

Escape the Heat: The Dynamics of Migration as Adaptation to Climate Change

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Climate change will continue to cause large-scale changes to global weather patterns and extreme events. Migration has the potential to be a powerful limiting force against the damages from such changes to the climate. Indeed, I find that climate damages from business-as-usual warming would be 28% higher if domestic migration within Mexico was no longer available as a tool of adaptation to climate change. Our ability to appropriately and dynamically adapt to the damages from climate change relies on our ability to correctly anticipate future warming. The fraction of the population that I estimate as forming naive expectations of the climate system would experience an average of 2% fewer climate damages over their lifetime if they were to become fully-informed on the climate transition. Given that much of the increased damages faced by this population come from a reduced propensity to migrate, one way to reduce these losses is to subsidize migration. I find that subsidizing migration at the value to the climate-naive from becoming fully-informed reduces their welfare losses by 8–19%, depending on the dynamic structure of the policy. The difference in the effectiveness of different policy options is driven by the value one places on the ability to move in the future at a reduced cost. This option value is positive for dynamically-available policies; is increasing in the rate of warming for the fully-informed, but not the climate-naive; and causes dynamically-available policies to positively select from the portion of the population marginal to a static reduction of moving costs. To provide answers to my research questions, I develop and estimate a dynamic lifecycle model of migration within Mexico. I combine this with a non-stationary and spatially-varying model of the climate, in which I allow for both fully-informed and naive expectations of the future progression of climate change. Estimation of the climate model uses daily-level historical weather data and output from high-quality climate simulations. Estimation of the lifecycle model uses a sample of life histories, covering the years 1950–2019, and follows a nested full solution pseudo-maximum likelihood routine.

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1 Introduction

Earth's climate is changing in ways that modern humanity has never experienced. We are already observing large-scale changes to the physical systems governing our oceans, the atmosphere, and as a result, extreme heat and precipitation events throughout the planet (IPCC, 2022). Although we have yet to understand the full extent of these changes on the habitability of the many different regions of the world (Sherwood and Huber, 2010; Mora et al., 2017), history tells us that humans will find a way to adapt to our new environment. One of the most prominent ways that we have ancestrally adapted to environmental change, and specifically climate change, is through migration (Willis et al., 2004; Carto et al., 2009; Beverly, 2023).

In this paper, I ask how valuable migration will be as a tool to reduce the damages of *future* changes to the climate. In order to provide answers to this question, it is necessary to first understand how the decision to migrate will be impacted by the changing climate. I develop and estimate, using a full solution routine, a microfounded dynamic model of migration within Mexico in response to climate change. Individuals are modeled as making a series of sequential decisions of where to live within Mexico, where each time that they make this decision they understand that they will live in their chosen destination until they have the opportunity to move again. The dynamic nature of the decision of where to live requires that individuals form expectations over the future climate in each of their potential destinations. Given the complexity of the climate system, I allow for individuals to form either fully-informed or naive expectations over a climate system which is changing at different rates throughout both the country and time.

Migration has been found to be both costly and a difficult decision to reverse (Kennan and Walker, 2011). Moreover, past work has highlighted that a sizeable fraction of the population would experience prohibitively high psychological costs from moving (Koşar et al., 2022). Completely accounting for these costs is critical when seeking to understand the effect that migration will have on one's well-being. As such, the model allows for individuals to have such high psychological frictions from migration that they never consider it as an option. Those that do consider the trade offs of migration make their decision of whether and where to move based on a comparison of the expected benefit of deciding to move to a particular location, accrued across the remainder of their lifetime, with the psychological and monetary cost associated with making such a move today.

I recover a high valuation for the ability to migrate within Mexico. As of the 2020 Census, the average individual values their lifetime ability to migrate domestically at around two years of their annual salary. This value is increasing in the level of warming. One way to see

this is through the change in the value of the ability to migrate across a single generation. I find that individuals who were less than two-years-old during the 2020 Census will value their ability to migrate domestically 70% more than a 17–19-year-old during the same time-period. Another way to see that the value of migration is increasing in warming is to compare it to the value of migration in a world without climate change. I find that the average value of migration is over 4-times higher under a business-as-usual climate scenario than it would have been in a world without climate change.

I find that accounting for demographic and spatial differences in the population is vital to understanding the value of migration in limiting climate damages. Across time, it is those individuals who live in the warmest regions of the country and those who have been historically marginalized who stand to gain the most from climate migration. Those in 2020 who live in the warm regions of the country value their ability to migrate at over 13 times the value of those in the temperate regions. Those employed in agriculture have values 1.8 times those employed in other sectors. Those with the lowest levels of education have values 1.6 times higher than those with at least 12 years of education. Time will only work to expand these differences. Individuals living in the warm regions of the country who were less than 2-years-old in 2020 will face a value of migration which is over 23 times higher than those in their age-cohort who were born in the temperate regions of the country. Similar comparisons hold across one's sector of employment and level of education.

Our ability to migrate in response, or more broadly adapt, to climate change rests on our ability to correctly anticipate future changes to the climate. This has the potential to be a cause for concern, as recent international surveys have highlighted that although a large fraction of the population understands that climate change is happening, they also report a lack of understanding of it (Leiserowitz et al., 2022). Using the estimated model, I find that the value of migration, to the climate-naive, would increase by an average of 5% from obtaining and using correct information on the climate transition. Put another way, I find that the projected lifetime climate damages, on average, would be 2% fewer for the climate-naive if they were to become fully-informed.

Relative to the decisions one would make with full information, I find that, for the climate-naive, mistakenly choosing to remain in one's location leads to an average loss to lifetime welfare of \$9,007. This is much larger than the average welfare losses from making a mistake during one's move: \$3,796. Moreover, the climate-naive are more than three-times as likely to mistakenly remain in their location as they are to mistakenly move. Taken together, this implies that one way to reduce the welfare losses to this population is to subsidize domestic migration.¹ I find that such policy has the potential to be quite effective:

¹Such policies are commonplace: Canada, the United Kingdom, and the United States have tax codes

subsidizing migration at the value to the climate-naive from becoming fully-informed reduces the lifetime welfare losses borne by the climate-naive by between 8% and 19%.

The differences in the effectiveness of policies which affect moving costs are driven by the dynamic structure of the different policy options. More specifically, policies which allow individuals to choose *when* to use them exploit the value that individuals place on the ability to move in the future at a reduced cost. When the option value of such future opportunities is positive, dynamically-available policies will positively select from the portion of the population marginal to a similarly-priced static policy. In other words, the increased value of future opportunities changes the decision made by some of the just-marginal movers, who would instead prefer to postpone the decision to move into the future under the dynamic policy. I find that the effects on the selection into using the policy are of first-order concern: differences in *who* decides to use the different policies causes dynamically-minded reductions to moving costs to be over twice as valuable to movers as similar static policies. Moreover, the value placed on the ability to move in the future drives a similar increase in the value of the policy to the climate-naive.

To provide answers to my research questions, I estimate the lifecycle model using rich microdata from the Mexican Migration Project (MMP). These data contain information on the life histories for a sample of over 11,000 individuals living and working in Mexico from 1950 to 2019. Given the large changes to the climate in Mexico over this time period, and similar to the long-difference strategy employed in the reduced form literature ([Burke and Emerick, 2016](#)), I exploit true change in the climate over time to identify the relevant primitives of the model.

Estimation of the model primitives follows a two-step procedure. In the first stage, I use high-quality gridded weather data to estimate a non-stationary model of the climate system that varies throughout the country. I then estimate heterogeneous mappings from the relevant climate variables (average daily temperature and agricultural-season precipitation) to various statistics of weather which are designed to differentially capture the impact of extreme heat on crop yields and human well-being, along with the risk of drought and flood. I supplement the MMP data with historical surveys on wages to estimate a wage equation that depends on individual demographics, sector of employment, weather, and the location of residence. I take these estimated mappings to the second stage, where I estimate the model primitives. I employ a nested full solution pseudo maximum likelihood routine ([Rust, 1987](#)), taken over a finite mixture ([Heckman and Singer, 1984](#)), to estimate the model primitives: the preference parameters and the mixing probabilities for the unobserved moving and belief types.

which historically allow for tax-deductible moving expenses.

The estimation framework is valuable in that it allows for an internally consistent model of climate-induced migration which avoids the need to calibrate anything other than the discount factor. I combine the estimated model with data from the 2020 full-count Census and high-quality future climate projections under a business-as-usual scenario ([Thrasher et al., 2022](#)) to simulate the choices for each Census-enumerated male living in Mexico in 2020, across the remainder of their lifetimes. This framework allows me to make changes to the environment, through changes to the climate system and the moving costs, to recover welfare metrics of interest. For example, once the model is estimated, to estimate the value of migration one need only simulate decisions, and calculate resulting individual lifetime-welfare, under both the status quo and in a counterfactual environment where moving is completely restricted.

Relation to the Literature

This article builds on three main literatures. First, I contribute to the climate economics literature. One branch of this work has begun questioning whether reduced form panel methods connecting weather shocks to economic outcomes identify the effect of climate change ([Dell et al., 2014](#); [Mendelsohn, 2019](#); [Lemoine, 2024](#)). Migration is an outcome which is likely to be particularly sensitive to the arguments raised in this literature. In particular, past work has generally found that short-term and unanticipated weather shocks can either encourage ([Bohra-Mishra et al., 2017](#); [Jessoe et al., 2018](#)) or inhibit migration ([Cattaneo and Peri, 2016](#)). Such moves are generally short-distance and short-term ([Cattaneo et al., 2019](#)), although this is not always the case ([Feng et al., 2010](#); [Missirian and Schlenker, 2017](#)). Climate change may cause the studied weather shocks to become more frequent in the future, however it will also manifest itself through slow-onset and non-marginal changes in the distribution of daily weather ([IPCC, 2022](#)). I contribute to this work by providing a novel empirical framework to study how forward-looking individuals forming heterogeneous expectations over a changing climate will respond to true changes in the climate, represented both through changes to current-period weather shocks and through movements in the distribution of weather.

Another branch of this literature has emphasized the existence of an “adaptation gap” with respect to climate change ([Carleton and Hsiang, 2016](#)). [Zappalà \(2024\)](#) is the first, and to my knowledge only, article to highlight that a portion of this gap is likely driven by inconsistent beliefs about the climate. They use cross-sectional data on reported farmer beliefs surrounding recent changes in the frequency of drought to estimate how such beliefs lead to differential irrigation responses to short-term dry spells in Bangladesh. I complement

their work by instead using data on individual decisions to recover beliefs on the climate system, similar to [Kala \(2017\)](#). As in [Zappalà \(2024\)](#), I find that naive expectations of the climate system will lead to lower-than-optimal rates of adaptation to climate change. I highlight the generality of this finding, and extend the discussed literature by arguing for the broad role that policy, by subsidizing adapting behaviors, can play in helping to close the “belief gap.”

Second, I contribute to the literature seeking to uncover the value of migration, irrespective of climate change. One branch of this literature has used reduced form techniques, generally in the form of quasi-experiments. More specifically, they study how arguably exogenous shifts in the propensity to migrate change lifetime outcomes ([Deryugina et al., 2018](#); [Nakamura et al., 2021](#); [Sarvimäki et al., 2022](#)). This literature has found that migration can be quite valuable, however almost all of this work has found that this value varies strongly throughout the population.

Another branch of the literature studying the value of migration has used dynamic discrete choice methods to directly model the moving decision to account for the fact that migration is an endogenous choice ([Bishop, 2008](#); [Thom, 2010](#); [Kennan and Walker, 2011](#); [Lessem, 2018](#)).² Past work has been primarily limited to high-income countries and has generally not focused on climate change, and has therefore generally assumed a stationary environment. I contribute to this literature by highlighting how the value of migration can change, across the population, over time, in a non-stationary environment, in this case, from climate change. Moreover, similar to a budding literature in industrial organization (e.g., [Aguirregabiria and Jeon, 2020](#)), I relax the assumption of full-rationality that is oftentimes present in this general body of work by allowing for different levels of sophistication surrounding expectations of the climate system.

Third, I contribute to the burgeoning economic geography literature studying how populations will sort through space as a result of the spatially heterogeneous effects of climate change on fundamental productivities and amenities ([Desmet and Rossi-Hansberg, 2015](#); [Rudik et al., 2022](#); [Cruz and Rossi-Hansberg, 2023](#)). This body of work has found mixed results as to the value of migration in reducing climate damages: some work has found negligible impacts in developed countries ([Rudik et al., 2022](#); [Bilal and Rossi-Hansberg, 2023](#)) while others find large impacts globally ([Desmet and Rossi-Hansberg, 2015](#); [Cruz and Rossi-Hansberg, 2023](#)). The discussed literature has focused on the general equilibrium channels through which migration will respond to climate change.³ In contrast, I develop a rich partial equilibrium

²Previous work ([Thom, 2010](#); [Lessem, 2018](#)) mentioned above have used an earlier version of the same data used in this project; however, they study very different questions and focus primarily on international as opposed to domestic migration.

³As discussed in Section 6, I design a stylized framework to check how general equilibrium forces, through

framework upon which I am able to exploit microdata to cleanly identify the drivers behind climate migration, including the large moving costs associated with migration, which are allowed to vary across the population. Moreover, this framework allows me to study how expectations of future warming shape today’s behavior, where like the moving costs, these expectations are allowed to vary across the population. The micro-focused framework also allows me to ex ante understand how implementable policy would change decision making and welfare through its impact on individual incentives, and how such impacts will change with the climate.

The remainder of this paper is structured as follows: Section 2 provides a summary of the data and various reduced form empirics and summary statistics that motivate the dynamic model. Section 3 presents the dynamic model. Section 4 provides intuition on the estimation strategy and sources of identifying variation of the preference parameters. Section 5 presents the estimated parameters and details how well the model fits reality. Section 6 provides the setup used in counterfactual simulations, as well as the results from such exercises. Section 7 discusses and contextualizes the results of the paper, concludes, and provides avenues for future research on this topic.

2 Data & Empirical Motivation

I use data from the Mexican Migration Project (MMP) surveys to estimate the model, as described in Section 3. I supplement these data with historical climate data from [Livneh et al. \(2015\)](#) and [Thornton et al. \(2022\)](#); climate projection data from [Thrasher et al. \(2022\)](#); and wage data from the ENIGH and ENE Mexican surveys. Furthermore, I use data on demographic distributions, at the municipality-level, from the 2020 Mexican census count in counterfactual simulations.

Data

Mexican Migration Project

I use the life-history files from the Mexican Migration Project surveys to estimate the preference parameters governing the structural model. These surveys contain annual information on the location of employment, at the municipality level, for a sample of over 28,000 individuals⁴

congestion impacts on wages, may affect my primary results. I find that these impacts do not seem to be of first-order concern in my context. Much of this is likely due to the relatively short time horizon that I analyze.

⁴The MMP data primarily contain the life histories for the head of the household. In Mexico, during this timeframe, this an overwhelmingly male population. For this reason, I focus my analysis on males.

living in Mexico for each year of their life. I also make use of both demographic and personal characteristics contained in these files, such as their birth location, level of education, and the year of birth for each of their children.

Most climate-induced migration to date has been within a country's border (IPCC, 2022). Therefore, in order to study only internal migration, I restrict the life histories to only include individuals who have remained within Mexico for their entire recorded history. The publicly available life history data censor many of the municipalities of employment and birth to maintain the confidentiality of the respondents; I received the censoring algorithm and used it to fill in these missing observations where possible. I therefore further restrict the data to only those individuals who either have a known municipality of birth, or whose municipality of birth was successfully decensored. Similarly, I remove all individuals who, after employing the decensoring and an imputation algorithm, have any observations with missing employment municipalities. After making such restrictions, I am left with 11,194 unique individuals, covering the years 1950 through 2019.⁵ After aggregating the data from annual to triennial observations,⁶ I am left with 90,578 person-year observations for estimation.

Wages

A limitation of the MMP data is that although information on annual location and sector of employment is provided, they do not contain such information on income. To account for this, external survey data are used to estimate a series of wage equations, which are then used when estimating the model. The wage information used in estimation is obtained from two Census surveys - La Encuesta Nacional de Ingresos y Gastos de los Hogares (ENIGH), and La Encuesta Nacional de Ocupación y Empleo (ENOE). I use the ENIGH surveys from 1984 through 2006, and the quarterly ENOE surveys from 2007 through 2019. I subset the above survey data to match the selected MMP sample: only males, who are the head of the household and not self-employed, between the ages of 17 and 60, who reside in one of the municipalities lived in by an individual in the MMP data, with a positive monthly income, and whose sector is not listed as missing. This leaves me with 5,841,519 observations in the repeated cross-section. I discuss more details on these data in Appendix A.2

⁵As will be made clear in Section 3, I begin modeling each individual's choices when they turn 17. As such, I model decisions from 1967-2064. The end point of this range represents the year of the final decision for a 17-year-old in 2018.

⁶The aggregation procedure is discussed in Appendix A.1

Historical Climate

I use historical weather data to understand the relationship between extreme weather outcomes, migration, and wages. To do so, I obtain historical climate data from two sources; [Livneh et al. \(2015\)](#) (Livneh) and Daymet ([Thornton et al., 2022](#)). Both of these weather products are gridded data based on daily data from weather stations.⁷ These data are used to obtain measures of the temperature and total level of precipitation across each municipality in Mexico, where Livneh is used for the early years of the panel described in the previous section (1930–1979) and Daymet for the latter years (1980–2018). I match the datasets on their 1980 values to ensure that the combined panel is corrected for differences in levels between the two.

I use data on the standardized precipitation-evapotranspiration index (SPEI) to calculate both drought and flooding at the municipality-year level ([Vicente-Serrano et al., 2010](#)). SPEI is calculated by applying a normalizing transformation to the difference between precipitation and potential evapotranspiration at a given timescale and point in space. Following the literature (e.g., [McKee et al., 1993](#)), drought is defined as occurring when the calculated SPEI is below -1.5. To focus on drought most relevant to agriculture, I define a location as being in a drought when the minimum daily SPEI is below the drought cutoff during the summer agricultural growing season, defined as May through October, as in [Jessoe et al. \(2018\)](#). Given that the SPEI can be interpreted as the approximate number of standard deviations that the climatic water balance is above or below the mean, I define flooding analogously to drought - that is, a maximum daily SPEI above 1.5.

Climate Projections

I use the output from twenty-five high-resolution, bias-corrected climate simulation models, provided by NASA Earth Exchange (NEX) ([Thrasher et al., 2022](#)). These data provide information on daily projected temperatures and precipitation levels on a 0.25×0.25 degree grid, at a global scale. Even after bias-correcting algorithms are used, individual climate models are generally biased, in that their predictions are model-specific ([Aufhammer et al., 2013](#)). To partially remedy this, I perform two additional debiasing exercises. First, I average the daily gridded values over many of the individual models ([Burke et al., 2015; Rudik et al., 2022](#)). Second, I match the resulting data with the 2020 values from the historical weather data to correct for biases in levels from the climate model output. I use a moderate warming scenario to represent climate change - the Shared Socioeconomic Pathway 2-4.5 (SSP2-4.5). This climate pathway is one of five primary scenarios used in the Intergovernmental Panel

⁷Livneh is gridded at at $1/16^\circ$ resolution ($\sim 6\text{km}$ at the equator) and Daymet at 1 km.

on Climate Change's sixth assessment report and represents a business-as-usual scenario following historical patterns (Riahi et al., 2017). Hsiang and Kopp (2018) provides a high-level overview of climate change science for economists, in which they discuss the different SSPs.

Daily Measures of Weather

Each of the sources of weather data are produced at the daily level; given the annual nature of the MMP data, these daily weather outcomes must be aggregated up to annual statistics. To do this, I construct three different measures of temperature outcomes, differentiated by their importance for agricultural productivity and amenity values.

Following the literature (e.g., Feng et al., 2010; Liu et al., 2023), I model the impact of temperature on wages primarily through an agricultural channel, and more specifically, through its impact on maize production through growing degree days.⁸ A growing degree day (GDD) is a piece-wise linear function, with a crop-specific upper and lower discontinuity. For the case of maize, the functional form is:

$$GDD(\text{temp}) = \begin{cases} 0 & \text{if temp} < 8 \\ \text{temp} - 8 & \text{if } 8 \leq \text{temp} \leq 32 \\ 24 & \text{if temp} \geq 32. \end{cases}$$

This mapping from daily heat is used throughout the agronomy literature (Herrero and Johnson, 1980; Wilson and Barnett, 1983; Bassetti and Westgate, 1993), and captures the fact that plants are able to absorb additional energy from heat, up until some threshold. I calculate growing degree days at the municipality level as the sum of all daily levels across the summer growing season (May through October).

Growing degree days are specifically designed with crops in mind. To capture amenities over extreme heat, or lack thereof, I also make use of degree days above and below a given threshold x . Degree days above x is a recentered and truncated linear function in daily temperature, which is zero until x then increases linearly without bound when the temperature is above x . Similarly, degree days below x is zero when the daily temperature is above x and increases linearly in the absolute difference between x and the daily temperature when the daily temperature is below x .⁹ Such statistics are computed using the average daily

⁸Maize is the most sown crop in Mexico (SIAP, 2022) and is grown during the summer months in every state in the country. The impact of weather on wages is modeled as entering through maize yields. This directly impacts the wages of agricultural employees, which spills over into non-agricultural sectors through general equilibrium channels. One of these is through decreases in local demand for non-agricultural products (Liu et al., 2023). As will be made apparent in Section 4, I capture such effects through a richly specified first stage wage equation.

⁹Cooling and heating degree days are generally used by the United States Energy Information Admin-

temperature. To bring these statistics to the annual level, I sum the calculated daily-level degree days over all days in a given year. Mathematically, daily-level degree days above x can be written as:

$$DD_x(\text{temp}) = \begin{cases} 0 & \text{if temp} \leq x \\ \text{temp} - x & \text{if temp} \geq x, \end{cases}$$

and daily-level degree days below x as:

$$DD^x(\text{temp}) = \begin{cases} x - \text{temp} & \text{if temp} \leq x \\ 0 & \text{if temp} \geq x. \end{cases}$$

I use degree days above 26°C, which I label DD26, and degree days below 14°C, which I label DD14. I justify these thresholds in the next section.

Empirical Motivation

Migration is a relatively rare event in one's life. This fact is readily apparent in the MMP data: just over 2.3% of person-year observations include an inter-municipality move from 1967 through 2018. Moreover, since many individuals never decide to move during their lifetime, the gains from such moves are concentrated in a small subsample of the overall population. The MMP data records that just under 25% of the observed population ever move in their recorded life-history. Panel (a) of Table 1 looks at how the subpopulations of movers and non differ along a few important dimensions. In particular, migrants, on average, have higher levels of education, are less likely to be employed in agriculture and are somewhat older when they have their first child.

Panel (b) of Table 1 studies the characteristics of the moves that migrants take. In particular, most moves are taken when an individual is relatively young, with the average age of a migrant on their initial move being just over 22 years old, and the average age for all moves being a bit older - 25 years old. Almost 70% of moves are to rural locations. To put this into perspective, 70% of moves originating in a rural location are to another rural location, while 65% of moves originating in an urban location are to a rural location. Among individuals born in urban locations, this proportion is much smaller - about 38% for both moves originating in rural and urban locations. Figure A1 plots the empirical density of the number of moves taken by migrants. One can readily see the high value associated with the consideration of future migration: the majority of migrants (65%) move more than once,

stration to predict energy requirements for cooling and heating. They use a cutoff of 65°F for these calculations.

Table 1: Descriptive Statistics

	Non-Migrants			Migrants			2020 Census		
	Mean	Std.	Dev.	Mean	Std.	Dev.	Mean	Std.	Dev.
Panel (a)									
Education	7.66	4.25		8.76	4.96	9.03	3.63		
Agricultural Worker	0.32	0.47		0.22	0.41	0.31	0.46		
Children	0.95	0.21		0.93	0.26	-	-		
Age at Birth of First Child	23.83	4.90		24.37	4.90	-	-		
Panel (b)									
Number of Moves	-	-		2.16	1.84	-	-		
Age at First Move	-	-		22.69	6.30	-	-		
Move to Urban Location	-	-		0.32	0.46	-	-		
Stay > 1 Year	-	-		0.74	0.44	-	-		
Panel (c)									
Born in Dry Climate	0.28	-		0.26	-	0.24	-		
Born in Temperate Climate	0.41	-		0.48	-	0.59	-		
Born in Warm Climate	0.32	-		0.26	-	0.17	-		
Born in Urban Location	0.25	-		0.24	-	0.41	-		
Number of Individuals	8406			2788					

Note: Education and age are measured in years, agricultural worker is one if an individual works in agriculture at least 25% of their observed life history, and zero otherwise. Children is one if an individual is a parent at the time of the survey, and zero otherwise.

with a large fraction (21%) moving more than twice. Moreover, the average number of moves is just over two and almost three quarters of observed moves last over one year, with the average length of stay being 7.4 years (median of 4 years). Given the long duration of most moves, a model which does not allow for dynamic considerations will mistakenly attribute the value of individual moves, across time, to a single year. Moreover, dynamics are necessary to rationalize both the high number of moves taken by migrants with the relative infrequency of overall migration.

Panel (c) of Table 1 shows the distribution of birth locations across both the three primary Köppen climate zones in Mexico and urban locations. There is a good amount of variation along these dimensions, however, more importantly, the proportions of births within each of these groups matches somewhat closely to the locations of individuals in the 2020 decennial census. In particular, the 2020 census places 24% of the population in dry climates (versus 28%), 59% in temperate climates (versus 42%), 17% in warm climates (versus 30%), and 41%

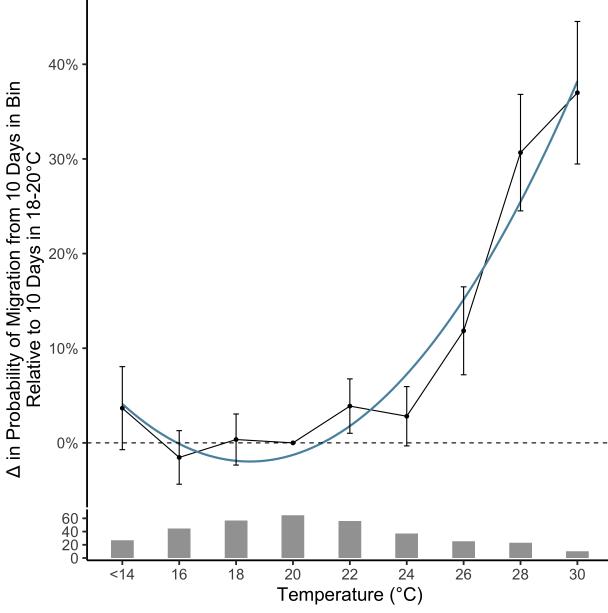


Figure 1: Nonlinear effect of daily temperatures on the probability of migration, conditional on expected wage, age, person and state-year fixed effects. Linear probability model using MMP data for migrants. The histogram plots the average number of days in each bin, taken across the estimation sample. 95% confidence bands, corresponding to standard errors clustered at the person and year-by-municipality level, are plotted around the point estimates.

in urban locations (versus 25%).

