

# To Do

**Read Sections 6.1 – 6.3.**

**Assignment 4 is due Friday  
November 25.**

# Last Class

- (1) Confidence Interval for Slope  $\beta$  and Testing  $H_0: \beta = \beta_0$
- (2) Test of No Relationship between Response and Explanatory Variates
- (3) Confidence interval for the mean response  $\mu(x) = \alpha + \beta x$
- (4) Prediction Interval for an Individual Response  $Y$

# Simple Linear Regression Model

For data  $(x_i, y_i)$ ,  $i = 1, 2, \dots, n$

we assume the model

$$Y_i \sim G(\alpha + \beta x_i, \sigma) \text{ for } i = 1, 2, \dots, n$$

independently and where the

$x_i$ 's,  $i = 1, 2, \dots, n$

are assumed to be known constants.

# Today's Class

- (1) General Form of a Gaussian Response Model**
- (2) Linear Regression Models**
- (3) Checking the Assumptions of the Simple Linear Regression Model**

# Simple Linear Regression Model

Simple linear regression model

$$Y_i \sim G(\alpha + \beta x_i, \sigma) \text{ for } i=1,2,\dots,n$$

independently and where the

$x_i$ 's,  $i = 1, 2, \dots, n$

are assumed to be known constants.

This model is a member of a larger family of models called **Gaussian response models**.

# Gaussian Response Models

The general form of a **Gaussian response model** is

$$Y_i \sim G(\mu(x_i), \sigma) \text{ for } i=1,2,\dots,n$$

independently and where the  $x_i$ 's,  $i = 1,2,\dots,n$  are assumed to be known constants (possibly vectors).

In this model

$$E(Y_i) = \mu(x_i)$$

depends on the explanatory variate  $x_i$ , but

$$\text{sd}(Y_i) = \sigma$$

does not.

# Gaussian Response Model

## The Gaussian Response Model

$Y_i \sim G(\mu(x_i), \sigma)$  for  $i=1,2,\dots,n$  independently  
can also be written in the form

$Y_i = \mu(x_i) + R_i$  where  $R_i \sim G(0, \sigma)$ ,  $i=1,2,\dots,n$   
independently.

$Y_i$  is a sum of two components.

The first component,  $\mu(x_i)$ , is a deterministic component (not a random variable) and the second component,  $R_i$ , is a random component or random variable.

# Linear Regression Models and STAT 331/STAT 371/STAT 373

In many examples the deterministic component takes the form

$$E(Y_i) = \mu(\mathbf{x}_i) = \beta_0 + \sum_{j=1}^k \beta_j x_{ij}$$

so  $E(Y_i)$  is a linear function of

$\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{ik})$ , a vector of explanatory variates for unit  $i$ , and the unknown parameters  $\beta_0, \beta_1, \dots, \beta_k$ .

These models are called **linear regression models**.

The  $\beta_j$ 's are called the regression coefficients.

The  $x_i$ 's are called covariates.



# Model Checking

There are two main assumptions for Gaussian linear response models:

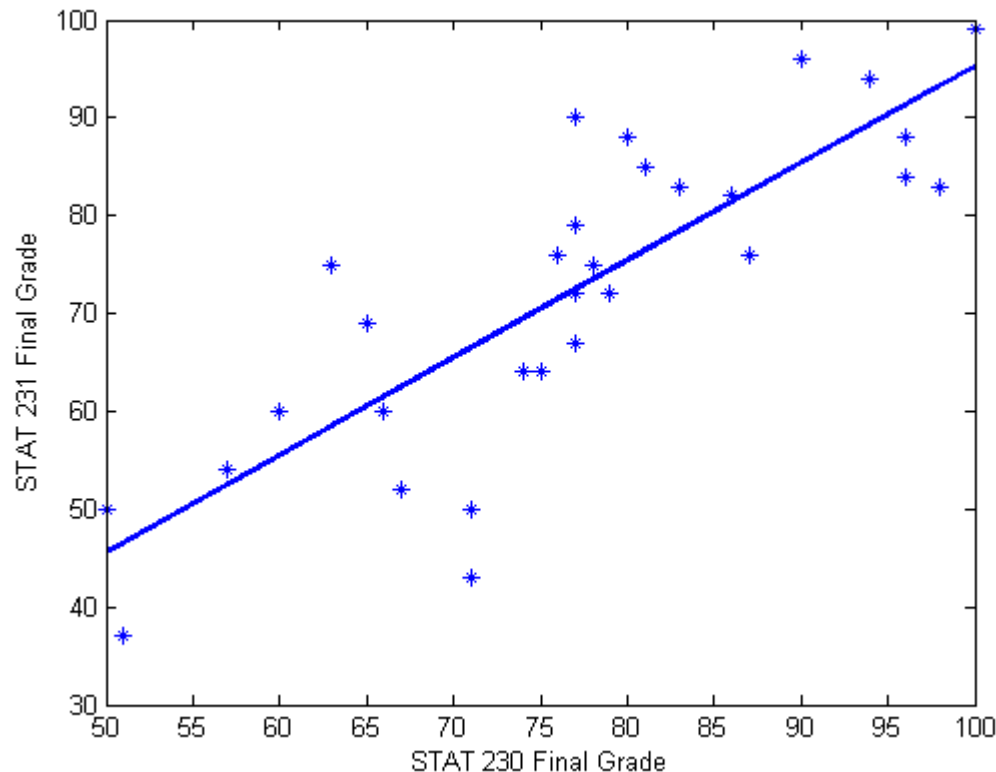
- (1)  $Y_i$  (given covariates  $x_i$ ) has a Gaussian distribution with standard deviation  $\sigma$  which does not depend on the covariates.
- (2)  $E(Y_i) = \mu(x_i)$  is a linear combination of known covariates  $x_i = (x_{i1}, x_{i2}, \dots, x_{ik})$ , and unknown regression coefficients  $\beta_0, \beta_1, \dots, \beta_k$ .

