Package 'coin'

July 18, 2017

Version 1.2-1				
Date 2017-07-16				
Title Conditional Inference Procedures in a Permutation Test Framework				
Description Conditional inference procedures for the general independence problem including two-sample, K-sample (non-parametric ANOVA), correlation, censored, ordered and multivariate problems.				
Depends R (>= 2.14.0), methods, survival				
Imports stats, modeltools (>= 0.2-9), mvtnorm (>= 1.0-5), multcomp				
Suggests parallel, xtable, e1071, vcd, TH.data (>= 1.0-7), libcoin (>= 0.9-0)				
LazyData yes				
NeedsCompilation yes				
ByteCompile yes				
License GPL-2				
<pre>URL http://coin.r-forge.r-project.org</pre>				
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Repository CRAN				
Date/Publication 2017-07-17 22:32:53 UTC				
R topics documented:				
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General Information on the coin Package

Description

The **coin** package provides an implementation of a general framework for conditional inference procedures commonly known as *permutation tests*. The framework was developed by Strasser and Weber (1999) and is based on a multivariate linear statistic and its conditional expectation, covariance and limiting distribution. These results are utilized to construct tests of independence between two sets of variables.

The package does not only provide a flexible implementation of the abstract framework, but also provides a large set of convenience functions implementing well-known as well as lesser-known classical and non-classical test procedures within the framework. Many of the tests presented in prominent text books, such as Hollander and Wolfe (1999) or Agresti (2002), are immediately available or can be implemented without much effort. Examples include linear rank statistics for the two- and K-sample location and scale problem against ordered and unordered alternatives including post-hoc tests for arbitrary contrasts, tests of independence for contingency tables, two- and K-sample tests for censored data, tests of independence between two continuous variables as well as tests of marginal homogeneity and symmetry. Approximations of the exact null distribution via the limiting distribution or conditional Monte Carlo resampling are available for every test procedure, while the exact null distribution is currently available for univariate two-sample problems only.

The salient parts of the Strasser-Weber framework are elucidated by Hothorn *et al.* (2006) and a thorough description of the software implementation is given by Hothorn *et al.* (2008).

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References

Agresti, A. (2002). *Categorical Data Analysis*, Second Edition. Hoboken, New Jersey: John Wiley & Sons.

Hollander, M. and Wolfe, D. A. (1999). *Nonparametric Statistical Methods*, Second Edition. New York: John Wiley & Sons.

Hothorn, T., Hornik, K., van de Wiel, M. A. and Zeileis, A. (2006). A Lego system for conditional inference. *The American Statistician* **60**(3), 257–263.

Hothorn, T., Hornik, K., van de Wiel, M. A. and Zeileis, A. (2008). Implementing a class of permutation tests: the coin package. *Journal of Statistical Software* **28**(8), 1–23. http://www.jstatsoft.org/v28/i08/

Strasser, H. and Weber, C. (1999). On the asymptotic theory of permutation statistics. *Mathematical Methods of Statistics* **8**(2), 220–250.

4 alpha

Examples

alpha

Genetic Components of Alcoholism

Description

Levels of expressed alpha synuclein mRNA in three groups of allele lengths of NACP-REP1.

Usage

alpha

Format

A data frame with 97 observations on 2 variables.

```
alength allele length, a factor with levels "short", "intermediate" and "long". elevel expression levels of alpha synuclein mRNA.
```

Details

Various studies have linked alcohol dependence phenotypes to chromosome 4. One candidate gene is NACP (non-amyloid component of plaques), coding for alpha synuclein. Bönsch *et al.* (2005) found longer alleles of NACP-REP1 in alcohol-dependent patients compared with healthy controls and reported that the allele lengths show some association with levels of expressed alpha synuclein mRNA.

Source

Bönsch, D., Lederer, T., Reulbach, U., Hothorn, T., Kornhuber, J. and Bleich, S. (2005). Joint analysis of the *NACP*-REP1 marker within the alpha synuclein gene concludes association with alcohol dependence. *Human Molecular Genetics* **14**(7), 967–971.

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References

Hothorn, T., Hornik, K., van de Wiel, M. A. and Zeileis, A. (2006). A Lego system for conditional inference. *The American Statistician* **60**(3), 257–263.

Examples

```
## Boxplots
boxplot(elevel ~ alength, data = alpha)
## Asymptotic Kruskal-Wallis test
kruskal_test(elevel ~ alength, data = alpha)
```

alzheimer

Smoking and Alzheimer's Disease

Description

A case-control study of smoking and Alzheimer's disease.

Usage

alzheimer

Format

A data frame with 538 observations on 3 variables.

```
smoking a factor with levels "None", "<10", "10-20" and ">20" (cigarettes per day). disease a factor with levels "Alzheimer", "Other dementias" and "Other diagnoses". gender a factor with levels "Female" and "Male".
```

Details

Subjects with Alzheimer's disease are compared to two different control groups with respect to smoking history. The data are given in Salib and Hillier (1997, Tab. 4).

Source

Salib, E. and Hillier, V. (1997). A case-control study of smoking and Alzheimer's disease. *International Journal of Geriatric Psychiatry* **12**(3), 295–300.

References

Hothorn, T., Hornik, K., van de Wiel, M. A. and Zeileis, A. (2006). A Lego system for conditional inference. *The American Statistician* **60**(3), 257–263.

6 asat

Examples

asat

Toxicological Study on Female Wistar Rats

Description

Measurements of the liver enzyme aspartate aminotransferase (ASAT) for a new compound and a control group of 34 female Wistar rats.

Usage

asat

Format

A data frame with 34 observations on 2 variables.

```
asat ASAT values.
group a factor with levels "Compound" and "Control".
```

Details

The aim of this toxicological study is the proof of safety for the new compound. The data were originally given in Hothorn (1992) and later reproduced by Hauschke, Kieser and Hothorn (1999).

Source

Hauschke, D., Kieser, M. and Hothorn, L. A. (1999). Proof of safety in toxicology based on the ratio of two means for normally distributed data. *Biometrical Journal* **41**(3), 295–304.

Hothorn, L. A. (1992). Biometrische analyse toxikologischer untersuchungen. In J. Adam (Ed.), *Statistisches Know-How in der Medizinischen Forschung*, pp. 475–590. Berlin: Ullstein Mosby.

References

Pflüger, R. and Hothorn, T. (2002). Assessing equivalence tests with respect to their expected *p*-value. *Biometrical Journal* **44**(8), 1015–1027.

Examples

ContingencyTests

Tests of Independence in Two- or Three-Way Contingency Tables

Description

Testing the independence of two nominal or ordered factors.

Usage

```
## S3 method for class 'formula'
chisq_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'table'
chisq_test(object, ...)
## S3 method for class 'IndependenceProblem'
chisq_test(object, ...)
## S3 method for class 'formula'
cmh_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'table'
cmh_test(object, ...)
## S3 method for class 'IndependenceProblem'
cmh_test(object, ...)
## S3 method for class 'formula'
lbl_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'table'
lbl_test(object, ...)
## S3 method for class 'IndependenceProblem'
lbl_test(object, distribution = c("asymptotic", "approximate", "none"), ...)
```

Arguments

formula a formula of the form y ~ x | block where y and x are factors and block is an optional factor for stratification.

data an optional data frame containing the variables in the model formula.

subset an optional vector specifying a subset of observations to be used. Defaults to

NULL.

weights an optional formula of the form ~ w defining integer valued case weights for

each observation. Defaults to NULL, implying equal weight for all observations.

object an object inheriting from classes "table" or "IndependenceProblem".

distribution a character, the conditional null distribution of the test statistic can be approx-

imated by its asymptotic distribution ("asymptotic", default) or via Monte Carlo resampling ("approximate"). Alternatively, the functions asymptotic or approximate can be used. Computation of the null distribution can be sup-

pressed by specifying "none".

... further arguments to be passed to independence_test.

Details

chisq_test, cmh_test and lbl_test provide the Pearson chi-squared test, the generalized Cochran-Mantel-Haenszel test and the linear-by-linear association test. A general description of these methods is given by Agresti (2002).

The null hypothesis of independence, or conditional independence given block, between y and x is tested.

If y and/or x are ordered factors, the default scores, 1:nlevels(y) and 1:nlevels(x) respectively, can be altered using the scores argument (see independence_test); this argument can also be used to coerce nominal factors to class "ordered". (lbl_test coerces to class "ordered" under any circumstances.) If both y and x are ordered factors, a linear-by-linear association test is computed and the direction of the alternative hypothesis can be specified using the alternative argument. For the Pearson chi-squared test, this extension was given by Yates (1948) who also discussed the situation when either the response or the covariate is an ordered factor; see also Cochran (1954) and Armitage (1955) for the particular case when y is a binary factor and x is ordered. The Mantel-Haenszel statistic was similarly extended by Mantel (1963) and Landis, Heyman and Koch (1978).

The conditional null distribution of the test statistic is used to obtain *p*-values and an asymptotic approximation of the exact distribution is used by default (distribution = "asymptotic"). Alternatively, the distribution can be approximated via Monte Carlo resampling or computed exactly for univariate two-sample problems by setting distribution to "approximate" or "exact" respectively. See asymptotic, approximate and exact for details.

Value

An object inheriting from class "IndependenceTest".

Note

The exact versions of the Pearson chi-squared test and the generalized Cochran-Mantel-Haenszel test do not necessarily result in the same p-value as Fisher's exact test (Davis, 1986).

References

Agresti, A. (2002). *Categorical Data Analysis*, Second Edition. Hoboken, New Jersey: John Wiley & Sons.

Armitage, P. (1955). Tests for linear trends in proportions and frequencies. *Biometrics* **11**(3), 375–386.

Cochran, W.G. (1954). Some methods for strengthening the common χ^2 tests. *Biometrics* **10**(4), 417–451.

Davis, L. J. (1986). Exact tests for 2×2 contingency tables. *The American Statistician* **40**(2), 139–141.

Landis, J. R., Heyman, E. R. and Koch, G. G. (1978). Average partial association in three-way contingency tables: a review and discussion of alternative tests. *International Statistical Review* **46**(3), 237–254.

Mantel, N. (1963). Chi-square tests with one degree of freedom: extensions of the Mantel-Haenszel procedure. *Journal of the American Statistical Association* **58**(303), 690–700.

Yates, F. (1948). The analysis of contingency tables with groupings based on quantitative characters. *Biometrika* **35**(1/2), 176–181.

Examples

```
## Example data
## Davis (1986, p. 140)
davis <- matrix(</pre>
   c(3, 6,
      2, 19),
   nrow = 2, byrow = TRUE
davis <- as.table(davis)</pre>
## Asymptotic Pearson chi-squared test
chisq_test(davis)
## Approximative (Monte Carlo) Pearson chi-squared test
ct <- chisq_test(davis,
                distribution = approximate(B = 10000))
                 # standard p-value
pvalue(ct)
midpvalue(ct)
                   # mid-p-value
pvalue_interval(ct) # p-value interval
## Exact Pearson chi-squared test (Davis, 1986)
## Note: disagrees with Fisher's exact test
ct <- chisq_test(davis,
               distribution = "exact")
pvalue(ct)
                 # standard p-value
midpvalue(ct)
                   # mid-p-value
pvalue_interval(ct) # p-value interval
fisher.test(davis)
## Laryngeal cancer data
```

```
## Agresti (2002, p. 107, Tab. 3.13)
cancer <- matrix(</pre>
    c(21, 2,
      15, 3),
    nrow = 2, byrow = TRUE,
    dimnames = list(
        "Treatment" = c("Surgery", "Radiation"),
           "Cancer" = c("Controlled", "Not Controlled")
)
cancer <- as.table(cancer)</pre>
## Exact Pearson chi-squared test (Agresti, 2002, p. 108, Tab. 3.14)
## Note: agrees with Fishers's exact test
(ct <- chisq_test(cancer,</pre>
                   distribution = "exact"))
midpvalue(ct)
                    # mid-p-value
pvalue_interval(ct) # p-value interval
fisher.test(cancer)
## Homework conditions and teacher's rating
## Yates (1948, Tab. 1)
yates <- matrix(</pre>
    c(141, 67, 114, 79, 39,
      131, 66, 143, 72, 35,
       36, 14, 38, 28, 16),
    byrow = TRUE, ncol = 5,
    dimnames = list(
        "Rating" = c("A", "B", "C"),
"Condition" = c("A", "B", "C", "D", "E")
    )
)
yates <- as.table(yates)</pre>
## Asymptotic Pearson chi-squared test (Yates, 1948, p. 176)
chisq_test(yates)
## Asymptotic Pearson-Yates chi-squared test (Yates, 1948, pp. 180-181)
## Note: 'Rating' and 'Condition' as ordinal
(ct <- chisq_test(yates,</pre>
                   alternative = "less",
                   scores = list("Rating" = c(-1, 0, 1),
                                  "Condition" = c(2, 1, 0, -1, -2)))
statistic(ct)^2 # chi^2 = 2.332
## Asymptotic Pearson-Yates chi-squared test (Yates, 1948, p. 181)
## Note: 'Rating' as ordinal
chisq_test(yates,
           scores = list("Rating" = c(-1, 0, 1))) # Q = 3.825
```

Change in clinical condition and degree of infiltration

```
## Cochran (1954, Tab. 6)
cochran <- matrix(</pre>
   c(11, 7,
     27, 15,
      42, 16,
      53, 13,
     11, 1),
    byrow = TRUE, ncol = 2,
    dimnames = list(
              "Change" = c("Marked", "Moderate", "Slight",
                            "Stationary", "Worse"),
        "Infiltration" = c("0-7", "8-15")
cochran <- as.table(cochran)</pre>
## Asymptotic Pearson chi-squared test (Cochran, 1954, p. 435)
chisq_test(cochran) # X^2 = 6.88
## Asymptotic Cochran-Armitage test (Cochran, 1954, p. 436)
## Note: 'Change' as ordinal
(ct <- chisq_test(cochran,</pre>
                  scores = list("Change" = c(3, 2, 1, 0, -1))))
statistic(ct)^2 # X^2 = 6.66
## Change in size of ulcer crater for two treatment groups
## Armitage (1955, Tab. 2)
armitage <- matrix(</pre>
   c( 6, 4, 10, 12,
     11, 8, 8, 5),
   byrow = TRUE, ncol = 4,
    dimnames = list(
        "Treatment" = c("A", "B"),
           "Crater" = c("Larger", "< 2/3 healed",
                        "=> 2/3 healed", "Healed")
)
armitage <- as.table(armitage)</pre>
## Approximative (Monte Carlo) Pearson chi-squared test (Armitage, 1955, p. 379)
chisq_test(armitage,
           distribution = approximate(B = 10000)) # chi^2 = 5.91
## Approximative (Monte Carlo) Cochran-Armitage test (Armitage, 1955, p. 379)
(ct <- chisq_test(armitage,</pre>
                  distribution = approximate(B = 10000),
                  scores = list("Crater" = c(-1.5, -0.5, 0.5, 1.5))))
statistic(ct)^2 # chi_0^2 = 5.26
## Relationship between job satisfaction and income stratified by gender
## Agresti (2002, p. 288, Tab. 7.8)
```

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CorrelationTests

Correlation Tests

Description

Testing the independence of two numeric variables.

Usage

```
## S3 method for class 'formula'
spearman_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'IndependenceProblem'
spearman_test(object, distribution = c("asymptotic", "approximate", "none"), ...)
## S3 method for class 'formula'
fisyat_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'IndependenceProblem'
fisyat_test(object, distribution = c("asymptotic", "approximate", "none"),
            ties.method = c("mid-ranks", "average-scores"), ...)
## S3 method for class 'formula'
quadrant_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'IndependenceProblem'
quadrant_test(object, distribution = c("asymptotic", "approximate", "none"),
              mid.score = c("0", "0.5", "1"), ...)
## S3 method for class 'formula'
koziol_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'IndependenceProblem'
koziol_test(object, distribution = c("asymptotic", "approximate", "none"),
            ties.method = c("mid-ranks", "average-scores"), ...)
```

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Arguments

formula	a formula of the form y $^\sim$ x $$ block where y and x are numeric variables and block is an optional factor for stratification.
data	an optional data frame containing the variables in the model formula.
subset	an optional vector specifying a subset of observations to be used. Defaults to \ensuremath{NULL} .
weights	an optional formula of the form \sim w defining integer valued case weights for each observation. Defaults to NULL, implying equal weight for all observations.
object	an object inheriting from class "IndependenceProblem".
distribution	a character, the conditional null distribution of the test statistic can be approximated by its asymptotic distribution ("asymptotic", default) or via Monte Carlo resampling ("approximate"). Alternatively, the functions asymptotic or approximate can be used. Computation of the null distribution can be suppressed by specifying "none".
ties.method	a character, the method used to handle ties: the score generating function either uses mid-ranks ("mid-ranks", default) or averages the scores of randomly broken ties ("average-scores").
mid.score	a character, the score assigned to observations exactly equal to the median: either 0 ("0", default), 0.5 ("0.5") or 1 ("1"); see median_test.
	further arguments to be passed to independence_test.

Details

spearman_test, fisyat_test, quadrant_test and koziol_test provide the Spearman correlation test, the Fisher-Yates correlation test using van der Waerden scores, the quadrant test and the Koziol-Nemec test. A general description of these methods is given by Hájek, Šidák and Sen (1999, Sec. 4.6). The Koziol-Nemec test was suggested by Koziol and Nemec (1979). For the adjustment of scores for tied values see Hájek, Šidák and Sen (1999, pp. 133–135).

The null hypothesis of independence, or conditional independence given block, between y and x is tested.

The conditional null distribution of the test statistic is used to obtain *p*-values and an asymptotic approximation of the exact distribution is used by default (distribution = "asymptotic"). Alternatively, the distribution can be approximated via Monte Carlo resampling by setting distribution to "approximate". See asymptotic and approximate for details.

Value

An object inheriting from class "IndependenceTest".

References

Hájek, J., Šidák, Z. and Sen, P. K. (1999). *Theory of Rank Tests*, Second Edition. San Diego: Academic Press.

Koziol, J. A. and Nemec, A. F. (1979). On a Cramér-von Mises type statistic for testing bivariate independence. *The Canadian Journal of Statistics* **7**(1), 43–52.

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Examples

```
## Asymptotic Spearman test
spearman_test(CONT ~ INTG, data = USJudgeRatings)
## Asymptotic Fisher-Yates test
fisyat_test(CONT ~ INTG, data = USJudgeRatings)
## Asymptotic quadrant test
quadrant_test(CONT ~ INTG, data = USJudgeRatings)
## Asymptotic Koziol-Nemec test
koziol_test(CONT ~ INTG, data = USJudgeRatings)
```

CWD

Coarse Woody Debris

Description

Carbon flux on six pieces of wood.

Usage

CWD

Format

A data frame with 13 observations on 8 variables.

sample2 carbon flux measurement for 2nd piece of wood. sample3 carbon flux measurement for 3rd piece of wood. sample4 carbon flux measurement for 4th piece of wood. sample6 carbon flux measurement for 6th piece of wood. sample7 carbon flux measurement for 7th piece of wood. sample8 carbon flux measurement for 8th piece of wood. trend measurement day (in days from beginning). time date of measurement.

Details

Coarse woody debris (CWD, dead wood greater than 10 cm in diameter) is a large stock of carbon in tropical forests, yet the flux of carbon out of this pool, via respiration, is poorly resolved (Chambers, Schimel and Nobre, 2001). The heterotrophic process involved in CWD respiration should respond to reductions in moisture availability, which occurs during dry season (Chambers, Schimel and Nobre, 2001).

