6.438 Final Exam

Collaborators: rdeits

Gaussian Random Variables

$$\begin{split} &\underline{x} \sim N(\underline{\mu}, \Lambda) \\ &E[\underline{x}] = \underline{\mu} \\ &E[(\underline{x} - \underline{\mu})(\underline{x} - \underline{\mu})^T] = \Lambda \\ &p_x(\underline{x}) = |2\pi\Lambda|^{-\frac{1}{2}} \exp\left[-\frac{1}{2}[(\underline{x} - \mu)^T\Lambda^{-1}[(\underline{x} - \mu)]\right] \end{split}$$

Transformations

$$\begin{split} \underline{z} \sim N(\underline{\mu}_z, \Lambda_z) &\quad \underline{x} = A\underline{z} + \underline{b} \rightarrow \underline{x} \sim N(\underline{\mu}_x, \Lambda_x) \\ \underline{\mu}_x = A\underline{\mu}_z + \underline{b} &\quad \Lambda_x = A\Lambda_z A^T \end{split}$$

Gaussian Information Form

$$\begin{aligned} p(x) &= \tfrac{1}{Z} \exp\left[-\tfrac{1}{2} x^T J x + h^T x\right] \\ J &= \Lambda^{-1} \quad h = J \mu \end{aligned}$$

Marginalization

$$p(x_1) \sim N(\mu_1, \Lambda_{11})$$
 $p(x_1) \sim N(h_1\prime, J_{11}\prime)$
 $h'_1 = h_1 - J_{12}J'_{22}h_2$ $J'_{11} = \Lambda_{11}^{-1} = J_{11} - J_{12}J_{22}^{-1}J_{21}$ (Schur complement)

Conditioning

$$\begin{split} p(x_1|x_2) &\sim N^{-1}(h_1'', J_{11}'') \\ J_{11}'' &= J_{11} \quad h_1'' = h_1 - J_{12}x_2 \\ \mu_1'' &= \mu_1 + \Lambda_{12}\Lambda_{22}^{-1}(x_2 - \mu_2) \quad \Lambda_{11}'' = \Lambda_{11} - \Lambda_{12}\Lambda_{22}^{-1}\Lambda_{21} \end{split}$$

Dependencies

G is an I-map of P if $CI(G) \subseteq CI(P)$ e.g. G is fully connected G is a D-map of P if $CI(G) \supseteq CI(P)$ e.g. G is unconnected G is a P-map of P if CI(G) = CI(P)

G is a minimal I-map if removing any edge would make it no longer an I-map

A directed model turned into a moralized undirected model is a P-map if moralization adds no edges.

Undirected graph G has a directed P-map iff G is chordal Directed graph H has an undirected P-map iff moralization adds no edges

Variable Elimination

$$m_i(x_{s_i}) = \sum_{x_i} \prod_{\varphi_i \in \Psi} \varphi_i(x_i, x_{s_i})$$

Sum-Product

$$\begin{split} p(x) &= \prod_i \phi_i(x_i) \prod_{(i,j) \in E} \psi_{ij}(x_i, x_j) \\ m_{i \to j}^{t+1}(x_j) &= \sum_{x_i} \phi_i(x_i) \psi_{ij}(x_i, x_j) \prod_{k \in N(i) \setminus \{j\}} m_{k \to i}^t(x_i) \\ p_{x_i}(x_i) &\propto \phi_i(x_i) \prod_{k \in N(i)} m_{k \to i}(x_i) \end{split}$$

