

## Recall

Continuous rvs, pdfs, cdfs, etc.

$$X \text{ w/ pdf } f_X(x) - F_X(x) = P\{X \leq x\} = \int_{-\infty}^x f_X(t) dt$$

Note: every random variable has a cdf, even discrete ones.

$$F_X(x) = P\{X \leq x\}$$

$X$  is Gaussian (normal) w/ mean  $\mu$ , variance  $\sigma^2$  when

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$$\text{CDF of a Gaussian} = \int_{-\infty}^x f_X(t) dt$$

Special Case:  $\mu=0, \sigma=1$  then get

$$\Phi(x) = \int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt$$

as cdf.

## Example - Communication Channel

Want to send one bit through a noisy channel. Encode a 1 as  $+C$ , a 0 as  $-C$  (signal levels).

Thus

$$X = \text{input to channel} = \begin{cases} C \text{ w/ probability } p \\ -C \text{ w/ probability } 1-p \end{cases} \quad (\text{after } p = \frac{1}{2} \text{ assumed})$$

Channel output

$$Y = X + N.$$

where  $N = \text{noise} - \text{model as a zero-mean, some variance } \sigma^2$

Decoding Rule: Decide  $+C$  was sent when  $Y > 0$ , decide  $-C$  was sent when  $Y < 0$  ← optimal for  $p = \frac{1}{2}$ . Else would need to skew towards one result.

What is  $P(\text{error in decoding})$ ?

It's

similar calculation gives  
 $1 - \Phi\left(\frac{C}{\sigma}\right)$  as well

$$P(\text{error} | C \text{ sent}) p + P(\text{error} | -C \text{ sent})(1-p)$$

$$= P(Y > 0 | X = C) = P(-C + N > 0) = P(N > -C)$$

Idea: get an error when  $N$  is positive enough to "push the  $-C$  into the 'decode as  $+C$ ' zone"

Get an appropriate value for  $P(N > -C)$  from  $\Phi$ -table -  $N$  zero mean,  $r^2 \Rightarrow N/r$  standard normal,  $\mu=0$ ,  $r=1$

Hence

$$P(N > -r^2) = P\left(\frac{N}{r} > \frac{-C}{r}\right) = 1 - \Phi\left(\frac{C}{r}\right)$$

**Comment:** Turns out that if  $X$  and  $Y$  are Gaussian defined on same probability space  $\Omega, \mathcal{P}$ , then any linear combo  $Z = c_1 X + c_2 Y$  is also Gaussian.  $\leftarrow$  More on this later

One other CDF-type thing

### Example - Experimental Meets Geometric Pt. 2

Consider first  $T$ , an  $\text{exponential}(\lambda)$  rv

Then

$$f_T(t) = \begin{cases} \lambda e^{-\lambda t}, & t \geq 0 \\ 0, & t < 0 \end{cases}$$

and

$$F_T(t) = \int_{-\infty}^t f_T(\tau) d\tau = \begin{cases} 1 - e^{-\lambda t}, & t \geq 0 \\ 0, & t < 0 \end{cases}$$

Next, consider the discrete  $\sim X$  obtained as follows:

- flip a  $p$ -coin in real time
- flip at times  $\delta, 2\delta, 3\delta, \dots$   $\delta$  small
- $X =$  the time of the first flip that is a heads

