Probability and Random Processes (15B11MA301)

Lecture-12



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Moment Generating Function

- The moment generating function of a random variable X denoted by $M_X(t)$ is defined as
 - $M_X(t) = E[e^{tX}]$ where t is a real variable.
- If X is a discrete random variable with PMF p(x), then $M_X(t) = E[e^{tX}] = \sum_x e^{tx} p(x)$
- If X is a continuous random variable with PDF f(x), then

$$M_X(t) = E[e^{tX}] = \int_{-\infty}^{\infty} e^{tx} f(x) dx$$

The coefficient $\frac{t'}{r!}$ in $M_X(t)$ is μ'_r , r = 1, 2, 3... and $\mu'_r = E[X^r]$ gives moments about the origin.

Proof: We know that
$$M_X(t) = E[e^{tX}]$$

$$= E\left[1 + \frac{tX}{1!} + \frac{(tX)^2}{2!} + \cdots\right]$$

$$= E(1) + \frac{t}{1!}E[X] + \frac{t^2}{2!}E[X^2] + \cdots + \frac{t^r}{r!}E[X^r] + \cdots$$

$$= 1 + \frac{t}{1!}\mu_1' + \frac{t^2}{2!}\mu_2' + \cdots + \frac{t^r}{r!}\mu_r' + \cdots$$
Hence $M_Y(t) = \sum_{i=1}^{\infty} \frac{t^r}{i!}\mu_1'$

Hence, $M_X(t) = \sum_{r=0}^{\infty} \frac{t^r}{r!} \mu'_r$(1)

this gives the MGF in terms of the moments.

The moments μ'_r can also be obtained as

Differentiating equation (1) with respect to t, r times and putting t = 0 provides moments

$$\mu'_r = \left[\frac{d^r}{dt^r} M_X(t)\right]_{t=0}, r=1, 2, 3...$$
 (2)

 $M_{aX}(t) = M_X(at)$, a being a constant.

Proof: By definition,
$$M_{aX}(t) = E[e^{taX}] = E[e^{(at)X}]$$

 $M_{aX}(t) = M_X(at)$

If Y = aX +b, then $M_Y(t) = e^{bt} M_X(at)$.

Proof: We know that,

$$M_Y(t) = E[e^{tY}]$$

$$= E[e^{t(aX+b)}]$$

$$= E[e^{t(aX)}]E[e^{tb}]$$

$$= e^{bt} E[e^{(ta)X}] = e^{bt} M_X(at).$$

The moment generating function of the sum of n independent random variables is equal to the product of their respective moment generating functions, i.e.

$$M_{X_1+X_2+\cdots+X_n}(t) = M_{X_1}(t) \cdot M_{X_2}(t) \cdot \dots \cdot M_{X_n}(t)$$

Proof: Using the definition of MGF, we have

$$M_{X_1+X_2+\cdots+X_n}(t) = E\left[e^{t(X_1+X_2+\cdots+X_n)}\right]$$

$$= E\left[e^{tX_1}\right]E\left[e^{tX_2}\right]E\left[e^{tX_3}\right]\dots E\left[e^{tX_n}\right] \text{ (since variables are independent)}$$

Therefore,

$$M_{X_1+X_2+\cdots+X_n}(t) = M_{X_1}(t) \cdot M_{X_2}(t) \cdot ... \cdot M_{X_n}(t)$$

Effect of Origin and Scale on MGF

Let the random variable X be transformed to a new variable U by changing both the origin and scale in X as $U = \frac{X-a}{h}$ where a and h are constants.

Then, the MGF of U (about origin) is given by

$$M_{U}(t) = E[e^{tU}]$$

$$= E\left[e^{t\left(\frac{X-a}{h}\right)}\right]$$

$$= e^{\left(\frac{-at}{h}\right)} E\left[e^{t\left(\frac{X}{h}\right)}\right]$$

$$= e^{\left(\frac{-at}{h}\right)} M_{X}(t/h)$$

Limitations of Moment Generating Function

4 A random variable X may have no moments although its moment generating function exists.

For example:
$$f(x) = \{\frac{1}{x(x+1)}, x = 1, 2, 3 ... \text{ and } f(x) = 0 \text{ otherwise.} \}$$

→ A random variable X can have MGF and some or all moments, yet the MGF does not generate the moments.

For example:
$$P(X = \pm 2^x) = \frac{e^{-1}}{x!}$$
, x=0,1,2,...

♣ A random variable X can have all or some moments, but MGF does not exist, except perhaps at one point.