Forward-Backward Probabilities

 $\begin{aligned} & \text{Markov chain with nodes } (x_1, \dots, x_N, \hat{y}_1, \dots, \hat{y}_N) \\ & \underbrace{p(y_{i+1}|x_{i+1})m_{i \to i+1}(x_{i+1})}_{\alpha_{i+1}(x_{i+1})} = \sum_{x_i} p(x_{i+1}|x_i) \underbrace{m_{y_i \to x_i}(x_i)m_{i-1 \to i}(x_i)}_{\alpha_i(x_i)} \\ & \underbrace{m_{i+1 \to i}(x_i)}_{\beta_i(x_i)} = \sum_{x_{i+1}} p(x_{i+1}|x_i)p(\hat{y}_{i+1}|x_{i+1}) \underbrace{m_{i+2 \to i+1}(x_{i+1})}_{\beta_{i+1}(x_{i+1})} \\ & \underbrace{p(x_i|\hat{y}_1, \dots, \hat{y}_N)}_{x_i'} = \sum_{x_i'} \underbrace{\alpha_i(x_i)\beta_i(x_i)}_{x_i'} \end{aligned}$

Sum-Product for Factor Tree

Factor \rightarrow node:

$$m_{a \to j}(x_j) = \sum_{\substack{x_k, k \in N(a) \setminus \{j\}}} f_a(x_{N(a)}) \prod_{k \in N(a) \setminus \{j\}} m_{k \to a}(x_k)$$

Node \rightarrow factor:

$$m_{j\to a}(x_j) = \prod_{b\in N(j)\setminus\{a\}} m_{b\to j}(x_j)$$

Kalman Filtering

$$\begin{split} x_{t+1} &= Ax_t + v_t, v \sim N(0, Q), x_0 \sim N(0, \Lambda_0) \\ y_t &= Cx_t + w_t, w_t \sim N(0, R) \\ x_0 &\sim N(0, \Lambda_0) \\ x_{t+1} | x_t \sim N(Ax_t, Q) \\ y_t | x_t \sim N(Cx_t, R) \end{split}$$

Filtering

$$\alpha(x_{i+1}) = \int \alpha(x_i) p(x_{i+1}|x_i) p(y_{i+1}|x_{i+1}) dx_i$$

Prediction

$$\begin{split} \mu_{i+1|i} &= A \mu_{i|i} \\ \Sigma_{i+1|i} &= A \Sigma_{i|i} A^T + Q \\ \mu_{0|-1} &= 0 \\ \Sigma_{0|-1} &= \Lambda_0 \end{split}$$

Update

$$\mu_{i+1|i+1} = \mu_{i+1|i} + G_{i+1}(y_{i+1} - C\mu_{i+1|i})$$

$$\Sigma_{i+1|i+1} = \Sigma_{i+1|i} - G_{i+1}C\Sigma_{i+1|i}$$

$$G_{i+1} = \Sigma_{i+1|i}C^{T}(C\Sigma_{i+1|i}C^{T} + R)^{-1}$$

Smoothing

$$\begin{split} &\gamma(x_i) = \int \gamma(x_{i+1}) \left[\frac{\alpha(x_i)p(x_{i+1}|x_i)}{\int \alpha(x_i')p(x_{i+1}|x_i')dx_i'} \right] dx_{i+1} \\ &\gamma(x_i) = \frac{\alpha(x_i)\beta(x_i)}{p(y_0^t)} \\ &\mu_{i|t} = \mu_{i|i} + F_i(\mu_{i+1|t} - \mu_{i+1|i}) \\ &\Sigma_{i|t} = F_i(\Sigma_{i+1|t} - \Sigma_{i+1|i})F_i^T + \Sigma_{i|i} \\ &F_i = \Sigma_{i|i}A^T \Sigma_{i+1|i}^{-1} \end{split}$$

Junction Trees

If a graph is chordal, then it has a junction tree.

