

Predicting Sepsis in the ICU

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May 11, 2015

Research Question

This study will focus on developing a prediction model for the onset of sepsis in the ICU using clinical history, demographic information, and non-invasive physiological data obtained within the first hours following ICU admission. Using a large dataset of patients, routine clinical measurements obtained during initial stages of care, imputation techniques, and data mining methodologies, the goal will not only be to build a classifier that surpasses the performance of current sepsis models, but also to provide a potential early warning system for sepsis in the general clinical care setting.

Introduction

Sepsis is systemic inflammatory response syndrome (SIRS), secondary to a documented infection. Sepsis can present itself on a continuum that ranges from sepsis, severe sepsis, and septic shock, resulting in multiple organ dysfunction. The symptoms of sepsis are often non-specific and involve difficulty breathing, hypoxemia, hypoperfusion, and hypotension [1].

Although sepsis is a common condition worldwide, the current understanding of the pathophysiology of sepsis has increased substantially, and sepsis mortality has declined in the last two decades [2]. The reason for the decline may be attributed to improved supportive care and the inherent symptomatology of patients who fall prey to sepsis. On the contrary, epidemiologic data suggests that sepsis incidence is increasing [2]. New treatments and therapies have failed to demonstrate efficacy. Sepsis affects approximately 700,000 people per year, and accounts for approximately 200,000 deaths per year in the United States [3], amassing an annual cost of 16.7 billion dollars [4].

The best form of treatment is preventive treatment. Early diagnosis and appropriate therapy must be typically delivered before laboratory test results are known, which bases the diagnosis on the co-presence of routine clinical measures. The SIRS criteria was developed in the 1991 International Sepsis Definition Conference to address these concerns and is still commonly used in the clinical care setting to flag patients for risk of sepsis [1]. Patients who meet the SIRS criteria exhibit two or more of the following symptoms:

- Temperature > 38 degrees Celsius or < 36 degrees Celsius
- Heart Rate > 90 bpm
- Respiratory Rate > 20 or $\text{PaCO}_2 < 32$ mm Hg
- White Blood Cell Count $> 12,000/\text{mm}^3$, $< 4,000/\text{mm}^3$, or $> 10\%$ bands

Unfortunately, the SIRS criteria has low discriminatory power in the intensive care unit as many critically ill patients who are not at risk for sepsis may also exhibit similar symptoms [5]. Previous studies demonstrated the poor utility of the SIRS criteria in identifying septic patients within a clinical care setting, in which SIRS exhibited both low sensitivity and specificity [6]. In the case of identifying patients at risk for sepsis, a test with poor sensitivity can be particularly harmful as false negatives may not receive the proper prophylactic care needed to prevent sepsis-related complications. As such, a high-recall prediction model (low false negatives) to identify patients with sepsis may provide benefits to caregivers in the form of an early warning system.

While previous studies have largely focused on predicting septic shock [7], few studies have focused on predicting earlier stages of the sepsis continuum. Multivariate logistic regression (Shavdia, 2007), decision

trees [8], and Dynamic Bayesian Networks [9] approaches have been used to predict sepsis in the intensive care unit. However, these studies tended to use a large number of invasive measurements - such as arterial blood pressure - in their feature set, reducing generalizability. Moreover, while other studies looked at the last measurements taken before the onset of sepsis [10], few models incorporated summary statistics (mean/sd or other pairs) of clinical features in the feature set to capture the centrality and dispersion of these measurements over time. Our study attempts to synthesize and add to previous approaches by applying: a) “modern” classification methods (naive bayes, regularized logistic regression, and random forest) to potentially improve model performance, b) summary statistics to routine, non-invasive clinical features to capture information from time-series data, and c) imputation methods to avoid pitfalls due to missing data.

Materials and Methods

Dataset

The data was obtained from the Multiparameter Intelligent Monitoring in Intensive Care Database (MIMIC II), a semi-public database which presents ICU patient records for approximately 25,000 adults at Boston’s Beth Israel Deaconess Medical Center. As a large, diverse dataset of ICU patients, MIMIC II is appropriate for building prediction models for critically ill patient populations.

Data for the analysis was sourced from a MIMIC II database instance running the PostgreSQL (version 9.2.10) engine. The package RPostgreSQL provided a Database Interface (DBI) compliant driver for R to access the PostgreSQL database system.

```
# devtools::add_path('~/.usr/local/bin/psql')
# devtools::install("~/Downloads/rpostgresql-read-only/RPostgreSQL/")
db <- src_postgres(dbname = "MIMIC2",
  host = "ec2-54-163-173-71.compute-1.amazonaws.com",
  port = 5432,
  user = "ec2-user",
  password = "thisisalongpassword1234")
```

Patient Selection

This study examined adults (≥ 16 years of age). Since our objective is to train a prediction model that will detect the onset of sepsis within the first few hours of admittance to the ICU, we included only patients who were admitted to the ICU for the first sequence of their hospital visit. To avoid bias introduced by censorship, we excluded samples who have not been in the ICU for longer than 24 hours, as patients will not have accrued enough data to make a risk assessment.

