

# ECE 4110 Homework 7

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*Due by 5pm on November 14*

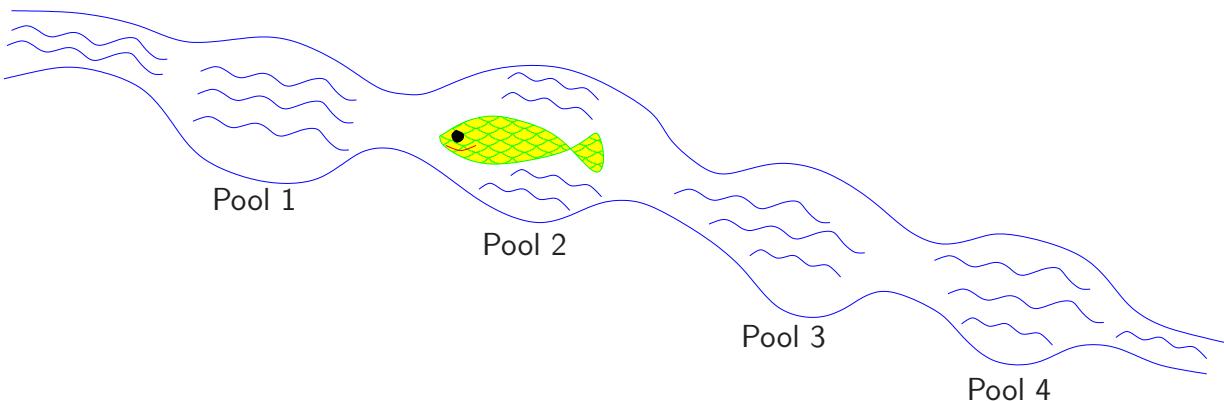
## 1 Reading Material

- The  $n$ -step transition probabilities and the state probabilities (Chapter 11.2.1 and 11.2.2).
- Stationary distribution (Chapter 11.2.3).
- Communication classes (Chapter 11.3.1).

## 2 Assignment

### 1. A Fish Called Wanda

A fish called Wanda is swimming among 4 pools (see the figure below). When Wanda changes pools (at discrete time instants), she goes downstream with probability  $p$  and upstream with probability  $1 - p$  when she has a choice. At pool 1, she stays with probability  $1 - p$  and moves to pool 2 with probability  $p$ . At pool 4, she stays with probability  $p$  and moves to pool 3 with probability  $1 - p$ .



- Suppose that Wanda is initially in Pool 1. Find out the distribution of Wanda's location after two time instants.
- Find the stationary distribution of Wanda's location.

## 2. Ferris Bueller's Day Off

Three out of every four trucks on the road are followed by a car, while one out of every five cars is followed by a truck. What fraction of vehicles on the road are trucks?

*Hint: Imagine sitting on the side of the road watching vehicles go by (sounds more interesting than going to a class on random processes!). Set the flow of the vehicles as a Markov chain.*

## 3. The Ehrenfest Model of Diffusion

A cubical volume is partitioned into two equal-sized chambers, a left chamber and a right chamber, by a barrier containing a small hole. In this volume there are  $M$  gas molecules. Suppose that at time 0, all of the molecules are in the left chamber. At each subsequent discrete time, one of the molecules, uniformly chosen from among the  $M$ , switches chambers. Let  $X_n$  denote the number of molecules in the left chamber at time  $n$ .

- (a) Argue that  $\{X_n\}_{n \geq 0}$  is Markov and draw its state transition diagram.
- (b) Show that  $\pi = [\pi_0, \dots, \pi_M]$  with

$$\pi_i = 2^{-M} \binom{M}{i}, \quad i = 0, \dots, M$$

is a stationary distribution of the chain.

## 4. Binomial Counting Process as a Markov Chain

Consider the Binomial counting process with parameter  $p$ .

