

# Introduction to Bayesian Statistics for Demographic Science

Day 2: Generalised linear models (GLMs) &  
hierarchical models

# Outline

- General linear model (refresher)
- Generalised linear models (GLMs)
- Mark-recapture
- Hierarchical models
- State-space

# Components of a model

1. Data model(s)

$$y_i \sim \text{Normal}(\mu_i, \sigma)$$

2. Process model(s)

$$\mu_i = \alpha + \beta x_i$$

3. Parameter model(s)

$$\alpha \sim \text{Normal}(0, 1)$$

$$\beta \sim \text{Normal}(0, 1)$$

$$\sigma \sim \text{HalfCauchy}(0, 1)$$

# Diagnostics

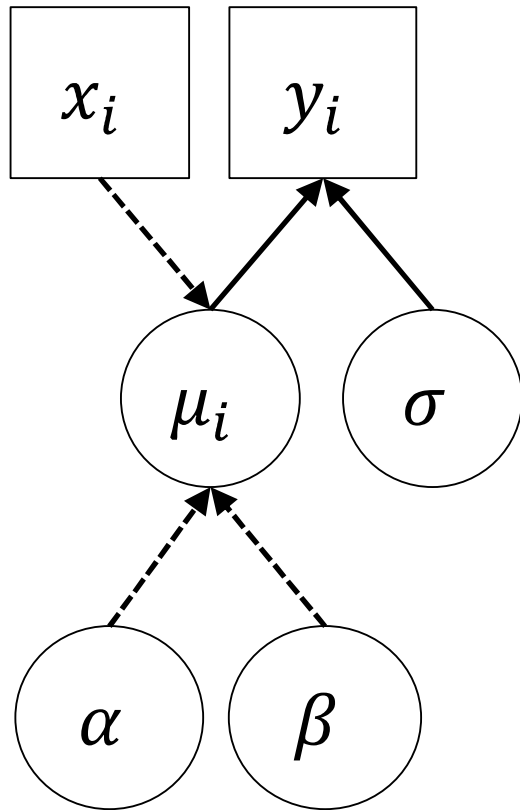
- MCMC Diagnostics
  - Trace plots
  - Rhat
- Model fit
  - Out-of-sample cross validation
  - Posterior predictive checks
  - (your normal way of assessing fit for any other statistical model)

# Linear regression

## General linear model

# General linear model

Refresher



$$y_i \sim \text{Normal}(\mu_i, \sigma)$$

$$\mu_i = \alpha + \beta x_i$$

$$\alpha \sim \text{Normal}(0, 1)$$

$$\beta \sim \text{Normal}(0, 1)$$

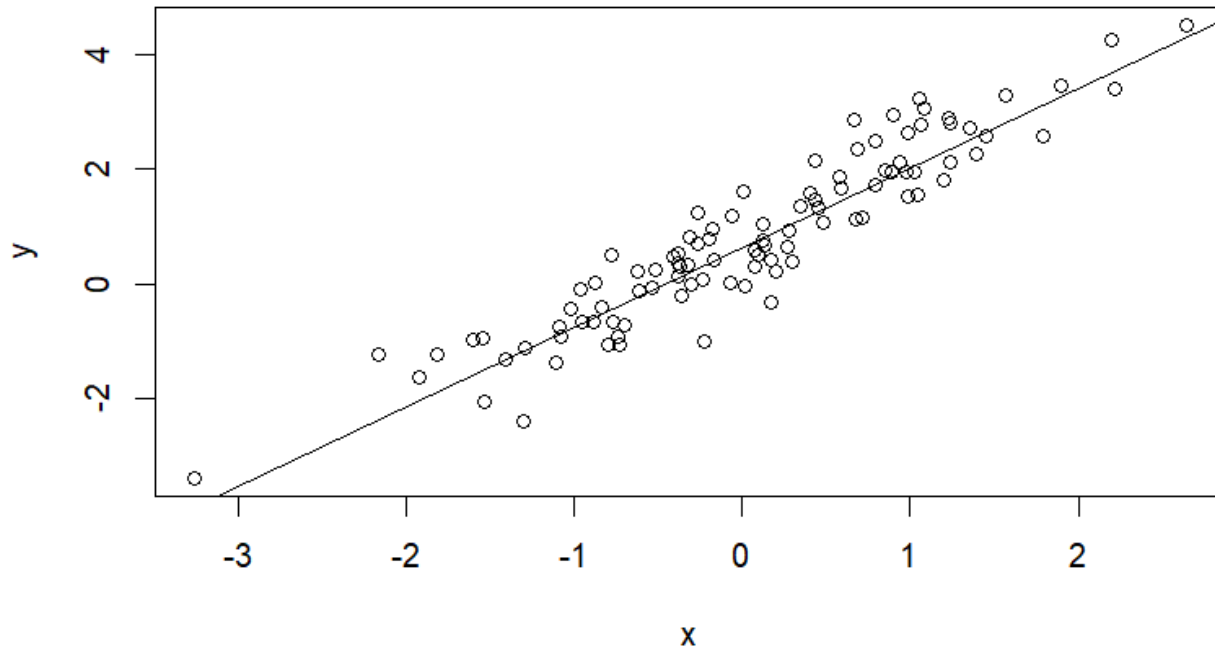
$$\sigma \sim \text{HalfCauchy}(0, 1)$$

# General linear model

## Refresher

What are examples of quantities that  $y_i$  could represent?