CWD respiration measurements were taken in a tropical forest in west French Guiana, which experiences extreme contrasts in wet and dry season (Bonal et al., 2008). An infrared gas analyzer

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and a clear chamber sealed to the wood surface were used to measure the flux of carbon out of the wood (Stahl *et al.*, 2011). Measurements were repeated 13 times, from July to November 2011, on six pieces of wood during the transition into and out of the dry season. The aim is to assess if there were shifts in the CWD respiration of any of the pieces in response to the transition into (early August) and out of (late October) the dry season.

Zeileis and Hothorn (2013) investigated the six-variate series of CO_2 reflux, aiming to find out whether the reflux had changed over the sampling period in at least one of the six wood pieces.

Source

The coarse woody debris respiration data were kindly provided by Lucy Rowland (School of Geo-Sciences, University of Edinburgh).

References

Bonal, D., Bosc, A., Ponton, S., Goret, J.-Y., Burban, B., Gross, P., Bonnefond, J.-M., Elbers, J., Longdoz, B., Epron, D., Guehl, J.-M. and Granier, A. (2008). Impact of severe dry season on net ecosystem exchange in the Neotropical rainforest of French Guiana. *Global Change Biology* **14**(8), 1917–1933.

Chambers, J. Q., Schimel, J. P. and Nobre, A. D. (2001). Respiration from coarse wood litter in central Amazon forests. *Biogeochemistry* **52**(2), 115–131.

Stahl, C., Burban, B., Goret, J.-Y. and Bonal, D. (2011). Seasonal variations in stem CO₂ efflux in the Neotropical rainforest of French Guiana. *Annals of Forest Science* **68**(4), 771–782.

Zeileis, A. and Hothorn, T. (2013). A toolbox of permutation tests for structural change. *Statistical Papers* **54**(4), 931–954.

Examples

16 expectation-methods

 ${\it expectation-methods} \qquad {\it Extraction of the Expectation, Variance and Covariance of the Linear} \\ {\it Statistic}$

Description

Methods for extraction of the expectation, variance and covariance of the linear statistic.

Usage

```
## S4 method for signature 'IndependenceLinearStatistic'
expectation(object, ...)
## S4 method for signature 'IndependenceTest'
expectation(object, ...)
## S4 method for signature 'CovarianceMatrix'
variance(object, ...)
## S4 method for signature 'IndependenceLinearStatistic'
variance(object, ...)
## S4 method for signature 'IndependenceTest'
variance(object, ...)
## S4 method for signature 'Variance'
variance(object, ...)
## S4 method for signature 'CovarianceMatrix'
covariance(object, ...)
## S4 method for signature 'IndependenceLinearStatistic'
covariance(object, ...)
## S4 method for signature 'IndependenceTest'
covariance(object, ...)
```

Arguments

object an object from which the expectation, variance or covariance of the linear statistic can be extracted.
... further arguments (currently ignored).

Details

The methods expectation, variance and covariance extract the expectation, variance and covariance, respectively, of the linear statistic.

Value

The expectation, variance or covariance of the linear statistic extracted from object. A numeric vector or matrix.

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Examples

```
## Example data
dta <- data.frame(</pre>
    y = g1(3, 2),
    x = sample(gl(3, 2))
)
## Asymptotic Cochran-Mantel-Haenszel Test
ct <- cmh_test(y \sim x, data = dta)
## The linear statistic, i.e., the contingency table...
(l <- statistic(ct, type = "linear"))</pre>
## ...and its expectation...
(El <- expectation(ct))</pre>
## ...and covariance
(Vl <- covariance(ct))</pre>
## The standardized contingency table...
(1 - El) / sqrt(variance(ct))
## ...is identical to the standardized linear statistic
statistic(ct, type = "standardized")
```

ExpectCovar-class

Class "ExpectCovar"

Description

Objects of class "ExpectCovar" represent the expectation and covariance of the linear statistic.

Objects from the Class

Objects can be created by calls of the form

```
new("ExpectCovar", ...).
```

Slots

```
expectation: Object of class "numeric". The expectation of the linear statistic. covariance: Object of class "matrix". The covariance matrix of the linear statistic. dimension: Object of class "integer". The dimension of the linear statistic.
```

Known Subclasses

```
Class "ExpectCovarInfluence", directly.
```

Methods

initialize signature(.Object = "ExpectCovar"): See the documentation for initialize (in package methods) for details.

Note

Normally, this class is used internally only.

ExpectCovarInfluence-class

Class "ExpectCovarInfluence"

Description

Objects of class "ExpectCovarInfluence" represent the expectation and covariance of the influence function.

Objects from the Class

Objects can be created by calls of the form

```
new("ExpectCovarInfluence", ...).
```

Slots

```
sumweights: Object of class "numeric". The sum of the case weights.
expectation: Object of class "numeric". The expectation of the influence function.
covariance: Object of class "matrix". The covariance matrix of the influence function.
dimension: Object of class "integer". The dimension of the influence function.
```

Extends

```
Class "ExpectCovar", directly.
```

Methods

initialize signature(.Object = "ExpectCovarInfluence"): See the documentation for initialize
 (in package methods) for details.

Note

Normally, this class is used internally only.

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glioma

Malignant Glioma Pilot Study

Description

A non-randomized pilot study on malignant glioma patients with pretargeted adjuvant radioim-munotherapy using yttrium-90-biotin.

Usage

glioma

Format

A data frame with 37 observations on 7 variables.

```
no. patient number.
age patient age (years).
sex a factor with levels "F" (Female) and "M" (Male).
histology a factor with levels "GBM" (grade IV) and "Grade3" (grade III).
group a factor with levels "Control" and "RIT".
event status indicator for time: FALSE for right-censored observations and TRUE otherwise.
time survival time (months).
```

Details

The primary endpoint of this small pilot study is survival. Since the survival times are tied, the classical asymptotic logrank test may be inadequate in this setup. Therefore, a permutation test using Monte Carlo resampling was computed in the original paper. The data are taken from Tables 1 and 2 of Grana *et al.* (2002).

Source

Grana, C., Chinol, M., Robertson, C., Mazzetta, C., Bartolomei, M., De Cicco, C., Fiorenza, M., Gatti, M., Caliceti, P. and Paganelli, G. (2002). Pretargeted adjuvant radioimmunotherapy with Yttrium-90-biotin in malignant glioma patients: A pilot study. *British Journal of Cancer* **86**(2), 207–212.

Examples

```
## Grade III glioma
g3 <- subset(glioma, histology == "Grade3")

## Plot Kaplan-Meier estimates
op <- par(no.readonly = TRUE) # save current settings
layout(matrix(1:2, ncol = 2))
plot(survfit(Surv(time, event) ~ group, data = g3),</pre>
```

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```
main = "Grade III Glioma", lty = 2:1,
    ylab = "Probability", xlab = "Survival Time in Month",
    xlim = c(-2, 72)
legend("bottomleft", lty = 2:1, c("Control", "Treated"), bty = "n")
## Exact logrank test
logrank_test(Surv(time, event) ~ group, data = g3,
             distribution = "exact")
## Grade IV glioma
gbm <- subset(glioma, histology == "GBM")</pre>
## Plot Kaplan-Meier estimates
plot(survfit(Surv(time, event) ~ group, data = gbm),
    main = "Grade IV Glioma", lty = 2:1,
    ylab = "Probability", xlab = "Survival Time in Month",
    xlim = c(-2, 72)
legend("topright", lty = 2:1, c("Control", "Treated"), bty = "n")
par(op) # reset
## Exact logrank test
logrank_test(Surv(time, event) ~ group, data = gbm,
             distribution = "exact")
## Stratified approximative (Monte Carlo) logrank test
logrank_test(Surv(time, event) ~ group | histology, data = glioma,
             distribution = approximate(B = 10000))
```

GTSG

Gastrointestinal Tumor Study Group

Description

A randomized clinical trial in gastric cancer.

Usage

GTSG

Format

A data frame with 90 observations on 3 variables.

```
time survival time (days).

event status indicator for time: 0 for right-censored observations and 1 otherwise.

group a factor with levels "Chemotherapy+Radiation" and "Chemotherapy".
```

GTSG 21

Details

A clinical trial comparing chemotherapy alone versus a combination of chemotherapy and radiation therapy in the treatment of locally advanced, nonresectable gastric carcinoma.

Note

There is substantial separation between the estimated survival distributions at 8 to 10 months, but by month 26 the distributions intersect.

Source

Stablein, D. M., Carter, W. H., Jr. and Novak, J. W. (1981). Analysis of survival data with nonproportional hazard functions. *Controlled Clinical Trials* **2**(2), 149–159.

References

Moreau, T., Maccario, J., Lellouch, J. and Huber, C. (1992). Weighted log rank statistics for comparing two distributions. *Biometrika* **79**(1), 195–198.

Shen, W. and Le, C. T. (2000). Linear rank tests for censored survival data. *Communications in Statistics – Simulation and Computation* **29**(1), 21–36.

Tubert-Bitter, P., Kramar, A., Chalé, J. J. and Moureau, T. (1994). Linear rank tests for comparing survival in two groups with crossing hazards. *Computational Statistics & Data Analysis* **18**(5), 547–559.

Examples

```
## Plot Kaplan-Meier estimates
plot(survfit(Surv(time / (365.25 / 12), event) \sim group, data = GTSG),
     lty = 1:2, ylab = "% Survival", xlab = "Survival Time in Months")
legend("topright", lty = 1:2,
       c("Chemotherapy+Radiation", "Chemotherapy"), bty = "n")
## Asymptotic logrank test
logrank_test(Surv(time, event) ~ group, data = GTSG)
## Asymptotic Prentice test
logrank_test(Surv(time, event) ~ group, data = GTSG, type = "Prentice")
## Asymptotic test against Weibull-type alternatives (Moreau et al., 1992)
moreau_weight <- function(time, n.risk, n.event)</pre>
    1 + log(-log(cumprod(n.risk / (n.risk + n.event))))
independence_test(Surv(time, event) ~ group, data = GTSG,
                  ytrafo = function(data)
                      trafo(data, surv_trafo = function(y)
                          logrank_trafo(y, weight = moreau_weight)))
## Asymptotic test against crossing-curve alternatives (Shen and Le, 2000)
shen_trafo <- function(x)</pre>
   ansari_trafo(logrank_trafo(x, type = "Prentice"))
```

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hohnloser

Left Ventricular Ejection Fraction

Description

Left ventricular ejection fraction in patients with malignant ventricular tachyarrhythmias including recurrence-free month and censoring.

Usage

hohnloser

Format

A data frame with 94 observations on 3 variables.

```
EF ejection fraction (%).
```

time recurrence-free month.

event status indicator for time: 0 for right-censored observations and 1 otherwise.

Details

The data was used by Lausen and Schumacher (1992) to illustrate the use of maximally selected statistics.

Source

Hohnloser, S. H., Raeder, E. A., Podrid, P. J., Graboys, T. B. and Lown, B. (1987). Predictors of antiarrhythmic drug efficacy in patients with malignant ventricular tachyarrhythmias. *American Heart Journal* **114**(1 Pt 1), 1–7.

References

Lausen, B. and Schumacher, M. (1992). Maximally selected rank statistics. *Biometrics* **48**(1), 73–85.

Examples

```
## Asymptotic maximally selected logrank statistics
maxstat_test(Surv(time, event) ~ EF, data = hohnloser)
```

```
IndependenceLinearStatistic-class
```

Class "IndependenceLinearStatistic"

Description

Objects of class "IndependenceLinearStatistic" represent the linear statistic and the transformed and original data structures corresponding to an independence problem.

Objects from the Class

Objects can be created by calls of the form

```
new("IndependenceLinearStatistic", object, varonly = FALSE, ...)
```

where object is an object of class "IndependenceTestProblem", varonly is a logical indicating that slot covariance (see 'Slots') should only contain the diagonal elements of the covariance matrix.

Slots

```
linearstatistic: Object of class "numeric". The linear statistic.
expectation: Object of class "Numeric". The expectation of the linear statistic.
covariance: Object of class "VarCovar". The covariance or variance of the linear statistic.
xtrans: Object of class "matrix". The transformed x.
ytrans: Object of class "matrix". The transformed y.
xtrafo: Object of class "function". The regression function for x.
ytrafo: Object of class "function". The influence function for y.
x: Object of class "data.frame". The variables x.
y: Object of class "data.frame". The variables y.
block: Object of class "factor". The block structure.
weights: Object of class "numeric". The case weights.
```

Extends

```
Class "IndependenceTestProblem", directly.
Class "IndependenceProblem", by class "IndependenceTestProblem", distance 2.
```

Known Subclasses

```
Class "IndependenceTestStatistic", directly.
Class "MaxTypeIndependenceTestStatistic", by class "IndependenceTestStatistic", distance 2.
Class "QuadTypeIndependenceTestStatistic", by class "IndependenceTestStatistic", distance 2.
Class "ScalarIndependenceTestStatistic", by class "IndependenceTestStatistic", distance 2.
```

Methods

```
covariance signature(object = "IndependenceLinearStatistic"): See the documentation
for covariance for details.
```

expectation signature(object = "IndependenceLinearStatistic"): See the documentation
for expectation for details.

initialize signature(.Object = "IndependenceLinearStatistic"): See the documentation
 for initialize (in package methods) for details.

statistic signature(object = "IndependenceLinearStatistic"): See the documentation for statistic for details.

variance signature(object = "IndependenceLinearStatistic"): See the documentation for variance for details.

IndependenceProblem-class

Class "IndependenceProblem"

Description

Objects of class "IndependenceProblem" represent the data structure corresponding to an independence problem.

Objects from the Class

Objects can be created by calls of the form

```
new("IndependenceProblem", x, y, block = NULL, weights = NULL, ...)
```

where x and y are data frames containing the variables X and Y respectively, block is an optional factor representing the block structure b and weights is an optional integer vector corresponding to the case weights w.

Slots

```
x: Object of class "data.frame". The variables x.y: Object of class "data.frame". The variables y.block: Object of class "factor". The block structure.weights: Object of class "numeric". The case weights.
```

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Known Subclasses

```
Class "IndependenceTestProblem", directly.
Class "SymmetryProblem", directly.
Class "IndependenceLinearStatistic", by class "IndependenceTestProblem", distance 2.
Class "IndependenceTestStatistic", by class "IndependenceTestProblem", distance 3.
Class "MaxTypeIndependenceTestStatistic", by class "IndependenceTestProblem", distance 4.
Class "QuadTypeIndependenceTestStatistic", by class "IndependenceTestProblem", distance 4.
Class "ScalarIndependenceTestStatistic", by class "IndependenceTestProblem", distance 4.
```

Methods

initialize signature(.Object = "IndependenceProblem"): See the documentation for initialize
 (in package methods) for details.

IndependenceTest

General Independence Test

Description

Testing the independence of two sets of variables measured on arbitrary scales.

Usage

Arguments

formula	a formula of the form $y1 + \ldots + yq \sim x1 + \ldots + xp \mid block$ where $y1, \ldots, yq$ and $x1, \ldots, xp$ are measured on arbitrary scales (nominal, ordinal or continuous with or without censoring) and block is an optional factor for stratification.
data	an optional data frame containing the variables in the model formula.
subset	an optional vector specifying a subset of observations to be used. Defaults to NULL.

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weights	an optional formula of the form ~ w defining integer valued case weights for each observation. Defaults to NULL, implying equal weight for all observations.
object	an object inheriting from classes "table" or "IndependenceProblem".
teststat	a character, the type of test statistic to be applied: either a maximum statistic ("maximum", default), a quadratic form ("quadratic") or a standardized scalar test statistic ("scalar").
distribution	a character, the conditional null distribution of the test statistic can be approximated by its asymptotic distribution ("asymptotic", default) or via Monte Carlo resampling ("approximate"). Alternatively, the functions asymptotic or approximate can be used. For univariate two-sample problems, "exact" or use of the function exact computes the exact distribution. Computation of the null distribution can be suppressed by specifying "none". It is also possible to specify a function with one argument (an object inheriting from "IndependenceTestStatistic") that returns an object of class "NullDistribution"
alternative	a character, the alternative hypothesis: either "two.sided" (default), "greater" or "less".
xtrafo	a function of transformations to be applied to the variables $x1, \ldots, xp$ supplied in formula; see 'Details'. Defaults to trafo.
ytrafo	a function of transformations to be applied to the variables y1,, yq supplied in formula; see 'Details'. Defaults to trafo.
scores	a named list of scores to be attached to ordered factors; see 'Details'. Defaults to NULL, implying equally spaced scores.
check	a function to be applied to objects of class "IndependenceTest" in order to check for specific properties of the data. Defaults to NULL.
	further arguments to be passed to or from other methods (currently ignored).

Details

independence_test provides a general independence test for two sets of variables measured on arbitrary scales. This function is based on the general framework for conditional inference procedures proposed by Strasser and Weber (1999). The salient parts of the Strasser-Weber framework are elucidated by Hothorn *et al.* (2006) and a thorough description of the software implementation is given by Hothorn *et al.* (2008).

The null hypothesis of independence, or conditional independence given block, between y1, ..., yq and x1, ..., xp is tested.

A vector of case weights, e.g., observation counts, can be supplied through the weights argument and the type of test statistic is specified by the teststat argument. Influence and regression functions, i.e., transformations of y1, ..., yq and x1, ..., xp, are specified by the ytrafo and xtrafo arguments respectively; see trafo for the collection of transformation functions currently available. This allows for implementation of both novel and familiar test statistics, e.g., the Pearson χ^2 test, the generalized Cochran-Mantel-Haenszel test, the Spearman correlation test, the Fisher-Pitman permutation test, the Wilcoxon-Mann-Whitney test, the Kruskal-Wallis test and the family of weighted logrank tests for censored data. Furthermore, multivariate extensions such as the multivariate Kruskal-Wallis test (Puri and Sen, 1966, 1971) can be implemented without much effort (see 'Examples').

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If, say, y1 and/or x1 are ordered factors, the default scores, 1:nlevels(y1) and 1:nlevels(x1) respectively, can be altered using the scores argument; this argument can also be used to coerce nominal factors to class "ordered". For example, when y1 is an ordered factor with four levels and x1 is a nominal factor with three levels, scores = list(y1 = c(1, 3:5), x1 = c(1:2, 4)) supplies the scores to be used. For ordered alternatives the scores must be monotonic, but non-montonic scores are also allowed for testing against, e.g., umbrella alternatives. The length of the score vector must be equal to the number of factor levels.

The conditional null distribution of the test statistic is used to obtain *p*-values and an asymptotic approximation of the exact distribution is used by default (distribution = "asymptotic"). Alternatively, the distribution can be approximated via Monte Carlo resampling or computed exactly for univariate two-sample problems by setting distribution to "approximate" or "exact" respectively. See asymptotic, approximate and exact for details.

Value

An object inheriting from class "IndependenceTest".

Note

Starting with **coin** version 1.1-0, maximum statistics and quadratic forms can no longer be specified using teststat = "maxtype" and teststat = "quadtype" respectively (as was used in versions prior to 0.4-5).

References

Hothorn, T., Hornik, K., van de Wiel, M. A. and Zeileis, A. (2006). A Lego system for conditional inference. *The American Statistician* **60**(3), 257–263.

Hothorn, T., Hornik, K., van de Wiel, M. A. and Zeileis, A. (2008). Implementing a class of permutation tests: the coin package. *Journal of Statistical Software* **28**(8), 1–23. http://www.jstatsoft.org/v28/i08/

Johnson, W. D., Mercante, D. E. and May, W. L. (1993). A computer package for the multivariate nonparametric rank test in completely randomized experimental designs. *Computer Methods and Programs in Biomedicine* **40**(3), 217–225.

Puri, M. L. and Sen, P. K. (1966). On a class of multivariate multisample rank order tests. *Sankhya* A **28**(4), 353–376.

Puri, M. L. and Sen, P. K. (1971). *Nonparametric Methods in Multivariate Analysis*. New York: John Wiley & Sons.

Strasser, H. and Weber, C. (1999). On the asymptotic theory of permutation statistics. *Mathematical Methods of Statistics* **8**(2), 220–250.

Examples

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```
## apply normal scores to asat$asat
                  ytrafo = function(data)
                      trafo(data, numeric_trafo = normal_trafo),
                  ## indicator matrix of 1st level of asat$group
                  xtrafo = function(data)
                      trafo(data, factor_trafo = function(x)
                         matrix(x == levels(x)[1], ncol = 1)))
## ...or more conveniently
normal_test(asat ~ group, data = asat,
            ## exact null distribution
            distribution = "exact",
            ## one-sided test
            alternative = "greater")
## Receptor binding assay of benzodiazepines
## Johnson, Mercante and May (1993, Tab. 1)
benzos <- data.frame(</pre>
      cerebellum = c(3.41, 3.50, 2.85, 4.43,
                      4.04, 7.40, 5.63, 12.86,
                      6.03, 6.08, 5.75, 8.09, 7.56),
       brainstem = c( 3.46, 2.73, 2.22, 3.16,
                     2.59, 4.18, 3.10, 4.49,
                      6.78, 7.54, 5.29, 4.57,
                                                 5.39),
          cortex = c(10.52, 7.52, 4.57, 5.48,
                     7.16, 12.00, 9.36, 9.35,
                     11.54, 11.05, 9.92, 13.59, 13.21),
    hypothalamus = c(19.51, 10.00, 8.27, 10.26,
                     11.43, 19.13, 14.03, 15.59,
                     24.87, 14.16, 22.68, 19.93, 29.32),
        striatum = c(6.98, 5.07, 3.57, 5.34,
                      4.57, 8.82, 5.76, 11.72,
                      6.98, 7.54, 7.66, 9.69, 8.09),
     hippocampus = c(20.31, 13.20, 8.58, 11.42,
                     13.79, 23.71, 18.35, 38.52,
                     21.56, 18.66, 19.24, 27.39, 26.55),
       treatment = factor(rep(c("Lorazepam", "Alprazolam", "Saline"),
                          c(4, 4, 5)))
)
## Approximative (Monte Carlo) multivariate Kruskal-Wallis test
## Johnson, Mercante and May (1993, Tab. 2)
independence_test(cerebellum + brainstem + cortex +
                  hypothalamus + striatum + hippocampus ~ treatment,
                  data = benzos,
                  teststat = "quadratic",
                  distribution = approximate(B = 10000),
                  ytrafo = function(data)
                      trafo(data, numeric_trafo = rank_trafo)) # Q = 16.129
```