**MODEL ASSUMPTIONS SHOULD ALWAYS BE CHECKED!!!**

We use graphical methods to do this.

# Model Checking Method 1

In simple linear regression, a scatterplot of the data with the fitted line superimposed indicates how well the model fits the data.



# Model Checking Method 2 - Residual Plots

Residual plots are very useful for model checking when there are 2 or more covariates. For the simple linear regression model let

$$\hat{\mu}_i = \hat{\alpha} + \hat{\beta}_i x_i$$

(often called the “fitted” response)

and

$$\hat{r}_i = y_i - \hat{\mu}_i$$

The  $\hat{r}_i$ 's are called residuals since  $\hat{r}_i$  represents what is “left” after the model has been “fitted” to the data.

# Residual Plots

The idea behind the  $\hat{r}_i$ 's is that they can be thought of as “observed”  $R_i$ 's in the model

$Y_i = \mu_i + R_i$  where  $R_i \sim G(0, \sigma)$ ,  $i = 1, 2, \dots, n$  independently.

This isn't exactly correct since we are using  $\hat{\mu}_i$  instead of  $\mu_i$ .

However if the model is correct, then the  $\hat{r}_i$ 's should behave roughly like a random sample from the  $G(0, \sigma)$  distribution.

# Residual Plots

## Recall

$$\hat{\alpha} = \bar{y} - \hat{\beta}\bar{x} \quad \text{or} \quad \bar{y} - \hat{\alpha} - \hat{\beta}\bar{x} = 0$$

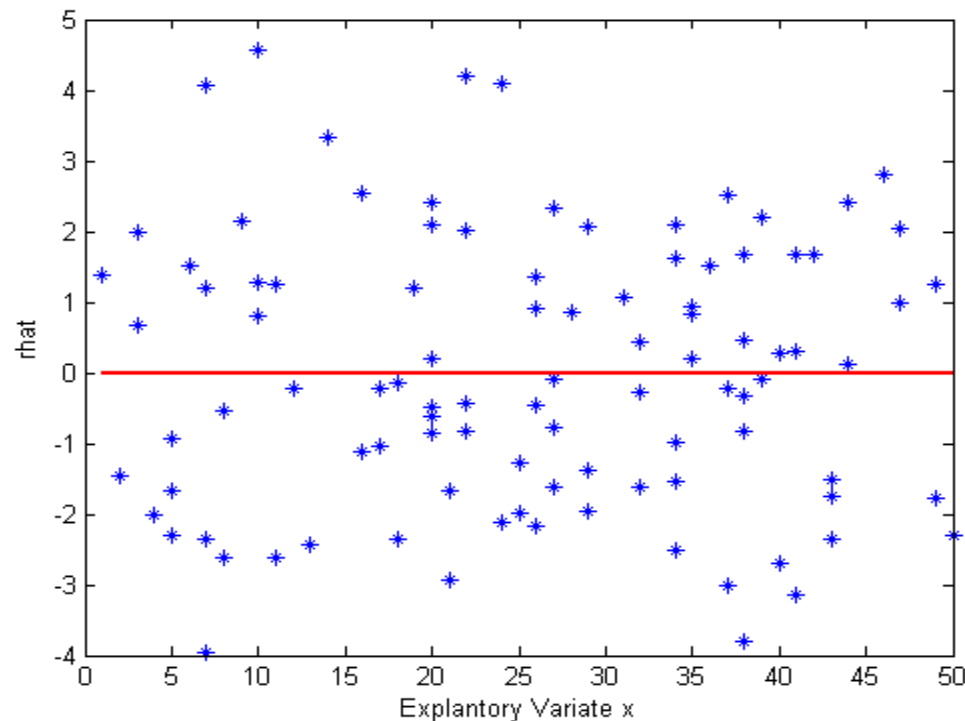
**which implies**

$$0 = \bar{y} - \hat{\alpha} - \hat{\beta}\bar{x} = \frac{1}{n} \sum_{i=1}^n (y_i - \hat{\alpha} - \hat{\beta}x_i) = \frac{1}{n} \sum_{i=1}^n \hat{r}_i$$