The outcome measure for this study was any instance of sepsis on the “sepsis continuum” - sepsis (995.91), severe sepsis (995.92), or septic shock (785.52) - as defined by International Classification of Diseases, 9th revision (ICD9). Severity of sepsis was not graded in this study, and thus any patient with sepsis-related ICD-9 codes was determined to have the same level of risk.

For our negative controls, we randomly sampled 5000 patients from the MIMIC II database who met the same exclusionary criteria.

Confounding Medical Interventions

The “ground truth” outcome for a patient is obscured by confounding medical interventions [11]. For example, we cannot determine the true sepsis risk for patients who receive antibiotics upon admission into the ICU because the treatment masks the “ground truth” - we don’t know whether the physician was right or wrong his

risk assessment of the patient. To avoid bias introduced by these so-called confounding medical interventions, patients who have undergone treatment with antibiotics within the first 24 hours will also be excluded.

Feature Selection

Since we are interested in creating a classifier that can predict the onset of sepsis using accessible clinical data available within the first few hours since admission, we restrict our feature set to the `demographic`, `icustay_details`, and `chartevents` tables in the MIMIC II database, which includes:

- demographic data (gender, age, etc.)
- chart events (SOFA score, SAPS-I score)
- basic health data (height, weight, etc.)

Demographic

We extract all variables from the `demographic` table, which includes items like religious affiliation, insurance information, and marriage status. From the `icustay_detail` table, we extract patient information that was recorded upon admission, such as age, admission time, weight, height, etc.

To avoid the pitfalls of categorical features with zero variance, we collapsed classes into a smaller number of sensible categories. To deal timestamps, we factorized time into “morning”, “afternoon”, “evening”, and “night” levels.

Chart Events

We determined “routine clinical measurements” to be variables in which over 80% of our population had at least 1 clinical measurement recorded on a per hourly basis during the first 24 hours. Not surprisingly, these variables were `respiratory rate`, `pulse oximetry`, `heart rate`, `non-invasive blood pressure`, `white blood cell count`, `sofa score`, `sapsi score`, and `temperature`.

Subsequent tables from the MIMIC II database were linked through a combination of the ICU stay IDs, subject IDs, and chart item IDs (for features). Routine clinical features were then discretized by time cutoffs, melted, and casted in order to create each hour time point as a feature in itself. To evaluate the performance of our prediction models with respect to time, we then created summary statistics (mean, min, max, sd) for each clinical feature using the discretized values.

Statistical Analysis

Following data extraction, 38 features were available for predictive analysis.

Missing values were subsequently identified and imputed using kNN imputation with $k = 10$. Standardization of continuous clinical measures was performed on numeric features in order to establish comparability (mean of 0 and a standard deviation of 1). Values falling outside the range will allow us to determine outliers and impossible values.

Six models were then selected for prediction of sepsis: logistic regression, regularized logistic regression, naïve Bayes, and C4.5-like decision trees (information gain), recursive partitioning trees (gini impurity), and random forest. Examination for collinearity was performed using linear correlations, resulting in 27 features. The remaining models were conducted on the full set of 39 features. All methodologies were evaluated using repeated 10-fold cross validation to obtain 30 resamples with results represented using area under the receiver operating characteristic curves (AUROC). AUROC was chosen as a performance statistic because it is insensitive to class balance issues.

Since our goal is to detect early onset of sepsis, we ideally want our predictive model to have high AUROC when predicting on data available within the first hour. We will also train our models on data collected at the following time intervals, e.g. 3 hours, 6 hours, 12 hours, and 24 hours after ICU admission. By doing this, we will be able to assess the stability of our performance statistics with respect to elapsed time in the ICU and determine whether the first hour of data is sufficient in a predictive model for sepsis.

Finally, to determine whether or not our prediction models perform better than the SIRS criteria mentioned above, we will create an implementation of the SIRS criteria and compare models using the “balanced accuracy” measure, which avoids inflated performance estimates on imbalanced datasets. It is defined as the arithmetic mean of sensitivity and specificity:

$$\text{balanced accuracy} = \frac{\text{sensitivity} + \text{specificity}}{2}$$

Results

A total of 2,783 patients were used for analysis, representing 2,783 unique ICU admissions. Among the cohort, 17.8% developed some form of sepsis during their ICU stay.