IndependenceTest-class

Class "IndependenceTest" and its subclasses

Description

Objects of class "IndependenceTest" and its subclasses "MaxTypeIndependenceTest", "QuadTypeIndependenceTest", "ScalarIndependenceTestConfint" represent an independence test including its original and transformed data structure, linear statistic, test statistic and reference distribution.

Objects from the Class

```
new("IndependenceTest", ...),
new("MaxTypeIndependenceTest", ...),
new("QuadTypeIndependenceTest", ...),
new("ScalarIndependenceTest", ...)
and
```

new("ScalarIndependenceTestConfint", ...).

Objects can be created by calls of the form

Slots

```
For objects of classes "IndependenceTest", "MaxTypeIndependenceTest", "QuadTypeIndependenceTest", "ScalarIndependenceTest" or "ScalarIndependenceTestConfint":

distribution: Object of class "PValue". The reference distribution.

statistic: Object of class "IndependenceTestStatistic". The test statistic, the linear statistic, and the transformed and original data structures.

estimates: Object of class "list". The estimated parameters.

method: Object of class "character". The test method.

call: Object of class "call". The matched call.

Additionally, for objects of classes "ScalarIndependenceTest" or "ScalarIndependenceTestConfint":

parameter: Object of class "character". The tested parameter.

nullvalue: Object of class "numeric". The hypothesized value of the null hypothesis.

Additionally, for objects of class "ScalarIndependenceTestConfint":

confint: Object of class "function". The confidence interval function.

conf.level: Object of class "numeric". The confidence level.
```

Extends

For objects of classes "MaxTypeIndependenceTest", "QuadTypeIndependenceTest" or "ScalarIndependenceTest": Class "IndependenceTest", directly.

For objects of class "ScalarIndependenceTestConfint":

Class "ScalarIndependenceTest", directly.

Class "IndependenceTest", by class "ScalarIndependenceTest", distance 2.

Known Subclasses

For objects of class "IndependenceTest":

Class "MaxTypeIndependenceTest", directly.

Class "QuadTypeIndependenceTest", directly.

Class "ScalarIndependenceTest", directly.

Class "ScalarIndependenceTestConfint", by class "ScalarIndependenceTest", distance 2.

For objects of class "ScalarIndependenceTest":

Class "ScalarIndependenceTestConfint", directly.

Methods

```
covariance signature(object = "IndependenceTest"): See the documentation for covariance
for details.
```

dperm signature(object = "IndependenceTest"): See the documentation for dperm for details.

expectation signature(object = "IndependenceTest"): See the documentation for expectation
for details.

midpvalue signature(object = "IndependenceTest"): See the documentation for midpvalue
 for details.

pperm signature(object = "IndependenceTest"): See the documentation for pperm for details.

pvalue signature(object = "IndependenceTest"): See the documentation for pvalue for details.

pvalue_interval signature(object = "IndependenceTest"): See the documentation for pvalue_interval
for details.

qperm signature(object = "IndependenceTest"): See the documentation for qperm for details.

rperm signature(object = "IndependenceTest"): See the documentation for rperm for details.

show signature(object = "IndependenceTest"): See the documentation for show (in package
 methods) for details.

show signature(object = "MaxTypeIndependenceTest"): See the documentation for show
 (in package methods) for details.

show signature(object = "QuadTypeIndependenceTest"): See the documentation for show (in package **methods**) for details.

```
show signature(object = "ScalarIndependenceTest"): See the documentation for show (in package methods) for details.
```

show signature(object = "ScalarIndependenceTestConfint"): See the documentation for show (in package methods) for details.

statistic signature(object = "IndependenceTest"): See the documentation for statistic
for details.

support signature(object = "IndependenceTest"): See the documentation for support for
 details.

variance signature(object = "IndependenceTest"): See the documentation for variance
for details.

IndependenceTestProblem-class

 ${\it Class}$ "IndependenceTestProblem"

Description

Objects of class "IndependenceTestProblem" represent the transformed and original data structures corresponding to an independence problem.

Objects from the Class

Objects can be created by calls of the form

```
new("IndependenceTestProblem", object, xtrafo = trafo, ytrafo = trafo, ...)
```

where object is an object of class "IndependenceProblem", xtrafo is the regression function $g(\mathbf{X})$ and ytrafo is the influence function $h(\mathbf{Y})$.

Slots

```
xtrans: Object of class "matrix". The transformed x. ytrans: Object of class "matrix". The transformed y. xtrafo: Object of class "function". The regression function for x. ytrafo: Object of class "function". The influence function for y. x: Object of class "data.frame". The variables x. y: Object of class "data.frame". The variables y. block: Object of class "factor". The block structure. weights: Object of class "numeric". The case weights.
```

Extends

Class "IndependenceProblem", directly.

Known Subclasses

```
Class "IndependenceLinearStatistic", directly.

Class "IndependenceTestStatistic", by class "IndependenceLinearStatistic", distance 2.

Class "MaxTypeIndependenceTestStatistic", by class "IndependenceTestStatistic", distance 3.

Class "QuadTypeIndependenceTestStatistic", by class "IndependenceTestStatistic", distance 3.

Class "ScalarIndependenceTestStatistic", by class "IndependenceTestStatistic", distance 3.
```

Methods

initialize signature(.Object = "IndependenceTestProblem"): See the documentation for initialize (in package methods) for details.

IndependenceTestStatistic-class

Class "IndependenceTestStatistic" and its subclasses

Description

Objects of class "IndependenceTestStatistic" and its subclasses "MaxTypeIndependenceTestStatistic", "QuadTypeIndependenceTestStatistic" and "ScalarIndependenceTestStatistic" represent the test statistic, the linear statistic, and the transformed and original data structures corresponding to an independence problem.

Objects from the Class

Class "IndependenceTestStatistic" is a *virtual* class, so objects cannot be created from it directly.

Objects can be created by calls of the form

where object is an object of class "IndependenceLinearStatistic", alternative is a character specifying the direction of the alternative hypothesis and paired is a logical indicating that paired data have been transformed in such a way that the (unstandardized) linear statistic is the sum of the absolute values of the positive differences between the paired observations.

Slots

```
For objects of classes "IndependenceTestStatistic", "MaxTypeIndependenceTestStatistic",
"QuadTypeIndependenceTestStatistic" or "ScalarIndependenceTestStatistic":
teststatistic: Object of class "numeric". The test statistic.
standardizedlinearstatistic: Object of class "numeric". The standardized linear statistic.
linear statistic: Object of class "numeric". The linear statistic.
expectation: Object of class "numeric". The expectation of the linear statistic.
covariance: Object of class "VarCovar". The covariance or variance of the linear statistic.
xtrans: Object of class "matrix". The transformed x.
ytrans: Object of class "matrix". The transformed y.
xtrafo: Object of class "function". The regression function for x.
ytrafo: Object of class "function". The influence function for y.
x: Object of class "data.frame". The variables x.
y: Object of class "data.frame". The variables y.
block: Object of class "factor". The block structure.
weights: Object of class "numeric". The case weights.
Additionally, for objects of classes "MaxTypeIndependenceTest" or "ScalarIndependenceTest":
alternative: Object of class "character". The direction of the alternative hypothesis.
Additionally, for objects of class "QuadTypeIndependenceTest":
covarianceplus: Object of class "matrix". The Moore-Penrose inverse of the covariance matrix.
df: Object of class "numeric". The rank of the Moore-Penrose inverse of the covariance matrix.
Additionally, for objects of classes "QuadTypeIndependenceTest" or "ScalarIndependenceTest":
paired: Object of class "logical". The indicator for paired test statistics.
```

Extends

```
For objects of class "IndependenceTestStatistic":
Class "IndependenceLinearStatistic", directly.
Class "IndependenceTestProblem", by class "IndependenceLinearStatistic", distance 2.
Class "IndependenceProblem", by class "IndependenceLinearStatistic", distance 3.
For objects of classes "MaxTypeIndependenceTestStatistic", "QuadTypeIndependenceTestStatistic"
or "ScalarIndependenceTestStatistic":
Class "IndependenceTestStatistic", directly.
Class "IndependenceLinearStatistic", by class "IndependenceTestStatistic", distance 2.
Class "IndependenceTestProblem", by class "IndependenceTestStatistic", distance 3.
Class "IndependenceProblem", by class "IndependenceTestStatistic", distance 4.
```

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Known Subclasses

For objects of class "IndependenceTestStatistic": Class "MaxTypeIndependenceTestStatistic", directly. Class "QuadTypeIndependenceTestStatistic", directly. Class "ScalarIndependenceTestStatistic", directly.

Methods

- **ApproxNullDistribution** signature(object = "MaxTypeIndependenceTestStatistic"): See the documentation for ApproxNullDistribution for details.
- **ApproxNullDistribution** signature(object = "QuadTypeIndependenceTestStatistic"): See the documentation for ApproxNullDistribution for details.
- **ApproxNullDistribution** signature(object = "ScalarIndependenceTestStatistic"): See the documentation for ApproxNullDistribution for details.
- **AsymptNullDistribution** signature(object = "MaxTypeIndependenceTestStatistic"): See the documentation for AsymptNullDistribution for details.
- **AsymptNullDistribution** signature(object = "QuadTypeIndependenceTestStatistic"): See the documentation for AsymptNullDistribution for details.
- **AsymptNullDistribution** signature(object = "ScalarIndependenceTestStatistic"): See the documentation for AsymptNullDistribution for details.
- ExactNullDistribution signature(object = "QuadTypeIndependenceTestStatistic"): See
 the documentation for ExactNullDistribution for details.
- **ExactNullDistribution** signature(object = "ScalarIndependenceTestStatistic"): See the documentation for ExactNullDistribution for details.
- initialize signature(.Object = "MaxTypeIndependenceTestStatistic"): See the documentation for initialize (in package methods) for details.
- initialize signature(.Object = "QuadTypeIndependenceTestStatistic"): See the documentation for initialize (in package methods) for details.
- initialize signature(.Object = "ScalarIndependenceTestStatistic"): See the documentation for initialize (in package methods) for details.
- statistic signature(object = "IndependenceTestStatistic"): See the documentation for statistic for details.

jobsatisfaction

Income and Job Satisfaction

Description

Income and job satisfaction by gender.

Usage

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Format

```
A contingency table with 104 observations on 3 variables.
```

```
Income a factor with levels "<5000", "5000-15000", "15000-25000" and ">25000".

Job.Satisfaction a factor with levels "Very Dissatisfied", "A Little Satisfied", "Moderately Satisfied" and "Very Satisfied".
```

Gender a factor with levels "Female" and "Male".

Details

The data are given in Agresti (2002, p. 288, Tab. 7.8).

Source

Agresti, A. (2002). *Categorical Data Analysis*, Second Edition. Hoboken, New Jersey: John Wiley & Sons.

Examples

```
## Approximative (Monte Carlo) linear-by-linear association test
lbl_test(jobsatisfaction, distribution = approximate(B = 10000))
```

LocationTests

Two- and K-Sample Location Tests

Description

Testing the equality of the distributions of a numeric response variable in two or more independent groups against shift alternatives.

Usage

```
## S3 method for class 'formula'
oneway_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'IndependenceProblem'
oneway_test(object, ...)

## S3 method for class 'formula'
wilcox_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'IndependenceProblem'
wilcox_test(object, conf.int = FALSE, conf.level = 0.95, ...)

## S3 method for class 'formula'
kruskal_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'IndependenceProblem'
kruskal_test(object, ...)
```

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Arguments

formula	a formula of the form $y \sim x \mid block$ where y is a numeric variable, x is a factor and block is an optional factor for stratification.
data	an optional data frame containing the variables in the model formula.
subset	an optional vector specifying a subset of observations to be used. Defaults to \ensuremath{NULL} .
weights	an optional formula of the form ~ w defining integer valued case weights for each observation. Defaults to NULL, implying equal weight for all observations.
object	an object inheriting from class "IndependenceProblem".
conf.int	a logical indicating whether a confidence interval for the difference in location should be computed. Defaults to FALSE.
conf.level	a numeric, confidence level of the interval. Defaults to 0.95.
ties.method	a character, the method used to handle ties: the score generating function either uses mid-ranks ("mid-ranks", default) or averages the scores of randomly broken ties ("average-scores").
mid.score	a character, the score assigned to observations exactly equal to the median: either 0 ("0", default), 0.5 ("0.5") or 1 ("1"); see 'Details'.
	further arguments to be passed to independence_test.

Details

oneway_test, wilcox_test, kruskal_test, normal_test, median_test and savage_test provide the Fisher-Pitman permutation test, the Wilcoxon-Mann-Whitney test, the Kruskal-Wallis test, the van der Waerden test, the Brown-Mood median test and the Savage test. A general description of these methods is given by Hollander and Wolfe (1999). For the adjustment of scores for tied values see Hájek, Šidák and Sen (1999, pp. 133–135).

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The null hypothesis of equality, or conditional equality given block, of the distribution of y in the groups defined by x is tested against shift alternatives. In the two-sample case, the two-sided null hypothesis is $H_0: \mu = 0$, where $\mu = Y_1 - Y_2$ and Y_s is the median of the responses in the sth sample. In case alternative = "less", the null hypothesis is $H_0: \mu \geq 0$. When alternative = "greater", the null hypothesis is $H_0: \mu \leq 0$. Confidence intervals for the difference in location are available (except for oneway_test) and computed according to Bauer (1972).

If x is an ordered factor, the default scores, 1:nlevels(x), can be altered using the scores argument (see independence_test); this argument can also be used to coerce nominal factors to class "ordered". In this case, a linear-by-linear association test is computed and the direction of the alternative hypothesis can be specified using the alternative argument.

The Brown-Mood median test offers a choice of mid-score, i.e., the score assigned to observations exactly equal to the median. In the two-sample case, mid-score = "0" implies that the linear test statistic is simply the number of subjects in the second sample with observations greater than the median of the pooled sample. Similarly, the linear test statistic for the last alternative, mid-score = "1", is the number of subjects in the second sample with observations greater than or equal to the median of the pooled sample. If mid-score = "0.5" is selected, the linear test statistic is the mean of the test statistics corresponding to the first and last alternatives and has a symmetric distribution, or at least approximately so, under the null hypothesis (see Hájek, Šidák and Sen, 1999, pp. 97–98).

The conditional null distribution of the test statistic is used to obtain *p*-values and an asymptotic approximation of the exact distribution is used by default (distribution = "asymptotic"). Alternatively, the distribution can be approximated via Monte Carlo resampling or computed exactly for univariate two-sample problems by setting distribution to "approximate" or "exact" respectively. See asymptotic, approximate and exact for details.

Value

An object inheriting from class "IndependenceTest". Confidence intervals can be extracted by confint.

Note

Starting with version 1.1-0, oneway_test no longer allows the test statistic to be specified; a quadratic form is now used in the K-sample case. Please use <code>independence_test</code> if more control is desired.

References

Bauer, D. F. (1972). Constructing confidence sets using rank statistics. *Journal of the American Statistical Association* **67**(339), 687–690.

Hájek, J., Šidák, Z. and Sen, P. K. (1999). *Theory of Rank Tests*, Second Edition. San Diego: Academic Press.

Hollander, M. and Wolfe, D. A. (1999). *Nonparametric Statistical Methods*, Second Edition. New York: John Wiley & Sons.

