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## Directed Polymers in Random Environment

### Chapter 3

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#### Abstract

These notes are based on the book "Directed Polymers in Random Environment" [2], and are intended for self-study and understand better these topics.

## 1 Martingale Approach and the $L^2$ region.

In this section we will study the partition function  $Z_n$ . Recall that  $\lambda(\beta) = \log \mathbb{P}(e^{\beta\omega})$ , so  $\mathbb{P}(e^{\beta\omega}) = e^{\lambda(\beta)}$ . Therefore,

$\mathbb{P}(Z_n) = \mathbb{P}(P[e^{\beta H_n}]) = P(\mathbb{P}(\exp\{\sum_{i=1}^n \omega(i, S_i)\})) = P(e^{n\lambda(\beta)}) = e^{n\lambda(\beta)}$ , where we used the fact that the environment is i.i.d under  $\mathbb{P}$ . Because this calculation, we will consider the normalized partition function

$$W_n := \frac{Z_n}{e^{n\lambda(\beta)}}, n \geq 1 \quad (1.0.1)$$

Now, fix a path  $x$ , and consider  $\bar{\xi}_n = \bar{\xi}_n(x) := e^{\beta H_n(x) - n\lambda(\beta)}$ . Then for that path,  $H_n(x)$  is a random walk. In particular,  $\xi_n$  is a martingale on  $(\Omega, \mathcal{F}, \mathbb{P})$  respect to the filtration  $(\mathcal{F}_n)_{n \in \mathbb{N}}$  given by  $\mathcal{F}_n := \sigma(\omega(j, x) : 1 \leq j \leq n, x \in \mathbb{Z}^d)$ . Then if  $x_1, \dots, x_k$  are paths and  $a_1, \dots, a_k$  are scalars,  $\sum_{i=1}^k a_i \bar{\xi}_n(x_i)$  is also a martingale. Because  $P(\bar{\xi}_n) = \sum_{x \in \mathbb{Z}^d} P(x) \bar{\xi}_n(x)$  then  $W_n = P(\bar{\xi}_n)$  also is a positive martingale.

Recall now the martingale convergence theorem, that basically says that if a martingale has  $L^1$  norm uniformly bounded, then the limit exists  $\mathbb{P}$  almost surely. In our case, the limit  $\lim_{n \rightarrow \infty} W_n$  will be called  $W_\infty$ . A natural question is ask if  $W_\infty = 0$  or not. In fact, note first that by

$$Z_{n+m} = Z_n \times P_n^{\beta, \omega}(Z_m \circ \theta_{n, S_n}) \quad (1.0.2)$$

we can write

$$W_{n+m} = P(\overline{\xi_n} \times W_m \circ \theta_{n,S_n})$$

Taking  $m \rightarrow \infty$  we deduce

$$\begin{aligned} W_\infty &= \lim_{m \rightarrow \infty} W_m = P(\overline{\xi_n} \times \lim_{m \rightarrow \infty} W_m \circ \theta_{n,S_n}) = P(\overline{\xi_n} \times W_\infty \circ \theta_{n,S_n}) \\ &= \sum_x P(S_n = x) \overline{\xi_n}(x) \times W_\infty \circ \theta_{n,x} = \sum_x P(S_n = x) e^{\beta H_n(x) - n\lambda(\beta)} \times W_\infty \circ \theta_{n,x} \\ &= Z_n e^{-n\lambda(\beta)} \sum_x \frac{1}{Z_n} (S_n = x) e^{\beta H_n(x)} \times W_\infty \circ \theta_{n,x} = W_n \times \sum_x P_n^{\beta,\omega}(S_n = x) \times W_\infty \circ \theta_{n,x} \end{aligned}$$

Because  $W_n > 0$  for all  $n$ , we have

$$\{W_\infty = 0\} = \bigcap_{x \in \mathbb{Z}^d: P(S_n=x)} \{W_\infty \circ \theta_{n,x} = 0\} \in \mathcal{F}_n$$

Thus, the event  $\{W_\infty = 0\}$  belongs to the tail  $\sigma$  algebra, and we can use the Kolmogorov's 0-1 law to deduce the following theorem:

**Theorem 1.1.** *The limit*

$$W_\infty = \lim_{n \rightarrow \infty} W_n$$

*exists  $\mathbb{P}$ - a.s. Moreover, the limit is strictly positive, or is zero  $\mathbb{P}$  almost surely.*

It's natural to introduce some terminology that distinguishes this dichotomy:

**Definition 1.2.** *We say that the polymer is in the weak disorder when  $W_\infty > 0$  - a.s and in the strong disorder if  $W_\infty = 0$  - a.s*

This definition is an analogous of the high/low temperature defined in the last section. In fact, there will be a similar statement related to a critical value.