Loopy BP

If we have a graph
$$\mathcal{G} = (V, E)$$
 with $p_x(x) \propto \prod_{i \in V} \exp(\phi_i(x_i)) \prod_{(i,j) \in E} \exp(\psi_{ij}(x_i, x_j))$ $m_{i \to j}^{t+1}(x_j) \propto \sum_{x_i \in \mathcal{X}} \exp(\phi_i(x_i)) \exp(\psi_{ij}(x_i, x_j)) \prod_{k \in N(i) \setminus j} m_{k \to i}^t(x_i)$ Node and edge marginals: $b_i^t(x_i) \propto \exp(\phi_i(x_i)) \prod_{k \in N(i)} m_{k \to i}^t(x_i)$ $b_{ij}^t(x_i, x_j) \propto \exp(\phi_i(x_i) + \phi_j(x_j) + \psi_{ij}(x_i, x_j)) \prod_{k \in N(i)} m_{k \to i}^t(x_i) \prod_{\ell \in N(i)} m_{\ell \to j}^t(x_j)$

Variational Methods

K-L Divergence

$$\mathrm{KL}(p||q) = \sum_{x} p(x) \log \frac{p(x)}{q(x)}$$

Bethe approximation

Constraints

$$\mu(x) = \prod_{i \in V} \mu_i(x_i) \prod_{(i,j) \in E} \frac{\mu_{ij}(x_i, x_j)}{\mu_i(x_i)\mu_j(x_j)}$$

$$\mu_i(x_i) \ge 0$$

$$\sum_{x_i \in \mathcal{X}} \mu_i(x_i) = 1$$

$$\mu_{ij}(x_i, x_j) \ge 0$$

$$\sum_{x_j \in \mathcal{X}} \mu_{ij}(x_i, x_j) = \mu_i(x_i)$$

$$\sum_{x_i \in \mathcal{X}} \mu_{ij}(x_i, x_j) = \mu_j(x_j)$$

Mean Field

$$\mu(x) = \prod_{i \in V} \mu_i(x_i)$$

$$\mu_i^{t+1}(x_i) \propto \exp\left[\phi_i(x_i) + \sum_{j \in N(i)} \sum_{x_j \in \mathcal{X}} \mu_j^t(x_j) \psi(x_i, x_j)\right]$$

Variational Objective

Maximize:
$$\mathcal{F}(\mu) = \sum_{x \in \mathcal{X}^N} \mu(x) \theta(x) - \sum_{x \in \mathcal{X}^N} \mu(x) \log \mu(x)$$
 where $P(x) = \frac{1}{Z(\theta)} e^{\theta(x)}$

Sampling

Markov Chain Monte Carlo

Metropolis-Hastings

Require a reversible Markov chain:

$$P(x)P(x \to x') = P(x' \to x)P(x')$$

and regular:
$$\exists n \text{ s.t. } P(X(n) = x \mid X(0) = x) > 0$$

Proposal distribution: $K(x \to x')$

Prob. of accepting a move from $x \to x'$:

$$A(x \to x\prime) = \min\left[1, \frac{K(x\prime \to x)P(x\prime)}{K(x \to x\prime)P(x)}\right]$$

$$P(x \to x\prime) = K(x \to x\prime) A(x \to x\prime)$$

$$P(x \to x) = 1 - \sum_{x \neq x} K(x \to x \prime) A(x \to x \prime)$$

Gibbs Sampling

Subclass of M-H with $A(x \to x') = 1$

- 0) select any x
- 1) pick k at random
- 2) Sample $x_{k'} \sim P(x_k|x_{-k}) = P(x_k|x_{N(k)})$

Importance Sampling

Instead of P, sample from q using weighting $w^k = \frac{P(x^k, y)}{q(x^k)}$

$$\frac{\sum_{k} w^{k} f(x^{k})}{\sum_{k} w^{k}} = E_{x \sim P_{x|y}} f(x)$$

Particle Filtering

samples
$$= k \in 1, \dots, K$$

$$x_0^k \sim p_{x_0}(\cdot)$$

$$w_0^k = \frac{1}{K} p_{y_0|x_0}(y_0|x_0)$$

$$x_{n+1}^k \sim p_{x_{n+1}|x}(\cdot|x_n^k)$$

$$w_{n+1}^k = w_n^k \times p_{u+1}|_{x_{n+1}}(y_{n+1}|x_{n+1}^k)$$

Bayesian Estimation

Treat θ as a random variable

Dirichlet prior: $P(\theta) = \frac{1}{Z} \prod_{x} \theta_{x}^{\alpha_{x}-1}$

$$Z = \frac{\prod_{x} \Gamma(\alpha_x)}{\Gamma(\sum_{x} \alpha_x)}$$

Inferring Structure

Bayesian Information Criterion

$$\ell(\hat{\theta}^{ML}; D) - \frac{\text{num. params}}{2} \log n$$

$$\frac{e(\theta^{n-1};D) - \frac{1}{2} - \log n}{\text{approximates } \log P(D;G) = \log \int P(\theta)L(\theta;D)d\theta}$$