Examples that it cannot be?



$$y_i \sim \text{Normal}(\mu_i, \sigma)$$

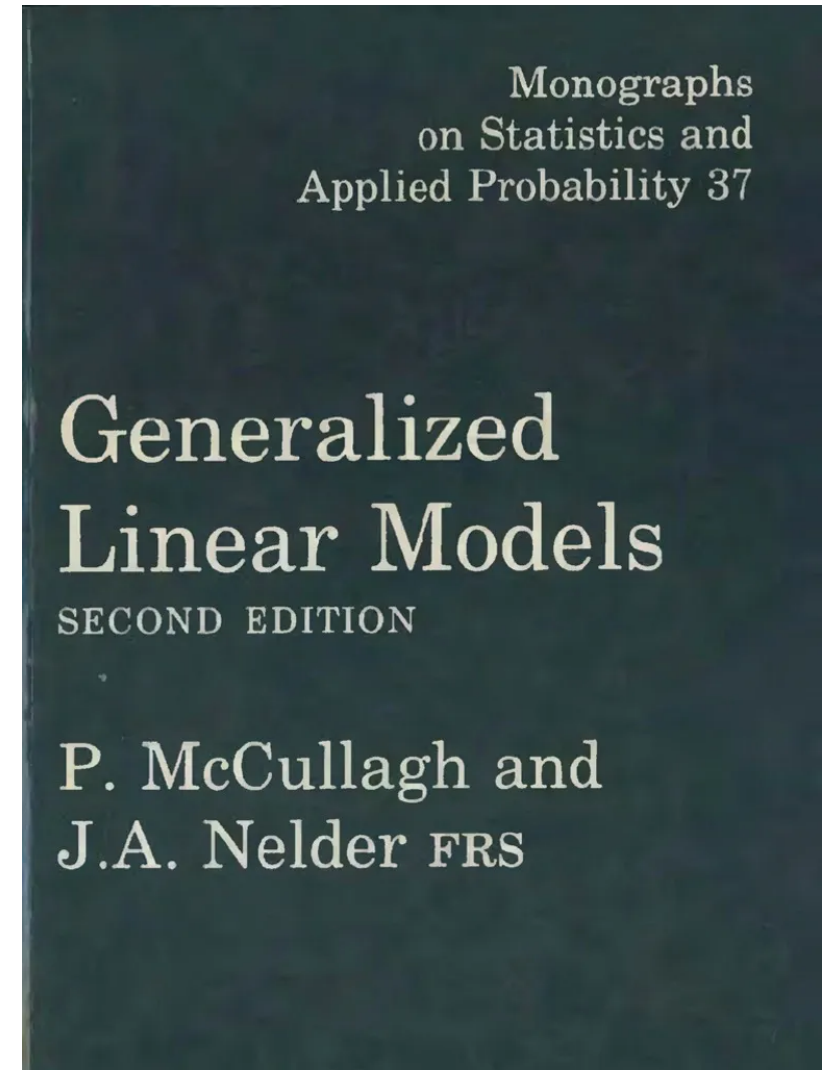
$$\mu_i = \alpha + \beta x_i$$

$$\alpha \sim \text{Normal}(0, 1)$$

$$\beta \sim \text{Normal}(0, 1)$$

$$\sigma \sim \text{HalfCauchy}(0, 1)$$

# Generalised linear models





# Generalised linear models

- Generalisation of a linear model to accommodate additional data types
- Three components:
  - Probability distribution
  - A linear predictor
  - Link function

# Can you sketch a model?

We have **counts of people**  $y_i$  from a sample of locations  $i$  across a country.

We want to infer the relationship of population size with the **unemployment rate**  $x_{1,i}$ , **cost of living**  $x_{2,i}$ , and **number of pubs**  $x_{3,i}$ .

# Generalised linear model

Poisson regression for count data

$$y_i \sim \text{Poisson}(\lambda_i)$$

$$\log(\lambda_i) = \alpha + \sum_{k=1}^K \beta_k x_{k,i}$$

$$\alpha \sim \text{Normal}(0, 15)$$

$$\beta \sim \text{Normal}(0, 10)$$

# Can we build a model for these data?

We have **counts of active Facebook users**  $y_i$  from one hundred countries  $i$ . We also know the **total population**  $N_i$  of each country.

We want to study how country characteristics influence the probability that a person uses Facebook.

We have data on **gross domestic product**  $x_{1,i}$ , **broadband internet speeds**  $x_{2,i}$ , and **rates of attainment for secondary education**  $x_{3,i}$ .