Figure 1 plots the results from a linear probability model of the decision to migrate on the number of days in the previous year that the average daily temperature falls within each 2°C bin, conditional on age, expected wage, individual and state-by-year fixed effects. Estimation of the marginal probabilities induced by this model makes comparisons only among migrants; comparing the number of days in each temperature bin in years that they do move to years that they don't, conditional on any macroeconomic trends that occur at the state-year level. The x-axis plots the rightward limit of the bin,¹⁰ with < 14 containing all days below 14°C, and > 30 containing all days with an average temperature greater than 30°C. The plotted effect shows the increase in the average propensity to migrate from replacing 10 days in the 18-20°C bin with 10 days in the given bin. One can see a clear nonlinear relationship between daily temperatures and migration decisions. In particular, days above 26°C are more than twice as important in predicting migration decisions as the very next bin containing days between 24 and 26°C. Although Figure 1 makes clear that heat is a much stronger correlate with the decision to migrate than cool temperatures, there is still a significant impact of cool days below 14°C. For this reason, I focus on degree days above 26°C and below 14°C.

¹⁰For example, “28” corresponds to the bin between 26 and 28°C.

Table 2: Daily Extreme Weather Increases the Probability of Migration

	<i>Dependent Variable: $\mathbb{1}\{\text{Migrate}\}$</i>			
	(1)	(2)	(3)	(4)
Average Temperature	0.00143*** (0.00032)	0.00142*** (0.00033)	0.0045*** (0.0013)	0.0106*** (0.0016)
Degree Days above 26°C	0.0129*** (0.0021)	0.0130*** (0.0021)	0.0059*** (0.0013)	0.0149*** (0.0026)
Degree Days below 14°C	0.00303*** (0.00055)	0.00318*** (0.00057)	0.0075*** (0.0014)	0.0083*** (0.0017)
Agricultural Worker	-0.0052*** (0.0010)			
Education	0.00084*** (0.00015)			
Children	-0.0136*** (0.0011)			
Num.Obs.	260472	260472	260472	260472
R2	0.052	0.050	0.168	0.201
FE: Year-State	✓	✓		✓
FE: Person			✓	✓

Note: DD26 (DD14) represents 100 annual degree days above (below) 26°C (14°C). Education is measured in years of schooling, agricultural worker and children are indicator variables representing whether an individual is employed in the agricultural sector and has children, respectively. Standard errors are two-way clustered at the person and year-by-municipality level. *p < 0.1; **p < 0.05; ***p < 0.01

Table 3: Reduced Form Mechanisms

	Dependent Variable: $\mathbb{1}\{\text{Migrate}\}$			
	(1)	(2)	(3)	(4)
Degree Days above 26°C	0.0113*** (0.0018)	0.0048*** (0.0010)	0.0180*** (0.0024)	0.0043** (0.0017)
Degree Days below 14°C	0.00189*** (0.00039)	0.0028*** (0.0011)	-0.000027 (0.001459)	-0.0012 (0.0013)
DD26: Away From Birth Loc	0.0071*** (0.0025)	0.017*** (0.004)	0.0123*** (0.0038)	
DD14: Away From Birth Loc	-0.00249*** (0.00078)	-0.0024 (0.0024)	-0.00029 (0.00233)	
Num.Obs.	260472	260472	260472	260472
R2	0.064	0.176	0.205	0.341
FE: Year-State	✓		✓	✓
FE: Person		✓	✓	
FE: Person-Location				✓

Notes: Degree days 26 (14) represents degree days above (below) 26 (14)° C. “Location” refers to the elements of the choice set, defined in Section 3 of the text. Standard errors are two-way clustered at the person and year-by-municipality level. *p < 0.1; **p < 0.05; ***p < 0.01

Table 2 presents the results from a series of linear probability models of the decision to migrate with different levels of fixed effects and controls. All four columns show that the primary weather variables of interest, average temperature, degree days above 26°C, and degree days below 14°C, are important correlates with the decision to migrate out of an area. This remains true even when making comparisons only among migrants (column 3), and even after making this comparison while additionally controlling for any local state-by-year-level macroeconomic shocks (column 4). The first column shows that the primary demographic variables used in the model strongly correlate with the decision to migrate. Agricultural workers are 23% less likely to move than those employed outside this sector; parents are 59% less likely to move, and each additional year of education, on average, correlates with a 4% increase in the likelihood of a move.

Although the above analysis shows that individual migration decisions are correlated with experienced heat and cold, it may be the case that the sampling locations from the MMP happen to be placed in locations which are of a different climate than the destinations of most moves. I look at two empirical robustness exercises in Table 3 to provide evidence against such possibilities. The first three columns study the differential impact of experienced degree days above 26°C and below 14°C on the migration propensity, when an individual is and is not currently residing in their location of birth. Column 3 is the preferred specification of

such analysis, including both person and state-year fixed effects. In particular, the estimates of the non-interacted degree day effects come from comparisons of the number of degree days when an individual lives in their birth location, in the years that they do and do not choose to migrate. I find that a 100 unit increase in degree days above 26°C increases their probability of migration by 78%. On the other hand, the estimate of the impact of degree days above 26°C when an individual is away from their location of birth compares the number of degree days above 26°C in all years that they are away from their birth location to the number of degree days in the years that they decide to move and are living in a location other than their birth location. I find that an increase of 100 degree days above 26°C increases the probability of migration by 53%. Most importantly, this exercise shows that individuals are responsive to the heat when they are both at and away from their location of birth.

The final column of Table 3 studies the impact of degree days above 26°C and degree days below 14°C on the probability of migration, controlling for person-location and state-year fixed effects. Locations are defined as the elements of the choice set, as given in Section 3, are areas of similar climates and financial opportunity. The estimation of this specification compiles the average number of degree days each year that an individual lives within a particular location and compares this number in years that they do not decide to move to the years that they do decide to move, holding fixed any macroeconomic shocks at the state-year level. I find that individuals are more likely to move in years when there happens to be higher than normal levels of degree days above 26°C, where normal is defined against the average number of degree days that this individual experienced in that particular location.

The above evidence has verified a high degree of correlation between experienced weather and the likelihood that an individual decides to move. I now turn to a discussion on both whether individuals are forward looking with respect to weather and then to how they form these beliefs over future outcomes. The first column of Table 4 shows that, conditional on person and state-year fixed effects, the current year's municipality-level experienced temperature is a strong correlate of the decision to migrate. The second column adds the previous year's temperature to the linear probability model. When doing so, the estimated coefficient in column 1 is practically split in half, with the current and previous year's temperature each receiving half of the importance in predicting migration decisions. One explanation for the importance placed on the previous temperature, conditional on the current temperature, is that individuals use this information in predicting next year's temperature, which is what is likely to drive this year's migration decisions. To provide evidence of this mechanism, the third column additionally adds next year's temperature to the model. One can now see that both the future and previous year's temperature, conditional on the current temperature are the important drivers behind the decision to migrate.

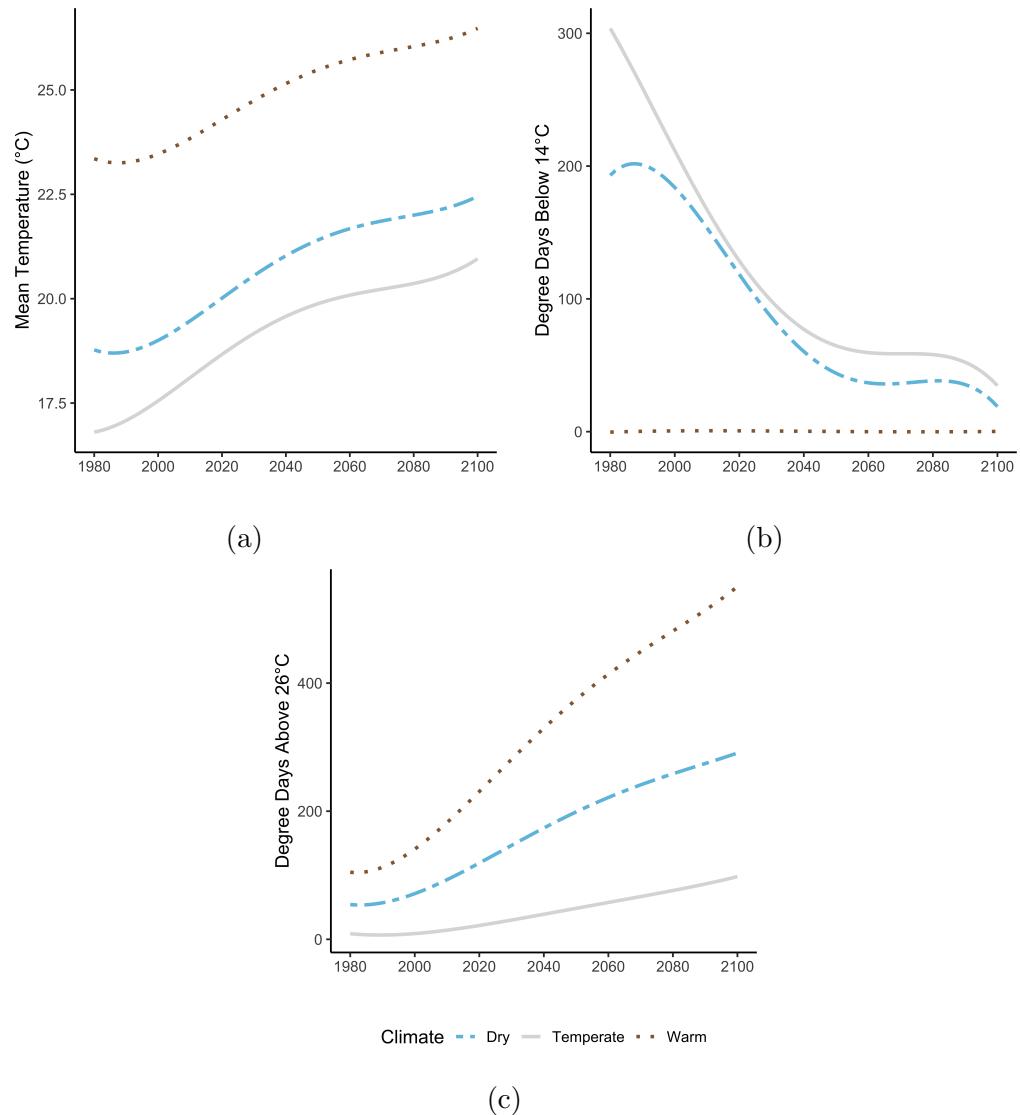


Figure 2: Climate transitions in Mexico under the CMIP6 Shared Socioeconomic Pathway 2-4.5, through 2100, and smoothed using a fourth-order polynomial in the year

Table 4: Model-Free Evidence of Forward Looking Beliefs over the Weather

	Dependent Variable: $\mathbb{1}\{\text{Migrate}\}$		
	(1)	(2)	(3)
Current Temperature	0.0097*** (0.0013)	0.0052*** (0.0014)	0.0011 (0.0016)
Last Year's Temperature		0.0050*** (0.0015)	0.0044*** (0.0015)
Next Year's Temperature			0.0052*** (0.0014)
Num.Obs.	260472	260472	260472
R2	0.200	0.200	0.200
FE: Year-State	✓	✓	✓
FE: Person	✓	✓	✓

Note: Temperature is measured in degrees Celsius. Standard errors are two-way clustered at the individual and year-by-municipality level. * $p < 0.1$; ** $p < 0.05$; *** $p < 0.01$

Table 5: Static Logit Parameters

Flow Utility	Moving Costs		
Income	0.059 (0.006)	Moving Intercept	2.891 (0.073)
Degree Days above 26°C	0.067 (0.058)	Distance	0.139 (0.011)
Degree Days above 26°C ²	-0.074 (0.021)	Children	0.129 (0.042)
Degree Days below 14°C	0.134 (0.011)	Age	0.209 (0.009)
Degree Days below 14°C ²	-0.022 (0.002)		
Urban Location	-0.342 (0.050)		
Birth Location	2.452 (0.024)		

Log-Likelihood: -25255.83

Note: Income is measured in annual hourly 2010 pesos, degree days are divided by 100, distance is measured in log kilometers. Standard errors are computed with the inverse Hessian.

I now turn to a discussion of whether individuals take into account not just the weather in their current location, but also that of their potential destinations, when deciding both whether and where to migrate. To do so, I estimate a static discrete choice model of migration within Mexico which is similar to the dynamic model discussed in the following section. In particular, I use a conditional logit framework to estimate the preference parameters corresponding to the utility of an individual living in location ℓ and then moving to location ℓ' :

$$\begin{aligned} u(\ell, \ell') = & \alpha_1 \text{income} + \alpha_2 \mathbb{1}\{\ell' \in \mathcal{U}\} + \alpha_3 \mathbb{1}\{\ell' = \nu^\ell\} \\ & + \alpha_4 DD26 + \alpha_5 DD26^2 + \alpha_6 DD14 + \alpha_7 DD14^2 \\ & - \mathbb{1}\{\ell' \neq \ell\} \times [\alpha_8 + \alpha_9 d(\ell, \ell') + \alpha_{10} \mathbb{1}\{\text{children}\} + \alpha_{11} \text{age}]. \end{aligned} \quad (1)$$

Here, $DD26$ represents last year's degree days above 26°C , $DD14$ last year's degree days below 14°C , \mathcal{U} the set of urban locations, ν^ℓ an individual's birth location, and d the Euclidean distance metric. Table 5 presents the estimated parameters. All the estimated parameters align with one's intuition. Individuals view higher incomes and living in the same area that they were born in as positive attributes in their location of residence. Moving is costly, and this cost increases with distance, parenthood, and age. Moreover, the estimated parameters on the weather show that individuals enjoy moderate amounts of both heat and cold, but dislike extreme levels of these outcomes. To put this into perspective, we can note that the change in utility from a change in the number of degree days above 26°C is:

$$\frac{\partial u}{\partial DD26} = \alpha_4 + 2\alpha_5 DD26.$$

This quantity is positive when $DD26 < -\frac{\alpha_4}{2\alpha_5} = 0.48$, or when the number of degree days above 26°C is less than 48. With this, the marginal impact of an additional degree day below the 68th percentile of the observed data has a positive impact on welfare. This is even stronger for degree days below 14°C ; an additional degree day below the 77th percentile of the observed data has a positive impact on welfare.

To provide further evidence of a forward looking dynamic over expectations of future weather on the decision to migrate, I modify the above static logit framework to additionally allow for preferences over future temperature. In particular, I include an additional preference parameter for the average of the daily temperature distribution over the next decade in the location chosen today. I present the estimated parameters in Appendix C, Table C2. I estimate a strongly significant preference for future expected temperatures, and take this as additional evidence that individuals seem to be forming forward looking expectations over

the climate when deciding where to live.

Given that heat and cold are enjoyed in moderation, but disliked in abundance, it is important to understand how the degree days of interest are projected to evolve in the future. Figure 2 plots changes in the relevant weather outcomes across the three primary Köppen climate zones and time under a projected moderate warming scenario.¹¹ Of particular note is the smooth increase in the average temperature across time, albeit at a reduced rate in the climate projections than under the historical observation. This increase in annual average temperatures maps strongly to increases in tail-events; Figure 2c highlights that a less than 2°C increase in average temperatures from 2020 through 2100 maps to a more than doubling in the number of annual degree days above 26°C. Similarly, Figure 2b shows that this increase maps to a sharp decrease in the number of degree days below 14°C in the dry and temperate climates in Mexico. The expected increase in the number of degree days and decrease in the number of degree days below 14°C work together to show the potential welfare losses via amenities from climate change.

This section has documented a correlation between experienced, future, and past temperatures with migration decisions. In the next section, I present the model framework, which allows me to study the role of dynamics in shaping the value of migration in response to climate change.

3 Model

Overview

Individuals are modeled as making a sequence of fifteen decisions of where to live. Each time they make this choice, they take into account both that the next time they will have the opportunity to move they will be a bit older and that they will be making that choice from an origin of today's choice. This decision of where to live takes place once every three years, beginning at age seventeen and ending at sixty-one.¹² These decisions are taken with respect to a choice set containing twenty-seven distinct options within Mexico, where each option is defined as a cluster of municipalities with similar climates and employment opportunities.

Individuals are modeled as varying, in unobserved ways, across two dimensions. The first deals with unobserved heterogeneity in the migration cost. An individual may be the kind of

¹¹Technically, there are four first level climates in Mexico, with the addition of a handful of areas in a polar climate. As there are no municipalities in Mexico with a majority of land area contained in polar climates, these areas are subsumed into nearby climates.

¹²The vast majority of observed moves are taken well before an individual is sixty-one years old - only a single individual moves (one time) after this in the MMP data used for estimation.

person who never truly considers the option of moving. This could be for a variety of reasons, including deep-rooted familial ties to their birth location, a strong fear of the unknown, or any other highly heterogeneous reason that a person absolutely desires to remain in the same location for their life. Or they can be the kind of person who actively considers the trade-offs of living in a different location, be it higher incomes, improved amenities and the like, from the monetary and psychological costs of having to move.

The second source of unobserved heterogeneity is with respect to the way in which individuals form beliefs about the climate in each of their potential choices. In the spirit of the literature studying the sophisticated ways that consumers form long-term expectations (e.g., [Busse et al., 2013](#)), I allow for individuals to form rich expectations about a spatially heterogeneous climate which is changing heterogeneously throughout time. Other work in economics has criticized the assumption that individuals are able to form correct expectations over complicated dynamics (e.g., [Simon, 1955](#)). For this reason, I additionally allow for individuals to form naive expectations about the future climate. I assume that these individuals see the weather in the current period and predict, with certainty, that all future weather will be the same. I use the model framework to estimate the unobserved portion of the population associated with each of the belief processes and moving types, at the same time as the other model primitives.

After observing the current year's weather and forming expectations of the implied future climate, but prior to making the decision of where to move to this period, an idiosyncratic, transitory, and location-specific preference shock affecting a combination of incomes and moving costs is realized. This individual then compiles the available information and chooses to live in the location which maximizes their lifetime welfare, as defined against their expectation process. If they decide to move to a new location, they then pay an upfront cost which depends on their demographic characteristics, as well as the characteristics of their origin and destination locations. After making this decision, agricultural season drought and flooding shocks affecting their income are realized.¹³ They then obtain three years worth of discounted flow utilities associated with their destination, and they begin the next period with an updated set of state variables, in their current location.

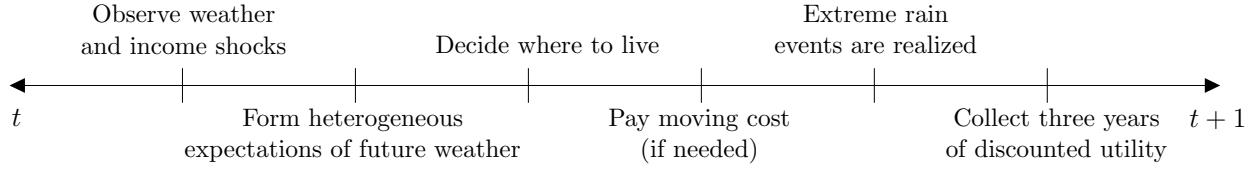


Figure 3: Within Period Timing

A Dynamic Model of Where to Live

More formally, take some individual, labeled by i , in some time period $t \in \{1, \dots, 15\}$. At the start of period t , i begins in their previous-period choice location, ℓ , and observes their vector of state variables, ω , along with a vector of location-specific utility shocks, ε . Given this framework, an individual receives a flow utility associated with the decision to move to any particular location, ℓ' , as given by $u(\ell, \ell', \omega, r) + \varepsilon_{\ell'}$, where r is a two-dimensional vector containing the drought and flooding realizations. Individual i is modeled as sequentially choosing locations ℓ'_t , consistent with maximizing the following discounted sum of future flow utilities, given their time t decision¹⁴:

$$\max_{k \in \mathcal{C}^{|T-t|}} \left\{ \sum_{s=0}^{T-t} \delta^{3s} \mathbb{E} [u_s(k_s, k_{s+1}, \omega_s, r; \theta, \tau) + \varepsilon_{k,s} | \omega, \tau] \right\}.$$

When making this decision i is assumed to be fully aware of their current state, as given by ω_τ , and the current period vector of idiosyncratic shocks. However, future realizations of these values are unknown from a time t perspective, as are the within-period rainfall shocks. Therefore, i makes use of their knowledge of the distribution over which these variables are realized in forming their decision of where to live today. Mild regularity conditions on this joint distribution, along with an assumption of the separability of the deterministic flow utility values from the stochastic utility shocks, allow for a translation from the above sequential representation of the modeled decision to that of a Bellman equation. Specifically,

¹³Modeling extreme rainfall events as occurring after the moving decision is done for two primary reasons. The first is that it better accounts for higher levels of uncertainty surrounding localized precipitation than localized temperatures. The second is that, conditional on the realized weather states, it is assumed to be uncorrelated across time (this is a computational requirement, in that these events are therefore not elements of the state space).

¹⁴I calibrate the annual discount rate, δ , to 0.95.

we can reformulate the time $t < T$ decision as:¹⁵

$$V_t(\ell, \omega; \theta, \tau) = \max_k \left\{ \mathbb{E}_R[u(\ell, k, \omega, r; \theta, \tau)] + \varepsilon_{tk} + \delta^3 \mathbb{E}_{\Omega, \varepsilon|\tau}[V_{t+1}(k, \omega'; \theta, \tau)|\omega] \right\}.$$

And the time T value function, representing the sum of expected current flow utility and a discounted retirement value as:

$$V_T(\ell, \omega; \theta, \tau) = \max_k \left\{ \mathbb{E}_R[u(\ell, k, \omega, r; \theta, \tau)] + \varepsilon_{tk} + \frac{\delta^3}{1-\delta} \mathbb{E}_{R, \Omega|\tau}[u(k, k, \omega', r'; \theta, \tau)|\omega] \right\},$$

where the final term of the above represents a retirement value associated with living indefinitely in location k , without the ability of further moves. Assuming that the state no longer evolves after an individual's final decision allows me to use a recursive strategy to fully solve the model, as in [Gowrisankaran et al. \(2024\)](#).

Formulating the individual decision in terms of a Bellman equation highlights the dynamic interplays present in such a model. Of particular note, each period an individual faces a high number of temporal trade-offs between current and future utilities, including an individual's awareness of future shocks which may unexpectedly change the future flow utilities associated with any decision made today. These shocks enter through both the idiosyncratic draws and future weather realizations throughout Mexico.

Components of Flow Utilities

Flow utilities associated with living in any particular location ℓ include three components: income, temperature as an amenity, and level shifters from both living in one's location of birth and living in an urban location. In particular, flow utility of an individual deciding to live in ℓ' is modeled as:

$$\begin{aligned} \bar{u}(\ell', \omega, r; \theta) = & \theta_1 \text{inc}(\ell', \omega, r) \\ & + \theta_2 DD26(\ell', \omega) + \theta_3 DD26(\ell', \omega)^2 + \theta_4 DD14(\ell', \omega) + \theta_5 DD14(\ell', \omega)^2 \\ & + \theta_6 \mathbb{1}\{\ell' = \nu^\ell\} + \theta_7 \mathbb{1}\{\ell' \in \mathcal{U}\}. \end{aligned} \quad (2)$$

The above parameterization is given for a single year. However, since each model period is three years long, an individual's period flow utility is instead given by the present discounted sum of the annual flow utility over three years. Specifically, each period an individual receives

¹⁵ \mathbb{E}_R is an expectation over the extreme rainfall events, while $\mathbb{E}_{\Omega, \varepsilon}$ is an expectation over the joint state variable and utility shock distribution. \mathbb{E} , with no subscript, is an expectation over the joint distribution of all three sources of stochasticity.

a flow payoff of $\sum_{s=0}^2 \delta^s \bar{u}(\ell', \omega)$.

Income

Income is parameterized as a function of one's location, individual-specific state variables and the weather, specifically, a flexible function of growing degree days, drought and flooding.¹⁶

In particular,

$$\text{inc}(\ell, \omega, r) = f(\ell, \omega, r).$$

Where f is assumed to be a known mapping to be estimated, and r contains indicators for drought and inundation, respectively. Within-period uncertainty over wages comes from uncertainty about the extreme rainfall outcomes, and future uncertainty comes from the transitory shocks, weather shocks, and extreme rainfall outcomes.

Temperature as an Amenity

Weather is allowed to affect flow utility both indirectly through wages and directly as an amenity. Amenity values are flexibly modeled as entering through extreme degree days, that is, through degree days below 14° C and degree days above 26° C. These values were chosen for two primary reasons. The first is that almost all the locations in the choice set have historically experienced nonzero levels of these statistics.¹⁷ The second is that, as per Figure 1, these thresholds are those at which a nonlinearity in the reduced form probability of migration are present. The specification presented in Equation 2 remains agnostic towards the sign and magnitude of the linear and quadratic terms for both extreme heat and cold. In particular, it allows for a bliss-level of both where utility is increasing until this point and then decreasing after it.

Flow Utility Shifters

In order to appropriately account for the high proportion of the population who either never leave their location of birth, or who leave and subsequently return, I allow for a level shifter of period utility from living in an individual's location of birth, as has been found to be important in past work (Kennan and Walker, 2011). Further, to account for differences in the populations living in urban and rural locations that are not explainable by income differentials, I also allow for a utility shifter from living in an urban location.

¹⁶This parameterization is discussed in Section 4.

¹⁷22 of 27 locations have experienced nonzero degree days above 26°C, while 26 have experienced nonzero degree days below 14°C.

Moving Costs

Given the relative infrequency of moves seen in data, it is important to accurately capture the costs associated with a given move. I model moving costs as depending on both observed and unobserved sources of individual, origin, and destination characteristics. This cost is paid upfront in the first year of a model period that an individual decided to move. I parameterize it as:

$$\bar{c}(\ell, \ell', \omega; \theta, \tau) = \gamma_1(\tau) + \gamma_2 d(\ell, \ell') + \gamma_3 \mathbb{1}\{k_\omega \geq 1\} + \gamma_4 \text{age}_\omega. \quad (3)$$

One source of unobserved, fixed within individual, heterogeneity enters through $\gamma_1(\tau)$; d represents the Euclidean distance metric, k_ω the number of children, and age_ω the current period time-step.

I allow for two unobserved permanent moving-types of individuals. The first type is composed of those individuals who consider the trade-offs associated with moving each period. These individuals are used to estimate the magnitude of γ_1 . The other type is composed of individuals who do not actively consider the option to move in any period; they have a γ_1 which is prohibitively high. Allowing for this source of unobserved heterogeneity is standard in the literature (Kennan and Walker, 2011; Lessem, 2018; Oswald, 2019), particularly, because past work has generally found that large masses of the population have very high psychological costs associated with moving (Koşar et al., 2022).

Observed individual heterogeneity enters through level-shifters in the cost function. Particularly, the moving cost is allowed to differ for individuals as they progress through their lifetime by allowing this cost to depend on their age. The moving cost also is allowed to be different once an individual has had their first child. Heterogeneity in costs across space are modelled through different moving costs by the distance from an individual's current location to that of their destination.

A Model of Heterogenous Climates in Mexico

The climate is generally defined as a high dimensional distribution linking physical systems, such as oceans and the atmosphere (e.g., Hsiang and Kopp, 2018). Weather, itself a high dimensional object, is just a realization of this distribution at a particular time and place. Computational feasibility requires that I take a summary statistic approach to modeling climate dynamics. I model the mean of the daily temperature distribution and the level of agricultural-season precipitation as following a multivariate Gaussian, with the mean and variance estimated using data from the previous 30 years, from the perspective of year y .