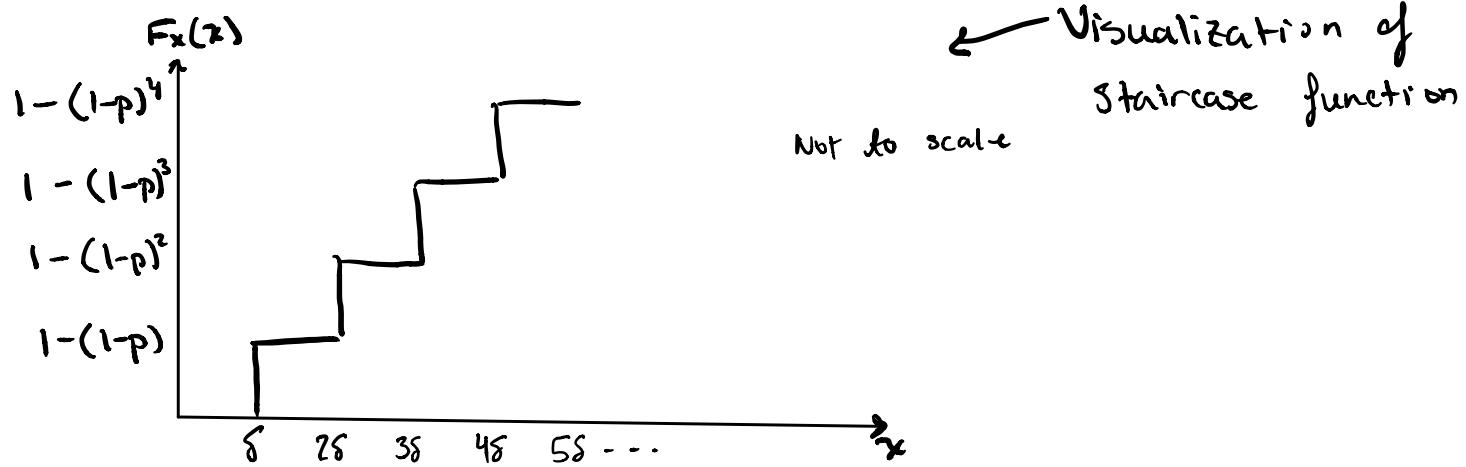
$\nwarrow k \text{ an integer}$

$$\forall k \geq 0, \quad \mathbb{P}(\{X=k\}) = p(1-p)^{k-1}.$$

CDF of  $X$ ?

$$F_X(x) = \mathbb{P}(\{X \leq x\}) = \sum_{m: m\delta < x} p(1-p)^m \quad \begin{array}{l} \nwarrow \text{Staircase function} \\ \text{of } x \\ \longleftarrow m \in \mathbb{Z} \end{array}$$

Value in the interval  $[k\delta, (k+1)\delta)$  is  $1 - (1-p)^k \quad \forall k \geq 0$



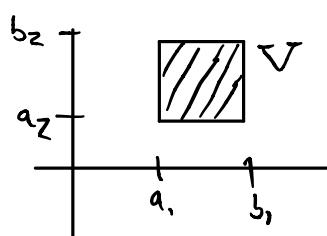
Observation: When  $\lambda = -\frac{\ln(1-p)}{s}$ , we have  $F_x(k\delta) = F_T(k\delta)$   $\forall k \geq 0$

Let  $\delta$  get smaller, staircase approximation gets better; etc.

Moving along - say  $X, Y$  rvs defined in same  $\Omega, \mathcal{P}$  are jointly continuous w/ joint pdf  $f_{X,Y}(x,y)$  when

$$\mathbb{P}(\{(X,Y) \in V\}) = \iint_V f_{X,Y}(x,y) dx dy \quad \forall V \subset \mathbb{R}^2$$

Special case of a  $V$ :  $[a_1, b_1] \times [a_2, b_2]$



Then

$$\mathbb{P}(\{(X,Y) \in V\}) = \int_{a_1}^{b_1} dx \int_{a_2}^{b_2} dy (f_{X,Y}(x,y))$$

Again, have marginals

$$f_X(x) = \int_{-\infty}^{+\infty} f_{X,Y}(x,y) dy \quad ; \quad f_Y(y) = \int_{-\infty}^{+\infty} f_{X,Y}(x,y) dx$$