# Generalised linear model

Binomial regression for count data

$$y_i \sim \text{Binomial}(N_i, \theta_i)$$

$$\text{logit}(\theta_i) = \alpha + \sum_{k=1}^K \beta_k x_{k,i}$$

$$\alpha \sim \text{Normal}(0, 1e3)$$

$$\beta \sim \text{Normal}(0, 1e3)$$

# How about this one?

We have GitHub data from individual **Russian programmers**  $i$  **indicating if they have migrated (or not)** to a new country since 2022.

We want to study how individual characteristics influence their probability of migration.

From their profile, we have their **number of followers**  $x_{1,i}$ , **number of contributions** in the past year  $x_{2,i}$ , and **number of popular code repositories**  $x_{3,i}$ .

# Generalised linear model

Logistic regression for binary outcomes

$$y_i \sim \text{Bernoulli}(\theta)$$

$$\text{logit}(\theta) = \alpha + \sum_{k=1}^K \beta_k x_{k,i}$$

$$\alpha \sim \text{Normal}(0, 1e3)$$

$$\beta \sim \text{Normal}(0, 1e3)$$

Some other useful models beyond GLMs...



# What do you think about this one?

We have data on the **proportion women** among employees of the top tech companies  $i$  globally.

We want to study how these ratios are influenced by company characteristics like **female CEO**  $x_{1,i}$ , **gender pay gap**  $x_{2,i}$ , and **days paid maternity leave**  $x_{3,i}$ .

# Beta regression

$$y_i \sim \text{Beta}(\mu\tau, \tau - \mu_i\tau)$$

$$\text{logit}(\mu_i) = \alpha + \sum_{k=1}^K \beta_k x_{k,i}$$

$$\alpha \sim \text{Normal}(0, 1e3)$$

$$\beta \sim \text{Normal}(0, 1e3)$$

$$\tau \sim \text{Uniform}(0, 1e3)$$

Last one...

We have data on **total fertility rates for every major city  $i$**  in the world.

We want to study how these rates are influenced by **population density  $x_{1,i}$ , hospitals per capita  $x_{2,i}$ , and 4G coverage  $x_{3,i}$** .

# Log-linear regression

$$y_i \sim \text{LogNormal}(\mu_i, \sigma)$$

$$\mu_i = \alpha + \sum_{k=1}^K \beta_k x_{k,i}$$

$$\alpha \sim \text{Normal}(0, 1e3)$$

$$\beta \sim \text{Normal}(0, 1e3)$$

$$\sigma \sim \text{HalfNormal}(0, 1)$$

What if we also wanted to include a **country effect**?  
How would you do that?

# Hierarchical Models

# Hierarchical models

- Multiple levels (i.e. sub-models) within the same model
- The outcome of one level is an input to another level

## **Why do we want to do this?**

- Sharing information among groups to increase statistical power
- Processes occurring at multiple scales (i.e. spatial, temporal, etc)
- Partition uncertainty into process error and observation error
- Different deterministic functions (i.e. regressions) for different levels of the model

# Random effects

- Effects (e.g. intercept, slope) are dependent on another factor
- Usually a grouping factor ( $g$ )
- The effects are not independent among groups
- Prior model and hyper-priors

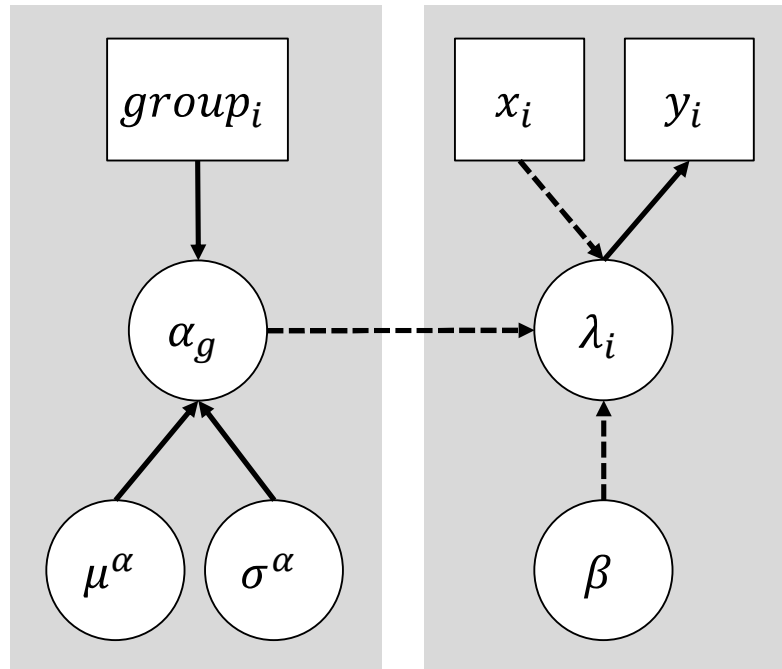
$$y_i \sim \text{Poisson}(\lambda_i)$$

$$\log(\lambda_i) = \alpha_g + \beta x_i$$

$$\alpha_g \sim \text{Normal}(\mu^\alpha, \sigma^\alpha)$$

$$\begin{aligned}\mu^\alpha &\sim \text{Normal}(0, 1e3) \\ \sigma^\alpha &\sim \text{Cauchy}(0, 1e3)\end{aligned}$$

