

# Machine Learning 1 WS18/19

## Submission for Exercise Sheet 2

Hendrik Makait 384968

Michael Hoppe 362514

Wai Tang Victor Chan 406094

Rudi Poepsel Lemaitre 373017

Jonas Piotrowski 399334

Aki Saksala 399293

## HL 1 Exercise 2

1a) for independence:  $p(x, y) = p(x)p(y) \quad \forall x, y \in \mathbb{R}^2$

$$\begin{aligned} p(x) &= \int_0^{\infty} p(x, y) dy = \int_0^{\infty} \lambda \eta e^{-\lambda x - \eta y} dy \\ &= \lambda \eta e^{-\lambda x} \int_0^{\infty} e^{-\eta y} dy = \lambda \eta e^{-\lambda x} \left[ \frac{-1}{\eta} e^{-\eta y} \right]_0^{\infty} = \lambda \eta e^{-\lambda x} \end{aligned}$$

$$\begin{aligned} p(y) &= \int_0^{\infty} p(x, y) dx = \int_0^{\infty} \lambda \eta e^{-\lambda x - \eta y} dx = \lambda \eta e^{-\eta y} \int_0^{\infty} e^{-\lambda x} dx \\ &= \lambda \eta e^{-\eta y} \left[ \frac{-1}{\lambda} e^{-\lambda x} \right]_0^{\infty} = \eta e^{-\eta y} \end{aligned}$$

$$\Rightarrow p(x, y) = p(x)p(y) = \lambda \eta e^{-\lambda x - \eta y} \quad \square$$

1b) Likelihood-funktion:  $p(D | \lambda, \eta) = \prod_{i=1}^N p(x_i, y_i | \lambda) = \prod_{i=1}^N \lambda \eta e^{-\lambda x_i - \eta y_i}$

$$\ell(\lambda) = \log(p(D | \lambda, \eta)) = -$$

$$= \sum_{i=1}^N (\log(\lambda) + \log(\eta) - \lambda x_i - \eta y_i)$$

$$= N(\log(\lambda) + \log(\eta)) - \sum_{i=1}^N (\lambda x_i + \eta y_i)$$

$$= N(\log(\lambda \eta)) - \sum_{i=1}^N (\lambda x_i + \eta y_i)$$

$$\arg \max_{\lambda} \ell(\lambda) \Rightarrow \frac{\partial \ell(\lambda)}{\partial \lambda} = \frac{N}{\lambda} - \sum_{i=1}^N x_i \stackrel{!}{=} 0$$

$$\Rightarrow \frac{N}{\lambda} = \sum_{i=1}^N x_i$$

$$\Rightarrow \hat{\lambda} = \frac{\sum_{i=1}^N x_i}{N}$$

A

$$1) c) \quad p(D | \lambda, \eta) = \prod_1^N p(x_i, y_i | \lambda) = \prod_1^N \lambda \eta e^{-\lambda x_i - \eta y_i}$$

$$\Rightarrow \eta = \frac{1}{\lambda} \Rightarrow p(D | \lambda) = \prod_1^N e^{-\lambda x_i - \frac{1}{\lambda} y_i}$$

$$\ell(\lambda) = \log p(D | \lambda) = \sum_1^N -\lambda x_i - \frac{1}{\lambda} y_i$$

$$\frac{d\ell(\lambda)}{d\lambda} = -\sum_1^N x_i + \frac{1}{\lambda^2} \sum_1^N y_i \stackrel{!}{=} 0$$

$$\lambda^2 = \frac{\sum_1^N y_i}{\sum_1^N x_i}$$

$$\lambda = \sqrt{\frac{\sum_1^N y_i}{\sum_1^N x_i}}$$

$$1d) \quad p(D | \lambda, \eta) = \prod_1^N p(x_i, y_i | \lambda) = \prod_1^N \lambda \eta e^{-\lambda x_i - \eta y_i}$$

$$\Rightarrow \eta = 1 - \lambda \Rightarrow p(D | \lambda) = \prod_1^N \lambda (1 - \lambda) e^{-\lambda x_i - y_i + \lambda y_i}$$

$$\ell(\lambda) = \sum_1^N \log(\lambda - \lambda^2) - \lambda x_i - y_i + \lambda y_i$$

$$= N \log(\lambda - \lambda^2) - N y_i + \sum_1^N \lambda y_i - \lambda x_i$$

$$\frac{d\ell(\lambda)}{d\lambda} = \frac{1}{\lambda} - \frac{1}{1-\lambda} + \sum_1^N y_i - x_i \stackrel{!}{=} 0$$