An official way to get this: Get  $F_X(x)$  first then take  $\frac{d}{dx} F_X(x)$

$$F_X(x) = \mathbb{P}(\{X \leq x\}) = \mathbb{P}(\{(x, y) \in (-\infty, x] \times (-\infty, +\infty)\})$$

$$= \int_{-\infty}^x dt \int_{-\infty}^{+\infty} dy (f_{x,y}(t,y))$$

then

$$\frac{d}{dx} F_X(x) = \int_{-\infty}^{+\infty} dy (f_{x,y}(x,y))$$

Could also derive marginal formulas as follows:

$$\forall V \subset \mathbb{R}, \mathbb{P}(\{X \in V\}) = \mathbb{P}(\{(x, y) \in V \times (-\infty, \infty)\})$$

$$= \int_V dx \int_{-\infty}^{+\infty} dy (f_{x,y}(x,y)) = \int_V \left( \int_{-\infty}^{+\infty} f_{x,y}(x,y) dy \right) dx$$

Must be  $f_X(x)$  integrate over  $V$  to  
get  $\mathbb{P}(\{X \in V\})$

Other stuff

$$- \int_{-\infty}^{+\infty} dx \int_{-\infty}^{+\infty} dy (f_{x,y}(x,y)) = 1$$

- Joint CDF;  $F_{x,y}(x,y) = \mathbb{P}(\{X=x\} \cap \{Y=y\}) = \int_{-\infty}^x ds \int_{-\infty}^t dt (f_{x,y}(s,t))$

- $f_{x,y}(x,y) = \frac{\partial}{\partial x} \frac{\partial}{\partial y} F_{x,y}(x,y)$

- generalization to  $> 2$  rvs pretty straightforward

As for discrete rvs, joint determines the marginals; NOT vice-versa.

### Example - Joint PDF

One example of a joint PDF: given any set  $\mathcal{V}_0 \subset \mathbb{R}^2$  w/ nonzero area, say  $(X, Y)$  jointly uniform on  $\mathcal{V}_0$  when

$$f_{X,Y} = \begin{cases} \frac{1}{\text{area}(\mathcal{V}_0)}, & (x, y) \in \mathcal{V}_0 \\ 0, & \text{otherwise} \end{cases}$$

Next,

### Conditional Stuff For Continuous Random Variables

Given a continuous rv  $X$  on  $\Omega, \mathbb{P}$  and some event  $A \subset \Omega$ , the conditional pdf of  $X$  given  $A$  "defined" as follows:

For any  $\mathcal{V} \subset \mathbb{R}$ , we have

$$\mathbb{P}(\{X \in \mathcal{V}\} | A) = \int_{\mathcal{V}} f_{X|A}(x) dx$$

In general, no decent formula for  $f_{X|A}(x)$  in terms of  $f_X(x)$ .

One way to compute it:

- First get conditional cdf of  $x$  given  $A$

$$F_{X|A} = \mathbb{P}(\{X \leq x\} | A)$$

- Then take  $d/dx$  to get  $f_X(x)$

However, if  $A$  is an event of the form  $\{X \in W\}$ , and  $P(A) > 0$ , we have

$$f_{X|A}(x) = \begin{cases} \frac{f_x(x)}{P(\{X \in W\})}, & \text{when } X \in W \\ 0, & \text{otherwise} \end{cases}$$

How does this arise?

$$P(\{X \in V\}|A) = \frac{P(\{X \in (V \cap W)\})}{P(\{X \in W\})} = \frac{\int_V f_x(x) dx}{P(\{X \in W\})}$$

Define the indicator function of  $W$  via

"chi"  $\rightarrow \chi_W(x) = \begin{cases} 1, & x \in W \\ 0, & x \notin W \end{cases}$

So,

$$\frac{\int_V f_x(x) dx}{P(\{X \in W\})} = \int_V \left( \frac{f_x(x) \chi_W(x)}{P(\{X \in W\})} \right) dx = \begin{cases} \frac{f_x(x)}{P(\{X \in W\})}, & x \in W \\ 0, & \text{else} \end{cases}$$

must be  
 $f_{X|A}(x)$