Table 1: Descriptive Statistics for Selected Features

```
## .
##
## 17 Variables      2783 Observations
## -----
## label
##      n missing  unique
##    2783         0      2
##
## case (496, 18%), control (2287, 82%)
## -----
## gender
##      n missing  unique
##    2783         0      2
##
## F (1106, 40%), M (1677, 60%)
## -----
## marital_status_descr
##      n missing  unique
##    2783         0      4
##
## divorced (187, 7%), married (1596, 57%)
## single (588, 21%), widowed (412, 15%)
## -----
## ethnicity_descr
##      n missing  unique
##    2783         0      6
##
##      asian black hispanic other unknown white
## Frequency   53   165      72   110     402   1981
## %           2     6        3     4      14     71
## -----
## overall_payor_group_descr
##      n missing  unique
```

```

##      2783      0      8
##
##          AUTO LIABILITY FREE CARE MEDICAID MEDICARE MEDICARE-PRIVATE
## Frequency          55      48      205      1227      235
## %              2      2      7      44      8
##          OTHER PRIVATE SELF-PAY
## Frequency      51      936      26
## %              2      34      1
## -----
## religion_descr
##      n missing  unique
##      2783      0      8
##
##          buddhist catholic jewish not specified orthodox other protestant
## Frequency      16      1126      236      451      21      182      421
## %              1      40      8      16      1      7      15
##          unobtainable
## Frequency      330
## %              12
## -----
## admission_type_descr
##      n missing  unique
##      2783      0      3
##
## ELECTIVE (692, 25%), EMERGENCY (1975, 71%)
## URGENT (116, 4%)
## -----
## admission_source_descr
##      n missing  unique
##      2783      0      6
##
## CLINIC REFERRAL/PREMATURE (194, 7%)
## EMERGENCY ROOM ADMIT (1082, 39%)
## PHYS REFERRAL/NORMAL DELI (813, 29%)
## TRANSFER FROM HOSP/EXTRAM (676, 24%)
## TRANSFER FROM OTHER HEALT (4, 0%)
## TRANSFER FROM SKILLED NUR (14, 1%)
## -----
## heart_rate_mean
##      n missing  unique      Info      Mean      .05      .10      .25      .50
##      2783      0      2694      1      84.7      61.8      66.3      74.7      83.8
##      .75      .90      .95
##      93.5      104.7      110.9
##
## lowest :      32.3      40.5      42.9      43.3      44.9
## highest:      135.2      135.6      137.4      139.7      144.5
## -----
## nbp_mean_mean
##      n missing  unique      Info      Mean      .05      .10      .25      .50
##      2665      118      2449      1      75.9      59.1      62.5      67.3      74.3
##      .75      .90      .95
##      83.3      91.9      98.6
##
## lowest :      31.1      32.0      39.6      40.2      43.0

```

```

## highest: 122.3 123.0 133.9 140.9 142.2
## -----
## nbp_mean_min
##      n missing  unique    Info    Mean    .05    .10    .25    .50
##    2665     118     578      1    64.2    46.7    50.2    56.0    62.7
##      .75     .90     .95
##    71.5     81.0     87.3
##
## lowest :   0.0  21.0  22.9  25.0  29.3
## highest: 115.0 117.7 118.0 122.3 142.2
## -----
## nbp_mean_max
##      n missing  unique    Info    Mean    .05    .10    .25    .50
##    2665     118     673      1    89.3    64.5    69.3    77.3    88.0
##      .75     .90     .95
##   100.0    110.4    117.3
##
## lowest :  43.0  43.3  43.5  46.7  46.7
## highest: 147.7 158.7 165.0 168.0 179.7
## -----
## nbp_mean_std
##      n missing  unique    Info    Mean    .05    .10    .25    .50
##    2681     102    2386      1    7.73    0.00    1.18    4.98    7.36
##      .75     .90     .95
##   10.22    13.36    15.94
##
## lowest :  0.000  0.118  0.157  0.236  0.236
## highest: 33.403 33.941 34.591 45.255 58.274
## -----
## respiratory_rate_mean
##      n missing  unique    Info    Mean    .05    .10    .25    .50
##    2781         2    2446      1    18.4    13.3    14.2    15.8    17.9
##      .75     .90     .95
##    20.4     23.7     25.5
##
## lowest :  0.60  8.41  8.58  9.20  9.31
## highest: 36.00 36.64 36.74 42.48 43.15
## -----
## spo2_mean
##      n missing  unique    Info    Mean    .05    .10    .25    .50
##    2782         1    2144      1    97.3    94.4    95.3    96.4    97.7
##      .75     .90     .95
##    98.7     99.4     99.7
##
## lowest :  55.6  59.1  59.2  65.5  65.8
## highest: 100.0 100.0 100.0 100.0 100.0
## -----
## temperature_mean
##      n missing  unique    Info    Mean    .05    .10    .25    .50
##    2656     127    2400      1    36.7    35.9    36.1    36.5    36.8
##      .75     .90     .95
##    37.2     37.6     37.8
##
## lowest :  2.02  2.22  2.28  2.56  2.61