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```
## Tritiated Water Diffusion Across Human Chorioamnion
## Hollander and Wolfe (1999, p. 110, Tab. 4.1)
diffusion <- data.frame(</pre>
    pd = c(0.80, 0.83, 1.89, 1.04, 1.45, 1.38, 1.91, 1.64, 0.73, 1.46,
           1.15, 0.88, 0.90, 0.74, 1.21),
    age = factor(rep(c("At term", "12-26 Weeks"), c(10, 5)))
)
## Exact Wilcoxon-Mann-Whitney test
## Hollander and Wolfe (1999, p. 111)
## (At term - 12-26 Weeks)
(wt <- wilcox_test(pd ~ age, data = diffusion,</pre>
                   distribution = "exact", conf.int = TRUE))
## Extract observed Wilcoxon statistic
## Note: this is the sum of the ranks for age = "12-26 Weeks"
statistic(wt, type = "linear")
## Expectation, variance, two-sided pvalue and confidence interval
expectation(wt)
covariance(wt)
pvalue(wt)
confint(wt)
## For two samples, the Kruskal-Wallis test is equivalent to the W-M-W test
kruskal_test(pd ~ age, data = diffusion,
             distribution = "exact")
## Asymptotic Fisher-Pitman test
oneway_test(pd ~ age, data = diffusion)
## Approximative (Monte Carlo) Fisher-Pitman test
pvalue(oneway_test(pd ~ age, data = diffusion,
                   distribution = approximate(B = 10000)))
## Exact Fisher-Pitman test
pvalue(ot <- oneway_test(pd ~ age, data = diffusion,</pre>
                         distribution = "exact"))
## Plot density and distribution of the standardized test statistic
op <- par(no.readonly = TRUE) # save current settings
layout(matrix(1:2, nrow = 2))
s <- support(ot)</pre>
d <- dperm(ot, s)</pre>
p <- pperm(ot, s)</pre>
plot(s, d, type = "S", xlab = "Test Statistic", ylab = "Density")
plot(s, p, type = "S", xlab = "Test Statistic", ylab = "Cum. Probability")
par(op) # reset
```

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```
## Example data
ex <- data.frame(</pre>
   y = c(3, 4, 8, 9, 1, 2, 5, 6, 7),
    x = factor(rep(c("no", "yes"), c(4, 5)))
)
## Boxplots
boxplot(y \sim x, data = ex)
## Exact Brown-Mood median test with different mid-scores
(mt1 <- median_test(y ~ x, data = ex, distribution = "exact"))</pre>
(mt2 \leftarrow median_test(y \sim x, data = ex, distribution = "exact",
                    mid.score = "0.5"))
(mt3 <- median_test(y ~ x, data = ex, distribution = "exact",</pre>
                    mid.score = "1")) # sign change!
## Plot density and distribution of the standardized test statistics
op <- par(no.readonly = TRUE) # save current settings
layout(matrix(1:3, nrow = 3))
s1 <- support(mt1); d1 <- dperm(mt1, s1)</pre>
plot(s1, d1, type = "h", main = "Mid-score: 0",
     xlab = "Test Statistic", ylab = "Density")
s2 <- support(mt2); d2 <- dperm(mt2, s2)
plot(s2, d2, type = "h", main = "Mid-score: 0.5",
     xlab = "Test Statistic", ylab = "Density")
s3 <- support(mt3); d3 <- dperm(mt3, s3)</pre>
plot(s3, d3, type = "h", main = "Mid-score: 1",
     xlab = "Test Statistic", ylab = "Density")
par(op) # reset
## Length of YOY Gizzard Shad
## Hollander and Wolfe (1999, p. 200, Tab. 6.3)
yoy <- data.frame(</pre>
    length = c(46, 28, 46, 37, 32, 41, 42, 45, 38, 44,
               42, 60, 32, 42, 45, 58, 27, 51, 42, 52,
               38, 33, 26, 25, 28, 28, 26, 27, 27, 27,
               31, 30, 27, 29, 30, 25, 25, 24, 27, 30),
    site = gl(4, 10, labels = as.roman(1:4))
)
## Approximative (Monte Carlo) Kruskal-Wallis test
kruskal_test(length ~ site, data = yoy,
             distribution = approximate(B = 10000))
## Approximative (Monte Carlo) Nemenyi-Damico-Wolfe-Dunn test (joint ranking)
## Hollander and Wolfe (1999, p. 244)
## (where Steel-Dwass results are given)
it <- independence_test(length ~ site, data = yoy,</pre>
                         distribution = approximate(B = 50000),
                         ytrafo = function(data)
                             trafo(data, numeric_trafo = rank_trafo),
                         xtrafo = mcp_trafo(site = "Tukey"))
```

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```
## Global p-value
pvalue(it)

## Sites (I = II) != (III = IV) at alpha = 0.01 (p. 244)
pvalue(it, method = "single-step") # subset pivotality is violated
```

malformations

Maternal Drinking and Congenital Sex Organ Malformation

Description

A subset of data from a study on the relationship between maternal alcohol consumption and congenital malformations.

Usage

malformations

Format

A data frame with 32574 observations on 2 variables.

consumption alcohol consumption, an ordered factor with levels "0", "<1", "1-2", "3-5" and ">=6".

malformation congenital sex organ malformation, a factor with levels "Present" and "Absent".

Details

Data from a prospective study undertaken to determine whether moderate or light drinking during the first trimester of pregnancy increases the risk for congenital malformations (Mills and Graubard, 1987). The subset given here concerns only sex organ malformation (Mills and Graubard, 1987, Tab. 4).

Graubard and Korn (1987) used this data set to illustrate that different choices of scores for ordinal variables can lead to conflicting conclusions.

Source

Mills, J. L. and Graubard, B. I. (1987). Is moderate drinking during pregnancy associated with an increased risk for malformations? *Pediatrics* **80**(3), 309–314.

References

Graubard, B. I. and Korn, E. L. (1987). Choice of column scores for testing independence in ordered $2 \times K$ contingency tables. *Biometrics* **43**(2), 471–476.

Examples

```
## Graubard and Korn (1987, Tab. 3)
## One-sided approximative (Monte Carlo) Cochran-Armitage test
## Note: midpoint scores (p < 0.05)
midpoints <-c(0, 0.5, 1.5, 4.0, 7.0)
chisq_test(malformation ~ consumption, data = malformations,
           distribution = approximate(B = 1000), alternative = "greater",
           scores = list(consumption = midpoints))
## One-sided approximative (Monte Carlo) Cochran-Armitage test
## Note: midrank scores (p > 0.05)
midranks <- c(8557.5, 24375.5, 32013.0, 32473.0, 32555.5)
chisq_test(malformation ~ consumption, data = malformations,
           distribution = approximate(B = 1000), alternative = "greater",
           scores = list(consumption = midranks))
## One-sided approximative (Monte Carlo) Cochran-Armitage test
## Note: equally spaced scores (p > 0.05)
chisq_test(malformation ~ consumption, data = malformations,
           distribution = approximate(B = 1000), alternative = "greater")
```

MarginalHomogeneityTests

Marginal Homogeneity Tests

Description

Testing the marginal homogeneity of a repeated measurements factor in a complete block design.

Usage

```
## S3 method for class 'formula'
mh_test(formula, data, subset = NULL, ...)
## S3 method for class 'table'
mh_test(object, ...)
## S3 method for class 'SymmetryProblem'
mh_test(object, ...)
```

Arguments

formula	a formula of the form $y \sim x \mid block$ where y and x are factors and block is an optional factor (which is generated automatically if omitted).
data	an optional data frame containing the variables in the model formula.
subset	an optional vector specifying a subset of observations to be used. Defaults to \ensuremath{NULL} .
object	an object inheriting from classes "table" (with identical dimnames components) or "SymmetryProblem".
	further arguments to be passed to symmetry_test.

Details

mh_test provides the McNemar test, the Cochran Q test, the Stuart(-Maxwell) test and the Madansky test of interchangeability. A general description of these methods is given by Agresti (2002).

The null hypothesis of marginal homogeneity is tested. The response variable and the measurement conditions are given by y and x, respectively, and block is a factor where each level corresponds to exactly one subject with repeated measurements.

This procedure is known as the McNemar test (McNemar, 1947) when both y and x are binary factors, as the Cochran Q test (Cochran, 1950) when y is a binary factor and x is a factor with an arbitrary number of levels, as the Stuart(-Maxwell) test (Stuart, 1955; Maxwell, 1970) when y is a factor with an arbitrary number of levels and x is a binary factor, and as the Madansky test of interchangeability (Madansky, 1963), which implies marginal homogeneity, when both y and x are factors with an arbitrary number of levels.

If y and/or x are ordered factors, the default scores, 1:nlevels(y) and 1:nlevels(x) respectively, can be altered using the scores argument (see symmetry_test); this argument can also be used to coerce nominal factors to class "ordered". If both y and x are ordered factors, a linear-by-linear association test is computed and the direction of the alternative hypothesis can be specified using the alternative argument. This extension was given by Birch (1965) who also discussed the situation when either the response or the measurement condition is an ordered factor; see also White, Landis and Cooper (1982).

The conditional null distribution of the test statistic is used to obtain p-values and an asymptotic approximation of the exact distribution is used by default (distribution = "asymptotic"). Alternatively, the distribution can be approximated via Monte Carlo resampling or computed exactly for univariate two-sample problems by setting distribution to "approximate" or "exact" respectively. See asymptotic, approximate and exact for details.

Value

An object inheriting from class "IndependenceTest".

Note

This function is currently computationally inefficient for data with a large number of pairs or sets.

References

Agresti, A. (2002). *Categorical Data Analysis*, Second Edition. Hoboken, New Jersey: John Wiley & Sons.

Birch, M. W. (1965). The detection of partial association, II: The general case. *Journal of the Royal Statistical Society* B **27**(1), 111–124.

Cochran, W. G. (1950). The comparison of percentages in matched samples. *Biometrika* 37(3/4), 256–266.

Madansky, A. (1963). Tests of homogeneity for correlated samples. *Journal of the American Statistical Association* **58**(301), 97–119.

Maxwell, A. E. (1970). Comparing the classification of subjects by two independent judges. *British Journal of Psychiatry* **116**(535), 651–655.

McNemar, Q. (1947). Note on the sampling error of the difference between correlated proportions or percentages. *Psychometrika* **12**(2), 153–157.

Stuart, A. (1955). A test for homogeneity of the marginal distributions in a two-way classification. *Biometrika* **42**(3/4), 412–416.

White, A. A., Landis, J. R. and Cooper, M. M. (1982). A note on the equivalence of several marginal homogeneity test criteria for categorical data. *International Statistical Review* **50**(1), 27–34.

```
## Performance of prime minister
## Agresti (2002, p. 409)
performance <- matrix(</pre>
    c(794, 150,
       86, 570),
    nrow = 2, byrow = TRUE,
    dimnames = list(
         "First" = c("Approve", "Disprove"),
        "Second" = c("Approve", "Disprove")
)
performance <- as.table(performance)</pre>
diag(performance) <- 0 # speed-up: only off-diagonal elements contribute
## Asymptotic McNemar Test
mh_test(performance)
## Exact McNemar Test
mh_test(performance, distribution = "exact")
## Effectiveness of different media for the growth of diphtheria
## Cochran (1950, Tab. 2)
cases <- c(4, 2, 3, 1, 59)
n <- sum(cases)</pre>
cochran <- data.frame(</pre>
    diphtheria = factor(
        unlist(rep(list(c(1, 1, 1, 1),
                         c(1, 1, 0, 1),
                         c(0, 1, 1, 1),
                         c(0, 1, 0, 1),
                         c(0, 0, 0, 0)),
                    cases))
    media = factor(rep(LETTERS[1:4], n)),
    case = factor(rep(seq_len(n), each = 4))
)
## Asymptotic Cochran Q test (Cochran, 1950, p. 260)
mh_test(diphtheria ~ media | case, data = cochran) # Q = 8.05
## Approximative Cochran Q test
mt <- mh_test(diphtheria ~ media | case, data = cochran,</pre>
```

```
distribution = approximate(B = 10000))
pvalue(mt)
                 # standard p-value
midpvalue(mt)
                     # mid-p-value
pvalue_interval(mt) # p-value interval
## Opinions on Pre- and Extramarital Sex
## Agresti (2002, p. 421)
opinions <- c("Always wrong", "Almost always wrong",
               "Wrong only sometimes", "Not wrong at all")
PreExSex <- matrix(</pre>
    c(144, 33, 84, 126,
        2, 4, 14, 29,
        0, 2, 6, 25,
        0, 0, 1, 5),
    nrow = 4,
    dimnames = list(
           "Premarital Sex" = opinions,
         "Extramarital Sex" = opinions
    )
)
PreExSex <- as.table(PreExSex)</pre>
## Asymptotic Stuart test
mh_test(PreExSex)
## Asymptotic Stuart-Birch test
## Note: response as ordinal
mh_test(PreExSex, scores = list(response = 1:length(opinions)))
## Vote intention
## Madansky (1963, pp. 107-108)
vote <- array(</pre>
    c(120, 1, 8, 2, 2, 1, 2, 1, 7,
        6, 2, 1, 1, 103, 5, 1, 4, 8,
       20, 3, 31, 1, 6, 30, 2, 1, 81),
    dim = c(3, 3, 3),
    dimnames = list(
         "July" = c("Republican", "Democratic", "Uncertain"),
"August" = c("Republican", "Democratic", "Uncertain"),
"June" = c("Republican", "Democratic", "Uncertain")
    )
)
vote <- as.table(vote)</pre>
## Asymptotic Madansky test (Q = 70.77)
mh_test(vote)
## Cross-over study
## http://www.nesug.org/proceedings/nesug00/st/st9005.pdf
dysmenorrhea <- array(</pre>
```

```
c(6, 2, 1, 3, 1, 0, 1, 2, 1,
     4, 3, 0, 13, 3, 0, 8, 1, 1,
     5, 2, 2, 10, 1, 0, 14, 2, 0),
   dim = c(3, 3, 3),
   dimnames = list(
         "Placebo" = c("None", "Moderate", "Complete"),
         "Low dose" = c("None", "Moderate", "Complete"),
        "High dose" = c("None", "Moderate", "Complete")
   )
)
dysmenorrhea <- as.table(dysmenorrhea)</pre>
## Asymptotic Madansky-Birch test (Q = 53.76)
## Note: response as ordinal
mh_test(dysmenorrhea, scores = list(response = 1:3))
## Asymptotic Madansky-Birch test (Q = 47.29)
## Note: response and measurement conditions as ordinal
mh_test(dysmenorrhea, scores = list(response = 1:3,
                                    conditions = 1:3))
```

 ${\tt MaximallySelectedStatisticsTests}$

Generalized Maximally Selected Statistics

Description

Testing the independence of two sets of variables measured on arbitrary scales against cutpoint alternatives.

Usage

Arguments

formula a formula of the form y1 + ... + yq ~ x1 + ... + xp | block where y1, ..., yq and x1, ..., xp are measured on arbitrary scales (nominal, ordinal or continuous with or without censoring) and block is an optional factor for stratification.

data an optional data frame containing the variables in the model formula.

subset

	NULL.
weights	an optional formula of the form \sim w defining integer valued case weights for each observation. Defaults to NULL, implying equal weight for all observations.
object	an object inheriting from classes "table" or "IndependenceProblem".
teststat	a character, the type of test statistic to be applied: either a maximum statistic ("maximum", default) or a quadratic form ("quadratic").
distribution	a character, the conditional null distribution of the test statistic can be approximated by its asymptotic distribution ("asymptotic", default) or via Monte Carlo resampling ("approximate"). Alternatively, the functions asymptotic or approximate can be used. Computation of the null distribution can be sup-

pressed by specifying "none".

an optional vector specifying a subset of observations to be used. Defaults to

minprob a numeric, a fraction between 0 and 0.5 specifying that cutpoints only greater

than the minprob · 100 % quantile of x1, ..., xp are considered. Defaults to

0.1.

a numeric, a fraction between 0.5 and 1 specifying that cutpoints only smaller maxprob

than the maxprob · 100 % quantile of x1, ..., xp are considered. Defaults to

1 - minprob.

further arguments to be passed to independence_test.

Details

maxstat_test provides generalized maximally selected statistics. The family of maximally selected statistics encompasses a large collection of procedures used for the estimation of simple cutpoint models including, but not limited to, maximally selected χ^2 statistics, maximally selected Cochran-Armitage statistics, maximally selected rank statistics and maximally selected statistics for multiple covariates. A general description of these methods is given by Hothorn and Zeileis (2008).

The null hypothesis of independence, or conditional independence given block, between y1, ..., yq and x1,..., xp is tested against cutpoint alternatives. All possible partitions into two groups are evaluated for each unordered covariate x1, ..., xp, whereas only order-preserving binary partitions are evaluated for ordered or numeric covariates. The cutpoint is then a set of levels defining one of the two groups.

If both response and covariate is univariable, say y1 and x1, this procedure is known as maximally selected χ^2 statistics (Miller and Siegmund, 1982) when y1 is a binary factor and x1 is a numeric variable, and as maximally selected rank statistics when y1 is a rank transformed numeric variable and x1 is a numeric variable (Lausen and Schumacher, 1992). Lausen et al. (2004) introduced maximally selected statistics for a univariable numeric response and multiple numeric covariates x1, ..., xp.

If, say, y1 and/or x1 are ordered factors, the default scores, 1:nlevels(y1) and 1:nlevels(x1) respectively, can be altered using the scores argument (see independence_test); this argument can also be used to coerce nominal factors to class "ordered". If both, say, y1 and x1 are ordered factors, a linear-by-linear association test is computed and the direction of the alternative hypothesis can be specified using the alternative argument. The particular extension to the case of a univariable binary factor response and a univariable ordered covariate was given by Betensky and Rabinowitz (1999) and is known as maximally selected Cochran-Armitage statistics.

The conditional null distribution of the test statistic is used to obtain *p*-values and an asymptotic approximation of the exact distribution is used by default (distribution = "asymptotic"). Alternatively, the distribution can be approximated via Monte Carlo resampling by setting distribution to "approximate". See asymptotic and approximate for details.

Value

An object inheriting from class "IndependenceTest".

Note

Starting with **coin** version 1.1-0, maximum statistics and quadratic forms can no longer be specified using teststat = "maxtype" and teststat = "quadtype" respectively (as was used in versions prior to 0.4-5).

References

Betensky, R. A. and Rabinowitz, D. (1999). Maximally selected χ^2 statistics for $k \times 2$ tables. *Biometrics* **55**(1), 317–320.

Hothorn, T. and Lausen, B. (2003). On the exact distribution of maximally selected rank statistics. *Computational Statistics & Data Analysis* **43**(2), 121–137.

Hothorn, T. and Zeileis, A. (2008). Generalized maximally selected statistics. *Biometrics* **64**(4), 1263–1269.

Lausen, B., Hothorn, T., Bretz, F. and Schumacher, M. (2004). Optimally selected prognostic factors. *Biometrical Journal* **46**(3), 364–374.

Lausen, B. and Schumacher, M. (1992). Maximally selected rank statistics. *Biometrics* **48**(1), 73–85.

Miller, R. and Siegmund, D. (1982). Maximally selected chi square statistics. *Biometrics* **38**(4), 1011–1016.

Müller, J. and Hothorn, T. (2004). Maximally selected two-sample statistics as a new tool for the identification and assessment of habitat factors with an application to breeding bird communities in oak forests. *European Journal of Forest Research* **123**(3), 219–228.

```
## Tree pipit data (Mueller and Hothorn, 2004)
## Asymptotic maximally selected statistics
maxstat_test(counts ~ coverstorey, data = treepipit)

## Asymptotic maximally selected statistics
## Note: all covariates simultaneously
mt <- maxstat_test(counts ~ ., data = treepipit)
mt@estimates$estimate

## Malignant arrythmias data (Hothorn and Lausen, 2003, Sec. 7.2)
## Asymptotic maximally selected statistics
maxstat_test(Surv(time, event) ~ EF, data = hohnloser,</pre>
```

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```
ytrafo = function(data)
                 trafo(data, surv_trafo = function(y)
                     logrank_trafo(y, ties.method = "Hothorn-Lausen")))
## Breast cancer data (Hothorn and Lausen, 2003, Sec. 7.3)
## Asymptotic maximally selected statistics
data("sphase", package = "TH.data")
maxstat_test(Surv(RFS, event) ~ SPF, data = sphase,
             ytrafo = function(data)
                 trafo(data, surv_trafo = function(y)
                     logrank_trafo(y, ties.method = "Hothorn-Lausen")))
## Job satisfaction data (Agresti, 2002, p. 288, Tab. 7.8)
## Asymptotic maximally selected statistics
maxstat_test(jobsatisfaction)
## Asymptotic maximally selected statistics
## Note: 'Job.Satisfaction' and 'Income' as ordinal
maxstat_test(jobsatisfaction,
             scores = list("Job.Satisfaction" = 1:4,
                           "Income" = 1:4))
```

mercuryfish

Chromosomal Effects of Mercury-Contaminated Fish Consumption

Description

The mercury level in blood, the proportion of cells with abnormalities, and the proportion of cells with chromosome aberrations in consumers of mercury-contaminated fish and a control group.

