

Recall

Multiple discrete rvs

- Joint pmfs $P_{X,Y}(x,y) = P(\{X=x\} \cap \{Y=y\})$ etc.
- Marginals in terms of joint
- ≥ 2 rvs

Started Conditional probability story for discrete rvs
 $P_{X|A}(x)$ for a discrete rv and event A

Total probability rule: If A_1, \dots, A_n partition Ω , then

$$P_X(x) = \sum_{k=1}^n P_{X|A_k}(x) P(A_k)$$

Given two rvs on same Ω, P - say X and Y - define

$$P_{X|Y}(x|y) = \frac{P(\{X=x\} \cap \{Y=y\})}{P(\{Y=y\})} = \frac{P_{X,Y}(x,y)}{P_Y(y)} ; \quad \forall x, y \\ w/ P_Y(y) > 0$$

For fixed y , $P_{X|Y}(x|y)$ defines a pmf "over x -values" -
 i.e.

$$P_{X|Y}(x|y) \geq 0 \quad \forall x \quad \text{and} \quad \sum_{x \in X} P_{X|Y}(x|y) = 1$$

" $P_{X|Y}(x|y)$ = conditional pmf of X given $Y=y$ "

Like the conditional "event-centered story", have a product rule of sorts

$$p_{X,Y}(x,y) = p_Y(y) p_{X|Y}(x|y) \quad \begin{matrix} x \in X \\ y \in Y \end{matrix}$$

OR

$$p_{X,Y}(x,y) = p_X(x) p_{Y|X}(y|x) \quad \begin{matrix} x \in X \\ y \in Y \end{matrix}$$

This expresses joint in terms of marginals + conditionals.

Also have a total-probability rule of sorts:

$$P_X(x) = \sum_{y \in Y} p_Y(y) p_{X|Y}(x|y)$$

$$p_Y(y) = \sum_{x \in X} P_X(x) p_{Y|X}(y|x)$$

Comment: These constructions generalize to ≥ 2 rvs. ^{i.e.} X, Y, Z defined on same \mathcal{S}, \mathbb{P}

Like

$$p_{X|Y,Z}(x|y,z) = \frac{\mathbb{P}(\{X=x\} \cap \{Y=y\} \cap \{Z=z\})}{\mathbb{P}(\{Y=y\} \cap \{Z=z\})} = \frac{p_{X,Y,Z}(x,y,z)}{p_{Y,Z}(y,z)}$$

$$p_{X,Y|Z}(x,y|z) = \frac{p_{X,Y,Z}(x,y,z)}{p_Z(z)}$$

⋮ etc

Get more involved product rules - $p_{X,Y,Z}(x,y,z) = p_Z(z) p_{Y|Z}(y|z) p_{X|Y,Z}(x|y,z)$
 ... etc

Example - Spinner + Hockey

Spinner has 4 equal wedges numbered 0, 1, 2, 3 (that's X).

- Take X shots on goal when spinner lands
- Shots you take are mutually independent
- Score with probability p each

Let $Y = \#(\text{goals you score})$

Here, $P_{Y|X}(y|x)$ and $p_X(x)$ is easy to figure out.

$$p_X(x) = \frac{1}{4} \text{ for } x = 0, 1, 2, 3$$

$$P_{Y|X}(y|0) = \begin{cases} 1, & y=0 \\ 0, & \text{otherwise} \end{cases}$$

$$P_{Y|X}(y|1) = \begin{cases} p, & y=1 \\ 1-p, & y=0 \end{cases}$$

make your one shot
miss your one shot

$$P_{Y|X}(y|2) = \begin{cases} (1-p)^2, & y=0 \\ p(1-p) + p(1-p), & y=1 \\ p^2, & y=2 \end{cases}$$

miss both
Two ways to score
1 goal!
make both

$$P_{Y|X}(y|3) = \begin{cases} \binom{3}{k} p^k (1-p)^{3-k}, & 0 \leq k \leq 3 \\ 0, & \text{else} \end{cases}$$

From these get joint pmf $P_{X,Y}(x,y) = P_X(x) P_{Y|X}(y|x)$.

