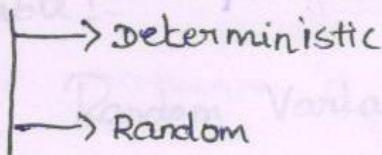


## UNIT - II ~ RANDOM VARIABLES

**Experiment :-**

→ It is an action which we conceive & do.

**Experiment**



**Deterministic Experiment :-**

An Experiment in which the outcomes can be predicted with certainty is called deterministic.

**Random Experiment :-**

It is an experiment in which the outcome cannot be predicted with certainty but all possible outcomes can be determined prior to performance of the experiment.

**Ex]**

1. Tossing a coin
2. Throwing a die
3. Number of phone calls in an hour.

Sample Space :-

(The collection of all possible outcomes.)

**Outcomes**

**Ex]**

1. In coin tossing Experiment
2. In die throw Experiment

$S = \{ \text{head, tail} \}$

$S = \{ 1, 2, 3, 4, 5, 6 \}$

**Event :-**

A part of sample space.

**Probability :-**

It is a chance value of something which will happen.

**Note :-**

Let  $S \rightarrow$  sample space

$A \rightarrow$  Event

- i)  $0 \leq P(A) \leq 1$
- ii)  $P(\emptyset) = 0$

Note:- The outcomes of a random Experiment to which our result

may be numerical or non-numerical in nature.

Ex]

The Number of telephone calls in an other

than the PPF is numerical but the result

a coin Experiment is non-numerical in

Nature.

To assign a number each non-numerical

outcome of a experiment we introduce the

concept of Random variable.

Random Variable:-

A Random Variable is a real

valued function defined on sample space.

Ex]

1) Consider a two coin toss experiment

S: H H T H T T

$x \rightarrow$  No. of tails

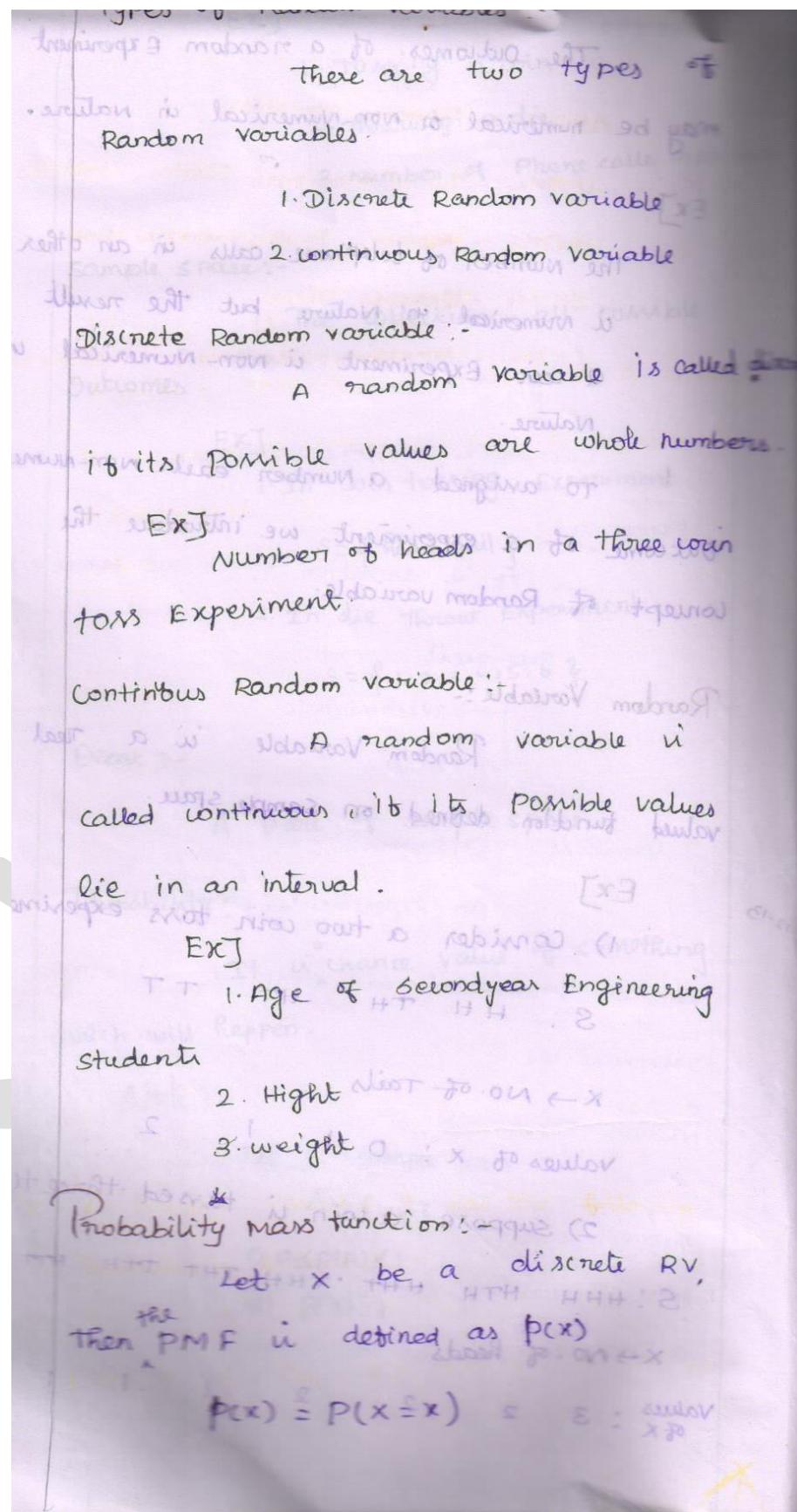
values of  $x$ : 0 1 2

2) Suppose a coin is tossed three times

S: HHH HTH HTT THH THT TTH HTT TT

$x \rightarrow$  No. of heads

values of  $x$ : 3 2 1 0



$p(x) \geq 0; \sum p(x) = 1$

Probability density Function: At a point  $x$ , the probability of finding a continuous RV,  $X$ , between  $x$  and  $x+dx$  is given by  $f(x)dx$ .

then the PDF is defined as  $f(x)$

$$f(x) = \lim_{dx \rightarrow 0} \frac{P(x \leq X \leq x+dx)}{dx}$$

$$f(x) \geq 0; \int_{-\infty}^{\infty} f(x)dx = 1$$

Cumulative distribution function (CDF): -

The CDF of a random variable  $X$  is given by

$$F(x) = P(X \leq x)$$

$$F(x) = P(X \leq x) = \sum_{x_i < x} p(x_i)$$

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(x)dx$$

Note: -

The relation between PDF,  $f(x)$ , and CDF,  $F(x)$ , is given by

$$f(x) = F'(x)$$

Probability Distribution:-

A random variable  $X$  together with their PMF (PDF) or CDF is called Probability distribution of  $X$ .

Problem:- A coin is tossed two times, if  $X$  denotes the number of heads, Find the Probability distribution of  $X$ .

Soln:-

When a coin is tossed two times.

S : HHTTHTTT

$X \rightarrow$  No. of heads

values of  $X$ :  $(x=0, 1, 2)$

The Probability distribution of  $X$ :

$x$	0	1	2
$P(x)$	$\frac{1}{4}$	$\frac{2}{4}$	$\frac{1}{4}$

✓ A random variable  $X$  has the following probability distribution:

$x$	0	1	2	3	4	5	6	7
$P(x)$	0	$k$	$2k$	$2k$	$3k$	$k^2$	$2k^2$	$7k^2+k$

= A

- Find the value of  $k$
- Evaluate  $P(X < b)$ ;  $P(X \geq b)$ ;  $P(0 < X < 5)$
- Find the cumulative distribution of  $X$ .
- Find the minimum value of ' $a$ ' such that  $P(X \leq a) > \frac{1}{2}$ .