Usage

mercuryfish

Format

A data frame with 39 observations on 4 variables.

```
group a factor with levels "control" and "exposed".

mercury mercury level in blood.

abnormal the proportion of cells with structural abnormalities.
```

ccells the proportion of C_u cells, i.e., cells with asymmetrical or incomplete-symmetrical chromosome aberrations.

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Details

Control subjects ("control") and subjects who are contaminated fish for more than three years ("exposed") are under study.

Rosenbaum (1994) proposed a coherence criterion defining a partial ordering, i.e., an observation is smaller than another when all responses are smaller, and a score reflecting the "ranking" is attached to each observation. The corresponding partially ordered set (POSET) test can be used to test if the distribution of the scores differ between the groups. Alternatively, a multivariate test can be applied.

Source

Skerfving, S., Hansson, K., Mangs, C., Lindsten, J. and Ryman, N. (1974). Methylmercury-induced chromosome damage in men. *Environmental Research* 7(1), 83–98.

References

Hothorn, T., Hornik, K., van de Wiel, M. A. and Zeileis, A. (2006). A Lego system for conditional inference. *The American Statistician* **60**(3), 257–263.

Rosenbaum, P. R. (1994). Coherence in observational studies. *Biometrics* **50**(2), 368–374.

```
## Coherence criterion
coherence <- function(data) {</pre>
   x <- as.matrix(data)</pre>
   matrix(apply(x, 1, function(y)
        sum(colSums(t(x) < y) == ncol(x)) -
            sum(colSums(t(x) > y) == ncol(x))), ncol = 1)
}
## Asymptotic POSET test
poset <- independence_test(mercury + abnormal + ccells ~ group,</pre>
                            data = mercuryfish, ytrafo = coherence)
## Linear statistic (T in the notation of Rosenbaum, 1994)
statistic(poset, type = "linear")
## Expectation
expectation(poset)
## Variance
## Note: typo in Rosenbaum (1994, p. 371, Sec. 2, last paragraph)
variance(poset)
## Standardized statistic
statistic(poset)
## P-value
pvalue(poset)
## Exact POSET test
```

50 neuropathy

neuropathy

Acute Painful Diabetic Neuropathy

Description

The logarithm of the ratio of pain scores measured at baseline and after four weeks in a control group and a treatment group.

Usage

neuropathy

Format

A data frame with 58 observations on 2 variables.

```
pain pain scores: ln(baseline / final).
group a factor with levels "control" and "treat".
```

Details

Data from Conover and Salsburg (1988, Tab. 1).

Source

Conover, W. J. and Salsburg, D. S. (1988). Locally most powerful tests for detecting treatment effects when only a subset of patients can be expected to "respond" to treatment. *Biometrics* **44**(1), 189–196.

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Examples

```
## Conover and Salsburg (1988, Tab. 2)
## One-sided approximative Fisher-Pitman test
oneway_test(pain ~ group, data = neuropathy,
            alternative = "less",
            distribution = approximate(B = 10000))
## One-sided approximative Wilcoxon-Mann-Whitney test
wilcox_test(pain ~ group, data = neuropathy,
            alternative = "less",
            distribution = approximate(B = 10000))
## One-sided approximative Conover-Salsburg test
oneway_test(pain ~ group, data = neuropathy,
            alternative = "less",
            distribution = approximate(B = 10000),
            ytrafo = function(data)
                trafo(data, numeric_trafo = consal_trafo))
## One-sided approximative maximum test for a range of 'a' values
it <- independence_test(pain ~ group, data = neuropathy,</pre>
                        alternative = "less",
                        distribution = approximate(B = 10000),
                        ytrafo = function(data)
                            trafo(data, numeric_trafo = function(y)
                                consal\_trafo(y, a = 2:7)))
pvalue(it, method = "single-step")
```

NullDistribution

Specification of the Reference Distribution

Description

Specification of the asymptotic, approximative (Monte Carlo) and exact reference distribution.

Usage

Arguments

```
maxpts an integer, the maximum number of function values. Defaults to 25000.

abseps a numeric, the absolute error tolerance. Defaults to 0.001.

releps a numeric, the relative error tolerance. Defaults to 0.
```

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В	a positive integer, the number of Monte Carlo replicates used for the computation of the approximative reference distribution. Defaults to 10000.
parallel	a character, the type of parallel operation: either "no" (default), "multicore" or "snow".
ncpus	an integer, the number of processes to be used in parallel operation. Defaults to 1.
cl	an object inheriting from class "cluster", specifying an optional parallel or snow cluster if parallel = "snow". Defaults to NULL.
algorithm	a character, the algorithm used for the computation of the exact reference distribution: either "auto" (default), "shift" or "split-up".
fact	an integer to multiply the response values with. Defaults to NULL.

Details

asymptotic, approximate and exact can be supplied to the distribution argument of, e.g., independence_test to provide control of the specification of the asymptotic, approximative (Monte Carlo) and exact reference distribution respectively.

The asymptotic reference distribution is computed using a randomised quasi-Monte Carlo method (Genz and Bretz, 2009) and is applicable to arbitrary covariance structures with dimensions up to 1000. See GenzBretz in package **mvtnorm** for details on maxpts, abseps and releps.

The approximative (Monte Carlo) reference distribution is obtained by a conditional Monte Carlo procedure, i.e., by computing the test statistic for B random samples from all admissible permutations of the response Y within each block (Hothorn *et al.*, 2008). By default, the distribution is computed using serial operation (parallel = "no"). The use of parallel operation is specified by setting parallel to either "multicore" (not available for MS Windows) or "snow". In the latter case, if cl = NULL (default) a cluster with ncpus processes is created on the local machine unless a default cluster has been registered (see setDefaultCluster in package parallel) in which case that gets used instead. Alternatively, the use of an optional parallel or snow cluster can be specified by cl. See 'Examples' and package parallel for details on parallel operation.

The exact reference distribution, currently available for univariate two-sample problems only, is computed using either the shift algorithm (Streitberg and Röhmel, 1986, 1987) or the split-up algorithm (van de Wiel, 2001). The shift algorithm handles blocks pertaining to, e.g., pre- and post-stratification, but can only be used with positive integer-valued scores $h(\mathbf{Y})$. The split-up algorithm can be used with non-integer scores, but does not handle blocks. By default, an automatic choice is made (algorithm = "auto") but the shift and split-up algorithms can be selected by setting algorithm to either "shift" or "split-up" respectively.

Note

Starting with coin version 1.1-0, the default for algorithm is "auto", having identical behaviour to "shift" in previous versions. In earlier versions of the package, algorithm = "shift" silently switched to the split-up algorithm if non-integer scores were detected, whereas the current version exits with a warning.

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References

Genz, A. and Bretz, F. (2009). *Computation of Multivariate Normal and t Probabilities*. Heidelberg: Springer-Verlag.

Hothorn, T., Hornik, K., van de Wiel, M. A. and Zeileis, A. (2008). Implementing a class of permutation tests: the coin package. *Journal of Statistical Software* **28**(8), 1–23. http://www.jstatsoft.org/v28/i08/

Streitberg, B. and Röhmel, J. (1986). Exact distributions for permutations and rank tests: an introduction to some recently published algorithms. *Statistical Software Newsletter* **12**(1), 10–17.

Streitberg, B. and Röhmel, J. (1987). Exakte verteilungen für rang- und randomisierungstests im allgemeinen c-stichprobenfall. *EDV in Medizin und Biologie* **18**(1), 12–19.

van de Wiel, M. A. (2001). The split-up algorithm: a fast symbolic method for computing p-values of distribution-free statistics. *Computational Statistics* **16**(4), 519–538.

```
## Approximative (Monte Carlo) Cochran-Mantel-Haenszel test
## Serial operation
set.seed(123)
cmh_test(disease ~ smoking | gender, data = alzheimer,
         distribution = approximate(B = 100000))
## Not run:
## Multicore with 8 processes (not for MS Windows)
set.seed(123, kind = "L'Ecuyer-CMRG")
cmh_test(disease ~ smoking | gender, data = alzheimer,
         distribution = approximate(B = 100000,
                                    parallel = "multicore", ncpus = 8))
## Automatic PSOCK cluster with 4 processes
set.seed(123, kind = "L'Ecuyer-CMRG")
cmh_test(disease ~ smoking | gender, data = alzheimer,
         distribution = approximate(B = 100000,
                                    parallel = "snow", ncpus = 4))
## Registered FORK cluster with 12 processes (not for MS Windows)
fork12 <- parallel::makeCluster(12, "FORK") # set-up cluster</pre>
parallel::setDefaultCluster(fork12) # register default cluster
set.seed(123, kind = "L'Ecuyer-CMRG")
cmh_test(disease ~ smoking | gender, data = alzheimer,
         distribution = approximate(B = 100000,
                                    parallel = "snow"))
parallel::stopCluster(fork12) # clean-up
## User-specified PSOCK cluster with 8 processes
psock8 <- parallel::makeCluster(8, "PSOCK") # set-up cluster</pre>
set.seed(123, kind = "L'Ecuyer-CMRG")
cmh_test(disease ~ smoking | gender, data = alzheimer,
         distribution = approximate(B = 100000,
                                    parallel = "snow", cl = psock8))
```

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```
parallel::stopCluster(psock8) # clean-up
## End(Not run)

NullDistribution-class

Class "NullDistribution" and its subclasses
```

Description

Objects of class "NullDistribution" and its subclasses "ApproxNullDistribution", "AsymptNullDistribution" and "ExactNullDistribution" represent the reference distribution.

Objects from the Class

Objects can be created by calls of the form

```
new("NullDistribution", ...),
new("ApproxNullDistribution", ...),
new("AsymptNullDistribution", ...)
and
new("ExactNullDistribution", ...).
```

Slots

For objects of classes "NullDistribution", "ApproxNullDistribution", "AsymptNullDistribution" or "ExactNullDistribution":

q: Object of class "function". The quantile function of the reference distribution.

d: Object of class "function". The density function of the reference distribution.

support: Object of class "function". The support of the reference distribution.

parameters: Object of class "list". Additional parameters.

pvalue: Object of class "function". The *p*-value function of the reference distribution.

midpvalue: Object of class "function". The mid-p-value function of the reference distribution.

pvalueinterval: Object of class "function". The *p*-value interval function of the reference distribution

p: Object of class "function". The distribution function of the reference distribution.

name: Object of class "character". The name of the reference distribution.

Additionally, for objects of classes "ApproxNullDistribution" or "AsymptNullDistribution":

seed: Object of class "integer". The random number generator state (i.e., the value of .Random. seed).

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Extends

```
For objects of class "NullDistribution": Class "PValue", directly.

For objects of classes "ApproxNullDistribution", "AsymptNullDistribution" or "ExactNullDistribution": Class "NullDistribution", directly.

Class "PValue", by class "NullDistribution", distance 2.
```

Known Subclasses

```
For objects of class "NullDistribution": Class "ApproxNullDistribution", directly. Class "AsymptNullDistribution", directly. Class "ExactNullDistribution", directly.
```

Methods

NullDistribution-methods

Computation of the Reference Distribution

Description

Methods for computation of the asymptotic, approximative (Monte Carlo) and exact reference distribution.

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Usage

```
## S4 method for signature 'MaxTypeIndependenceTestStatistic'
AsymptNullDistribution(object, ...)
## S4 method for signature 'QuadTypeIndependenceTestStatistic'
AsymptNullDistribution(object, ...)
## S4 method for signature 'ScalarIndependenceTestStatistic'
AsymptNullDistribution(object, ...)
## S4 method for signature 'MaxTypeIndependenceTestStatistic'
ApproxNullDistribution(object, B = 10000, ...)
## S4 method for signature 'QuadTypeIndependenceTestStatistic'
ApproxNullDistribution(object, B = 10000, ...)
## S4 method for signature 'ScalarIndependenceTestStatistic'
ApproxNullDistribution(object, B = 10000, ...)
## S4 method for signature 'QuadTypeIndependenceTestStatistic'
ExactNullDistribution(object, algorithm = c("auto", "shift", "split-up"), ...)
## S4 method for signature 'ScalarIndependenceTestStatistic'
ExactNullDistribution(object, algorithm = c("auto", "shift", "split-up"), ...)
```

Arguments

object	an object from which the asymptotic, approximative (Monte Carlo) or exact reference distribution can be computed.
В	a positive integer, the number of Monte Carlo replicates used for the computation of the approximative reference distribution. Defaults to 10000.
algorithm	a character, the algorithm used for the computation of the exact reference distribution: either "auto" (default), "shift" or "split-up".
	further arguments to be passed to or from methods.

Details

The methods AsymptNullDistribution, ApproxNullDistribution and ExactNullDistribution compute the asymptotic, approximative (Monte Carlo) and exact reference distribution respectively.

Value

An object of class "AsymptNullDistribution", "ApproxNullDistribution" or "ExactNullDistribution".

Description

Survival times of 35 women suffering from ovarian carcinoma at stadium II and IIA.

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Usage

ocarcinoma

Format

A data frame with 35 observations on 3 variables.

```
time time (days).
stadium a factor with levels "II" and "IIA".
event status indicator for time: FALSE for right-censored observations and TRUE otherwise.
```

Details

Data from Fleming *et al.* (1980) and Fleming, Green and Harrington (1984). Reanalysed in Schumacher and Schulgen (2002).

Source

Fleming, T. R., Green, S. J. and Harrington, D. P. (1984). Considerations of monitoring and evaluating treatment effects in clinical trials. *Controlled Clinical Trials* **5**(1), 55–66.

Fleming, T. R., O'Fallon, J. R., O'Brien, P. C. and Harrington, D. P. (1980). Modified Kolmogorov-Smirnov test procedures with applications to arbitrarily censored data. *Biometrics* **36**(4), 607–625.

References

Schumacher, M. and Schulgen, G. (2002). *Methodik Klinischer Studien: Methodische Grundlagen der Planung, Durchführung und Auswertung.* Heidelberg: Springer.

PermutationDistribution-methods

Computation of the Permutation Distribution

Description

Methods for computation of the density function, distribution function, quantile function, random numbers and support of the permutation distribution.

Usage

```
## S4 method for signature 'IndependenceTest'
dperm(object, x, ...)
## S4 method for signature 'NullDistribution'
dperm(object, x, ...)
## S4 method for signature 'IndependenceTest'
pperm(object, q, ...)
## S4 method for signature 'NullDistribution'
pperm(object, q, ...)
## S4 method for signature 'IndependenceTest'
qperm(object, p, ...)
## S4 method for signature 'NullDistribution'
qperm(object, p, ...)
## S4 method for signature 'IndependenceTest'
rperm(object, n, ...)
## S4 method for signature 'NullDistribution'
rperm(object, n, ...)
## S4 method for signature 'IndependenceTest'
support(object, ...)
## S4 method for signature 'NullDistribution'
support(object, ...)
```

Arguments

object	an object from which the density function, distribution function, quantile function, random numbers or support of the permutation distribution can be computed.
x, q	a numeric vector, the quantiles for which the density function or distribution function is computed.
p	a numeric vector, the probabilities for which the quantile function is computed.
n	a numeric vector, the number of observations. If $length(n) > 1$, the length is taken to be the number required.
	further arguments to be passed to methods.

Details

The methods dperm, pperm, qperm, rperm and support compute the density function, distribution function, quantile function, random deviates and support, respectively, of the permutation distribution.

Value

The density function, distribution function, quantile function, random deviates or support of the permutation distribution computed from object. A numeric vector.

Note

The density of asymptotic permutation distributions for maximum-type tests or exact permutation distributions obtained by the split-up algoritm is reported as NA. The quantile function of asymptotic permutation distributions for maximum-type tests cannot be computed for p less than 0.5, due to limitations in the **mvtnorm** package. The support of exact permutation distributions obtained by the split-up algorithm is reported as NA.

In versions of **coin** prior to 1.1-0, the support of asymptotic permutation distributions was given as an interval containing 99.999 % of the probability mass. It is now reported as NA.

```
## Two-sample problem
dta <- data.frame(</pre>
    y = rnorm(20),
    x = g1(2, 10)
)
## Exact Ansari-Bradley test
at <- ansari_test(y ~ x, data = dta, distribution = "exact")
## Support of the exact distribution of the Ansari-Bradley statistic
supp <- support(at)</pre>
## Density of the exact distribution of the Ansari-Bradley statistic
dens <- dperm(at, x = supp)
## Plotting the density
plot(supp, dens, type = "s")
## 95 % quantile
qperm(at, p = 0.95)
## One-sided p-value
pperm(at, q = statistic(at))
## Random number generation
rperm(at, n = 5)
```

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photocar

Multiple Dosing Photococarcinogenicity Experiment

Description

Survival time, time to first tumor, and total number of tumors in three groups of animals in a photo-cocarcinogenicity study.

Usage

photocar

Format

A data frame with 108 observations on 6 variables.

group a factor with levels "A", "B", and "C".

ntumor total number of tumors.

time survival time.

event status indicator for time: FALSE for right-censored observations and TRUE otherwise.

dmin time to first tumor.

tumor status indicator for dmin: FALSE when no tumor was observed and TRUE otherwise.

Details

The animals were exposed to different levels of ultraviolet radiation (UVR) exposure (group A: topical vehicle and 600 Robertson–Berger units of UVR, group B: no topical vehicle and 600 Robertson–Berger units of UVR and group C: no topical vehicle and 1200 Robertson–Berger units of UVR). The data are taken from Tables 1 to 3 in Molefe *et al.* (2005).

The main interest is testing the global null hypothesis of no treatment effect with respect to survival time, time to first tumor and number of tumors. (Molefe *et al.*, 2005, also analysed the detection time of tumors, but that data is not given here.) In case the global null hypothesis can be rejected, the deviations from the partial null hypotheses are of special interest.

Source

Molefe, D. F., Chen, J. J., Howard, P. C., Miller, B. J., Sambuco, C. P., Forbes, P. D. and Kodell, R. L. (2005). Tests for effects on tumor frequency and latency in multiple dosing photococarcinogenicity experiments. *Journal of Statistical Planning and Inference* **129**(1–2), 39–58.

References

Hothorn, T., Hornik, K., van de Wiel, M. A. and Zeileis, A. (2006). A Lego system for conditional inference. *The American Statistician* **60**(3), 257–263.

PValue-class 61

Examples

```
## Plotting data
op <- par(no.readonly = TRUE) # save current settings
layout(matrix(1:3, ncol = 3))
with(photocar, {
   plot(survfit(Surv(time, event) ~ group),
         lty = 1:3, xmax = 50, main = "Survival Time")
   legend("bottomleft", lty = 1:3, levels(group), bty = "n")
   plot(survfit(Surv(dmin, tumor) ~ group),
         lty = 1:3, xmax = 50, main = "Time to First Tumor")
   legend("bottomleft", lty = 1:3, levels(group), bty = "n")
    boxplot(ntumor ~ group, main = "Number of Tumors")
})
par(op) # reset
## Approximative multivariate (all three responses) test
it <- independence_test(Surv(time, event) + Surv(dmin, tumor) + ntumor ~ group,</pre>
                        data = photocar,
                        distribution = approximate(B = 10000))
## Global p-value
pvalue(it)
## Why was the global null hypothesis rejected?
statistic(it, type = "standardized")
pvalue(it, method = "single-step")
```

PValue-class

Class "PValue"

Description

Objects of class "PValue" represent the p-value, mid-p-value and p-value interval of the reference distribution.

Objects from the Class

Objects can be created by calls of the form