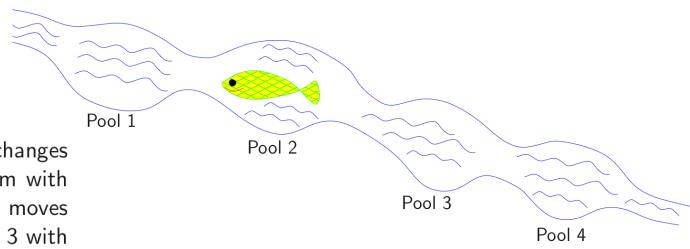
- (a) Show that it is a Markov chain and write out the transition matrix.
- (b) What are the communication classes? Is the chain irreducible?

## 5. Communication Class

For each of the Markov chains with transition probability matrices given below, draw the state transition diagrams and identify all communication classes. Is the Markov chain irreducible?

$$\begin{aligned} \mathbf{P}_1 &= \begin{pmatrix} 0 & 1 & 0 \\ \frac{1}{2} & 0 & \frac{1}{2} \\ 1 & 0 & 0 \end{pmatrix} & \mathbf{P}_2 &= \begin{pmatrix} 1 & 0 & 0 \\ 0 & 0 & 1 \\ 0 & 1 & 0 \end{pmatrix} & \mathbf{P}_3 &= \begin{pmatrix} \frac{1}{2} & \frac{1}{2} & 0 \\ 0 & 1 & 0 \\ \frac{1}{2} & 0 & \frac{1}{2} \end{pmatrix} \\ \mathbf{P}_4 &= \begin{pmatrix} 0 & \frac{1}{2} & \frac{1}{2} & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 1 & 0 \\ 1 & 0 & 0 & 0 \end{pmatrix} & \mathbf{P}_5 &= \begin{pmatrix} \frac{1}{2} & \frac{1}{2} & 0 & 0 \\ 1 & 0 & 0 & 0 \\ \frac{1}{2} & 0 & \frac{1}{4} & \frac{1}{4} \\ 0 & \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{pmatrix} & \mathbf{P}_6 &= \begin{pmatrix} \frac{1}{2} & 0 & 0 & 0 & \frac{1}{2} \\ 0 & \frac{1}{2} & 0 & \frac{1}{2} & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{2} & 0 & 0 & 0 & \frac{1}{2} \end{pmatrix} \end{aligned}$$

# ① A Fish Called Wanda



A fish called Wanda is swimming among 4 pools (see the figure below). When Wanda changes pools (at discrete time instants), she goes downstream with probability  $p$  and upstream with probability  $1 - p$  when she has a choice. At pool 1, she stays with probability  $1 - p$  and moves to pool 2 with probability  $p$ . At pool 4, she stays with probability  $p$  and moves to pool 3 with probability  $1 - p$ .

(a) Suppose that Wanda is initially in pool 1. Find out the distribution of Wanda's location after two time instants

## Scenarios

$$\text{pool 1} \xrightarrow{1-p} \text{pool 1} \xrightarrow{1-p} \text{pool 1}$$

$$\text{pool 1} \xrightarrow{p} \text{pool 2} \xrightarrow{1-p} \text{pool 1}$$

$$\text{pool 1} \xrightarrow{1-p} \text{pool 1} \xrightarrow{p} \text{pool 2}$$

$$\text{pool 1} \xrightarrow{p} \text{pool 2} \xrightarrow{p} \text{pool 3}$$

$$\text{Wanda is at } \begin{cases} \text{pool 1 w.p. } (1-p)^2 + p(1-p) \\ \text{pool 2 w.p. } (1-p)p \\ \text{pool 3 w.p. } p^2 \end{cases}$$

We verify these probabilities sum to 1, since Wanda can only be at these 3 locations.

$$\rightarrow (1-p)^2 + p(1-p) + (1-p)p + p^2$$

$$= 1 - 2p + p^2 + p - p^2 + p - p^2 + p^2 = 1$$

Nice

It's nice going through cases and all, but now we do this more cleverly.