# Random effects



$$y_i \sim \text{Poisson}(\lambda_i)$$

$$\log(\lambda_i) = \alpha_g + \beta x_i$$

$$\alpha_g \sim \text{Normal}(\mu^\alpha, \sigma^\alpha)$$

$$\mu^\alpha \sim \text{Normal}(0, 1e3)$$

$$\sigma^\alpha \sim \text{Cauchy}(0, 1e3)$$



# Mark-recapture

- A method originally designed for estimating unobservable animal populations
- Useful in demography for estimating populations that are impossible to census
- **Lincoln-Peterson mark recapture** is the oldest and simplest approach (assuming a closed population)
- **Cormack-Jolly-Seber method** is an open-population method in which the population can change due to mortality.
- **“Multi-systems estimation”** is based on these approaches and has been used to estimate casualties from armed conflicts (including unobserved casualties).

Manrique-Vallier et al. 2021. Capture-recapture for casualty estimation and beyond: Recent advances and research directions. Chapter in *Statistics in the Public Interest*. Pp. 15-31. Springer.

Stan User's Guide. Mark-recapture models. <https://mc-stan.org/docs/stan-users-guide/latent-discrete.html#mark-recapture-models>

# Mark-recapture model (Lincoln-Peterson)

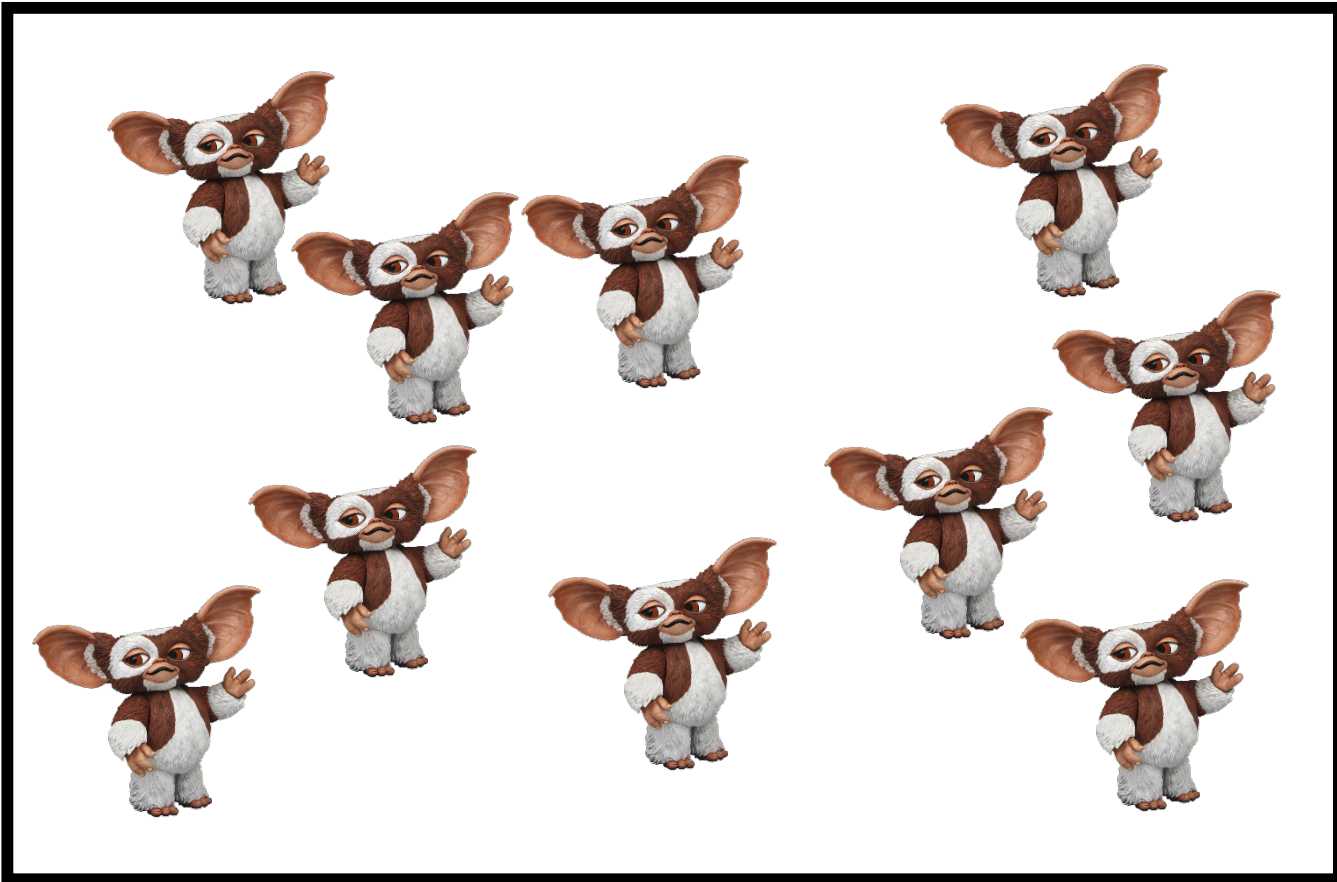
- Estimate an unobservable population size from mark-recapture data
- Conduct a survey randomly sampling  $M$  “marked” individuals from the population
- Conduct a second round of the survey randomly sampling  $C$  “captured” individuals
- Count the number of “recaptured” individuals  $R$  in the second round