**Remark 1.3.** *We can prove that  $W_\infty > 0 \Rightarrow p(\beta) = \lambda(\beta)$ . Indeed, we have a.s*

$$p(\beta) = \lim_{n \rightarrow \infty} \frac{1}{n} \log(Z_n(\beta, \omega)) = \lambda(\beta) + \lim_{n \rightarrow \infty} \frac{1}{n} \log(W_n) = \lambda(\beta)$$

*because  $\log(W_n) \rightarrow \log(W_\infty) < \infty$*

Now we state the analogous of Theorem 1.19, chapter 2:

**Proposition 1.4.** *There exists  $\bar{\beta}_c = \bar{\beta}_c(\mathbb{P}, d) \in [0, \infty]$  such that*

$$\begin{cases} W_\infty > 0 & \beta \in (0, \bar{\beta}_c) \\ W_\infty = 0 & \beta > \bar{\beta}_c \end{cases}$$

*Proof.* Let  $\delta \in (0, 1)$  arbitrary. The idea is to prove the result for  $(W_n^\delta)_n$  for every  $\delta$ . Note that since  $\mathbb{P}(W_n) = 1$ , then  $(W_n^\delta)_n$  is uniformly integrable (e.g. Hölder). We have then the a.s. convergence of  $W_n^\delta \rightarrow W_\infty^\delta$  and  $\mathbb{P}(W_n^\delta) \rightarrow \mathbb{P}(W_\infty^\delta)$ , which is 0 in the strong disorder case, and strictly positive in the weak one. Is enough to prove that the map  $\beta \rightarrow \mathbb{P}(W_n^\delta)$  is non-increasing on  $\mathbb{R}_+$ . This implies that the map  $\beta \rightarrow \mathbb{P}(W_\infty^\delta)$  also is non-increasing on  $\mathbb{R}_+$ . To finish, define  $\bar{\beta}_c := \inf\{\beta \geq 0 : \mathbb{P}(W_n^\delta) = 0\} \in [0, \infty]$ . We compute

$$\begin{aligned} \frac{d}{d\beta} \mathbb{P}(W_n^\delta) &= \mathbb{P}\left(\frac{d}{d\beta} W_n^\delta\right) = \delta \mathbb{P}\left(W_n^{\delta-1} \frac{d}{d\beta} Z_n e^{-n\lambda}\right) = \delta \mathbb{P}\left(W_n^{\delta-1} \left[e^{-n\lambda} \frac{d}{d\beta} Z_n - n\lambda' e^{-n\lambda} Z_n\right]\right) \\ &= \delta \mathbb{P}\left(W_n^{\delta-1} [e^{-n\lambda} P(H_n \bar{\xi}_n) - n\lambda' P(\bar{\xi}_n)]\right) = \delta \mathbb{P}(W_n^{\delta-1} P[\{H_n - n\lambda'\} \bar{\xi}_n]) \\ &= \delta P(\mathbb{P}[\bar{\xi}_n W_n^{\delta-1} \{H_n - n\lambda'\}]) \end{aligned} \quad (1.0.3)$$

Now, we apply the FGK-Harris inequality as follows: we consider the measure  $\bar{\xi}_n d\mathbb{P}$  for fixed path  $x$ , the decreasing function (on  $\omega$ )  $W_n^{\delta-1}$  and the non-decreasing one  $H_n - n\lambda'$ . Then we obtain

$$\frac{d}{d\beta} \mathbb{P}(W_n^\delta) = \delta P(\mathbb{P}[\bar{\xi}_n W_n^{\delta-1} \{H_n - n\lambda'\}]) \leq \delta P(\mathbb{P}[\bar{\xi}_n W_n^{\delta-1}]) \mathbb{P}[\bar{\xi}_n \{H_n - n\lambda'\}] = 0$$

The last equality is by

$$\mathbb{P}[\omega e^{\beta\omega}] = \lambda'(\beta) e^{\lambda(\beta)} \quad (1.0.4)$$

□

Because the last theorem and Theorem 1.19, chapter 2, it's expected that  $\beta_c = \bar{\beta}_c$ , but at the moment this question remains open. However, if  $d \in \{1, 2\}$ , the equality is satisfied. Note that always  $\bar{\beta}_c \leq \beta_c$ . In effect, if  $\beta = \beta_c + \epsilon$  with  $\epsilon > 0$ , then  $\beta \geq \bar{\beta}_c$ , otherwise  $W_\infty(\beta) > 0 \Rightarrow p(\beta) = \lambda(\beta)$  (by Remark 1.3)  $\Rightarrow \beta \leq \beta_c$ , a contradiction.