Soln:

- Find the value of  $k$
- Evaluate  $P(X < 2)$ ;  $P(X \geq 0)$

Given  $\sum p(x) = 1$  because it must be 1

$$0 + k + 2k + 3k + k^2 + 2k^2 + 7k^2 + k = 1$$

$$10k^2 + 9k - 1 = 0$$

$$(10k^2 + 10k - k - 1) = 0$$

$$(10k(k+1) - (k+1)) = 0$$

$$(k+1)(10k-1) = 0$$

$$\begin{cases} k+1 = 0 \\ k = -1 \end{cases} \quad \begin{cases} 10k-1 = 0 \\ k = \frac{1}{10} \end{cases}$$

$$K = -1 \text{ is impossible}$$

$$K = \frac{1}{10}$$

$$P(X < 2) = P(0, 1, 2, 3, 4, 5)$$

$$= 0 + k + 2k + 3k + k^2 + 2k^2 + 7k^2$$

$$= k^2 + 8k$$

$$= \frac{1}{100} + \frac{8}{10} = \frac{81}{100} = 0.81$$

$$P(X \leq b) = 0.81$$

$P(X \geq 6) = P(6, 7) \text{ or } \text{event } A \in \{\omega : x > 6\}$ $\geq K^2 + 7K + K = 19K$ $x \text{ is random variable so it lies between } 0 \text{ and } 10$ $\text{minimum value } 0 \text{ and maximum value } 10$ $= \frac{9}{10} + \frac{1}{10} = \frac{19}{10}$																											
$P(X \geq 6) = \frac{19}{10}$																											
$P(0 < X < 5) = P(1, 2, 3, 4)$ $= K + 2K + 2K + 3K = 8K$																											
$P(0 < X < 5) = \frac{8}{10}$																											
iii) cumulative distribution of $X$ .																											
<table border="1" style="width: 100%; border-collapse: collapse;"> <thead> <tr> <th style="text-align: center;"><math>X</math></th> <th style="text-align: center;">0</th> <th style="text-align: center;">1</th> <th style="text-align: center;">2</th> <th style="text-align: center;">3</th> <th style="text-align: center;">4</th> <th style="text-align: center;">5</th> <th style="text-align: center;">6</th> <th style="text-align: center;">7</th> </tr> </thead> <tbody> <tr> <td style="text-align: center;"><math>P(X)</math></td> <td style="text-align: center;"><math>0</math></td> <td style="text-align: center;"><math>\frac{1}{10}</math></td> <td style="text-align: center;"><math>\frac{2}{10}</math></td> <td style="text-align: center;"><math>\frac{2}{10}</math></td> <td style="text-align: center;"><math>\frac{3}{10}</math></td> <td style="text-align: center;"><math>\frac{1}{100}</math></td> <td style="text-align: center;"><math>\frac{2}{100}</math></td> <td style="text-align: center;"><math>\frac{17}{100}</math></td> </tr> <tr> <td style="text-align: center;"><math>F(x)</math></td> <td style="text-align: center;"><math>0</math></td> <td style="text-align: center;"><math>\frac{1}{10}</math></td> <td style="text-align: center;"><math>\frac{3}{10}</math></td> <td style="text-align: center;"><math>\frac{5}{10}</math></td> <td style="text-align: center;"><math>\frac{8}{10}</math></td> <td style="text-align: center;"><math>\frac{8}{10} + \frac{1}{100}</math></td> <td style="text-align: center;"><math>\frac{81}{100}</math></td> <td style="text-align: center;"><math>1</math></td> </tr> </tbody> </table>	$X$	0	1	2	3	4	5	6	7	$P(X)$	$0$	$\frac{1}{10}$	$\frac{2}{10}$	$\frac{2}{10}$	$\frac{3}{10}$	$\frac{1}{100}$	$\frac{2}{100}$	$\frac{17}{100}$	$F(x)$	$0$	$\frac{1}{10}$	$\frac{3}{10}$	$\frac{5}{10}$	$\frac{8}{10}$	$\frac{8}{10} + \frac{1}{100}$	$\frac{81}{100}$	$1$
$X$	0	1	2	3	4	5	6	7																			
$P(X)$	$0$	$\frac{1}{10}$	$\frac{2}{10}$	$\frac{2}{10}$	$\frac{3}{10}$	$\frac{1}{100}$	$\frac{2}{100}$	$\frac{17}{100}$																			
$F(x)$	$0$	$\frac{1}{10}$	$\frac{3}{10}$	$\frac{5}{10}$	$\frac{8}{10}$	$\frac{8}{10} + \frac{1}{100}$	$\frac{81}{100}$	$1$																			
iv) $P(X \leq a) \geq \frac{1}{2}$ $F(a) \geq \frac{1}{2}$ $a = 4, 5, 6, 7$ Minimum value of $a = 4$																											

3. A random variable  $X$  has the following Probability distribution

$X$	-2	-1	0	1	2	3
$P(X)$	0.1	$k$	0.2	$2k$	0.3	$3k$

- i) find the value of  $k$
- ii) Evaluate  $P(X < 2)$ ;  $P(-2 < X \leq 1)$
- iii) Find the CDF of  $X$ .

Soln:-

$$i) 0.1 + k + 0.2 + 2k + 0.3 + 3k = 1$$

$$0.1 + k + 0.2 + 2k + 0.3 + 3k = 1$$

$$0.6 + 6k = 1$$

$$6k + 0.6 - 1 = 0$$

$$6k - 0.4 = 0$$

$$6k = 0.4$$

$$k = \frac{0.4}{6}$$

$$k = \frac{1}{15}$$

Now we can find the probability distribution of  $X$

$$ii) P(X < 2) = P(X = -2, -1, 0, 1)$$

$$(iii) F(x) = P(X \leq x) = 0.1 + k + 0.2 + 2k$$

$$= 0.3 + 3k$$

$$= 0.3 + (3 \times \frac{1}{15})$$

$$= 0.3 + 0.2$$

$$\begin{aligned}
 & P(-2 < x < 1) = P(-1, 0) \\
 & = \frac{7.5}{15} = 0.5 \\
 & P(-2 < x < 1) = 0.5 \\
 & P(-2 < x < 1) = P(-1, 0) \\
 & = \frac{1}{15} + 0.2 = 0.26 \\
 & I = \frac{4}{15} \\
 & P(-2 < x < 1) = 0.26
 \end{aligned}$$

iii) Cumulative Distribution of  $X$

$x$	-2	-1	0	1	2	3
$p(x)$	0.1	$\frac{1}{15}$	0.2	$\frac{2}{15}$	0.3	$\frac{3}{15}$
$F(x)$	0.1	0.16	0.36	0.5	0.8	1

• Check whether the following is a Probability density function or not.

$$f(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0, \lambda > 0 \\ 0, & \text{elsewhere.} \end{cases}$$

Soln:-

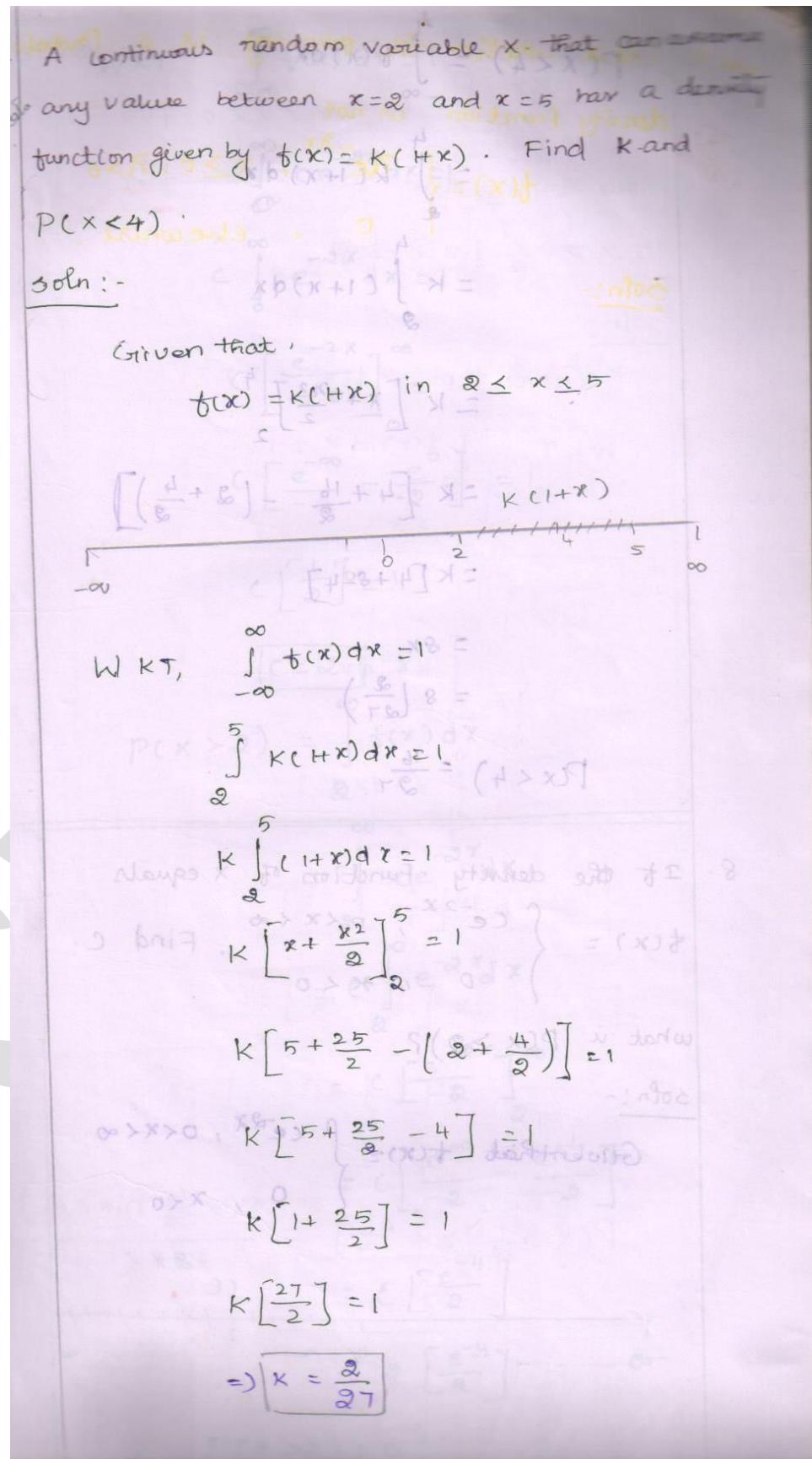
$$\int_{-\infty}^{\infty} f(x) dx = \int_{0}^{\infty} \lambda e^{-\lambda x} dx$$