**so that the average of the residuals is always zero.**

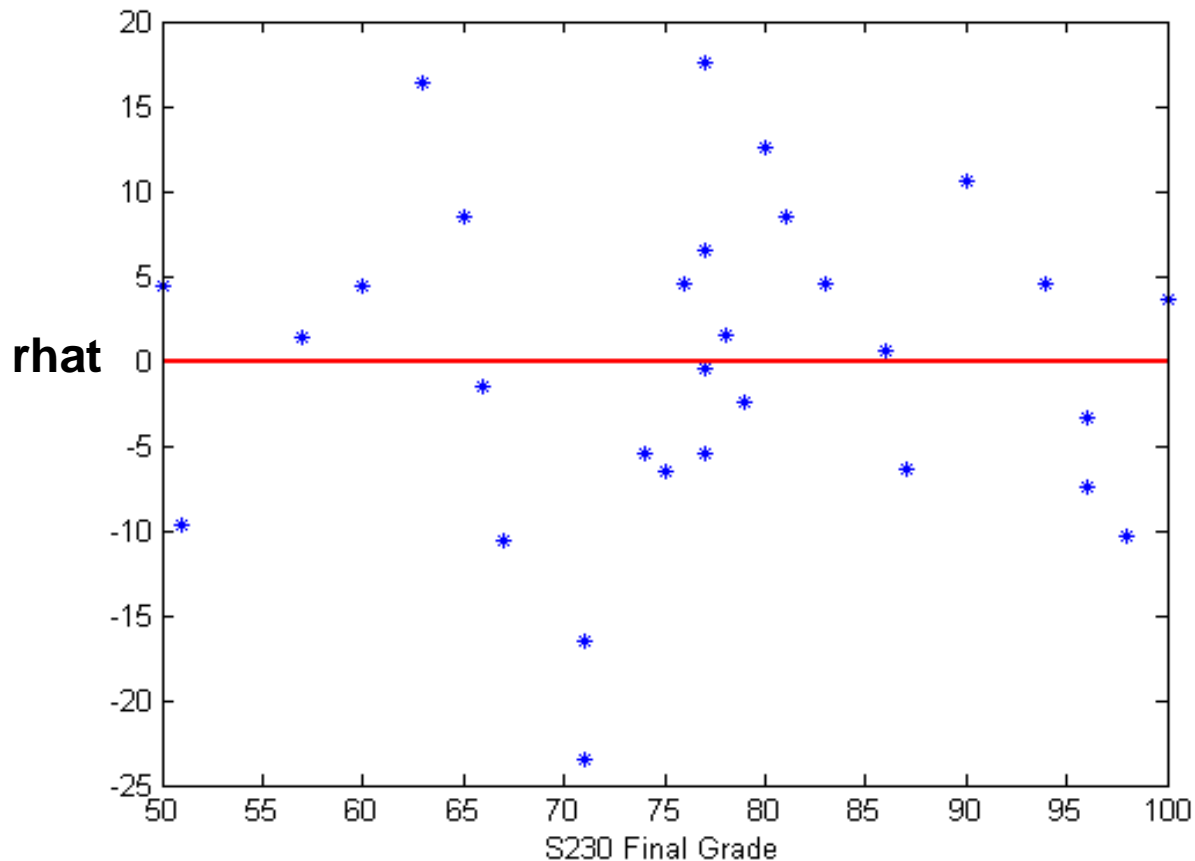
# Residual Plots

If the model assumptions hold then a plot of the points  $(x_i, \hat{r}_i)$ ,  $i = 1, 2, \dots, n$  should lie more or less within a horizontal band or belt around the line  $\hat{r}_i = 0$  showing no obvious pattern.



# STAT 231/230 Residual Plot

What would you conclude?



# Standardized Residual Plots

Define the standardized residuals

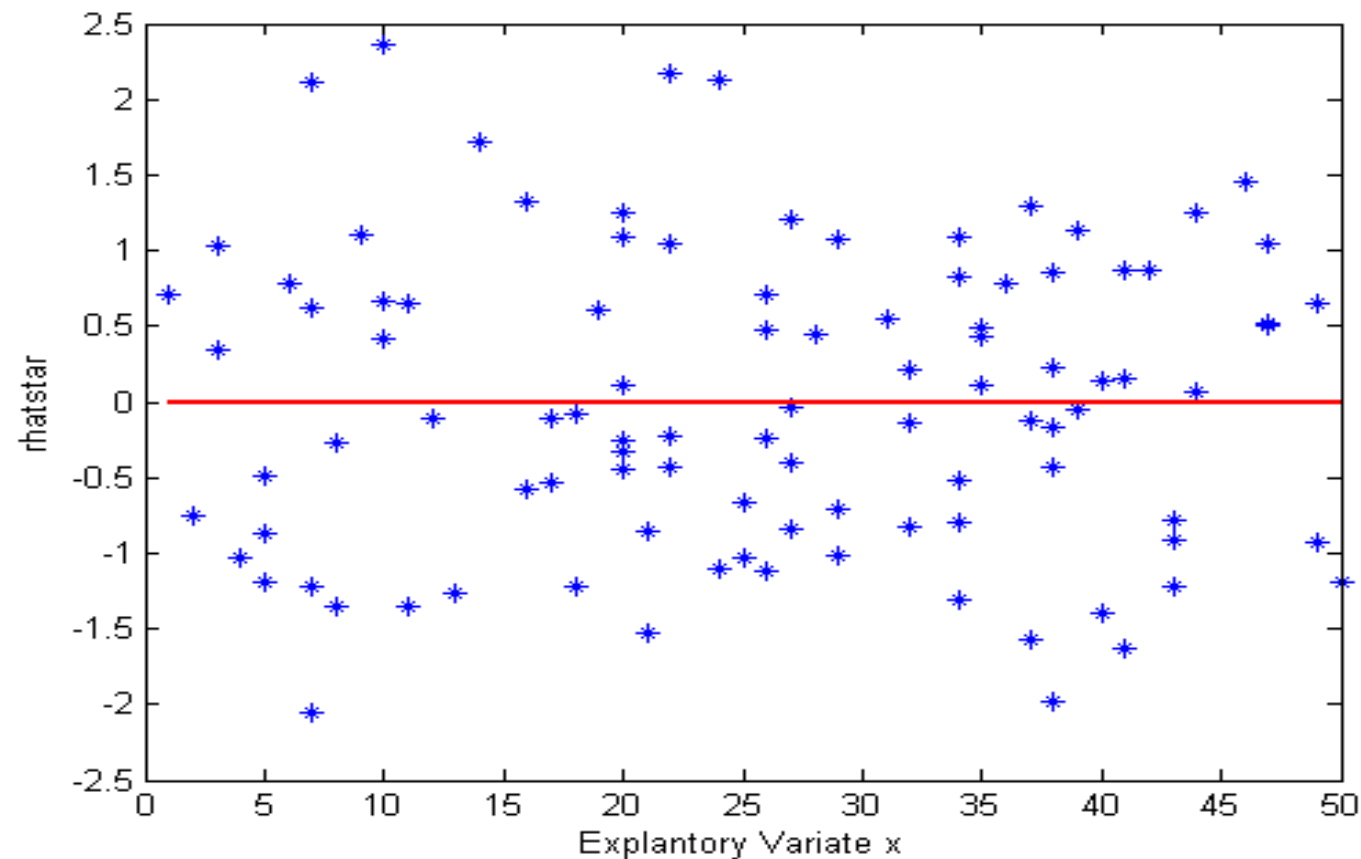
$$\hat{r}_i^* = \frac{\hat{r}_i}{s_e} = \frac{y_i - \hat{\mu}_i}{s_e} \quad i = 1, 2, \dots, n$$

What is the only difference between a plot of the points  $(x_i, \hat{r}_i)$ ,  $i = 1, 2, \dots, n$  and a plot of the points  $(x_i, \hat{r}_i^*)$ ,  $i = 1, 2, \dots, n$  ?