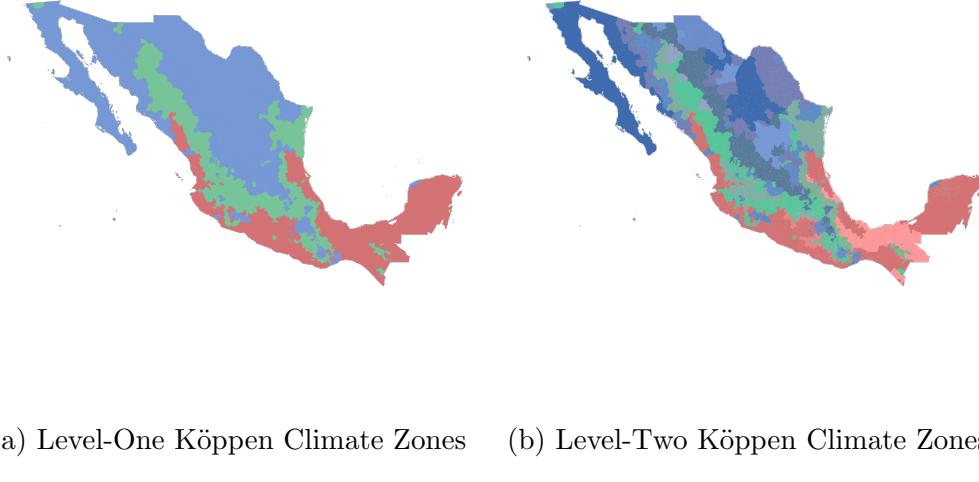


Figure 4: Köppen Climates in Mexico

That is,

$$\begin{pmatrix} \text{temp} \\ \text{prec} \end{pmatrix} \sim \mathcal{N}(\mu^y, \Sigma^y), \quad (4)$$

where temp and prec are both three-dimensional vectors containing the relevant realizations across the three level-one Köppen climates in Mexico. Figure 4a plots the level-one Köppen climate zones in Mexico, with blue representing the “dry”, teal the “temperate”, and red the “warm” climates. As is discussed in Section 4, individuals, in part, choose over locations defined against level-two climates, which are nested within the level-one climates. Figure 4b plots these throughout space. Appendix A.3 provides a detailed discussion of this classification system and how it is used in this paper.

The model allows for two different belief formation processes, and as discussed in Section 4, the mass of the population associated with each is estimated along with the other model primitives. The first process is one of naivety. Individuals aligned with this framework observe the weather each period and assume that all future weather will be drawn from a degenerate distribution at this point. In particular, I only assume that these individuals are aware of the realized weather in the current period for all locations in their choice set.

The second process is one of fully-informed expectations. More specifically, individuals who form expectations under this process understand the functional form of μ^y and Σ^y , as given in 4. Where:

$$\mu_y^w = \alpha_w^y + \alpha_{w1}^y \mu_{y-1}^w + \alpha_{w2}^y y,$$

and w represents the level-one average temperature or rainy-season precipitation. I model

fully-informed individuals as using realized weather in the previous 30 years, as is generally accepted as the length of a “climate”, to form their understanding of α^y and Σ^y .

Although computational tractability requires only tracking summary statistics of the climate system, past work has found that local weather impacts outcomes in a nonlinear fashion (Schlenker and Roberts, 2009; Aragón et al., 2021; Miller et al., 2021). To appropriately account for this in the dynamic model, conceptually, I allow for heterogeneous climate distributions for each element of the choice set, where the particular climate distribution is determined by the relevant level-one weather realizations.¹⁸ I assume that, on average, the individuals who form fully-informed expectations about the climate system are aware of these distributions and that they use the relevant realized level-one climate temperature to pin down the daily temperature distribution for a particular location. They use the expected daily distribution to integrate out any necessary degree day calculation. Operationally, I assume that they use location-specific mappings from the level-one climate mean temperature into the expected number of degree days in a location. Mathematically, they understand that the expected number of a particular type of degree day (DD) is given by:

$$DD_\ell = \mathbb{1}\{\lambda_\ell^D + \lambda_{1\ell}^D \text{temp}_{C(\ell)} > 0\} \times [\lambda_\ell^D + \lambda_{1\ell}^D \text{temp}_{C(\ell)}], \quad (5)$$

where $C(\ell)$ represents the level-one climate for location ℓ and temp the realized mean of it’s daily temperature distribution.

Precipitation levels are generally of first order concern in agriculture (Deschênes and Greenstone, 2007; Seo and Mendelsohn, 2008; Dell et al., 2014), however, their impact on outcomes has proven difficult to capture empirically. This is especially true when precipitation is aggregated to large spatial areas. To account for this, I instead model that individuals use a mapping from level-one climate agricultural season precipitation levels to location-specific probabilities of drought and flooding. In particular, they form expectations based on the following:

$$\text{prob}(R_\ell) = \Phi(\eta_\ell^R + \eta_1^R P_{C(\ell)} + \eta_2^R \text{temp}_{C(\ell)}), \quad (6)$$

where, R is either drought or flooding, and again, $C(\ell)$ represents the level-one climate for location ℓ , P the agricultural-season precipitation level, and temp the mean of the daily temperature distribution. Appendix B.1 discusses these mappings and their estimation in greater detail.

Figure 5 provides a graphical representation of the climate model used by the individuals forming fully-informed expectations. Starting at the top, these individuals understand that

¹⁸For example, both “dry warm dry” and “dry warm temperate” are different level-two climates (and therefore define different locations in the choice set) who share a level-one climate of “dry”.

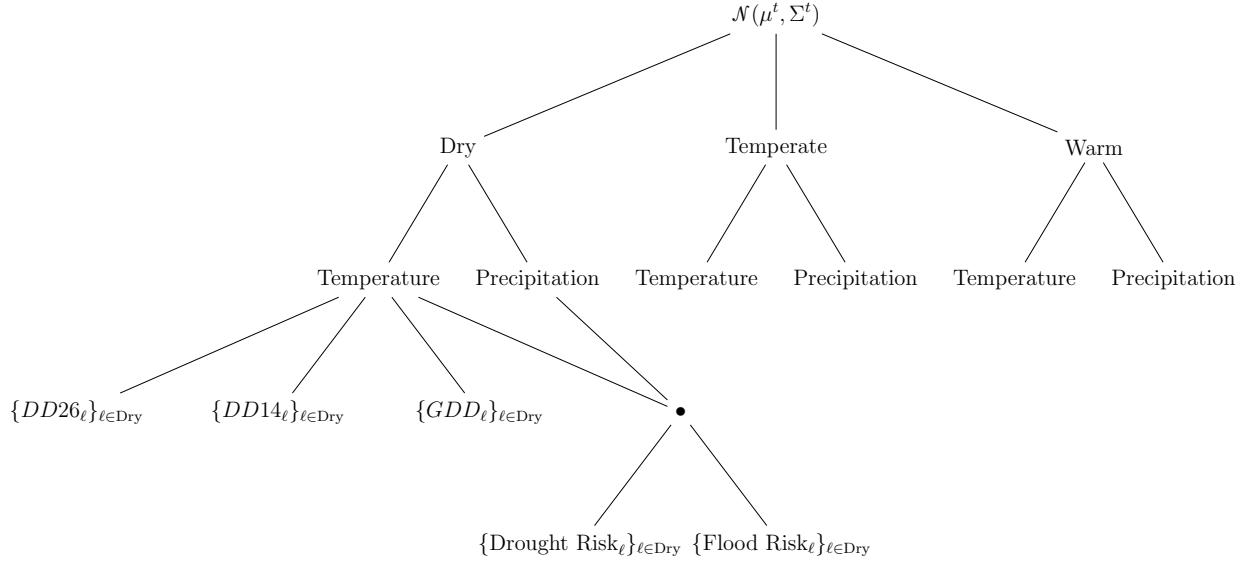


Figure 5: Graphical Representation of the Climate Model

the weather is drawn from a Gaussian characterized by the mean and variance described earlier in this section. A draw from this distribution gives the mean daily temperature and level of agricultural-season precipitation for each of the three primary Köppen climates: dry, temperate, and warm. These individuals then use these values to form expectations about the number of degree days and the probability of drought and flooding in each location. They form these expectations based on Equations 5 and 6, respectively. Given that these equations are estimated in a first stage, this allows for a rich set of different weather variables to influence migration decisions. In particular, I use six different weather variables to characterize an individual's state. After combining these variables with the estimated mappings, I am able to use 141 different weather variables in the dynamic model.

4 Identification & Estimation

Identification

I briefly discuss the intuition behind the identifying variation for each of the model primitives; a more formal argument is presented in Appendix B.3. Simply put, the income and temperature flow utility parameters are identified from two sources of variation in the life history data. Conditional on making a move, the first source of identifying variation compares differences between the given quality in the observed choice location and that in all other options. For example, part of the income parameter comes directly from comparisons of the expected

income in an individual's choice location to that across the menu of choices an individual faced in the year of their move. The second source of variation identifying these parameters instead comes from comparisons in average rates of migration, across differences in these qualities, conditional on all other state variables. More specifically, this source of identification holds constant all state variables (and other preference parameters) except that which one is seeking to identify and compares how changes in this variable influence the expected migration rate within each state variable grouping.

The variation identifying the preference for living in one's home location also comes from two facets of the data. The first of these compares the aggregate migration rate away from an individual's home location to the migration rate away from all other locations. The second compares the duration of stay when an individual is and is not in their birth location. Similar comparisons are made when identifying the preference for living in urban locations.

Identification of the mixing probability for the different moving types comes from a comparison of the migration rate for individuals on their second or later move to those who would be making their first move. Intuitively, the first-time migration rate contains a mixture of information about both moving types, while the migration rate for later moves only contains information about the movers. The estimation routine exploits this difference, conditional on all other state variables to back out the mass of the population associated with either type. The moving cost intercept is identified from the migration rate across the population, conditional on all state variables and other parameters, most importantly, on the mixing probabilities. This means that its value is identified from comparisons of migration rates within first and later moves. The distance parameter is identified from comparisons of the distance traveled for each move in the data, the different cost born by parents comes from migration rate comparisons across individuals with and without children, and the age parameter comes from migration rate comparisons across the age distribution.

Finally, identification of the differential masses of each climate belief type is identified from the variation, caused by climate change, in the future relative values of living in the different locations in an individual's choice set. To build intuition, imagine that an individual can only choose between two locations, a and b . Further, suppose that today a is preferable to b , but that a 's future desirability is decreasing at a faster rate than that of b . A fully-informed individual would choose b with a higher probability than a naive individual. The estimation framework, discussed later this section, makes comparisons such as this one across the model framework and the migration decisions observed in the MMP data to identify the mass associated with either belief process.

As a final point of emphasis, much of the model's ability to identify point estimates associated with the preference parameters rests on normalizations imposed on the stochastic

shocks. In particular, I normalize the scale of this distribution to one, which allows for separate identification for each preference parameter.¹⁹

Choice Set

The choice set is defined as the product of the level-two Köppen climate zones, in locations where individuals from the MMP data were seen to reside, and an indicator for urban status, all defined at the municipality level. Figure 4 plots the location of each level one and two Köppen climate group, where the level-one climates are shown as salmon for the “dry” climates, green for the “temperate” climates and blue for the “warm” climates.

The second element of the choice set is the classification as either rural or urban. I use the Mexican government’s metro zone classification to make this distinction. In particular, a municipality is considered “urban” if there is a metro zone within its boundaries, and “rural”, otherwise. Each element of the choice set is therefore a cluster of municipalities, where a given cluster is defined such that it contains all municipalities of the same level-two climate grouping and “urban” status.²⁰

Transition Probabilities

I estimate the state variable transition distribution in a first stage. The estimated distribution is used in the second stage to integrate over the value of future states within each of the available choices. For numerical tractability, I assume that occupation (in or out of agriculture), education level, and one’s birth location are constant within individual. Appendix B provides evidence in support of these restrictions. Except for one’s moving-belief type, the remaining state variables transition across an individual’s lifetime.

Given the triennial structure imposed on the model timing, an individual’s age increases by 3 years each time step with probability one. If an individual has had a child up to the current period, then they do not transition out of the parenthood state, however, if they have not had a child, then they become a parent with probability $p_c(a)$. This probability depends on a second-order polynomial of their age and is estimated using a probit regression in a first stage on the MMP data.

I compress each year into its decade, which transitions to the next element with probability 0.3. The decade is used to track the relevant climate distribution, as described in Section 3.

¹⁹An alternative strategy would instead normalize the marginal flow utility of income to one and instead estimate the variance of the stochastic shocks. This identification would follow from comparisons of migration outcomes across individuals with identical states who make different decisions.

²⁰There is one climate group with no metro zones within any municipality in its borders; therefore, there are 27 different locations in the choice set.

The model's solution procedure, discussed later in this section, requires that the model be solved through the year 2064. For all years beyond 2019, the parameters are estimated using data from the CMIP6 business-as-usual climate scenario. Estimates of all first-stage transitions are provided in Appendix B.2.

Wages

Expected wages are estimated in a first step, using external census survey data, as discussed in Section 2. I assume that expected wages are of the following form:

$$\begin{aligned} \text{inc}(\ell, \omega, r) = & \sum_a \sum_e \beta_{a,e} + \beta_{ag} ag(\omega) \\ & + \sum_{c=1}^3 [\beta_c^{GDD,C} GDD(\omega) + \beta_c^{GDD2,C} GDD(\omega)^2 + \beta_c^{D,C} D(r) + \beta_c^{I,C} I(r)] \\ & + \sum_{ag=0}^1 [\beta_{ag}^{GDD,Ag} GDD(\omega) + \beta_{ag}^{GDD2,Ag} GDD(\omega)^2 + \beta_{ag}^{D,Ag} D(r) + \beta_{ag}^{I,Ag} I(r)] \\ & + \xi_\ell + \eta_y(\omega), \end{aligned} \quad (7)$$

where a indexes age, e education level, and ag whether an individual is employed in agriculture. This functional form allows for a large degree of flexibility, while still allowing for a structural interpretation of the resulting parameter estimates. Each location has its own intercept, ξ_ℓ . The age-education intercepts allow for a flexible return to experience and education, which, given the year fixed effects $\eta_y(\omega)$, accounts for potential biases arising from differences across birth-cohorts. The remaining terms represent the differential impact across level-one climates and agricultural employment of different measures of weather on wages. $GDD(\omega)$ represent agricultural-season growing degree days, entering into wages quadratically, have been shown to be important indicators of wages and migration decisions in this context (Jessoe et al., 2018). D represents an indicator for drought and I for inundation (flooding), both measured during the agricultural growing season. Appendix A.2 provides details on the estimated coefficients from Equation 7.

Estimation Framework

I estimate the model primitives using a nested full solution routine within a pseudo-maximum likelihood framework (Rust, 1987). I briefly discuss the format of this procedure here, and provide more details in Appendix B. The likelihood is a mixture over the unobserved moving-belief types. In particular, for a given unobserved type τ and a vector of structural

parameters θ , the log-likelihood is the sum of the individual log-likelihoods, where each individual likelihood integrates over the distribution of unobserved heterogeneity: the moving and belief types, represented by τ . Mathematically:

$$\begin{aligned}\Lambda(\theta, \pi) &= \sum_i \log (\mathbb{E}_\tau[\mathcal{L}_i(\theta; \tau)]) \\ &= \sum_i \log \left(\sum_\tau \pi_\tau \cdot \mathcal{L}_i(\theta; \tau) \right).\end{aligned}\quad (8)$$

Given θ and π_τ , Λ can be computed directly from the data, with the interpretation of it being the log-likelihood of observing the choices made in the MMP data, given the model structure. The assumption that ε_ℓ follows a type-1 extreme value distribution provides a closed form of $\mathcal{L}_i(\cdot)$. To see this, we can first define the alternative-specific value function ($\bar{V}_t(\ell, k, \omega; \theta, \tau)$) as the value afforded to an individual with state ω , who currently lives in ℓ , from choosing option k , net of the current period stochastic shock. Mathematically,

$$\bar{V}_t(\ell, k, \omega; \theta, \tau) = \mathbb{E}_R[u(\ell, k, \omega, r; \theta, \tau)] + \delta^3 \mathbb{E}_{\Omega, \varepsilon|\tau}[V_{t+1}(k, \omega'; \theta, \tau)|\omega]. \quad (9)$$

Conditional on the structural parameters and the unobserved type, the individual's choice problem can then be rewritten as:

$$\max_k \{\bar{V}_t(\ell, k, \omega; \theta, \tau) + \varepsilon_k\}.$$

Following [McFadden \(1973\)](#), rewriting the choice problem in this way allows one to write the model-generated choice probabilities as:

$$\mathcal{L}_{it}(\theta; \tau) = \frac{\exp(\bar{V}_t(\ell_{it}, \ell_{it+1}, \omega; \theta, \tau))}{\sum_{j \in \mathcal{C}} \exp(\bar{V}_t(\ell_{it}, j, \omega; \theta, \tau))}.$$

This formulation of the choice probabilities makes clear that one need only obtain the set of alternative-specific value functions at each time step and location in the state space.

The distributional assumption on the preference shocks additionally assists with this endeavor. In particular, Equation 9 can be rewritten as:

$$\begin{aligned}\bar{V}_t(\ell, k, \omega; \theta, \tau) &= \mathbb{E}_R[u(\ell, k, \omega, r; \theta, \tau)] + \delta^3 \mathbb{E}_{\Omega, \varepsilon|\tau}[V_{t+1}(k, \omega'; \theta, \tau)|\omega] \\ &= \mathbb{E}_R[u(\ell, k, \omega, r; \theta, \tau)] + \delta^3 \int \max_j \{\bar{V}_{t+1}(k, j, \omega'; \theta, \tau) + \varepsilon_j\} dF_{\omega', \varepsilon|\omega, \tau}\end{aligned}$$

$$= \mathbb{E}_R[u(\ell, k, \omega, r; \theta, \tau)] + \delta^3 \int \log \left(\sum_{j \in \mathcal{C}} \exp(\bar{V}_{t+1}(k, j, \omega'; \theta, \tau)) \right) dF_{\omega'|\omega, \tau},$$

where the second line uses the definition of the alternative-specific value function to replace the value of beginning next period in location k with state ω' with the maximum alternative-specific value, inclusive of each preference shock, across the different choices available.²¹ The third line exploits the iid nature of the T1EV preference draws to break apart the joint expectation into a sequential expectation over the state variable transitions and the preference shocks. It then uses the properties of the T1EV distribution to rewrite the expectation of this maximum in a closed form.²²

To obtain the set of alternative-specific value functions, given a vector of preference values, I employ a backward-recursion strategy (Wolpin, 1984). Briefly, given some vector of preference parameters, one can start at the terminal period and calculate the set of alternative-specific values associated with making any decision from the perspective of an individual at every spot in the state space. The time T values can then be used to calculate the time T-1 alternative-specific value functions. One can work backward in this fashion until they recover the full set of alternative-specific value functions. With these in hand, one can directly calculate Equation 8 and maximize its value over the structural parameters and mixing probabilities to obtain consistent estimates of their values (Rust, 1988).

Substitution Patterns

As a minor point of emphasis, the described framework additionally relaxes the independence of irrelevant alternatives (IIA) assumption that is ubiquitous in much of spatial literature to allow for a more accurate representation of the substitution patterns that a potential migrant in Mexico faces when they make the decision of where and whether to move (Chipman, 1960; de Dios Ortuzar, 1983; Brownstone and Train, 1998). To see why this is important in the context of migration and climate change, I provide the following stylized example. Suppose that individuals choose from three locations in the country: dry, temperate, and warm. Further, suppose that in relative terms, the temperate location is becoming more desirable with time, while the dry and warm locations are becoming less desirable with time. Moreover, suppose that the desirability of the warm location is declining faster than that of

²¹Conditional on one's unobserved type, the integral on the second line is thirty-six dimensional. This comes from 27 location shocks, and the state transition for: one's age and parental status, the year, and the six weather state variables. I don't include one's birth location, their education, and their sector of employment in this calculation, as computational tractability requires that those are assumed constant within individual.

²²The resulting integral is now nine dimensional.

the dry location. Notice that the increase in the value of living in the temperate location would cause a higher proportion of the population to decide to live there. IIA substitution patterns would require that this increase is proportionately driven by the warm and dry locations. In reality, we may expect that relatively more people would move to the temperate location from the warm than the dry location. The dynamic framework that I employ allows for such substitution patterns to occur.

5 Results

Model Estimates

Table 6: Structural Parameter Estimates

Description	Parameter	Coefficient	Standard Error
Flow utility			
Income	θ_1	0.006	(0.0007)
100 Degree days above 26°C	θ_2	0.034	(0.005)
100 Degree days above 26°C squared	θ_3	-0.015	(0.002)
100 Degree days below 14°C	θ_4	0.020	(0.002)
100 Degree days below 14°C squared	θ_5	-0.003	(0.0003)
Living in an urban location	θ_6	-0.041	(0.003)
Living in location of birth	θ_7	0.183	(0.002)
Moving costs			
Fixed migration cost, for movers	γ_1	2.599	(0.073)
Distance of move	γ_2	0.140	(0.011)
Cost shifter: parenthood	γ_3	0.072	(0.022)
Cost shifter: age	γ_4	0.172	(0.006)
Unobserved heterogeneity			
Mass of movers	π_m	0.628	(0.013)
Mass of fully-informed expectations	π_b	0.673	(0.206)

Note: Income is measured in annual hourly 2010 pesos. Distance is measured in log kilometers. The likelihood contains 90,578 individual-year observations from 11,194 individuals. Asymptotic standard errors are calculated using the score of the likelihood.

The estimated preference parameters are provided in Table 6. Similar to past work in the United States (Bishop, 2008; Kennan and Walker, 2011), I estimate a strong preference

for living in one's location of birth. The estimate implies that individuals are indifferent between spending one year away from their location of birth with an income roughly twice that received in their birth location. Even with such strong attachments to this location, I still find that a large portion of the population actively considers moving each period - 63%. Similarly, I estimate a negative preference for living in an urban area. This is reminiscent of a compensating differential associated with living and working in an urban center (Rosen, 1986). In particular, the average individual requires a roughly 24% higher income to be indifferent between living in an urban area and a rural one.

I estimate that the behavior of more than half of the population (67%) is more consistent with a model of full-information than one of naivety. I am unable, however, to reject the null hypothesis of a unit-mass of fully-informed individuals.²³ I find that individuals dislike extreme amounts of both heat and cold, as measured by degree days above 26°C and below 14°C. To put these preference estimates into perspective, one can study how flow utilities change with respect to changes in the number of degree days. In particular, the marginal flow utility of an additional degree day above 26° C is above zero when the number of degree days is below ~ 120 . Similarly, the marginal impact on flow utilities of an additional degree day below 14° C is positive when the number of degree days is below ~ 260 . The quadratic preference specification over degree days allows for a “bliss-point” interpretation, where this value is defined as the peak of the relevant quadratic.

To put these values into perspective, 17% of the life history data experience annual degree days above 26°C which are above their bliss-point, and 28% above that for degree days below 14°C. If these same individuals experienced the 2050 business-as-usual climate, 49% of them would face a number of annual degree days above 26°C to the right of its bliss point and less than 1% would face annual degree days below 14°C to the right of the same. I explore the consequences of such in the following section.

Similar to past work in developed countries (Kennan and Walker, 2011; Oswald, 2019; Ransom, 2022), I estimate a high migration cost which is increasing in age and the distance between the destination and origin. I also find that these costs increase once one becomes a parent. One way to understand the magnitude of the moving cost is by comparing the marginal utility of income to the intercept of the moving cost equation to translate from utils to dollars.²⁴ This exercise reveals a fixed migration cost²⁵ of \$99,000, or roughly 10

²³I have also estimated two different versions of the model where I assume that the climate-naive expect that future weather draws are drawn from either a stationary or non-stationary degenerate distribution at the average weather over the last thirty years. The estimated preference parameters and the mass of movers are statistically indistinguishable from those presented here. I estimate a smaller mass of the climate-naive (8% and 12% vs 33%), but with much more precision (a standard error of 0.019 and 0.014).

²⁴Appendix B.6 discusses the mapping I use to translate from utils to dollars.

²⁵I am defining the fixed cost, here, simply by the moving intercept. The level shifters, children and age,

years worth of income. This is the full cost, inclusive of all psychological and monetary costs associated with a forced move, in an arbitrary period, to an arbitrary location. Since migration is an endogenous decision, individuals decide to move, many times, because they received a high preference shock (represented by ε) in their chosen destination or a low preference shock in their origin. I incorporate this fact into the moving cost calculation by then subtracting off the expected maximum of $N - 1$ iid type-one extreme value draws.²⁶ This reduces the fixed moving cost to a moving *subsidy* of \$50,000.²⁷ Kennan and Walker (2011) provide another way to contextualize the moving costs by defining them as the difference in expected flow utilities between the origin and the destination, inclusive of moving costs, net of the expected difference in preference shocks, given the decision made. This incorporates that when an individual is willing to leave a relatively desirable location for one that is less so, it must have been that their unobserved preference shock in the destination was sufficiently high, as compared to that in their origin. Calculating moving costs in this way provides for an average *subsidy* of \$50,000. Given the strong preference for living in one's birth location, the calculated moving cost back home is much higher - a cost of \$29,000. I discuss these calculations in more detail in Appendix B.4.

The above calculations highlight the highly heterogeneous nature of migration. In particular, model-observables are not, by themselves, able to fully capture the incentives associated with moving in any given period to any given location. However, combining these observables with the underlying model structure, including the distributional assumptions on the preference shocks, allows one to obtain a more accurate picture of the importance placed on each component of the migration decision process. The model uses the variation described in Section 4, in conjunction with the normalized preference shocks, to provide insights into the relevance of climate on migration decisions, through both incomes and amenities. In the next section, I provide evidence that the estimated model does a good job of predicting migration decisions within Mexico.

Model Fit

To understand how well my estimated model fits the observed data, I use it to simulate each of the decisions of the individuals used in its estimation. Table 7 shows the average migration rates across a series of demographics in the data and across 100 simulations of the model. The model generally slightly overestimates migration rates across the population.

will, in practice, increase the magnitude of this fixed cost.

²⁶ N , here, represents the number of elements in the choice set.

²⁷I am defining a moving subsidy as a negative moving cost. Given that the stochastic shocks can also be interpreted as entering through incomes, another way to view this would be through the lens of a transitory (3-year-long) positive shock to incomes in their destination.

Table 7: Model Fit

	Migration Rate	
	Model	Data
Overall	5.4%	4.7%
Children		
Yes	5.2%	3.4%
No	5.4%	7.2%
Agricultural Worker		
Yes	4.5%	3.5%
No	5.8%	5.3%
Years of Education		
0-5	5.0%	4.2%
6-11	5.3%	3.8%
12+	6.1%	7.5%

Notes: For each category, I calculate the average migration predicted by 100 model simulations of the life trajectories of the individual-year observations used in estimation. The data migration rates are calculated as the average migration rate for each category.