```

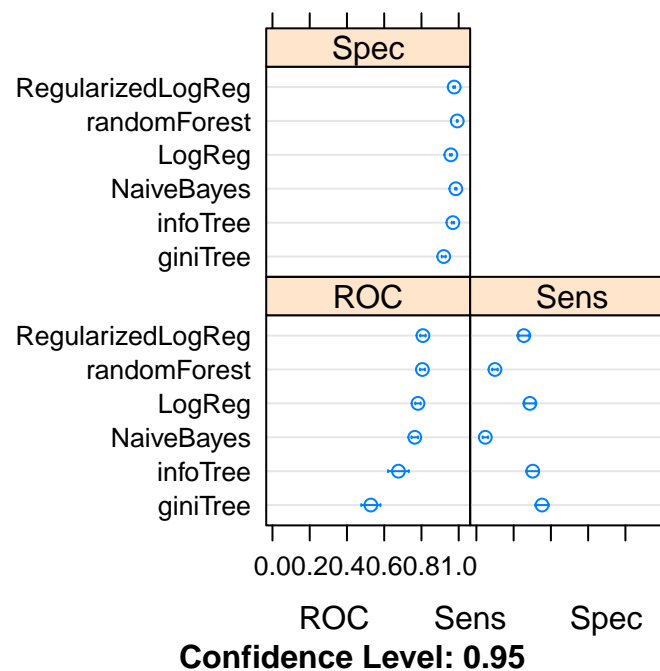
```
## highest: 38.86 38.88 38.89 39.08 39.76
## -----
## wbc_mean
##      n missing  unique    Info    Mean    .05    .10    .25    .50
##    2116     667     886      1   13.1   5.04   6.30   8.44  11.15
##      .75     .90     .95
##    14.75   19.30   22.84
##
## lowest :    0.10    0.20    0.30    0.40    0.45
## highest:   53.80   57.05   61.70  138.68 1567.88
## -----
```

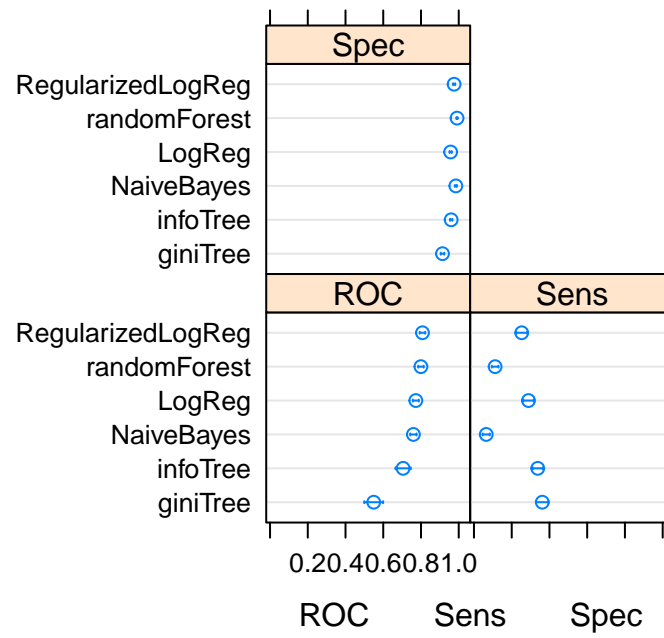
Model Selection

```
## note: only 8 possible values of the max tree depth from the initial fit.
## Truncating the grid to 8 .
##
## note: only 8 possible values of the max tree depth from the initial fit.
## Truncating the grid to 8 .
```