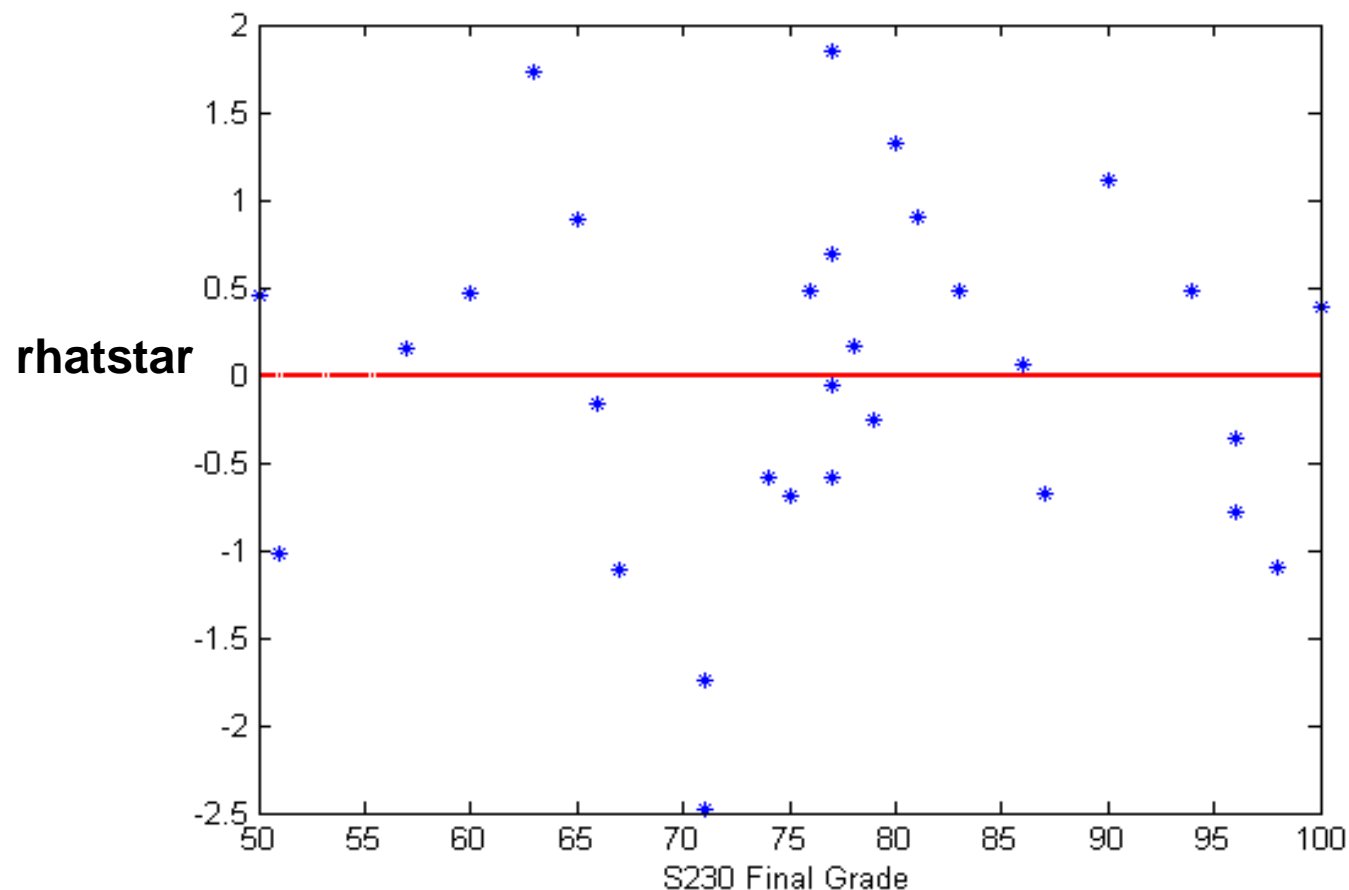
If the model is correct then the  $\hat{r}_i^*$  values will lie in the range  $(-3, 3)$ . Why is this?



# Example – Standardized Residual Plot



# STAT 231/230 Standardized Residual Plot



# Residual Plot Type 2

Another type of residual plot consists of plotting the points

$$(\hat{\mu}_i, \hat{r}_i^*), \quad i = 1, 2, \dots, n$$

Such a plot can be used to check the assumption about the form of the mean  $E(Y_i) = \mu(x_i)$ .