This is primarily because of the retrospective nature of the MMP data. In particular, I only reproduce exactly those decisions in the data, and as such, I do not predict decisions made at points in life later than those observed (the age at which an individual is surveyed). This leads to an overprediction of migration from a higher mass of young individuals than the estimation routine accounts for. Appendix C replicates Table 7 when instead simulating the entire life history of each individual. This leads to a much closer matching of the demographic moments.

Figure 6a plots model predicted versus observed migration rates, broken apart by age and education levels. The dashed curve represents the model-predicted migration rate across ages and education levels; the solid curve represents that from the estimation data. The model does a good job of capturing trends over time and education levels. Perhaps reassuringly, the model only recognizes differences across levels of education through differences in the wage process, and even still, it is able to correctly capture the initial gap in migration rates by young individuals with high levels of education compared to those with lower levels.²⁸ It also correctly predicts both the convergence of migration rates across education levels as

²⁸Some of these individuals are likely moving to college. The model captures this via a higher wage opportunity in all locations.

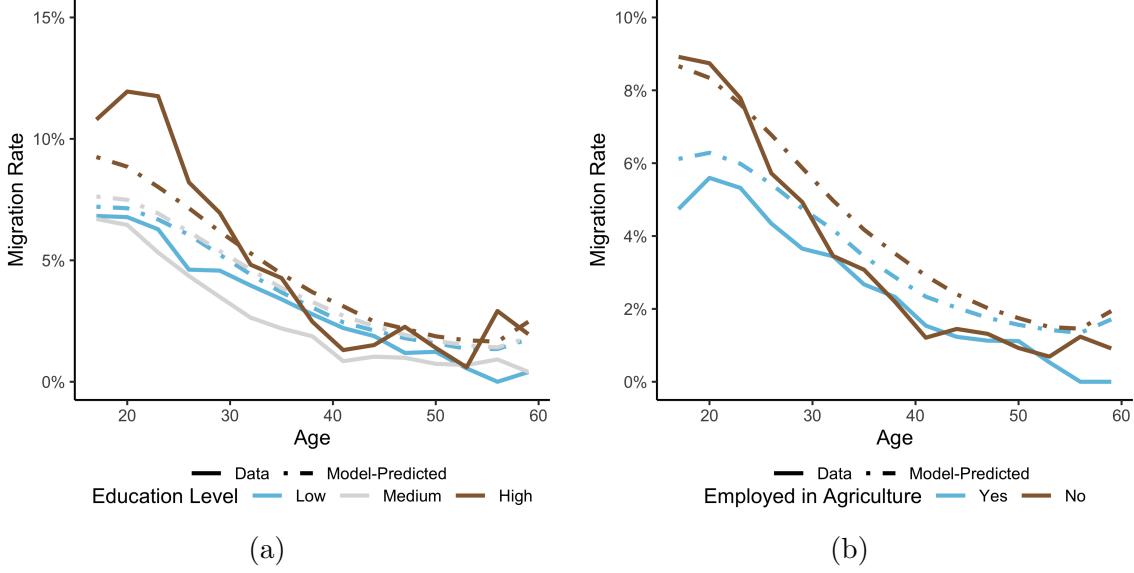


Figure 6: (a) plots the average migration rates in the data against those predicted by the model, broken apart by education levels. Low, mid, and high education refers to less than 6 years, 6–11 years, and 12 or more years of education, respectively. (b) plots the average migration rates in the data against those predicted by the model, broken apart by agricultural employment.

individuals age and the eventual uptick in migration rates at the end of the life cycle.

Figure 6b tells a similar story as Figure 6a, this time broken apart by agricultural employment. As in the setup for education, the model only incorporates an individual's sector of employment through the wage process. It is still able to correctly match many of the differential life-cycle trends across these groups. In particular, it matches the initial uptick in migration at the beginning of the life cycle for agricultural workers, but not for non-agricultural workers. It also correctly predicts the eventual convergence in rates across the two groups and the increase in rates for non-agricultural workers at the end of the life cycle. Appendix C plots the same figures but from simulating decisions across the entire life cycle for all individuals, rather than just those decisions made and observed in the data.

6 Counterfactuals

I use the estimated model to simulate forward the life histories of all Census-enumerated Mexican males through 2080 under a variety of counterfactual scenarios. To do so, I combine data from the 2020 Census count and the MMP life history data. In particular, I use Census data on the number of individuals in each age bin residing in a municipality assigned to each location in 2020, along with the location-specific education and agricultural employment distributions to create an individual-level cross-section. Through 2038, I incorporate new

cohorts of young men into the simulation, where the number and locations of new entrants comes from the age distribution for children in the 2020 census. Appendix E discusses this approach, and the data used, in more detail.

I use the counterfactual simulations to understand the effectiveness of migration in limiting future climate damages. In particular, I design a series of counterfactual exercises which modify the climate system, moving costs, and the ability to migrate. I use these to understand both how the value of the ability to migrate varies with the degree of warming as well as how migration directly affects expected climate damages

The baseline scenario, a business-as-usual climate change projection, is modeled directly through an extension to the framework used in estimation. More specifically, the full solution procedure used in estimation requires solving the model through 2064, to account for the terminal period of those who turn 17 in 2018. Throughout the estimation framework, I assume that the individuals who form fully-informed expectations of the future climate use the previous thirty years of weather data to form their expectations. I use historical data through 2020, and use daily projected weather data from the SSP2-4.5 climate projections for the remaining years. This helps maintain an internally-consistent framework upon which to study counterfactual scenarios.

To understand the difference between migration decisions under business-as-usual and a world without climate change, I use historical weather data from 1950–1979 to construct a stationary climate system with limited climate change. For this scenario, I model that the climate system is governed by a Gaussian distribution with a constant mean and variance, both of which are set to the observed values from 1950–1979. As in estimation, in both the business-as-usual and stationary climate scenarios, I assume that there are both movers and non-movers and those who form fully-informed expectations and those who form naive expectations about future weather draws, with the mass of each of these groups set to the estimated proportions from the MMP data.

Business-as-Usual Warming Increases the Value of Migration

The first set of simulations shed light on how climate change will affect the value that different individuals place on the ability to choose whether or not to migrate throughout their lifetime. To study this question, I use the estimated model to simulate the choices of each Census-enumerated male for the remainder of their lives. When doing so, I track their model-predicted welfare, when they can and cannot choose to move domestically as well as when the climate is characterized by the business-as-usual scenario and the counterfactual scenario of limited warming. Mathematically, the average value of migration under scenario

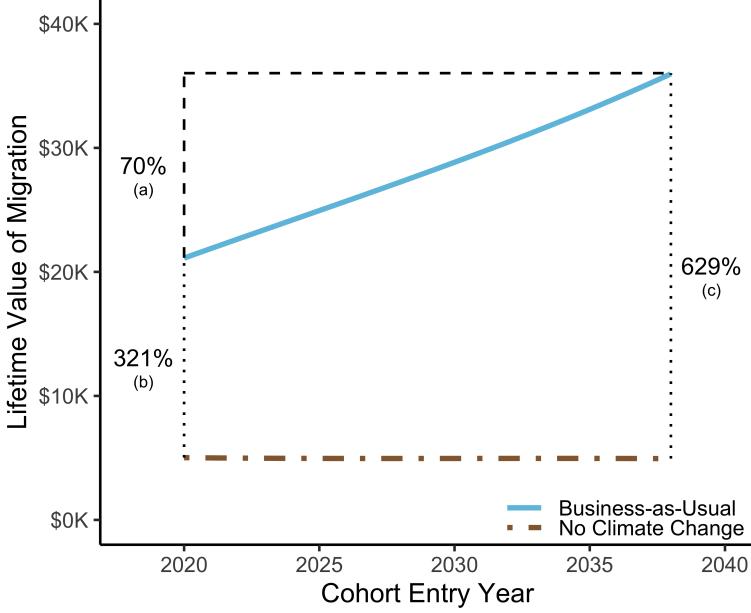


Figure 7: Value of Migration over Time

Notes: Plotted is the average value of migration for 17–19-year-olds over time, with the percentage change in this value from 2020 to 2038. (a) represents the percentage change in this value under the business-as-usual climate scenario from 2020 through 2038. (b) represents the percent difference in this value between the business-as-usual and stationary climate scenarios in 2020. (c) represents the same as (b), but in 2038. All curves are smoothed with a third-order polynomial. The results are averaged over 100 simulations.

s , in the base year (2020), is given by:

$$\frac{1}{n_y} \sum_{i(y)} \left\{ \tilde{V}(\ell_i, \omega_i | s) - \tilde{V}(\ell_i, \omega_i | s, \gamma_1 = \infty) \right\},$$

where $\tilde{V}(\ell_i, \omega_i | s)$ is the observed value of individual i 's lifetime value under scenario s , appropriately scaled to 2024 dollars, and y indexes the year.

Figure 7 plots the results. I find that migration is valuable: the average 17–19-year-old in 2020 values their ability to move domestically at approximately 2-years of annual income. The entering cohorts consistently have values of migration which are greater than those of the last – as panel (a) shows, by 2038 this quantity increases by 70% relative to its 2020 value. Much of both the initial level and the steep change are driven by welfare damages from increased warming in the near future. Panel (b) highlights this finding: projected warming increases the lifetime value of the ability to migrate by just over 4 times, relative to the stationary historical climate scenario. Panel (c) combines (a) and (b) to understand how the value of migration, driven by climate change, changes with time. I find a fairly flat representation for

the value of migration in a no warming scenario in the near future.²⁹ This, combined with the acceleration in values under business-as-usual, are represented in panel (c). The value of migration under projected warming stands to increase the value of migration by over 7 times, relative to a world without climate change.

The first two columns of Table 8 decompose this finding across a variety of spatial and demographic dimensions. I find that the value of migration, for the average 17–19-year-old in 2020, across Mexico is \$21,139, or roughly 2 years worth of income, under the business-as-usual climate scenario – about 4 times higher than that under the limited climate change scenario. These two columns additionally highlight the progressive nature of migration as an adapting tool for climate damages: it is around 60% more valuable to those with the lowest relative levels of education than those with the highest.³⁰ The same can be said for individuals employed in agriculture, relative to those employed in other sectors.

The upper panel of Table 8 presents the spatial heterogeneity value throughout Mexico. I find that individuals living in the temperate parts of the country face a relatively stable value of migration across the different climate scenarios, while those in the warm regions face an almost 15-fold increase in their value under the business-as-usual scenario relative to the no climate change scenario. Decomposing these values across the demographic distribution of the population shows that low-skilled and agricultural workers stand to gain the most from the ability to move under a business-as-usual pathway, while the gains from migration are spread more evenly across the population under the stationary climate scenario.

I note that dynamic considerations are vital to analyzing the value of migration in this context. To see why this is true, I compare the estimated value of migration under a business-as-usual climate scenario with and without dynamics. In particular, I use the static logit framework presented in Equation 1 to simulate the decisions of the male population of Mexico through 2080. I estimate a lifetime value of domestic migration which is two orders of magnitude lower than that given by the dynamic framework: \$520. This difference is driven by a lower static-predicted rate of migration. As displayed in Figure C5, the dynamic model predicts migration rates for young men in Mexico as being between 8% and 10% from 2020 through 2038. This is compared to the migration rate predicted by the static framework: between 3% and 5% over the same timeframe, as seen in Figure C6. These differences arise from the fact that the dynamic model is able to capture the long-term benefits from migration, accrued over one’s lifetime, whereas the static model only captures the value from migration

²⁹This, mainly, is a model-driven result, as the only non-stationary state distribution is the joint climate distribution.

³⁰The differential wage premium of urban employment for those with high education is well established in the literature (c.f., [Gould, 2007](#)). I purposely abstract from such effects to aid in comparing how the relative values of migration will change with the climate across groups such as education.

Table 8: Heterogenous Values of Migration

	Business-as-Usual Climate		
	New workers today	Children born today	Limited Climate Change
Average	21,139	36,026	5,023
Climate			
Dry	15,059	20,909	4,323
Temperate	5,812	5,415	5,292
Warm	75,687	130,389	5,197
Agricultural Worker			
Yes	35,699	62,206	4,998
No	19,210	31,959	5,026
Years of Education			
0–5	31,285	54,802	5,148
6–11	24,719	42,469	5,049
12+	19,304	32,458	5,007

Note: These are the average results of 100 simulations of the model. Values are for 17–19 year olds and are measured in 2024 dollars. The value of migration is defined as the difference between average lifetime welfare for an individual who can migrate and one who cannot. Business-as-usual refers to the SSP2-4.5 climate projections, limited climate change refers to a climate system centered around 1950–1979 averages. Children born today and new workers today refer to individuals who were less than 2-years-old and between 17–19-years-old during the 2020 Census count, respectively.

for a single year when an individual is deciding whether or not to move.

What is driving the large differences in welfare across the two climate pathways? The model incorporates two primary channels for climate to affect welfare: directly through the flow utility of temperature as an amenity and indirectly through local wages. I use the model to understand how much of the reported heterogeneous values are driven by the separate channels. To do so, I individually turn on and off the wage and amenity channels along with the ability to migrate. For each of these counterfactuals, I resimulate behavior under the business-as-usual climate scenario and present the results in Table C3 in Appendix C. The first column (Only Wages) looks at the value of migration when the amenity value of weather is turned off, i.e., by setting θ_2 – θ_6 equal to zero. I find that the resulting option value is much more homogeneous across space and quite similar to those reported for the no climate change scenario. The second column (Only Amenities) instead turns off the wage impact in flow utilities, i.e., by setting θ_1 equal to 0. I find values that are much closer to those in the business-as-usual scenario. This highlights the high premium that individuals place on the amenities of their location and reaffirms past findings in the United States (e.g., Kennan

and Walker, 2011; Ransom, 2022). Moreover, I find that it is through amenities, rather than wages, that the heterogeneous value of migration is most noticeable. This is likely driven by past sorting behavior of non-agricultural workers and those with higher levels of education to areas less affected by climate change.

Stylized Framework to Incorporate Congestion Impacts on Wages

Given that migration is modeled in a partial equilibrium framework, the above analysis discounts any general equilibrium effects that may lead to decreases in wages in desirable locations. For example, it likely will be the case that individuals are cognizant of the fact that local wages will respond to future changes in population. In other words, the places that are desirable today will attract higher portions of the population in the future, which will drive future wages down in these locations, making them less desirable today. This would work to decrease the value of migration.

It is not computationally feasible to allow for these dynamics, in their base form, in the dynamic model. I design a simple and stylized general equilibrium exercise to understand how serious of a concern this may be. To establish intuition, imagine a world in which individuals begin in one of two locations, a and b . For simplicity suppose that individuals all agree that a is the more desirable location. However, individuals also understand that wages respond to the proportion of the population that is in a . In particular, take some individual, i , who is deciding whether to move to a . i 's immediate reaction is that they would like to move to a , however after pondering the decision they recognize that many other people will likely also move to a , which will drive wages down from w to $w' < w$. They are familiar with much of the literature on the topic, and combine the expected change in populations with Borjas (2003)'s estimated elasticity of wages to labor supply to update what they expect their wages to be in location a : w' . But then they recognize that everybody else is likely performing the same set of calculations, and as such some fraction of the population will actually decide to stay in b . This means that wages in a will not fall all the way to w' , but rather to $w'' > w'$. This process can be iterated on until $w^{(k)} \approx w^{(k+1)}$.

Convergence of this framework is similar in spirit to the iterative logic of level-k reasoning theory, as used in behavioral economics to explain why equilibrium theory fails to accurately predict outcomes in many games (Camerer et al., 2004). In each iteration, some mass of the individuals who were marginal to moving to a given location decide to instead either stay in their current location or move to another location. If all of these marginal individuals decide to stay in their current location, and the set of marginal individuals is decreasing across iterations, then convergence defined against changes in population counts in each location across iterations, is guaranteed. However, given that most individuals who decide to

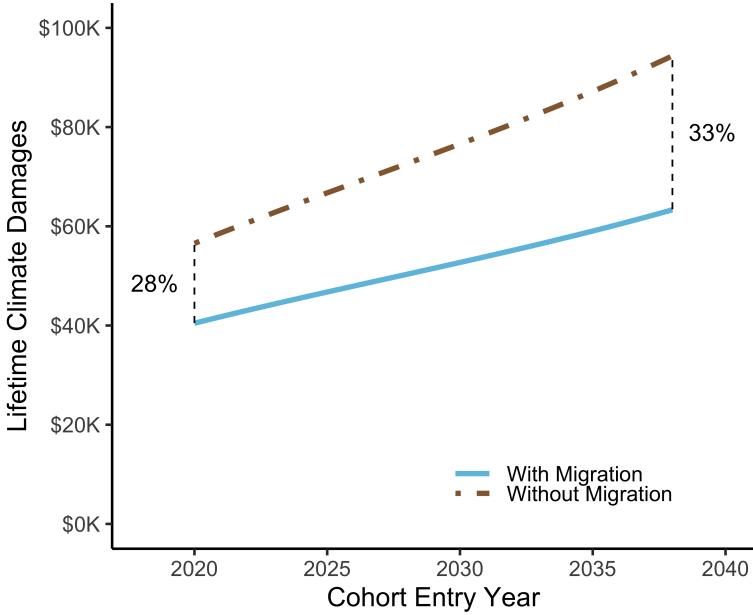


Figure 8: Migration and climate damages over time

Notes: This figure plots model-predicted climate damages, for 17–19-year-olds over time. The solid curve represents damages in a world with migration; the dashed curve represents damages in a world without. The results are averaged over 100 simulations.

move are doing so not only because of the specific destination but also because of the origin, convergence is not guaranteed.

More concretely, I simulate behavior under the baseline scenario and track the percentage change in population levels in each of the locations from 2023 and 2038. I modify the wage equation to respond to these changes, using an elasticity of -0.3 (Borjas, 2003), linearly from 2023 through 2038. This marks the end of the first iteration. In the second iteration, I resimulate behavior with the updated wage equation. I again keep track of population changes to arrive at a further updated wage equation to be used in the third iteration and so on. In practice, it does not take many iterations for the algorithm to converge: by iteration 12 the maximum difference in relative population changes across iterations is below 1%. The final column of Table C3 shows the updated values of migration after accounting for congestion impacts.³¹ I find that the average option value across Mexico decreases by around 10%. Appendix E.4 provides more details on the implementation of this exercise.

Migration Limits Climate Damages

The above analysis has highlighted that warming in the near future will lead to a stark increase in the value that individuals place on their ability to move domestically. I now ask

³¹There is a maximum difference of less than 0.09%, across locations, between the 16th and 17th iterations.

how effective we should expect migration to be in limiting overall climate damages. I use Figure 8 to study this question. The figure plots average lifetime climate damages for the cohorts of 17–19-year-olds, from 2020 through 2038. Average expected climate damages are calculated as:

$$\frac{1}{n_y} \sum_{i(y)} \left\{ \tilde{V}(\ell_i, \omega_i | \text{no cc}) - \tilde{V}(\ell_i, \omega_i | \text{BAU}) \right\},$$

where $\tilde{V}(\cdot | \text{no cc})$ represents the lifetime value for an individual, in 2024 dollars, when they are in a world with limited climate change, and $\tilde{V}(\cdot | \text{BAU})$ represents the same for an individual in a world governed by the SSP2-4.5 climate projections. The solid curve represents the trajectory of these damages through the 2038 cohort, in the baseline scenario with estimated migration frictions. The dashed curve represents the same trajectory in the world where migration is completely restricted.

I find that climate damages are increasing over time, both under the status quo and under a world without migration. Taken together with the relative importance of amenities and wages in driving the differential values of migration, this is driven by the stark shift in the distribution of extreme heat events across Mexico, as signified in Figure 2. Moreover, the figure shows that it is not only the level of climate damages which are increasing with time, but also the gap between damages in a world with migration and one without: average climate damages experienced by a 17–19-year-old in 2020 are 28% lower because of their ability to migrate within Mexico. For an individual who was less than 2-years-old in 2020, this migration-induced reduction in climate damages is projected to increase to 33%.

The first column of Table 9 directly measures the importance of migration in limiting climate damages for young men in 2020. The spatial heterogeneity in the decrease in damages from migration speaks directly to the mechanism through which migration limits climate damages. In particular, individuals living in the temperate regions of the country face a 4% reduction in their expected lifetime climate damages due to their ability to migrate. This is in stark contrast to the 35% reduction in damages faced by those living in the warmest regions of the country, who are able to move to the temperate regions of the country to escape damages.

The second column studies the same reduction in climate damages from the ability to migrate, but for the 2038 cohort. I find that the decrease in damages from migration is expected to decline for the individuals living in temperate regions of the country, whereas it is expected to increase very slightly for those in the dry regions and to increase by more than 4 percentage points for those in the warm regions of the country. The differences in the change in the reduction in damages from migration across the country comes from two facts. First, even the temperate regions of the country are expected to face increased climate

Table 9: Migration Reduces Expected Climate Damages

Age in 2020:	% ↓ in Damages from Migration	
	17–19-years-old	<2-years-old
Average	28%	33%
Climate		
Dry	27%	27%
Temperate	3%	1%
Warm	35%	40%
Agricultural Worker		
Yes	32%	37%
No	28%	32%
Years of Education		
0–5	32%	36%
6–11	30%	34%
12+	28%	32%

Note: These are the average results of 100 simulations of the model. Damages are measured in lifetime values. The percentage decrease in damages from migration is the average difference in climate damages from a world without migration to one with migration.

damages in the future. Second, the other locations in their choice set are also worsening, with the dry regions warming even faster than the temperate one. Taken together, the value that an individual who already lives in a temperate location will place on their ability to move within Mexico will attenuate over time with the degree of warming.

The Value of Climate Information to the Climate-Naive

The previous two sections have highlighted that migration will be an increasingly valuable form of adaptation to climate change. That said, given the dynamic interplays present in the decision to move, its value and effectiveness as an adapting strategy will depend on our understanding of future warming. Here, I use the estimated model to uncover the importance of understanding the climate system when deciding whether to dynamically adapt to climate change. I then present evidence that policy can be used to dampen the welfare losses due to naivety on the climate transition across the population.³²

In the first two columns of Table 10, I highlight the value to the climate-naive of becoming fully-informed on the climate system through the change in their expected lifetime climate

³²In Appendix B.7, I derive a closed form, recursive, representation of the alternative-specific value function for the climate-naive individuals, which allows me to calculate the true welfare for the climate-naive.

Table 10: Value of Climate Information to the Climate-Naive

	% ↓ in Climate Damages		Value of Climate Information in 2020		
	New workers today	Children born today	Dollar Value	% ↓ from \$1,100 Dynamic Subs.	% ↓ from \$1,100 Static Subs.
Average	2.2%	1.2%	1,099	18.7%	7.8%
Climate					
Dry	0.4%	0.6%	544	29.7%	12.8%
Temperate	1.3%	2.2%	298	49.0%	21.4%
Warm	3.2%	1.2%	4,300	10.4%	4.1%
Agricultural Worker					
Yes	2.9%	1.0%	2,019	13.7%	5.3%
No	2.1%	1.3%	976	20.2%	8.6%
Years of Education					
0–5	2.7%	1.1%	1,700	13.7%	6.0%
6–11	2.5%	1.2%	1,313	18.8%	7.0%
12+	2.1%	1.3%	987	19.2%	8.5%

Note: These are the average results of 100 simulations of the model. Values are for 17-19 year olds, measured in 2024 dollars. The percent decrease in climate damages (as in Table 9) represents that for the climate-naive individuals from becoming fully-informed. The value of climate information refers to the difference in average lifetime welfare between the population of fully-informed and climate-naive individuals, in 2020. The decrease in the value of climate information from a dynamic and static subsidy refers to the decrease of the dollar value of climate information from a one-time \$1,100 subsidy to be used at the first time an individual moves and a take-it-or-leave-it offer, both in 2020.

damages. In 2020, I find that obtaining and using a full understanding of the climate system would lead to an expected decrease in lifetime climate damages of just over 2%. This corresponds to a dollar value of \$1,099 over one's lifetime (column 3). By 2038, the expected reduction in lifetime climate damages is projected to fall to just over 1%. As Figure 2a highlights, this is because of the flattening of the rate of warming around 2040 under the business-as-usual climate scenario, which causes a narrowing in the gap between the decisions made by the climate-naive and the fully-informed, as seen in Figure C5.

Where do the increased climate damages due to naivety on the climate system come from? Compared to the decisions one would make with full information, I find that the climate-naive are over three-times as likely to mistakenly choose to stay in their current location as they are to mistakenly decide to move. Moreover, across the simulated time-frame, mistakenly deciding to remain in one's location induces average losses to lifetime welfare of \$9,006. The average lifetime-losses from making a mistake during a move are less than half of this: \$3,796. Taken together, this implies that subsidizing migration has the potential to be one way of reducing the heightened expected climate damages borne by the climate-naive.

The fourth and fifth columns of Table 10 highlight the effectiveness of subsidizing migration at reducing the welfare losses of the climate naive. I present the results from two potential

policy options. The first is a dynamic subsidy which can be used at the first time an individual decides to move. The second is a static policy which is only available in the policy year. Both policies correspond to a subsidy roughly worth the average lifetime value to the climate-naive of becoming fully-informed. Importantly, I highlight that it is possible that subsidizing migration could actually *increase* the welfare losses to the climate-naive if their losses outweighed their gains from their policy-induced moves. I find that this is not the case. Moreover, I find that dynamic policies are substantially more valuable than one-time offers.

The difference between the welfare implications of the two policy options is driven by the dynamic policy's influence on the value that each individual places on the ability to move at a reduced cost in the future. This increases the relative value of the dynamic policy through two channels. The first is that the dynamic policy will affect the welfare of everyone who considers the option of migration. Notice that even those who would not move in the policy year, with or without the policy, will receive positive value from the knowledge that they will have the option to move at a reduced cost in the future. The second is that the dynamic policy will generate a strictly higher value-per-mover than a static policy. In other words, the dynamic policy positively selects from the mass of the population marginal to moving under a static policy.

Figure 9 provides intuition for why dynamically-available policies will positively select from the set of individuals marginal to a static reduction in moving costs. I plot a conceptual distribution taken across the population. m , here, represents the expected cutoff of the maximum over the preference shocks, in the absence of policy, for which individuals would decide to move. Individuals to the left of m are inframarginal to the policy, and therefore receive a transfer of s from any offered subsidy. The static subsidy of s shifts the portion of the population who decide to move to the right, as represented by m' . m' is determined by an individual's indifference condition: this represents the individual who receives \$0 from moving with the subsidy, that is, they receive a welfare-loss of exactly s from moving without the policy. w represents the option value of future moves with a subsidy of s . Notice that under the dynamic policy, individuals would not move today if their value of moving was below w , as that is their value of withholding the subsidy into the future. The average value of the different subsidies to movers is the weighted average of these values across the population who decides to move, with the weights given by the mass of the population with that particular value for the policy. The fact that the dynamic policy has a positive option value is what drives the positive selection results I highlight.