Also can get marginal $P_Y(y)$ via

$$P_Y(y) = \sum_{x \in X} P_X(x) P_{Y|X}(y|x)$$

use this to answer
"what is IP(score 2 goals)?"

$$P_Y(y=2) = 0 + 0 + \frac{1}{4} p^2 + \frac{1}{4} \binom{3}{2} p^2 (1-p)$$

can't score 2
 goals in 0 or 1
 shots

$P_X(2)$ $P_{Y|X}(2|2)$ $P_X(3)$ $P_{Y|X}(2|3)$

Next, Conditional expectation (or conditional expected value)

Given Ω, \mathcal{P} ; discrete rv X ; event A :

$$\mathbb{E}(X|A) = \sum_{x \in X} x P_{X|A}(x)$$

just plug the pmf
 $P_{X|A}(x)$ into standard
 IE - formula

note how it's NOT
 $P_{X|A}(x|A)$. This is
 b/c A is an event & NOT
 a rv.

conditional expectation
 of X given event A

If A_1, A_2, \dots, A_n partition Ω , then we know that

$$P_X(x) = \sum_k P_{X|A_k}(x|A_k) P(A_k)$$

Hence,

$$\begin{aligned}\mathbb{E}(X) &= \sum_{x \in X} x p_X(x) = \sum_x \sum_k x p_{X|A_k}(x) P(A_k) \\ &= \sum_k \left(\sum_x x p_{X|A_k}(x) \right) P(A_k)\end{aligned}$$

Given Ω, \mathbb{P} , rvs X, Y :

$\mathbb{E}(X|Y=y) = \mathbb{E}(X|A)$; A is the event $\{Y=y\}$

So

"conditional expectation of X given $Y=y$ "

$$\mathbb{E}(X|A) = \sum_{x \in X} x p_{X|A}(x) = \sum_{x \in X} x p_{X|Y}(x|y)$$

Since all events $\{Y=y\}$ partition Ω , get

$$\mathbb{E}(X) = \sum_{y \in Y} \mathbb{E}(X|Y=y) P(\{Y=y\})$$

That's the second instance of

total expectation = avg of conditional expectations

Example - Spinner & Shots on Goal... continued

Easy to find

$$\mathbb{E}(Y|X=0) = 0$$

$$\mathbb{E}(Y|X=2) = 2p$$

$$\mathbb{E}(Y|X=1) = p$$

$$\mathbb{E}(Y|X=3) = 3p$$

Thus

expected number of goals scored

$$\mathbb{E}(Y) = \frac{1}{4}p + \frac{1}{4}2p + \frac{1}{4}3p = \frac{3}{2}p$$

Example - 2 envelopes

Have 2 envelopes. One has twice the amount of money as the other.

i.e. one has m dollars

one has $2m$ dollars

Call the envelopes X and Y .

let Z = total amount of money between the two

Say you find

$$X = m$$

Get to keep m if you want. Should you?

Figure out $\mathbb{E}(Y|X=m)$ using following **NAIVE** approach:

$$P(\{Y=2m\} | \{X=m\}) = P\left\{Y=\frac{m}{2}\right\} | \{X=m\} = \frac{1}{2}$$

This line of reasoning is incorrect

$$\mathbb{E}(Y|X=m) = 2m\left(\frac{1}{2}\right) + \frac{m}{2}\left(\frac{1}{2}\right) = \frac{5}{4}m > m$$

This leads you to believe you should switch.

The reasoning is specious because it is based on a woefully incomplete probability model.

How's it incomplete?

- Are X, Y integers? Continuous valued?
 - Eg if integers, if X is 19 you should definitely switch
- What's $P_Z(z)$? Any complete model would enable you to find this...

Let us make it more complete.

- Dealer picks Z "at random" - a multiple of 3 dollars
- Divides into two piles. $\frac{z}{3}$; $\frac{2z}{3}$.
- Flip a fair coin to decide how to allot piles to 2 envelopes, X, Y .

Say you open X and find m dollars.

$$Z = \begin{cases} 3m & , \text{ if } X \text{ is smaller envelope} \\ \frac{3}{2}m & , \text{ if } X \text{ is larger envelope} \end{cases}$$

If m is odd, can't have $Z = \frac{3}{2}m$ - but say m is even.

Fact: It's impossible to have

$$P_{Z|X}(3m|m) = P_{Z|X}\left(\frac{3}{2}m|m\right) = \frac{1}{2}$$