MATH 323 - Tutorial 7 (Feb. 22nd /17)



Dof: The Kth moment of a random variable Y taken about the origin is defined to be $E(Y^k) = \mu_k$.

The moment-generating function $m(t) := \mathbb{E}(e^{tY})$ for random variable Y exists if there exists a possitive constant b such that m(t) is finite for $|t| \le b$.

Theorem: If m(t) exists, then for any positive integer k, $\frac{d^{k}}{dt^{k}} = m^{(k)}(0) = E(y^{(k)}) = u^{(k)}$ $\frac{d^{k}}{dt^{k}} = m^{(k)}(0) = E(y^{(k)}) = u^{(k)}$

I. Derive the moment generating function for the Binomial distribution and use it to calculate the mean and variance.

Solution:



Let $Y \sim B$ inomial (n, p), then $P(Y=y) = (y) p^y (1-p)^{-y}$ where $y \in \{0,1\},...,n\}$ and $0 \leq p \leq 1$.

$$m(t) = E(e^{tY}) = \sum_{y=0}^{n} (y) p^{3} (1-p)^{n-3} e^{ty}$$

$$= \sum_{y=0}^{n} (y) (e^{t}p)^{3} (1-p)^{n-3}$$

which by the binomial expansion

$$= (e^{t}p + (1-p))^{n}$$

$$= (1-p+pe^{t})^{n}$$

d m(t) = n (1-p+pet) n-1 pet

 $\frac{d^{2} m(t) = np \left\{ e^{t} (1-p+pe^{t})^{n-1} + e^{t} (n-1)(1-p+pe^{t})^{n-2} pe^{t} \right\}$

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\frac{d}{dt} m(t) = n(1-p+pe^{0})^{n-1}pe^{0}
= n(1-p+p)^{n-1}p
= np = E(Y)
   \frac{d^{2} m(t)}{dt^{2}} = np \left\{ e^{\circ} (1-p+pe^{\circ})^{n-1} + e^{\circ} (n-1) (1-p+pe^{\circ})^{n-2} pe^{\circ} \right\}
                 = np { 1 + (n-1)p} = np (1+np-p)
                          = np + n^2p^2 - np^2 = E(Y^2)
    V(Y) = \mathbb{E}(Y^2) - (\mathbb{E}(Y))^2
         = np + n^2p^2 - np^2 - (np)^2
= np + n^2p^2 - np^2 - n^2p^2 = np - np^2 = np(1-p)
2. (Exercise 3.156,8): Suppose that Y is a random variable with
   a) what is m(0)?
   b) If W=3Y, show the moment-generating function of W is m(3t).
c) If X=Y-2, show the moment-generating function of X is e^{-2t}m(t).
d) If Z=aY+b, show the moment-generating function of Z is e^{tb}m(at).
   Solution:

(t) = E(ety) - 840) 9 =
      => m(o) = E(eoy) = E(I) = I
      2/m= [P(14-0.50/ LO.03) X& rolling
     b) Let m_w(t) be the mgf of W.

m_w(t) = E(e^{tW}) = E(e^{t(3Y)}) = E(e^{(3t)Y})
     c) Let mx (t) be the mgf of X.

mx (t) = E(e<sup>tX</sup>) = E(e<sup>t(Y-2)</sup>) = E(e<sup>tY-2t</sup>)
                                                     = E(e-2t et9)
                                                    = e-2+ E(e+4)
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d) Let
$$W_Z(t)$$
 be the map of Z .

 $W_Z(t) = E(e^{tZ}) = JE(e^{t(49+b)})$
 $= E(e^{(45)Y})$
 $= e^{tb}E(e^{(45)Y})$
 $= e^{tb}M(at)$

Then for any constant $k > 0$,

 $P(14-\mu | 2k\sigma) \ge 1 - \frac{1}{k^2}$ or $P(14-\mu | 2k\sigma) \le \frac{1}{k^2}$.

3. (Exercise 3.170) The U.S. mint produces dimes with an average diameter of 0.5 inch and standard deviation 0.01. Using The large shifts theorem, find a lower bound for the number of coins in a lot of 400 coins that are expected to have a diameter between 0.4% and 0.52.

Solution:

Let Y be the diameter of a coin. Then

 $P(0.48 \ge 4 \le 0.50 \ge 0.52 = 0.50$
 $= P(-0.02 \le 4 \le 0.50 \ge 0.02)$
 $= P(14-0.50| \le 0.02)$
 $= P(14-0.50| \le 0.02)$
 $= P(14-0.50| \le 0.02)$

Then by Telebyshift's theorem,

 $P(14-0.50| \le 2(0.01)) \ge 1 - \frac{1}{a^2}$
 $= \frac{3}{4}$

Thus, a lower bound on the probability is 3/4 which implies a lower bound for the number of coins out of 400 is $400 \times 3/4 = 300$.