$$R \sim \text{Binomial}\left(C, \frac{M}{N}\right)$$

$$N \sim \text{Uniform}(C - R + M, 1e6)$$

# Mark-recapture model (Lincoln-Peterson)

survey = 1, mogwai = 10

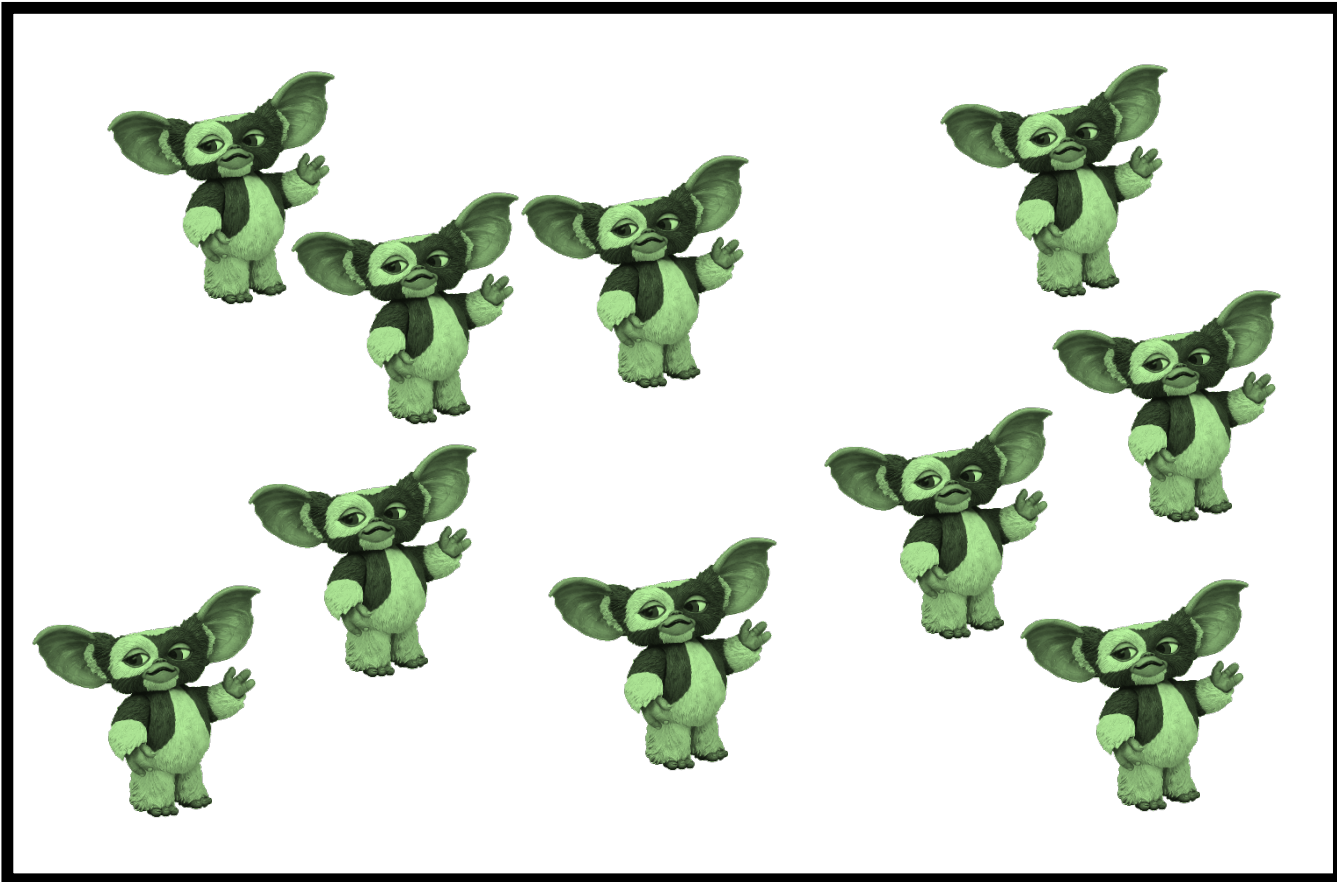


$$R \sim \text{Binomial}\left(C, \frac{M}{N}\right)$$

$$N \sim \text{Uniform}(C - R + M, 1e6)$$

# Mark-recapture model (Lincoln-Peterson)

survey = 1, mogwai = 10



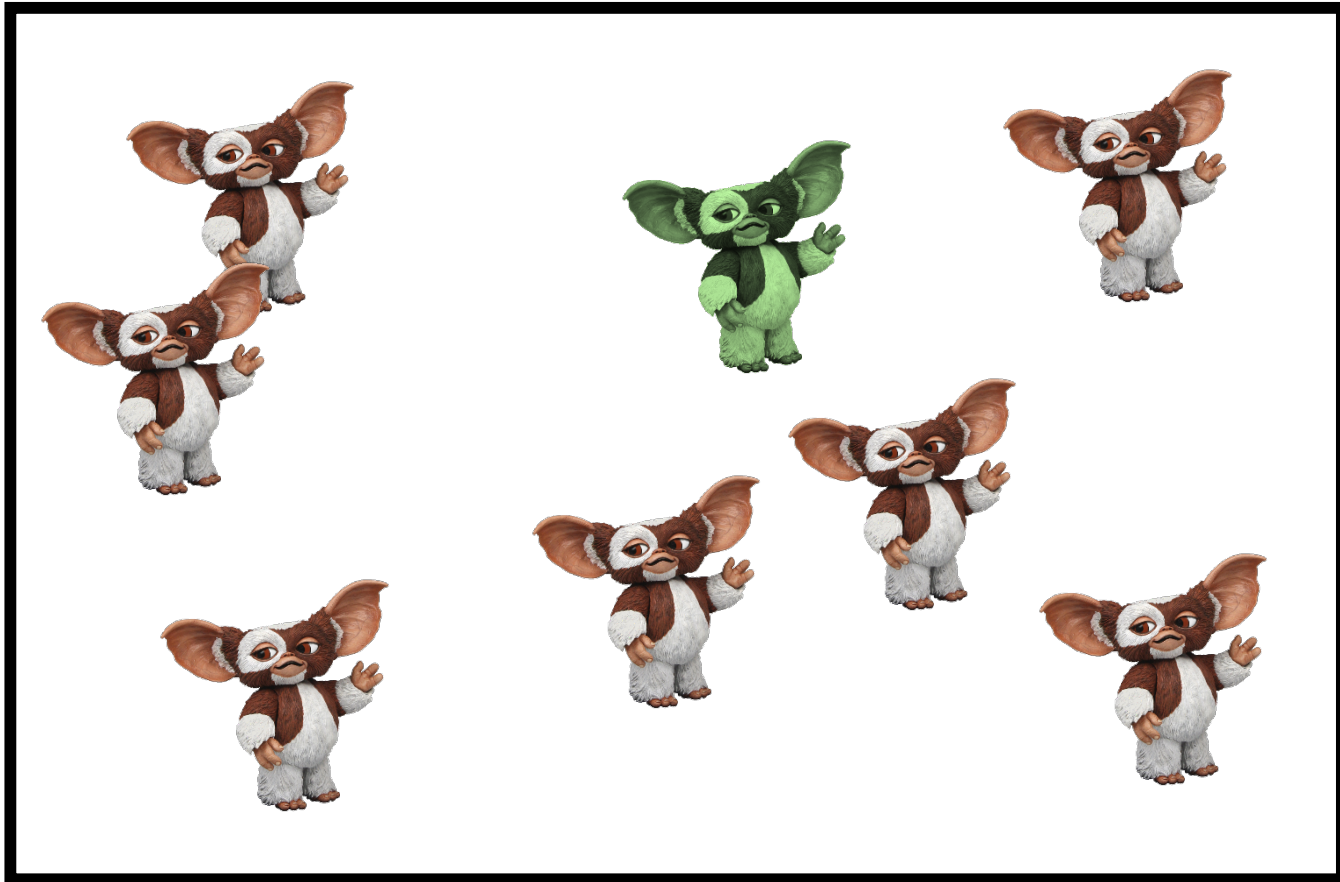
$$R \sim \text{Binomial}\left(C, \frac{M}{N}\right)$$

$$N \sim \text{Uniform}(C - R + M, 1e6)$$

Marked (M) = 10

# Mark-recapture model (Lincoln-Peterson)

survey = 2, mogwai = 8



$$R \sim \text{Binomial}\left(C, \frac{M}{N}\right)$$

$$N \sim \text{Uniform}(C - R + M, 1e6)$$

Marked (M) = 10

Captured (C) = 8

Recaptured (R) = 1

# Mark-recapture model (Lincoln-Peterson)

What is the big assumption in this model?

$$R \sim \text{Binomial}\left(C, \frac{M}{N}\right)$$

The detection probability in round 1 is equal to the probability of round 2 captures being recaptures.

$$N \sim \text{Uniform}(C - R + M, 1e6)$$

$$\frac{M}{N} = \frac{R}{C}$$

# Sometimes we can't “mark” individuals

A “mark” could simply be recording a person's name and birthdate, but sometimes this is not possible.