Keeping with Figure 9, climate change will cause two simultaneous changes to the above figure. First, m and m' will both shift to the right. This is represented by the increasing migration rate due to warming. Second, w will shift closer to m and away from m' , as

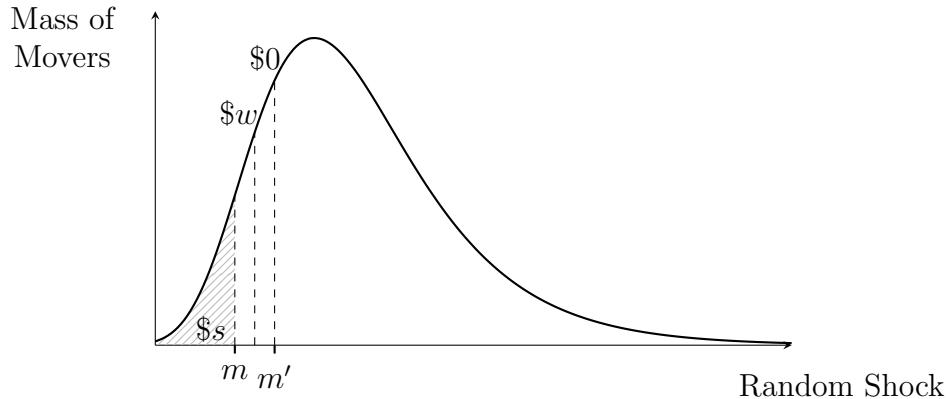


Figure 9: Dynamic and Static Policy

Notes: m is the cutoff for inframarginal movers, s is the size of the subsidy, w is the size of the option value for future moving opportunities with a subsidy of s . The shaded area is the portion of the population inframarginal to any policy.

represented by the increasing option value of future moves at a reduced cost. This highlights that as the climate becomes increasingly hostile, individuals will increasingly value the option to choose *when* to invoke the decision to move at a reduced cost. Figure C8 highlights this finding across time, from warming.

7 Conclusion

This paper has begun to understand the dynamic relationship between climate change and individual migration decisions. In particular, the value that individuals place on the option to migrate is both substantial and increasing in the degree of warming. Put another way, as time progresses individuals will be increasingly exposed to extreme heat from the seemingly small shifts in average temperature. This fact will lead many to seek out less affected areas, even within their own country.

Our ability to adapt our way through climate change relies heavily on our ability to correctly anticipate future warming. As such I use the estimated model to understand the value of correct information about the climate system. To do so, I exploit variation in individual decisions resulting from variation in rates of warming across Mexico to identify the proportion of the population forming fully-informed and naive expectations on the climate transition. I estimate that a non-negligible fraction of the population makes decisions that are more closely consistent with a framework of naive expectations than one of full-information. I find that such individuals would stand to gain an increase of 5% to their value of migration, or 2% fewer climate damages, over the course of their lifetime from obtaining information on

the climate transition system to be used when forming expectations over the climate to aid in adaptation. The average masks a high degree of heterogeneity: those living in the warm regions of the country and those who have been historically marginalized stand to gain the most.

I find that much of the value of correct information on the climate system comes from correcting otherwise lower predicted rates of migration by the climate-naive. As such, one way to close this gap is to subsidize migration. I find that one-time reductions to moving costs at the level of the internality, taken when an individual first decides to move, reduce the loss of welfare to the climate-naive by 19%, whereas take-it-or-leave-it offers reduce the same by 8%. Given these large differences, I use the estimated model to highlight the selection effects that are caused by the different policies. In particular, dynamic reductions to moving costs generate option value from the ability to use them in the future, which reduces the contemporaneous set of individuals marginal to the policy. This value is increasing, both in the likelihood of one deciding to move in the future, and in the value associated with such a move. Warming will cause both of these factors to increase simultaneously.

I highlight two potential avenues for future work. First, the analysis provided in this paper assumes that the only way that one is able to adapt to climate change is through migration. Practically, this means that the estimated valuations of migration in limiting the damages of climate change will be biased upwards because some individuals who the model predicts to migrate would instead prefer to engage in in-situ forms of adaptation. That said, since I abstract away from these other forms of adapting behavior, I am underestimating the overall value of adaptation, generally defined, in limiting the damages of climate change. Future work could study the role and limits of in-situ adaptation to climate change. One way to do this would be to slightly modify the framework I design for this paper through allowing for a joint decision to migrate and switch occupations or to invest in additional education.

Second, future work could study the value of climate information in a setting with a higher degree of variation in the rates of warming across space to better identify the fraction of the population associated with different belief structures. Moreover, in such a setting future work may be able to identify variation in the mass of each belief type across demographics, personal experience with extreme weather, and other observable characteristics which can be used to better target policy to aid in individual adaptation decisions.

References

- AGUIRREGABIRIA, V. AND J. JEON (2020): “Firms’ Beliefs and Learning: Models, Identification, and Empirical Evidence,” *Review of Industrial Organization*, 56, 203–235.

ARAGÓN, F. M., F. OTEIZA, AND J. P. RUD (2021): “Climate Change and Agriculture: Subsistence Farmers’ Response to Extreme Heat,” *American Economic Journal: Economic Policy*, 13, 1–35.

AUFFHAMMER, M., S. M. HSIANGY, W. SCHLENKER, AND A. SOBELZ (2013): “Using weather data and climate model output in economic analyses of climate change,” *Review of Environmental Economics and Policy*, 7, 181–198.

BASSETTI, P. AND M. E. WESTGATE (1993): “Senescence and Receptivity of Maize Silks,” *Crop Science*, 33.

BEVERLY, E. J. (2023): “Using climate to model ancient human migration,” *Science*, 381, 605–606.

BILAL, A. AND E. ROSSI-HANSBERG (2023): “Anticipating Climate Change Across the United States,” .

BISHOP, K. C. (2008): “A Dynamic Model of Location Choice and Hedonic Valuation,” .

BOHRA-MISHRA, P., M. OPPENHEIMER, R. CAI, S. FENG, AND R. LICKER (2017): “Climate variability and migration in the Philippines,” *Population and Environment*, 38, 286–308.

BORJAS, G. J. (2003): “The Labor Demand Curve is Downward Sloping: Reexamining the Impact of Immigration on the Labor Market,” *The Quarterly Journal of Economics*, 118, 1335–1374.

BROWNSTONE, D. AND K. TRAIN (1998): “Forecasting new product penetration with flexible substitution patterns,” *Journal of Econometrics*, 89, 109–129.

BURKE, M., J. DYKEMA, D. B. LOBELL, E. MIGUEL, AND S. SATYANATH (2015): “Incorporating Climate Uncertainty into Estimates of Climate Change Impacts,” *The Review of Economics and Statistics*, 97, 461–471.

BURKE, M. AND K. EMERICK (2016): “Adaptation to Climate Change: Evidence from US Agriculture,” *American Economic Journal: Economic Policy*, 8, 106–140.

BUSSE, M. R., C. R. KNITTEL, AND F. ZETTELMEYER (2013): “Are Consumers Myopic? Evidence from New and Used Car Purchases,” *American Economic Review*, 103.

CAMERER, C. F., T.-H. HO, AND J.-K. CHONG (2004): “A Cognitive Hierarchy Model of Games,” *The Quarterly Journal of Economics*, 119, 861–898.

- CARLETON, T. A. AND S. M. HSIANG (2016): “Social and economic impacts of climate,” *Science*, 353.
- CARTO, S. L., A. J. WEAVER, R. HETHERINGTON, Y. LAM, AND E. C. WIEBE (2009): “Out of Africa and into an ice age: on the role of global climate change in the late Pleistocene migration of early modern humans out of Africa,” *Journal of Human Evolution*, 56, 139–151.
- CATTANEO, C., M. BEINE, C. J. FRÖHLICH, D. KNIVETON, I. MARTINEZ-ZARZOSO, M. MASTRORILLO, K. MILLOCK, E. PIGUET, AND B. SCHRAVEN (2019): “Human migration in the era of climate change,” *Review of Environmental Economics and Policy*, 13, 189–206.
- CATTANEO, C. AND G. PERI (2016): “The migration response to increasing temperatures,” *Journal of Development Economics*, 122, 127–146.
- CHIPMAN, J. S. (1960): “The Foundations of Utility,” *Econometrica*, 28, 193–224.
- CRUZ, J.-L. AND E. ROSSI-HANSBERG (2023): “The Economic Geography of Global Warming,” *The Review of Economic Studies*, 91, 899–939.
- DE DIOS ORTUZAR, J. (1983): “Nested logit models for mixed-mode travel in urban corridors,” *Transportation Research Part A: General*, 17, 283–299.
- DELL, M., B. F. JONES, AND B. A. OLKEN (2014): “What do we learn from the weather? The new climate-economy literature,” *Journal of Economic Literature*, 52, 740–798.
- DERYUGINA, T., L. KAWANO, AND S. LEVITT (2018): “The Economic Impact of Hurricane Katrina on Its Victims: Evidence from Individual Tax Returns,” *American Economic Journal: Applied Economics*, 10, 202–33.
- DESCHÈNES, O. AND M. GREENSTONE (2007): “The Economic Impacts of Climate Change: Evidence from Agricultural Output and Random Fluctuations in Weather,” *American Economic Review*, 97, 354–385.
- DESMET, K. AND E. ROSSI-HANSBERG (2015): “On the spatial economic impact of global warming,” *Journal of Urban Economics*, 88, 16–37.
- FENG, S., A. B. KRUEGER, AND M. OPPENHEIMER (2010): “Linkages among climate change, crop yields and Mexico-US cross-border migration,” *Proceedings of the National Academy of Sciences of the United States of America*, 107, 14257–14262.

GARCÍA, E. (2004): *Modificaciones al Sistema de Clasificación Climática de Köppen (Para Adaptarlo a las Condiciones de la República Mexicana)*, no. 6 in Serie Libros, Instituto de Geografía, Universidad Nacional Autónoma de México, quinta edición: corregida y aumentada ed.

GOULD, E. D. (2007): “Cities, Workers, and Wages: A Structural Analysis of the Urban Wage Premium,” *The Review of Economic Studies*, 74, 477–506.

GOWRISANKARAN, G., A. LANGER, AND M. REGUANT (2024): “Energy Transitions in Regulated Markets,” Working Paper 32088, National Bureau of Economic Research.

HECKMAN, J. AND B. SINGER (1984): “A Method for Minimizing the Impact of Distributional Assumptions in Econometric Models for Duration Data,” *Econometrica*, 52, 271–320.

HERRERO, M. P. AND R. R. JOHNSON (1980): “High Temperature Stress and Pollen Viability of Maize,” *Crop Science*, 20, 796–800.

HOTZ, V. J. AND R. A. MILLER (1993): “Conditional Choice Probabilities and the Estimation of Dynamic Models,” *The Review of Economic Studies*, 60, 497–529.

HSIANG, S. AND R. E. KOPP (2018): “An economist’s guide to climate change Science,” *Journal of Economic Perspectives*, 32, 3–32.

IFRACH, B. AND G. Y. WEINTRAUB (2016): “A Framework for Dynamic Oligopoly in Concentrated Industries,” *The Review of Economic Studies*, 84, 1106–1150.

IPCC (2022): *Climate Change 2022: Impacts, Adaptation and Vulnerability*, Summary for Policymakers, Cambridge, UK and New York, USA: Cambridge University Press.

JESSOE, K., D. T. MANNING, AND J. E. TAYLOR (2018): “Climate Change and Labour Allocation in Rural Mexico: Evidence from Annual Fluctuations in Weather,” *Economic Journal*, 128, 230–261.

KALA, N. (2017): “Learning, adaptation, and climate uncertainty: Evidence from Indian agriculture,” .

KENNAN, J. (2008): “Logit Switching Costs,” Accessed: 2024-01-05.

KENNAN, J. AND J. WALKER (2011): “The Effect of Expected Income on Individual Migration Decisions,” *Econometrica*, 79, 211–251.

KOŞAR, G., T. RANSOM, AND W. VAN DER KLAUW (2022): “Understanding migration aversion using elicited counterfactual choice probabilities,” *Journal of Econometrics*, 231, 123–147.

KÖPPEN, W. (1884): “Die Wärmezonen der Erde, nach der Dauer der heissen, gemässigten und kalten Zeit und nach der Wirkung der Wärme auf die organische Welt betrachtet,” *Meteorologische Zeitschrift*, 1, 215–226.

LEISEROWITZ, A., J. CARMAN, N. BUTTERMORE, L. NEYENS, S. ROSENTHAL, J. MARLON, J. SCHNEIDER, AND K. MULCAHY (2022): “International Public Opinion on Climate Change, 2022,” .

LEMOINE, D. (2024): “Estimating the Consequences of Climate Change from Variation in Weather,” Working Paper 25008, National Bureau of Economic Research.

LESSEM, R. (2018): “Mexico-U.S. immigration: Effects of wages and border enforcement,” *Review of Economic Studies*, 85, 2353–2388.

LIU, M., Y. SHAMDASANI, AND V. TARAZ (2023): “Climate Change and Labor Reallocation: Evidence from Six Decades of the Indian Census,” *American Economic Journal: Economic Policy*, 15, 395–423.

LIVNEH, B., T. J. BOHN, D. W. PIERCE, F. MUÑOZ-ARRIOLA, B. NIJSSEN, R. VOSE, D. R. CAYAN, AND L. BREKKE (2015): “A spatially comprehensive, hydrometeorological data set for Mexico, the U.S., and Southern Canada 1950–2013,” *Scientific Data*, 2.

MCFADDEN, D. (1973): “Conditional Logit Analysis of Qualitative Choice Behavior,” *Frontiers in Econometrics*, 105–142.

MCKEE, T. B., N. J. DOESKEN, J. KLEIST, ET AL. (1993): “The relationship of drought frequency and duration to time scales,” in *Proceedings of the 8th Conference on Applied Climatology*, California, vol. 17, 179–183.

MENDELSOHN, R. (2019): “An Examination of Recent Revealed Preference Valuation Methods and Results,” *Review of Environmental Economics and Policy*, 13, 267–282.

MILLER, S., K. CHUA, J. COGGINS, AND H. MOHTADI (2021): “Heat Waves, Climate Change, and Economic Output,” *Journal of the European Economic Association*, 19, 2658–2694.

MISSIRIAN, A. AND W. SCHLENKER (2017): “Asylum applications respond to temperature fluctuations,” *Science*, 358, 1610–1614.

MONTOISON, A., P. PASCAL, AND L. SALOMON (2020): “NOMAD.jl: A Julia interface for the constrained blackbox solver NOMAD,” <https://github.com/bbopt/NOMAD.jl>.

MORA, C., B. DOUSSET, I. R. CALDWELL, F. E. POWELL, R. C. GERONIMO, C. R. BIELECKI, C. W. W. COUNSELL, B. S. DIETRICH, E. T. JOHNSTON, L. V. LOUIS, M. P. LUCAS, M. M. MCKENZIE, A. G. SHEA, H. TSENG, T. W. GIAMBELLUCA, L. R. LEON, E. HAWKINS, AND C. TRAUERNICHT (2017): “Global risk of deadly heat,” *Nature Climate Change*, 7, 501–506.

NAKAMURA, E., J. SIGURDSSON, AND J. STEINSSON (2021): “The Gift of Moving: Intergenerational Consequences of a Mobility Shock,” *The Review of Economic Studies*, 89, 1557–1592.

OSWALD, F. (2019): “The effect of homeownership on the option value of regional migration,” *Quantitative Economics*, 10, 1453–1493.

RANSOM, T. (2022): “Labor Market Frictions and Moving Costs of the Employed and Unemployed,” *Journal of Human Resources*, 57, S137–S166.

RIAHI, K., D. P. VAN VUUREN, E. KRIEGLER, J. EDMONDS, B. C. O’NEILL, S. FUJIMORI, N. BAUER, K. CALVIN, R. DELLINK, O. FRICKO, W. LUTZ, A. POPP, J. C. CUARESMA, S. KC, M. LEIMBACH, L. JIANG, T. KRAM, S. RAO, J. EMMERLING, K. EBI, T. HASEGAWA, P. HAVLIK, F. HUMPPENÖDER, L. A. DA SILVA, S. SMITH, E. STEHFEST, V. BOSETTI, J. EOM, D. GERAAT, T. MASUI, J. ROGELJ, J. STREFLER, L. DROUET, V. KREY, G. LUERER, M. HARMSEN, K. TAKAHASHI, L. BAUMSTARK, J. C. DOELMAN, M. KAINUMA, Z. KLIMONT, G. MARANGONI, H. LOTZE-CAMPEN, M. OBERSTEINER, A. TABEAU, AND M. TAVONI (2017): “The Shared Socioeconomic Pathways and their energy, land use, and greenhouse gas emissions implications: An overview,” *Global Environmental Change*, 42, 153–168.

ROSEN, S. (1986): “Chapter 12 The theory of equalizing differences,” Elsevier, vol. 1 of *Handbook of Labor Economics*, 641–692.

RUDIK, I., G. LYN, W. TAN, AND A. ORTIZ-BOBEA (2022): “The Economic Effects of Climate Change in Dynamic Spatial Equilibrium,” .

RUST, J. (1987): “Optimal Replacement of GMC Bus Engines: An Empirical Model of Harold Zurcher,” *Econometrica*, 55, 999–1033.

——— (1988): “Maximum likelihood estimation of discrete control processes,” *SIAM Journal of Control and Optimization*, 26.

- SARVIMÄKI, M., R. UUSITALO, AND M. JÄNTTI (2022): “Habit Formation and the Misallocation of Labor: Evidence from Forced Migrations,” *Journal of the European Economic Association*, 20, 2497–2539.
- SCHLENKER, W. AND M. J. ROBERTS (2009): “Nonlinear temperature effects indicate severe damages to U.S. crop yields under climate change,” *Proceedings of the National Academy of Sciences of the United States of America*, 106, 15594–15598.
- SEO, S. N. AND R. MENDELSOHN (2008): “An analysis of crop choice: Adapting to climate change in South American farms,” *Ecological Economics*, 67, 109–116.
- SHERWOOD, S. C. AND M. HUBER (2010): “An adaptability limit to climate change due to heat stress,” *Proceedings of the National Academy of Sciences*, 107, 9552–9555.
- SIAPI (2022): *Panorama Agroalimentario 2022*, Ciudad de México: Servicio de Información Agroalimentaria y Pesquera, edición 2022 ed.
- SIMON, H. A. (1955): “A Behavioral Model of Rational Choice,” *The Quarterly journal of economics*, 69, 99–118.
- THOM, K. (2010): “Repeated Circular Migration: Theory and Evidence from Undocumented Migrants,” .
- THORNTON, M., R. SHRESTHA, Y. WEI, P. THORNTON, S.-C. KAO, AND B. WILSON (2022): “Daymet: Daily Surface Weather Data on a 1-km Grid for North America, Version 4 R1,” .
- THRASHER, B., W. WANG, A. MICHAELIS, F. MELTON, T. LEE, AND R. NEMANI (2022): “NASA Global Daily Downscaled Projections, CMIP6,” *Scientific Data*, 9.
- VICENTE-SERRANO, S. M., S. BEGUERÍA, AND J. I. LÓPEZ-MORENO (2010): “A Multiscalar Drought Index Sensitive to Global Warming: The Standardized Precipitation Evapotranspiration Index,” *Journal of Climate*, 23, 1696 – 1718.
- WILLIS, K. J., K. D. BENNETT, D. WALKER, AND P. FORSTER (2004): “Ice Ages and the Mitochondrial DNA Chronology of Human Dispersals: A Review,” *Philosophical Transactions of the Royal Society of London. Series B: Biological Sciences*, 359, 255–264.
- WILSON, L. AND W. BARNETT (1983): “Degree-days: An Aid in Crop and Pest Management,” *California Agriculture*, 37, 4–7.

WOLPIN, K. I. (1984): “An Estimable Dynamic Stochastic Model of Fertility and Child Mortality,” *Journal of Political Economy*, 92, 852–874.

ZAPPALÀ, G. (2024): “Adapting to climate change accounting for individual beliefs,” *Journal of Development Economics*, 169, 103289.

A Data

Table A1: Data Sources

Data Source	Purpose	Years
<i>Mexican Migration Project (MMP)</i>	Estimation (Life-histories)	1950–2019
<i>ENIGH, ENE surveys</i>	Estimation (Income)	1984–2019
<i>Linveh et al. (2015)</i>	Estimation (Weather)	1930–1979
<i>Daymet</i>	Estimation (Weather)	1980–2019
<i>NASA NEX-GDDP CMIP6</i>	Estimation/Simulations (Weather)	2020–2100
<i>Mexican full-count Census</i>	Simulations	2020

Table A1 presents the data sources used in this paper.

A.1 Mexican Migration Project

The MMP is a joint data collection effort between Princeton University and Universidad de Guadalajara that “aims to gather and maintain high-quality data on the characteristics and behavior of documented and undocumented Mexican migrants to the United States”³³. To accomplish this, the MMP has sent surveyors out to between three and five pre-selected communities each Winter since 1982, where data are collected from a random sample of 200 households.

During these interviews, the surveyors collect various types of information from each household head. These include demographic and health information about both the head of the household and their spouse, detailed information on the first and last domestic and international migratory trips by both of them as well as any other migrants in their household, and a variety of different economic variables for these individuals, such as their current (time of the survey) wage. Importantly, each household head also provides the surveyor with their life history at the yearly level and that of their spouse. In this history, they are asked to provide information about where they have lived and their primary occupation during every year of their life. The publicly available dataset containing these variables censors some of the location data to preserve the confidentiality of those individuals surveyed;³⁴ however, I was provided access to the censoring algorithm and a crosswalk necessary for de-censoring the data so that I am able to verify the exact municipality of residence for the majority of individuals surveyed. This is the primary section of the survey that I exploit in this project.

³³Quote is taken from the [MMP’s home webpage](#), accessed 03/20/23

³⁴1.9% of person-year observations were either censored or left blank due to issues with recall on the part of the surveyed individual.

The following restrictions are made in selecting the sample used to estimate the structural model:

- Birth date not before 1950
- Birth location within Mexico
- After the de-censoring algorithm is employed, known birth location
- After the de-censoring and imputation algorithms are employed³⁵, municipality of residence for each year of the life history is known and is in Mexico.
- No years with missing state variables

The first condition is for two reasons. First, I use the previous 30 years' of weather realizations when estimating state transitions. Restricting births to not before 1950 implies that I begin modeling choices in 1967, which means that I require weather data from 1930 (I discretize years into decades which transition stochastically). Moreover, it helps ensure that the estimated wage data are not used to predict too far out sample. Further, I require that each individual be born in Mexico and have a known municipality of birth because past work has shown that individuals strongly prefer to live in their home location (e.g., [Kennan and Walker, 2011](#)). The remaining requirements are enforced to guarantee a complete panel for each individual in the data, which is necessary to estimate the dynamic model.

The life history data has an annual frequency; however, the model is instead designed with a triennial structure. To convert the data to this frequency, I employ the following algorithm, given individual i at time-step t :

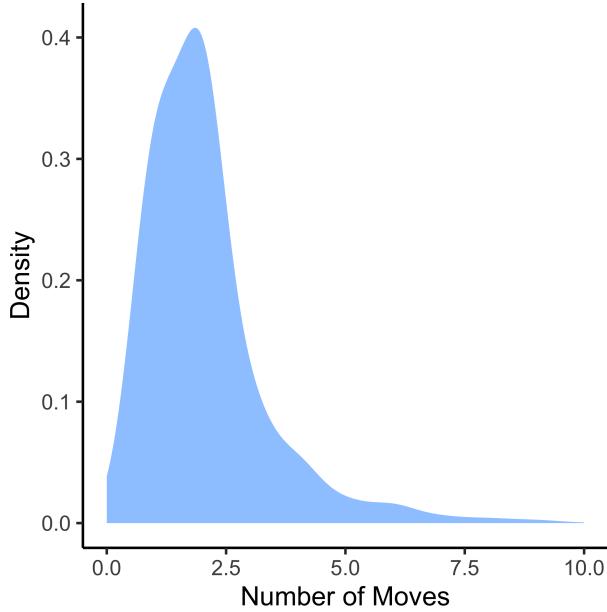
- If there are no moves during the three-year period, i 's starting location at time $t + 1$ is the same as their location at time t . Assign state variables according to the first year

³⁵Even after employing the de-censoring algorithm, there are still some individuals (834) who live in an unknown municipality at some point in their life. I drop any of these individuals who live in an unknown municipality and an unknown state. For the remaining individuals, given knowledge of the state, but not municipality, of residence, I use the following algorithm to impute their location in those years:

- If they live in a state that they do not again live in with a known municipality, drop them from the data
- If they live in a state where they either have previously lived or will live in the future, and at that time there is a single known municipality of residence, assign the years that they reside in that state, but with an unknown municipality, as the same known municipality in that state.
- If they live in a state where they either have previously lived or will live in the future, and also have lived/will live in multiple known municipalities in that state, assign the years that they reside in that state, but with an unknown municipality, with the municipality that they live in the year nearest the year with missing information.

This algorithm removes 517 individuals from the panel.

Figure A1: Density of Number of Moves in the MMP Data, For Migrants



in this period.³⁶

- If there is a move during the three-year period, i 's starting location at time $t + 1$ is the location of their first move during this period. Assign state variables according to the year of this move.

A.2 Wages

Given a lack of robust wage data in the MMP life histories, I use Census surveys to estimate a wage equation. More specifically, I use the 1984, 1989, 1992, 1994, 1996, 1998, 2000, 2002, 2004, and 2006 versions of La Encuesta Nacional de Ingresos y Gastos de los Hogares (ENIGH), and the quarterly 2008 through 2019 versions of La Encuesta Nacional de Ocupación y Empleo (ENOE). I use these data to create a repeated cross section of wages for each of the above years. When forming this cross section, I restrict the sample to males, who are the head of their household between the ages of 17 and 61, with positive hours worked and wages, who are not self-employed or working for a cooperative, and who live in one of the municipalities visited by an individual in the MMP data. To adjust for outliers, I remove

³⁶For example, the first time period covers ages 17-20. If i does not move during these years, then their state variables for this period are those when they are 17.

the top and bottom 5% of the wage data within each location-year. This leaves me with 3,626,618 observations. Table A2 presents the estimated coefficients corresponding to the weather impacts on wages.

A.3 Köppen Climates

The Köppen climate classification system is a widely used system for classifying locations in space by their differing local climates (Köppen, 1884). This system is global in scale, and as such is not able to capture local idiosyncrasies in climate. For this reason, El Instituto de Geografía and La Universidad Nacional Autónoma de México have developed a more detailed classification system for Mexico (García, 2004). This system is based on the Köppen system, but is more detailed with respect to the local nuances of the climates within Mexico. I use this updated classification system in this paper, and I discuss the specifics of it here.

I assign each municipality their climate class based on the classification which contains the most area of the municipality. The classification system is broken into four main groups: warm (cálido), dry (seco), temperate (templado), and cold (frío). None of the municipalities have a majority of their area in the cold region, so I do not discuss this group further. Each of the three remaining groups is further broken down into their subgroup: warm into warm warm and warm semi-warm, dry into dry very dry, dry dry, and dry semi-dry, temperate into temperate semi-warm, temperate temperate and temperate semi-cold. The classification is then broken down one more time: warm and temperate groups into those with humid and subhumid climates, and the dry group into very warm, warm, semi-warm with warm winter, semi-warm with cool winter, temperate with warm summer, temperate with cool and long summer and semi-cold.

There are no municipalities in the warm semi-warm humid class, so there are three warm climates: warm warm humid, warm warm subhumid, and warm semi-warm subhumid.

I combine temperate semicold humid (one municipality) and temperate semicold subhumid (nine municipalities) into a single class, temperate semi-cold. There are therefore five temperate climates: temperate semi-warm humid, temperate semi-warm subhumid, temperate temperate humid, temperate temperate subhumid, and temperate semi-cold.