**Question 1.**  $\beta_c = \bar{\beta}_c$ ?

**Question 2.** What happens at  $\bar{\beta}_c$ ?

It's expected that  $W_\infty(\bar{\beta}_c) = 0$

## 1.1 The second moment method and the $L^2$ region

We recall some known facts about the return probability of the simple random walk, denoted by  $\pi_d$ . More precisely,

$$\pi_d := P(S_n = 0 \text{ for some } n \geq 1) \quad (1.1.1)$$

Then  $\pi_d = 1$  if  $d \in \{1, 2\}$  and  $< 1$  for  $d \geq 3$ . In fact,  $\pi_d < 1$  if  $d \geq 3$ . For more information, see [5]. Now we define the important  $L^2$  condition:

**Definition 1.5** ( $L^2$  condition). *Given  $d \geq 3$ , we say that  $\beta$  satisfies the  $L^2$  condition if*

$$\lambda_2(\beta) := \lambda(2\beta) - 2\lambda(\beta) < \log(1/\pi_d) \quad (1.1.2)$$

Note that  $\lambda_2(0) = 0$  and  $\lambda'_2(\beta) = 2(\lambda'(2\beta) - \lambda'(\beta)) > 0$ , since  $\lambda'$  is increasing. Therefore,  $\lambda_2(\beta)$  is increasing in  $\mathbb{R}_+$ , and if  $d \geq 3$ , the  $L^2$  condition is equivalent to  $\beta < \beta_{L^2}$ , with

$$\beta_{L^2} := \inf\{\beta \geq 0 : \lambda_2 \leq \log(1/\pi_d)\} \quad (1.1.3)$$

**Theorem 1.6.** *Suppose that  $d \geq 3$  and the  $L^2$  condition is satisfied. Then  $W_\infty > 0$  a.s. In particular,  $p = \lambda$  if  $\beta \leq \beta_{L^2}$*

*Proof.* Let's consider the product space  $(\Omega^2, \mathcal{F}^{\otimes 2})$  and the probability measure  $P^{\otimes 2}(dx, d\tilde{x})$ . This is the distribution of two independent walks  $S, \tilde{S}$ . Note that

$$W_n^2 = P \frac{Z_n^2}{e^{2n\lambda(\beta)}} = P(e^{\beta H_n(S)}) P(e^{\beta H_n(\tilde{S})}) e^{-2n\lambda(\beta)} = P^{\otimes 2}(e^{\beta[H_n(S) - H_n(\tilde{S})] - 2n\lambda(\beta)})$$

Now we compute  $\mathbb{P}(W_n^2)$ :

$$\begin{aligned} \mathbb{P}(W_n^2) &\stackrel{\text{Fubini}}{=} P^{\otimes 2} \mathbb{P} \left[ \prod_{t=1}^n e^{\beta[\omega(t, S_t) + \omega(t, \tilde{S}_t)] - 2\lambda(\beta)} \right] = P^{\otimes 2} \left[ \prod_{t=1}^n \mathbb{P} \left( e^{\beta[\omega(t, S_t) + \omega(t, \tilde{S}_t)] - 2\lambda(\beta)} (1_{S_t = \tilde{S}_t} + 1_{S_t \neq \tilde{S}_t}) \right) \right] \\ &= P^{\otimes 2} \left[ \prod_{t=1}^n e^{\lambda(2\beta) - 2\lambda(\beta)} (1_{S_t = \tilde{S}_t} + 1_{S_t \neq \tilde{S}_t}) \right] \\ &= P^{\otimes 2} \left[ \prod_{t=1}^n e^{\lambda_2(\beta) 1_{S_t = \tilde{S}_t}} \right] = P^{\otimes 2} [e^{\lambda_2(\beta) N_n}] \end{aligned}$$