$$\sum_1^N y_i - x_i = \frac{1}{1-\lambda} - \frac{1}{\lambda}$$

$$\sum_1^N y_i - x_i = \frac{\lambda - 1 + \lambda}{\lambda(1-\lambda)}$$

$$\lambda \sum_1^N y_i - x_i - \lambda^2 \sum_1^N y_i - x_i = 2\lambda - 1$$

$$\lambda^2 \sum_1^N (y_i - x_i) - \lambda (2 - \sum_1^N (y_i - x_i)) + 1 = 0$$

$$\lambda = \frac{\sum_1^N (y_i - x_i) - 2 \pm \sqrt{4 - 4\sum_1^N (y_i - x_i) + \sum_1^N (y_i - x_i)^2 + 4 \cdot \sum_1^N (y_i - x_i) \cdot 1}}{2 \sum_1^N (y_i - x_i)}$$

$$\lambda = \frac{\sum_1^N (y_i - x_i) - 2 \pm \sqrt{4 + (\sum_1^N (y_i - x_i))^2}}{2 \sum_1^N (y_i - x_i)}$$

## ML Exercise 2

②

a)

$$\begin{aligned}P(D|\theta) &= \prod_{n=1}^N p(x_n|\theta) \\&= p(\text{head}|\theta)^5 p(\text{tail}|\theta)^2 \\&= \theta^5 (1-\theta)^2 \\&= \theta^5 \cdot (\theta^2 - 2\theta + 1)\end{aligned}$$

$$P(D|\theta) = \theta^7 - 2\theta^6 + \theta^5$$

b)

~~$\hat{\theta}$~~

$$0 = 7\theta^6 - 12\theta^5 + 5\theta^4$$

$$\hat{\theta} = \frac{5}{7}$$

$$\begin{aligned}P(x_8 = \text{head}, x_9 = \text{head} | \hat{\theta}) &= p(\text{head} | \hat{\theta})^2 \\&= \left(\frac{5}{7}\right)^2\end{aligned}$$

$$= \frac{25}{49} \approx \underline{\underline{0,51}}$$

c)

$$p(\theta) = \begin{cases} 1 & \text{if } 0 \leq \theta \leq 1 \\ 0 & \text{else} \end{cases}$$

$$p(\theta|D) = \frac{P(D|\theta) p(\theta)}{P(D)}$$

$$p(\theta|D) = \frac{P(D|\theta) p(\theta)}{\int P(D|\theta) p(\theta) d\theta}$$

$$p(\theta|D) = \frac{\prod_{n=1}^N p(x_n|\theta) \cdot 1}{\int \prod_{n=1}^N p(x_n|\theta) \cdot 1 d\theta} = \frac{p(\text{head}|\theta)^5 p(\text{tail}|\theta)^2}{\int p(\text{head}|\theta)^5 p(\text{tail}|\theta)^2 d\theta}$$

$$= \frac{\theta^5 (1-\theta)^2}{\int_0^1 \theta^5 (1-\theta)^2 d\theta}$$

$$\int_0^1 \theta^5 (1-\theta)^2 d\theta$$

$$= \int_0^1 \theta^7 - 2\theta^6 + \theta^5$$

$$= \frac{1}{8} \cdot 1^8 - \frac{2}{7} \cdot 1^7 + \frac{1}{6} \cdot 1^6$$

$$= \frac{1}{8} - \frac{2}{7} + \frac{1}{6}$$

$$= \frac{1}{168} \approx 0,0059524$$

$$p(\theta|D) = \frac{\theta^5 (1-\theta)^2}{\frac{1}{168}} = \underline{\underline{168 \theta^5 (1-\theta)^2}}$$

$$p(x|D) = \int p(x|\theta) p(\theta|D) d\theta$$

$$\cancel{p(x_2)} \quad \int P(x_8 = \text{head}, x_9 = \text{head} | \theta) p(\theta|D) d\theta$$

$$= \int p(\text{head}|\theta)^2 \cdot 168 \theta^5 (1-\theta)^2 d\theta$$

$$= \int 168 \theta^5 (1-\theta)^2 \theta^2 d\theta$$

$$= \int 168 \cdot (\theta^7 - 2\theta^6 + \theta^5) \cdot \theta^2 d\theta$$

$$= \int (168 \theta^9 - 336 \theta^8 + 168 \theta^7) d\theta$$

$$= \int_0^1 168 \theta^9 - 336 \theta^8 + 168 \theta^7 d\theta$$

$$= \frac{168}{10} 1^{10} - \frac{336}{9} 1^9 + \frac{168}{8} 1^8$$

$$= \frac{84}{5} - \frac{112}{3} + 21$$

$$\underline{\underline{\approx 0,4667}}$$



Exercise 3:

a) Show:  $\sigma_n^2 \leq \min\left(\frac{\sigma^2}{n}, \sigma_0^2\right)$

It is given:  $\frac{1}{\sigma_n^2} = \frac{n}{\sigma^2} + \frac{1}{\sigma_0^2}$ ,  $\sigma^2, \sigma_0^2 > 0$

