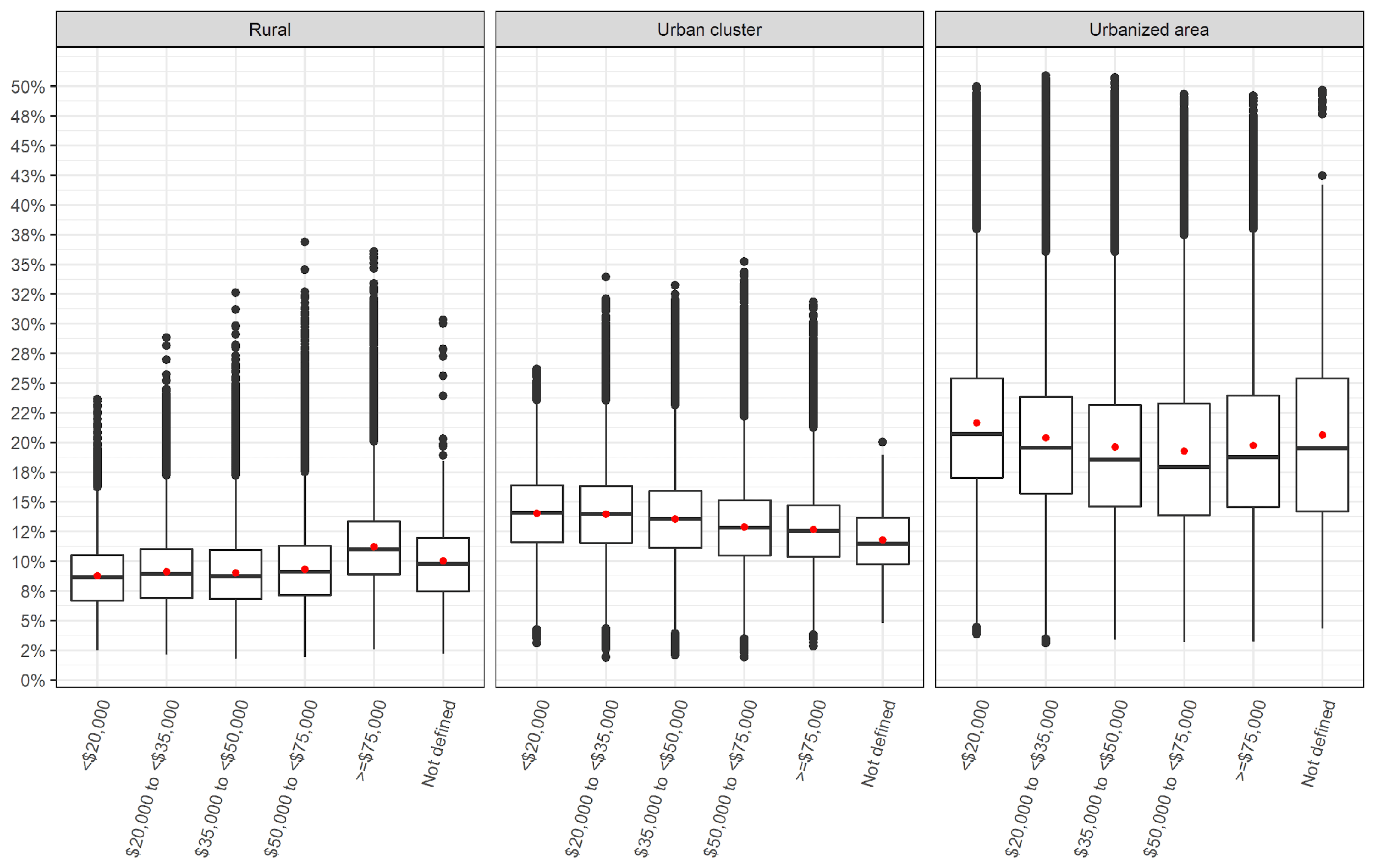
*Figure S8: Distribution of attributable fractions by living location*

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*Figure S9: Distribution of attributable fractions by median household income group*

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*Figure S10: Distribution of attributable fractions by median household income group stratified into living location*

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