$$= \lambda \left[ \frac{e^{-\lambda x}}{-\lambda} \right]_0^{\infty}$$

$$= \lambda \left[ 0 + \frac{1}{\lambda} \right]$$

$$= 1$$

$f(x)$  is a pdt.



$$\begin{aligned}
 P(x < 4) &= \int_{-\infty}^4 f(x) dx \\
 &= \int_2^4 K(1+x) dx \\
 &= K \int_2^4 (1+x) dx \\
 &= K \left[ x + \frac{x^2}{2} \right]_2^4 \\
 &= K \left[ 4 + \frac{16}{2} - \left( 2 + \frac{4}{2} \right) \right] \\
 &= K [4 + 8 - 4] \\
 &= 8K \\
 &= 8 \left( \frac{1}{27} \right) \\
 P(x < 4) &= \frac{16}{27}.
 \end{aligned}$$

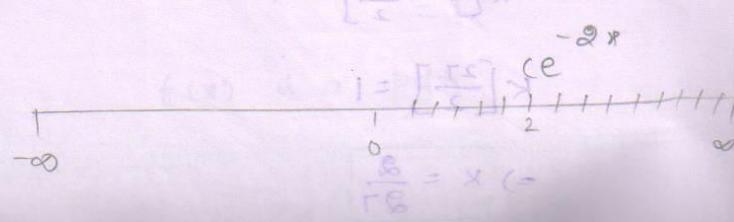
If the density function of  $x$  equals

$$f(x) = \begin{cases} ce^{-2x}, & 0 < x < \infty \\ 0, & x \leq 0 \end{cases}, \text{ Find } c.$$

what is  $P(x > 2)$ ?

Soln:-

$$\text{Given that } f(x) = \begin{cases} ce^{-2x}, & 0 < x < \infty \\ 0, & x \leq 0 \end{cases}$$



WKT  $\int_{-\infty}^{\infty} f(x) dx = 1$  continuous

$$\text{at } b=0 \Rightarrow x > -1 \Rightarrow x = (0) \Rightarrow \text{Ans}$$

$$\int_{-\infty}^{\infty} ce^{-\alpha x} dx = 1$$

$$P\left(\frac{1}{2} < x\right) = \int_{-\infty}^{\infty} f(x) dx$$

$$c \int_{-\infty}^{\infty} e^{-\alpha x} dx = 1$$

$$\therefore c \left[ \frac{e^{-\alpha x}}{-\alpha} \right]_{-\infty}^{\infty} = 1 \quad \text{at required}$$

$$1 = x f(x) \Big|_{-\infty}^{\infty}$$

$$c \left[ \frac{e^{-\infty}}{-\alpha} - \frac{e^0}{-\alpha} \right] = 1$$

$$c \left[ \frac{1}{2} \right] = 1$$

$$\boxed{c = 2}$$

$$P(x > 2) = \int_0^{\infty} f(x) dx$$

$$= \int_0^{\infty} ce^{-\alpha x} dx$$

$$= \left[ -\frac{c}{\alpha} e^{-\alpha x} \right]_0^{\infty}$$

$$= \left[ -\frac{c}{\alpha} e^{-2\alpha} \right]$$

$$= \left[ -\frac{2}{\alpha} e^{-4} \right]$$

$$\frac{(x > n \geq x)^q}{(n > x)^q} = \left( \frac{1}{\alpha} > x \right)^{\frac{1}{\alpha} < x}^q$$

$$= \left[ -\frac{2}{\alpha} e^{-4} \right]$$

$$= 2 \left[ \frac{e^{-4}}{\alpha} \right]$$

$$P(x > 2) = e^{-4}$$

Q4. If random variable  $X$  has the CDF

$$F(x) = \begin{cases} 1 - e^{-\alpha x}, & x > 0 \\ 0, & x \leq 0 \end{cases}$$

Find  $P(1 < X < 2)$ .

Soln:-

Given that  $F(x) = \begin{cases} 1 - e^{-\alpha x}, & x > 0 \\ 0, & x \leq 0 \end{cases}$

$$P(1 < X < 2) = F(2) - F(1)$$

$$= -e^{-2\alpha} - (-e^{-\alpha})$$

$$= e^{-\alpha} - e^{-2\alpha}$$

QUESTION Mathematical Expectation

Let  $X$  be a random variable then the mathematical expectation of  $X$  is given by

$$\frac{d}{dx} P(X=x) = f(x)$$

$$E[X] = \sum x \cdot P(x), \quad x \text{ is discrete}$$

$$E[X] = \int x f(x) dx, \quad x \text{ is continuous}$$

Moments about origin (Raw moments)

The  $r^{\text{th}}$  moment about origin is given by

$M_r = E[x^r] = \begin{cases} \sum x^r p(x), & x \text{ is discrete} \\ \int x^r f(x) dx, & x \text{ is continuous} \end{cases}$

Note :-

Mean :  $\bar{x} = E[x]$

Variance =  $E[x^2] - \text{Mean}^2$

Moments about Mean (Central Moments) :-

The  $r^{\text{th}}$  moment about Mean is given by.

$$M_r = E[(x - \bar{x})^r]$$

15. The density function of a continuous random variable  $x$  is given by  $f(x) = Kx(2-x)$  for  $0 \leq x \leq 2$ .

- Find  $K$ .
- $r^{\text{th}}$  moment about origin
- Mean
- Variance

Soln:-

Given that  $\frac{f(x)}{K+1} = Kx(2-x)$ ,  $0 \leq x \leq 2$

Then  $\frac{K+1}{K+1} = \frac{Kx(2-x)}{K+1}$

WKT,  $\int f(x) dx = 1$

$$\int_0^2 Kx(2-x) dx = 1$$

Find  $E[x]$

$$\int_0^2 (2x - x^2) dx = E[x]$$

$$K \left[ \frac{2x^2}{2} - \frac{x^3}{3} \right]_0^2 = 1$$

$E[x] = \frac{8}{3}$

$$K \left[ \frac{4}{3} - \frac{8}{3} \right] = 1$$

$K = \frac{3}{4}$

$E[x] = \frac{8}{3}$

The  $n^{th}$  moment about origin

$$\mu_r = E[x^r] = \int x^r f(x) dx$$

$$= \int_0^2 x^r \cdot Kx(2-x) dx$$

$$= K \int_0^2 x^r \cdot x(2-x) dx$$

$$= K \int_0^2 (2x^{r+1} - x^{r+2}) dx$$

$$= K \left[ \frac{2x^{r+2}}{r+2} - \frac{x^{r+3}}{r+3} \right]_0^2$$

$$= K \left[ \frac{2 \cdot 2^{r+2}}{r+2} - \frac{2^{r+3}}{r+3} \right]$$

$$\begin{aligned}
 &= K \left[ \frac{2}{r+2} - \frac{2^{r+3}}{r+3} \right] \\
 &= K 2^{r+3} \left[ \frac{1}{r+2} - \frac{1}{r+3} \right] \\
 &= K 2^{r+3} \left[ \frac{r+3 - r-2}{(r+2)(r+3)} \right] \\
 &= K 2^{r+3} \left[ \frac{1}{(r+2)(r+3)} \right] \\
 &\boxed{E[x^r] = x^r \frac{3}{4} \times \frac{2^{r+3}}{(r+2)(r+3)}}
 \end{aligned}$$

Put  $r=1$

$$E[x] = \frac{3}{4} \times \frac{2^4}{3 \times 4} = \frac{3 \times 16}{4 \times 3 \times 4} = 1$$

