# Covariance Matching Kalman Filter for observable LTI Systems

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## Overview

#### Introduction

Motivation Literature Review

#### Derivation of the Filter

Stacked Dynamics
Time Series Formulation
Covariance Estimation

#### Outline of Proof

Convergence of Covariance Convergence of State Error Covariance

#### Results

Simulation Problem Matrix Estimation Error State Estimation Error



#### Motivation

#### System Description

$$x_k = Fx_{k-1} + w_{k-1}, \quad w_{k-1} \sim \mathcal{N}(0, Q), \ x_0 \sim \mathcal{N}(m_x, P_0)$$
  
 $y_k = Hx_k + v_k, \quad v_k \sim \mathcal{N}(0, R)$ 

Common issues with a linear Kalman Filter:

- Q and R matrices are inaccurately known
- F and H matrices are uncertain
- Either  $w_k$  or  $v_k$  are non Gaussian

Adaptive filters simultaneously estimate the state and the unknown quantities on-line.

#### Literature Review

#### Existing adaptive filters:

- Application specific algorithms
- Restrictive assumptions on F and H matrix
- Interdependence between the state and covariance estimate

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#### Important properties of our algorithm:

- Either one of the Q or R matrices are completely known
- Does not need the state estimate for covariance estimation
- Assumes observability of the LTI system
- We provide a proof of convergence

# Stacked Dynamics

Propagating the state dynamics in time,

$$x_{k+1} = Fx_k + w_k$$

$$x_{k+2} = F^2 x_k + Fw_k + w_{k+1}$$

$$x_{k+n-1} = F^{n-1} x_k + F^{n-2} w_k + \dots + w_{k+n-1}$$

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Predicting future measurements from the current state,

$$y_{k} = Hx_{k} + v_{k}$$

$$y_{k+1} = HFx_{k} + Hw_{k} + v_{k+1}$$

$$y_{k+2} = HF^{2}x_{k} + HFw_{k} + Hw_{k+1} + v_{k+2}$$

$$y_{k+n-1} = HF^{n-1}x_{k} + HF^{n-2}w_{k} + Hw_{k+n-2} + v_{k+n-1}$$

# Coagulating all the measurements

The observability matrix  $M_o$ , formed below has full column rank.

$$\begin{bmatrix}
y_{k+n-1} \\
y_{k+n-2} \\
\vdots \\
y_k
\end{bmatrix} = \begin{bmatrix}
HF^{n-1} \\
HF^{n-2} \\
\vdots \\
H
\end{bmatrix} x_k + \begin{bmatrix}
H & HF & HF^2 & \cdots & HF^{n-2} \\
0 & H & HF & \cdots & HF^{n-3} \\
\vdots & \vdots & \ddots & \vdots & \vdots \\
0 & 0 & 0 & 0 & H \\
0 & 0 & 0 & 0 & 0
\end{bmatrix} \begin{bmatrix}
w_{k+n-2} \\
w_{k+n-3} \\
\vdots \\
w_k
\end{bmatrix} \\
\triangleq W_k$$

$$+ \underbrace{\begin{bmatrix}
v_{k+n-1}^T & v_{k+n-2}^T & \cdots & v_k^T\end{bmatrix}^T}_{\triangleq V_k}$$

Propagating the stacked measurements in time,

$$\mathcal{Y}_k = M_o x_k + M_w W_k + V_k$$

Propagating the stacked measurements in time,

$$y_k = M_o x_k + M_w W_k + V_k y_{k+1} = M_o x_{k+1} + M_w W_{k+1} + V_{k+1}$$

Propagating the stacked measurements in time,

$$y_k = M_o x_k + M_w W_k + V_k$$
$$y_{k+1} = M_o x_{k+1} + M_w W_{k+1} + V_{k+1}$$

Eliminating the state,

$$FM_o^{\dagger} \mathcal{Y}_k = Fx_k + FM_o^{\dagger} M_w W_k + FM_o^{\dagger} V_k$$
  

$$M_o^{\dagger} \mathcal{Y}_{k+1} = Fx_k + w_k + M_o^{\dagger} M_w W_{k+1} + M_o^{\dagger} V_{k+1}$$

Propagating the stacked measurements in time,

$$\mathcal{Y}_{k} = M_{o}x_{k} + M_{w}W_{k} + V_{k}$$
  
 $\mathcal{Y}_{k+1} = M_{o}x_{k+1} + M_{w}W_{k+1} + V_{k+1}$ 

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State independent linear strictly stationary time series,

$$\underbrace{\frac{M_o^{\dagger} \mathcal{Y}_{k+1} - F M_o^{\dagger} \mathcal{Y}_k}_{\triangleq \mathcal{Z}_k}}_{\triangleq \mathcal{Z}_k} = \underbrace{\frac{w_k + M_o^{\dagger} M_w W_{k+1} - F M_o^{\dagger} M_w W_k}_{\triangleq \mathcal{W}_k} + \underbrace{M_o^{\dagger} V_{k+1} - F M_o^{\dagger} V_k}_{\triangleq \mathcal{Y}_k}$$