For example:
$$P(X = \pm 2^x) = \frac{e^{-1}}{2x!}$$
, x=0,1,2,... and $P(X = \pm 2^x) = 0$, otherwise

Example 1: If a random variable X has the MGF $M_X(t) = \frac{3}{3-t}$, obtain the standard deviation of X.

Solution:
$$M_X(t) = \frac{3}{3-t} = 1 + \frac{t}{3} + \frac{t^2}{9} + \dots$$

$$E(X) = \text{coefficient of } \frac{t}{1!} = 1/3$$

$$E(X^2)$$
 = coefficient of $\frac{t^2}{2!}$ = 2/9

$$Var(X) = E(X^{2}) - (E(X))^{2}$$
$$= \frac{2}{9} - \frac{1}{9} = \frac{1}{9}$$

Standard Deviation = $\sigma_X = 1/3$

Example 2: Find the MGF of the random variable X whose probability function $(X = x) = \frac{1}{2^x}$, x = 1, 2, 3... hence, find its mean.

Solution:
$$M_X(t) = E(e^{tX}) = \sum_{x=0}^{\infty} e^{tx} P(X = x)$$

$$= \sum_{x=1}^{\infty} e^{tx} \frac{1}{2^x} = \sum_{x=1}^{\infty} \left(\frac{e^t}{2}\right)^x$$

On expanding the above summation, we get

$$M_X(t) = \frac{e^t}{2} \left(\frac{2}{2-e^t}\right) = \frac{e^t}{2-e^t}$$

Mean =
$$\mu'_1 = \frac{d}{dt} M_X(t)$$
 at t =0.

$$=\frac{\mathrm{d}}{\mathrm{dt}}\left(\frac{e^t}{2-e^t}\right)$$
 at $t=0$

$$Mean = 2$$

Example 3: A random variable X has the PDF given by $f(x) = \{2e^{-2x}, x \ge 0 \text{ and } 0 \text{ if } x < 0.$

Find (i) MGF and

(ii) the first four moments of X about the origin.

Solution:
$$M_X(t) = E[e^{tX}] = \int_{-\infty}^{\infty} e^{tx} f(x) dx$$

$$M_X(t) = \int_0^\infty |e^{tx} e^{-2x} dx = \left[\frac{e^{-(2-t)x}}{-(2-t)}\right] \text{ from } 0 \text{ to } \infty$$

Therefore,
$$M_X(t) = \frac{2}{2-t}$$

We know that,
$$M_X(t) = \sum_{r=0}^{\infty} \frac{t^r}{r!} \mu'_r = \frac{2}{2-t} = \frac{2}{2(1-\frac{t}{2})}$$

$$1 + \frac{t}{1!} \mu'_1 + \frac{t^2}{2!} \mu'_2 + \dots + \frac{t^r}{r!} \mu'_r + \dots = \left(1 - \frac{t}{2}\right)^{-1}$$

$$= 1 + \frac{t}{2} + \frac{t^2}{2^2} + \frac{t^3}{2^3} + \frac{t^4}{2^4} + \dots$$

$$= 1 + \frac{1}{2} \frac{t}{1!} + \frac{2!}{4} \frac{t^2}{2!} + \dots$$

On equating the coefficients of $\frac{t}{1!}$, $\frac{t^2}{2!}$, and so on, we have..

$$\mu_{1}^{'}=1/2$$
, $\mu_{2}^{'}=\frac{1}{2}$, $\mu_{3}^{'}=\frac{3}{4}$, $\mu_{2}^{'}=3/2$

Practice Questions

- 1. A random variable X has the density function given by $f(x) = \begin{cases} \frac{1}{k}, 0 < x < k \\ 0, otherwise \end{cases}$. Find (i) MGF, (ii) rth moment, (iii) Mean and (iv) Variance
 - [Ans. (i) $\frac{(e^{tk}-1)}{kt}$ (ii) $\frac{k^r}{(r+1)!}$ (iii) $\frac{k}{2}$ (iv) $\frac{k^2}{12}$]
- 2. Let X be a random variable with PDF $f(x) = \begin{cases} \frac{e^{\frac{-x}{3}}}{3}, 0 < x \\ 0, otherwise \end{cases}$.
 - Find (i) P(X>3) (ii) MGF of X, (iii) E(X) and Var(X).
 - [Ans. (i) 1/e (ii) $(1-3t)^{-1}$ (iii) E(X) = 3, Var(X) = 9]

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