Put  $r=2$

$$E[x^2] = \frac{3}{4} \times \frac{25}{4 \times 5} = \frac{3 \times 32}{4 \times 4 \times 5} = \frac{6}{5}$$

Mean  $= \frac{1}{5} E[x^2] = 1$

Variance  $= E[x^2] - \text{Mean}^2 = \frac{6}{5} - 1 = \frac{1}{5}$

Mean  $= 1$   
 $\sqrt{\text{Var}} = \frac{1}{\sqrt{5}}$

If the Pdt of  $f(x) = \begin{cases} 2(1-x), & 0 < x < 1 \\ 0, & \text{otherwise} \end{cases}$

then show that the  $n$ th moment about origin

$$\mu_n = \frac{2}{(n+1)(n+2)}$$

Soln: Given that  $f(x) = \begin{cases} \frac{\alpha}{\sigma} e^{-\frac{|x-\mu|}{\sigma}}, & \text{if } |x-\mu| \leq \alpha \\ 0, & \text{otherwise} \end{cases}$

The  $r^{\text{th}}$  moment about origin

$$M_r = E[x^r] = \int_{-\infty}^{\infty} x^r f(x) dx$$

$$= \int_{-\infty}^{\mu - \alpha} x^r \cdot 0 dx + \int_{\mu - \alpha}^{\mu + \alpha} x^r \cdot \frac{\alpha}{\sigma} e^{-\frac{|x-\mu|}{\sigma}} dx + \int_{\mu + \alpha}^{\infty} x^r \cdot 0 dx$$

$$= \int_{\mu - \alpha}^{\mu + \alpha} x^r \cdot \frac{\alpha}{\sigma} e^{-\frac{|x-\mu|}{\sigma}} dx$$

$$= \int_{\mu - \alpha}^{\mu + \alpha} x^r \cdot \frac{\alpha}{\sigma} e^{-\frac{\alpha - x}{\sigma}} dx$$

$$= \int_{\mu - \alpha}^{\mu + \alpha} x^r \cdot \frac{\alpha}{\sigma} e^{-\frac{\alpha}{\sigma} + \frac{x}{\sigma}} dx$$

$$= \int_{\mu - \alpha}^{\mu + \alpha} x^r \cdot \frac{\alpha}{\sigma} e^{-\frac{\alpha}{\sigma}} e^{\frac{x}{\sigma}} dx$$

$$= \frac{\alpha}{\sigma} e^{-\frac{\alpha}{\sigma}} \int_{\mu - \alpha}^{\mu + \alpha} x^r e^{\frac{x}{\sigma}} dx$$

$$= \frac{\alpha}{\sigma} e^{-\frac{\alpha}{\sigma}} \left[ \frac{x^{r+1}}{\sigma} \right]_{\mu - \alpha}^{\mu + \alpha}$$

$$= \frac{\alpha}{\sigma} e^{-\frac{\alpha}{\sigma}} \left[ \frac{(\mu + \alpha)^{r+1} - (\mu - \alpha)^{r+1}}{\sigma} \right]$$

$$= \frac{\alpha}{\sigma} \left[ \frac{1}{r+1} - \frac{1}{r+2} \right]$$

$$= 1 - \frac{d}{\sigma} = \text{moment} - \left[ -2 \left[ \frac{1}{r+1} - \frac{1}{r+2} \right] \right]$$

$$= -2 \left[ \frac{r+2 - r-1}{(r+1)(r+2)} \right]$$

$$= 2 \left[ \frac{1}{(r+1)(r+2)} \right]$$

$$\boxed{M_r = \frac{2}{(r+1)(r+2)}}$$

Moment Generating Function (MGF)

Let  $X$  be a random variable  
then the moment generating function of  $X$  is given by

$$M_X(t) = E[e^{tX}] = \begin{cases} \sum x e^{tx} p(x), & X \text{ is discrete} \\ \int e^{tx} f(x) dx, & X \text{ is continuous} \end{cases}$$

Q1. If the density function of a random variable  $X$  is  $f(x) = \begin{cases} \frac{1}{3}, & -1 < x < 2 \\ 0, & \text{otherwise} \end{cases}$ , then find  $X$ 's MGF.

Soln:-

Given that,

$$f(x) = \begin{cases} \frac{1}{3}, & -1 < x < 2 \\ 0, & \text{otherwise} \end{cases}$$

$$M_X(t) = E[e^{tX}] = \int_{-\infty}^{\infty} e^{tx} f(x) dx = \int_{-1}^{2} e^{tx} \frac{1}{3} dx = \frac{1}{3} [e^{tx}] \Big|_{-1}^{2} = \frac{1}{3} [e^{2t} - e^{-t}]$$

The MGF is,

$$M_X(t) = \frac{1}{3} [e^{2t} - e^{-t}]$$

Soln:-

$$M_X(t) = E[e^{tX}] = \int e^{tx} f(x) dx$$

Given that,

$$f(x) = \begin{cases} \frac{1}{3}, & -1 < x < 2 \\ 0, & \text{otherwise} \end{cases}$$

$$M_X(t) = \int_{-1}^{2} e^{tx} \left(\frac{1}{3}\right) dx = \frac{1}{3} \left[ \frac{e^{tx}}{t} \right] \Big|_{-1}^{2} = \frac{1}{3} \left[ \frac{e^{2t}}{t} - \frac{e^{-t}}{t} \right]$$

$$\begin{aligned} M_X(t) &= \frac{1}{3} \left[ e^{tx} \right]_{-1}^2 = (t+M) \\ M_X(t) &= \frac{1}{3} \left[ \frac{e^{2t}}{t} - \frac{e^{-t}}{t} \right]. \\ Xb(x)dx &= [x] \end{aligned}$$

Find the MGIF of a random variable  $X$  having the PDF  $f(x) = \begin{cases} x, & 0 < x < 1 \\ 2-x, & 1 < x < 2 \\ 0, & \text{otherwise.} \end{cases}$ . Also Find mean, var.

Soln:- Given that  $f(x) = \begin{cases} x, & 0 < x < 1 \\ 2-x, & 1 < x < 2 \\ 0, & \text{otherwise.} \end{cases}$

Note:-  $\left(\frac{x}{\varepsilon}\right) - \frac{1}{\varepsilon} + \frac{1}{\varepsilon} = \frac{1}{\varepsilon} > 0$ , otherwise.

(i) Mean  $\frac{1+1}{\varepsilon} = \frac{2}{\varepsilon} - 1 + \frac{1}{\varepsilon} = \frac{2}{\varepsilon} - 1$

(ii) Var  $\frac{\varepsilon}{\varepsilon} = \frac{1+8-8-0+1+1}{2} = \frac{1}{2}$

The MGIF is,

$$\begin{aligned} M_X(t) &= E[e^{tx}] = \int e^{tx} f(x) dx \\ &= \int_0^2 x e^{tx} dx + \int_1^2 (2-x) e^{tx} dx \\ &= \left[ x \frac{e^{tx}}{t} - 1 \frac{e^{tx}}{t^2} \right]_0^1 + \left[ (2-x) \frac{e^{tx}}{t} - (-1) \frac{e^{tx}}{t^2} \right]_1^2 \\ &= \left[ \frac{x e^{tx}}{t} - \frac{e^{tx}}{t^2} \right]_0^1 + \left[ \frac{(2-x) e^{tx}}{t} + \frac{e^{tx}}{t^2} \right]_1^2 \\ &= \frac{e^t}{t} - \frac{e^t}{t^2} + \frac{1}{t^2} + \frac{e^{2t}}{t^2} - \left( \frac{e^t}{t} + \frac{e^t}{t^2} \right) \\ &= \frac{e^t}{t} - \frac{e^t}{t^2} + \frac{1}{t^2} + \frac{e^{2t}}{t^2} - \frac{e^t}{t} - \frac{e^t}{t^2} \\ &= -\frac{2e^{t-2}}{t^2} + \frac{1}{t^2} + \frac{e^{2t}}{t^2} \end{aligned}$$

$M_X(t) = \frac{e^{2t} + 1 - e^t}{t^2}$

$E[X] = \int x \cdot f(x) dx$

$\text{given by } = \int_0^1 x \cdot x dx + \int_1^2 x(2-x) dx$

$= \int_0^1 x^2 dx + \int_1^2 (2x-x^2) dx$

$= \left[ \frac{x^3}{3} \right]_0^1 + \left[ \frac{2x^2}{2} - \frac{x^3}{3} \right]_1^2$