I combine all dry climates based on their second grouping with “warm” and “temperate” third level classifications (there are no temperate semi-cold municipalities). This results in six dry climates: dry warm semi-dry, dry temperate semi-dry, dry warm dry, dry temperate dry, dry warm very dry, and dry temperate very dry.

Figure 4, in the main text, plots the resulting climates throughout Mexico.

Table A2: Wage Equation Estimates, Weather

	Hourly Income
Ag worker	0.74 (3.55)
Growing Degree Days	7.5*** (2.8)
Growing Degree Days ²	-0.95** (0.44)
Drought	-0.38* (0.20)
Inundation	-0.036 (0.310)
Ag worker × GDD	-3.0 (2.8)
Ag worker × GDD ²	0.58 (0.52)
Ag worker × Drought	0.38 (0.25)
Ag worker × Inundation	-0.44* (0.24)
Temperate × GDD	-5.1 (3.8)
Warm × GDD	-6.4 (14.6)
Temperate × GDD ²	0.75 (0.75)
Warm × GDD ²	0.87 (2.17)
Temperate × Drought	0.77*** (0.22)
Warm × Drought	0.66** (0.33)
Temperate × Inundation	-0.56* (0.32)
Warm × Inundation	-0.13 (0.36)
Num.Obs.	3626618
R2	0.231
FE: Location, Year	X
Dep. Var. Mean	27.06

Note: Estimates correspond to the coefficients from Equation 7 in the main text. Wages are measured in 2010 hourly pesos. Standard errors, in parentheses, are clustered at the municipality level. *p < 0.1; **p < 0.05; ***p < 0.01

A.4 Metropolitan Zones

La Secretaría de Desarrollo Social, El Consejo Nacional de Población, and El Instituto Nacional de Estadística y Geografía have developed a classification system for municipalities in Mexico based on their population size and economic activity into metropolitan zones (las zonas metropolitanas). This system is explicitly designed to determine which parts of the country are “urban”. I use their 2004 definitions, based on 2000–2010 populations, to classify each municipality as either urban, if there is a metropolitan zone in the municipality, and rural, otherwise. There are 63 such metropolitan zones, covering 1,021 municipalities.

B Methods & Empirical Implementation

B.1 Climate & Weather

The climate is a high dimensional physical system upon which conventional measures of weather are drawn. Given the complexity of the climate system and the need for a high degree of parsimony in the model to maintain computational feasibility, I assume that individuals keep track of two summary statistics in each of the 3 climates in Mexico: the average of the daily temperature distribution and the total level of agricultural-season precipitation. In total, individuals are assumed to keep track of 6 weather statistics, each of which contribute to their observed state.³⁷ Even though individuals are assumed to keep track of summary statistics of the climates, past work has shown that it is oftentimes extremes which drive economic outcomes. For this reason, I use rich reduced-form mappings of the relationship between these summary statistics and the relevant extreme outcomes to ensure that the model is able to capture the relevant variation driving migration decisions.

Individuals use the following mapping to determine the expected value of degree days above 26 degrees Celsius, degree days below 14 degrees Celsius, and growing degree days:

$$DD_\ell = \mathbb{1}\{\alpha_\ell^D + \alpha_{1\ell}^D \text{temp}_{C(\ell)} > 0\} \times [\alpha_\ell^D + \alpha_{1\ell}^D \text{temp}_{C(\ell)}]$$

Where each location ℓ has a different set of coefficients, α_ℓ^D for each of the degree day types. These coefficients are estimated using a censored Tobit framework to account for the fact that the degree days are non-negative and censored at 0. Table B1 provides summary statistics for these mappings, showing the number of zeros in each climate for each type of degree day,

³⁷This is similar, in spirit, to a moment-based Markov equilibrium (Ifrach and Weintraub, 2016), where individuals keep track of the states of a few dominant firms and a few summary statistics for the mass of fringe firms.

the average fixed effect (α_ℓ^D), and the average linear coefficient ($\alpha_{1\ell}^D$). Similarly, individuals

Table B1: Weather Mapping Statistics

	Growing	DD26	DD14
Dry			
N Zeros	0	1	0
Avg. FE	-580.75	-2222.66	1539.43
Avg. Temp.	173.14	115.92	-70.63
Temperate			
N Zeros	0	4	0
Avg. FE	-929.01	-1733.00	2035.91
Avg. Temp.	172.86	44.89	-96.88
Warm			
N Zeros	0	0	1
Avg. FE	-442.82	-1936.38	-192.89
Avg. Temp.	169.49	138.17	57.38

Note: Estimates correspond to the coefficients from the Tobit regression, as described in Section 3. N Zeros represents the number of locations with no degree days in that category. Average fixed effect (FE) and temperature are the average of the (non-zero) fixed effects and level-one climate impacts, respectively, for locations in that climate. DD26 and DD14 are degree days above 26 and below 14 degrees Celsius, respectively.

use the following mapping to determine the probability of either drought or flooding in each of the locations, as predicted by the level of precipitation in the agricultural season and the average temperature in the relevant level-one Köppen climate:

$$\text{prob}(R_\ell) = \Phi(\gamma_\ell^R + \gamma_1^R P_{C(\ell)} + \gamma_2^R \text{temp}_{C(\ell)})$$

Where each location ℓ has its own fixed effect but a shared set of precipitation and temperature coefficients, for each type of extreme rain event. These coefficients are estimated using probit regression. Table B2 provides summary statistics for this mapping, showing the average fixed effect (γ_ℓ^R), and the shared precipitation and temperature coefficients.³⁸

³⁸It is important to note that drought and flood are defined using the SPEI, which accounts for differences in potential and realized evapotranspiration. This means that these events are defined relative to the local climate, rather than in absolute terms.

Table B2: Rainfall Probability Estimates

	Drought	Flood
Precipitation	-1.34	2.31
Temperature	0.11	-0.15
Climate (Average FE)		
Dry	-2.54	1.29
Temperate	-2.09	0.52
Warm	-2.21	0.68

Note: Estimates correspond to the coefficients from the probit regressions of level-one climate variables on drought and flood risk, described in Section 3. The bottom panel reflects the average fixed effect across climates for the particular rainfall event.

Table B3: Weather Mapping Fit

Growing Degree Days	Degree Days above 26	Degree Days below 14	Drought	Flood
R ² : Full	R ² : No FE		Pseudo R ²	
0.985	0.647	0.239	0.239	0.041 0.101

Note: Growing degree days are for maize, and refer to degree days with a lower kink at 8°C and an upper kink at 32°C. Drought and flood refer to an SPEI below (above) 1.5. I use the McFadden pseudo R². All statistics refer to the mappings presented in the main text: Equation 5 and Equation 6.

B.2 Transitions

I employ a two-stage estimation routine, where in the first stage I estimate the distribution of state variable transitions, which I feed into the second stage nested-full solution maximum likelihood algorithm. I discuss here the estimates from the first stage.

An individual's state is described by the following set:

$$\left\{ \underbrace{\nu^\ell, \text{educ}, \text{ag}}_{\text{Constant}}, \underbrace{\text{age}}_{\text{Deterministic Trans.}}, \underbrace{\text{dec}, k, \mathbf{w}}_{\text{Stochastic Trans.}}, \underbrace{\ell}_{\text{Endogenous}}, \underbrace{\tau}_{\text{Unobserved}}, \varepsilon \right\}$$

Where ν^ℓ is an individual's birth location, educ is either low, medium, or high representing less than 6, between 6 and 12, and at least 12 years of education, respectively, ag is 1 if the individual works in agriculture at least 25% of their observed life history and 0 otherwise. dec transitions to the next decade with probability 0.3 and remains the same with probability 0.7. age transitions to the next time-step with probability 1. τ represents the moving-belief type,

Table B4: First Stage Parenthood Transitions

	$\mathbb{1}\{\text{New Child}\}$
Intercept	-1.785*** (0.018)
Time Step	0.377*** (0.011)
Time Step Squared	-0.0373*** (0.0014)
Num.Obs.	82256
RMSE	0.32

Notes: New Child is a binary variable representing whether an individual becomes a new parent in the following period. That is, they have no children in the current period and at least one child in the following. Estimation is done using a probit regression on the subsample of non-parents in the MMP data. *** $p < 0.01$.

ℓ the current location, and ε the vector of location-individual-time-specific T1EV draws.

k is an indicator for whether or not an individual is currently a parent. This is an absorbing state; that is, once an individual is a parent, they will always be a parent: $k_t = 1 \Rightarrow k_s = 1 \forall s \geq t$. If $k_t = 0$, then $k_{t+1} = 1$ with some probability p , where $p = \Phi(\gamma_0 + \gamma_1 age + \gamma_2 age^2)$. I estimate γ using the MMP data with a probit regression - Table B4 presents the estimated coefficients.

w represents a six-dimensional vector containing the average daily temperature and level of agricultural-season precipitation in each of the six climates in Mexico. I assume that this vector follows a multivariate Gaussian:

$$w \sim \mathcal{N}(\mu_y, \Sigma_y)$$

Where $\mu_{iy} = \rho_{i0} + \rho_{i1}y + \rho_{i2}\mu_{it-1}$. The mean of each component of this Gaussian is assumed to be an AR(1) process, augmented with a time trend to account for climate change over time. Table B5 presents the estimates of ρ for the historical climate used in estimation. Σ represents the covariance matrix, obtained via a seemingly unrelated regression of each μ_i .

B.3 Identification of Preference Parameters

Section 4 of the main text provides examples of the sources of variation in the estimation data which identify the preference parameter estimates. Here, I highlight two sources of variation,

whose existence can be verified using reduced form statistics. In particular, the first column of Table B6 highlights differences in average migration rates across differences in wages, age, experienced weather, whether an individual has migrated in the past, and whether they were born in their current location. All of these differences are strongly significant conditional on the others and differences across rural and urban locations. The second column of Table B6 highlights the relationship between the length of stay for migrants, measured in years, and the same set of variables. This set of comparisons highlights a source of identification for the preference for living in an urban location, conditional on all other variables, including expected wages.

B.4 Moving Costs

The estimated structural model allows for a calculation of frictions associated with any particular move seen in the estimation data. Section 5 discusses three different ways of contextualizing such estimates. The moving cost for an individual observed moving from location ℓ to location ℓ' , with current-period state variables given by ω ³⁹ ($\bar{\Delta}(\ell, \ell', \omega)$), the moving cost, is defined as the difference in flow utilities, inclusive of the estimated moving cost, given by 3, net of the expected difference in the preference shock, conditional on making choice ℓ' . Mathematically:

$$\bar{\Delta}(\ell, \ell', \omega) = u(\ell, \ell, \omega) - u(\ell, \ell', \omega) - \mathbb{E}[\varepsilon_{\ell'} - \varepsilon_\ell | c = \ell', \ell, \omega] \quad (10)$$

Borrowing from the conditional choice probability literature (Hotz and Miller, 1993), Kennan and Walker (2011) show that Equation 10 can be rewritten as something equivalent to:

$$\bar{\Delta}(\ell, \ell', \omega) = \frac{\log(p(\ell, \ell, \omega))}{1 - p(\ell, \ell, \omega)} + \delta \int \left[\log \left(\sum_{k=1}^{27} \exp(\bar{v}_{t+1}(\ell', k, \omega')) \right) - \log \left(\sum_{k=1}^{27} \exp(\bar{v}_{t+1}(\ell, k, \omega')) \right) \right] dF_{\omega'|\omega, \tau} \quad (11)$$

Equation 11 contains three ingredients worth discussing. The first of these is $p(\ell, \ell, \omega)$. This is the model-predicted probability of an individual residing in location ℓ not moving in their current period, conditional on their current state. The moving cost is increasing in this value. Intuitively, the probability of not moving is positively related to the observable difference in flow utility from staying in the current location and moving to any other location, including ℓ' . For this reason, it accounts for part of the first difference in Equation 10. The second and third ingredient represent the discounted continuation value from starting next period in ℓ' and ℓ , conditional on ω , respectively. The moving cost is increasing in the continuation value

³⁹Note, here I include an individual's time-step, t in ω . This is different from the definition given the main text.

of starting in location ℓ' . To see why this is, note that conditional on a given difference in flow utilities and ε_ℓ , a higher expected continuation value from choosing location ℓ' means that the conditional ε' has less work to do to convince an individual to move to ℓ' . This causes Equation 10 to be larger, again, all else held equal. Similar logic can explain the presence of the final term: the continuation value associated with remaining in ℓ . Table B7 provides statistics on the expected moving costs of the moves used in estimating the structural model.

B.5 Estimation

I discuss here a few technical details regarding estimation of the dynamic model. The first of these is the discretization procedure used. As discussed in Appendix B.2, the only continuous state variable in the model is the set of weather states. I discretize each of the continuous weather variables into 3 bins, where the lower and upper limits are determined such that practically all of the observed data lie within them. The dry and temperate climates each have a temperature bin-width of 1 degree Celsius, while the warm climate has a bin-width of 0.5 degrees Celsius.⁴⁰ Similarly, the dry climate has a rainy-season bin-width of 0.1 meters, while the temperate and warm climates have a bin-width of 0.2 meters.

I use multilinear interpolation to interpolate the value function outside/inside of these bins. In particular, for a given weather state, \mathbf{w} , I find the nearest bin above and below w_i for each i in 1 through 6.⁴¹ I then form the set of all combinations of these grid points, there are 64 (2^6) such combinations. I then calculate the Euclidean distance from \mathbf{w} to each combination, and set the relevant value, at \mathbf{w} , as the inverse-distance weighted average of the value at each combination.

The above discretization leads to a state space with roughly 4.6 billion different points. The model solution explicitly solves the Bellman at each of these points. The MLE framework requires that the model be solved thousands of times.⁴² The large and non-stationary state space is necessary to study two intrinsically dynamic objects: migration and climate change. In estimation, I exploit separability across many large swathes of the state space to parallelize the computation across a cluster of 9 high performance GPU nodes on the University of Arizona's High Performance Computing cluster. I also exploit sparsity of many sections of the transition distribution to further reduce the computational burden⁴³. Moreover, I derive a

⁴⁰There is much less variation in mean temperature in the warm climate than in the other two.

⁴¹If w_i is above (below) the upper (lower) limit, then I use the upper (lower) limit for both of the points described.

⁴²For example, the primary version of the model required 2,639 evaluations of the solution procedure.

⁴³I employ a custom compressed row storage (CRS) routine for the transition matrix. In particular, I modify the generic CRS algorithm to allow for a sparse tensor. This transition object is a tensor rather than a matrix as each decade has its own 2-d transition matrix.

closed form representation of the model solution which avoids numerical integration over the preference shocks. Using NVIDIA P100 GPUs, each likelihood evaluation takes around three minutes, with the likelihood maximized using a highly-effective black-box solver ([Montoisin et al., 2020](#)).

Standard Errors

To obtain standard errors on the structural parameter estimates, I calculate central differences around the optimal parameter vector. To be clear, I calculate standard errors using the following algorithm:⁴⁴

1. Perturb θ_k^* above and below by $|\theta_k^*|\sqrt{\epsilon}$
2. Calculate $\log(\text{prob}_i|\theta_k^p)$ for each perturbation direction (p), k , and i
3. Calculate $\frac{\partial \log(\text{prob}_i)}{\partial \theta}$ using finite central differences
4. Create individual contributions to the score as $\frac{\partial \log(\text{prob}_i)}{\partial \theta} \frac{\partial \log(\text{prob}_i)}{\partial \theta}'$
5. Calculate information matrix as $\mathcal{J} = \frac{1}{n} \sum_{i=1}^n \frac{\partial \log(\text{prob}_i)}{\partial \theta} \frac{\partial \log(\text{prob}_i)}{\partial \theta}'$
6. Invert \mathcal{J}
7. Divide \mathcal{J}^{-1} by the number of people (n)
8. The square root of the resulting diagonal converges to the vector of standard errors

The resulting standard errors are clustered at the individual level, due to the finite-mixture structure imposed on the problem.

B.6 Translation to Dollars

I discuss here the mapping I use to translate values, as measured in utils, into dollars. Mathematically, suppose that we have some difference in value functions, u , which is measured in utils. Notice that the estimated marginal utility of income, θ_1 from [Equation 2](#), is measured in utils per 2010 peso per hour per year. Then normalizing u by θ_1 delivers u_1 , which is measured in 2010 pesos per hour per year. I multiply this by the average number of hours worked by an individual living in Mexico in 2010: 2244.⁴⁵ This gives the value of u_2 , which is measured in 2010 pesos. I bring u_2 to 2010 dollars using the exchange rate on January 1,

⁴⁴ ϵ , here, represents machine precision for 16-bit floats

⁴⁵I get this figure from: [OECD Data](#), accessed 7/9/2024.

2010: 12.8096 pesos/dollar⁴⁶ and then to 2024 dollars using the consumer price index for all urban consumers: 1 2010 dollar to 1.45 2024 dollars.

B.7 Calculating Biased Welfare

I discuss here how I calculate the welfare of individuals forming naive expectations of the climate system. Before doing so, I highlight that for the fully-informed individuals, their remaining lifetime welfare is given by the sum of their alternative specific value function associated with the state and particular period t choice and the logit draw for that decision. This is because their expectations mirror the true climate system. On the other hand, the naive individuals make their choices based off a biased Bellman, and as such, their welfare will be lower than that given by this Bellman.

I highlight that in the terminal period a naive and fully-informed individual share the same true alternative-specific value function. This is because after this period there are no further choices to be made. I use this fact to recursively calculate the welfare for naive individuals. Mathematically, at time T , a naive individual's true welfare (net of their current period stochastic shock) is given by:

$$\bar{V}_T^{nt}(\ell, \ell', \omega) = \bar{V}_T(\ell, \ell', \omega),$$

where $\bar{V}_T(\ell, \ell', \omega)$ is as given by Equation 9. For time $t < T$, the naive individuals welfare (net of the stochastic logit draw) is given by:

$$\begin{aligned} \bar{V}_t^{nt}(\ell, \ell', \omega) &= u(\ell, \ell', \omega) + \delta^3 \mathbb{E} [V_{t+1}^{nt}(\ell', \omega') | \omega] \\ &= u(\ell, \ell', \omega) + \delta^3 \iint \left\{ \bar{V}_{t+1}^{nt}(\ell', j, \omega') + \mathbb{E} [\varepsilon_j | d^n = j] \right\} dG^n(j | \ell', \omega') dF(\omega' | \omega) \\ &= u(\ell, \ell', \omega) + \delta^3 \iint \left\{ \bar{V}_{t+1}^{nt}(\ell', j, \omega') + \gamma + \log \left(\sum_{k=1}^{27} \exp (\bar{V}_{t+1}^n(\ell', k, \omega') - \bar{V}_{t+1}^n(\ell', j, \omega')) \right) \right\} \\ &\quad dG^n(j | \ell', \omega') dF(\omega' | \omega), \end{aligned}$$

where \bar{V}^n is the biased alternative specific value function for naive individuals and d^n is the decision made based on this biased view of the world. The expectation operator on the first line of this equation is taken over the joint distribution of the state transitions (F), the stochastic shocks, and the naive choice probabilities (G^n). The second line uses the definition of the alternative-specific value function to rewrite the true value of a naive individual from living in location ℓ' with state ω' as a weighted sum of the true expected value of living in

⁴⁶[FRED Data](#), accessed 7/9/2024.

location ℓ' with state ω' and choosing location j , where the weight is given by the naive choice probability.⁴⁷ The third line borrows from Kennan (2008) to rewrite the expected value of the logit draw, conditional on the choice made based on the biased alternative-specific value function (\bar{V}^n), in the form given above.

B.8 Stylized Policy Example

Suppose that we have some individual i who lives for two periods and, without loss, begins in location b with the opportunity to move at the beginning of each period, t . Further, suppose for simplicity that they do not discount the future ($\delta = 1$). They receive utility u^a from living in location a and u^b from living in location b . Assume that moving costs in period t are distributed according to some distribution: $c_t \sim F(\cdot)$, and that they make their period t moving decision after observing c_t , but before receiving flow utility for period t . We can solve for the period 1 decision using backward recursion. In particular, notice that if they find themselves in location b at the start of $t = 2$, they will move to location a if:

$$u^a - c_2 > u^b, \quad (12)$$

and if they find themselves in location a at the beginning of period 2, they will move back to location b if:

$$u^b - c_2 > u^a, \quad (13)$$

We can use Inequality 12 and Inequality 13 to solve for their period 1 choice. In particular, i will move to location a in period 1 if:

$$u^a - c_1 + \underbrace{u^a[1 - F(u^b - u^a)] + (u^b - E[c_2])F(u^b - u^a)}_{E[V(a)]} > u^b + \underbrace{u^b(1 - F(u^a - u^b)) + (u^a - E[c_2])F(u^a - u^b)}_{E[V(b)]}. \quad (14)$$

Where $E[V(\cdot)]$ represents the expected value of beginning period 2 in location \cdot . Inequality 14 holds when:

$$c_1 < (u^a + E[V(a)]) - (u^b + E[V(b)]).$$

Or when the current period moving cost is less than the gain from moving. We can use this to define the ex ante probability that i moves in the first period from location b to location a as:

$$p(m_1) = \int_{-\infty}^{(u^a + E[V(a)]) - (u^b + E[V(b])]} dF.$$

⁴⁷Another way to see this is through an application of the law of iterated expectations taken across the joint distribution of the naive choice probabilities and the state transitions.

Let $z = (u^a + E[V(a)]) - (u^b + E[V(b)])$. Notice that we can write the expected number of moves as:

$$\begin{aligned} & \underbrace{F(z)}_{\text{Period 1 moves}} + \underbrace{F(u^a - u^b)[1 - F(z)]}_{\text{Period 2 moves from } b} + \underbrace{F(u^b - u^a)F(z)}_{\text{Period 2 moves from } a} \\ &= F(z) + \underbrace{F(u^a - u^b) + F(z)[F(u^b - u^a) - F(u^a - u^b)]}_{\text{Period 2 moves}}. \end{aligned} \quad (15)$$

Where m_t represents movers in period t .

Static Policy in Period One

Imagine now that there is a static reduction in moving costs for the first period of $s > 0$. This only shifts the left hand side of Equation 14 to the right by s . In particular, they now move to a in period 1 if:

$$c_1 < (u^a + E[V(a)]) - (u^b + E[V(b)]) + s,$$

which gives the ex ante probability that i moves in the first period as:

$$\begin{aligned} p(m_1 | s_1) &= \int_{-\infty}^{(u^a + E[V(a)]) - (u^b + E[V(b)]) + s} dF \\ &= \underbrace{\int_{-\infty}^{(u^a + E[V(a)]) - (u^b + E[V(b)])} dF}_{\text{Inframarginal movers}} + \underbrace{\int_{(u^a + E[V(a)]) - (u^b + E[V(b)])}^{(u^a + E[V(a)]) - (u^b + E[V(b)]) + s} dF}_{\text{Marginal movers}}. \end{aligned}$$

We can write the expected number of moves as:

$$\begin{aligned} & \underbrace{F(z + s)}_{\text{Period 1 moves}} + \underbrace{F(u^a - u^b)[1 - F(z + s)]}_{\text{Period 2 moves from } b} + \underbrace{F(u^b - u^a)F(z + s)}_{\text{Period 2 moves from } a} \\ &= F(z + s) + \underbrace{F(u^a - u^b) + F(z + s)[F(u^b - u^a) - F(u^a - u^b)]}_{\text{Period 2 moves}}. \end{aligned} \quad (16)$$

Given that $s > 0$, the static policy will induce weakly more moves than the base setting with no policy intervention. Moreover, we can see from the second part of Equation 16 that the static policy detracts from the overall number of second period migrants from b when $F(z + s) > F(z)$. That said, it also increases the number of second period migrants back to b , by increasing the number of individuals in a . When $u^b > u^a$, the second channel overpowers the first: the static subsidy will induce weakly more second period migrants than the baseline.

When $u^a > u^b$, the difference in second period migrants between the static policy and the baseline is ambiguously signed.

Static Policy in Period Two

Now suppose instead that there is a static reduction in *second* period moving costs of $s > 0$. i 's expected value of moving to location a in period 1 is given by:

$$u^a + \underbrace{u^a[1 - F(u^b - u^a + s)] + (u^b - E[c_2] + s)F(u^b - u^a + s)}_{E[V(a)|s_2]} - c_1,$$

and of staying in location b by:

$$u^b + \underbrace{u^b(1 - F(u^a - u^b + s)) + (u^a - E[c_2] + s)F(u^a - u^b + s)}_{E[V(b)|s_2]}.$$

They will move to location a at the beginning of the first period if:

$$\begin{aligned} c_1 &< (u^a + E[V(a)|s_2]) - (u^b + E[V(b)|s_2]) \\ &= (u^a + u^a[1 - [F(u^b - u^a) + p(u^b - u^a \leq c_2 \leq u^b - u^a + s)]] \\ &\quad + (u^b - E[c_2] + s)[F(u^b - u^a) + p(u^b - u^a \leq c_2 \leq u^b - u^a + s)]) \\ &\quad - (u^b + u^b(1 - [F(u^a - u^b) + p(u^a - u^b \leq c_2 \leq u^a - u^b + s)])) \\ &\quad + (u^a - E[c_2])[F(u^a - u^b) + p(u^a - u^b \leq c_2 \leq u^a - u^b + s)] + sF(u^a - u^b + s)) \\ &= (u^a + u^a[1 - F(u^b - u^a)] + (u^b - E[c_2])F(u^b - u^a)) \\ &\quad - (u^b + u^b[1 - F(u^a - u^b)] + (u^a - E[c_2])F(u^a - u^b)) \\ &\quad + (-u^a p(u^b - u^a \leq c_2 \leq u^b - u^a + s) + (u^b - E[c_2])p(u^b - u^a \leq c_2 \leq u^b - u^a + s) + sF(u^b - u^a + s)) \\ &\quad - (-u^b p(u^a - u^b \leq c_2 \leq u^a - u^b + s) + (u^a - E[c_2])p(u^a - u^b \leq c_2 \leq u^a - u^b + s) + sF(u^a - u^b + s)) \\ &= \underbrace{(u^a + E[V(a)]) - (u^b + E[V(b)])}_{\text{Inframarginal movers}} \\ &\quad + \underbrace{((u^b - u^a - E[c_2])p(u^b - u^a \leq c_2 \leq u^b - u^a + s) + sF(u^b - u^a + s))}_{\text{Increase in period 1 movers, from subsidy to move back}} \\ &\quad - \underbrace{((u^a - u^b - E[c_2])p(u^a - u^b \leq c_2 \leq u^a - u^b + s) + sF(u^a - u^b + s))}_{\text{Reduction in period 1 movers, from subsidy in period 2}}. \end{aligned}$$

This highlights that there are two countervailing forces at play with a static policy in the second period. In particular, some individuals will decide to move in the first period because they know they can move back at a reduced cost. On the other hand, some individuals will

decide not to move in the first period because they would prefer to wait until the following period to move to a when the subsidy is available.