where  $N_n = N_n(S, \tilde{S}) := \sum_{t=1}^n 1_{S_t = \tilde{S}_t}$  is the number of intersections of  $S, \tilde{S}$  up to time  $n$ . Taking  $n \rightarrow \infty$ , we have  $N_n \rightarrow \infty$ , and by the Monotone Convergence Theorem,

$$\mathbb{P}(W_n^2) \rightarrow P^{\otimes 2}[e^{\lambda_2(\beta) N_\infty}]$$

Also note that if  $X_n = S_n - \tilde{S}_n$ , then  $N_\infty$  is the number of visits to zero of the symmetric random walk  $X_n$  (this walk is not a nearest-neighbor one). Then  $N_\infty$  is geometrically distributed with success probability  $\pi_d$ . Then we have

$$P^{\otimes 2}[e^{\lambda_2(\beta)N_\infty}] = \sum_{k=0}^{\infty} (1 - \pi_d) \pi_d^k e^{k\lambda_2} = \begin{cases} \frac{1-\pi_d}{1-\pi_d e^{\lambda_2}} & \pi_d e^{\lambda_2} < 1 \\ +\infty & \pi_d e^{\lambda_2} \geq 1 \end{cases}$$

Therefore,  $\sup_n \mathbb{P}(W_n^2) < \infty \Leftrightarrow \lambda_2 + \log(\pi_d) < 0 \Leftrightarrow \lambda_2 < \log(1/\pi_d)$ . In that case, the martingale  $W_n$  is bounded in  $L^2$ . This implies not only the convergence to  $W_\infty$ , but also convergence in  $L^1$ . Thus,  $\mathbb{P}(W_\infty) = 1$  and necessarily  $W_\infty$  must be strictly positive.  $\square$

**Corollary 1.7.** *Let  $s = \text{ess sup}_{\mathbb{P}} \omega(t, x)$ . The function  $\lambda_2(\beta)$  is increasing on  $\mathbb{R}_+$ , with  $\lambda_2(+\infty) = -\log(\mathbb{P}(\omega(t, x) = s))$ . Thus, the  $L^2$  condition holds for all  $\beta \geq 0$  as soon as  $\mathbb{P}(\omega(t, x) = s) > \pi_d$ .*

*Proof.* Let  $q$  be the law of  $\omega(t, x)$ . By the [Theorem 1.6](#), it's enough to prove that

$$\lim_{\beta \rightarrow \infty} \lambda_2(\beta) = \begin{cases} \infty & s = \infty \\ -\log(q(\{s\})) & s < \infty \end{cases}$$

The claim is clear for  $s = \infty$ , so we consider the case  $s < \infty$ . To prove this case, observe that when  $q(\{s\}) > 0$ , it's enough to prove

$$\lambda(\beta) = \beta s + \log(q(\{s\})) + \epsilon(\beta) \tag{1.1.4}$$

for some  $\epsilon(\beta) \xrightarrow{\beta \rightarrow \infty} 0$ . This implies that  $\lambda_2(\beta) \xrightarrow{\beta \rightarrow \infty} -\log(q(\{s\}))$ . To prove (1.1.4), we write for  $h > 0$

$$\begin{aligned} \mathbb{P}(e^{\beta\omega} : \omega = s) &:= \mathbb{P}(e^{\beta\omega} 1_{\omega=s}) \leq \mathbb{P}(e^{\beta\omega}) = \mathbb{P}(e^{\beta\omega} : \omega \in [s-h, s]) + \mathbb{P}(e^{\beta\omega} : \omega \leq s-h) \\ &= e^{\beta s} (q([s-h, s]) + e^{-\beta h} q((-\infty, s-h])) \\ &\leq e^{\beta s} q(\{s\}) (q([s-h, s]) + e^{-\beta h}) / q(\{s\}) \end{aligned}$$

Taking logarithms, we deduce

$$\beta s + \log(q(\{s\})) \leq \lambda(\beta) \leq \beta s + \log(q(\{s\})) + \log\left(\frac{q([s-h, s]) + e^{-\beta h}}{q(\{s\})}\right)$$

Finally observe that  $\inf \left\{ \log \left( \frac{q([s-h, s]) + e^{-\beta h}}{q(\{s\})} \right) : h > 0 \right\}$  decreases to 1 when  $\beta \rightarrow +\infty$   $\square$

**Example 1.8.** If  $\omega \sim N(0, 1)$ , then  $\lambda(\beta) = \beta^2/2$ . Therefore,  $\lambda_2(\beta) = \beta^2$ . Hence, the  $L^2$  condition holds if  $\beta < \sqrt{\log(1/\pi_d)}$

**Definition 1.9** ( $L^2$  region). We call  $L^2$  region to the set of  $\beta$ 's such that the  $L^2$  condition (1.1.2) holds.