$= \frac{1}{3} + 4 - \frac{8}{3} - 1 \left( 1 - \frac{1}{3} \right)$

$= \frac{1}{3} + 4 - \frac{8}{3} - 1 + \frac{1}{3}$

$\text{Find } E[X^2]$

$= \frac{1+12-8-8+1}{3} = \frac{3}{3}$

$E[X^2] = \int x^2 f(x) dx$

$= \int_0^1 x^2 \cdot x dx + \int_1^2 x^2(2-x) dx$

$= \int_0^1 x^3 dx + \int_1^2 (2x^2-x^3) dx$

$= \left[ \frac{x^4}{4} \right]_0^1 + \left[ \frac{2x^3}{3} - \frac{x^4}{4} \right]_1^2$

$= \frac{1}{4} + \frac{16}{3} - \frac{16}{4} - \left( \frac{2}{3} - \frac{1}{4} \right)$

$= \frac{3+64}{3} - 48 - 8 + 3 = \frac{67}{3} - 48 = \frac{67-144}{3} = \frac{-77}{3}$

$$\text{Mean} = \frac{70 - 56}{12} = \frac{14}{12} = \frac{7}{6}$$

$$\text{Var} = E[x^2] - \text{Mean}^2 = \frac{7}{6} - 1 = \frac{1}{6}$$

$$\boxed{E[x^2] = \frac{7}{6}}$$

$$\text{Mean} = E[x] = 1$$

$$\text{Var} = E[x^2] - \text{Mean}^2 = \frac{7}{6} - 1 = \frac{1}{6}$$

Note :-

$$(i) \text{ Mean } (x) = M_x'(0)$$

$$(ii) \text{ Var } = M_x''(0) - \text{Mean}^2$$

## Binomial Distribution ( $n, p$ )

A discrete random variable  $x$  follows binomial distribution if its PMF is

$$P(x) = n C_x p^x q^{n-x}, \quad x = 0, 1, 2, 3, \dots, n$$

$p \rightarrow$  Probability of success

$q \rightarrow$  Probability of failure

$$p+q = 1$$

3. Find the MGF of binomial distribution and hence find its mean and variance.

Soln:-

In Binomial the PMF is  $P(x) = n C_x p^x q^{n-x}$ ,  $x = 0, 1, 2, \dots, n$

The MGF is  $M_x(t) = E[e^{tx}]$ .

$$= \sum e^{tx} P(x)$$

$$= \sum_{x=0}^n e^{tx} \cdot n C_x p^x q^{n-x}$$

$$\therefore t \leftarrow x = q^n \leftarrow 8 = n C_0 M$$

$$\begin{aligned}
 &= \sum_{x=0}^n nCx (pe^t)^x q^{n-x} \\
 \therefore (a+b)^n &= \sum_{x=0}^n (nCx a^x b^{n-x}) \\
 M_x(t) &= (pe^t + q)^n \\
 M_x'(t) &= np e^t (pe^t + q)^{n-1} \\
 M_x'(0) &= np (1 + p + q)^{n-1} = np \quad p+q=1 \\
 M_x''(t) &= np [e^t pe^t (n-1) (pe^t + q)^{n-2} + (pe^t + q)^{n-1} e^t] \\
 M_x''(0) &= np [np - p + 1] \\
 &= np [np - p + 1] \\
 P(X > 2) &= n^2 p^2 - np^2 + np - np^2 \\
 \text{Mean} &= M_x'(0) = np \\
 \text{Var} &= M_x''(0) - \text{Mean}^2 = n^2 p^2 - np^2 + np - np^2 \\
 \therefore q &= 1-p \quad 1 = p+q \quad = np - np^2 = np(1-p) \\
 &= npq. \\
 \boxed{\text{Mean} = np} \\
 \boxed{\text{Var} = npq}
 \end{aligned}$$

The Mean and Variance of a binomial variate

Mean and Variance are 8 and 6. Find  $(X \geq 2)$

$\Rightarrow 8q = 6 \quad (\text{By } \textcircled{1})$   
 $\Rightarrow q = \frac{6}{8} = \frac{3}{4} \quad (1 \leq x) q$   
 $\Rightarrow q = \frac{3}{4}$   
 $\boxed{q = 0.75}$   
 $\text{HKT, } P+q=1 \quad \frac{P}{8} = 0.25$   
 $P + \frac{3}{4} = 1$   
 $\boxed{P = \frac{1}{4}} = 0.25$   
 $\text{From } \textcircled{1} \Rightarrow n \times \frac{1}{4} = 8$   
 $\Rightarrow \boxed{n = 32}$   
 $\text{In Binomial, } P(x) = {}^n C_x p^x q^{n-x}, x=0,1,2,\dots$   
 $P(x) = 32 \times \left(\frac{1}{4}\right)^x \left(\frac{3}{4}\right)^{32-x}, x=0,1,2,\dots$   
 $P(x \geq 2) = 1 - P(x < 2)$   
 $= 1 - P(0,1)$   
 $\therefore P(x \geq 2) = 1 - \left\{ 32 \left(0.25\right)^0 \left(0.75\right)^{32} + 32 \left(1\right) \left(0.25\right)^1 \left(0.75\right)^{31} \right\}$   
 $= 1 - \left(0.75\right)^{32} + 32 \times 0.25 \times \left(0.75\right)^{31}$   
 $\therefore P(x \geq 2) \approx 0.9988$

Poisson Distribution ( $\lambda$ )

A discrete random variable  $X$

follows Poisson distribution if its PMF is

$$P(X) = \frac{e^{-\lambda} \cdot \lambda^X}{X!}, X = 0, 1, 2, \dots$$

29. Find the MGF of Poisson distribution.  
Hence find Mean and variance.

Soln:-

In Poisson,  $P(X) = \frac{e^{-\lambda} \cdot \lambda^X}{X!}, X = 0, 1, 2, \dots$

MGF is  $M_X(t) = E[e^{tX}] = \sum e^{tx} P(x)$

$$\begin{aligned} &= \sum_{x=0}^{\infty} e^{tx} \frac{e^{-\lambda} \cdot \lambda^x}{x!} \\ &= e^{-\lambda} \cdot e^{\lambda t} \cdot \sum_{x=0}^{\infty} \frac{(\lambda e^t)^x}{x!} \\ &= e^{-\lambda} \cdot e^{\lambda t + \lambda t} \\ M_X(t) &= e^{\lambda t + \lambda t} \\ M_X'(t) &= \lambda e^t e^{\lambda t} = \lambda e^{\lambda t + \lambda t} \\ M_X''(t) &= \lambda e^t (1 + \lambda e^t) e^{\lambda t} = \lambda (1 + \lambda e^t) e^{\lambda t + \lambda t} \\ &= \lambda + \lambda^2 e^{2\lambda t} \end{aligned}$$

(A) Poisson Distribution

If  $x$  is a poission variate such that

$$P(x=2) = 9 \cdot P(x=4) + 90 \cdot P(x=6), \text{ Find}$$

- Mean
- $E[x^2]$
- $P(x \geq 2)$

Soln:-

The Probability of a poission variate is given by  $P(x) = \frac{e^{-\lambda} \cdot \lambda^x}{x!}, x=0, 1, 2, \dots$

$\therefore P(x=2) = 9 \cdot P(x=4) + 90 \cdot P(x=6)$

$$\frac{e^{-\lambda} \cdot \lambda^2}{2!} = 9 \left[ \frac{e^{-\lambda} \cdot \lambda^4}{4!} \right] + 90 \left[ \frac{e^{-\lambda} \cdot \lambda^6}{6!} \right]$$

$$\therefore e^{-\lambda} \cdot \lambda^2 \cdot \frac{1}{2!} = \frac{9 \lambda^4 \cdot 1}{4!} + \frac{90 \lambda^6 \cdot 1}{6!}$$

$$\frac{\lambda^4}{8} + \frac{3 \lambda^2 - 4}{8} = 0$$

$$\lambda^4 + 3\lambda^2 - 4 = 0$$

$$\lambda^4 + 3\lambda^2 - 3 - 1 = 0$$

$$(\lambda^4 + 1) + 3(\lambda^2 - 1) = 0$$

$$(\lambda^2 + 1)(\lambda^2 - 1) + 3(\lambda^2 - 1) = 0$$

$$(\lambda^2 - 1)(\lambda^2 + 1 + 3) = 0$$

$$(x^2 - 1)(x^2 + 4) = 0 \Rightarrow x^2 = 1 \text{ or } x^2 = -4$$

$$x^2 = 1 \Rightarrow x = \pm 1$$

$$x^2 = -4 \text{ has no real roots}$$

$$\therefore x = \pm 1$$

$$P(x) = \frac{e^{-1} \cdot 1^x}{x!} = \frac{e^{-1}}{x!}, x = 0, 1, 2, \dots$$

i) Mean =  $\lambda = 1$

$E[x^2] = \sum x^2 P(x)$

$E[x^2] = 1^2 \cdot \frac{e^{-1}}{0!} + 2^2 \cdot \frac{e^{-1}}{1!} + 3^2 \cdot \frac{e^{-1}}{2!} + \dots$

ii)  $E[x^2] = 2$

iii)  $P(x \geq 2) = 1 - P(x < 2)$

$= 1 - \left\{ \frac{e^{-1}}{0!} + \frac{e^{-1}}{1!} \right\}$

$= 1 - \left\{ e^{-1} + e^{-1} \right\}$

$= 1 - 2e^{-1}$

$= 0.264$

$P(x > 1) = (1 - x)^{-1} = \frac{(2-1)^{-1}}{1!} = \frac{1}{1!} = 1$

Find the MGF of Geometric distribution.