$$P(2) = p(0) \bar{P}^2$$

where

$$\bar{P} = \begin{bmatrix} 1-p & p & 0 & 0 \\ 1-p & 0 & p & 0 \\ 0 & 1-p & 0 & p \\ 0 & 0 & 1-p & p \end{bmatrix}$$

is the transition matrix

$$\bar{P}^2 = \begin{bmatrix} (1-p)^2 + p(1-p) & (1-p)p & p^2 & 0 \\ (1-p)^2 & 2p(1-p) & 0 & p^2 \\ (1-p)^2 & 0 & 2p(1-p) & p^2 \\ 0 & (1-p)^2 & p(1-p) & p(1-p) + p^2 \end{bmatrix}$$

$$p(0) = [1 \ 0 \ 0 \ 0]$$

$$p(0) \bar{P}^2 = (1-p)^2 + p(1-p) + p(1-p) + p^2$$

the pmf at time 2 is obtained by multiplying initial state by  $\bar{P}^2$ .

(b) Find the Stationary distribution of Warela's location

Need

Know  $\bar{\pi} P = \bar{\pi}$   
and thus

$$P^T \bar{\pi}^T = \bar{\pi}^T$$

$$P = \begin{bmatrix} 1-p & p & 0 & 0 \\ 1-p & 0 & p & 0 \\ 0 & 1-p & 0 & p \\ 0 & 0 & 1-p & p \end{bmatrix}$$

i.e.  $P^T$  has eigenvalue 1 w/ eigenvector  $\bar{\pi}^T$ .

$$(P^T - \lambda I) = 0 \quad \lambda=1$$

$$P^T - I = 0$$

i.e. want  $E_{\lambda=1}(P^T) = \ker(P^T - I)$   
 $= \bar{\pi}^T$

$$P^T - I = \begin{bmatrix} -p & 1-p & 0 & 0 \\ p & -1 & 1-p & 0 \\ 0 & p & -1 & 1-p \\ 0 & 0 & p & p-1 \end{bmatrix} \begin{bmatrix} a \\ b \\ c \\ d \end{bmatrix} \xleftarrow{\bar{\pi}^T} = \begin{bmatrix} 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}$$

$$P^T \bar{\pi} = 0$$

Want non-trivial solution

$$\bar{\pi} = [a \ b \ c \ d]$$

$$-pa + (1-p)b = 0$$

$$pa - b + (1-p)c = 0$$

$$pb - c + (1-p)d = 0$$

$$pc + (p-1)d = 0$$

$$(1-p)b = pb \quad a = \frac{1-p}{p}b$$

$$P\left(\frac{1-p}{p}\right)b - b + (1-p)c = 0$$

$$(1-p)b - b + (1-p)c = 0$$

$$b(1-p-1) + c(1-p) = 0$$

$$c = b \frac{p}{1-p}$$

$$P\left(\frac{p}{1-p}\right)b + (p-1)d = 0$$

$$d = -\frac{p^2 b}{(1-p)(p-1)} = \frac{p^2}{(1-p)^2} b$$

$\therefore \bar{\pi} = \left[ \frac{1-p}{p}b \quad b \quad \frac{p}{1-p}b \quad \frac{p^2}{(1-p)^2}b \right]$

Now

$$\frac{1-p}{p}b + b + \frac{p}{1-p}b + \frac{p^2}{(1-p)^2}b = 1$$

$$b \left( 1 + \frac{1-p}{p} + \frac{p}{1-p} + \frac{p^2}{(1-p)^2} \right) = 1$$

$$b = \frac{1}{1 + \frac{1-p}{p} + \frac{p}{1-p} + \frac{p^2}{(1-p)^2}}$$

## 2. Ferris Bueller's Day Off

Three out of every four trucks on the road are followed by a car, while one out of every five cars is followed by a truck. What fraction of vehicles on the road are trucks?

