

FRE-GY 6233: Assignment 2

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Problem 1:

- (i) Suppose that $(\Omega, \mathcal{F}, \mathbb{P})$ is a probability space where $\Omega = \{a, b, c, d, e, f\}$, \mathcal{F} is σ -algebra, and \mathbb{P} is uniform (so $\mathbb{P}(a) = \mathbb{P}(b) = \dots = \frac{1}{6}$).
- (ii) Let X, Y, Z be r.v. given by

$$\begin{aligned} X(a) = 1, X(b) = X(c) = 3, X(d) = X(e) = 5, X(f) = 7 \\ Y(a) = Y(b) = 2, Y(c) = Y(d) = 1, Y(e) = Y(f) = 7 \\ Z(a) = Z(b) = Z(c) = Z(d) = 3, Z(e) = Z(f) = 2 \end{aligned} \quad (1)$$

Solve the follow questions:

1. Write down $\sigma(X), \sigma(Y), \sigma(Z)$. Are there any relationships between them?

Solution:

$$\begin{aligned} \sigma(X) &= \{\emptyset, \Omega, \{a\}, \{b, c\}, \{d, e\}, \{f\}, \{a, b, c\}, \{a, d, e\}, \{a, f\}, \\ &\quad \{b, c, d, e\}, \{b, c, f\}, \{d, e, f\}, \{a, b, c, d, e\}, \{b, c, d, e, f\}, \{a, d, e, f\}, \{a, b, c, f\}\} \\ \sigma(Y) &= \{\emptyset, \Omega, \{a, b\}, \{c, d\}, \{e, f\}, \{a, b, c, d\}, \{a, b, e, f\}, \{c, d, e, f\}\} \\ \sigma(Z) &= \{\emptyset, \Omega, \{a, b, c, d\}, \{e, f\}\} \end{aligned}$$

We note that $\sigma(Z) \subset \sigma(Y)$.

2. Define a r.v. $\mathbb{E}[X \mid \sigma(Y)]$

Solution: As $\mathbb{E}[X|Y] = \sum_{i \in \{1,2,7\}} \mathbb{E}[X|Y=i] \mathbb{1}_i(y)$, we consider the following:

If $Y = 1$, $\mathbb{P}(Y^{-1}(1) = c) = \frac{\mathbb{P}(c)}{\mathbb{P}(c)+\mathbb{P}(d)} = \frac{1}{2}$, $\mathbb{P}(Y^{-1}(1) = d) = \frac{1}{2}$

So, $\mathbb{E}[X|Y=1] = (X(c) \cdot \frac{1}{2} + X(d) \cdot \frac{1}{2}) = 3 \cdot \frac{1}{2} + 5 \cdot \frac{1}{2} = 4$

If $Y = 2$, $\mathbb{P}(Y^{-1}(2) = a) = \frac{\mathbb{P}(a)}{\mathbb{P}(a)+\mathbb{P}(b)} = \frac{1}{2}$, $\mathbb{P}(Y^{-1}(2) = b) = \frac{1}{2}$

So, $\mathbb{E}[X|Y = 2] = (X(a) \cdot \frac{1}{2} + X(b) \cdot \frac{1}{2}) = 2$
 If $Y = 7$, $\mathbb{P}(Y^{-1}(7) = e) = \frac{\mathbb{P}(e)}{\mathbb{P}(e) + \mathbb{P}(f)} = \frac{1}{2}$, $\mathbb{P}(Y^{-1}(7) = f) = \frac{1}{2}$
 So, $\mathbb{E}[X|Y = 7] = (X(e) \cdot \frac{1}{2} + X(f) \cdot \frac{1}{2}) = 6$
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3. Check directly the averaging property

$$\mathbb{E}[\mathbb{E}[X | \sigma(Y)]] = \mathbb{E}[X]$$

Solution: $\mathbb{E}[\mathbb{E}[X | \sigma(Y)]] = \sum_{i \in \{1,2,7\}} \mathbb{E}[X|Y = i] \cdot \mathbb{P}[Y = i]$. As $\mathbb{P}[Y = i] = \frac{1}{3}$ for $i \in \{1, 2, 7\}$, we have that $\mathbb{E}[\mathbb{E}[X | \sigma(Y)]] = \frac{1}{3} \cdot (4 + 2 + 6) = 4$.
 We then note that:
 $\mathbb{E}[X] = \sum_{i \in \{a,b,c,d,e,f\}} \mathbb{P}[i] \cdot X(i) = \frac{1}{6}(1 + 3 + 3 + 5 + 5 + 7) = \frac{24}{6} = 4$.
 From this we have checked the averaging property.