Also find its mean & variance

Soln:-

Expo  $x$ ,  $\frac{e^{kx}-1}{tx}$  =  $(x)$  exponential in I

In Geometric  $P(x) = pq^{x-1}$ ,  $x = 1, 2, 3, \dots$

$$\text{M}_x(t) = E[e^{tx}]$$

$$\sum_{x \in S} e^{\frac{t}{\gamma} x} P(x)$$

$$\sum_{x=0}^{\infty} e^{tx} = p q^x$$

$$(1, \varepsilon, s, 1, 0)q = P \sum_{n=0}^{\infty} (qe^n)^*$$

$$= P \{ 1 + (qe^t) + (qe^t)^2 + (qe^t)^3 + \dots \}$$

$$\left[ 1 + x + x^2 + \dots \right] = (1-x)^{-1}$$

$$S_0 = P(1 - qe^t)^{-1}$$

$$(g) M_x(t) = \frac{P}{(1-ge^t)}$$

$$M(t) = E[e^{tX}]$$

$$f(x) = \frac{1}{Z} e^{-\beta H(x)} p(x)$$

$$= \sum_{p,q} e^{tx^p} p q^{x-1}$$

Meine Tweak für  $P \sum e^{\frac{tx}{q}} q^{x-1}$

$$= P \left\{ e^{t+} e^{2t} q + e^{st} q^2 + \dots \right\}$$

$$\begin{aligned}
 P(x) &= Pe^t \left\{ 1 + e^t q + e^{2t} q^2 + \dots \right\} \\
 &\stackrel{\text{WKT, } 1+x+x^2+\dots}{=} (1-x)^{-1} \\
 \Rightarrow P(x) &= Pe^t (1-qe^t)^{-1} \\
 M_x(t) &= \boxed{\frac{Pe^t}{1-qe^t}} \\
 M_x'(t) &= \frac{(1-qe^t)Pe^t - Pe^t(-qe^t)}{(1-qe^t)^2} \\
 &= \frac{Pe^t (1-qe^t + qe^t)}{(1-qe^t)^2} \\
 M_x'(t) &= \boxed{\frac{Pe^t}{(1-qe^t)^2}} \\
 M_x'(0) &= \frac{P}{(1-q)^2} = \frac{P}{P^2} = \frac{1}{P} \\
 M_x'(0) &= \boxed{\frac{1}{P}} \\
 M_x''(t) &= \frac{(1-qe^t)^2 Pe^t - Pe^t \cdot 2(1-qe^t)(-qe^t)}{(1-qe^t)^4} \\
 \textcircled{1} \leftarrow M_x''(t) &= \frac{Pe^t (1-qe^t) \{ 1-qe^t + 2qe^t \}}{(1-qe^t)^4} \\
 M_x''(t) &= \frac{Pe^t (1+q)e^t}{(1-qe^t)^3} \\
 M_x''(0) &= \frac{P(1+q)}{P-1(1-q)^3} = \frac{P(1+q)}{P\beta_2} = \frac{1}{\beta_2} \\
 M_x''(0) &= \frac{1+q}{P^2} = 
 \end{aligned}$$

$\text{Mean} = Mx'(0) = \frac{1}{P} q =$   
 $\text{Var} = Mx''(0) - \text{Mean}^2 = \frac{1}{P^2} - \frac{1}{P^2} = \frac{q}{P^2}$

---

State and Prove memory less Property  
 on Geometric distribution  $= (f) \times M$

If  $X$  is a Geometric variate, then  $X$  lacks memory (i.e.)

(a)  $P(X > s+t | X > s) = P(X > t), s, t > 0$

Proof :-

In Geometric  $P(X) = pq^{x-1}, x = 1, 2, 3, \dots$

$P(X > s+t | X > s) = \frac{P(X > s+t \cap X > s)}{P(X > s)}$

$$= \frac{q}{pq} = \frac{q}{\frac{s}{pq}} = \frac{q}{\frac{s(p+1)}{pq}} = (0)' \times M$$

$P(X > s+t | X > s) = \frac{P(X > s+t)}{P(X > s)} \rightarrow \textcircled{1}$

$P(X > s) = \frac{P(s+1, s+2, s+3, \dots)}{P(X)}$

$$= \frac{(pq)^s + (pq)^{s+1} + (pq)^{s+2} + \dots}{pq} = (0)' \times M$$

$$= pq^s \{ 1 + q + q^2 + \dots \}$$

$$= \frac{(p+1)q^s}{q} = \frac{(p+1)q^s}{pq(1-q)} = (0)' \times M$$

$$= pq^{s-1} p^{-1} = q^s \times M$$

$P(X > s) = q^s$   
 $P(X > s+t) = q^{s+t}$   
 $\Rightarrow P(X > s+t | X > s) = \frac{q^{s+t}}{q^s} = \frac{q^s \cdot q^t}{q^s} = q^t$   
 $\therefore P(X > t) = P(X > s+t | X > s)$

It is the Probability that an applicant for a drivers licence will pass the road test on any given trial is 0.8, what is the Probability that he will finally pass the road test?

1. On fourth trial
2. In fewer than four trials?

Soln:  $P(X > 4) = (0.2)^3$  (completing first three trials and failing at 4th trial)

Success  $\rightarrow$  Pass the road test  
 $(X \leq x) = 1 - P(X > x)$

~~$P = 0.8$~~   
 $q = 1 - P = 0.2$   
 $P(X \leq x) = 1 - (0.2)^{x-1}, x = 1, 2, 3, \dots$   
 $P(X \leq 4) = 1 - (0.2)^3 = 0.8$   
 $P(X \leq 3) = 1 - (0.2)^2 = 0.96$   
 $P(X \leq 2) = 1 - (0.2)^1 = 0.98$   
 $P(X \leq 1) = 1 - (0.2)^0 = 0.99$

$P(\text{Pass the test on 4th trial}) = P(X = 4)$   
 $= (0.8)(0.2)^3 = 0.0064$

$P(\text{Pass the test in fewer than 4 trials}) = P(X < 4)$   
 $= P(X = 1, 2, 3) = (0.8)(0.2)^0 + (0.8)(0.2)^1 + (0.8)(0.2)^2 = 0.992$

A die is tossed until 6 appears. What is the probability that it must be tossed more than 5 times?

Soln:-

Solve & Prove memory less Property

Success  $\rightarrow$  Getting 6

$P = \frac{1}{6}$  for standard 6 faces  
So are not continuous is a Geometric series

 $q = 1 - P = \frac{5}{6}$ 

In Geometric  $P(x) = pq^{x-1}, x = 1, 2, 3, \dots$

$P(x) = \left(\frac{1}{6}\right)\left(\frac{5}{6}\right)^{x-1}, x = 1, 2, \dots$

$P(\text{die is tossed more than 5 times}) = P(x \geq 5)$

$= 1 - P(x \leq 4)$

$= 1 - P(1, 2, 3, 4, 5)$

$= 1 - \left\{ \frac{1}{6} + \left(\frac{1}{6}\right)\left(\frac{5}{6}\right) + \left(\frac{1}{6}\right)\left(\frac{5}{6}\right)^2 + \dots \right\}$

$(1 - 0.5) = 1 - 0.5968$

$(0.5)(0.5) = 0.4018$

Q. A coin is tossed until the first head

④ Occurs assuming that the tosses are independent. The probability of head occurring is  $P(H)$ . Find the value of  $P(H)$ .

(Uniform distribution  $(a,b)$ )

A continuous Random variable  $x$

follows Uniform distribution in  $(a,b)$  if its pdf is

$$f(x) = \frac{1}{b-a}, \quad a < x < b$$

Find the MGF of Uniform distribution. Also find its mean and variance.