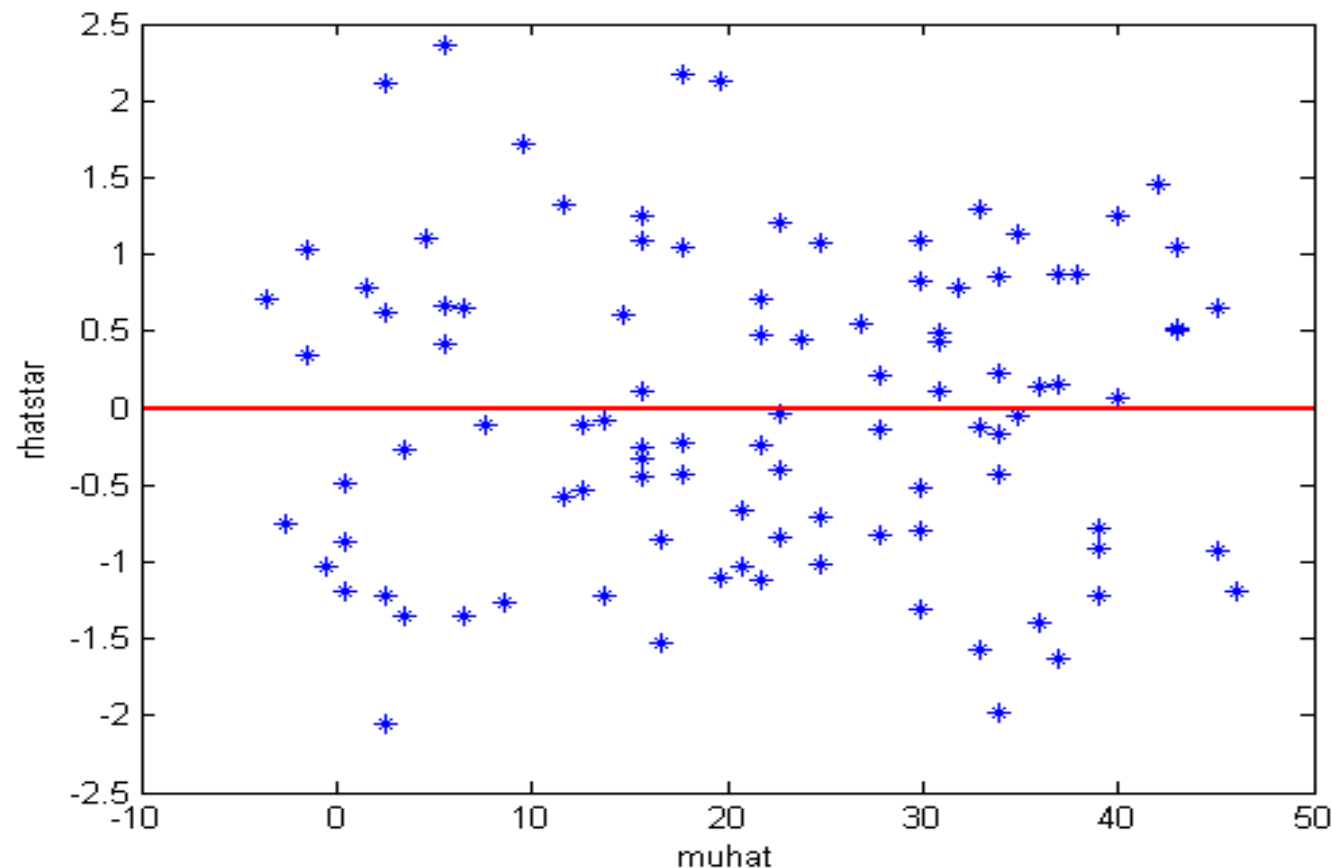
For the simple linear regression model we are checking whether the assumed mean

$E(Y_i) = \mu(x_i) = \alpha + \beta x_i$  is reasonable.

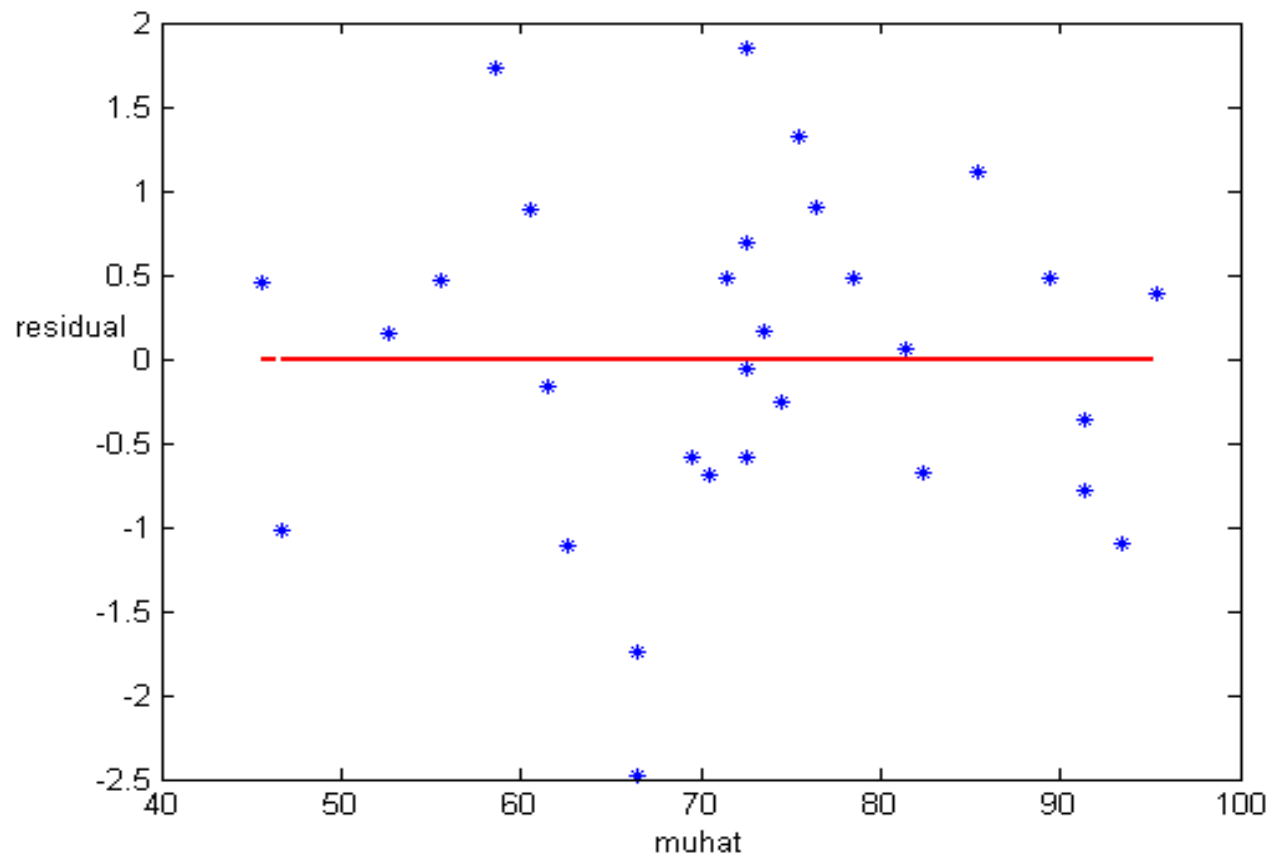
If the assumed mean is reasonable we should see approximately a horizontal band around the line

$$\hat{r}_i^* = 0 \quad .$$

# Example – Standardized Residual Plot Using Muhat



# STAT 231/230 Standardized Residual Plot with Muhat



# Qqplot of Residuals

To check the Gaussian assumption we use a qqplot of the standardized residuals.