4. Show directly (by calculating) that

$$\mathbb{E}[\mathbb{E}[X | \sigma(Y)] | \sigma(Z)] = \mathbb{E}[X | \sigma(Z)]$$

Solution: We first show the RHS:
 Let us denote events $\omega_{z1} = \{a, b, c, d\}$, $\omega_{z2} = \{e, f\}$ to partition Ω to form $\sigma(Z)$.
 We then denote events $\omega_{x1} = \{a\}$, $\omega_{x2} = \{b, c\}$, $\omega_{x3} = \{d, e\}$, $\omega_{x4} = \{f\}$ to partition Ω to form $\sigma(X)$.
 We then note that $\omega_{z1} \cap \omega_{x1} = \{a\}$, $\omega_{z1} \cap \omega_{x2} = \{b, c\}$, $\omega_{z1} \cap \omega_{x3} = \{d\}$. We then have
 $\mathbb{E}[X | \omega_{z1}] = X(\omega \in \{a\}) \frac{\mathbb{P}(\{a\})}{\mathbb{P}(\omega_{z1})} + X(\omega \in \{b, c\}) \frac{\mathbb{P}(\{b, c\})}{\mathbb{P}(\omega_{z1})} + X(\omega \in \{d\}) \frac{\mathbb{P}(\{d\})}{\mathbb{P}(\omega_{z1})} =$
 $1 \cdot \frac{1}{4} + 3 \cdot \frac{2}{4} + 5 \cdot \frac{1}{4} = 3$
 We also note that $\omega_{z2} \cap \omega_{x3} = \{e\}$, $\omega_{z2} \cap \omega_{x4} = \{f\}$ so that we have
 $\mathbb{E}[X | \omega_{z2}] = X(\omega \in \{e\}) \frac{\mathbb{P}(\{e\})}{\mathbb{P}(\omega_{z2})} + X(\omega \in \{f\}) \frac{\mathbb{P}(\{f\})}{\mathbb{P}(\omega_{z2})} = 5 \cdot \frac{1}{2} + 7 \cdot \frac{1}{2} = 6$
 We write this as:

$$\mathbb{E}[X | \sigma(Z)] = \begin{cases} 3, & \text{for } \omega \in \omega_{z1} = \{a, b, c, d\} \\ 6, & \text{for } \omega \in \omega_{z2} = \{e, f\} \end{cases}$$

We then look at the LHS:

In part (ii), we had subtly defined $X | \sigma(Y)$ over events $\omega_{y1} = \{a, b\}$, $\omega_{y2} = \{c, d\}$, $\omega_{y3} = \{e, f\}$ that partition Ω to form $\sigma(Y)$. From that we recieved the

following random variable:

$$\mathbb{E}[X \mid \sigma(Y)] = \begin{cases} 2, & \text{for } \omega \in \omega_{y1} = \{a, b\} \\ 4, & \text{for } \omega \in \omega_{y2} = \{c, d\} \\ 6, & \text{for } \omega \in \omega_{y3} = \{e, f\} \end{cases}$$

We also noted that $\sigma(Z) \subset \sigma(Y)$. It follows that as $\omega_{y1} \cup \omega_{y2} = \omega_{z1}$ and $\omega_{y3} = \omega_{z3}$,

$$\begin{aligned} & \mathbb{E}[X \mid \sigma(Y)] \mid \omega_{z1} \\ &= (\mathbb{E}[X \mid \omega_{y1}] \mid \omega_{z1}) \cdot \frac{\mathbb{P}(\omega_{y1} \cap \omega_{z1})}{\omega_{z1}} + (\mathbb{E}[X \mid \omega_{y2}] \mid \omega_{z1}) \cdot \frac{\mathbb{P}(\omega_{y2} \cap \omega_{z1})}{\omega_{z1}} \\ &= 2 \cdot \frac{1}{2} + 4 \cdot \frac{1}{2} = 3 \end{aligned}$$

$$\begin{aligned} & \mathbb{E}[X \mid \sigma(Y)] \mid \omega_{z2} \\ &= (\mathbb{E}[X \mid \omega_{y3}] \mid \omega_{z2}) \cdot \frac{\mathbb{P}(\omega_{y3} \cap \omega_{z2})}{\omega_{z2}} = \mathbb{E}[X \mid \omega_{z3}] = 6 \end{aligned}$$

We then have:

$$\mathbb{E}[X \mid \sigma(Y)] \mid \sigma(Z) = \begin{cases} 3, & \text{for } \omega \in \omega_{z1} = \{a, b, c, d\} \\ 6, & \text{for } \omega \in \omega_{z2} = \{e, f\} \end{cases}$$

So it is evident that LHS = RHS.

