## 1 Exercise 1

We have N samples  $(x_i, y_i)$   $\pi = (\pi_1, ..., \pi_M), \theta = (\theta_{1,1}, ...\theta_{M,K})$   $a_m = \mid \{i \mid z_i = m, \forall i \in [1, N]\} \mid, b_{m,k} = \mid \{i \mid z_i = m \text{ and } x_i = k, \forall i \in [1, N]\} \mid$ 

$$l(\boldsymbol{\pi}, \boldsymbol{\theta}) = \sum_{i=0}^{n} log(p(x_i, z_i))$$
$$= \sum_{i=0}^{n} log(p(x_i \mid z_i)p(z_i))$$
$$= \sum_{i=0}^{n} (log(\theta_{z_i, x_i}) + log(\pi_{x_i}))$$

 $l(\boldsymbol{\pi}, \boldsymbol{\theta})$  is concave. We want to minimize  $-l(\boldsymbol{\pi}, \boldsymbol{\theta})$  subjected to  $\sum_{k=1}^K \pi_k = 1$  and  $\sum_{k=1}^K \sum_{m=1}^M \theta_{m,k} = 1$  Lets introduce the langrangian.

$$L(\pmb{\pi}, \pmb{\theta}, \lambda_1, \lambda_2) = -(\sum_{i=0}^n (log(\theta_{z_i, x_i}) + log(\pi_{x_i}))) + \lambda_1(\sum_{k=1}^K \pi_k - 1) + \lambda_2(\sum_{k=1}^K \sum_{m=1}^M \theta_{m,k} - 1)$$

The Slaters constraint qualification are trivialy verified and therefore the problem has strong duality property. Therefore we have

$$\min_{\boldsymbol{\pi},\boldsymbol{\theta}} -l(\boldsymbol{\pi},\boldsymbol{\theta}) = \max_{\lambda_1,\lambda_2} L(\boldsymbol{\pi},\boldsymbol{\theta},\lambda_1,\lambda_2)$$

Moreover the lagrangian is convex with respect to  $\pi$  and  $\theta$ 

$$\frac{\partial L}{\partial \pi_m} = 0 \Rightarrow \tilde{\pi}_m = \frac{a_m}{\lambda_1}$$

$$\frac{\partial L}{\partial \theta_{m,k}} = 0 \Rightarrow \tilde{\theta}_{m,k} = \frac{b_{m,k}}{\lambda_2}$$

Using the constrains we can calculate  $\lambda_1, \lambda_2$ .

$$\tilde{\pi}_m = \frac{a_m}{N}$$

$$\tilde{\theta}_{m,k} = \frac{b_{m,k}}{N}$$

## 2 Exercicse 2

## Generative model LDA

We have N samples and  $n = |\{i, y_i = 0, \forall i \in [1, N]\}|$ 

$$\begin{split} l(\omega, \Sigma, \mu_0, \mu_1) &= \sum_{i=1}^{N} log(p(x_i, y_i)) \\ &= \sum_{i=1}^{N} log(p(x_i \mid y_i)p(y_i)) \\ &= \sum_{i=1, y_i=0}^{N} log(p(x_i \mid y_i=0)) + nlog(\omega) + \sum_{i=1, y_i=0}^{N} log(p(x_i \mid y_i=1)) + (N-n)log(1-\omega) \\ &= -\frac{Nd}{2} log(2\pi) + \frac{N}{2} log(|\Sigma^{-1}|) - \sum_{\substack{i=1, y_i=0}}^{N} \frac{1}{2} (x_i - \mu_0)^T \Sigma^{-1} (x_i - \mu_0) \\ &- \sum_{\substack{i=1, y_i=0}}^{N} \frac{1}{2} (x_i - \mu_1)^T \Sigma^{-1} (x_i - \mu_1) + nlog(\omega) + (N-n)log(1-\omega) \end{split}$$

This log likelyhood is not concave in  $(\omega, \Sigma, \mu_0, \mu_1)$ . It is concave in  $(\omega, \mu_0, \mu_1)$ with  $\Sigma$  fixed.

$$\nabla_{\omega} l = \frac{n}{\omega} - \frac{N - n}{1 - \omega}$$

 $\nabla_{\omega} l = 0$  gives us:

$$\tilde{\omega} = \frac{n}{N}$$

Calculating the gradient in  $\mu_0$ ,  $\mu_1$  and equalating it to 0 gives us.

$$\tilde{\mu}_0 = \frac{1}{n} \sum_{\substack{i=1, \\ y_i = 0}}^{N} x_i$$

$$\tilde{\mu}_1 = \frac{1}{N - n} \sum_{\substack{i=1, \\ y_i = 1}}^{N} x_i$$

Let us now differentiate 
$$l$$
 w.r.t.  $\Sigma^{-1}$ .  
Let  $A = \Sigma^{-1}$ ,  $\Sigma_0 = \frac{1}{n} \sum_{\substack{i=1 \ y_i=0}}^{N} (x_i - \mu_0)^T (x_i - \mu_0)$ ,  
 $\Sigma_1 = \frac{1}{N-n} \sum_{\substack{i=1 \ y_i=1}}^{N} (x_i - \mu_1)^T (x_i - \mu_1)$ 

$$\Sigma_1 = \frac{1}{N-n} \sum_{\substack{i=1, \ y_i=1}}^{N} (x_i - \mu_1)^T (x_i - \mu_1)$$