# N-mixture models of population size

- Repeated  $j$  counts  $y_{i,j}$  of individuals at each location  $i$ .
- The repeated counts allow us to estimate detection probability  $\rho_{i,j}$
- Population size  $N_i$  determined by characteristics of each location  $x_{k,i}$ .
- Assumes a closed population across  $j$  surveys

$$y_{i,j} \sim \text{Binomial}(N_i, \rho_{i,j})$$

$$\text{logit}(\rho_{i,j}) = \alpha + \beta x_{i,j}$$

$$N_i \sim \text{Poisson}(\lambda_i)$$

$$\log(\lambda_i) = \alpha + \beta x_i$$



# N-mixture models of population size

location (i) = 1, survey (j) = 1

mogwai = 6

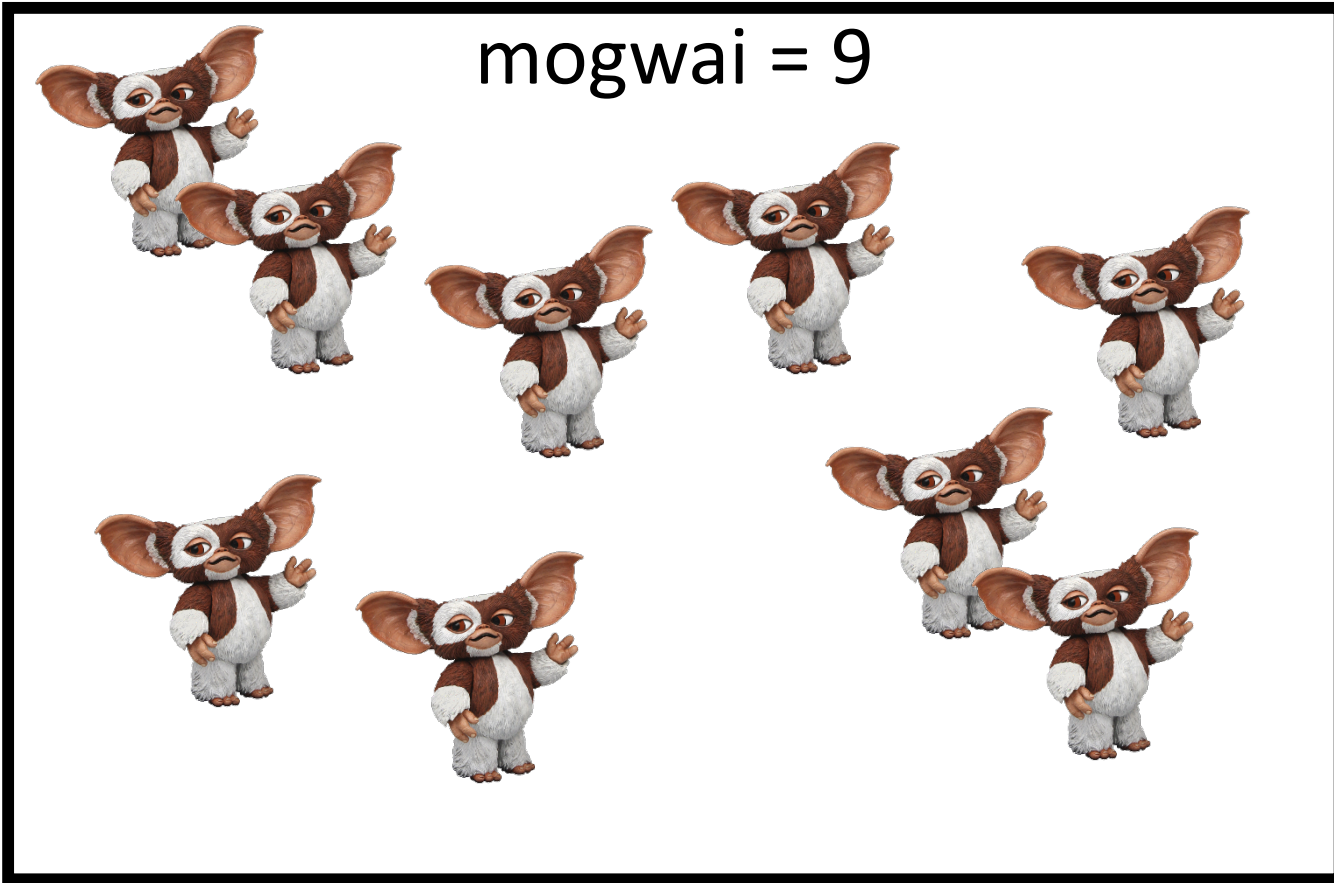
$$y_{1,1} = 6$$



# N-mixture models of population size

location (i) = 1, survey (j) = 2

mogwai = 9



$$y_{1,1} = 6$$

$$y_{1,2} = 9$$

# N-mixture models of population size

location (i) = 1, survey (j) = 3

mogwai = 4



$$y_{1,1} = 6$$

$$y_{1,2} = 9$$

$$y_{1,3} = 4$$

# N-mixture models of population size

Observation model:  $y_{i,j} \sim \text{Binomial}(N_i, \rho_{i,j})$

$$\text{logit}(\rho_{i,j}) = \alpha + \beta x_{i,j}$$

Process model:

$$N_i \sim \text{Poisson}(\lambda_i)$$

$$\log(\lambda_i) = \alpha + \sum_{k=1}^K \beta_k x_{k,i}$$

# State-space models

- Unobserved process in which the state  $z_t$  of the system at time  $t$  depends on the state in the previous time step with stochastic variation  $\epsilon$
- Observed data  $y_t$  that are dependent on the process state  $z_t$  with observation error  $\sigma$
- Make inferences about a process that you cannot observe directly
- Estimate process error separately from observation error