*Hint: Imagine sitting on the side of the road watching vehicles go by (sounds more interesting than going to a class on random processes!). Set the flow of the vehicles as a Markov chain.*

Think of sampling infinite flow at some random time to get initial state

$$P = \begin{matrix} & T & C \\ T & \left[ \begin{matrix} 1/4 & 3/4 \\ 1/5 & 4/5 \end{matrix} \right] \\ C & \end{matrix}$$

Macaulay<sup>2</sup>

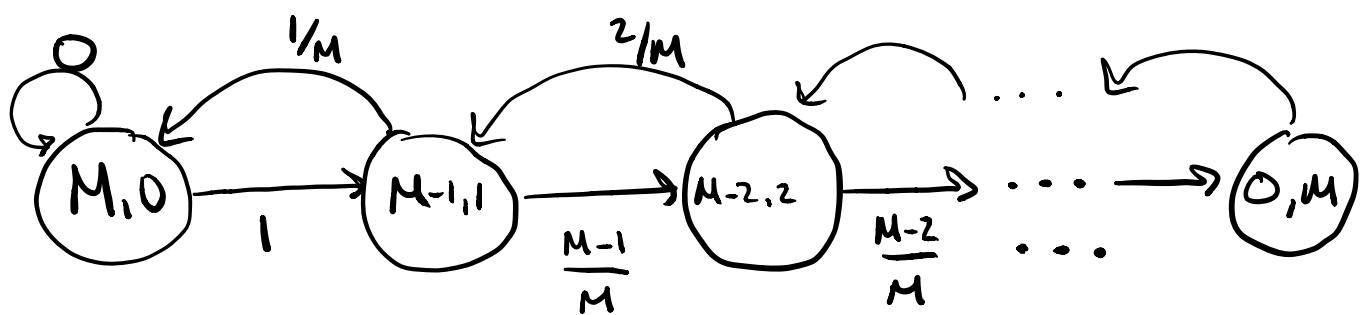
$$\lim_{n \rightarrow \infty} P^n = \begin{bmatrix} .210526 & .789474 \\ .210526 & .789474 \end{bmatrix}$$

### 3. The Ehrenfest Model of Diffusion

A cubical volume is partitioned into two equal-sized chambers, a left chamber and a right chamber, by a barrier containing a small hole. In this volume there are  $M$  gas molecules. Suppose that at time 0, all of the molecules are in the left chamber. At each subsequent discrete time, one of the molecules, uniformly chosen from among the  $M$ , switches chambers. Let  $X_n$  denote the number of molecules in the left chamber at time  $n$ .

(a) Argue that  $\{X_n\}_{n \geq 0}$  is Markov and draw a state-transition diagram.

The state diagram is as follows



$\{X_n\}_{n \geq 0}$  is intuitively Markov.

Let  $\{L_n\}$  denote amount in left half  
 $\{R_n\}$  denote amount in right half

Then

$$X_n = (L_n, R_n)$$

$$X_{n+1} = \begin{cases} (L_{n-1}, R_{n+1}) & \text{w.p. } \frac{L_n}{M} \\ (L_{n+1}, R_{n-1}) & \text{w.p. } \frac{R_n}{M} \end{cases}$$

Independence of uniform selection  $\Rightarrow X_n \perp\!\!\!\perp X_{n+1}$ .

(b) Show that  $\pi = [\pi_0, \dots, \pi_m]$  with

$$\pi_i = 2^{-M} \binom{M}{i} \quad i=0, \dots, M$$

is a stationary distribution of the chain.

$$P = \begin{bmatrix} 0 & 1 & 0 & \cdots & \cdots & 0 \\ \frac{1}{M} & 0 & \frac{M-1}{M} & \cdots & \cdots & 0 \\ 0 & \frac{2}{M} & 0 & \frac{M-2}{M} & \cdots & 0 \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & 0 & 1 & 0 \end{bmatrix}$$

$(m+1) \times (m+1)$   
matrix

Need  $\bar{\pi} P = \bar{\pi} \quad 2^{-M} \binom{M}{i}$

Base case,  $M=1$

$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix}$$

$$\pi_1 = \frac{1}{2}, \pi_2 = 0 \quad \bar{\pi} \text{ is length } M+1$$

$$\bar{\pi} P \stackrel{?}{=} \bar{\pi} = \begin{bmatrix} 1/2 \\ 1/2 \end{bmatrix} \quad \checkmark$$

Assume true for  $M=k \geq 1$

Need to show true for  $M=k+1$ .

