

# Math 522 Hw1

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January 22, 2025

1. Let

$$\xi_n = \begin{cases} -1 & \text{w.p. } 1 - \frac{1}{n^2} \\ n^2 - 1 & \text{w.p. } \frac{1}{n^2} \end{cases}$$

be independent. Then clearly  $\mathbb{E}[\xi_n] = 0$ . With  $X_n = \xi_1 + \cdots + \xi_n$ , we know that  $X_n$  is a martingale. I claim that  $\mathbb{P}(\limsup_{n \rightarrow \infty} X_n = -\infty) \geq 1/2$ . This is because  $\limsup_{n \rightarrow \infty} X_n = -\infty$  occurs certainly whenever all the  $\xi_n$  are  $-1$ . Thus, by independence:

$$\mathbb{P}\left(\limsup_{n \rightarrow \infty} X_n = -\infty\right) \geq \mathbb{P}\left(\bigcap_{n=1}^{\infty} \{\xi_n = -1\}\right) = \prod_{n=1}^{\infty} \left(1 - \frac{1}{n^2}\right) = \frac{1}{2}$$

Now the event  $\{\limsup X_n = -\infty\}$  is tail measurable, since if we throw out the first  $m$  many terms, the limsup will still be  $-\infty$ . Thus, by Kolmogorov 0-1 law, we know that  $\mathbb{P}(\limsup X_n = -\infty) = 0$  or  $1$ . But we have shown that  $\mathbb{P}(\limsup X_n = -\infty) \geq 1/2$ , so we must have  $\mathbb{P}(\limsup X_n = -\infty) = 1$ .

2. We know that  $X_n = \prod_{m \leq n} Y_m$  is a non-negative martingale, thus by Martingale convergence theorem,  $X_n \rightarrow X$  a.s. with  $X$  finite a.s., as  $\mathbb{E}[X] \leq \mathbb{E}[X_0] = 1$ . In particular,  $X_n$  is Cauchy in the following sense:

$$\mathbb{P}(|X_{n+1} - X_n| > \eta) \rightarrow 0$$

This is because  $\eta \leq |X_{n+1} - X_n|$  implies  $\eta \leq |X_n - X| + |X_{n+1} - X|$  which in turn implies  $\eta/2 \leq |X_n - X|$  or  $\eta/2 \leq |X_{n+1} - X|$ . Since  $X_n \rightarrow X$  almost surely, it converges in probability, so it is easy to see that those last 2 events converge to 0 as  $n \rightarrow \infty$ . Thus

the first event must also converge to 0. Since  $\mathbb{P}(Y_m = 1) < 1$ , and  $\mathbb{E}[Y_m] = 1$ , we can find some  $\delta > 0$  so that  $\mathbb{P}(Y_m > 1 - \delta) > 0$ . This is because otherwise  $Y_m \geq 1$  almost surely, and  $\mathbb{P}(Y_m > 1) > 0$ , so  $Y_m - 1$  is a non-negative r.v. with positive expectation, a contradiction.

3. First, we prove that if  $y_n > -1$  for all  $n$  and  $\sum |y_n| < \infty$ , then  $\prod_{m=1}^{\infty} (1 + y_m)$  exists. By Taylor's theorem for remainders,  $\log(1 + x) = 0 + \frac{1}{1+\zeta}x$  for some  $\zeta \in (x, 0)$  for  $x < 0$  and  $(0, x)$  for  $x > 0$ . Since  $|y_n| \rightarrow 0$  since the series converges, we know that  $|y_n| < 1/2$  eventually, and since  $y_n > -1$ , this means that  $y_n > -1/2$  eventually. For these values, we then have  $|\log(1 + y)| \leq 2|y|$  by the calculation above. Thus, for  $n$  sufficiently large:

$$\left| \sum_{m \geq n} \log(1 + y_m) \right| \leq 2 \sum_{m \geq n} |y_m| \rightarrow 0$$

So  $\sum \log(1 + y_m)$  converges as it is cauchy, so  $\exp(\sum \log(1 + y_m)) = \prod (1 + y_m)$  converges.