```
new("PValue", ...).
```

Slots

```
pvalue: Object of class "function". The p-value function of the reference distribution.midpvalue: Object of class "function". The mid-p-value function of the reference distribution.pvalueinterval: Object of class "function". The p-value interval function of the reference distribution.
```

p: Object of class "function". The distribution function of the reference distribution. name: Object of class "character". The name of the reference distribution.

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Methods

No methods defined with class "PValue" in the signature.

pvalue-methods Computation of the p-Value, Mid-p-Value and p-Value Interval

Description

Methods for computation of the p-value, mid-p-value and p-value interval.

Usage

```
## S4 method for signature 'IndependenceTest'
pvalue(object, ...)
## S4 method for signature 'MaxTypeIndependenceTest'
pvalue(object, method = c("global", "single-step", "step-down", "unadjusted"),
       distribution = c("joint", "marginal"),
       type = c("Bonferroni", "Sidak"), ...)
## S4 method for signature 'NullDistribution'
pvalue(object, q, ...)
## S4 method for signature 'IndependenceTest'
midpvalue(object, ...)
## S4 method for signature 'NullDistribution'
midpvalue(object, q, ...)
## S4 method for signature 'IndependenceTest'
pvalue_interval(object, ...)
## S4 method for signature 'NullDistribution'
pvalue_interval(object, q, ...)
```

Arguments

object	an object from which the p -value, mid- p -value or p -value interval can be computed.
method	a character, the method used for the p -value computation: either "global" (default), "single-step", "step-down" or "unadjusted".
distribution	a character, the distribution used for the computation of adjusted p -values: either "joint" (default) or "marginal".
type	a character, the type of p -value adjustment when the marginal distributions are used: either "Bonferroni" (default) or "Sidak".
q	a numeric, the quantile for which the p -value, mid- p -value or p -value interval is computed.
	further arguments (currently ignored).

pvalue-methods 63

Details

The methods pvalue, midpvalue and pvalue_interval compute the p-value, mid-p-value and p-value interval respectively.

For pvalue, the global p-value (method = "global") is returned by default and is given with an associated 99 % confidence interval when resampling is used to determine the null distribution (which for maximum statistics may be true even in the asymptotic case).

The familywise error rate (FWER) is always controlled under the global null hypothesis, i.e., in the *weak* sense, implying that the smallest adjusted *p*-value is valid without further assumptions. Control of the FWER under any partial configuration of the null hypotheses, i.e., in the *strong* sense, as is typically desired for multiple tests and comparisons, requires that the *subset pivotality* condition holds (Westfall and Young, 1993, pp. 42–43; Bretz, Hothorn and Westfall, 2011, pp. 136–137). In addition, for methods based on the joint distribution of the test statistics, failure of the *joint exchangeability* assumption (Westfall and Troendle, 2008; Bretz, Hothorn and Westfall, 2011, pp. 129–130) may cause excess Type I errors.

Assuming *subset pivotality*, single-step or *free* step-down adjusted *p*-values using max-*T* procedures are obtained by setting method to "single-step" or "step-down" respectively. In both cases, the distribution argument specifies whether the adjustment is based on the joint distribution ("joint") or the marginal distributions ("marginal") of the test statistics. For procedures based on the marginal distributions, Bonferroni- or Šidák-type adjustment can be specified through the type argument by setting it to "Bonferroni" or "Sidak" respectively.

The *p*-value adjustment procedures based on the joint distribution of the test statistics fully utilizes distributional characteristics, such as discreteness and dependence structure, whereas procedures using the marginal distributions only incorporate discreteness. Hence, the joint distribution-based procedures are typically more powerful. Details regarding the single-step and *free* step-down procedures based on the joint distribution can be found in Westfall and Young (1993); in particular, this implementation uses Equation 2.8 with Algorithm 2.5 and 2.8 respectively. Westfall and Wolfinger (1997) provide details of the marginal distributions-based single-step and *free* step-down procedures. The generalization of Westfall and Wolfinger (1997) to arbitrary test statistics, as implemented here, is given by Westfall and Troendle (2008).

Unadjusted p-values are obtained using method = "unadjusted".

For midpvalue, the global mid-p-value is given with an associated 99 % mid-p confidence interval when resampling is used to determine the null distribution. The two-sided mid-p-value is computed according to the minimum likelihood method (Hirji et al., 1991).

The p-value interval $(p_0, p_1]$ obtained by pvalue_interval was proposed by Berger (2000, 2001), where the upper endpoint p_1 is the conventional p-value and the mid-point, i.e., $p_{0.5}$, is the mid-p-value. The lower endpoint p_0 is the smallest p-value obtainable if no conservatism attributable to the discreteness of the null distribution is present. The length of the p-value interval is the null probability of the observed outcome and provides a data-dependent measure of conservatism that is completely independent of the significance level.

Value

The p-value, mid-p-value or p-value interval computed from object. A numeric vector or matrix.

64 pvalue-methods

Note

The mid-p-value and p-value interval of asymptotic permutation distributions for maximum-type tests or exact permutation distributions obtained by the split-up algoritm is reported as NA.

In versions of **coin** prior to 1.1-0, a min-P procedure computing Šidák single-step adjusted p-values accounting for discreteness was available when specifying method = "discrete". **This is now deprecated and will be removed in a future release** due to the introduction of a more general max-T version of the same algorithm.

References

Berger, V. W. (2000). Pros and cons of permutation tests in clinical trials. *Statistics in Medicine* **19**(10), 1319–1328.

Berger, V. W. (2001). The p-value interval as an inferential tool. The Statistician 50(1), 79-85.

Bretz, F., Hothorn, T. and Westfall, P. (2011). *Multiple Comparisons Using R.* Boca Raton: CRC Press.

Hirji, K. F., Tan, S.-J. and Elashoff, R. M. (1991). A quasi-exact test for comparing two binomial proportions. *Statistics in Medicine* **10**(7), 1137–1153.

Westfall, P. H. and Troendle, J. F. (2008). Multiple testing with minimal assumptions. *Biometrical Journal* **50**(5), 745–755.

Westfall, P. H. and Wolfinger, R. D. (1997). Multiple tests with discrete distributions. *The American Statistician* **51**(1), 3–8.

Westfall, P. H. and Young, S. S. (1993). *Resampling-Based Multiple Testing: Examples and Methods for p-Value Adjustment*. New York: John Wiley & Sons.

```
## Two-sample problem
dta <- data.frame(</pre>
   y = rnorm(20),
    x = g1(2, 10)
## Exact Ansari-Bradley test
(at <- ansari_test(y ~ x, data = dta, distribution = "exact"))</pre>
pvalue(at)
midpvalue(at)
pvalue_interval(at)
## Bivariate two-sample problem
dta2 <- data.frame(</pre>
   y1 = rnorm(20) + rep(0:1, each = 10),
   y2 = rnorm(20),
    x = g1(2, 10)
)
## Approximative (Monte Carlo) bivariate Fisher-Pitman test
(it <- independence_test(y1 + y2 \sim x, data = dta2,
```

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```
distribution = approximate(B = 10000)))
## Global p-value
pvalue(it)
## Joint distribution single-step p-values
pvalue(it, method = "single-step")
## Joint distribution step-down p-values
pvalue(it, method = "step-down")
## Sidak step-down p-values
pvalue(it, method = "step-down", distribution = "marginal", type = "Sidak")
## Unadjusted p-values
pvalue(it, method = "unadjusted")
## Length of YOY Gizzard Shad (Hollander and Wolfe, 1999, p. 200, Tab. 6.3)
yoy <- data.frame(</pre>
    length = c(46, 28, 46, 37, 32, 41, 42, 45, 38, 44,
               42, 60, 32, 42, 45, 58, 27, 51, 42, 52,
               38, 33, 26, 25, 28, 28, 26, 27, 27, 27,
               31, 30, 27, 29, 30, 25, 25, 24, 27, 30),
    site = gl(4, 10, labels = as.roman(1:4))
)
## Approximative (Monte Carlo) Fisher-Pitman test with contrasts
## Note: all pairwise comparisons
(it <- independence_test(length ~ site, data = yoy,</pre>
                         distribution = approximate(B = 10000),
                         xtrafo = mcp_trafo(site = "Tukey")))
## Joint distribution step-down p-values
pvalue(it, method = "step-down") # subset pivotality is violated
```

rotarod

Rotating Rats

Description

The endurance time of 24 rats in two groups on a rotating cylinder.

Usage

rotarod

66 Scale Tests

Format

A data frame with 24 observations on 2 variables.

```
time endurance time (seconds).
group a factor with levels "control" and "treatment".
```

Details

The rats were randomly assigned to receive a fixed oral dose of a centrally acting muscle relaxant ("treatment") or a saline solvent ("control"). The animals were placed on a rotating cylinder and the endurance time of each rat, i.e., the length of time each rat remained on the cylinder, was measured up to a maximum of 300 seconds.

This dataset is the basis of a comparison of 11 different software implementations of the Wilcoxon-Mann-Whitney test presented in Bergmann, Ludbrook and Spooren (2000).

Note

The empirical variance in the control group is 0 and the group medians are identical. The exact conditional *p*-values are 0.0373 (two-sided) and 0.0186 (one-sided). The asymptotic two-sided *p*-value (corrected for ties) is 0.0147.

Source

Bergmann, R., Ludbrook, J. and Spooren, W. P. J. M. (2000). Different outcomes of the Wilcoxon-Mann-Whitney test from different statistics packages. *The American Statistician* **54**(1), 72–77.

Examples

ScaleTests

Two- and K-Sample Scale Tests

Description

Testing the equality of the distributions of a numeric response variable in two or more independent groups against scale alternatives.

ScaleTests 67

Usage

```
## S3 method for class 'formula'
taha_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'IndependenceProblem'
taha_test(object, conf.int = FALSE, conf.level = 0.95, ...)
## S3 method for class 'formula'
klotz_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'IndependenceProblem'
klotz_test(object, ties.method = c("mid-ranks", "average-scores"),
           conf.int = FALSE, conf.level = 0.95, ...)
## S3 method for class 'formula'
mood_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'IndependenceProblem'
mood_test(object, ties.method = c("mid-ranks", "average-scores"),
          conf.int = FALSE, conf.level = 0.95, ...)
## S3 method for class 'formula'
ansari_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'IndependenceProblem'
ansari_test(object, ties.method = c("mid-ranks", "average-scores"),
            conf.int = FALSE, conf.level = 0.95, ...)
## S3 method for class 'formula'
fligner_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'IndependenceProblem'
fligner_test(object, ties.method = c("mid-ranks", "average-scores"),
             conf.int = FALSE, conf.level = 0.95, ...)
## S3 method for class 'formula'
conover_test(formula, data, subset = NULL, weights = NULL, ...)
## S3 method for class 'IndependenceProblem'
conover_test(object, conf.int = FALSE, conf.level = 0.95, ...)
```

Arguments

formula	a formula of the form $y \sim x \mid$ block where y is a numeric variable, x is a factor and block is an optional factor for stratification.
data	an optional data frame containing the variables in the model formula.
subset	an optional vector specifying a subset of observations to be used. Defaults to \ensuremath{NULL} .
weights	an optional formula of the form ~ w defining integer valued case weights for each observation. Defaults to NULL, implying equal weight for all observations.
object	an object inheriting from class "IndependenceProblem".

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conf.int a logical indicating whether a confidence interval for the ratio of scales should be computed. Defaults to FALSE.

conf.level a numeric, confidence level of the interval. Defaults to 0.95.

ties.method a character, the method used to handle ties: the score generating function either uses mid-ranks ("mid-ranks", default) or averages the scores of randomly broken ties ("average-scores").

... further arguments to be passed to independence_test.

Details

taha_test, klotz_test, mood_test, ansari_test, fligner_test and conover_test provide the Taha test, the Klotz test, the Mood test, the Ansari-Bradley test, the Fligner-Killeen test and the Conover-Iman test. A general description of these methods is given by Hollander and Wolfe (1999). For the adjustment of scores for tied values see Hájek, Šidák and Sen (1999, pp. 133–135).

The null hypothesis of equality, or conditional equality given block, of the distribution of y in the groups defined by x is tested against scale alternatives. In the two-sample case, the two-sided null hypothesis is $H_0: V(Y_1)/V(Y_2) = 1$, where $V(Y_s)$ is the variance of the responses in the sth sample. In case alternative = "less", the null hypothesis is $H_0: V(Y_1)/V(Y_2) \geq 1$. When alternative = "greater", the null hypothesis is $H_0: V(Y_1)/V(Y_2) \leq 1$. Confidence intervals for the ratio of scales are available and computed according to Bauer (1972).

The Fligner-Killeen test uses median centering in each of the samples, as suggested by Conover, Johnson and Johnson (1981), whereas the Conover-Iman test, following Conover and Iman (1978), uses mean centering in each of the samples.

The conditional null distribution of the test statistic is used to obtain p-values and an asymptotic approximation of the exact distribution is used by default (distribution = "asymptotic"). Alternatively, the distribution can be approximated via Monte Carlo resampling or computed exactly for univariate two-sample problems by setting distribution to "approximate" or "exact" respectively. See asymptotic, approximate and exact for details.

Value

An object inheriting from class "IndependenceTest". Confidence intervals can be extracted by confint.

Note

In the two-sample case, a *large* value of the Ansari-Bradley statistic indicates that sample 1 is *less* variable than sample 2, whereas a *large* value of the statistics due to Taha, Klotz, Mood, Fligner-Killeen, and Conover-Iman indicate that sample 1 is *more* variable than sample 2.

References

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```
## Serum Iron Determination Using Hyland Control Sera
## Hollander and Wolfe (1999, p. 147, Tab 5.1)
sid <- data.frame(</pre>
    serum = c(111, 107, 100, 99, 102, 106, 109, 108, 104, 99,
              101, 96, 97, 102, 107, 113, 116, 113, 110, 98,
              107, 108, 106, 98, 105, 103, 110, 105, 104,
              100, 96, 108, 103, 104, 114, 114, 113, 108, 106, 99),
    method = gl(2, 20, labels = c("Ramsay", "Jung-Parekh"))
)
## Asymptotic Ansari-Bradley test
ansari_test(serum ~ method, data = sid)
## Exact Ansari-Bradley test
pvalue(ansari_test(serum ~ method, data = sid,
                   distribution = "exact"))
## Platelet Counts of Newborn Infants
## Hollander and Wolfe (1999, p. 171, Tab. 5.4)
platelet <- data.frame(</pre>
    counts = c(120, 124, 215, 90, 67, 95, 190, 180, 135, 399,
               12, 20, 112, 32, 60, 40),
    treatment = factor(rep(c("Prednisone", "Control"), c(10, 6)))
)
## Approximative (Monte Carlo) Lepage test
## Hollander and Wolfe (1999, p. 172)
lepage_trafo <- function(y)</pre>
    cbind("Location" = rank_trafo(y), "Scale" = ansari_trafo(y))
independence_test(counts ~ treatment, data = platelet,
                  distribution = approximate(B = 10000),
                  ytrafo = function(data)
                      trafo(data, numeric_trafo = lepage_trafo),
                  teststat = "quadratic")
## Why was the null hypothesis rejected?
## Note: maximum statistic instead of quadratic form
ltm <- independence_test(counts ~ treatment, data = platelet,</pre>
                         distribution = approximate(B = 10000),
                         ytrafo = function(data)
```

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```
trafo(data, numeric_trafo = lepage_trafo))
## Step-down adjustment suggests a difference in location
pvalue(ltm, method = "step-down")
## The same results are obtained from the simple Sidak-Holm procedure since the
## correlation between Wilcoxon and Ansari-Bradley test statistics is zero
cov2cor(covariance(ltm))
pvalue(ltm, method = "step-down", distribution = "marginal", type = "Sidak")
```

statistic-methods

Extraction of the Test Statistic and Linear Statistic

Description

Methods for extraction of the test statistic and linear statistic.

Usage

```
## S4 method for signature 'IndependenceLinearStatistic'
statistic(object, type = c("test", "linear", "standardized"), ...)
## S4 method for signature 'IndependenceTest'
statistic(object, type = c("test", "linear", "standardized"), ...)
## S4 method for signature 'IndependenceTestStatistic'
statistic(object, type = c("test", "linear", "standardized"), ...)
```

Arguments

object an object from which the test statistic or linear statistic can be extracted.

type a character, the type of statistic: either "test" (default) for the test statistic,
 "linear" for the unstandardized linear statistic or "standardized" for the
 standardized linear statistic.

further arguments (currently ignored).

Details

The method statistic extracts the test statistic or the, possibly multivariate, linear statistic in its unstandardized or standardized form.

The test statistic (type = "test") is returned by default. The unstandardized or standardized linear statistic is obtained by setting type to "linear" or "standardized" respectively.

Value

The test statistic or the unstandardized or standardized linear statistic extracted from object. A numeric vector or matrix.

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Examples

```
## Example data
dta <- data.frame(</pre>
   y = gl(4, 5),
   x = gl(5, 4)
)
## Asymptotic Cochran-Mantel-Haenszel Test
ct <- cmh_test(y \sim x, data = dta)
## Test statistic
statistic(ct)
## The unstandardized linear statistic...
statistic(ct, type = "linear")
## ...is identical to the contingency table
xtabs(~x + y, data = dta)
## Illustrating departures from the null hypothesis of independence using the
## standardized linear statistic
statistic(ct, type = "standardized")
```

SurvivalTests

Two- and K-Sample Tests for Censored Data

Description

Testing the equality of the survival distributions in two or more independent groups.

Usage

Arguments

formula a formula of the form y ~ x | block where y is a survival object (see Surv in package survival), x is a factor and block is an optional factor for stratification.

data an optional data frame containing the variables in the model formula.

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subset an optional vector specifying a subset of observations to be used. Defaults to

NULL.

weights an optional formula of the form ~ w defining integer valued case weights for

each observation. Defaults to NULL, implying equal weight for all observations.

object an object inheriting from class "IndependenceProblem".

ties.method a character, the method used to handle ties: the score generating function either

uses mid-ranks ("mid-ranks", default), the Hothorn-Lausen method ("Hothorn-Lausen")

or averages the scores of randomly broken ties ("average-scores"); see 'De-

tails'.

type a character, the type of test: either "logrank" (default), "Gehan-Breslow",

"Tarone-Ware", "Prentice", "Prentice-Marek", "Andersen-Borgan-Gill-Keiding",

"Fleming-Harrington", "Gaugler-Kim-Liao" or "Self"; see 'Details'.

rho a numeric, the ρ constant when type is "Tarone-Ware", "Fleming-Harrington",

"Gaugler-Kim-Liao" or "Self"; see 'Details'. Defaults to NULL, implying 0.5

for type = "Tarone-Ware" and 0 otherwise.

gamma a numeric, the γ constant when type is "Fleming-Harrington", "Gaugler-Kim-Liao"

or "Self"; see 'Details'. Defaults to NULL, implying 0.

... further arguments to be passed to independence_test.