(i)  $\sigma_0^2 > 0 \Rightarrow \frac{1}{\sigma_n^2} \geq \frac{n}{\sigma^2} \xLeftrightarrow{\sigma_n^2, \sigma_0^2 > 0} \frac{1}{n} \geq \sigma_n^2$

as well as

(ii)  $\sigma_0^2 > 0 \Rightarrow \frac{1}{\sigma_n^2} \geq \frac{1}{\sigma_0^2} \xLeftrightarrow{\sigma_n^2, \sigma_0^2 > 0} \sigma_0^2 \geq \sigma_n^2$

(i), (ii)  $\Rightarrow \sigma_0^2 \geq \sigma_n^2$  and  $\sigma^2 \geq \sigma_n^2 \Rightarrow \min(\sigma^2, \sigma_0^2) \geq \sigma_n^2$  q.e.d.

b) Show:  $\min(\hat{\mu}_n, \mu_0) \leq \mu_n \leq \max(\hat{\mu}_n, \mu_0)$

with given (1):  $\frac{1}{\sigma_n^2} = \frac{n}{\sigma^2} + \frac{1}{\sigma_0^2}$ , (2):  $\frac{\mu_n}{\sigma_n^2} = \frac{n}{\sigma^2} \hat{\mu}_n + \frac{\mu_0}{\sigma_0^2}$

Case 1:  $\hat{\mu}_n \geq \mu_0 \Leftrightarrow \hat{\mu}_n = \max(\hat{\mu}_n, \mu_0), \mu_0 = \min(\hat{\mu}_n, \mu_0)$

It is given: (1)  $\frac{\mu_n}{\sigma_n^2} = \frac{n}{\sigma^2}$

(i) (2):  $\frac{\mu_n}{\sigma_n^2} = \frac{n}{\sigma^2} \hat{\mu}_n + \frac{\mu_0}{\sigma_0^2} \geq \frac{n}{\sigma^2} \mu_0 + \frac{1}{\sigma_0^2} \mu_0 \stackrel{(1)}{=} \frac{1}{\sigma_n^2} \mu_0$

$\Leftrightarrow \frac{\mu_n}{\sigma_n^2} \geq \frac{\mu_0}{\sigma_n^2} \Leftrightarrow \mu_n \geq \mu_0 = \min(\hat{\mu}_n, \mu_0)$

(ii) (2):  $\frac{\mu_n}{\sigma_n^2} = \frac{n}{\sigma^2} \hat{\mu}_n + \frac{\mu_0}{\sigma_0^2} \leq \frac{n}{\sigma^2} \hat{\mu}_n + \frac{1}{\sigma_0^2} \hat{\mu}_n \stackrel{(1)}{=} \frac{1}{\sigma_n^2} \hat{\mu}_n$

$\Leftrightarrow \frac{\mu_n}{\sigma_n^2} \leq \frac{\hat{\mu}_n}{\sigma_n^2} \Leftrightarrow \mu_n \leq \hat{\mu}_n = \max(\hat{\mu}_n, \mu_0)$



Case 2:  $\hat{\mu}_n < \mu_0 \Leftrightarrow \hat{\mu}_n = \min(\hat{\mu}_n, \mu_0)$ ,  $\mu_0 = \max(\hat{\mu}_n, \mu_0)$

$$(iii) (2) \frac{\mu_n}{\sigma_n^2} = \frac{n}{\sigma_n^2} \hat{\mu}_n + \frac{\mu_0}{\sigma_0^2} \geq \frac{n}{\sigma_n^2} \hat{\mu}_n + \frac{1}{\sigma_0^2} \hat{\mu}_n \stackrel{(i)}{=} \frac{1}{\sigma_n^2} \hat{\mu}_n$$

$$\Leftrightarrow \frac{\mu_n}{\sigma_n^2} \geq \frac{\hat{\mu}_n}{\sigma_n^2} \Leftrightarrow \mu_n \geq \hat{\mu}_n = \min(\hat{\mu}_n, \mu_0)$$

$$(iv) (2) \frac{\mu_n}{\sigma_n^2} = \frac{n}{\sigma_n^2} \hat{\mu}_n + \frac{\mu_0}{\sigma_0^2} \leq \frac{n}{\sigma_n^2} \mu_0 + \frac{1}{\sigma_0^2} \mu_0 \stackrel{(i)}{=} \frac{1}{\sigma_n^2} \mu_0$$

$$\Leftrightarrow \frac{\mu_n}{\sigma_n^2} \leq \frac{\mu_0}{\sigma_n^2} \Leftrightarrow \mu_n \leq \mu_0 = \max(\hat{\mu}_n, \mu_0)$$

From (i) to (iv) it follows that

$$\min(\hat{\mu}_n, \mu_0) \leq \mu_n \leq \max(\hat{\mu}_n, \mu_0) \quad \text{q.e.d.}$$