Now define a new r.v.

$$Z_n = \frac{X_n}{\prod_{m \leq n-1} (1 + Y_m)}$$

where the empty product has value 1. Then:

$$\mathbb{E}(Z_{n+1} | \mathcal{F}_n) = \frac{1}{\prod_{m \leq n} (1 + Y_m)} \mathbb{E}(X_{n+1} | \mathcal{F}_n) \leq \frac{1}{\prod_{m \leq n} (1 + Y_m)} (1 + Y_n) X_n = \frac{X_n}{\prod_{m \leq n-1} (1 + Y_m)} = Z_n$$

so  $Z_n$  is a non-negative supermartingale, thus it converges almost surely. Since  $\sum Y_n \leq \infty$  a.s., the product  $\prod_{m \leq n} (1 + Y_m)$  converges a.s. so  $Z_n \cdot \prod_{m \leq n} (1 + Y_m) = X_n$  converges a.s. which completes the proof.

4. Let  $X_n^1, X_n^2$  be supermartingales adapted to  $\mathcal{F}_n$  and let  $N = \inf \{m : X_m^1 \geq X_m^2\}$ . Then  $N$  is a stopping time, and let:

$$\begin{aligned} Y_n &= X_n^1 1_{N > n} + X_n^2 1_{N \leq n} \\ Z_n &= X_n^1 1_{N \geq n} + X_n^2 1_{N < n} \end{aligned}$$

We claim that  $Y_n, Z_n$  are supermartingales. First, we show that  $Y_n \leq Z_n$  everywhere. This is because when  $N > n$ ,  $Y_n = X_n^1$  and  $Z_n = X_n^1$ , when  $N < n$ ,  $Y_n = X_n^2$  and  $Z_n = X_n^2$ , and when  $N = n$ ,  $X_n^1 \geq X_n^2$  by definition of  $N$ , and  $Y_n = X_n^2$  while  $Z_n = X_n^1$ , so  $Y_n \leq Z_n$

in all cases.

Now,

$$\mathbb{E}(Y_{n+1} | \mathcal{F}_n) \leq \mathbb{E}(Z_{n+1} | \mathcal{F}_n) = \mathbb{E}(X_{n+1}^1 1_{N \geq n+1} + X_{n+1}^2 1_{N < n+1} | \mathcal{F}_n)$$

From here,  $1_{N \geq n+1}$  and  $1_{N < n+1}$  are  $\mathcal{F}_n$  measurable, since the first event is the complement of  $N \leq n$  and the second is equal to  $N \leq n$ . Thus we can take them out of the conditional expectation, getting:

$$\mathbb{E}(X_{n+1}^1 | \mathcal{F}_n) 1_{N \geq n+1} + \mathbb{E}(X_{n+1}^2 | \mathcal{F}_n) 1_{N < n+1} \leq X_n^1 1_{N > n} + X_n^2 1_{N \leq n} = Y_n \leq Z_n$$

where we used that  $X_n^1$  and  $X_n^2$  are supermartingales in the middle step, and just rewrote the indices of the indicator functions. By then re-using that  $Y_n \leq Z_n$ , we get that both  $Y_n, Z_n$  are supermartingales when all the equations are put together.

5. We can obviously write:

$$\begin{aligned} Z_n^{2^{j-1}} &= 1_{N_1 > n} + (X_n/a) 1_{N_1 \leq n < N_2} \\ &\quad + (b/a) 1_{N_2 \leq n < N_3} + (b/a)(X_n/a) 1_{N_3 \leq n < N_4} \\ &\quad \vdots \\ &\quad + (b/a)^{j-1} 1_{N_{2^{j-2}} \leq n < N_{2^{j-1}}} + (b/a)^{j-1} (X_n/a) 1_{N_{2^{j-1}} \leq n} \end{aligned}$$

and

$$\begin{aligned} Z_n^{2^j} &= 1_{N_1 > n} + (X_n/a) 1_{N_1 \leq n < N_2} \\ &\quad + (b/a) 1_{N_2 \leq n < N_3} + (b/a)(X_n/a) 1_{N_3 \leq n < N_4} \\ &\quad \vdots \\ &\quad + (b/a)^{j-1} 1_{N_{2^{j-2}} \leq n < N_{2^{j-1}}} + (b/a)^{j-1} (X_n/a) 1_{N_{2^{j-1}} \leq n < N_{2^j}} \\ &\quad + (b/a)^j 1_{N_{2^j} \leq n} \end{aligned}$$

