

Stationarity:  $F_{X(t_1+\Delta), X(t_2+\Delta), \dots, X(t_n+\Delta)}(x_1, \dots, x_n)$

$$= F_{X(t_1), \dots, X(t_n)}(x_1, \dots, x_n) \quad \forall n$$

$$\{t_i\}_{i=1}^n \quad \Delta$$

Wide-Sense Stationary:

$$E[X(t)] = \mu(t) \equiv \mu$$

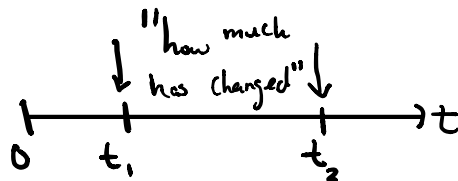
$$R_X(t_1, t_2) \triangleq E[X(t_1)X(t_2)]$$

$$= R_X(\underbrace{t_2 - t_1}_{\tau})$$

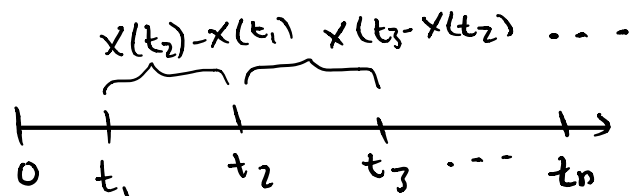
Note:  $R(0) \triangleq E[X^2(t_1)] \leftarrow$  will be deterministic

Increments:

$$X(t_2) - X(t_1)$$



Independent Increments



if all increments (which are rv)

are independent we say we have independent increments.

Common Trick: Have

write  $x(t_2)$  as  $\underbrace{g(x(t_2), x(t_1))}_{(t_1, t_2]} + x(t_1)$   $t_2 > t_1$   $\swarrow [0, t_1]$

to make analysis easier

# Counting Process

$$f(t), \quad t > 0$$
$$f(0) = 0$$

$f$  takes non-negative integers

$f$  is nondecreasing

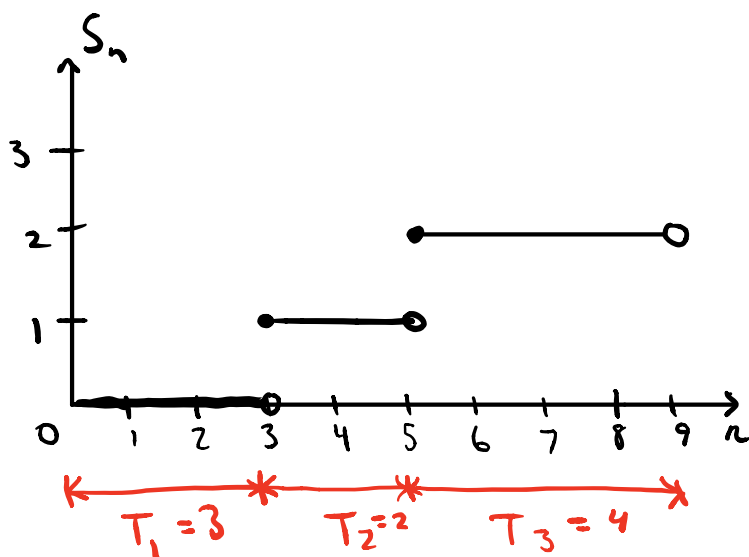
## Binomial Counting Process

$$X_i \sim \text{Bernoulli}(p)$$

$$S_n = \sum_{i=1}^n x_i$$

$$S_n \sim \text{Binomial}(n, p)$$

$$\Pr[S_n = k] = \binom{n}{k} p^k (1-p)^{n-k}$$



$T_1, T_2, T_3$  are a sequence of inter-arrival times

Interarrival Time  $\{T_i\}_{i=1}^{\infty}$

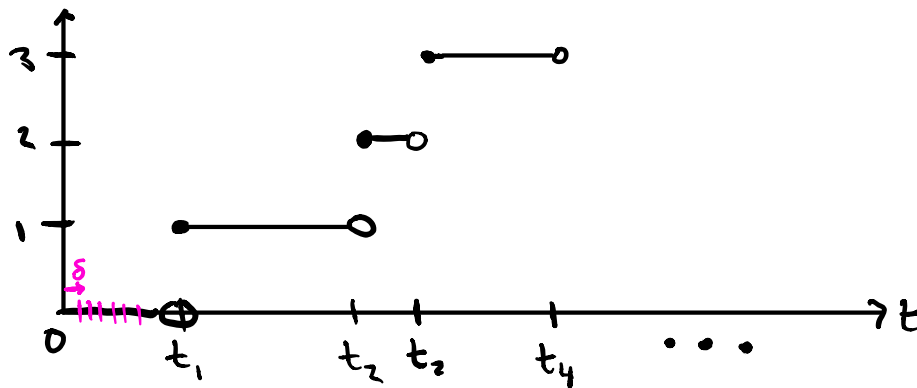
In above example,

$$T_i \sim \text{Geometric}(p)$$

$$\Pr[T_1 = k] = (1-p)^{k-1} p$$

What about continuous counting processes?

Say events arrive continuously w/ rate  $\lambda$  <sup>← avg arrival/unit time</sup>



Partition time into  $\delta$ -length intervals.

Assumptions.

For  $\delta \rightarrow 0$

- ① The probability of having more than one arrivals in  $\delta$ -interval is negligible (i.e.  $\rightarrow 0$ )
- ② Whether there is an arrival within a  $\delta$ -interval is **independent** of arrivals in other  $\delta$ -intervals

$$\boxed{p = \lambda \delta}$$

These assumptions make this similar to the binomial discrete case w/  $p = \lambda \delta$

$$\delta \rightarrow 0$$

$$p = \lambda \delta \rightarrow 0$$

$$n = \frac{t}{\delta} \rightarrow \infty$$

$$S_n \sim \text{Binomial}\left(\frac{t}{\delta}, \lambda \delta\right) \rightarrow \text{Poisson}(\lambda t)$$

## Poisson Process:

Definition:  $\{N(t)\}_{t \geq 0}$  is Poisson process w/ rate  $\lambda$  if it is a counting process w/ independent increments and  $N(t) - N(s) \sim \text{Pois}(\lambda(t-s)) \quad \forall t \geq s$ .

## Interarrival Time (note: $T_i$ iid $\forall i$ )

Look at  $T_1$

$$\Pr[T_1 > t] = (1 - \lambda \delta)^{\frac{t}{\delta}}$$

$$\lim_{\delta \rightarrow 0} (1 - \lambda \delta)^{t/\delta} = e^{-\lambda t}$$

$$\text{i.e. } T_1 \sim \text{exp}(\lambda)$$

$$\mathbb{E}[T_1] = \frac{1}{\lambda}$$

leads to ANOTHER definition

Definition:  $\{N(t)\}_{t \geq 0}$  is Poisson process w/ rate  $\lambda$  if it is a counting process w/ interarrival times iid exponentially distributed.

Also have a third definition (proved in HW/5)

Definition:  $\{N(t)\}_{t \geq 0}$  is a counting process w/ rate  $\lambda$  if it

is a counting process such that  $\forall \tau > 0$

any interval  $\tau$

$N(\tau) \sim \text{Pois}(\lambda\tau)$ , and given  $N(\tau) = n$  ( $n$  arrival

times in interval  $\tau$ ), are iid  $U[0, \tau]$