#### Covariance Estimation

 $\mathcal{Z}_k$  is zero mean and has a constant covariance

$$\mathcal{Z}_k = N_0 y_k + N_1 y_{k+1} + \dots + N_n y_{k+n} = A_1 w_k + A_2 w_{k+1} + \dots + A_{n-1} w_{k+n-1} + B_0 v_k + B_1 v_{k+1} + \dots + B_n v_{k+n}$$

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#### Covariance calculation

$$Cov(\mathcal{Z}_{k}) = A_{1}QA_{1}^{T} + \dots + A_{n-1}QA_{n-1}^{T} + B_{0}RB_{0}^{T} + \dots + B_{n}RB_{n}^{T}$$

$$Cov(\mathcal{Z})_{i} = \sum_{k=1}^{k=i} \frac{1}{k} \bar{\mathcal{Z}}_{k} \bar{\mathcal{Z}}_{k}^{T}$$

$$Cov(\mathcal{Z})_{i} - A_{1}QA_{1}^{T} - \dots - A_{n-1}QA_{n-1}^{T} = B_{0}\hat{R}_{i}B_{0}^{T} + \dots + B_{n}\hat{R}_{i}B_{n}^{T}$$

# Q or R Estimation

Vectorizing the equation,

$$vec(C_k) = \underbrace{(B_0 \otimes B_0 + \cdots + B_n \otimes B_n)}_{\triangleq S} vec(\hat{R}_k)$$
$$vec(D_k) = \underbrace{(A_1 \otimes A_1 + \cdots + A_n \otimes A_n)}_{\triangleq T} vec(\hat{Q}_k)$$

Repetitive elements of the symmetric matrix and corresponding columns of S and T matrix are averaged out.

#### Outline of Proof

Time Series

$$\mathcal{Z}_k = N_0 y_k + N_1 y_{k+1} + \dots + N_n y_{k+n} = A_1 w_k + A_2 w_{k+1} + \dots + A_{n-1} w_{k+n-1} + B_0 v_k + B_1 v_{k+1} + \dots + B_n v_{k+n}$$

- We prove that  $Cov(\mathcal{Z})_k \xrightarrow{\mathcal{P}} Cov(\mathcal{Z})$  using the CLT for time series with decaying autocovariance
- Then,  $\hat{R}_k \xrightarrow{\mathcal{P}} R$  follows from the full rank assumption

# State Error Covariance Convergence

• Formulate the following 3 Covariance matrices:

$$\hat{P}_{k+1} = \hat{\bar{F}}_k \hat{P}_k \hat{\bar{F}}_k^T + \hat{K}_k \hat{R}_k \hat{K}_k^T + Q$$

$$P_{k+1} = \hat{\bar{F}}_k P_k \hat{\bar{F}}_k^T + \hat{K}_k R \hat{K}_k^T + Q$$

$$P_{k+1,opt} = \bar{F}_k P_{k,opt} \bar{F}_k^T + K_k R K_k^T + Q$$

- We prove that  $\hat{P}_k \stackrel{\mathcal{P}}{\longrightarrow} P_k$
- We prove that  $P_k \xrightarrow{\mathcal{P}} P_{k,opt}$

#### Simulation Problem

A marginally stable system with T = 0.05,

$$x_{k} = \begin{bmatrix} 1 & T & 0.5T^{2} \\ 0 & 1 & T \\ 0 & 0 & 1 \end{bmatrix} x_{k-1} + w_{k-1}$$
$$y_{k} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \end{bmatrix} x_{k} + v_{k}$$

We use the same example to estimate R and then Q

$$R = \begin{bmatrix} 0.25 & 0 \\ 0 & 0.25 \end{bmatrix} \text{ and } Q = \begin{bmatrix} 0.25 & 0.04 & 0.04 \\ 0.04 & 0.25 & 0.04 \\ 0.04 & 0.04 & 0.25 \end{bmatrix}$$

#### Simulation for covariance matrices

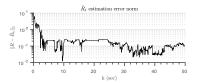


Figure: *R* estimate vs. time

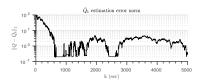


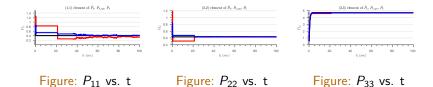
Figure: Q estimate vs. time

The Q matrix was assumed to be a constant

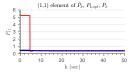
The *R* matrix was assumed to be a constant

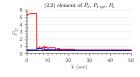
In both the cases, the <u>initial</u> estimate of the covariance was 5I with suitable dimensions.

# Simulation for state estimation (unknown R case)



# Simulation for state estimation (unknown Q case)





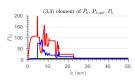


Figure:  $P_{11}$  vs. t

Figure:  $P_{22}$  vs. t

Figure:  $P_{33}$  vs. t

#### Discussion

- Linearly dependent measurements (non-invertible S matrix) introduce ambiguity in R matrix estimation
- Methods like linear matrix inequalities and weighting the initial estimate can be used for covariance estimation
- The rate of convergence depends on difference in order of magnitude of the known and the unknown covariance

$$Cov(\mathcal{Z}_k) = (\text{terms with R}) + (\text{terms with Q})$$