We prove that  $Z_n^j$  is a supermartingale by induction. First,  $Z_n^1 = 1_{N_1 > n} + (X_n/a) 1_{N_1 \leq n}$ . Let  $X_n^1$  be the supermartingale (constant) 1, and  $X_n^2 = X_n/a$ . Then the stopping time  $N$  that is  $1 \geq X_n/a$  is just the first time  $X_n \leq a$ , which is precisely  $N_1$ . Thus by switching principle,  $Z_n^1$  is a supermartingale. Then clearly,  $Z_n^{2^j} = 1_{N_{2^j} > n} Z_n^{2^{j-1}} + (b/a)^j 1_{N_{2^j} \leq n}$ .

We now investigate the first time when  $Z_n^{2^{j-1}} \geq (b/a)^j$ . If  $n < N_{2^{j-1}}$ , then precisely one term of the form  $(b/a)^\ell$  or  $(b/a)^{\ell-1} X/a$  is non-zero for  $\ell \leq j-1$ . Then the condition

would become either  $(b/a)^\ell \geq (b/a)^j$ , obviously impossible, or  $(b/a)^{\ell-1} X_n/a \geq (b/a)^j$ , which is also impossible since in this case  $N_{2\ell-1} \leq n < N_{2\ell}$ , and the condition would become  $X_n \geq a(b/a)^j(a/b)^{\ell-1} \geq b$ , but this would mean that we would already be over  $b$ , which contradicts the definition of  $N_{2\ell}$  since we should be in between an upcrossing. Thus the only way this can happen is if  $N_{2j-1} \leq n$ . The term would become  $(b/a)^{j-1} X_n/a \geq (b/a)^j$ , which is equivalent to  $X_n \geq b$ , which is precisely how  $N_{2j}$  is defined. Thus  $N_{2j}$  is the stopping time in the switching theorem with  $X_n^1 = Z_n^{2j-1}$  and  $X_n^2 = (b/a)^j$ , so  $Z_n^{2j}$  is a supermartingale as well.

Similarly,  $Z_n^{2j+1} = 1_{N_{2j+1} > n} Z_n^{2j} + (b/a)^j (X/a) 1_{N_{2j+1} \leq n}$ , and the same logic applies and the only way  $Z_n^{2j} \geq (b/a)^j (X/a)$  is when  $N_{2j} \leq n$  and in this case the only term that shows up is  $(b/a)^j$ , so the condition becomes equivalent to  $(b/a)^j \geq (b/a)^j (X/a)$ , which is thus the first time after  $N_{2j}$  that  $X_n \leq a$ , which is how  $N_{2j+1}$  is defined. Thus  $Z_n^{2j+1}$  is a supermartingale as well.

From martingale convergence theorem, we know that:

$$\mathbb{E}(Y_{n \wedge N_{2k}}) \leq \mathbb{E}(Y_0)$$

We claim that  $Y_0 = X_0/a \wedge 1$ . This is because if  $X_0 \leq a$  then  $N_1$  is 0, and  $Y_0$  will be  $X_0/a$ , otherwise  $X_0 > a$  and  $X_0 = 1$ . In conclusion,

$$\left(\frac{b}{a}\right)^{2k} \mathbb{P}(N_{2k} \leq n) = \mathbb{E}(Y_{n \wedge N_{2k}} 1_{N_{2k} \leq n}) \leq \mathbb{E}(Y_{n \wedge N_{2k}}) \leq \mathbb{E}(X_0/a \wedge 1)$$

Sending  $n \rightarrow \infty$  shows that:

$$\mathbb{P}(N_{2k} < \infty) \leq \left(\frac{a}{b}\right)^{2k} \mathbb{E}(1 \wedge X_0/a)$$

Since  $N_{2k} < \infty$  iff the number of upcrossings is at least  $k$ , we conclude Dubins that:

$$\mathbb{P}(U \geq k) \leq \left(\frac{a}{b}\right)^{2k} \mathbb{E}(1 \wedge X_0/a)$$

which completes the proof. ■