True for  $m=k$

$$P_{k+1,k+1} = \begin{bmatrix} 0 & 1 & 0 & \cdots & 0 \\ \frac{1}{k} & 0 & \frac{k-1}{k} & \cdots & 0 \\ 0 & \frac{2}{k} & 0 & \frac{k-2}{k} & \cdots 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ 0 & 0 & 0 & \cdots 1 & 0 \end{bmatrix}$$

$$\bar{\pi} P = \bar{\pi} \quad \text{true by assumption}$$

thus

$$\bar{\pi} = \left[ 2^{-M}(1) \quad 2^{-M}\left(\frac{k}{2}\right) \cdots \cdot 2^{-M}\left(\frac{k}{k+1}\right) \right]$$

For  $M=k+1$ ,

$$\bar{\pi} = \left[ 2^{-M}(1) \cdots 2^{-M}\left(\frac{k}{k+1}\right) \quad 2^{-M}\left(\frac{k+1}{k+2}\right) \right]$$

$P_{k+2,k+2}$  has same structure as  $P_{k,k}$  but one more row and column since one more molecule.

It follows that  $\bar{\pi}_{k+2} = \bar{\pi}_{k+2} P_{k+2,k+2}$ .

(Could have really done nitty gritty but quite intuitive to see).

#### 4. Binomial Counting Process as a Markov Chain

Consider the Binomial counting process with parameter  $p$ .

(a) Show that this is a Markov chain and write out the transition matrix.

$$X_{n+1} | X_n, X_{n-1}, \dots, X_1 \sim \text{Bernoulli}(p)$$

$$X_{n+1} | X_n \sim \text{Bernoulli}(p)$$

Thus  $\Pr[X_{n+1} = x_{n+1} | X_n = x_n, \dots, X_1 = x_1] = \Pr[X_{n+1} = x_{n+1}] = \Pr[X_{n+1} = x_{n+1} | X_n = x_n]$

$$P = \begin{bmatrix} 1-p & p & 0 & \dots & \dots & \dots & 0 & \dots \\ 0 & 1-p & p & 0 & \dots & 0 & \dots & \dots \\ 0 & 0 & \ddots & \ddots & \ddots & \ddots & \ddots & \ddots \\ 0 & 0 & 0 & 1-p & p & \dots & \dots & \dots \\ 0 & 0 & 0 & 0 & \ddots & \ddots & \ddots & \ddots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \ddots & \ddots & \ddots \end{bmatrix}$$

(b) What are the communication classes? Is the chain irreducible?

Communication classes are

$$\{0\}, \{1\}, \{2\}, \dots$$

thus the chain is NOT reducible.

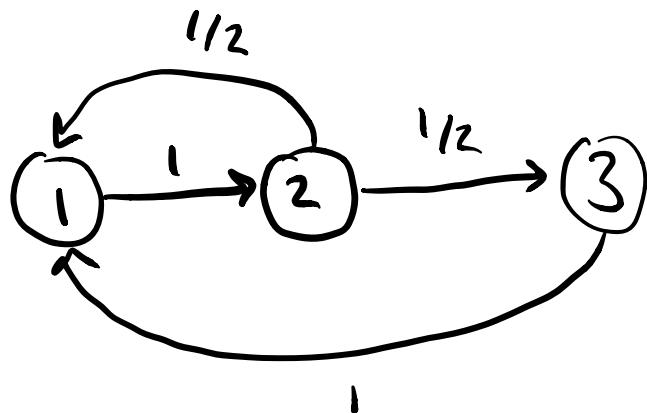
## 5. Communication Class

For each of the Markov chains with transition probability matrices given below, draw the state transition diagrams and identify all communication classes. Is the Markov chain irreducible?