## 1.2 Diffusive Behavior in $L^2$ Region

In this section we assume  $d \geq 3, \beta$  in the  $L^2$  region. The next theorem states that in this region, the environment does not change the transversal fluctuations of the polymer for large  $d$  and small  $\beta$ .

**Theorem 1.10.** Under the assumptions of Theorem 1.6, we have

$$\lim_{n \rightarrow \infty} P_n^{\beta, \omega}(|S_n|^2)/n = 1 \quad \mathbb{P} \text{ a.s.} \quad (1.2.1)$$

and for all  $f \in C(\mathbb{R}^d)$  with at most polynomial growth at infinity

$$\lim_{n \rightarrow \infty} P_n^{\beta, \omega}[f(S_n/\sqrt{n})] = (2\pi)^{-d/2} \int_{\mathbb{R}^d} f(x/\sqrt{d}) \exp(-|x^2|/2) dx \quad \mathbb{P} \text{ a.s.} \quad (1.2.2)$$

In particular, if  $Z \sim N_d(0, d^{-1}I_d)$ , we have

$$P_n^{\beta, \omega}(S_n/\sqrt{n} \in \cdot) \rightarrow P(Z \in \cdot) \quad \mathbb{P} \text{ a.s.}$$

Before giving a proof of this theorem, we will define a family of martingales  $(M_n)_{n \geq 1}$  on  $(\Omega, \mathcal{F}, \mathbb{P})$  of the form

$$M_n = P(\phi(n, S_n) \bar{\xi}_n) \quad (1.2.3)$$

Where we recall that  $\bar{\xi}_n = e^{\beta H_n(x) - n\lambda(\beta)}$  and  $\phi : \mathbb{N} \times \mathbb{Z}^d \rightarrow \mathbb{R}$  is a function for which we assume the following:

(P1) There are constants  $C_i, p \in \mathbb{N}, i = 0, 1, 2$  such that

$$|\phi(n, x)| \leq C_0 + C_1|x|^p + C_2n^{p/2} \text{ for all } (n, x) \in \mathbb{N} \times \mathbb{Z}^d \quad (1.2.4)$$

(P2)  $\Phi_n := \phi(n, S_n)$  is a martingale on  $(\Omega_{\text{traj}}, \mathcal{G}, P)$  with respect to the filtration  $\mathcal{G}_n = \sigma(S_j : j \leq n)$

Note that  $M_n$  is a  $(\mathcal{F}_n)$  martingale on  $(\Omega, \mathcal{F}, \mathbb{P})$ . In effect, if  $\mathbb{P}^{\mathcal{F}_n}, P^{\mathcal{G}_n}$  denote the conditional expectations, we have

$$\begin{aligned}\mathbb{P}^{\mathcal{F}_n} M_{n+1} &= P[\phi(n+1, S_{n+1}) \bar{\xi}_n \mathbb{P}^{\mathcal{F}_n} e^{\beta \omega(n+1, S_{n+1}) - \lambda}] = P[\phi(n+1, S_{n+1}) \bar{\xi}_n] \\ &= P[\bar{\xi}_n P^{\mathcal{G}_n} \phi(n+1, S_{n+1})] = P[\bar{\xi}_n \phi(n, S_n)] = M_n\end{aligned}\tag{1.2.5}$$

We will need the following proposition before proving [Theorem 1.10](#).

**Proposition 1.11.** *Consider the martingale  $(M_n)_{n \geq 1}$  defined in [1.2.3](#). Suppose that  $d \geq 3$  the  $L^2$  condition [\(1.1.2\)](#), (P1) and (P2) are satisfied. Then there exists  $\kappa \in [0, p/2)$  such that*

$$\max_{0 \leq j \leq n} |M_j| = O(n^\kappa) \quad \text{as } n \rightarrow \infty \tag{1.2.6}$$

If in addition,  $p < \frac{1}{2}d - 1$ , then

$$\lim_{n \rightarrow \infty} M_n \text{ exists } \mathbb{P} \text{ a.s and in } L^2(\mathbb{P}) \tag{1.2.7}$$