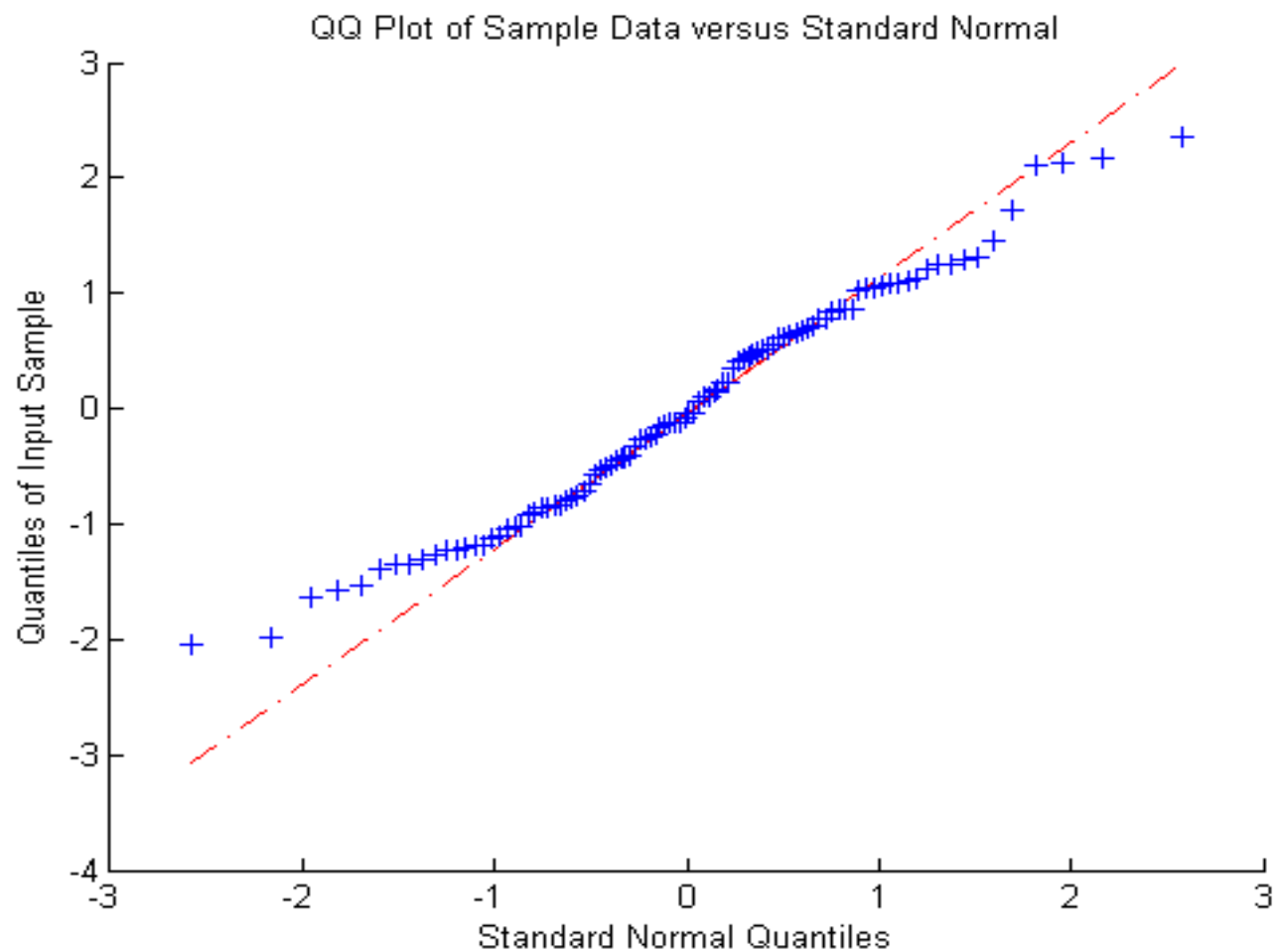
Since our assumed model is

$$R_i / \sigma = (Y_i - \mu_i) / \sigma \sim G(0,1)$$

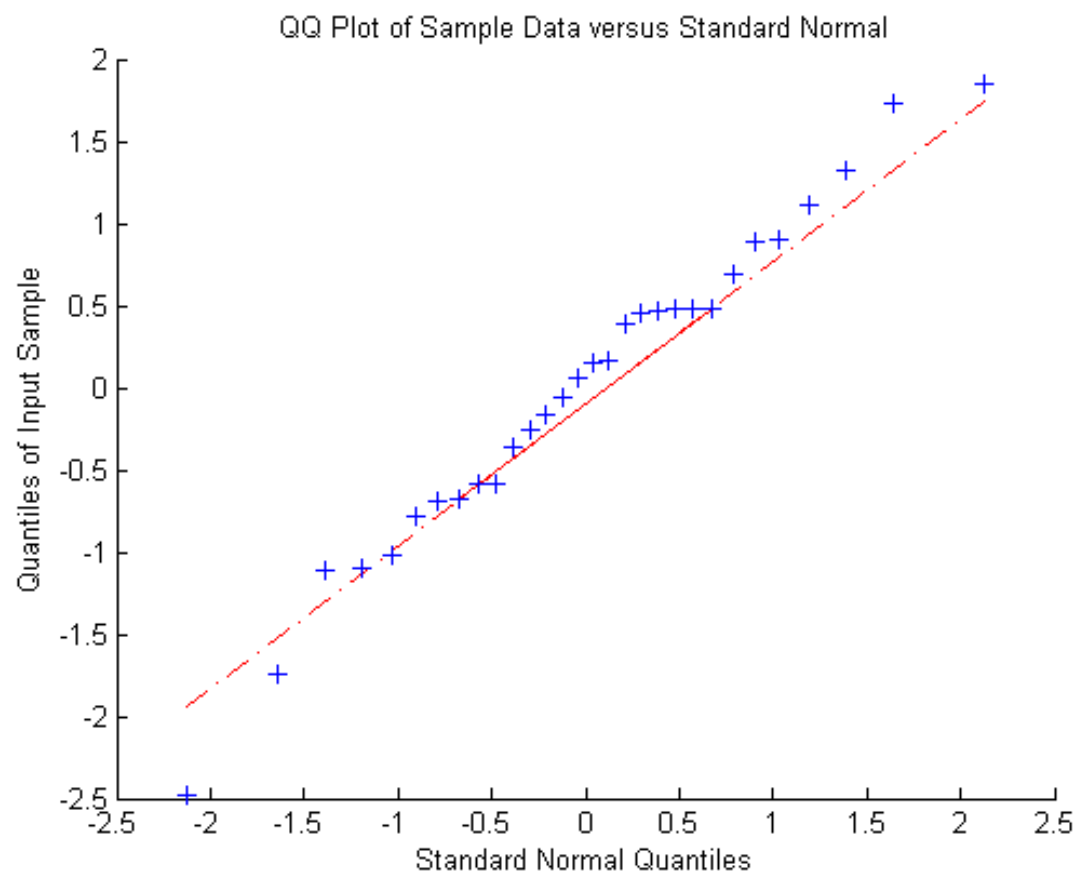
the  $\hat{r}_i^*$ 's should roughly represent a sample from the  $G(0,1)$  distribution.

Therefore a qqplot of the  $\hat{r}_i^*$ 's should give approximately a straight line if the model assumptions hold.

# Example - Qqplot



# STAT 231/230 Qqplot





# Interpreting Residual Plots

If a plot of the points

$$(\hat{\mu}_i, \hat{r}_i^*), \quad i = 1, 2, \dots, n$$

or

$$(\hat{\mu}_i, \hat{r}_i), \quad i = 1, 2, \dots, n$$

shows a distinctive pattern then this suggests the assumed form for  $E(Y_i) = \mu(x_i)$  may be inappropriate.

