

# COMP3206: Exercises on Matrix calculus for optimisation

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## I Learning the parameters of a Gaussian by Maximum Likelihood Estimation (MLE)

**Estimation:** Data  $\mathcal{X} = \{\mathbf{x}^1, \dots, \mathbf{x}^n, \dots, \mathbf{x}^N\}$  Find mean vector  $\boldsymbol{\mu}$  and variance-covariance matrix  $\boldsymbol{\Sigma}$ .

Assuming data are drawn i.i.d. from Gaussian  $\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Lambda})$ ,  $\boldsymbol{\Lambda} \triangleq \boldsymbol{\Sigma}^{-1}$ , the log likelihood  $\mathcal{L}(\boldsymbol{\mu}, \boldsymbol{\Lambda})$  is

$$\mathcal{L}(\boldsymbol{\mu}, \boldsymbol{\Lambda}) = - \sum_{n=1}^N \frac{1}{2} (\mathbf{x}^n - \boldsymbol{\mu})^T \boldsymbol{\Lambda} (\mathbf{x}^n - \boldsymbol{\mu}) + \frac{N}{2} \log \det(\boldsymbol{\Lambda}) + \text{constant}.$$

The exercises below will enable you to find the optimal mean and covariance matrix using maximum likelihood estimation.

- Optimal  $\boldsymbol{\mu}$ :  $\frac{\partial}{\partial \boldsymbol{\mu}} \mathcal{L}(\boldsymbol{\mu}, \boldsymbol{\Lambda}) = 0 = \sum_n \boldsymbol{\Lambda} (\mathbf{x}^n - \boldsymbol{\mu})$

$$\sum_n \boldsymbol{\Lambda} \mathbf{x}^n = \sum_n \boldsymbol{\Lambda} \boldsymbol{\mu} = N \boldsymbol{\Lambda} \boldsymbol{\mu} \Rightarrow \boldsymbol{\mu} = \frac{1}{N} \sum_n \mathbf{x}^n.$$

- Optimal  $\boldsymbol{\Lambda}$ :  $\frac{\partial}{\partial \boldsymbol{\Lambda}} \mathcal{L}(\boldsymbol{\mu}, \boldsymbol{\Lambda}) = 0$ .

$$\boldsymbol{\Sigma} = \boldsymbol{\Lambda}^{-1} = \frac{1}{N} \sum_{n=1}^N (\mathbf{x}^n - \boldsymbol{\mu})(\mathbf{x}^n - \boldsymbol{\mu})^T.$$

### I.I Exercises

- i. Verify that the log likelihood function for

$$p(\mathbf{x}^n | \boldsymbol{\mu}, \boldsymbol{\Sigma}) = (\sqrt{(2\pi)^p |\boldsymbol{\Sigma}|})^{-1} \exp \left( -\frac{1}{2} (\mathbf{x}^n - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x}^n - \boldsymbol{\mu}) \right)$$

is as shown above.

2. For  $p \times p$  matrices  $\mathbf{A}, \mathbf{B}$  with matrix elements  $(\mathbf{A})_{ij} = a_{ij}$  and  $(\mathbf{B})_{ij} = b_{ij}$ , show that  $\text{tr}(\mathbf{AB}) = \text{tr}(\mathbf{BA})$  by writing  $\text{tr}(\mathbf{A}) = \sum_i a_{ii}$  and the product of matrices as

$$(\mathbf{AB})_{ij} = \sum_k a_{ik} b_{kj}.$$

3. For  $n \times p$  matrix  $\mathbf{A}$  with matrix elements  $(\mathbf{A})_{ij} = a_{ij}$ , show that the sum of the squares of the matrix elements

$$\sum_{ij} a_{ij}^2 = \text{tr}(\mathbf{AA}^T), \text{ where tr is the matrix trace.}$$

4. For the Kronecker delta  $\delta_{ij}$  defined as

$$\delta_{ij} = \begin{cases} 1 & \text{for } i = j \\ 0 & \text{otherwise.} \end{cases}$$

show that

- (i)  $\sum_j a_{ij} \delta_{kj} = a_{ik},$
  - (ii) The diagonal elements of the product of matrices  $\mathbf{A}, \mathbf{B}$  is  $\sum_{jk} a_{ij} b_{jk} \delta_{ki},$
  - (iii) Trace  $\text{tr}(\mathbf{A}) = \sum_{ij} a_{ij} \delta_{ij},$
  - (iv) Make sure you grok  $\frac{\partial}{\partial x_i} x_j = \delta_{ij}.$
5. For  $p \times p$  matrix  $\mathbf{A}$  with matrix elements  $(\mathbf{A})_{ij} = a_{ij}$   $1 \leq i, j \leq p$  and vector  $\mathbf{x} = (x_1, \dots, x_p)^T$  the  $i$ -th element of vector  $(\mathbf{Ax})$  is  $(\mathbf{Ax})_i = \sum_{j=1}^p a_{ij} x_j$ . Show that  $\frac{\partial}{\partial \mathbf{x}} (\mathbf{Ax}) = \mathbf{A}^T$  by writing out the indices explicitly:

$$\left( \frac{\partial}{\partial \mathbf{x}} (\mathbf{Ax}) \right)_{ij} = \frac{\partial}{\partial x_i} (\mathbf{Ax})_j = \frac{\partial}{\partial x_i} \sum_{k=1}^p a_{jk} x_k.$$

6. Show, by writing out the matrix elements as above, that the gradient of the scalar quadratic form  $\mathbf{xAx}$  is  $\frac{\partial}{\partial \mathbf{x}} \mathbf{x}^T \mathbf{Ax} = (\mathbf{A} + \mathbf{A}^T) \mathbf{x}$ . *Hint:* the  $i$ -th matrix element of the gradient is

$$\frac{\partial}{\partial x_i} \left( \sum_{r,s=1}^p x_r a_{rs} x_s \right).$$

This should lead to the expression for the MLE of the mean.

