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% Chapter 04 — Randomized Algorithms % CS161 Reader (Plotkin F10) %

## 1 Chapter 04 — Randomized Algorithms

Randomization is a *proof technique disguised as an implementation trick*. The key discipline is to state **exactly** what is random, what is adversarial, and what kind of guarantee you are proving.

**Primary patterns:** expectation (indicators + linearity), amplification (repeat/boost), induction (recurrences).

**Executable witnesses:** `cs161lab.algorithms.sorting.rand_quicksort`, `cs161lab.algorithms.sorting.quickselect`, `cs161lab.algorithms.mincut.karger`.

### 1.1 4.1 Model and guarantee types

We analyze algorithms in the randomized RAM model: the algorithm may draw independent random bits at any step.

#### 1.1.1 Las Vegas vs Monte Carlo

- **Las Vegas:** always correct; randomness affects running time (Randomized QuickSort).
- **Monte Carlo:** fixed running time; randomness affects correctness (Karger’s min-cut).

### 1.1.2 Common guarantee traps

Expected-time bounds do **not** automatically imply “fast with high probability.” When you need high probability, use explicit tail bounds or amplification (repetition plus a correctness filter or “best-of” selection).

## 1.2 4.2 Linearity of expectation and indicator variables

### 1.2.1 Theorem 4.1 (Linearity of expectation)

For any random variables  $X_1, \dots, X_m$  (independence not required),

$$E\left[\sum_{i=1}^m X_i\right] = \sum_{i=1}^m E[X_i].$$

**Proof.** Expand  $E[\cdot]$  by definition (sum/integral) and swap summations/integrals.  $\square$

### 1.2.2 Lemma 4.2 (Indicators)

If  $I$  is an indicator (0/1) for an event  $E$ , then  $E[I] = P(E)$ .

**Proof.**  $E[I] = 0 \cdot P(I = 0) + 1 \cdot P(I = 1) = P(I = 1) = P(E)$ .  $\square$

This is the standard CS161 move: reduce expected counts to sums of event probabilities.

## 1.3 4.3 Randomized QuickSort

### 1.3.1 Algorithm

Given  $n$  distinct keys: 1. Pick a pivot uniformly at random from the subarray. 2. Partition around the pivot. 3. Recurse on both sides.

**Executable witness:** `cs161lab.algorithms.sorting.rand_quicksort`

**Trace events:** `pivot`, `partition`.

### 1.3.2 Theorem 4.3 (Expected comparisons)

Let  $C_n$  be the number of comparisons. Then

$$E[C_n] = 2(n+1)H_n - 4n = O(n \log n),$$

where  $H_n = \sum_{k=1}^n \frac{1}{k}$  is the  $n$ -th harmonic number.

**1.3.2.1 Proof (indicator-variable proof)** Label keys in sorted order  $z_1 < z_2 < \dots < z_n$ . For each  $i < j$ , define  $I_{ij} = 1$  iff QuickSort compares  $z_i$  and  $z_j$  at some point. Then

$$C_n = \sum_{i < j} I_{ij} \quad \Rightarrow \quad E[C_n] = \sum_{i < j} E[I_{ij}] = \sum_{i < j} P(I_{ij} = 1).$$

Consider the set  $S_{ij} = \{z_i, z_{i+1}, \dots, z_j\}$ . Keys  $z_i$  and  $z_j$  are compared **iff** the first pivot chosen from  $S_{ij}$  is one of the endpoints ( $z_i$  or  $z_j$ ). If the first pivot in  $S_{ij}$  is some  $z_k$  with  $i < k < j$ , the recursion separates  $z_i$  and  $z_j$  forever.

The first pivot drawn from  $S_{ij}$  is uniformly distributed among its  $|S_{ij}| = j - i + 1$  elements, so

$$P(I_{ij} = 1) = \frac{2}{j - i + 1}.$$

Therefore,

$$E[C_n] = \sum_{i < j} \frac{2}{j - i + 1} = \sum_{\ell=2}^n \sum_{i=1}^{n-\ell+1} \frac{2}{\ell} = \sum_{\ell=2}^n \frac{2(n-\ell+1)}{\ell} = 2(n+1)(H_n - 1) - 2(n-1) = 2(n+1)H_n - 4n.$$

□

## 1.4 4.4 Randomized Selection (QuickSelect)

QuickSelect finds the  $k$ -th smallest element by partitioning around a random pivot and recursing on one side.