*Sketch of the proof:* As is shown in [\[4\]](#), we have

$$\mathbb{P}(M_n^2) = O(b_n), \quad b_n = \sum_{j=1}^n j^{p-d/2} \tag{1.2.8}$$

Let  $M_n^* := \max_{0 \leq j \leq n} |M_j|$ . Note that  $\sqrt{b_n} = O(n^{p/2-d/4+1/2})$ . If  $d \geq 3$ , then  $\frac{1}{2} - \frac{d}{4} \leq -\frac{1}{4}$ . Thus,  $b_n = O(n^{p/2-1/4})$ . Therefore, to prove the first part of the proposition, is enough to show that if  $\delta > 0$  is small enough, then

$$M_n^* = O(n^\delta \sqrt{b_n}) \tag{1.2.9}$$

when  $n \rightarrow \infty, \mathbb{P}$  a.s. Because the monotonicity of  $M_n^*$  and the polynomial growth of  $n^\delta \sqrt{b_n}$ , it's enough to prove [\(1.2.9\)](#) along a subsequence  $\{n^k : n \in \mathbb{N}\}$  with  $k \geq 2$ . If  $k > 1/\delta$ , using the Doob's inequality for martingales, we have

$$\mathbb{P}(M_{n^k}^* > n^{k\delta} \sqrt{b_{n^k}}) \leq \mathbb{P}(M_{n^k}^* > n \sqrt{b_{n^k}}) \leq \mathbb{P}[(M_{n^k}^*)^2]/(n^2 b_{n^k}) \leq 4\mathbb{P}[(M_{n^k})^2]/(n^2 b_{n^k}) \leq Cn^{-2}$$

Where the last inequality comes from [\(1.2.8\)](#). Using Borel-Cantelli, we have that

$$\mathbb{P}(M_{n^k}^* \leq n^{k\delta} \sqrt{b_{n^k}} \text{ for large enough } n\text{'s}) = 1$$

That concludes the first part of the proposition. The second part comes from the Martingale Convergence Theorem.

□

*Proof of Theorem 1.10.* We the first part. Take  $\phi(n, x) := |x|^2 - n$ . In that case,  $p = 2$ . By Proposition 1.11, there exists  $\kappa \in [0, 1)$  such that

$$\begin{aligned} P_n^{\beta, \omega}[|S_n|^2] - n &= \frac{1}{Z_n} \sum_x |x|^2 e^{\beta H_n(x)} P(S_n = x) - n = \frac{1}{W_n} \sum_x |x|^2 \bar{\xi}_n(x) P(S_n = x) - n \\ &= P[|S_n|^2 \bar{\xi}_n] / W_n - n = P[\phi(S_n, n) \bar{\xi}_n] / W_n = O(n^\kappa) \end{aligned}$$

Dividing by  $n$  and taking  $n \rightarrow \infty$ , we conclude the first part.

Now we sketch the second part. We will use the standard multi-index notation, that is, if  $a = (a_j)_{j=1}^d$ , then  $|a|_1 = a_1 + \dots + a_d$ ,  $x^a = x_1^{a_1} \dots x_d^{a_d}$ ,  $(\frac{\partial}{\partial x})^a = (\frac{\partial}{\partial x_1})^{a_1} \dots (\frac{\partial}{\partial x_d})^{a_d}$  for  $x \in \mathbb{R}^d$ .

It's enough to prove the second part if  $f(x) = x^a$ . We prove this by induction in  $|a|_1$ . Let define

$$\begin{aligned} \phi(n, x) &:= \left( \frac{\partial}{\partial \theta} \right)^a \exp(\theta \cdot x - n\rho(\theta))|_{\theta=0} \\ \psi(n, x) &:= \left( \frac{\partial}{\partial \theta} \right)^a \exp(\theta \cdot x - n\frac{|\theta|^2}{2d})|_{\theta=0} \end{aligned}$$

where  $\rho(\theta) := \log(\frac{1}{d} \sum_{j=1}^d \cosh(\theta_j))$ .