5. Check if X and Y , or Y and Z are independent under given probability.

Solution: It is straightforward to see that X and Y are not independent; otherwise, $\mathbb{E}[X \mid \sigma(Y)] = \mathbb{E}[X]$. We have shown that $\mathbb{E}[X] = 4$, which is clearly not what we have shown to be $\mathbb{E}[X \mid \sigma(Y)]$ in the problems above.

To show that Y and Z are not independent, we proceed by using the definition of independence. We check that for some $\omega_1 \in \sigma(Y), \omega_2 \in \sigma(Z)$, that $\mathbb{P}(\omega_1 \cap \omega_2) \neq \mathbb{P}(\omega_1) \cdot \mathbb{P}(\omega_2)$. It is most evident that as $\omega_{z2} = \omega_{y3}$ we would have $\mathbb{P}(\omega_{z2} \cap \omega_{y3}) = \mathbb{P}(\omega_{z2}) = \mathbb{P}(\omega_{z2})\mathbb{P}(\omega_{y3}) = \mathbb{P}(\omega_{z2})^2$. As $\mathbb{P}(\omega_{z2}) = \frac{1}{3}$, this cannot be true. As the two sigma-algebra are not independent, random variables Y and Z are not independent.

Problem 2: Prove Markov and Tchebyshev inequalities.

Solution: Markov Inequality: For $\lambda, p > 0$ and nonnegative random variable X , we have:

$$\mathbb{P}(\omega : |X(\omega)| \geq \lambda) \leq \frac{1}{\lambda^p} \mathbb{E}[|X|^p]$$

Proof. We first by showing the above inequality for $p = 1$. This follows directly from the definition of expectation:

If we fix $\lambda > 0$ and define set $A = \{\omega : X(\omega) \geq \lambda\}$

$$\begin{aligned} \mathbb{E}[X] &= \int_{\mathbb{R}} X(\omega) dP(\omega) = \int_A X(\omega) dP(\omega) + \int_{\mathbb{R} \setminus A} X(\omega) dP(\omega) \\ &\geq \int_A X(\omega) dP(\omega) \geq \int_A \lambda dP(\omega) \text{ [this follows from the condition that } X(\omega) \geq \lambda \text{ in set } A \text{]} \\ &= \lambda \int_A dP(\omega) = \lambda \mathbb{P}(\omega : |X(\omega)| \geq \lambda). \end{aligned}$$

\Rightarrow

$$\mathbb{P}(\omega : |X(\omega)| \geq \lambda) \leq \frac{1}{\lambda^p} \mathbb{E}[|X|^p]$$

To extend it to all other $p > 0$, we consider function $\phi(x) = |x|^p$. If ϕ is positive and non-decreasing, we have that $\mathbb{P}[X \geq \lambda] \leq \mathbb{P}[\phi(X) \geq \phi(\lambda)] \leq \mathbb{E}[\phi(X)]/\phi(\lambda)$. As ϕ is positive and non-decreasing, we are done. \square

Tchebychev Inequality: If X is a r.v. with mean μ and variance σ^2 , then

$$\mathbb{P}(\omega : |X(\omega) - \mu| \geq \lambda) \leq \frac{\sigma^2}{\lambda^2}$$

Proof. The above inequality follows from Markov's inequality on $(X - \mu)^2$ with $p = 2$. We observe that $\mathbb{P}(\omega : |X(\omega) - \mu| \geq \lambda) \leq \frac{\mathbb{E}[(X - \mu)^2]}{\lambda^2} = \frac{\mathbb{E}[X^2 - 2\mu \cdot X + \mu^2]}{\lambda^2}$
 $= \frac{1}{\lambda^2} \cdot (\mathbb{E}[X^2] - 2\mu^2 + \mu^2) = \frac{1}{\lambda^2} \cdot (\mathbb{E}[X^2] - \mathbb{E}[X]^2) = \frac{\sigma^2}{\lambda^2}$ \square

Problem 3: Let X be a r.v. and $\lambda > 0$. Prove that the following bound holds:

$$\mathbb{P}(X \geq \lambda) \leq \frac{\mathbb{E}[e^{tX}]}{e^{\lambda t}}, \forall t > 0$$

Use Markov inequality.

Solution: We note that the function $\phi(x) = e^{tx}$ for $t > 0$ is a positive and non-decreasing function. We first have by Markov's inequality that $\mathbb{P}(X \geq \lambda) = \mathbb{P}(\phi(x) \geq \phi(\lambda))$
 $\leq \frac{1}{\phi(\lambda)} \mathbb{E}[\phi(X)] \Rightarrow \mathbb{P}(X \geq \lambda) \leq \frac{\mathbb{E}[e^{tX}]}{e^{\lambda t}}, \forall t > 0.$