**Executable witness:** `cs161lab.algorithms.sorting.quickselect`

**Trace events:** `pivot`, `partition`, `shrink`, `found`.

### 1.4.1 Theorem 4.4 (Expected linear time)

The expected running time of QuickSelect is  $O(n)$ .

**1.4.1.1 Proof (good pivot with constant probability)** Partitioning costs  $cn$ . A pivot is **good** if its rank lies in the middle half (between  $n/4$  and  $3n/4$ ), which happens with probability at least  $1/2$ . With a good pivot, the recursion continues on size at most  $3n/4$ . With a bad pivot, size is at most  $n - 1$ . Hence,

$$E[T(n)] \leq cn + \frac{1}{2}E[T(3n/4)] + \frac{1}{2}E[T(n-1)].$$

A standard induction shows this solves to  $E[T(n)] \leq an$  for a sufficiently large constant  $a$ . □

## 1.5 4.5 Amplification and the union bound

### 1.5.1 Theorem 4.5 (Union bound)

For events  $E_1, \dots, E_m$ ,

$$P\left(\bigcup_{i=1}^m E_i\right) \leq \sum_{i=1}^m P(E_i).$$

### 1.5.2 Amplification lemma

If an independent trial succeeds with probability at least  $p > 0$ , repeating  $t$  times yields failure probability at most  $(1-p)^t \leq e^{-pt}$ . To get failure at most  $\delta$ , choose  $t \geq \frac{1}{p} \ln(1/\delta)$ .

## 1.6 4.6 Karger's randomized min-cut

Karger's algorithm is Monte Carlo: one run can fail, but it is fast, and repetition makes failure negligible.

**Executable witness:** `cs161lab.algorithms.mincut.karger`

**Trace events:** `contract`, `self_loop`, `trial`, `best`, `result`.

### 1.6.1 Algorithm (Random Contraction)

While more than 2 supernodes remain: 1. pick a uniformly random edge, 2. contract its endpoints, 3. delete self-loops. Return the cut induced by the final 2 supernodes.

### 1.6.2 Theorem 4.6 (One-run success probability)

For an  $n$ -vertex multigraph with min-cut value  $\lambda$ , one run returns a min-cut with probability at least  $\frac{2}{n(n-1)}$ .

**1.6.2.1 Proof** Fix a particular minimum cut  $C$  of size  $\lambda$ . The run succeeds if no contraction ever contracts an edge crossing  $C$ . When there are  $k$  supernodes left and we have not crossed  $C$ , the cut still has exactly  $\lambda$  crossing edges. Also, the number of edges  $m_k$  satisfies  $m_k \geq k\lambda/2$  because the minimum degree is at least  $\lambda$  (otherwise we'd have a smaller cut). Thus,

$$P(\text{avoid } C \text{ at stage } k) = 1 - \frac{\lambda}{m_k} \geq 1 - \frac{\lambda}{k\lambda/2} = 1 - \frac{2}{k}.$$

Multiply from  $k = n$  down to 3:

$$P(\text{avoid } C \text{ for all contractions}) \geq \prod_{k=3}^n \left(1 - \frac{2}{k}\right) = \prod_{k=3}^n \frac{k-2}{k} = \frac{2}{n(n-1)}.$$

□

### 1.6.3 Corollary 4.7 (Repetition schedule)

After  $t$  independent runs and taking the best cut,

$$P(\text{all fail}) \leq \left(1 - \frac{2}{n(n-1)}\right)^t \leq \exp\left(-\frac{2t}{n(n-1)}\right).$$

To make failure  $\leq \delta$ , it suffices to take  $t = \Theta(n^2 \log(1/\delta))$ .

## 1.7 4.7 What can go wrong

- Running a Monte Carlo algorithm without amplification (and then believing it).
- Assuming tail bounds from expectations without proof.
- Non-uniform pivot selection (implementation bias).
- PRNG seeding mistakes that destroy independence assumptions.

## 1.8 Appendix — Trace events as proof witnesses

- QuickSort: `pivot`, `partition` support the indicator proof narrative (who gets compared).
- QuickSelect: `shrink` is a termination witness; the good-pivot argument predicts frequent shrinkage.
- Karger: `contract` and `trial` let you empirically validate the repetition schedule.