We have:

$$\begin{split} l(\omega, \Sigma, \mu_0, \mu_1) &= -\frac{Nd}{2}log(2\pi) + \frac{N}{2}log(\mid \Sigma^{-1}\mid) - \frac{1}{2}\mathrm{Trace}(A(n\Sigma_0 + (N-n)\Sigma_1) \\ &+ nlog(\omega) + (N-n)log(1-\omega) \\ \nabla_A l &= \frac{N}{2}A^{-1} - \frac{1}{2}(n\Sigma_0 + (N-n)\Sigma_1) \end{split}$$

Which leads to

$$\tilde{\Sigma} = \frac{n}{N} \Sigma_0 + \frac{N-n}{N} \Sigma_1$$

We have found a unique stationnary point for the likelyhood. To be sure it is a maximum we would have to calculate the Hessian. Now we will calculate the  $p(y = 1 \mid x)$ .

$$p(y = 1 \mid x) = \frac{p(x \mid y = 1)p(y = 1)}{p(x)}$$

$$\begin{split} \log(\frac{p(y=1\mid x)}{p(y=0\mid x)}) &= \log(\frac{1-\omega}{\omega}) - \frac{1}{2}(x-\mu_1)\Sigma^{-1}(x-\mu_1) + \frac{1}{2}(x-\mu_0)\Sigma^{-1}(x-\mu_0) \\ \log(\frac{p(y=1\mid x)}{p(y=0\mid x)}) &= \log(\frac{1-\omega}{\omega}) + \frac{1}{2}(\mu_0^T\Sigma^{-1}\mu_0 - \mu_1^T\Sigma^{-1}\mu_1) + x^T\Sigma^{-1}(\mu_1 - \mu_0) \\ p(y=1\mid x) &= \frac{1}{1 + \frac{w}{1-w}\exp\left(\frac{1}{2}(\mu_1^T\Sigma^{-1}\mu_1 - \mu_0^T\Sigma^{-1}\mu_0)\right)\exp\left(-x^T\Sigma^{-1}(\mu_1 - \mu_0)\right)} \end{split}$$

It is of the form

$$p(y = 1 \mid x) = \frac{1}{1 + \exp(-(x^T a + \alpha))}$$

It is the formula of logistic regression

## 2.2 QDA model

We have N samples and  $n = |\{i, y_i = 0, \forall i \in [1, N]\}|$ 

$$\begin{split} l(\omega, \Sigma_0, \Sigma_1, \mu_0, \mu_1) &= \sum_{i=1}^N log(p(x_i, y_i)) \\ &= \sum_{i=1}^N log(p(x_i \mid y_i) p(y_i)) \\ &= \sum_{i=1}^N log(p(x_i \mid y_i = 0)) + nlog(\omega) + \sum_{i=1}^N log(p(x_i \mid y_i = 1)) + (N-n)log(1-\omega) \\ &= nlog(\omega) - \frac{Nd}{2} log(2\pi) + \frac{n}{2} log(\mid \Sigma_0^{-1} \mid) - \sum_{\substack{i=1, \ y_i = 0}}^N \frac{1}{2} (x_i - \mu_0)^T \Sigma_0^{-1} (x_i - \mu_0) \\ &+ (N-n)log(1-\omega) + \frac{N-n}{2} log(\mid \Sigma_1^{-1} \mid) - \sum_{\substack{i=1, \ y_i = 1}}^N \frac{1}{2} (x_i - \mu_1)^T \Sigma_1^{-1} (x_i - \mu_1) \end{split}$$

This log likelyhood is not concave in  $(\omega, \Sigma_0, \Sigma_1, \mu_0, \mu_1)$ . It is concave in  $(\omega, \mu_0, \mu_1)$  with  $\Sigma_0$  and  $\Sigma_1$  fixed. We obtain like in the previous questions.

$$\tilde{\omega} = \frac{n}{N}$$

$$\tilde{\mu}_0 = \frac{1}{n} \sum_{\substack{i=1, \\ y_i = 0}}^{N} x_i$$

$$\tilde{\mu}_1 = \frac{1}{N-n} \sum_{\substack{i=1, \\ y_i = 1}}^{N} x_i$$

Differentiating l w.r.t.  $\Sigma_0^{-1}$  with the rest fixed and equalizing to 0 (and then doing the same with  $\Sigma_1^{-1}$ ) gives us.

$$\tilde{\Sigma}_0 = \frac{1}{n} \sum_{\substack{i=1, \\ y_i = 0}}^{N} (x_i - \mu_0)^T (x_i - \mu_0)$$

$$\tilde{\Sigma}_1 = \frac{1}{N - n} \sum_{i=1, }^{N} (x_i - \mu_1)^T (x_i - \mu_1)$$

We did not provide the calculations because it is amlost the same as above.