ML4 Supplemental Results in rMarkdown

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Author note

This script generates supplemental analyses not reported in the main ML4 manuscript. To knit this document you must install the papaja package from GitHub.

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# 1 Supplemental Results

This document presents supplementary analyses that consider ratings of the pro-US and anti-US authors as two separate DVs, rather than as a single difference score DV as in the main text.

The preregistered three-level meta-analyses frequently failed to converge because most sites ran only one study. In a divergence from the preregistration, we fit random-effects meta-analyses instead.

## 1.1 Pro-US Author Ratings and Anti-US Author Ratings separately

The primary finding of interest from Greenberg et al. (1994) was that participants who underwent the mortality salience treatment showed greater preference for the pro-US essay author compared to the anti-US essay author. In the main article, we report our DVs with these separately averaged, and then anti-US scores subtracted from pro-US scores. In this section we instead treat these as two separate DVs and repeat the primary analyses reported in the manuscript.

### 1.1.1 PRO-US RATINGS ONLY:.

#### 1.1.1.1 Research question 1: Meta-analytic results across all labs.

A random-effects meta-analysis was conducted using the MetaSEM package (Cheung, 2014) in R.[[1]](#footnote-21) This analysis produces the grand mean effect size across all sites and versions.

First we examine if there was an overall effect of the mortality salience manipulation on ratings of the pro-US author. Regardless of which exclusion criteria were used, we did not observe an effect: Exclusion set 1: *Hedges’ g* = 0.00, 95% CI = [-0.10, 0.10], *SE* = 0.05, *Z* = -0.03, *p* = .977. Exclusion set 2: *Hedges’ g* = 0.04, 95% CI = [-0.07, 0.16], *SE* = 0.06, *Z* = 0.72, *p* = .470. Exclusion set 3: *Hedges’ g* = 0.04, 95% CI = [-0.08, 0.16], *SE* = 0.06, *Z* = 0.60, *p* = .551.

We also examined how much variation was observed among effect sizes (e.g., heterogeneity). For example, there may have been a mortality salience effect at some sites and not others. For all exclusion sets, this variability did not exceed variability we would expect by chance: Exclusion set 1: *Q*(16) = 13.80, *p* = .614; Exclusion set 2: *Q*(16) = 12.68, *p* = .696; Exclusion set 3: *Q*(16) = 13.97, *p* = .601.

In sum, we observed no evidence for an overall effect of mortality salience on pro-US author ratings. Additionally, results indicate no detected heterogeneity in effect sizes across sites. This lack of variation suggests that it is unlikely we will observe an effect of Author Advised versus In House protocols or other moderators such as differences in samples or TMT knowledge. Even so, the plausible moderation by Author Advised/In House protocol is examined in the following section.

#### 1.1.1.2 Research Question 2: Moderation by Author Advised/In House protocol.

A covariate of protocol type was added to the random effects model to create a mixed-effects meta-analysis. This was pre-registered as our primary analysis.[[2]](#footnote-23)

This analysis again produces an overall grand mean effect size, and those were again not statistically significant across all three exclusion sets: Exclusion set 1: *Hedges’ g* = 0.00, 95% CI = [-0.10, 0.10], *SE* = 0.05, *Z* = 0.01, *p* = .989. Exclusion set 2: *Hedges’ g* = 0.08, 95% CI = [-0.04, 0.21], *SE* = 0.06, *Z* = 1.30, *p* = .193. Exclusion set 3: *Hedges’ g* = 0.12, 95% CI = [-0.03, 0.27], *SE* = 0.08, *Z* = 1.58, *p* = .114.

In addition, protocol version did not significantly predict replication effect size. Exclusion set 1: *b* = 0.02, 95% CI = [-0.08, 0.12], *SE* = 0.05, *Z* = 0.44, *p* = .661; Exclusion set 2: *b* = 0.11, 95% CI = [-0.02, 0.23], *SE* = 0.06, *Z* = 1.68, *p* = .093; Exclusion set 3: *b* = 0.14, 95% CI = [-0.01, 0.29], *SE* = 0.08, *Z* = 1.86, *p* = .063.

We again did not observe heterogeneity between labs in any of the exclusion sets: exclusion set 1: *Q*(16) = 13.80, *p* = .614; exclusion set 2: *Q*(16) = 12.68, *p* = .696; exclusion set 3: *Q*(16) = 13.97, *p* = .601.

#### 1.1.1.3 Research Question 3: Effect of Standardization.

Finally, we tested whether In House protocols displayed greater variability in effect size than Author Advised protocols. To test this hypothesis, we ran the mixed-effects models but constrained the variance to 0, creating fixed-effects models. These models were then compared with a chi-squared differences test to assess whether the fit significantly changed. In this case, none of the three models significantly decreased in fit: Exclusion set 1: *²* (1) = 0.00, *p* = 1; Exclusion set 2: *²* (1) = 0.00, *p* = 1; Exclusion set 3: *²* (1) = 0.00, *p* = 1. Overall, there was no evidence that In House protocols elicited greater variability than Author Advised protocols.