7. The partial derivative of the quadratic form  $\mathbf{x}\mathbf{A}\mathbf{x}$  with respect to  $\mathbf{A}$  can be evaluated for each matrix element  $a_{ij}$ ,  $1 \leq i, j \leq p$ :

$$\frac{\partial}{\partial a_{ij}} \left( \sum_{r,s=1}^p x_r a_{rs} x_s \right).$$

Show that the result is  $\mathbf{x}\mathbf{x}^T$  (a  $p \times p$  matrix).

8. Remember that the determinant of a matrix can be written as a sum

$$\det(\mathbf{A}) = \sum_{j=1}^p a_{ij} \text{cof}(a_{ij})$$

where  $\text{cof}(a_{ij})$  is  $(-1)^{i+j}$  times the determinant of the submatrix of  $\mathbf{A}$  obtained by deleting the  $i$ -th row and  $j$ -th column. In particular,  $\text{cof}(a_{ij})$  does not contain  $a_{ij}$ . Show that

$$\left( \frac{\partial}{\partial \mathbf{A}} \ln \det \mathbf{A} \right)_{ij} = \frac{\partial}{\partial a_{ij}} \ln \left( \sum_{s=1}^p a_{rs} \text{cof}(a_{rs}) \right) = (\mathbf{A}^{-1})_{ij}.$$

This and the previous problem should help you derive the MLE of the covariance matrix.

## 2 Regularised linear regression

In regression problems, we have been minimising the residual sum of errors (RSS) with respect to the parameters  $\theta$  that are the weight vectors  $\mathbf{w}$  in

$$f(\mathbf{x}; \mathbf{w}) = w_0 + \sum_{i=1}^p w_i \phi_i(\mathbf{x}).$$

If we introduce  $x_0 = 1$ , we can write, for each data point  $(\mathbf{x}^n, y^n) = (1, x_1^n, \dots, x_p^n, y^n)$ , and the RSS is

$$\text{RSS} = \sum_{n=1}^N (r^n)^2 = \sum_{n=1}^N (y^n - f(\mathbf{x}^n; \mathbf{w}))^2.$$

When we introduce a  $L_2$  regularisation term  $\|\mathbf{w}\|_2 = \mathbf{w}^T \mathbf{w}$  for the weights  $\mathbf{w}$ , the minimisation is then over a loss function  $\ell(\mathbf{w})$ :

$$\ell(\mathbf{w}) = -\mathcal{L}(\mathbf{w}) = \sum_{n=1}^N (y^n - f(\mathbf{x}^n; \mathbf{w}))^2 + \lambda \mathbf{w}^T \mathbf{w},$$

where  $\lambda$  controls the trade-off between where the data wants the learnt functions to go and how small the modeller wants to keep  $\|\mathbf{w}\|_2$ . Minimising  $\ell(\mathbf{w})$  is equivalent to maximising  $\mathcal{L}(\mathbf{w})$ . This  $L_2$  regularised version of linear regression is called **ridge regression**.

## 2.1 Exercises

1. Take the gradient of the loss function  $\ell(\mathbf{w})$  with respect to the weight vector  $\mathbf{w}$  and set it equal to zero. For each vector component  $w_i$  compute

$$\frac{\partial}{\partial w_i} \sum_{n=1}^N \left( y^n - \sum_{j=0}^p w_j \phi_j(\mathbf{x}^n) \right) \left( y^n - \sum_{k=0}^p w_k \phi_k(\mathbf{x}^n) \right) + \lambda \sum_{j,k=0}^p w_j w_k \delta_{jk}.$$

Show that the derivative reduces to  $-2$  multiplied by

$$\sum_{n=1}^N \left\{ \phi_i(\mathbf{x}^n) y^n - \sum_{j=0}^p (w_j + \lambda \delta_{ij}) \phi_j(\mathbf{x}^n) \phi_i(\mathbf{x}^n) \right\},$$

a quantity that we will set to zero for max/minimisation.

2. Keep in mind that the data index  $n$  in the superscript is a row index while the  $i, j, k$  indices for weights stand for columns. Introduce the matrix  $\Phi$  with matrix elements  $(\Phi)_{nj} = \phi_j(\mathbf{x}^n)$ . Also, the column vector of  $y$  values is  $\mathbf{y}$ . Use this to rewrite the above as a matrix equation

$$\Phi^T \mathbf{y} = (\Phi^T \Phi + \lambda \mathbb{I}) \mathbf{w} \Rightarrow \mathbf{w} = (\Phi^T \Phi + \lambda \mathbb{I})^{-1} \Phi^T \mathbf{y}.$$

3. Identify the negative of the loss function  $\ell(\mathbf{w})$  with the quantity you worked with for the maximum likelihood estimation problem for the Gaussian. Think about the correspondence. The  $\lambda \|\mathbf{w}\|_2^2$  term becomes a prior distribution on weights in a Bayesian interpretation.