$$\mathbf{P}_1 = \begin{pmatrix} 0 & 1 & 0 \\ \frac{1}{2} & 0 & \frac{1}{2} \\ 1 & 0 & 0 \end{pmatrix} \quad \mathbf{P}_2 = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 0 & 1 \\ 0 & 1 & 0 \end{pmatrix} \quad \mathbf{P}_3 = \begin{pmatrix} \frac{1}{2} & \frac{1}{2} & 0 \\ 0 & 1 & 0 \\ \frac{1}{2} & 0 & \frac{1}{2} \end{pmatrix}$$

$$\mathbf{P}_4 = \begin{pmatrix} 0 & \frac{1}{2} & \frac{1}{2} & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 1 & 0 \\ 1 & 0 & 0 & 0 \end{pmatrix} \quad \mathbf{P}_5 = \begin{pmatrix} \frac{1}{2} & \frac{1}{2} & 0 & 0 \\ 1 & 0 & 0 & 0 \\ \frac{1}{2} & 0 & \frac{1}{4} & \frac{1}{4} \\ 0 & \frac{1}{4} & \frac{1}{4} & \frac{1}{2} \end{pmatrix} \quad \mathbf{P}_6 = \begin{pmatrix} \frac{1}{2} & 0 & 0 & 0 & \frac{1}{2} \\ 0 & \frac{1}{2} & 0 & \frac{1}{2} & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{2} & 0 & 0 & 0 & \frac{1}{2} \end{pmatrix}$$

