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Least square

Estimation

Prediction

Regression Analysis Simple Linear Regression

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Simple linear regression

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- Consider a data set $D = \{(x_i, y_i) | x_i \in \mathbb{R}, y_i \in \mathbb{R}, \forall i = 1, 2, \dots, n\}$
- x_i s are non stochastic
- y_i s are stochastic and realized values of random variable Y_i s
- ϵ_i s are iid $N(0, \sigma^2)$ random variables
- Regression parameter (β_0, β_1) and error parameter σ^2 are unknown

Problem statement

Considering the linear model

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i \quad \forall \quad i = 1, 2, \dots, n$$

we want to estimate (β_0, β_1) which will minimize least squared condition.



Least square estimate

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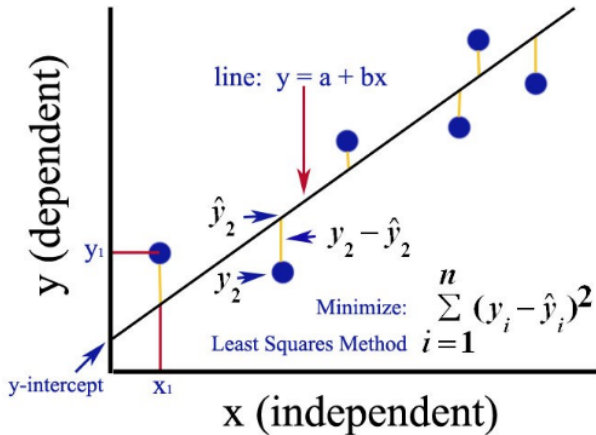


Figure: Least square



Parameter estimation

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Least Squared condition

The least squared condition to estimate the model parameters is to minimize

$$S(\beta_0, \beta_1) = \sum_i (y_i - \beta_0 - \beta_1 x_i)^2. \quad (1)$$

with respect to β_0 and β_1 .

If $(\hat{\beta}_0, \hat{\beta}_1)$ minimizes $S(\beta_0, \beta_1)$ then their values can be obtained by solving the **normal equations**

$$\frac{\partial S(\beta_0, \beta_1)}{\partial \beta_0} = 0 \implies n\hat{\beta}_0 + \hat{\beta}_1 \sum_i x_i = \sum_i y_i \quad (2)$$

$$\frac{\partial S(\beta_0, \beta_1)}{\partial \beta_1} = 0 \implies \hat{\beta}_0 \sum_i x_i + \hat{\beta}_1 \sum_i x_i^2 = \sum_i y_i x_i \quad (3)$$



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Estimated parameters

Defining $S_{xy} = \sum_i (y_i - \bar{y})(x_i - \bar{x})$ we have the solutions as

$$\hat{\beta}_1 = \frac{S_{xy}}{S_{xx}} \text{ and } \hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$$

Least squared prediction line

For any x such as old x_i s or some x_{new} the prediction line is

$$\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x$$



Prediction Error

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The **prediction error or residual** is defined as $e_i = y_i - \hat{y}_i$ and hence the residual sum of square (SSR) is

$$\begin{aligned}\sum_i e_i^2 &= \sum_i (y_i - \hat{y}_i)^2 \\ &= \sum_i (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2 \\ &= \sum_i [(y_i - (\bar{y} - \hat{\beta}_1 \bar{x} - \hat{\beta}_1 x_i))]^2 \\ &= \sum_i [(y_i - \bar{y}) - \hat{\beta}_1 (x_i - \bar{x})]^2 \\ &= S_{yy} + \hat{\beta}_1^2 S_{xx} - 2\hat{\beta}_1 S_{xy} \\ &= S_{yy} - \hat{\beta}_1 S_{xy} \\ &= S_{yy} - \frac{S_{xy}^2}{S_{xx}}\end{aligned}\tag{4}$$

NOTE: We estimate σ^2 by $\hat{\sigma}^2 = \frac{SSR}{n-2} \equiv MSR$, mean residual sum of square. [Explanation will be studied under Multiple Linear Regression]



Prediction Error

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- We estimate σ^2 by $\hat{\sigma}^2 = \frac{SSR}{n-2} \equiv MSR$, mean residual sum of square. [Explanation will be studied under Multiple Linear Regression]
- We can estimate σ^2 even without computing individual errors e_i .
- Prediction error $\mathbf{e} = (e_1, e_2, \dots, e_n)$ and prediction vector $\hat{\mathbf{y}} = (\hat{y}_1, \hat{y}_2, \dots, \hat{y}_n)$ are independently distributed.

Linear estimator

If an estimator $T(\mathbf{y})$ can be expressed as a linear combination of \mathbf{y} with non random coefficients i.e. $T(\mathbf{y}) = \sum_i \alpha_i y_i$ then $T(\mathbf{y})$ is called a linear estimator.



Note

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1 $\hat{\beta}_1$ is a linear estimator of β_1 . and $\hat{\beta}_1 \sim N\left(\beta_1, \frac{\sigma^2}{S_{xx}}\right)$

2 $\hat{\beta}_0$ is a linear estimator of β_0 and $\hat{\beta}_0 \sim N\left(\beta_0, \sigma^2\left(\frac{1}{n} + \frac{\bar{x}^2}{S_{xx}}\right)\right)$

3 Regression line passes through (\bar{x}, \bar{y})

4 $SSR \sim \sigma^2 \chi_{n-2}^2$



What more?

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- Can we test for $H_0 : \beta_0 = b_0$ vs $H_1 : \beta_0 \neq b_0$?
- Can we test for $H_0 : \beta_1 = b_1$ vs $H_1 : \beta_1 \neq b_1$?
- Can we test for $H_0 : \sigma^2 = \sigma_0^2$ vs $H_1 : \sigma^2 \neq \sigma_0^2$?
- Can we have a prediction interval for \hat{y} ?

YES !!!! Now we can...



When regressor is stochastic

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Bivariate normal($\mu_x, \mu_y, \sigma_x^2, \sigma_y^2, \rho$)

If (X, Y) has joint density function $f(x, y)$ as

$$f(x, y) = \frac{e^{-\frac{1}{2(1-\rho^2)} \left[\left(\frac{x-\mu_x}{\sigma_x} \right)^2 + \left(\frac{y-\mu_y}{\sigma_y} \right)^2 - 2\rho \left(\frac{x-\mu_x}{\sigma_x} \right) \left(\frac{y-\mu_y}{\sigma_y} \right) \right]}}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}}$$

then (X, Y) is said to follow Bivariate normal.

- If (X, Y) follow Bivariate normal($\mu_x, \mu_y, \sigma_x^2, \sigma_y^2, \rho$) then show that $Y|X = x$ follows $N(\mu_y + \rho \frac{\sigma_y}{\sigma_x}(x - \mu_x), (1 - \rho^2)\sigma_y^2)$
- Perform a large and small sample test for $H_0 : \rho = 0$ Vs $H_1 : \rho \neq 0$.
- Perform a large sample test for $H_0 : \rho = \rho_0$ Vs $H_1 : \rho \neq \rho_0$.



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