

Recall

- Central Limit Theorem

- X_1, \dots, X_n iid; common mean μ , variance σ^2

- $M_n = \frac{X_1 + \dots + X_n}{n}$

- Set $Z_n = \frac{\sqrt{n}}{\sigma} (M_n - \mu)$

- Z_n has mean 0, variance 1 $\forall n > 0$

- As $n \rightarrow \infty$, $Z_n \rightarrow$ Gaussian mean 0, variance 1
in the sense that

$$F_{Z_n}(z) \rightarrow \Phi(z) \text{ if } z \text{ as } n \rightarrow \infty$$

Caution: people often use CLT "carelessly" to generate quantitative info

Let's see where care is needed.

Example - Be Careful

Say n large; given $c > 0$, what is $P(X_1 + \dots + X_n \geq c)$?

Provided n is large enough so Z_n above are "close to Gaussian", estimate this as follows:

Note

$$X_1 + \dots + X_n \geq c$$

is the same as

$$M_n \geq c/n$$

which is the same as

$$M_n - \mu > C_n - \mu$$

which is the same as

$$\frac{\sqrt{n}}{\sigma} (M_n - \mu) > \frac{\sqrt{n}}{\sigma} \left(\frac{C}{n} - \mu \right)$$

which is the same as

$$Z_n > \frac{\sqrt{n}}{\sigma} \left(\frac{C}{n} - \mu \right)$$

Thus

$$\begin{aligned} P(X_1 + \dots + X_n \geq c) &\approx P(Z_n \geq \frac{\sqrt{n}}{\sigma} \left(\frac{c}{n} - \mu \right)) \\ &\approx 1 - \Phi \left(\frac{\sqrt{n}}{\sigma} \left(\frac{c}{n} - \mu \right) \right) \end{aligned}$$

The "caution" thing here: n must be large enough so Gaussian approximation of Z_n holds. Unfortunately, no systematic way - even given complete pmf/pdf info about the X_k 's - to determine how large n has to be.

Example - Unknown p-coin

Have a coin; don't know $p = P(\{H\})$

Want to estimate p by repeatedly flipping

Let $X_k = 1$ if flip k is H

From

$$M_n = \frac{X_1 + \dots + X_n}{n} \quad \left. \begin{array}{l} \text{by WLLN, converges to} \\ p \text{ (in probability)} \end{array} \right\}$$

Question: How big does n have to be so

$$P(|M_n - p| \geq 0.01) \leq 0.05 \quad \begin{array}{l} \text{Chebyshev says} \\ \text{need } n \approx 50,000 \end{array}$$

How to use CLT to get a smaller n that will ensure this?

First, note that for large n , M_n distributed roughly symmetrically around p , so

$$P(|M_n - p| \geq 0.01) \approx 2P(M_n - p \geq 0.01)$$

Next, proceed as in previous example - turn this into a statement about Z_n .

$$M_n - p \geq 0.01$$

is the same as

$$\frac{\sqrt{n}}{\sigma} (M_n - p) \geq 0.01 \frac{\sqrt{n}}{\sigma}$$

Thus

$$2P(M_n - p \geq 0.01)$$

is equal to

$$2P(Z_n \geq 0.01 \frac{\sqrt{n}}{\sigma})$$

unlike last example,
or unknown!
 $\sigma = \sqrt{p(1-p)}$; don't know p .

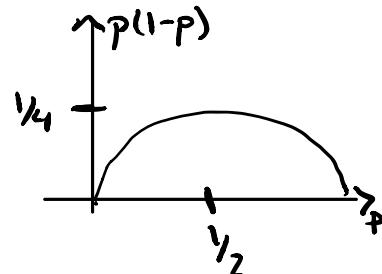
Since you don't know σ , use the fact $p(1-p) \geq \frac{1}{4}$
to bound σ by $\frac{1}{2}$

This yields a higher, more conservative, estimate of

$$P(Z_n \geq 0.01 \frac{\sqrt{n}}{\sigma})$$

Bottom line: Find how big n has to be so that

$$2P(Z_n \geq 0.02\sqrt{n}) \leq 0.05 \quad \xrightarrow{\text{from II-table}} n \approx 9604$$



Last thing: convergence w/ probability 1 of a sequence Y_1, Y_2, \dots of random variables.

Consider the sequence $\{Y_n : n > 0\}$.

Given some random variable Y ,

$$\left\{ \lim_{n \rightarrow \infty} Y_n = Y \right\}$$

is an event — need to refer back to Ω, \mathcal{P} , etc.

Say $Y_n \xrightarrow{\text{w.p.1}} Y$ with probability 1 (w.p.1)

$$Y_n \xrightarrow{\text{w.p.1}} Y$$

$$Y_n \xrightarrow{\text{a.s.}} Y \quad \text{a.s.} \rightarrow \text{almost surely}$$

when this event has probability 1.

Turns out: Convergence with probability 1 \Rightarrow convergence in probability

Example - Convergence With Probability 1

Say W_1, W_2, \dots are iid Uniform $[0, 1]$ random variables.

let

$$Y_n = \min\{W_1, \dots, W_n\}$$

Turns out

- Y_n is decreasing in n in the sense that $P(Y_{n+1} \leq Y_n) = 1 \forall n$
- Y_n is bounded below by 0.

Thus Y_n converges to a limit as $n \rightarrow \infty$ ← NOT obvious. See appendix

This is a random variable Y

- Can show that $P(Y=0) = 1$
 - This is because for any $\delta > 0$

$$P(Y \geq \delta) \leq P(X_n \geq \delta) = (1-\delta)^n \xrightarrow{n} 0$$

meaning $P(Y \geq \delta) = 0$

Because $P(Y=0) = 1$, we have

$$Y_n \xrightarrow{w.p.1} 0 \quad \text{as } n \rightarrow \infty$$

Strong Law of Large Numbers: When X_n iid, common mean μ , common variance σ^2 , and $M_n = \frac{X_1 + \dots + X_n}{n}$, we have

$$M_n \xrightarrow{w.p.1} \mu \quad \text{as } n \rightarrow \infty$$

Appendix (NOT done in lecture; added for my own clarity)

Definition: A sequence $\{s_n\}$ of real numbers is said to be

(i) monotonically increasing if $s_n \leq s_{n+1}$ ($n=1, 2, 3, \dots$)

(ii) monotonically decreasing if $s_n \geq s_{n+1}$ ($n=1, 2, 3, \dots$)

Theorem: Suppose $\{s_n\}$ is monotonic. Then $\{s_n\}$ converges if and only if it is bounded.

Proof

(\Leftarrow) Suppose $s_n \leq s_{n+1}$

let E be the range of $\{s_n\}$.

If $\{s_n\}$ is bounded, let $s = \sup E$.

Then

$$s_n \leq s \quad (n=1, 2, 3, \dots)$$

For every $\epsilon > 0$, $\exists N \in \mathbb{N}$ such that

$$s - \epsilon < s_N < s$$

for otherwise $s - \epsilon$ would be an upper bound of E .

Since $\{s_n\}$ increases,

$$n \geq N \Rightarrow s - \epsilon < s_n \leq s$$

Thus $\{s_n\}$ converges to s . [Analogous proof for decreasing]

(\Rightarrow) Suppose $p_n \rightarrow p$. $\exists N \in \mathbb{N}$ such that $n > N \Rightarrow d(p, p_n) < M$.

Put

$$r = \max(M, d(p_1, p), \dots, d(p_N, p))$$

then $d(p_n, p) \leq r$ for $n=1, 2, 3, \dots$

Q.E.D