In Uniform,  $f(x) = \frac{1}{b-a}$ ,  $a < x < b$

MGF is  $M_x(t) = E[e^{tx}]$

$$= \int_a^b e^{tx} f(x) dx$$

$$= \int_a^b e^{tx} \frac{1}{b-a} dx$$

$$= \frac{1}{b-a} \left[ \frac{e^{tx}}{t} \right]_a^b$$

$$= \frac{1}{b-a} \left[ \frac{e^{bt}}{t} - \frac{e^{at}}{t} \right]$$

$$= \frac{1}{b-a} \left[ \frac{e^{bt} - e^{at}}{t} \right]$$

$$E[x^r] = \int_a^b x^r f(x) dx$$

$$= \int_a^b x^r \frac{1}{b-a} dx$$

$$= \frac{1}{b-a} \left[ \frac{x^{r+1}}{r+1} \right]_a^b$$

$E[x^r] = \frac{1}{b-a} \left[ \frac{x^{r+1}}{r+1} \right]_a^b$

$$E[x^r] = \frac{1}{b-a} \left[ \frac{b^{r+1} - a^{r+1}}{r+1} \right]$$

Put  $r=1$

$$E[x] = \frac{1}{b-a} \left[ \frac{b^2 - a^2}{2} \right]$$

$$= \frac{1}{b-a} \left[ \frac{(b-a)(b+a)}{2} \right]$$

$$E[x] = \frac{b+a}{2}$$

Put  $r=2$

$$E[x^2] = \frac{1}{b-a} \left[ \frac{b^3 - a^3}{3} \right]$$

$$= \frac{1}{b-a} \left[ \frac{(b-a)(b^2 + ba + a^2)}{3} \right]$$

$$E[x^2] = \frac{b^2 + ba + a^2}{3}$$

Mean =  $E[x] = \frac{b+a}{2}$

Variance =  $E[x^2] - \text{Mean}^2$

$$= \frac{b^2 + ba + a^2}{3} - \frac{(b+a)^2}{4}$$

$$= \frac{4b^2 + 4ba + 4a^2 - (b^2 + 2ab + a^2)}{12}$$

$$= \frac{b^2 - 2ab + a^2}{12}$$

$$= \frac{(b-a)^2}{12}$$

$\therefore \text{Var} = \frac{(b-a)^2}{12}$

If  $x$  is uniformly distributed with mean 1 and variance  $\frac{4}{3}$ , Find  $P(x < 0)$ .

Soln:-

In Uniform,  $f(x) = \frac{1}{b-a}$ ,  $a < x < b$

Given Mean = 1  $\Rightarrow \frac{b+a}{2} = 1 \Rightarrow b+a = 2 \rightarrow \textcircled{1}$

$\text{Var} = \frac{4}{3} \Rightarrow \frac{(b-a)^2}{12} = \frac{4}{3} \Rightarrow (b-a)^2 = 12 \times \frac{4}{3}$

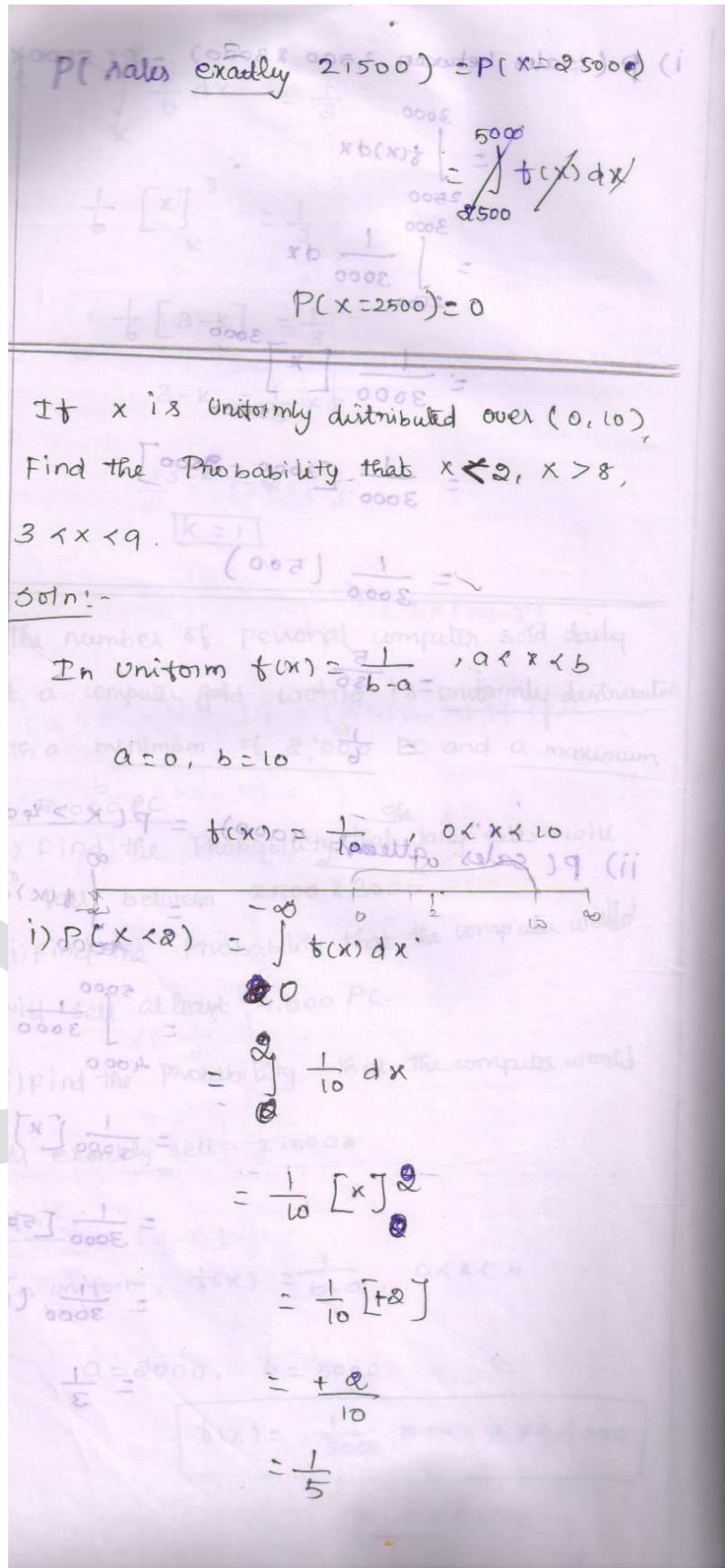
$(b-a)^2 = 16 \Rightarrow b-a = 4 \rightarrow \textcircled{2}$

$\textcircled{1} + \textcircled{2} \Rightarrow b+a+b-a = 2+4 \Rightarrow 2b = 6 \Rightarrow b = 3$

$\textcircled{1} \Rightarrow 3+a = 2 \Rightarrow a = -1$

$f(x) = \frac{1}{4}, -1 < x < 3$

$P(x < 0) = \int_{-\infty}^0 f(x) dx = \int_{-\infty}^0 \frac{1}{4} dx = \left[ \frac{x}{4} \right]_{-\infty}^0 = \frac{0}{4} - \left( \frac{-\infty}{4} \right) = \frac{\infty}{4} = \infty$



Exponential Distribution

$$\text{i)} P(x > 8) = \int_{8}^{\infty} f(x) dx$$

$$= \int_{8}^{\infty} \frac{1}{10} dx$$

then  $x > 8 \Rightarrow x - 8 > 0 \Rightarrow (x-8)^{-1} e^{-\lambda(x-8)} = p(x>8)$

$$= \frac{1}{10} [x]_{8}^{\infty}$$

$$= \frac{1}{10} [\infty - 8] = \frac{1}{10} \times 10 = 1$$

$$x b(x) \int_{0}^{\infty} = \frac{1}{10} (2)$$

$$\text{In } x b(x) \int_{0}^{\infty} = \frac{2}{10} = \frac{1}{5}$$

$$\text{iii)} P(8 < x < 9) = \int_{8}^{9} \frac{1}{10} dx$$

$$x b(x) \int_{0}^{\infty} = \frac{1}{10} [x]_{0}^{\infty}$$

$$= \frac{1}{10} [\infty - 0] = \frac{1}{10} \infty$$

$$\int_{0}^{\infty} \frac{x(\lambda - \lambda)}{(\lambda - \lambda) - 1} \left[ \frac{1}{\lambda} \right] = \frac{1}{10} [9 - 8] = \frac{1}{10} [1] = \frac{1}{10}$$

$$\left[ \frac{0}{\lambda - \lambda} - \frac{\infty}{\lambda - \lambda} \right] \lambda = \frac{1}{10} [6] = \frac{6}{10} = \frac{3}{5}$$

$$\left[ \frac{1}{\lambda - \lambda} \right] P(8 < x < 9) = \frac{3}{5}$$

$$(7+\lambda) A + \frac{\lambda}{\lambda - \lambda} = (1) \times 9$$

$$\Rightarrow \text{Exponential Distribution } (\lambda)$$

$$\frac{1}{\lambda} = R = \lambda \cdot R = (0)^{\lambda} \times 1 = (0)^{\lambda} \lambda = (0)^{\lambda} \lambda$$

A continuous Random variable  $x$  follows Exponential distribution if its PDF is

$$f(x) = \lambda e^{-\lambda x}, x \geq 0$$

$$\frac{1}{\lambda} = \frac{1}{R} = \frac{0}{R} = 0$$

Find the MGF of Exponential distribution.