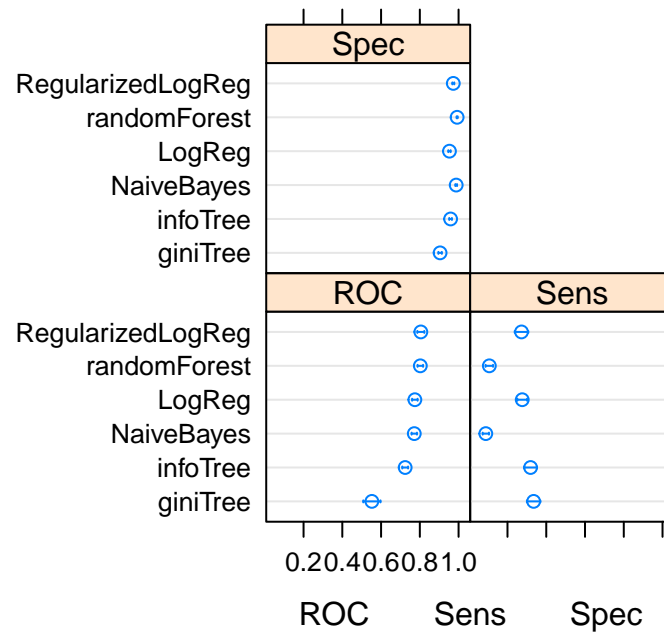
Cross-Validation

Performance statistics for repeated 10-fold cross validation did not vary significantly when trained on data available at 1 hour, 3 hours, 6 hours, 12 hours, and 24 hours (shown in sequential order).

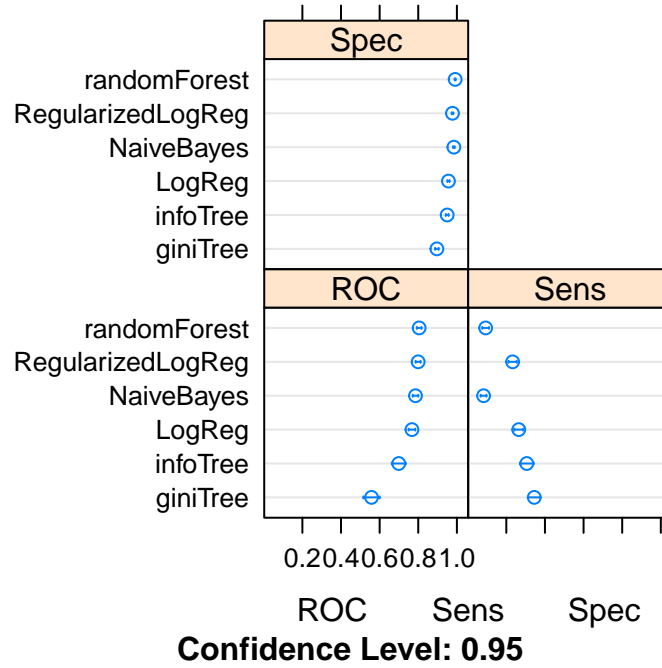
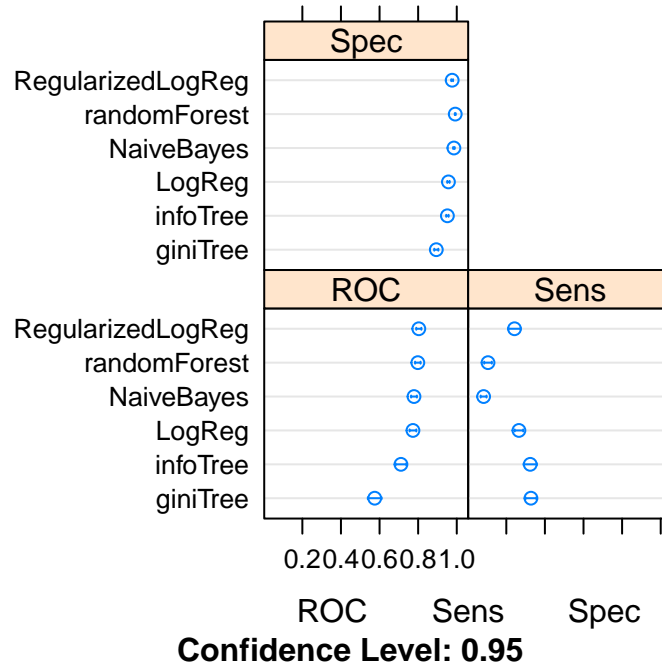




Confidence Level: 0.95



Confidence Level: 0.95



Given the results of our sensitivity analysis, we opted for the models fitted on the first hour of data made available in the ICU. Resampled ROC statistics for each of these models are shown below.