One-Time Dynamic Policy

Now instead suppose that there is a dynamic subsidy which is only available to be used at the first time an individual decides to move. In this case, i 's expected value from moving to a in the first period is given by:

$$u^a + \underbrace{u^a[1 - F(u^b - u^a)] + (u^b - E[c_2])F(u^b - u^a)}_{E[V(a)]} - c_1 + s,$$

and their expected value from staying in location b is given by:

$$u^b + \underbrace{u^b(1 - F(u^a - u^b + s)) + (u^a - E[c_2] + s)F(u^a - u^b + s)}_{E[V(b)|s_2]}.$$

Notice that the expected value of moving to location a in the first period is the same as that of the static offering in the first period, while their expected value of staying in location b is the same as that of the static offering in the second period. This means that they will move when:

$$c_1 < (u^a + E[V(a)]) - (u^b + E[V(b)|s_2]) + s,$$

or when:

$$\begin{aligned} c_1 &< \underbrace{(u^a + E[V(a)]) - (u^b + E[V(b)]) + s}_{\text{Movers under static policy in Period 1}} \\ &\quad - \underbrace{((u^a - u^b - E[c_2])p(u^a - u^b \leq c_2 \leq u^a - u^b + s) + sF(u^a - u^b + s))}_{\text{Change in period 1 movers, from subsidy in period 2}}. \end{aligned}$$

The ex ante probability of i choosing to move in the first period under this policy is therefore given by:

$$p(m_1|s_{d1}) = \underbrace{\int_{-\infty}^z dF}_{\text{Inframarginal movers}} + \underbrace{\int_z^{z+s} dF}_{\text{Marginal movers under static policy}} + \underbrace{\int_{z+s}^{z+s-w} dF}_{\text{Change in movers under dynamic policy}},$$

where $z = (u^a + E[V(a)]) - (u^b + E[V(b)])$ and $w = ((u^a - u^b - E[c_2])p(u^a - u^b \leq c_2 \leq u^a - u^b + s) + sF(u^a - u^b + s))$. Note the following:

$$w > 0 \Leftrightarrow \underbrace{(u^a - u^b - E[c_2])}_{\text{Expected gain (or loss) from period 2 move}} \underbrace{[F(u^a - u^b + s) - F(u^a - u^b)]}_{\Delta \text{ in the probability of moving because of the subsidy}} + \underbrace{sF(u^a - u^b + s)}_{\text{Expected value of the subsidy}} > 0. \quad (17)$$

When the value of moving is larger than the expected cost, a dynamic policy will always induce fewer individuals to move in the first period. This is because some individuals who received a poor cost draw understand that their next period draw will likely be more favorable. How many individuals find themselves in this position depends both on how strongly the subsidy shifts their behavior and the size of the gain from migration.

Determining the sign of w is more nuanced when the value of moving is *smaller* than the expected cost. Suppose that the expected change in utility from moving in period 2 ($E[\Delta u]$) is less than 0. Notice that w is positive when:⁴⁸

$$\begin{aligned} E[\Delta u] [p(m) - p(m|s)] &< sp(m|s) \\ \frac{p(m) - p(m|s)}{p(m|s)} &> \frac{s}{E[\Delta u]} \\ \underbrace{\text{abs } (\% \Delta p(m|s))}_{\text{Percentage change in period 2 movers from losing the subsidy}} &< \underbrace{\text{abs } \left(\frac{s}{E[\Delta u]} \right)}_{\text{Size of subsidy, relative to expected loss from moving}} . \end{aligned}$$

That is, w is positive when the shift in the probability of moving in the second period, as brought about by removing subsidy, is less than the value of the subsidy relative to the expected loss from moving. For a sufficiently small subsidy, this highlights that the slope of the density of moving costs at the endpoint of the inframarginal individuals drives whether a dynamic subsidy will induce more or less individuals, than a static policy, to move in the first period. This will also determine whether a dynamic policy will induce positive or negative selection as compared to a static policy.

We can write the expected number of moves as:

$$\begin{aligned} &\underbrace{F(z + s - w)}_{\text{Period 1 moves}} + \underbrace{F(u^a - u^b) [1 - F(z + s - w)]}_{\text{Period 2 moves from } b} + \underbrace{F(u^b - u^a) F(z + s - w)}_{\text{Period 2 moves from } a} \\ &= F(z + s - w) + \underbrace{F(u^a - u^b) + F(z + s - w) [F(u^b - u^a) - F(u^a - u^b)]}_{\text{Period 2 moves}} \end{aligned} \quad (18)$$

⁴⁸The first line only rewrites 17. The second line uses the fact that $E[\Delta u] < 0$ when flipping the inequality. The third line multiplies each side by -1 to make both sides positive.

Notice that $w > 0$ implies that the dynamic policy will reduce the number of first period movers and the number of second period movers from a back to b . Mathematically, this is because $F(z + s - w) [1 + F(u^b - u^a)] \leq F(z + s) [1 + F(u^b - u^a)]$. Intuitively, this is because the dynamic policy induces some individuals who would have moved from b to a in period 1 under the static policy to instead spend their period in location b in the hopes that next period's draw will be better. That said, these individuals are not entirely giving up on the prospect of moving to a - they are only delaying the expected move. This means that the dynamic policy will induce *more* individuals to move to a in the second period, as compared to the static offering. Mathematically, this enters through the fact that $F(u^a - u^b)[1 - F(z + s - w)] \geq F(u^a - u^b)[1 - F(z + s)]$, for $w > 0$. Which of these two forces prevails depends on the particular distribution from which moving costs are drawn and the gain from migration.

Further, notice that the dynamic subsidy will weakly generate higher welfare than the first-period static policy. To see why this is, notice that the dynamic policy can be broken apart into two distinct policies. In the first period, individuals can take exactly the static policy, however they further have the option of choosing the static policy in the second period (albeit only from a starting position of b , rather than a). This ability to choose generates option value, in itself. Moreover, when $w > 0$, the reduction in the fraction of the population marginal to moving gives the selection results I highlight in the main text. In particular, the individuals who decide not to move under the dynamic policy are those who were marginal to the policy, which means that they received a relatively high cost draw in the current period, which is why they required the subsidy to induce their move. The differential effects of selection into the policy is what drives the large differences in the welfare outcomes of movers under the different dynamic structures of policy.

Finally, notice that w is exactly the option value of the dynamic subsidy. To see why this is true, I highlight that the value of having the option to migrate in the second period is given by:

$$\underbrace{F(u^a - u^b + s(p))}_{\text{Probability of moving in period 2}} \underbrace{[u^a - u^b - E[c_2] + s(p)]}_{\text{Value of moving in period 2}},$$

where $s(p)$ is equal to s for the dynamic policy and 0 for the static policy and the baseline. The option value of the dynamic subsidy is then the difference between the option value of moving in period two with the subsidy from the option value of moving in period 2 without

the subsidy. Mathematically:

$$\begin{aligned}
& F(u^a - u^b + s)[u^a - u^b - E[c_2] + s] - F(u^a - u^b)[u^a - u^b - E[c_2]] \\
& = [F(u^a - u^b + s) - F(u^a - u^b)][u^a - u^b - E[c_2]] + sF(u^a - u^b + s) \\
& = w.
\end{aligned}$$

This tells us that the dynamic policy will induce positive selection from those marginal to moving exactly when the option value of moving tomorrow, as brought about by the subsidy, is positive. The magnitude of this selection is increasing with the magnitude of the option value from the subsidy. Notice that we can very simply move away from a two period model and retain similar intuition to the above. In particular, we can generalize the above where now the value of moving to a in period t is given by $u^a + EV_a^{t+1}$ and the value of staying in b is given by $u^b + EV_b^{t+1}$. The value of having the option to move in the following period is therefore given by:

$$F((u^a + EV_a^{t+2}) - (u^b + EV_b^{t+2}) + s(p))[(u^a + EV_a^{t+2}) - (u^b + EV_b^{t+2}) - E[c_2] + s(p)].$$

The remainder of the calculations follow through as given above.

Fully Dynamic Policy

For completeness, now suppose instead that there is a policy which reduces one's moving costs by $s > 0$ in both periods. Under this scenario, i values moving to a in period 1 at:

$$u^a + \underbrace{u^a[1 - F(u^b - u^a + s)] + (u^b - E[c_2] + s)F(u^b - u^a + s)}_{E[V(a)|s_2]} - c_1 + s,$$

and of staying in b in period 1 at:

$$u^b + \underbrace{u^b(1 - F(u^a - u^b + s)) + (u^a - E[c_2] + s)F(u^a - u^b + s)}_{E[V(b)|s_2]}.$$

Notice that the expected value of moving to a in the first period is the same as that of the static policy in the second period, net of the subsidy received from moving. Moreover, the value of staying in b is exactly the same as that of the static policy in the second period. We can use this fact to derive the cutoff to move to a in the first period as:

$$c_1 < (u^a + E[V(a)|s_2]) - (u^b + E[V(b)|s_2]) + s$$

$$\begin{aligned}
&= \underbrace{(u^a + E[V(a)]) - (u^b + E[V(b)]) + s}_{\text{Marginal movers w/ static policy}} \\
&\quad + \underbrace{((u_b - u_a - E[c_2])p(u^b - u^a \leq c_2 \leq u^b - u^a + s) + sF(u^b - u^a + s))}_{\text{Change in period 1 movers, from subsidy to move back}} \\
&\quad - \underbrace{((u^a - u_b - E[c_2])p(u^a - u^b \leq c_2 \leq u^a - u^b + s) + sF(u^a - u^b + s))}_{\text{Change in period 1 movers, from subsidy in period 2}}.
\end{aligned}$$

The ex ante probability of i deciding to move in the first period is therefore given by:

$$p(m_1|s_d) = \underbrace{\int_{-\infty}^z dF}_{\star_1} + \underbrace{\int_z^{z+s} dF}_{\star_2} + \underbrace{\int_{z+s}^{z+s+v} dF}_{\star_3} + \underbrace{\int_{z+s+v}^{z+s+v-w} dF}_{\star_4},$$

where \star_1 are the inframarginal movers, \star_2 are the movers induced by the static first-period subsidy, \star_3 represents the movers induced to move by the prospect of a reduced return moving cost, and \star_4 represent individuals who prefer to wait until the following period to potentially move, $z = (u^a + E[V(a)]) - (u^b + E[V(b)])$, $v = (u_b - u_a - E[c_2])p(u^b - u^a \leq c_2 \leq u^b - u^a + s) + sF(u^b - u^a + s)$, and $w = (u^a - u_b - E[c_2])p(u^a - u^b \leq c_2 \leq u^a - u^b + s) + sF(u^a - u^b + s)$.

Table B5: Average First Stage Historical Climate Transitions

	Temperature	Precipitation
Dry		
Intercept	5.854507*** (1.4242)	0.297038*** (0.0321)
Time Trend	0.002950** (0.0013)	-0.000977*** (0.0002)
Lag	0.687234*** (0.0760)	0.067172 (0.0975)
Temperate		
Intercept	3.361396*** (0.9733)	0.641326*** (0.0800)
Time Trend	0.003898** (0.0016)	-0.000001 (0.0005)
Lag	0.801928*** (0.0574)	0.154281 (0.1027)
Warm		
Intercept	8.793137*** (1.8607)	0.672060*** (0.0893)
Time Trend	0.001561 (0.0012)	0.000517 (0.0006)
Lag	0.624611*** (0.0794)	0.245836** (0.0977)

Notes: Temperature is the average of the daily temperature distribution in the relevant climate. Precipitation is the average the total level of agricultural-season precipitation across the locations composing the relevant climate. The time trend is measured in years from 1950. The data used in estimation come from [Livneh et al. \(2015\)](#) for years 1950-1979 and [Thornton et al. \(2022\)](#) for years 1980-2020.

Table B6: Comparisons for Identification

	$\mathbb{1}\{\text{Migrate}\}$	Length of Stay
Previous Migrant	0.05262*** (0.00086)	-0.64*** (0.23)
$\mathbb{1}\{\text{Origin is Birth Location}\}$	-0.04401*** (0.00092)	-0.53** (0.22)
Predicted Wage	-0.000185*** (0.000049)	0.163*** (0.017)
$\mathbb{1}\{\text{Origin is Urban}\}$	-0.00083 (0.00070)	1.50*** (0.22)
Age	-0.000929*** (0.000036)	-0.108*** (0.017)
Degree days above 26°C	0.00418*** (0.00066)	0.48** (0.23)
Degree days above 26°C ²	-0.00068*** (0.00013)	-0.043 (0.048)
Degree Days below 14°C	0.0016*** (0.0004)	0.58*** (0.13)
Degree Days below 14°C ²	-0.000271*** (0.000046)	-0.024 (0.015)
Num.Obs.	260472	6012
R2	0.043	0.069

Column 1 represents a linear probability model for individual migration decisions, using the MMP data. Column 2 represents the OLS estimates for the number of years of a given move, recorded in the MMP data, as a function of the characteristics of the first year of that move. *p < 0.1;
p < 0.05; *p < 0.01

Table B7: Expected Moving Costs

	Expected Cost		
	Average	Minimum	Maximum
Overall	-50,194	-247,082	152,622
Towards Birth Location	29,167	-32,009	152,622
Away From Birth Location	-96,618	-247,082	-33,453
Number of Moves	4,262		

Notes: Costs are measured in 2024 dollars. The expected cost is defined in Appendix B.4. Negative costs represent net positive expected contemporaneous utility.

C Robustness & Additional Results

Additional Model Fit Statistics

Table C1 shows the data and model-predicted moments discussed in the main text, Figures C1 and C2 replicate the relevant figures in the main text, when simulating the entire life-history of each individual in the estimation data.

Table C1: Model Fit - Full Lifecycle

	Migration Rate	
	Model	Data
Overall	4.1%	4.7%
Children		
Yes	3.9%	3.4%
No	4.1%	7.2%
Agricultural Worker		
Yes	3.5%	3.5%
No	4.3%	5.3%
Years of Education		
0-5	3.8%	4.2%
6-11	4.0%	3.8%
12+	4.6%	7.5%

Note: For each category, I calculate the average migration predicted by 100 model simulations of the life trajectories of the individual-year observations used in estimation. The data migration rates are calculated as the average migration rate for each category.

Additional Results

Table C4 presents the estimated preference and moving cost parameters when allowing for a different moving cost for agricultural workers. Figure C3 present the value of migration for 17–19-year-olds, broken apart by climate, urban/rural location of birth, education level, and sector of employment. Figure C4 presents the climate damages for the same group, broken apart by the same characteristics.

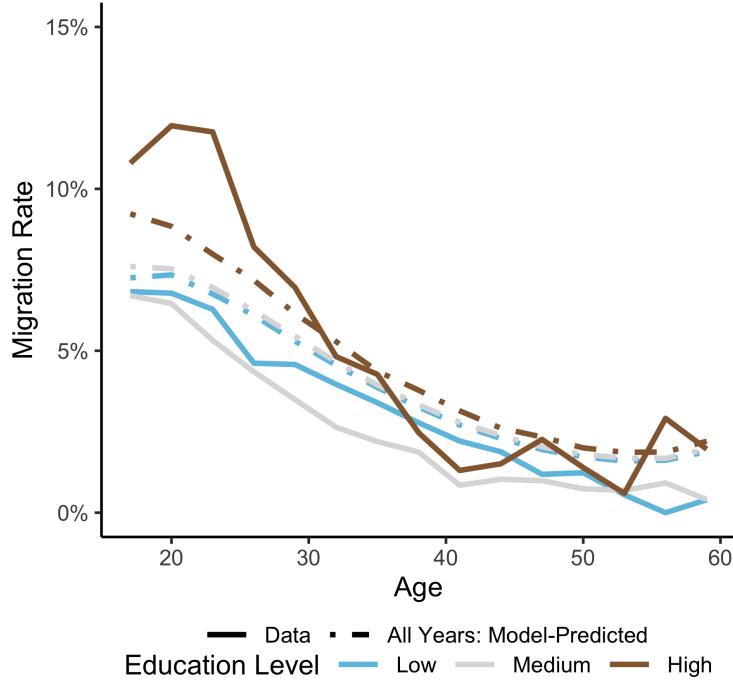


Figure C1: Model-Predicted Migration Rates by Education

Notes: This figure plots the average migration rates in the data against those predicted by the model, broken apart by education levels. It represents the model-simulations for the entire lifecycle for all individuals used in estimation. Low, mid, and high-education refers to less than 6 years, 6–11 years, and 12 or more years of education, respectively

Additional Policy Results

I discuss here the different effects on migration across the age distribution from the dynamic structure of policy-induced reductions in moving costs. I study the effectiveness of a 1% reduction in moving costs, broken apart across the age distribution. Figure C9 plots the results of this exercise. The solid curve represents the change in the ex ante probability than an individual migrates in 2020 from a permanent 1% reduction in moving costs. The dashed curve plots the same for a one-time 1% reduction in moving costs; that is, a policy which reduces moving costs by 1% in 2020, but for which all future costs would return to the baseline levels.

I find that an individual's relative responsiveness to the two schemes changes throughout their life-cycle. This is because of two conflicting forces. On the one hand, individuals are more likely to move when they know that they will be able to leave their destination at a reduced cost in the future. On the other hand, individuals understand that the cost of future moves will be higher than they are today, simply because they will be older in the future than they are today. Under the one-time subsidy, moving costs will be even higher in the future, relative to the year of the subsidy. The same is not true for the permanent subsidy.

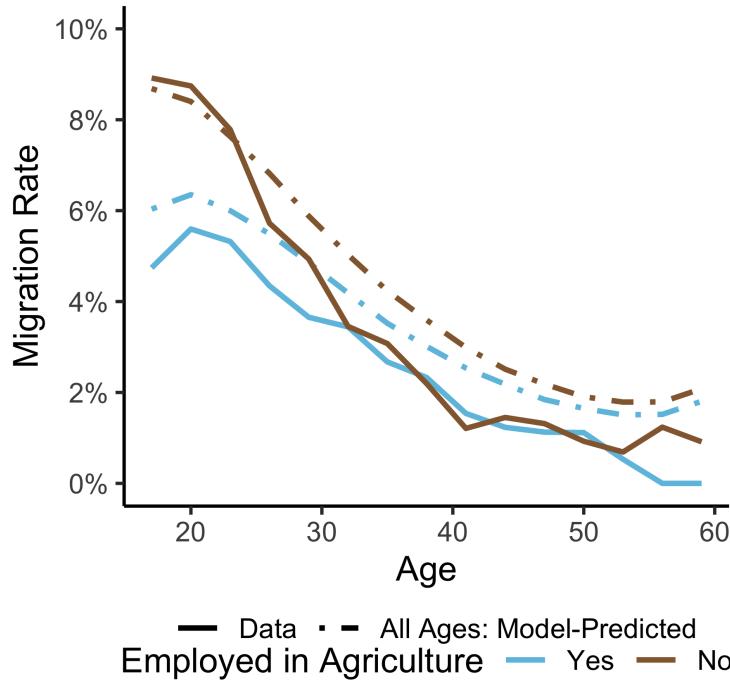


Figure C2: Model-Predicted Migration Rates by Sector of Employment

Notes: This figure plots the average migration rates in the data against those predicted by the model, broken apart by sector of employment. It represents the model-simulations for the entire lifecycle for all individuals used in estimation. Low, mid, and high-education refers to less than 6 years, 6–11 years, and 12 or more years of education, respectively

If individuals are planning on making a move at some point in their lives, this would lead to a higher proportion of migrants under the one-time subsidy than under the permanent subsidy. At the beginning of an individual's life, the first force dominates the second, leading to a higher rate of migration under the permanent subsidy. However, as individuals age, the probability that they will make more than one additional move in their lifetime falls, as signified by the decreasing difference between the two curves. Eventually, the second force dominates the first, leading to a higher rate of migration under the one-time subsidy.

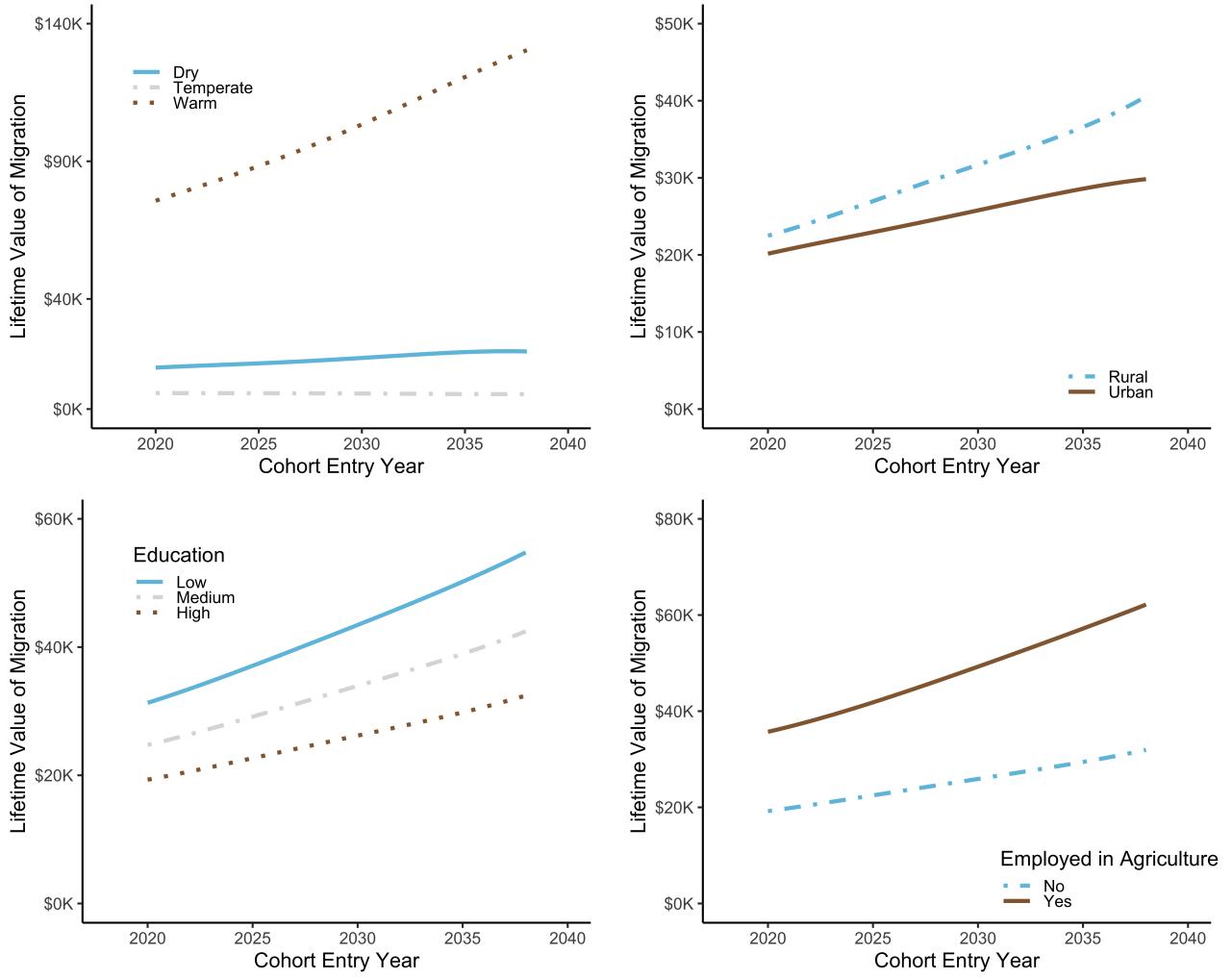


Figure C3: This figure plots the average value of migration for 17–19-year-olds, from 2020 through 2038. (a) breaks this value apart by climate of birth, (b) by urban/rural status of the location of birth, (c) by education level, and (d) by sector of employment

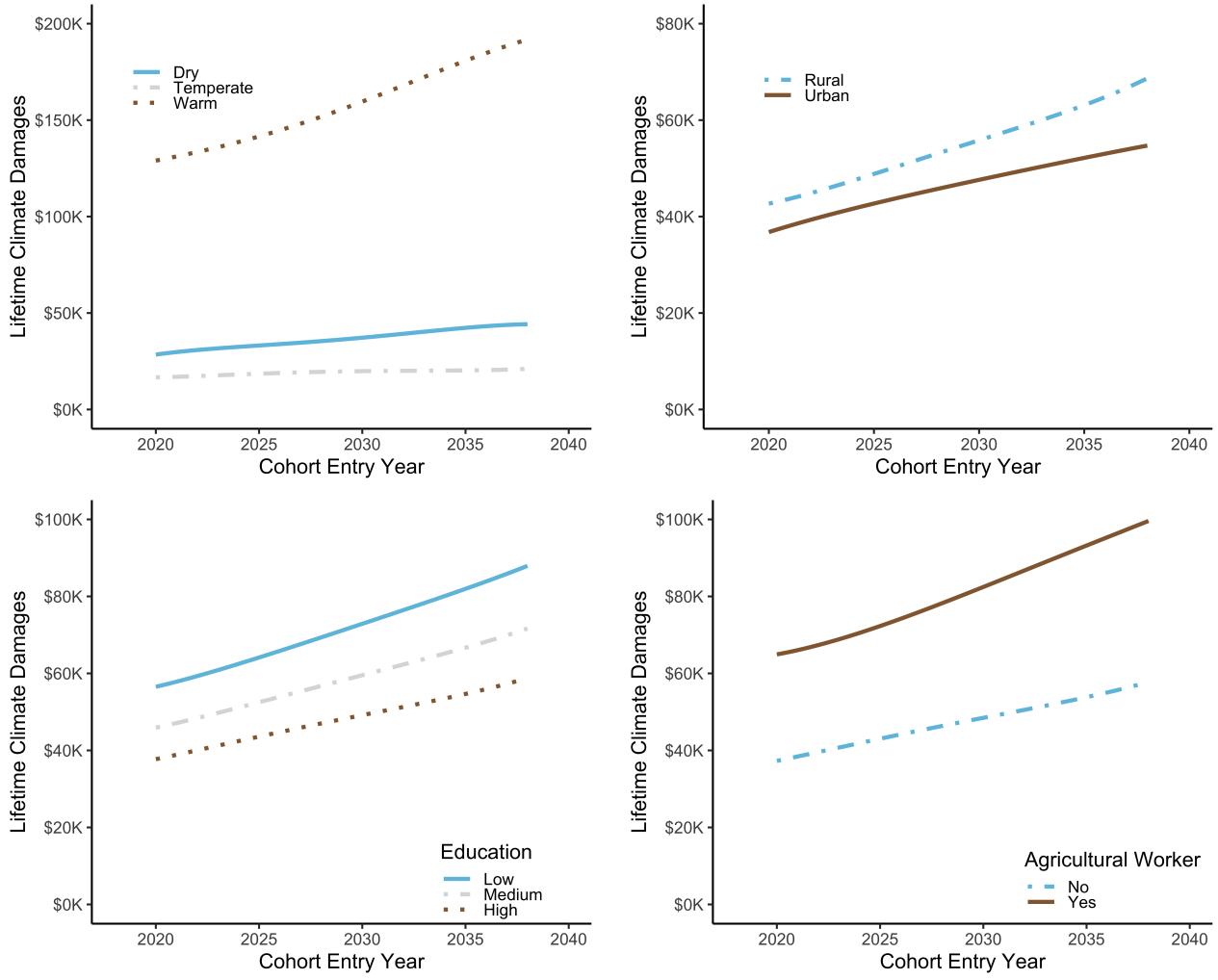


Figure C4: This figure plots the average level of climate damages for 17–19-year-olds, from 2020 through 2038. (a) breaks this value apart by climate of birth, (b) by urban/rural status of the location of birth, (c) by education level, and (d) by sector of employment

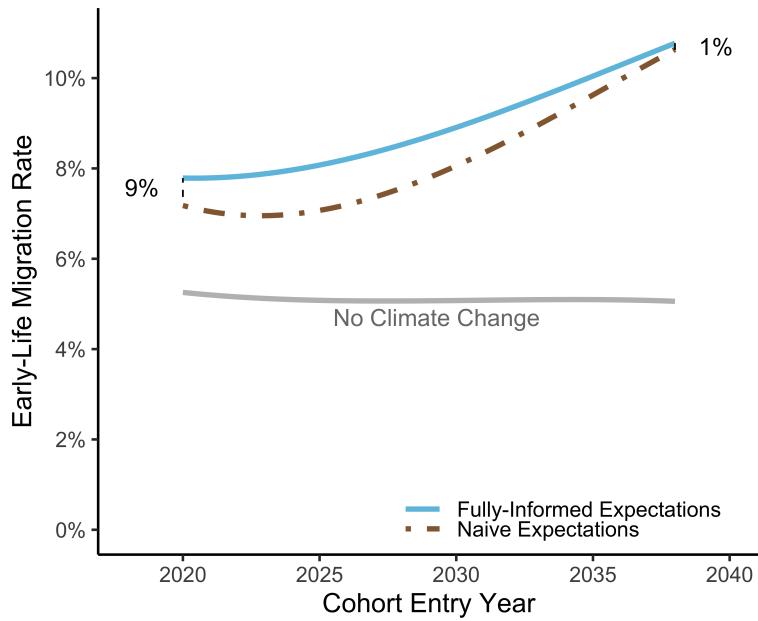


Figure C5: Migration rates across time and beliefs

Notes: This figure plots model-predicted migration rates, for 17–19-year-olds by time. The solid curve represents the migration rate for fully-informed individuals; the dashed curve for climate-naive individuals, and the light solid curve the average migration rate in the limited climate change scenario. The results are averaged over 100 simulations.