Details

logrank_test provides the weighted logrank test reformulated as a linear rank test. The family of weighted logrank tests encompasses a large collection of tests commonly used in the analysis of survival data including, but not limited to, the standard (unweighted) logrank test, the Gehan-Breslow test, the Tarone-Ware class of tests, the Prentice test, the Prentice-Marek test, the Andersen-Borgan-Gill-Keiding test, the Fleming-Harrington class of tests and the Self class of tests. A general description of these methods is given by Klein and Moeschberger (2003, Ch. 7). See Letón and Zuluaga (2001) for the linear rank test formulation.

The null hypothesis of equality, or conditional equality given block, of the survival distribution of y in the groups defined by x is tested. In the two-sample case, the two-sided null hypothesis is $H_0: \theta = 1$, where $\theta = \lambda_2/\lambda_1$ and λ_s is the hazard rate in the sth sample. In case alternative = "less", the null hypothesis is $H_0: \theta \geq 1$, i.e., the survival is lower in sample 1 than in sample 2. When alternative = "greater", the null hypothesis is $H_0: \theta \leq 1$, i.e., the survival is higher in sample 1 than in sample 2.

If x is an ordered factor, the default scores, 1:nlevels(x), can be altered using the scores argument (see independence_test); this argument can also be used to coerce nominal factors to class "ordered". In this case, a linear-by-linear association test is computed and the direction of the alternative hypothesis can be specified using the alternative argument. This type of extension of the standard logrank test was given by Tarone (1975) and later generalized to general weights by Tarone and Ware (1977).

Let (t_i, δ_i) , $i = 1, 2, \ldots, n$, represent a right-censored random sample of size n, where t_i is the observed survival time and δ_i is the status indicator $(\delta_i$ is 0 for right-censored observations and 1 otherwise). To allow for ties in the data, let $t_{(1)} < t_{(2)} < \cdots < t_{(m)}$ represent the $m, m \leq n$, ordered distinct event times. At time $t_{(k)}$, $k = 1, 2, \ldots, m$, the number of events and the number of subjects at risk are given by $d_k = \sum_{i=1}^n I(t_i = t_{(k)} | \delta_i = 1)$ and $n_k = n - r_k$, respectively, where r_k depends on the ties handling method.

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Three different methods of handling ties are available using ties.method: mid-ranks ("mid-ranks", default), the Hothorn-Lausen method ("Hothorn-Lausen") and average-scores ("average-scores"). The first and last method are discussed and contrasted by Callaert (2003), whereas the second method is defined in Hothorn and Lausen (2003). The mid-ranks method leads to

$$r_k = \sum_{i=1}^n I(t_i < t_{(k)})$$

whereas the Hothorn-Lausen method uses

$$r_k = \sum_{i=1}^n I(t_i \le t_{(k)}) - 1.$$

The scores assigned to right-censored and uncensored observations at the kth event time are given by

$$C_k = \sum_{j=1}^k w_j \frac{d_j}{n_j}$$
 and $c_k = C_k - w_k$,

respectively, where w is the logrank weight. For the average-scores method, used by, e.g., the software package StatXact, the d_k events observed at the kth event time are arbitrarily ordered by assigning them distinct values $t_{(k_l)}$, $l=1,2,\ldots,d_k$, infinitesimally to the left of $t_{(k)}$. Then scores C_{k_l} and c_{k_l} are computed as indicated above, effectively assuming that no event times are tied. The scores C_k and c_k are assigned the average of the scores C_{k_l} and c_{k_l} respectively. It then follows that the score for the ith subject is

$$a_i = \begin{cases} C_{k'} & \text{if } \delta_i = 0\\ c_{k'} & \text{otherwise} \end{cases}$$

where $k' = \max\{k : t_i \ge t_{(k)}\}.$

The type argument allows for a choice between some of the most well-known members of the family of weighted logrank tests, each corresponding to a particular weight function. The standard logrank test ("logrank", default) was suggested by Mantel (1966), Peto and Peto (1972) and Cox (1972) and has $w_k=1$. The Gehan-Breslow test ("Gehan-Breslow") proposed by Gehan (1965) and later extended to K samples by Breslow (1970) is a generalization of the Wilcoxon rank-sum test, where $w_k=n_k$. The Tarone-Ware class of tests ("Tarone-Ware") discussed by Tarone and Ware (1977) has $w_k=n_k^\rho$, where ρ is a constant; $\rho=0.5$ (default) was suggested by Tarone and Ware (1977), but note that $\rho=0$ and $\rho=1$ lead to the the standard logrank test and Gehan-Breslow test respectively. The Prentice test ("Prentice") is another generalization of the Wilcoxon rank-sum test proposed by Prentice (1978), where

$$w_k = \prod_{j=1}^k \frac{n_j}{n_j + d_j}.$$

The Prentice-Marek test ("Prentice-Marek") is yet another generalization of the Wilcoxon ranksum test discussed by Prentice and Marek (1979), with

$$w_k = \tilde{S}_k = \prod_{j=1}^k \frac{n_j + 1 - d_j}{n_j + 1}.$$

The Andersen-Borgan-Gill-Keiding test ("Andersen-Borgan-Gill-Keiding") suggested by Andersen *et al.* (1982) is a modified version of the Prentice-Marek test using

$$w_k = \frac{n_k}{n_k + 1} \prod_{j=0}^{k-1} \frac{n_j + 1 - d_j}{n_j + 1}$$

where $n_0 \equiv n$ and $d_0 \equiv 0$. The Fleming-Harrington class of tests ("Fleming-Harrington") proposed by Fleming and Harrington (1991) uses $w_k = \hat{S}_k^{\rho} (1 - \hat{S}_k)^{\gamma}$, where ρ and γ are constants and

$$\hat{S}_k = \prod_{j=0}^{k-1} \frac{n_j - d_j}{n_j}, \quad \hat{S}_0 \equiv 1$$

is the *left-continuous* Kaplan-Meier estimator of the survival function; $\rho=0$ and $\gamma=0$ lead to the standard logrank test. The Gaugler-Kim-Liao class of tests ("Gaugler-Kim-Liao") discussed by Gaugler *et al.* (2007) is a modified version of the Fleming-Harrington class of tests, replacing \hat{S}_k with \tilde{S}_k so that $w_k=\tilde{S}_k^\rho(1-\tilde{S}_k)^\gamma$, where ρ and γ are constants; $\rho=0$ and $\gamma=0$ lead to the standard logrank test. The Self class of tests ("Self") suggested by Self (1991) has $w_k=v_k^\rho(1-v_k)^\gamma$, where

$$v_k = \frac{1}{2} \frac{t_{(k-1)} + t_{(k)}}{t_{(m)}}, \quad t_{(0)} \equiv 0$$

is the standardized mid-point between the (k-1)th and the kth event time. (This is a slight generalization of Self's original proposal in order to allow for non-integer follow-up times.) Again, ρ and γ are constants and $\rho=0$ and $\gamma=0$ lead to the standard logrank test.

The conditional null distribution of the test statistic is used to obtain *p*-values and an asymptotic approximation of the exact distribution is used by default (distribution = "asymptotic"). Alternatively, the distribution can be approximated via Monte Carlo resampling or computed exactly for univariate two-sample problems by setting distribution to "approximate" or "exact" respectively. See asymptotic, approximate and exact for details.

Value

An object inheriting from class "IndependenceTest".

Note

Peto and Peto (1972) proposed the test statistic implemented in logrank_test and named it the *logrank test*. However, the Mantel-Cox test (Mantel, 1966; Cox, 1972), as implemented in survdiff (in package survival), is also known as the logrank test. These tests are similar, but differ in the choice of probability model: the (Peto-Peto) logrank test uses the permutational variance, whereas the Mantel-Cox test is based on the hypergeometric variance.

Combining independence_test or symmetry_test with logrank_trafo offers more flexibility than logrank_test and allows for, among other things, maximum-type versatile test procedures (e.g., Lee, 1996; see 'Examples') and user-supplied logrank weights (see GTSG for tests against Weibull-type or crossing-curve alternatives).

Starting with version 1.1-0, logrank_test replaced surv_test which was made **defunct** in version 1.2-0. Furthermore, logrank_trafo is now an increasing function for all choices of ties.method, implying that the test statistic has the same sign irrespective of the ties handling method. Consequently, the sign of the test statistic will now be the opposite of what it was in earlier versions unless

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ties.method = "average-scores". (In versions of **coin** prior to 1.1-0, logrank_trafo was a decreasing function when ties.method was other than "average-scores".)

Starting with version 1.2-0, mid-ranks and the Hothorn-Lausen method can no longer be specified with ties.method = "logrank" and ties-method = "HL" respectively.

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Examples

```
## Example data (Callaert, 2003, Tab.1)
callaert <- data.frame(</pre>
    time = c(1, 1, 5, 6, 6, 6, 6, 2, 2, 2, 3, 4, 4, 5, 5),
    group = factor(rep(0:1, c(7, 8)))
)
## Logrank scores using mid-ranks (Callaert, 2003, Tab.2)
with(callaert.
     logrank_trafo(Surv(time)))
## Asymptotic Mantel-Cox test (p = 0.0523)
survdiff(Surv(time) ~ group, data = callaert)
## Exact logrank test using mid-ranks (p = 0.0505)
logrank_test(Surv(time) ~ group, data = callaert, distribution = "exact")
## Exact logrank test using average-scores (p = 0.0468)
logrank_test(Surv(time) ~ group, data = callaert, distribution = "exact",
             ties.method = "average-scores")
## Lung cancer data (StatXact 9 manual, p. 213, Tab. 7.19)
lungcancer <- data.frame(</pre>
    time = c(257, 476, 355, 1779, 355,
             191, 563, 242, 285, 16, 16, 16, 257, 16),
    event = c(0, 0, 1, 1, 0,
              1, 1, 1, 1, 1, 1, 1, 1, 1),
    group = factor(rep(1:2, c(5, 9)),
                   labels = c("newdrug", "control"))
)
## Logrank scores using average-scores (StatXact 9 manual, p. 214)
     logrank_trafo(Surv(time, event), ties.method = "average-scores"))
## Exact logrank test using average-scores (StatXact 9 manual, p. 215)
logrank_test(Surv(time, event) ~ group, data = lungcancer,
             distribution = "exact", ties.method = "average-scores")
## Exact Prentice test using average-scores (StatXact 9 manual, p. 222)
logrank_test(Surv(time, event) ~ group, data = lungcancer,
             distribution = "exact", ties.method = "average-scores",
             type = "Prentice")
## Approximative (Monte Carlo) versatile test (Lee, 1996)
rho.gamma \leftarrow expand.grid(rho = seq(0, 2, 1), gamma = seq(0, 2, 1))
lee_trafo <- function(y)</pre>
    logrank_trafo(y, ties.method = "average-scores",
                  type = "Fleming-Harrington",
                  rho = rho.gamma["rho"], gamma = rho.gamma["gamma"])
```

SymmetryProblem-class Class "SymmetryProblem"

Description

Objects of class "SymmetryProblem" represent the data structure corresponding to a symmetry problem.

Objects from the Class

Objects can be created by calls of the form

```
new("SymmetryProblem", x, y, block = NULL, weights = NULL, ...)
```

where x and y are data frames containing the variables X and Y respectively, block is an optional factor representing the block structure b and weights is an optional integer vector corresponding to the case weights w.

Slots

```
x: Object of class "data.frame". The variables x.y: Object of class "data.frame". The variables y.block: Object of class "factor". The block structure.weights: Object of class "numeric". The case weights.
```

Extends

```
Class "IndependenceProblem", directly.
```

Methods

initialize signature(.Object = "SymmetryProblem"): See the documentation for initialize
 (in package methods) for details.

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General Symmetry Test

Description

Testing the symmetry of set of repeated measurements variables measured on arbitrary scales in a complete block design.

Usage

Arguments

guments	
formula	a formula of the form y1 + + yq \sim x block where y1,, yq are measured on arbitrary scales (nominal, ordinal or continuous with or without censoring), x is a factor and block is an optional factor (which is generated automatically if omitted).
data	an optional data frame containing the variables in the model formula.
subset	an optional vector specifying a subset of observations to be used. Defaults to NULL.
weights	an optional formula of the form ~ w defining integer valued case weights for each observation. Defaults to NULL, implying equal weight for all observations.
object	an object inheriting from classes "table" (with identical dimnames components) or "SymmetryProblem".
teststat	a character, the type of test statistic to be applied: either a maximum statistic ("maximum", default), a quadratic form ("quadratic") or a standardized scalar test statistic ("scalar").
distribution	a character, the conditional null distribution of the test statistic can be approximated by its asymptotic distribution ("asymptotic", default) or via Monte Carlo resampling ("approximate"). Alternatively, the functions asymptotic or approximate can be used. For univariate two-sample problems, "exact" or use of the function exact computes the exact distribution. Computation of the null distribution can be suppressed by specifying "none". It is also

possible to specify a function with one argument (an object inheriting from "IndependenceTestStatistic") that returns an object of class "NullDistribution".

alternative	a character, the alternative hypothesis: either "two.sided" (default), "greater" or "less".
xtrafo	a function of transformations to be applied to the factor x supplied in formula; see 'Details'. Defaults to trafo.
ytrafo	a function of transformations to be applied to the variables y1,, yq supplied in formula; see 'Details'. Defaults to trafo.
scores	a named list of scores to be attached to ordered factors; see 'Details'. Defaults to NULL, implying equally spaced scores.
check	a function to be applied to objects of class "IndependenceTest" in order to check for specific properties of the data. Defaults to NULL.
paired	a logical, indicating that paired data have been transformed in such a way that the (unstandardized) linear statistic is the sum of the absolute values of the positive differences between the paired observations. Defaults to FALSE.
	further arguments to be passed to or from other methods (currently ignored).

Details

symmetry_test provides a general symmetry test for a set of variables measured on arbitrary scales. This function is based on the general framework for conditional inference procedures proposed by Strasser and Weber (1999). The salient parts of the Strasser-Weber framework are elucidated by Hothorn *et al.* (2006) and a thorough description of the software implementation is given by Hothorn *et al.* (2008).

The null hypothesis of symmetry is tested. The response variables and the measurement conditions are given by y1, ..., yq and x, respectively, and block is a factor where each level corresponds to exactly one subject with repeated measurements.

A vector of case weights, e.g., observation counts, can be supplied through the weights argument and the type of test statistic is specified by the teststat argument. Influence and regression functions, i.e., transformations of y1, ..., yq and x, are specified by the ytrafo and xtrafo arguments respectively; see trafo for the collection of transformation functions currently available. This allows for implementation of both novel and familiar test statistics, e.g., the McNemar test, the Cochran Q test, the Wilcoxon signed-rank test and the Friedman test. Furthermore, multivariate extensions such as the multivariate Friedman test (Gerig, 1969; Puri and Sen, 1971) can be implemented without much effort (see 'Examples').

If, say, y1 and/or x are ordered factors, the default scores, 1:nlevels(y1) and 1:nlevels(x) respectively, can be altered using the scores argument; this argument can also be used to coerce nominal factors to class "ordered". For example, when y1 is an ordered factor with four levels and x is a nominal factor with three levels, scores = list(y1 = c(1, 3:5), x = c(1:2, 4)) supplies the scores to be used. For ordered alternatives the scores must be monotonic, but non-montonic scores are also allowed for testing against, e.g., umbrella alternatives. The length of the score vector must be equal to the number of factor levels.

The conditional null distribution of the test statistic is used to obtain *p*-values and an asymptotic approximation of the exact distribution is used by default (distribution = "asymptotic"). Alternatively, the distribution can be approximated via Monte Carlo resampling or computed exactly for univariate two-sample problems by setting distribution to "approximate" or "exact" respectively. See asymptotic, approximate and exact for details.

Value

An object inheriting from class "IndependenceTest".

Note

Starting with **coin** version 1.1-0, maximum statistics and quadratic forms can no longer be specified using teststat = "maxtype" and teststat = "quadtype" respectively (as was used in versions prior to 0.4-5).

References

Gerig, T. (1969). A multivariate extension of Friedman's χ_r^2 test. *Journal of the American Statistical Association* **64**(328), 1595–1608.

Hothorn, T., Hornik, K., van de Wiel, M. A. and Zeileis, A. (2006). A Lego system for conditional inference. *The American Statistician* **60**(3), 257–263.

Hothorn, T., Hornik, K., van de Wiel, M. A. and Zeileis, A. (2008). Implementing a class of permutation tests: the coin package. *Journal of Statistical Software* **28**(8), 1–23. http://www.jstatsoft.org/v28/i08/

Puri, M. L. and Sen, P. K. (1971). *Nonparametric Methods in Multivariate Analysis*. New York: John Wiley & Sons.

Strasser, H. and Weber, C. (1999). On the asymptotic theory of permutation statistics. *Mathematical Methods of Statistics* **8**(2), 220–250.

Examples

```
## One-sided exact Fisher-Pitman test for paired observations
y1 \leftarrow c(1.83, 0.50, 1.62, 2.48, 1.68, 1.88, 1.55, 3.06, 1.30)
y2 <- c(0.878, 0.647, 0.598, 2.05, 1.06, 1.29, 1.06, 3.14, 1.29)
dta <- data.frame(</pre>
    y = c(y1, y2),
    x = gl(2, length(y1)),
    block = factor(rep(seq_along(y1), 2))
)
symmetry_test(y \sim x \mid block, data = dta,
              distribution = "exact", alternative = "greater")
## Alternatively: transform data and set 'paired = TRUE'
delta <- y1 - y2
y <- as.vector(rbind(abs(delta) * (delta >= 0), abs(delta) * (delta < 0)))
x \leftarrow factor(rep(0:1, length(delta)), labels = c("pos", "neg"))
block <- gl(length(delta), 2)</pre>
symmetry_test(y \sim x \mid block,
              distribution = "exact", alternative = "greater",
              paired = TRUE)
```

Example data

```
### Gerig (1969, p. 1597)
gerig <- data.frame(</pre>
   y1 = c(0.547, 1.811, 2.561,
            1.706, 2.509, 1.414,
           -0.288, 2.524, 3.310,
            1.417, 0.703, 0.961,
            0.878, 0.094, 1.682,
           -0.680, 2.077, 3.181,
            0.056, 0.542, 2.983,
            0.711, 0.269, 1.662,
           -1.335, 1.545, 2.920,
            1.635, 0.200, 2.065),
   y2 = c(-0.575, 1.840, 2.399,
            1.252, 1.574, 3.059,
           -0.310, 1.553, 0.560,
            0.932, 1.390, 3.083,
            0.819, 0.045, 3.348,
            0.497, 1.747, 1.355,
           -0.285, 0.760, 2.332,
            0.089, 1.076, 0.960,
           -0.349, 1.471, 4.121,
            0.845, 1.480, 3.391),
   x = factor(rep(1:3, 10)),
   b = factor(rep(1:10, each = 3))
)
### Asymptotic multivariate Friedman test
### Gerig (1969, p. 1599)
symmetry_test(y1 + y2 \sim x \mid b, data = gerig, teststat = "quadratic",
              ytrafo = function(data)
                  trafo(data, numeric_trafo = rank_trafo,
                        block = gerig$b)) # L_n = 17.238
### Asymptotic multivariate Page test
(st <- symmetry_test(y1 + y2 ~ x | b, data = gerig,
                     ytrafo = function(data)
                          trafo(data, numeric_trafo = rank_trafo,
                                block = gerig$b),
                     scores = list(x = 1:3)))
pvalue(st, method = "step-down")
```

SymmetryTests

Symmetry Tests

Description

Testing the symmetry of a numeric repeated measurements variable in a complete block design.