Hence find its Mean and variance

Soln:-

$$\text{MGF } M_x(t) = E[e^{tx}]$$

$$\text{In exponential, } f(x) = \lambda e^{-\lambda x}, x > 0$$

$$M_x(t) = \int_0^{\infty} e^{tx} \lambda e^{-\lambda x} dx$$

$$= \lambda \int_0^{\infty} e^{tx - \lambda x} dx$$

$$= \lambda \int_0^{\infty} e^{-(\lambda-t)x} dx$$

$$= \lambda \left[ \frac{e^{-(\lambda-t)x}}{-(\lambda-t)} \right]_0^{\infty}$$

$$= \lambda \left[ \frac{0}{-(\lambda-t)} - \frac{1}{-(\lambda-t)} \right]$$

$$= \lambda \left[ \frac{1}{\lambda-t} \right]$$

$$M_x(t) = \frac{\lambda}{\lambda-t} = \lambda(\lambda-t)^{-1}$$

(R) residual probability

$$M_x'(t) = \lambda(\lambda-t)^{-2}, M_x'(0) = \lambda \cdot \lambda^{-2} = \lambda = \frac{1}{\lambda}$$

$$M_x''(t) = +2\lambda(\lambda-t)^{-3}, M_x''(0) = +2\lambda \cdot \lambda^{-3} = \frac{2}{\lambda^2}$$

$$\text{Mean } \mu = M_x'(0) = \frac{1}{\lambda} = \lambda t$$

$$\text{Var} = M_x''(0) - \text{Mean}^2 = \frac{2}{\lambda^2} - \frac{1}{\lambda^2} = \frac{1}{\lambda^2}$$

- State and Prove Memoryless Property

Exponential distribution:

If  $x$  is a Exponential variate,  
 $f(x) = \lambda e^{-\lambda x}$

then  $x$  lacks memory (i)  $P(x > s+t | x > s) = P(x > t)$

(ii)  $P(t < x) = 1 - e^{-\lambda x}$

Proof:-

In exponential,  $f(x) = \lambda e^{-\lambda x}, x > 0$

$$P(x > s+t | x > s) = \frac{P(x > s+t \cap x > s)}{P(x > s)}$$

$$(t < x) \Leftrightarrow x > s$$

$$\text{Now } P(x > s+t) = \frac{P(x > s+t \cap x > s)}{P(x > s)} \rightarrow \text{D}$$

$$P(x > s) = \int_s^{\infty} \lambda e^{-\lambda x} dx$$

$$= \lambda \left[ \frac{e^{-\lambda x}}{-\lambda} \right]_s^{\infty}$$

$$= \lambda \left[ \frac{e^{-\infty}}{-\lambda} - \frac{e^{-s\lambda}}{-\lambda} \right]$$

$$\begin{aligned}
 &= \lambda \left[ \frac{e^{-\lambda s}}{\lambda} \right] \\
 P(X > s) &= e^{-\lambda s} \\
 P(X > s+t) &= e^{-\lambda(s+t)} \\
 &= e^{-\lambda s - \lambda t} \\
 &= e^{-\lambda s} \cdot e^{-\lambda t} \\
 \text{① } \Rightarrow P(X > s+t | X > s) &= \frac{e^{-\lambda s} \cdot e^{-\lambda t}}{e^{-\lambda s}} \\
 &= e^{-\lambda t} \\
 &= P(X > t)
 \end{aligned}$$

**Gamma Distribution( $X; k$ )**

A continuous random variable  $X$  follows gamma distribution if its PDF

$$f(x) = \frac{\lambda^k x^{k-1} e^{-\lambda x}}{\Gamma(k)}, \quad x > 0$$

Find the MGF, mean and variance of gamma distribution.

Soln:-

$$\text{In Gamma, } f(x) = \frac{\lambda^k x^{k-1} e^{-\lambda x}}{\Gamma(k)}, x > 0$$

$$\text{MGF is } M_x(t) = E[e^{tx}]$$

$$= \int_{-\infty}^{\infty} e^{tx} f(x) dx$$

$$= \int_0^{\infty} e^{tx} \frac{\lambda^k x^{k-1} e^{-\lambda x}}{\Gamma(k)} dx$$

$$e^{tx - \lambda x} = e^{-(\lambda - t)x}$$

$$= \frac{\lambda^k}{\Gamma(k)} \int_0^{\infty} x^{k-1} e^{-(\lambda - t)x} dx$$

$$u = (\lambda - t)x \Rightarrow x = \frac{u}{\lambda - t}$$

$$\Rightarrow dx = \frac{du}{\lambda - t}$$

$$= \frac{\lambda^k}{\Gamma(k)} \int_0^{\infty} \frac{u^{k-1} e^{-u}}{(\lambda - t)^k} du$$

$$= \frac{\lambda^k}{\Gamma(k)} \int_0^{\infty} u^{k-1} e^{-u} du$$

Now (using Beta function) go

$$= \frac{\lambda^k}{\Gamma(k)} \int_0^{\infty} u^{k-1} e^{-u} du$$

$$\left[ \begin{array}{l} \text{Formula} \\ x^n e^{-x} dx = \Gamma(n+1) \end{array} \right]$$

$m_x(t) = \frac{\lambda^k t^k}{(\lambda-t)^k}$  TRM 2013

$\Rightarrow m_x'(t) = k\lambda^k (\lambda-t)^{-k-1}$  -! note

$m_x'(0) = k\lambda^k \lambda^{-k-1} = k\lambda^{-1}$  =  $\frac{k}{\lambda}$

$m_x''(t) = -k\lambda^k (-k-1)(\lambda-t)^{-k-2}$  TRM

when  $x = 0$ ,  $m_x''(0) = k(k+1)\lambda^k (\lambda-t)^{-k-2}$

$m_x''(0) = k(k+1)\lambda^k \cdot \lambda^{-k-2}$

$= k(k+1)\lambda^{-2}$

$m_x'''(0) = \frac{k^2+k}{\lambda^2}$

Mean  $= m_x'(0) = \frac{k}{\lambda}$

Vari  $= m_x''(0) \Rightarrow \text{mean}^2 = \frac{k^2+k}{\lambda^2} - \frac{k^2}{\lambda^2}$

$\Rightarrow \frac{k^2+k}{\lambda^2} = \lambda^2 - (2\lambda^2 + \lambda^2) = \lambda^2$

$\Rightarrow -k = \frac{\lambda^2}{\lambda^2}$

b. In a city, the daily consumption of electric power is a random variable with gamma distribution having parameter  $\lambda = \frac{1}{2}$ ,  $k = 3$ . If the power plant of the city has daily capacity of 18 MWh (million kilo watt hour) what is the probability that the power supply be inadequate on any given day.

# 2-2-2

In Gamma  $f(x) = \frac{\lambda^x x^{k-1} e^{-\lambda x}}{\Gamma(k)}$ ,  $x > 0$

$\therefore f(x) = \frac{(1/2)^3 x^2 e^{-x/2}}{\Gamma(3)}$

It is the joint  $f(x) = \frac{1}{16} x^2 e^{-x/2}$ ,  $x > 0$

$P(X > 12) = P(\lambda > 12)$

then find

- i) joint density function
- ii) Marginal distribution
- iii) conditional Probability  $P(X \leq 12 | \lambda = 12)$
- iv) conditional Probability  $P(\lambda > 12 | X \geq 12)$

$\therefore P(X > 12) = \frac{1}{16} \int_{12}^{\infty} x^2 e^{-x/2} dx$

$= \frac{1}{16} \left[ -2x^2 e^{-x/2} \Big|_{12}^{\infty} - 2x \left[ \frac{e^{-x/2}}{-1/4} \right] \Big|_{12}^{\infty} + 2 \left[ \frac{e^{-x/2}}{1/8} \right] \Big|_{12}^{\infty} \right]$

$= \frac{1}{16} \left\{ -2x^2 e^{-x/2} \Big|_{12}^{\infty} - 8x e^{-x/2} \Big|_{12}^{\infty} - 16 e^{-x/2} \Big|_{12}^{\infty} \right\}$

$= \frac{1}{16} \left\{ 0 - (-288 e^{-6} - 96 e^{-6} - 16 e^{-6}) \right\}$

$= \frac{1}{16} \times 400 e^{-6} = 25 e^{-6}$

$= 25 e^{-6}$