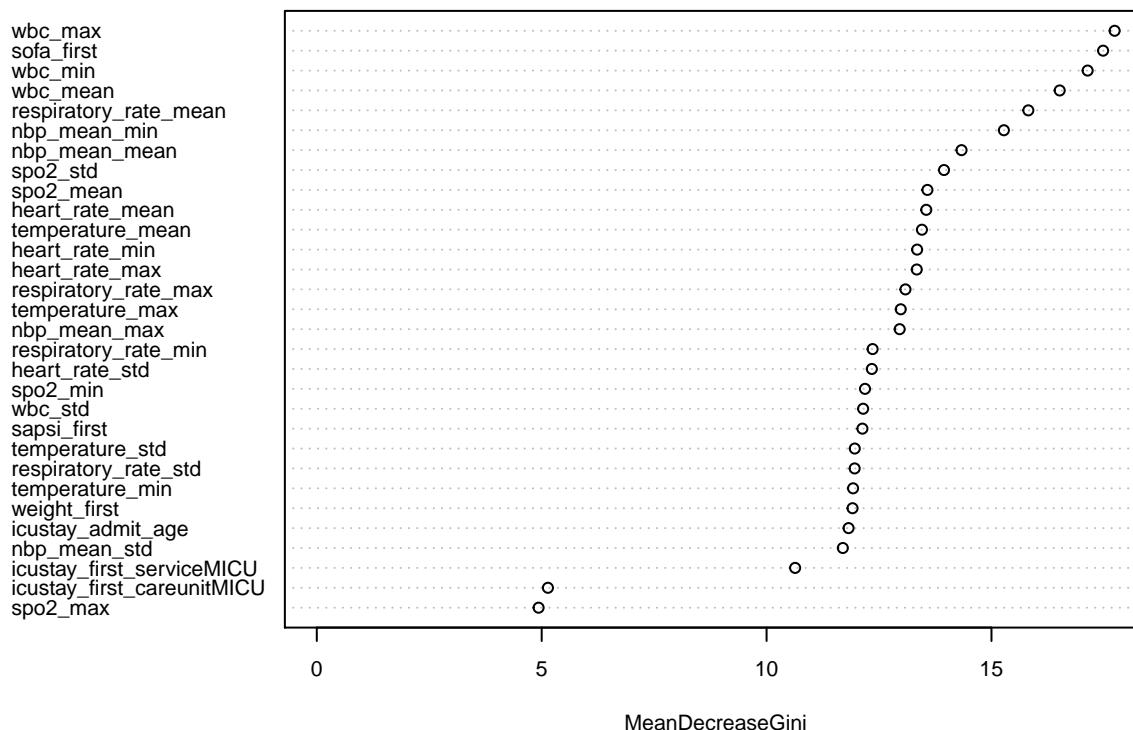
	Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
LogReg	0.7119	0.7553	0.7769	0.7813	0.8093	0.8562
RegularizedLogReg	0.7152	0.7810	0.8139	0.8079	0.8316	0.8912
NaiveBayes	0.6545	0.7389	0.7719	0.7641	0.7845	0.8592
infoTree	0.3337	0.7015	0.7359	0.6764	0.7544	0.8113
giniTree	0.3386	0.3914	0.4858	0.5283	0.6646	0.7255
randomForest	0.7121	0.7843	0.8096	0.8052	0.8208	0.8699

Under 10-fold cross validation, regularized logistic regression had an ROC of 0.808, random forest had an ROC of 0.804, logistic regression had an ROC of 0.781, Naive Bayes had an ROC of 0.764, C4.5-like decision trees had an ROC of 0.67, and CART decision trees had an ROC of 0.528. With respect to the tradeoff between performance and interpretability, logistic regression was not too shabby.

Variable Importance

For logistic regression, top features selected for classification included SAPS-I score, MICU as First Service, Minimum Non-invasive Blood Pressure, and Maximum Pulse Oximetry. For regularized logistic regression, 10 features remained after coefficients were penalized, which included CSRU as First Service, Private Medical Insurance, and Morning ICU Admission. As determined by the decrease in the mean Gini in our CART model, First SOFA score, Minimum White Blood Cell Count, and Maximum Respiratory Rate were the most important features. As determined by the decrease in the mean Gini in our random forest model, Mean White Blood Cell Count, First SOFA Score, and Mean Heart Rate were determined to be the most important features. SOFA, SAPS-I, heart rate, white blood cell count, pulse oximetry, and respiratory rate have been previously described as risk factors for sepsis (see SIRS criteria above).