# Interpreting Residual Plots

If a plot of the points

$$(\hat{\mu}_i, \hat{r}_i^*), i = 1, 2, \dots, n$$

indicates that the variability in the  $\hat{r}_i^*$ 's is bigger for large values of  $\hat{\mu}_i$  than for small values of  $\hat{\mu}_i$  (or vice versa) then there is evidence to suggest that the assumption of constant variance,  $\text{Var}(Y_i) = \text{Var}(R_i) = \sigma^2, i=1, 2, \dots, n$  does not hold.

# Interpreting Residual Plots

**If the points in the qqplot do not lie roughly in a straight line then this suggests the Gaussian assumption may not hold.**

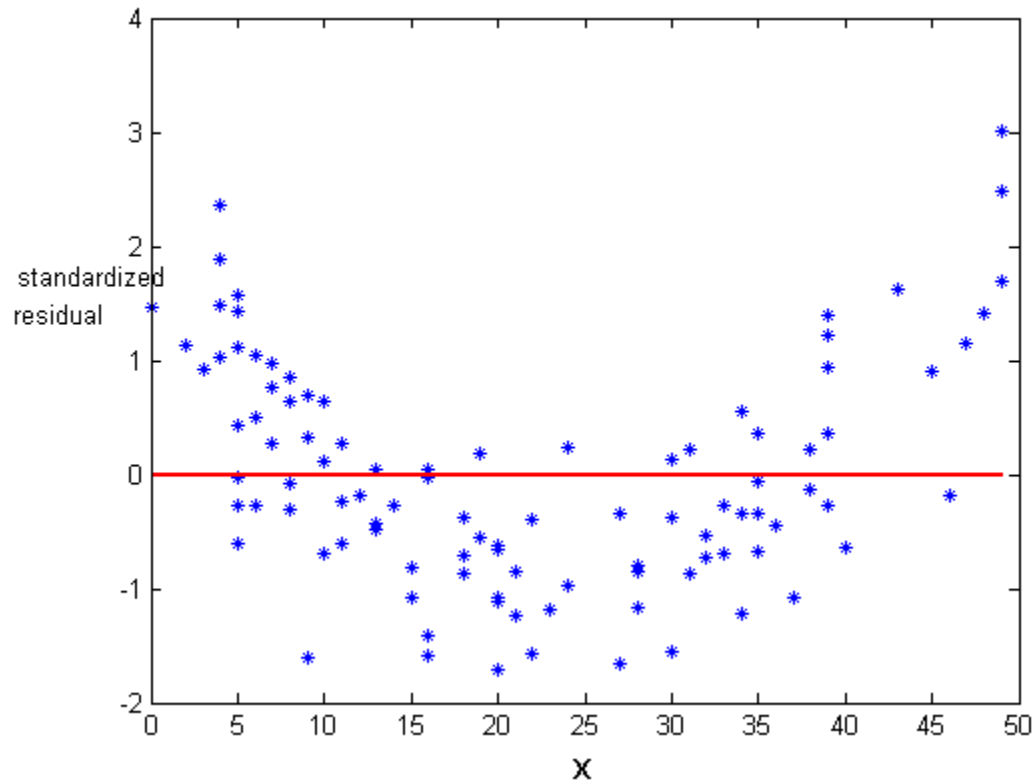
# **Interpreting Residual Plots - Warning**