$$P(z_t \mid \theta z_{t-1}, \epsilon)$$

$$P(y_t \mid \beta z_t, \sigma)$$

# State-space models

Hidden Markov model is a specific case of this

$$\text{Process model: } P(z_t \mid \theta z_{t-1}, \epsilon)$$

$$\text{Observation model: } P(y_t \mid \beta z_t, \sigma)$$

# Dail-Madsen open population models

- Apparent survival  $S_{i,t}$  is the number of individuals surviving and not emigrating
- Recruitment  $G_{i,t}$  is the number of new individuals entering the population
- N-mixture observation model

$$S_{i,t} \sim \text{Binomial}(N_{i,t-1}, \omega)$$

$$G_{i,t} \sim \text{Poisson}(\lambda N_{i,t-1})$$

$$N_{i,t} = S_{i,t} + G_{i,t}$$

$$y_{i,j,t} \sim \text{Binomial}(N_{i,t}, \rho_j)$$

$$\text{logit}(\rho_j) \sim \text{Normal}(\mu^\rho, \sigma^\rho)$$

# Mark-recapture state-space model

- A population process
- Why a log-normal? Hmm....
- Priors not shown
- Mark-recapture observations
- A prior is only needed to initialise N in the first time step

$$N_t \sim \text{LogNormal}(\log(N_{t-1}\lambda_t), \sigma)$$

$$\log(\lambda_t) = \alpha + \sum_{k=1}^K \beta_k x_{k,t}$$

$$R_t \sim \text{Binomial}\left(C_t, \frac{M_t}{N_t}\right)$$

$$N_{t=1} \sim \text{Uniform}(C - R + M, 1e6)$$



# Stan and discrete latent parameters

- Stan does not allow parameters to be integers. Only data can be integers.
- This is infinitely annoying, but the Stan developers are firm in their crusade against discrete latent parameters in the noble pursuit of accurately estimating the tails of distributions and increasing MCMC efficiency.
- If you have a model with a latent discrete parameter (e.g. unobserved population size), then you have three options:
  1. Re-parameterise your model to “marginalise out” the integer parameter.
  2. Use a continuous distribution for your integer parameter (e.g. LogNormal).
  3. Use JAGS instead of Stan.

# Recap

- Ordinary linear models
- Generalised linear models
  - Logistic regression
  - Binomial regression
  - Poisson regression
- Log-linear models
- Beta regression
- Mark-recapture

## **Hierarchical Models**

- Random effects
- N-mixture models
- State-space
  - Dail-Madsen N-mixture
  - Mark-recapture state-space

# Components of a model

1. Data model (likelihood)

$$y_i \sim \text{Normal}(\mu_i, \sigma)$$

2. Deterministic process

$$\mu_i = \alpha + \beta x_i$$

3. Parameter model (*e.g.* priors)

$$\alpha \sim \text{Normal}(0, 1)$$

$$\beta \sim \text{Normal}(0, 1)$$

$$\sigma \sim \text{HalfCauchy}(0, 1)$$

# Tie it all together...

Ordinary linear regression

$$y_i \sim \text{Normal}(\mu_i, \sigma)$$

$$\mu_i = \alpha + \beta x_i$$

# Tie it all together...

Generalised linear model

$$y_i \sim \text{Poisson}(\lambda_i)$$

$$\log(\lambda_i) = \alpha + \beta x_i$$

# Tie it all together...

Generalised linear model

$$y_i \sim \text{Binomial}(N_i, \rho_i)$$

$$\text{logit}(\rho_i) = \alpha + \beta x_i$$

# Tie it all together...

Random effects (hierarchical model)

$$y_i \sim \text{Binomial}(N_i, \rho_i)$$

$$\text{logit}(\rho_i) = \alpha_g + \beta x_i$$

$$\alpha_g \sim \text{Normal}(\mu, \sigma)$$

# Tie it all together...

Hierarchical models (e.g. N-mixture model)

Latent process model:

$$N_i \sim \text{Poisson}(\lambda_i)$$

$$\log(\lambda_i) = \alpha + \beta x_i$$

Observation model:

$$y_{i,j} \sim \text{Binomial}(N_i, \rho_{i,j})$$

$$\text{logit}(\rho_{i,j}) = \alpha + \beta x_{i,j}$$



# Tie it all together...

State space model

$$N_{i,t} \sim \text{Poisson}(N_{i,t-1}\lambda_{i,t})$$

$$\log(\lambda_{i,t}) = \alpha + \beta x_{i,t}$$

$$y_{i,j,t} \sim \text{Binomial}(N_{i,t}, \rho_{i,j,t})$$

$$\text{logit}(\rho_{i,j,t}) = \alpha + \beta x_{i,j,t}$$

# Tie it all together...

State space model

$$P(z_t \mid \lambda z_{t-1})$$

$$P(y_t \mid \rho z_t)$$