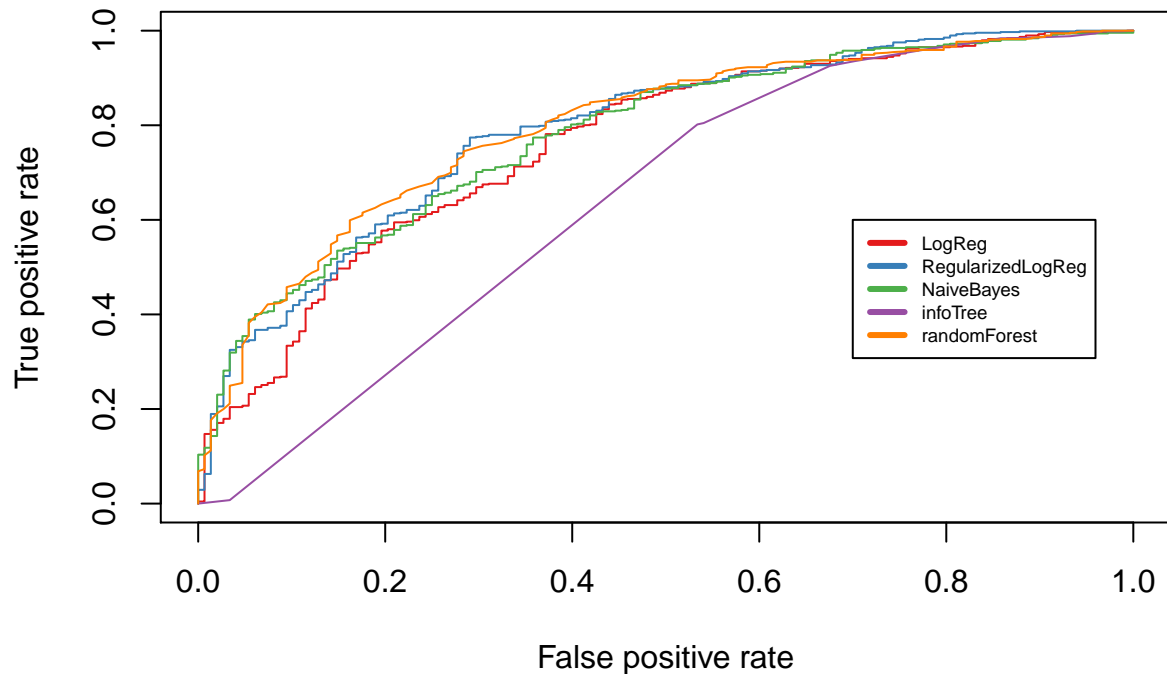
RF Variable Importance



Hold-Out Performance

Between-Model Comparison

Most of the classifiers showed strong predictive power in predicting sepsis in the ICU. Because our positive class is in the minority (cases ~ 18%), accuracy statistics were not suitable for model assessment. Instead we used ROC statistics, which are agnostic to class imbalance issues. C4.5-like decision trees were not included in the figure below because we were unable to debug the prediction task.



Baseline Comparison

So did our models perform better than the current SIRS criteria in predicting sepsis in the ICU?

	Sensitivity	Specificity	Pos.Pred.Value	Neg.Pred.Value	Balanced.Accuracy
SIRS	0.5878	0.4942	0.2005	0.8475	0.5410
LogReg	0.2297	0.9650	0.5862	0.8531	0.5974
RegLogReg	0.2162	0.9825	0.7273	0.8532	0.5994
NaiveBayes	0.0608	0.9942	0.6923	0.8307	0.5275
GiniTree	0.2027	0.9665	0.5660	0.8489	0.5846
RandomForest	0.0676	0.9956	0.7692	0.8319	0.5316

Yes, but not by much. Moreover, training on data available at the 3 hour, 6 hour, 12 hour, and 24 hour time slices only marginally improved the performance of our prediction models with regularized logistic regression leading the pack. At the end of the day, however, logistic regression had the best balanced accuracy using one hour's worth of data on our hold-out set at 59.5%.

Discussion

When we began this project, we sought to answer the following question: could we develop a prediction model for the onset of sepsis in the ICU using clinical history and non-invasive physiological data obtained in the first hour of admission? The short answer: yes, but predictive performance is only marginally better than current (and simpler) methodologies.

Not suprisingly, fitting more complex models does not necessarily lead to better performance. While we used repeated 10-fold cross validation to fine-tune complexity parameters, we very likely overfit our training data using random forest - which performed worse than the SIRS criteria on our hold-out set. In some instances, regularized logistic regression performed marginally better than other models because 10-fold cross validation tuned the alpha parameter to 1 (LASSO), which allowed the model to undergo automated feature selection

by shrinking some coefficients zero. However, the tradeoff in the interpretability of the model in its native form may not be worth the marginal improvements in predictive power. After all, clinical decision support systems should have some descriptive transparency. Penalized regression models and random forest are not likely to provide that. “Vanilla” logistic regression - with well-understood parameters, wide usage, and one of the top balanced accuracy measures in our study - is.

Moreover, it is important to note that performance statistics provided in this study are based on a 50% probability threshold for classification. In other words, samples are classified as “sepsis” only if the probability of class is greater than 50%. Since false positives (patients classified as sepsis when they are not at risk) and false negatives (patients classified as not sepsis when they are at risk) are not equal in this scenario, we may want to adjust the threshold downwards and attempt to improve the recall of our prediction models.