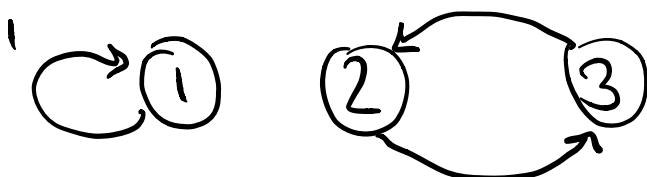
P<sub>1</sub>



Communication Classes  
 $\{1, 2, 3\}$

$\therefore$  irreducible

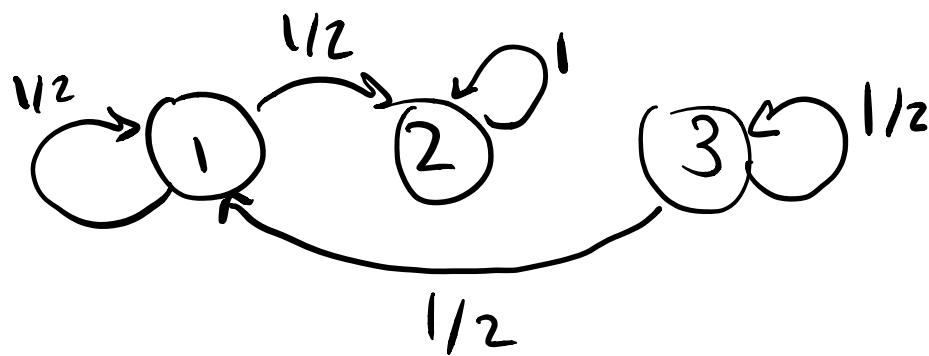
P<sub>2</sub>



$\{1\}, \{2, 3\}$

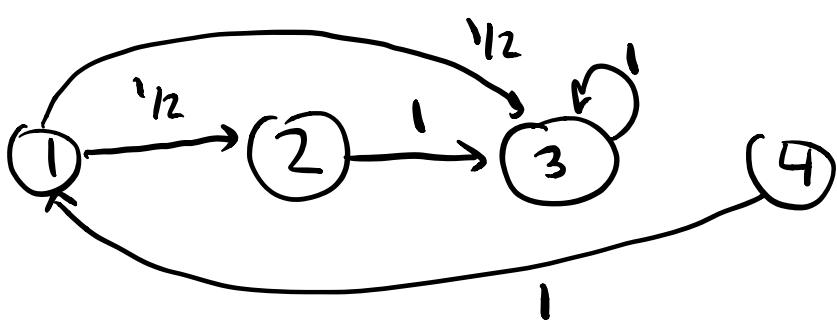
$\therefore$  not irreducible

P<sub>3</sub>



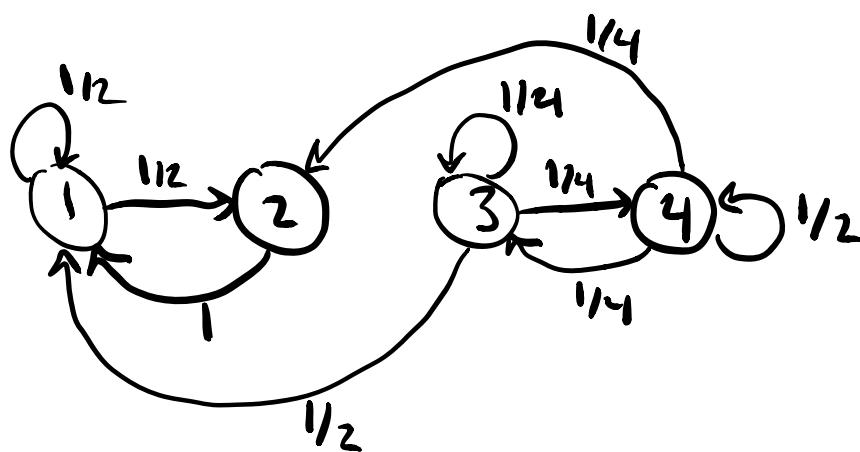
$\{1\}, \{2\}, \{3\}$   
 $\therefore$  not irreducible

P<sub>4</sub>



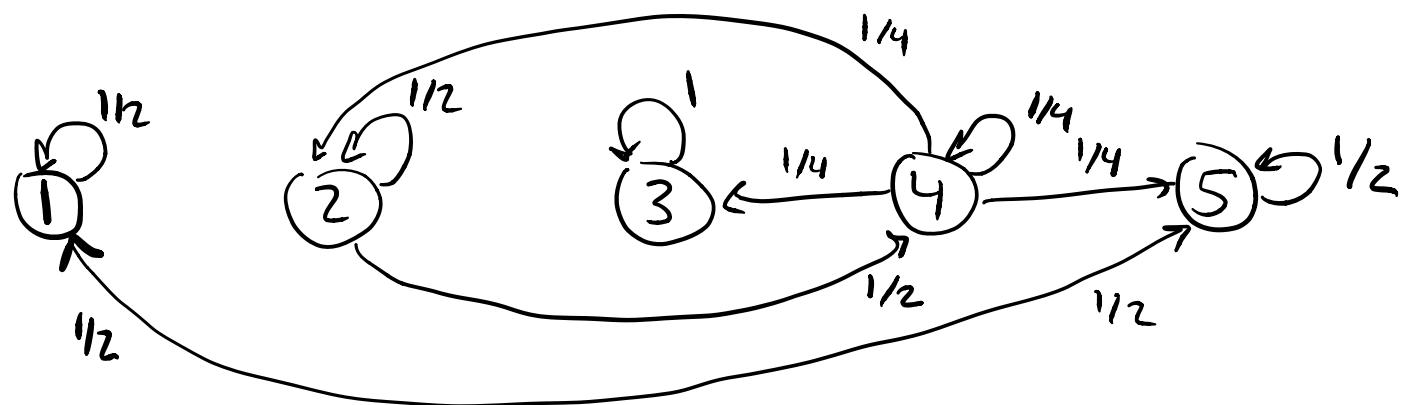
$\{1\}, \{2\}, \{3\}, \{4\}$   
 $\therefore$  not irreducible

P<sub>5</sub>



$\{1, 2\}$   
 $\{3, 4\}$   
 $\therefore$  not irreducible

P<sub>6</sub>



$$\{1\}, \{2, 4\}, \{3\}, \{5\}$$

∴ not reducible