### 1.1.2 ANTI-US RATINGS ONLY:.

#### 1.1.2.1 Research question 1: Meta-analytic results across all labs.

A random-effects meta-analysis was conducted using the MetaSEM package (Cheung, 2014) in R.[[3]](#footnote-27) This analysis produces the grand mean effect size across all sites and versions.

First we examine if there was an overall effect of the mortality salience manipulation on ratings of the anti-US author. There was no significant effect across exclusion sets: Exclusion set 1: *Hedges’ g* = -0.09, 95% CI = [-0.21, 0.02], *SE* = 0.06, *Z* = -1.62, *p* = .105. Exclusion set 2: *Hedges’ g* = -0.09, 95% CI = [-0.22, 0.04], *SE* = 0.07, *Z* = -1.34, *p* = .179. Exclusion set 3: *Hedges’ g* = -0.09, 95% CI = [-0.23, 0.06], *SE* = 0.07, *Z* = -1.18, *p* = .239.

We also examined how much variation was observed among effect sizes (e.g., heterogeneity). For example, there may have been a mortality salience effect at some sites and not others. We generally observed little evidence for heterogeneity: exclusion set 1: *Q*(16) = 21.11, *p* = .174; exclusion set 2: *Q*(16) = 23.15, *p* = .110; exclusion set 3: *Q*(16) = 20.92, *p* = .182.

In sum, we observed no evidence for an overall effect of mortality salience on anti-us author ratings. Additionally, overall results suggest that there was minimal or no heterogeneity in effect sizes across sites. This lack of variation suggests that it is unlikely we will observe an effect of Author Advised versus In House protocols or other moderators such as differences in samples or TMT knowledge. Even so, the plausible moderation by Author Advised/In House protocol is examined in the following section.

#### 1.1.2.2 Research Question 2: Moderation by Author Advised/In House protocol.

A covariate of protocol type was added to the random effects model to create a mixed-effects meta-analysis. This was pre-registered as our primary analysis.[[4]](#footnote-29)

This analysis again produces an overall grand mean effect size, and those were again null across all three exclusion sets: Exclusion set 1: *Hedges’ g* = -0.09, 95% CI = [-0.21, 0.02], *SE* = 0.06, *Z* = -1.61, *p* = .107. Exclusion set 2: *Hedges’ g* = -0.09, 95% CI = [-0.23, 0.05], *SE* = 0.07, *Z* = -1.23, *p* = .220. Exclusion set 3: *Hedges’ g* = -0.08, 95% CI = [-0.24, 0.09], *SE* = 0.09, *Z* = -0.90, *p* = .369.

In addition, protocol version did not significantly predict replication effect size. Exclusion set 1: *b* = 0.00, 95% CI = [-0.11, 0.11], *SE* = 0.06, *Z* = 0.02, *p* = .982; exclusion set 2: *b* = 0.01, 95% CI = [-0.14, 0.15], *SE* = 0.07, *Z* = 0.08, *p* = .936; exclusion set 3: *b* = 0.02, 95% CI = [-0.15, 0.19], *SE* = 0.09, *Z* = 0.21, *p* = .837.

We again did not observe heterogeneity between labs, in any of the exclusion sets: exclusion set 1: *Q*(16) = 21.11, *p* = .174; exclusion set 2: *Q*(16) = 23.15, *p* = .110; exclusion set 3: *Q*(16) = 20.92, *p* = .182.

#### 1.1.2.3 Research Question 3: Effect of Standardization.

Finally, we tested whether In House protocols displayed greater variability in effect size than Author Advised protocols. To test this hypothesis, we ran the mixed-effects models but constrained the variance to 0, effectively creating fixed-effects models. These models were then compared with a chi-squared differences test to assess whether the fit significantly changed. In this case, none of the three models significantly decreased in fit: Exclusion set 1: *²* (1) = 0.46, *p* = 0.50; Exclusion set 2: *²* (1) = 0.93, *p* = 0.33; Exclusion set 3: *²* (1) = 1.11, *p* = 0.29. Overall, there was no evidence that In House protocols elicited greater variability than Author Advised protocols.

1. Sample code to run this analysis is: meta(y=es, v=var, data=dataset). In this sample code, “y=es” directs the program to the column of effect sizes, and “v=var” indicates the variable to be used as the sampling variance for each effect size. [↑](#footnote-ref-21)
2. The addition of the argument “x = expert.ctr” to the prior metaSEM R code can be seen here: meta(y=es, v=var, x=expert.ctr, data=dataset) [↑](#footnote-ref-23)
3. Sample code to run this analysis is: meta(y=es, v=var, data=dataset). In this sample code, “y=es” directs the program to the column of effect sizes, and “v=var” indicates the variable to be used as the sampling variance for each effect size. [↑](#footnote-ref-27)
4. The addition of the argument “x = version” to the prior metaSEM R code can be seen here: meta(y=es, v=var, x=version, data=dataset) [↑](#footnote-ref-29)