

IE 7374 ST: Machine Learning in Engineering  
Spring, 2021, Mid-Term  
Solution Sheet

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1. (1) **True.** For any continuous random variable  $x$  with a probability distribution function  $p(x)$  the probability is bounded as  $0 \leq p(x) \leq 1$  as  $\int_{-\infty}^{\infty} p(x)dx = 1$
  - (2) **False.** If  $X$  and  $Y$  are independent, then  $\mathbb{E}[XY^2] = \mathbb{E}[X]\mathbb{E}[Y^2]$
  - (3) **True.** True error is a hypothetical error value on the entire population. Hence, as  $n \rightarrow \infty$  the training error will converge to the true error.
  - (4) **False.** The linear regression is likely to have high bias and low variance whereas the polynomial regression with degree 3 is likely to have a low bias and high variance.
  - (5) **False.** The empirical risk of  $f_1$  will be greater than  $f_2$ .
  - (6) **True.** Overfitting can occur due to a high approximation error as a result of overly simplistic hypothesis space or small quantity of data.
  - (7) **False.** The approximation error will always be deterministic.
  - (8) **True.** Linear regression has the least variance.
  - (9) **False.** The log likelihood of logistic regression is a convex function and will not have multiple local optima.
  - (10) **False.** The correspondence implies that the posterior Gaussian Naive Bayes can be written in the same form as Logistic Regression.
2. (a)

$$p(y|x_1 = 0) = \frac{p(x_1 = 0|y)p(y)}{p(x_1 = 0)}$$

We know that,

$$\begin{aligned}
p(x_1 = 0|y = 1) &= 1 - \theta_1 = 0.5 \\
p(x_1 = 0|y = 2) &= 1 - \theta_2 = 0.5 \\
p(x_1 = 0|y = 3) &= 1 - \theta_3 = 0.5 \\
p(y = 1) &= 0.5, p(y = 2) = p(y = 3) = 0.25 \\
p(x_1 = 0) &= \sum_{m=1}^3 p(x_1 = 0)p(y = m) = 0.25 + 0.125 + 0.125 = 0.5 \\
p(y = 1|x_1 = 0) &= \frac{0.5 \cdot 0.5}{0.5} = 0.5
\end{aligned}$$

Similarly,

$$\begin{aligned}
p(y = 2|x_1 = 0) &= \frac{0.5 \cdot 0.25}{0.25} = 0.25 \\
p(y = 3|x_1 = 0) &= \frac{0.5 \cdot 0.25}{0.25} = 0.25
\end{aligned}$$

(b)

$$\begin{aligned}
p(y|x_2 = 0) &= \frac{p(x_2 = 0|y)p(y)}{p(x_2 = 0)} \\
p(x_2 = 0|y) &= \frac{1}{\sqrt{2\pi\sigma^2}} \exp^{-\frac{1}{2\sigma^2}(x-\mu)^2} = \begin{bmatrix} 0.242 \\ 0.399 \\ 0.242 \end{bmatrix} \\
p(x_2 = 0) &= \sum_{m=1}^3 p(x_2 = 0)p(y = m) = 0.2815
\end{aligned}$$

Therefore,

$$\begin{aligned}
p(y = 1|x_2 = 0) &= \frac{0.242 \cdot 0.5}{0.2815} = 0.43 \\
p(y = 2|x_2 = 0) &= \frac{0.399 \cdot 0.5}{0.2815} = 0.354 \\
p(y = 3|x_2 = 0) &= \frac{0.242 \cdot 0.5}{0.2815} = 0.215
\end{aligned}$$

(c)  $p(y = 2|x_1) = p(y = 3|x_1)$ . Probability of classifying y as 2 or 3, given  $x_1$ , are equal .

3. (a)

$$\begin{aligned}
p(z_{ik} = 1|p(w_i)) &= p(w_i) \\
p(z_{ik} = 0|p(w_i)) &= 1 - p(w_i) \\
p(z_{ik}|w_i) &= p(w_i)^{z_{ik}}(1 - p(w_i))^{1-z_{ik}}
\end{aligned}$$

Using the law of independence, we have,

$$\begin{aligned}
p(z_{ik}|w_i) &= p(w_i)^{z_{ik}}(1 - p(w_i))^{1-z_{ik}} \\
&= \prod_{k=1}^n p(z_{ik}|p(w_i)) \\
&= \prod_{k=1}^n (p(w_i)^{z_{ik}}(1 - p(w_i))^{1-z_{ik}})
\end{aligned}$$

- (b) From the likelihood function derived above we can take log on both sides and rewrite the equation as

$$\begin{aligned}
\ln p(w_i) &= \ln \left( \prod_{k=1}^n (p(w_i)^{z_{ik}}(1 - p(w_i))^{1-z_{ik}}) \right) \\
&= \sum_{k=1}^n z_{ik} \ln(p(w_i)) + (1 - z_{ik}) \ln(1 - p(w_i))
\end{aligned}$$

Taking derivative with respect to  $p(w_i)$  and equating it to zero, we get

$$\begin{aligned}
\frac{\partial \ln p(w_i)}{\partial p(w_i)} &= \frac{1}{p(w_i)} \sum_{i=1}^n z_{ik} - \frac{1}{1 - p(w_i)} \sum_{k=1}^n (1 - z_{ik}) \\
&\quad \frac{1}{p(w_i)} \sum_{i=1}^n z_{ik} - \frac{1}{1 - p(w_i)} \sum_{k=1}^n (1 - z_{ik}) = 0 \\
\therefore (1 - p(w_i)) \sum_{i=1}^n z_{ik} &= p(w_i) \sum_{k=1}^n (1 - z_{ik}) \\
\sum_{i=1}^n z_{ik} &= p(w_i) \sum_{i=1}^n z_{ik} + np(w_i) - p(w_i) \sum_{i=1}^n z_{ik} \\
\therefore p(w_i) &= \frac{1}{n} \sum_{i=1}^n z_{ik}
\end{aligned}$$

4.

$$p(y = 1|x) = \frac{p(x|y = 1)p(y = 1)}{p(y = 1)p(x|y = 1) + p(y = 2)p(x|y = 2) + p(y = 3)p(x|y = 3)}$$

5. Using Bayes' theorem we can write

$$p(\theta|x < 3) = \frac{p(x < 3|\theta)p(\theta)}{p(x < 3)}$$

Prior can be written as the pdf of Beta distribution =  $Beta(\theta, 1, 1) = \frac{1}{Beta(1,1)}\theta(1 - \theta)^{1-1} = Uniform(0, 1)$

Further, the likelihood can be given as

$$\begin{aligned}
 p(x < 3|\theta) &= \sum_{i=1}^3 p(x = i|\theta) \\
 &= \binom{5}{0} \theta^0 (1 - \theta)^5 + \binom{5}{1} \theta^1 (1 - \theta)^4 + \binom{5}{2} \theta^2 (1 - \theta)^3 \\
 &= \sum_{k=0}^2 \text{Binomial}(k|\theta, 5)
 \end{aligned}$$

Therefor, the posterior can then be approximated as

$$\begin{aligned}
 p(\theta|x < 3) &\approx p(x < 3|\theta)p(\theta) \\
 &= \left( \sum_{k=0}^2 \text{Binomial}(k|\theta, 5) \right) \cdot \text{Uniform}(0, 1)
 \end{aligned}$$