Limitations and Future Work

Better features would have likely improved the performance of our prediction models. For one, our decision to restrict features to routine clinical variables was based heavily on domain knowledge rather than a purely data-driven methodology. This meant that our models suffered from substantial information loss. Determining a better definition for “routine” and translating that definition into an automated feature selection process could have given us more (and better) features to work with from the start. Second, valuable information from temporal features was not fully extracted using summary statistics (mean, min, max, sd). While these simple statistical features are a good starting point, future prediction models could incorporate parameters of ARIMA models or frequencies of the k peaks in amplitude in the Discrete Fourier Transforms for the detrended d dimensions.

Moreover, our decision to define positive cases of sepsis from ICD-9 codes presented a number of problems. For one, our so-called “ground truth” labels may, in fact, be mislabeled due to administrative errors. This would mean that we were training our classifiers on incorrectly labeled samples, essentially making this work non-generalizable and effectively useless. Moreover, we did not consider where the rank of the sepsis code within the ICD-9 sequence for each individual patient. For example, we considered a patient with a sepsis code in the #1 slot (indicating a primary diagnosis) and the #10 slot to be the same. Future studies could use prior weights based on this ICD-9 code sequence or avoid ICD-9 codes altogether and determine a physiological cutoff for the onset of sepsis.

Conclusions

In this project, we have shown that a) supervised classification techniques b) summary statistics for routine, non-invasive clinical features and c) imputation methods can be implemented in order to predict sepsis in the ICU, albeit only marginally better than current systems. Logistic regression - after feature elimination based on collinearity - provided the best performance while maintaining interpretability. This work presented in this report may serve to guide the development of an early warning system for sepsis in the intensive care setting.

Acknowledgements

Thanks Noemie, Andrew, and William for making our last semester at CUMC a blast! And for making access to the MIMIC II database a lot more tolerable than it otherwise would have been.

Session Information

```
## R version 3.2.0 (2015-04-16)
## Platform: x86_64-apple-darwin13.4.0 (64-bit)
```

```
## Running under: OS X 10.10.3 (Yosemite)
##
## locale:
## [1] en_US.UTF-8/en_US.UTF-8/en_US.UTF-8/C/en_US.UTF-8/en_US.UTF-8
##
## attached base packages:
## [1] parallel  grid      stats      graphics  grDevices  utils      datasets
## [8] methods   base
##
## other attached packages:
## [1] randomForest_4.6-10 rpart_4.1-9      RWeka_0.4-24
## [4] klaR_0.6-12         MASS_7.3-40      glmnet_2.0-2
## [7] Matrix_1.2-0        doMC_1.3.3       iterators_1.0.7
## [10] foreach_1.4.2       dplyr_0.4.1      plyr_1.8.2
## [13] ROCR_1.0-7          gplots_2.17.0    caret_6.0-47
## [16] RColorBrewer_1.1-2  lubridate_1.3.3  magrittr_1.5
## [19] stringr_1.0.0       reshape2_1.4.1   Hmisc_3.16-0
## [22] ggplot2_1.0.1       Formula_1.2-1    survival_2.38-1
## [25] lattice_0.20-31     knitr_1.10.5     knitcitations_1.0.5
##
## loaded via a namespace (and not attached):
## [1] httr_0.6.1          splines_3.2.0     gtools_3.4.2
## [4] assertthat_0.1      highr_0.5         latticeExtra_0.6-26
## [7] yaml_2.1.13         quantreg_5.11     digest_0.6.8
## [10] RefManager_0.8.45   minqa_1.2.4       colorspace_1.2-6
## [13] htmltools_0.2.6     XML_3.98-1.1      bibtex_0.4.0
## [16] BradleyTerry2_1.0-6 SparseM_1.6        scales_0.2.4
## [19] gdata_2.16.1        brglm_0.5-9       lme4_1.1-7
## [22] combinat_0.0-8      mgcv_1.8-6        car_2.0-25
## [25] nnet_7.3-9          pbkrtest_0.4-2    proto_0.3-10
## [28] RJSONIO_1.3-0       memoise_0.2.1     evaluate_0.7
## [31] nlme_3.1-120        class_7.3-12      foreign_0.8-63
## [34] tools_3.2.0         RWekajars_3.7.12-1 formatR_1.2
## [37] munsell_0.4.2       cluster_2.0.1     e1071_1.6-4
## [40] caTools_1.17.1      RCurl_1.95-4.6    nloptr_1.0.4
## [43] bitops_1.0-6        rmarkdown_0.6.1   gtable_0.1.2
## [46] codetools_0.2-11    DBI_0.3.1         gridExtra_0.9.1
## [49] KernSmooth_2.23-14  rJava_0.9-6       stringi_0.4-1
## [52] Rcpp_0.11.6         acepack_1.3-3.3
```

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