$$score(G) = \sum_{i=1}^{N} score(i|pa_i; D) = \log \prod_{j=1}^{q_i} \frac{\Gamma(\alpha_{ij})}{\Gamma(\alpha_{ijk})} \prod_{k=1}^{r_i} \frac{\Gamma(\alpha_{ijk} + n_{ijk})}{\Gamma(\alpha_{ijk})}$$
for BN

Learning Models

Maximum Likelihood estimation

$$\hat{\theta}^{ML} = \arg\max_{\theta} L(\theta; x)$$

$$I(u;v) \triangleq \sum_{u,v} p_{u,v}(u,v) \log \frac{p_{u,v}(u,v)}{p_u(u)p_v(v)}$$

$$H(u) \triangleq -\sum_{u} p_u(u) \log p_u(u) \ge 0$$

$$\hat{\ell}(G, D) = \sum_{i=1}^{N} \hat{I}(x_i; x_{\pi_i}) - \sum_{i=1}^{N} \hat{H}(x_i)$$

Expectation Maximization Algorithm

$$y = (y_1, \dots, y_N)$$
 observed

$$x = (x_1, \dots, x_N)$$
 latent

assume we know $p_{y,x}(\cdot,\cdot;\theta)$ w/ param θ , want $\hat{\theta}^{ML}$

 $\ell(\theta; y)$ incomplete log likelihood

 $\ell(\theta; y, x)$ complete log likelihood

Choose distribution q over x: $q(\cdot|y)$

$$\ell(\theta; y) \ge \sum_{x} q(x|y) \log \frac{P_{y,x}(y,x;\theta)}{q(x|y)} \stackrel{\triangle}{=} \tilde{\ell}(q,\theta)$$
, maximize $\tilde{\ell}(q,\theta)$

E-step:
$$q^{(i+1)} = \arg\max_{q} \tilde{\ell}(q, \theta^{(i)})$$

M-step:
$$\theta^{(i+1)} = \arg \max_{\theta} \tilde{\ell}(q^{(i+1)}, \theta)$$

Solving those steps gives:

E-step:
$$q^{(i+1)} = p_{x|y}(\cdot|y;\theta^{(i)})$$

M-step:
$$\theta^{(i+1)} = \arg \max_{\theta} \mathbb{E} \left[\log p_{y,x}(y,x;\theta) | \mathbf{y} = y; \theta^{(i)} \right]$$

Estimating Undirected Models

$$\ell(\theta; D) = \frac{1}{n} \sum_{t=1}^{n} \log P(x^{t}; \theta)$$

$$\hat{p}_c(x_c) = \frac{1}{n} \sum_{t=1}^{n} \mathbb{1}(x_c, x_c^t)$$

$$p_c(x_c; \theta) = \frac{\partial}{\partial \theta_c(x_c)} \log Z(\theta)$$

Iterative Proportionality Fitting

For each clique c, eval $P_c^{(i)}(x_c)$

$$P^{(i+1)}(x_1, \dots, x_N) = P^{(i)}(x_1, \dots, x_N) \frac{\hat{p}_c(x_c)}{p_c^{(i)}(x_c)}$$

Closed-form solution for a chordal graph:

$$p(x) = \frac{\prod_c \psi_c(x_c)}{\prod_s \phi_s(x_s)} = \frac{\prod_c \hat{p}_c(x_c)}{\prod_s \hat{p}_s(x_s)}$$
 with cliques c and separators s