**Reading these plots takes practice and you should try not to read too much into plots especially if the plots are based on a small number of points.**

**The plots on the next slides exhibit patterns.**

# Examples of Residual Plots with Patterns

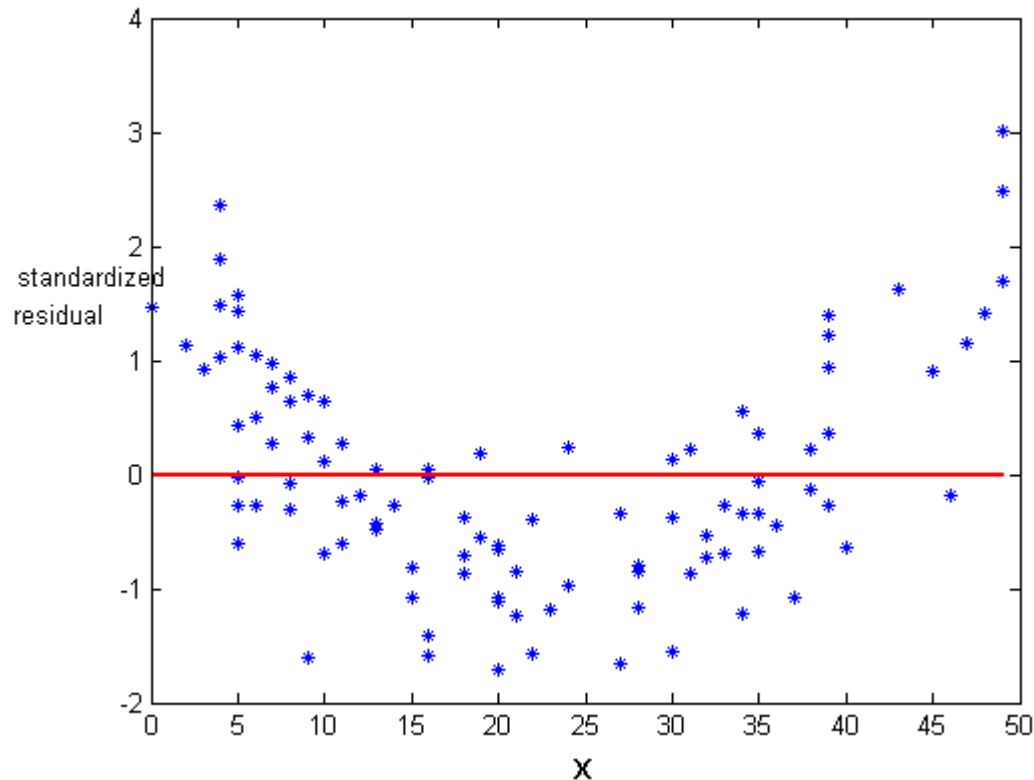
This plot suggests that the function  $\mu(x_i)$  is not correctly specified. Can you suggest a better model?



# Examples of Residual Plots with Patterns

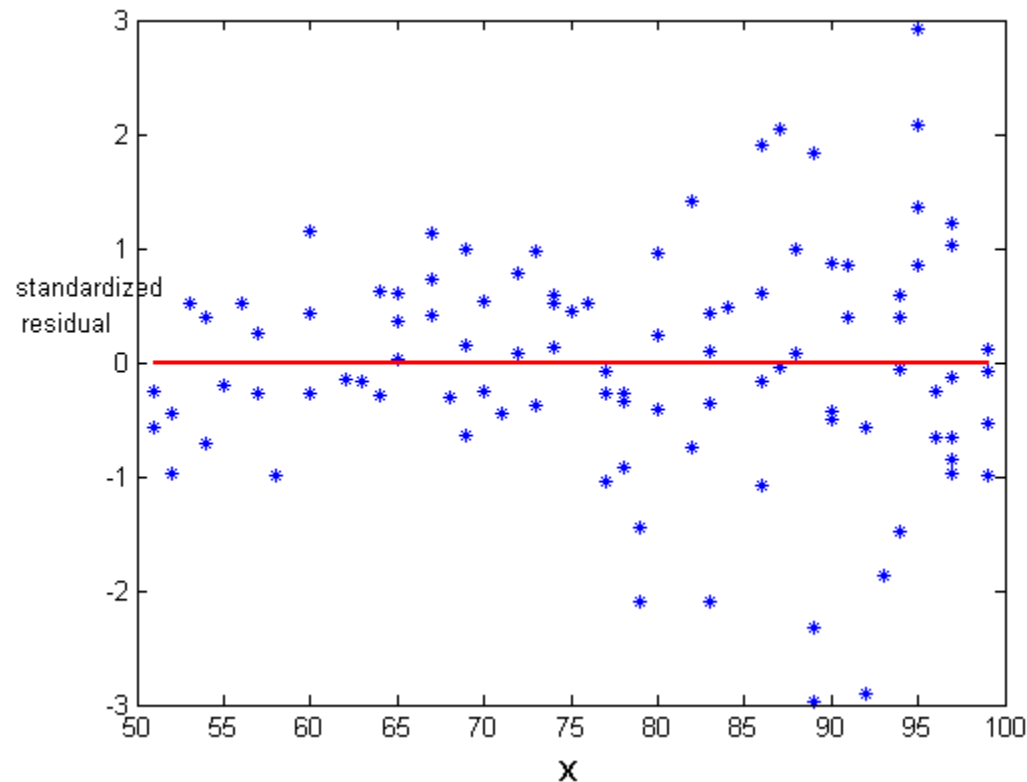
Assume a quadratic model for the mean:

$$\mu(x_i) = \alpha + \beta x_i + \gamma x_i^2 \text{ rather than } \mu(x_i) = \alpha + \beta x_i$$



# Examples of Residual Plots with Patterns

What do you notice?



# Scatterplot for same data