Usage

```
## S3 method for class 'formula'
sign_test(formula, data, subset = NULL, ...)
## S3 method for class 'SymmetryProblem'
sign_test(object, ...)
## S3 method for class 'formula'
wilcoxsign_test(formula, data, subset = NULL, ...)
## S3 method for class 'SymmetryProblem'
wilcoxsign_test(object, zero.method = c("Pratt", "Wilcoxon"), ...)
## S3 method for class 'formula'
friedman_test(formula, data, subset = NULL, ...)
## S3 method for class 'SymmetryProblem'
friedman_test(object, ...)
## S3 method for class 'formula'
quade_test(formula, data, subset = NULL, ...)
## S3 method for class 'SymmetryProblem'
quade_test(object, ...)
```

Arguments

formula	a formula of the form y ~ x block where y is a numeric variable, x is a factor with two (sign_test and wilcoxsign_test) or more levels and block is an optional factor (which is generated automatically if omitted).	
data	an optional data frame containing the variables in the model formula.	
subset	an optional vector specifying a subset of observations to be used. Defaults to NULL.	
object	an object inheriting from class "SymmetryProblem".	
zero.method	a character, the method used to handle zeros: either "Pratt" (default) or "Wilcoxon"; see 'Details'.	
	further arguments to be passed to symmetry_test.	

Details

sign_test, wilcoxsign_test, friedman_test and quade_test provide the sign test, the Wilcoxon signed-rank test, the Friedman test, the Page test and the Quade test. A general description of these methods is given by Hollander and Wolfe (1999).

The null hypothesis of symmetry is tested. The response variable and the measurement conditions are given by y and x, respectively, and block is a factor where each level corresponds to exactly one subject with repeated measurements. For sign_test and wilcoxsign_test, formulae of the form $y \sim x \mid block$ and $y \sim x$ are allowed. The latter form is interpreted as y is the first and x the second measurement on the same subject.

If x is an ordered factor, the default scores, 1:nlevels(x), can be altered using the scores argument (see symmetry_test); this argument can also be used to coerce nominal factors to class

"ordered". In this case, a linear-by-linear association test is computed and the direction of the alternative hypothesis can be specified using the alternative argument. For the Friedman test, this extension was given by Page (1963) and is known as the Page test.

For wilcoxsign_test, the default method of handling zeros (zero.method = "Pratt"), due to Pratt (1959), first rank transforms the absolute differences (including zeros) and then discards the ranks corresponding to the zero-differences. The proposal by Wilcoxon (1949) first discards the zero-differences and then rank transforms the remaining absolute differences (zero.method = "Wilcoxon").

The conditional null distribution of the test statistic is used to obtain *p*-values and an asymptotic approximation of the exact distribution is used by default (distribution = "asymptotic"). Alternatively, the distribution can be approximated via Monte Carlo resampling or computed exactly for univariate two-sample problems by setting distribution to "approximate" or "exact" respectively. See asymptotic, approximate and exact for details.

Value

An object inheriting from class "IndependenceTest".

Note

Starting with **coin** version 1.0-16, the zero.method argument replaced the (now removed) ties.method argument. The current default is zero.method = "Pratt" whereas earlier versions had ties.method = "HollanderWolfe" which is equivalent to zero.method = "Wilcoxon".

References

Hollander, M. and Wolfe, D. A. (1999). *Nonparametric Statistical Methods*, Second Edition. New York: John Wiley & Sons.

Page, E. B. (1963). Ordered hypotheses for multiple treatments: a significance test for linear ranks. *Journal of the American Statistical Association* **58**(301), 216–230.

Pratt, J. W. (1959). Remarks on zeros and ties in the Wilcoxon signed rank procedures. *Journal of the American Statistical Association* **54**(287), 655–667.

Quade, D. (1979). Using weighted rankings in the analysis of complete blocks with additive block effects. *Journal of the American Statistical Association* **74**(367), 680–683.

Wilcoxon, F. (1949). Some rapid approximate statistical procedures. *Annals of the New York Academy of Sciences* **52**(6), 808–814.

Examples

```
(wt <- wilcoxsign_test(y1 ~ y2, distribution = "exact",</pre>
                       alternative = "greater"))
statistic(wt, type = "linear")
midpvalue(wt) # mid-p-value
## Comparison with R's wilcox.test() function
wilcox.test(y1, y2, paired = TRUE, alternative = "greater")
## Data with explicit group and block information
dta <- data.frame(y = c(y1, y2), x = gl(2, length(y1)),
                  block = factor(rep(seq_along(y1), 2)))
## For two samples, the sign test is equivalent to the Friedman test...
sign_test(y ~ x | block, data = dta, distribution = "exact")
friedman_test(y ~ x | block, data = dta, distribution = "exact")
\#\# ...and the signed-rank test is equivalent to the Quade test
wilcoxsign_test(y ~ x | block, data = dta, distribution = "exact")
quade_test(y ~ x | block, data = dta, distribution = "exact")
## Comparison of three methods ("round out", "narrow angle", and "wide angle")
## for rounding first base.
## Hollander and Wolfe (1999, p. 274, Tab. 7.1)
rounding <- data.frame(</pre>
    times = c(5.40, 5.50, 5.55,
              5.85, 5.70, 5.75,
              5.20, 5.60, 5.50,
              5.55, 5.50, 5.40,
              5.90, 5.85, 5.70,
              5.45, 5.55, 5.60,
              5.40, 5.40, 5.35,
              5.45, 5.50, 5.35,
              5.25, 5.15, 5.00,
              5.85, 5.80, 5.70,
              5.25, 5.20, 5.10,
              5.65, 5.55, 5.45,
              5.60, 5.35, 5.45,
              5.05, 5.00, 4.95,
              5.50, 5.50, 5.40,
              5.45, 5.55, 5.50,
              5.55, 5.55, 5.35,
              5.45, 5.50, 5.55,
              5.50, 5.45, 5.25,
              5.65, 5.60, 5.40,
              5.70, 5.65, 5.55,
              6.30, 6.30, 6.25),
   methods = factor(rep(1:3, 22),
                     labels = c("Round Out", "Narrow Angle", "Wide Angle")),
   block = gl(22, 3)
)
```

```
## Asymptotic Friedman test
friedman_test(times ~ methods | block, data = rounding)
## Parallel coordinates plot
with(rounding, {
    matplot(t(matrix(times, ncol = 3, byrow = TRUE)),
            type = "1", lty = 1, col = 1, ylab = "Time", xlim = c(0.5, 3.5),
            axes = FALSE)
    axis(1, at = 1:3, labels = levels(methods))
    axis(2)
})
## Where do the differences come from?
## Wilcoxon-Nemenyi-McDonald-Thompson test (Hollander and Wolfe, 1999, p. 295)
## Note: all pairwise comparisons
(st <- symmetry_test(times ~ methods | block, data = rounding,</pre>
                     ytrafo = function(data)
                          trafo(data, numeric_trafo = rank_trafo,
                                block = rounding$block),
                     xtrafo = mcp_trafo(methods = "Tukey")))
## Simultaneous test of all pairwise comparisons
## Wide Angle vs. Round Out differ (Hollander and Wolfe, 1999, p. 296)
pvalue(st, method = "single-step") # subset pivotality is violated
## Strength Index of Cotton
## Hollander and Wolfe (1999, p. 286, Tab. 7.5)
cotton <- data.frame(</pre>
    strength = c(7.46, 7.17, 7.76, 8.14, 7.63,
                 7.68, 7.57, 7.73, 8.15, 8.00,
                 7.21, 7.80, 7.74, 7.87, 7.93),
    potash = ordered(rep(c(144, 108, 72, 54, 36), 3),
                     levels = c(144, 108, 72, 54, 36)),
    block = gl(3, 5)
)
## One-sided asymptotic Page test
friedman_test(strength ~ potash | block, data = cotton, alternative = "greater")
## One-sided approximative (Monte Carlo) Page test
friedman_test(strength ~ potash | block, data = cotton, alternative = "greater",
              distribution = approximate(B = 10000))
## Data from Quade (1979, p. 683)
dta <- data.frame(</pre>
    y = c(52, 45, 38,
          63, 79, 50,
          45, 57, 39,
          53, 51, 43,
          47, 50, 56,
          62, 72, 49,
```

```
49, 52, 40),
     x = factor(rep(LETTERS[1:3], 7)),
     b = factor(rep(1:7, each = 3))
)
## Approximative (Monte Carlo) Friedman test
## Quade (1979, p. 683)
friedman_test(y \sim x | b, data = dta,
              distribution = approximate(B = 10000)) # chi^2 = 6.000
## Approximative (Monte Carlo) Quade test
## Quade (1979, p. 683)
(qt \leftarrow quade_test(y \sim x \mid b, data = dta,
                   distribution = approximate(B = 10000))) # W = 8.157
## Comparison with R's quade.test() function
quade.test(y \sim x \mid b, data = dta)
## quade.test() uses an F-statistic
b <- nlevels(qt@statistic@block)</pre>
A <- sum(qt@statistic@y^2)
B <- sum(statistic(qt, type = "linear")^2) / b
(b - 1) * B / (A - B) # F = 8.3765
```

Transformations

Functions for Data Transformation

Description

Transformations for factors and numeric variables.

Usage

Arguments

x	an object of class "numeric", "factor", "ordered" or "Surv".
ties.method	a character, the method used to handle ties. The score generating function either uses the mid-ranks ("mid-ranks", default) or, in the case of rank_trafo, randomly broken ties ("random"). Alternatively, the average of the scores resulting from applying the score generating function to randomly broken ties are used ("average-scores"). See logrank_test for a detailed description of the methods used in logrank_trafo.
mid.score	a character, the score assigned to observations exactly equal to the median: either 0 ("0", default), 0.5 ("0.5") or 1 ("1"); see median_test.
а	a numeric vector, the values taken as the constant \boldsymbol{a} in the Conover-Salsburg scores. Defaults to 5.
j	a numeric, the value taken as the constant j in the Koziol-Nemec scores. Defaults to 1.
weight	a function where the first three arguments must correspond to time, n.risk, and n.event given below. Defaults to logrank_weight.
time	a numeric vector, the ordered distinct time points.
n.risk	a numeric vector, the number of subjects at risk at each time point specified in time.
n.event	a numeric vector, the number of events at each time point specified in time.
type	a character, one of "logrank" (default), "Gehan-Breslow", "Tarone-Ware", "Prentice", "Prentice-Marek", "Andersen-Borgan-Gill-Keiding", "Fleming-Harrington" or "Self"; see logrank_test.
rho	a numeric vector, the ρ constant when type is "Tarone-Ware", "Fleming-Harrington" or "Self"; see logrank_test. Defaults to NULL, implying 0.5 for type = "Tarone-Ware" and 0 otherwise.
gamma	a numeric vector, the γ constant when type is "Fleming-Harrington" or "Self"; see logrank_test. Defaults to NULL, implying 0.
scores	a numeric vector or list, the scores corresponding to each level of an ordered factor. Defaults to NULL, implying 1:nlevels(x).
minprob	a numeric, a fraction between 0 and 0.5; see maxstat_test. Defaults to 0.1.

a numeric a fraction between 0.5 and 1, see may stat, test. Defaults to 1 - minnrah
a numeric, a fraction between 0.5 and 1; see maxstat_test. Defaults to 1 - minprob.
an object of class "data.frame".
a function to be applied to elements of class "numeric" in data, returning a matrix with nrow(data) rows and an arbitrary number of columns. Defaults to id_trafo.
a function to be applied to elements of class "factor" in data, returning a matrix with nrow(data) rows and an arbitrary number of columns. Defaults to f_trafo.
a function to be applied to elements of class "ordered" in data, returning a matrix with nrow(data) rows and an arbitrary number of columns. Defaults to of_trafo.
a function to be applied to elements of class "Surv" in data, returning a matrix with nrow(data) rows and an arbitrary number of columns. Defaults to logrank_trafo.
an optional named list of functions to be applied to the corresponding variables in data. Defaults to NULL.
an optional factor whose levels are interpreted as blocks. trafo is applied to each level of block separately. Defaults to NULL.
logrank_trafo: further arguments to be passed to weight. mcp_trafo: factor name and contrast matrix (as matrix or character) in a 'tag = value' format for multiple comparisons based on a single unordered factor; see mcp in package multcomp.

Details

The utility functions documented here are used to define specialized test procedures.

id_trafo is the identity transformation.

rank_trafo, normal_trafo, median_trafo, savage_trafo, consal_trafo and koziol_trafo compute rank scores, normal scores, median scores, Savage scores, Conover-Salsburg scores (see neuropathy) and Koziol-Nemec scores, respectively, for location problems.

klotz_trafo, mood_trafo, ansari_trafo and fligner_trafo compute Klotz scores, Mood scores, Ansari-Bradley scores and Fligner-Killeen scores, respectively, for scale problems.

logrank_trafo computes weighted logrank scores for right-censored data, allowing for a user-defined weight function through the weight argument (see GTSG).

f_trafo computes dummy matrices for factors and of_trafo assigns scores to ordered factors. For ordered factors with two levels, the scores are normalized to the [0, 1] range.

maxstat_trafo, fmaxstat_trafo and ofmaxstat_trafo compute scores for cutpoint problems (see maxstat_test).

trafo applies its arguments to the elements of data according to the classes of the elements. A trafo function with modified default arguments is usually supplied to independence_test via the xtrafo or ytrafo arguments. Fine tuning, i.e., different transformations for different variables, is possible by supplying a named list of functions to the var_trafo argument.

mcp_trafo computes contrast matrices for factors.

Value

A numeric vector or matrix with nrow(x) rows and an arbitrary number of columns. For trafo, a named matrix with nrow(data) rows and an arbitrary number of columns.

Note

Starting with **coin** version 1.1-0, all transformation functions are now passing through missing values (i.e., NAs). Furthermore, median_trafo and logrank_trafo are now *increasing* functions (in conformity with most other transformations in this package).

Examples

```
## Dummy matrix, two-sample problem (only one column)
f_{trafo(gl(2, 3))}
## Dummy matrix, K-sample problem (K columns)
x < -gl(3, 2)
f_trafo(x)
## Score matrix
ox <- as.ordered(x)</pre>
of_trafo(ox)
of_trafo(ox, scores = c(1, 3:4))
of_trafo(ox, scores = list(s1 = 1:3, s2 = c(1, 3:4)))
## Normal scores
y \leftarrow runif(6)
normal_trafo(y)
## All together now
trafo(data.frame(x = x, ox = ox, y = y), numeric\_trafo = normal\_trafo)
## The same, but allows for fine-tuning
trafo(data.frame(x = x, ox = ox, y = y), var_trafo = list(y = normal_trafo))
## Transformations for maximally selected statistics
maxstat_trafo(y)
fmaxstat_trafo(x)
ofmaxstat_trafo(ox)
## Apply transformation blockwise (as in the Friedman test)
trafo(data.frame(y = 1:20), numeric_trafo = rank_trafo, block = gl(4, 5))
## Multiple comparisons
dta <- data.frame(x)</pre>
mcp_trafo(x = "Tukey")(dta)
## The same, but useful when specific contrasts are desired
K \leftarrow rbind("2 - 1" = c(-1, 1, 0),
           "3 - 1" = c(-1, 0, 1),
           "3 - 2" = c(0, -1, 1)
mcp_trafo(x = K)(dta)
```

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treepipit

Tree Pipits in Franconian Oak Forests

Description

Data on the population density of tree pipits, *Anthus trivialis*, in Franconian oak forests including variables describing the forest ecosystem.

Usage

treepipit

Format

A data frame with 86 observations on 10 variables.

counts the number of tree pipits observed.

age age of the overstorey oaks taken from forest data.

coverstorey cover of canopy overstorey (%). The crown cover is described relative to a fully stocked stand. Very dense overstorey with multiple crown cover could reach values greater than 100 %.

coverregen cover of regeneration and shrubs (%).

meanregen mean height of regeneration and shrubs.

coniferous coniferous trees (% per hectare).

deadtree number of dead trees (per hectare).

cbpiles number of crowns and branch piles (per hectare). All laying crowns and branch piles were counted. These were induced by logging and the creation of wind breaks.

ivytree number of ivied trees (per hectare).

fdist distance to the forest edge. The closest distance to the forest edge was measured from the centre of each grid.

Details

This study is based on fieldwork conducted in three lowland oak forests in the Franconian region of northern Bavaria close to Uffenheim, Germany. Diurnal breeding birds were sampled five times, from March to June 2002, using a quantitative grid mapping. Each grid was a one-hectare square. In total, 86 sample sites were established in 9 stands. All individuals were counted in time intervals of 7 min/grid during slow walks along the middle of the grid with a stop in the centre. Environmental factors were measured for each grid.

References

Müller, J. and Hothorn, T. (2004). Maximally selected two-sample statistics as a new tool for the identification and assessment of habitat factors with an application to breeding bird communities in oak forests. *European Journal of Forest Research* **123**(3), 219–228.

VarCovar-class 91

Examples

VarCovar-class

Class "VarCovar" and its subclasses

Description

Objects of class "VarCovar" and its subclasses "CovarianceMatrix" and "Variance" represent the covariance and variance, respectively, of the linear statistic.

Objects from the Class

Class "VarCovar" is a *virtual* class defined as the class union of "CovarianceMatrix" and "Variance", so objects cannot be created from it directly.

Objects can be created by calls of the form

```
new("CovarianceMatrix", covariance, ...)
and
new("Variance", variance, ...).
```

where covariance is a covariance matrix and variance is numeric vector containing the diagonal elements of the covariance matrix.

Slots

```
For objects of class "CovarianceMatrix":

covariance: Object of class "matrix". The covariance matrix.

For objects of class "Variance":

variance: Object of class "numeric". The diagonal elements of the covariance matrix.
```

Extends

```
For objects of classes "CovarianceMatrix" or "Variance": Class "VarCovar", directly.
```

Known Subclasses

```
For objects of class "VarCovar":
Class "CovarianceMatrix", directly.
Class "Variance", directly.
```

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Methods

covariance signature(object = "CovarianceMatrix"): See the documentation for covariance
 for details.

initialize signature(.Object = "CovarianceMatrix"): See the documentation for initialize
 (in package methods) for details.

initialize signature(.Object = "Variance"): See the documentation for initialize (in package methods) for details.

variance signature(object = "CovarianceMatrix"): See the documentation for variance
 for details.

variance signature(object = "Variance"): See the documentation for variance for details.

vision

Unaided Distance Vision

Description

Assessment of unaided distance vision of women in Britain.

Usage

vision

Format

A contingency table with 7477 observations on 2 variables.

Right.Eye a factor with levels "Highest Grade", "Second Grade", "Third Grade" and "Lowest Grade". Left.Eye a factor with levels "Highest Grade", "Second Grade", "Third Grade" and "Lowest Grade".

Details

Paired ordered categorical data from case-records of eye-testing of 7477 women aged 30–39 years employed by Royal Ordnance Factories in Britain during 1943–46, as given by Stuart (1953).

Stuart (1955) used this data set to illustrate a test of marginal homogeneity.

Source

Stuart, A. (1953). The estimation and comparison of strengths of association in contingency tables. *Biometrika* **40**(1/2), 105–110.

References

Stuart, A. (1955). A test for homogeneity of the marginal distribution in a two-way classification. *Biometrika* **42**(3/4), 412–416.

vision 93

Examples

```
## Asymptotic Stuart(-Maxwell) test (Q = 11.96)
diag(vision) <- 0 # speed-up
mh_test(vision)</pre>
```

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