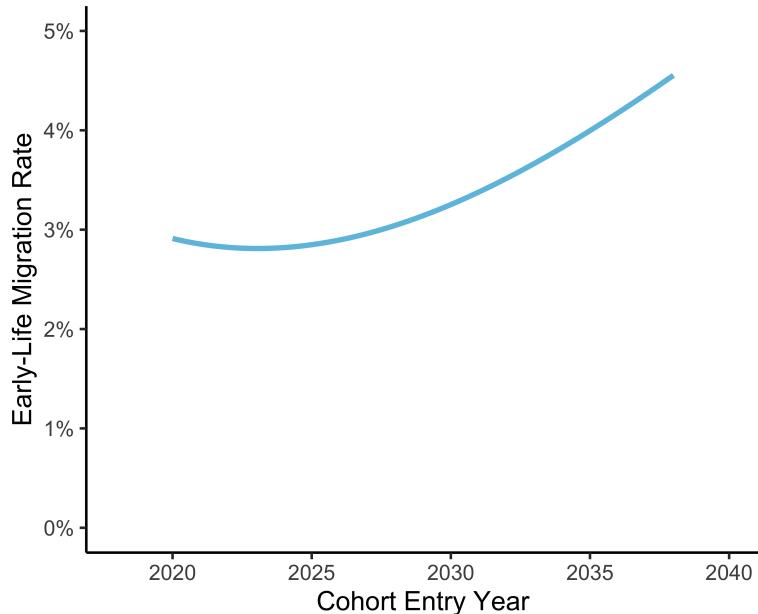


Figure C6: Migration rates from static logit

Notes: This figure plots the static logit model-predicted migration rates, for 17–19-year-olds by time, under a business-as-usual warming scenario. The results are averaged over 100 simulations.

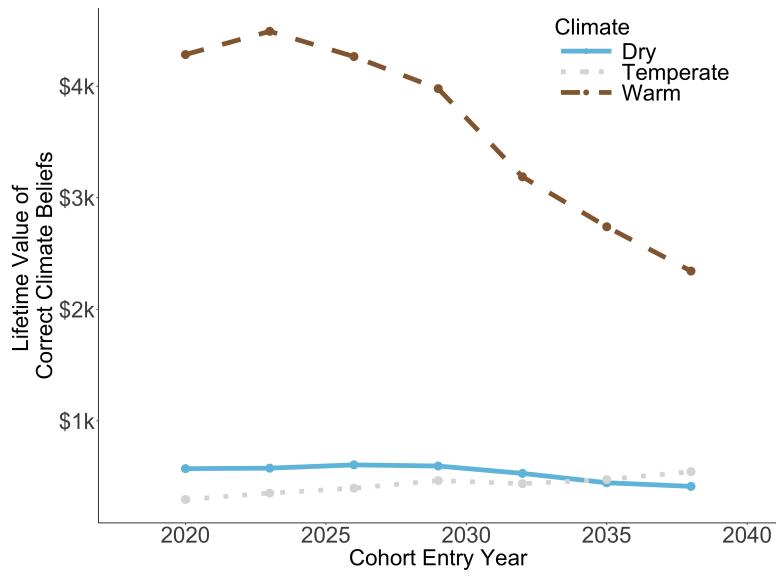


Figure C7: Value of information to climate-naive

Notes: This figure plots the value of becoming fully-informed, to the climate-naive, over time and space. The solid curve represents the value to those living in dry location, the dashed curve in warm locations, and the dotted curve in temperate locations. Results are averaged over 100 simulations.

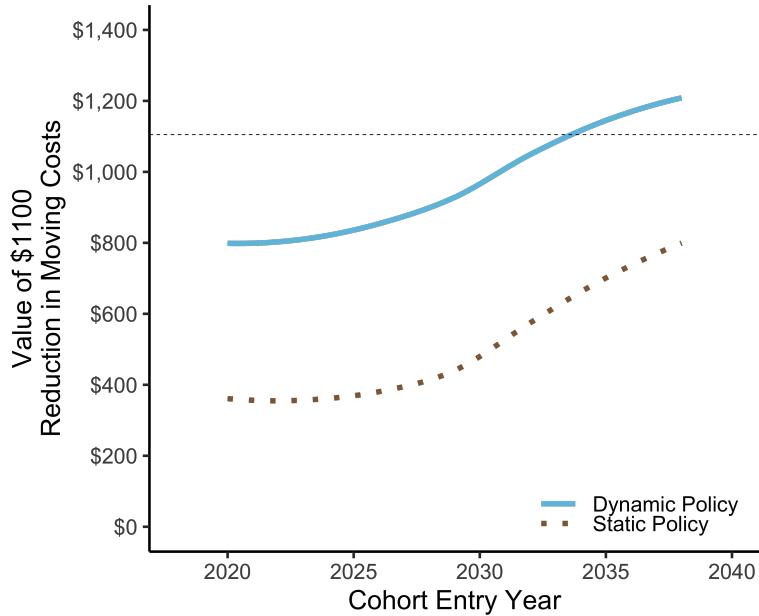


Figure C8: Value of migration subsidy, across time and policy-structure

Notes: This figure plots the value, to 17–19-year-old movers, of \$1,100 migration subsidies, broken apart by the dynamic structure of the policy. The solid curve represents the value of a dynamic policy which is used at the first time an individual decides to move. The dotted curve represents the value of a static policy which is only available in the given year. The dashed line represents the level of the subsidy. Results are averaged over 100 simulations.

Table C2: Static Logit Parameters: Forward Looking Evidence

	Flow Utility	Moving Costs	
Income	0.063 (0.006)	Moving Intercept (0.095)	2.911
Degree Days above 26°C	0.023 (0.042)	Distance	0.136 (0.014)
Degree Days above 26°C ²	-0.070 (0.014)	Children	0.128 (0.044)
Degree Days below 14°C	0.148 (0.014)	Age	0.209 (0.009)
Degree Days below 14°C ²	-0.024 (0.002)		
Average Temperature Next Decade	0.019 (0.006)		
Urban Location	-0.349 (0.027)		
Birth Location	2.450 (0.022)		

Log-Likelihood: -25255.83

Note: Income is measured in annual hourly 2010 pesos, degree days are divided by 100, distance is measured in log kilometers. Average temperature over the next decade refers to the mean of the daily temperature distribution at the level-one Köppen climate level.

Future temperatures are drawn from the historical data through 2020 and the business-as-usual simulated weather data for 2020–2030. Standard errors are computed with the inverse Hessian.

Table C3: Supplemental Values of Migration

	Value of Migration		
	Only Wages	Only Amenities	General Equilibrium
Average	5,773	21,271	19,205
Climate			
Dry	4,947	16,587	13,660
Temperate	6,306	5,909	5,363
Warm	5,336	73,963	68,544
Agricultural Worker			
Yes	5,851	34,604	32,652
No	5,851	19,504	17,423
Years of Education			
0–5	5,501	30,530	29,527
6–11	5,682	24,649	23,136
12+	5,821	19,555	17,232

Note: These are the average results of 10 simulations of the model. Values are for 17-19 year olds forming fully-informed climate expectations, in 2020, measured in 2024 dollars. The value of migration is defined as the difference between average lifetime welfare for an individual who can migrate and one who cannot. All values are for a business-as-usual climate scenario: the SSP2-4.5 climate projections. General equilibrium refers to the business-as-usual scenario with a congestion cost in the income equation, iterated to convergence. The amenity impact is from a counterfactual where weather does not impact wages. The wage impact is from a counterfactual where weather does not impact amenities.

Table C4: Utility Parameters, Agricultural Movers

	Flow Utility	Moving Costs	
Income	0.006 (0.0007)	Moving Intercept	2.781 (0.074)
Degree Days above 26°C	0.023 (0.005)	Distance	0.109 (0.011)
Degree Days above 26°C ²	-0.010 (0.002)	Children	0.091 (0.030)
Degree Days below 14°C	0.021 (0.002)	Age	0.172 (0.006)
Degree Days below 14°C	-0.004 (0.0003)	Agricultural Worker	0.078 (0.027)
Urban Location	-0.041 (0.003)		
Birth Location	0.179 (0.002)		
Mass of Movers	0.627 (0.013)		
Log-Likelihood: -24866.35			

Note: Income is measured in annual hourly 2010 pesos. Degree days are divided by 100, and measured in the destination location. Distance is measured in log kilometers. The likelihood contains 90,578 individual-year observations from 11,194 individuals. Standard errors in parentheses.

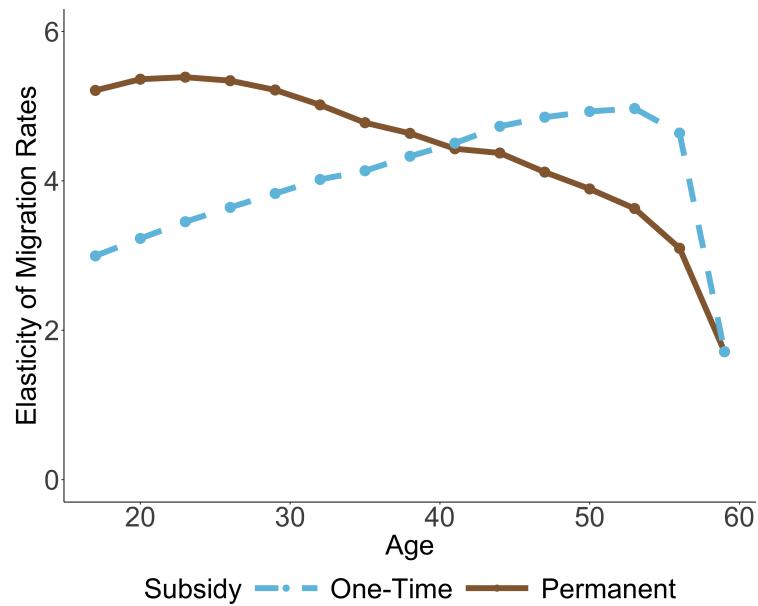


Figure C9: Responsiveness of Migration Rates to Migration Subsidies

Notes: This figure plots the percent increase in migration rates from a 1% decrease in migration costs, across the age distribution and duration of the subsidy. The solid curve represents a permanent subsidy, while the dashed curve represents a one-time subsidy.

D Extensions

Within-Location Across-Municipality Moves

Table C5: Utility Parameters: Within Location Moves

Flow Utility		Moving Costs	
Income	0.006 (0.0006)	Moving Intercept	3.185 (0.066)
Degree Days above 26°C	0.059 (0.005)	Within Location	4.486 (0.025)
Degree Days above 26°C ²	-0.023 (0.002)	Distance	0.166 (0.010)
Degree Days below 14°C	0.027 (0.002)	Children	0.375 (0.283)
Degree Days below 14°C ²	-0.004 (0.0003)	Age	0.226 (0.050)
Urban Location	-0.043 (0.003)		
Birth Location	0.214 (0.002)		
Mass of Movers	0.641 (0.013)		
Mass of Believers	0.766 (0.176)		
Log-Likelihood: -30858.60			

Note: Income is measured in annual hourly 2010 pesos. Degree days are divided by 100 and measured in the destination location. Distance is measured in log kilometers. The likelihood contains 90,578 individual-year observations from 11,194 individuals. Standard errors in parentheses.

Table C5 presents the estimated preference and moving cost parameters when additionally modeling the decision to move across municipalities but within the same location - that is, individuals choose over 28 options: 26 across location moves, 1 within location move, and the option to remain in the same municipality-location. In particular, I modify the baseline moving costs as follows:

$$\tilde{c}(\ell, \ell', \omega) = \tilde{\gamma}_1 \omega + \tilde{\gamma}_2 \mathbb{1}\{\ell = \ell' \wedge \text{move}\} + \tilde{\gamma}_3 d(\ell, \ell') + \tilde{\gamma}_4 \mathbb{1}\{k_\omega \geq 1\} + \tilde{\gamma}_5 \text{age}_\omega,$$

where, now, $\tilde{\gamma}_2$ represents the within location moving cost. $\tilde{\gamma}_3$ is identified only from moves

across locations, and $\tilde{\gamma}_4$ and $\tilde{\gamma}_5$ are identified from both within and across location moves. The estimated moving costs are higher than in the baseline specification, but the preference parameters are quite similar. The within location moving cost is quite high. This is because most moves are long-distance in the estimation data.

Intensification effects of climate change

The model discussed in the primary section of this paper assumes that migration frictions in the future will resemble those of today. This is an assumption that is ubiquitous in the literature. Past work has pointed out that the role that migration (or, more broadly, adaptation) plays in limiting future climate damages will depend on whether intensification effects of climate change outweigh technological progress related to migration costs. In this section, I present a modified version of the model to account for these conflicting forces.

Throughout the duration of the MMP survey used in estimation, 1960s-2010s, Mexico experienced significant technological change as well as a noticeable rightward shift in its climate distribution. I modify the moving cost function to account for potential differences in the moving cost arising because of differences in local agricultural weather ([Cattaneo and Peri, 2016](#)). The baseline model remains identical to that presented in Section 3, with the exception that now the moving cost additionally is allowed to depend on a quadratic in the level of local growing degree days in the origin. Mathematically,

$$\bar{c}(\ell, \ell', \omega) = \gamma_1\omega + \gamma_2GDD(\ell, \omega) + \gamma_3GDD(\ell, \omega)^2 + \gamma_4d(\ell, \ell') + \gamma_5\mathbb{1}\{k_\omega \geq 1\} + \gamma_6age_\omega,$$

Identification of γ_2 and γ_3 relies on differences in migration rates away from areas with different levels of growing degree days and on the choices of where individuals move to, acknowledging that future moving costs will be a function of expected future growing degree days in the destination today. I remain agnostic as to the sign and magnitude of the linear and quadratic terms in the moving cost specification. In particular, one could imagine that a good agricultural harvest, as proxied for by growing degree days, could lead to either increased migration because of increased incomes (e.g., [Liu et al., 2023](#)) or decreased migration due to a lower need to supplement current incomes. Given that liquidity constraints enter into the model through the moving cost, it is also possible that sufficiently poor agricultural outcomes, again, as proxied for by growing degree days may be associated with lower migration rates.

Table C6 presents the estimated preference parameters from this updated model specification. Similar to the baseline model for the amenity values of extreme heat, this model allows for a bliss point interpretation of the coefficients for growing degree days in the moving costs. The model estimated bliss point is 2,463 growing degree days, with a marginal degree day

Table C6: Utility Parameters, Intensification

Flow Utility		Moving Costs	
Income	0.006 (0.0007)	Moving Intercept	3.018 (0.118)
Degree Days above 26°C	0.040 (0.005)	Growing Degree Days	-0.202 (0.084)
Degree Days above 26°C ²	-0.016 (0.002)	Growing Degree Days ²	0.041 (0.016)
Degree Days below 14°C	0.023 (0.002)	Distance	0.111 (0.011)
Degree Days below 14°C ²	-0.004 (0.0003)	Children	0.072 (0.028)
Urban Location	-0.040 (0.003)	Age	0.172 (0.006)
Birth Location	0.183 (0.002)		
Mass of Movers	0.629 (0.013)		
Log-Likelihood: -24859.345			

Note: Income is measured in annual hourly 2010 pesos. Degree days are divided by 100 and measured in the destination location. Growing degree days have a lower discontinuity at 8 degrees Celsius and an upper discontinuity at 32 degrees Celsius, are divided by 1000, and are measured in the origin location. Distance is measured in log kilometers. The likelihood contains 90,578 individual-year observations from 11,194 individuals. Standard errors in parentheses.

below this value reducing the migration cost and one above it increasing this cost. 40% of the data used in estimation experienced growing degree days above this threshold; 30% of those living in dry locations, 15% in temperate locations, and 86% in warm locations. If these same individuals were transplanted to 2050, the share of individuals living in dry locations who experience growing degree days beyond this threshold would increase to 97%, the share living in temperate locations to 45%, and the share living in warm locations to 100%.

I find that accounting for this mechanism leads to a 10% increase in lifetime climate damages experienced by the average 17–19-year-old in 2020. Those living in temperate climates stand to relatively lose the most - with an increase in lifetime damages of 27%, while those living in warm climates experience an 8% increase, and those living in dry climates experience an 5% increase. Individuals living in temperate climates face the highest increase to climate damages because of their destinations. Particularly, some of those living in these locations decide to move to locations in the dry and warm regions of the country, which

experience the brunt of climate change. The model now separately incorporates the effect of climate change through reductions in their lifetime value of moving to these places, but also through the increased cost of moving away. This causes their lifetime damages to increase the most, relative to the modest baseline damages. On the other hand, I find that climate damages in the warm parts of the country increase to 40% of lifetime value, or a 3 percentage point increase on the baseline. This is because the model accurately captures that not only are they unhappy in their current location, but they are also less likely to be able to leave.

E Simulations

E.1 Census Data

The counterfactual simulations use data on the entire male population living in Mexico, as of 2020. I discuss here how I use publicly available Census data to create an individual-level dataset to simulate population movements.

I use the 2020 full-count Census (*El Censo de Población y Vivienda 2020*) to recover municipality-level distributions of education, age, and agricultural employment. More specifically, I use the *características económicas* module to recover education levels in each municipality. These data contain information on the count of individuals, by sex, in each of the following education categories: no schooling, complete primary education, incomplete secondary education, complete secondary education, complete technical education, completed upper secondary, and completed higher education. I define low education as those with no schooling (less than 6 years), medium education as those with complete primary or incomplete secondary education (6–11 years), and high education as those with complete secondary or higher education (12 or more years). When forming the municipality-level education distribution, I subset to only include males.

The above survey is based on the basic questionnaire, which is administered to all individuals living in Mexico. The basic survey does not contain information on the sector of employment. For this reason, I additionally use the extended questionnaire (*Tabulados del Cuestionario Apliado*) to recover the proportion of each municipality's population that works in agriculture, livestock, forestry, fishing, and hunting. I use this proportion to define agricultural employment. A drawback of this extended questionnaire is that it only covers the population living in private households; it misses all individuals living in collective housing, the homeless, and those in the Foreign Service. Moreover, there are three municipalities in Mexico that are not covered by the extended questionnaire. I assume that their agricultural employment rates are the same as the weighted average of the other municipalities in their

assigned location.

I use the Características de la Población module to recover raw population counts, broken apart by sex and age, at the municipality level. I subset these data to only males who are less than 62 years old (the terminal age group in the model). I keep all children younger than 17 in the data, as I will bring these individuals into the model as they come of age.

The above three data sources are merged to create a set of municipality-level state-variable distributions and population-by-age counts. I use the Köppen maps, discussed in Appendix A.3, and the metropolitan-zones, discussed in Appendix A.4 to aggregate this information to the 27 model-locations. In particular, I sum population-ages across each municipality in a given location, and use the total population, at the municipality level, to weight the education and agricultural employment distributions.

This framework allows me to create an individual-level dataset. To do so, for each location and age group, I draw the number of individuals as given by that location-age grouping. I assign their agricultural employment status from a Bernoulli draw with a location-specific success probability. I then assign their education level from a multinomial draw with a location-specific probability vector. Since the census data does not contain information on individual birth municipalities, I assume that everyone in the data currently resides in their birth location. Finally, I assign children to parents using the population counts below 17 and fertility rates (number of children by age) from the MMP data. I hold the resulting dataset fixed for the counterfactual simulations, across both counterfactual scenarios and within scenarios and across iterations.

Table 1, in the main text, presents statistics on the relevant variables across the 2020 census and the MMP data.

E.2 Simulation Framework

I use the individual-level dataset described above to simulate the migration choices of the entire male population through 2038 under a variety of counterfactual scenarios. Within each counterfactual presented, I simulate behavior for 100 iterations. Within each iteration, I draw new shocks from:

- Moving types, from a Bernoulli distribution with probability given it's MLE estimate
- Belief types, from a Bernoulli distribution with probability given it's MLE estimate
- Weather shocks across the three primary Köppen climates for the six weather states, from a mean-zero Gaussian with covariance given by that from the historical weather data

- Preference shocks for each individual-year, from a T1EV distribution
- Fertility shocks for each individual-year that is not a parent, from a standard normal

These shocks, along with the individual-level data described above, completely characterize the state of each simulated person. I use the MLE parameter estimates to solve the relevant Bellman equations for each individual to simulate forward their life choices.

E.3 Counterfactual Scenarios

As discussed in the paper, I model a business-as-usual climate scenario as one following the SSP2-4.5 scenario. This scenario is characterized by a moderate degree of warming under current climate policies. On the other hand, I model a scenario of no *further* climate change as a climate centered around recent historical averages, those from 1950 to 1980. Table D1 presents the estimates of ρ for the business-as-usual scenario, averaged across the entire period.⁴⁹

E.4 General Equilibrium

This paper models migration in a partial equilibrium setting. One potential concern is that general equilibrium forces will respond to shifts in populations through space through a suppression of the wage level in desirable locations. Some percent of the population modeled as migrating will sufficiently marginal in their decision such that this reduction would be sufficient to reverse their decision; either causing them to move to a different location or not to move at all. This would imply that both my estimated value of migration, my migration rate elasticity would be biased upwards, and the estimated percent reduction in climate damages because of migration, would be biased upwards. I find that these forces are relevant, but that the primary conclusions of the paper remain outstanding after accounting for the described congestion impacts.

Before describing the relevant algorithm, I first discuss how such impacts should enter into the model. In particular, define ξ as the elasticity of wages w with respect to labor supply Q . Then if Q_ℓ rises by $x\%$, w_ℓ will increase by $x\xi\%$. I calibrate $\xi = -0.3$ (Borjas, 2003).

I iterate on the following algorithm until population shares converge:

1. Simulate migration decisions with the estimated model and wage equation through 2038

⁴⁹Each year of the simulation has its own set of coefficients corresponding to the previous 30 years of weather outcomes. I present the coefficients when estimating the transition across the entire simulation period. These are not the coefficients I use in the counterfactuals.

2. Calculate the difference in population levels from 2023 to 2038 throughout Mexico
3. Assume that wages respond to this change in population⁵⁰
4. Resimulate behavior under the new wage regime

For example, in the first step of this process I use the estimated model and the setup described in Section 6 to simulate the behavior of each male in Mexico, beginning in 2020 and ending in 2038. This allows me to calculate the percent change in populations between 2023 and 2038 for each of the locations in the choice set.⁵¹ I use these changes in populations to calculate updated wages, where the new wage is calculated according to the change in populations and the calibrated elasticity. I then simulate behavior according to this update in wages, and repeat steps 2-4 until the maximum difference in population differences, across locations, between iterations, is sufficiently small.

More formally, define $m^{(k)}(t) \in \mathbb{R}^{27 \times 27}$ as the migration matrix at year t in iteration k . Then, $m^{(k)}(t)_{ij} = \int \mathbb{1}\{\ell = i \cap \ell' = j\} dP(k)$, where $P(k)$ is the distribution of individuals at iteration k . Let $n^{(k)}(t) \in \mathbb{R}^{27}$ be the population vector at year t in iteration k . Then, $n_i^{(k)}(t) = \sum_j m_{ij}^{(k)}(t)$. Finally, define $a^{(k)} \in \mathbb{R}^{27}$ as the percentage change in populations from 2023 to 2038 in iteration k . Then, $a_i^{(k)} = \frac{n_i^{(k)}(2038) - n_i^{(k)}(2023)}{n_i^{(k)}(2023)}$. This algorithm converges when $d(a^{(k)}, a^{(k-1)}) < \varepsilon$, where $n^{(k)}$ comes from simulated population movements, as described in Appendix E, with wages updated such that $w^{(k)} = w^{(k-1)} * (1 + \xi a(k-1))$. A sufficient condition for convergence is that $\frac{\partial a_p^{(k+1)}}{\partial a_j^{(k)}} \rightarrow 0 \forall p, j$. If all marginal movers across simulations decide to either move to their initial destination or to not move at all, then this condition will be met. In reality, most chosen destinations are close substitutes to other locations within the choice set. This implies that the sufficient condition is unlikely to hold in practice.

⁵⁰In particular, I assume that wages respond linearly to this change in populations up to 2038, and then remain at 2038 levels.

⁵¹Beginning in 2023 only aids in speeding up convergence of this algorithm. Similar results are drawn when using 2020 as the base year.

Table D1: Average First Stage SSP2–4.5 Climate Transitions

	Temperature	Precipitation
Dry		
Intercept	16.264189*** (3.6974)	0.291821*** (0.0588)
Time Trend	0.028378*** (0.0069)	-0.000175 (0.0007)
Lag	0.199050 (0.1824)	0.142953 (0.1732)
Temperate		
Intercept	16.632479*** (2.9162)	0.989206*** (0.1339)
Time Trend	0.031055*** (0.0063)	-0.002027 (0.0013)
Lag	0.112766 (0.1559)	-0.159349 (0.1550)
Warm		
Intercept	18.750155*** (2.5778)	0.996126*** (0.1177)
Time Trend	0.018471*** (0.0028)	-0.001678*** (0.0004)
Lag	0.238298** (0.1049)	0.087531 (0.1069)

Notes: Temperature is the average of the daily temperature distribution in the relevant climate. Precipitation is the average the total level of agricultural-season precipitation across the locations composing the relevant climate. The time trend is measured in years from 2020. The data used in estimation come from [Thrasher et al. \(2022\)](#), are subsetted to years 2020–2050.