We have the following related to these functions:

1.  $\phi$  satisfies (P1) – (P2) with  $p = |a|_1$

2.

$$(2\pi)^{-d/2} \int_{\mathbb{R}^d} \psi(1, x/\sqrt{d}) e^{-|x|^2/2} dx = 0 \quad (1.2.10)$$

Following [1], he proves that

$$\begin{aligned} \phi(n, x) &= x^a + \phi_0(n, x) \\ \psi(n, x) &= x^a + \psi_0(n, x) \end{aligned} \quad (1.2.11)$$

where

$$\begin{aligned} \phi_0(n, x) &= \sum_{\substack{|b|_1 + 2j \leq |a|_1 \\ j \geq 1}} A_a(b, j) x^b n^j \\ \psi_0(n, x) &= \sum_{\substack{|b|_1 + 2j = |a|_1 \\ j \geq 1}} A_a(b, j) x^b n^j \end{aligned}$$



for some  $A_a(b, j) \in \mathbb{R}$ . In particular,  $\phi_0, \psi_0$  have the same coefficients for  $x^b n^j$  with  $|b|_1 + 2j = |a|_1$ . Thus, we can write

$$(x/\sqrt{n})^a = \phi(n, x)n^{-|a|_1/2} - \psi_0(1, x/\sqrt{n}) + [\psi_0(n, x) - \psi_0(n, x)]n^{-|a|_1/2}$$

because  $(x/\sqrt{n})^a = x^a n^{-|a|_1/2}$  and  $\psi_0(1, x/\sqrt{n}) = \psi_0(n, x)n^{-|a|_1/2}$ . Therefore, we have

$$P_n^{\beta, \omega}[(S_n/\sqrt{n})^a] = \frac{1}{W_n} P[\phi(n, S_n)\bar{\xi}_n]n^{-|a|_1/2} - \frac{1}{W_n} P[\psi_0(1, S_n/\sqrt{n})\bar{\xi}_n] + \frac{1}{W_n} [\psi_0(n, x) - \psi_0(n, x)]n^{-|a|_1/2}$$

The second term converges to  $(2\pi)^{-d/2} \int_{\mathbb{R}^d} (x/\sqrt{d})^a e^{-|x|^2/2} dx$  by (1.2.10), (1.2.11) and the induction hypothesis. The first term converges to zero by Theorem 1.6 and Proposition 1.11. A similar argument applies to the third term.  $\square$

### 1.3 Local Limit Theorem in the $L^2$ region.

In this section we will consider the point to point partition function

$$W_n(y) := P(e^{\beta H_n(S) - n\lambda(\beta)} | S_n = y) \quad , y \in \mathbb{Z}^d \quad (1.3.1)$$

That is, we fix the last point to  $(n, y)$ . To make a distinction, we usually call point to level partition function to  $W_n$ .

Let be  $n \in \mathbb{N} \setminus \{0\}$ ,  $x \in \mathbb{Z}^d$  such that  $P(S_n = x) > 0$ , so  $|x|_1 \leq n$  and  $|x|_1 \equiv n \pmod{2}$ . We write

$$x \leftrightarrow n \Leftrightarrow P(S_n = x) > 0 \quad (1.3.2)$$

In that case,

$$\mathbb{P}(W_n(x)) = P[\mathbb{P}(e^{\beta H_n(S) - n\lambda(\beta)} | S_n = y)] = P[\mathbb{P}(e^{\beta H_n(S) - n\lambda(\beta)})] = 1$$

and we have

$$W_n = \sum_y W_n(y) P(S_n = y)$$

Define the reflection operator  $\theta_{n,x}^{\leftarrow}(\omega) = \omega'$ , with  $\omega'(u, y) := \omega(n - u, x + y)$ . In this section we will write for  $0 \leq k \leq n$

$$\bar{\xi}_{k,n}(x) = e^{\beta \sum_{i=k}^n \omega(i, x_i) - (n-k+1)\lambda(\beta)}$$

Note that  $\bar{\xi}_{1,n} = \bar{\xi}_n$ . We also write

$$\bar{\xi}_{k,n}^{\leftarrow} = e^{\beta \sum_{i=k-1}^{n-1} \omega(n-i, x_i) - (n-k+1)\lambda(\beta)}$$

Note that, because the  $\omega$ 's are i.i.d then  $P_y(\bar{\xi}_{1,n}^{\leftarrow}) \stackrel{law}{=} W_n$ . Also, the random variable  $W_\infty \circ \theta_{n,x}^{\leftarrow}$  is well defined for all  $n, x$ , thus we have

$$P_y(\bar{\xi}_{1,n}^{\leftarrow}) - W_\infty \circ \theta_{n,x}^{\leftarrow} \rightarrow 0$$

in probability, and in  $L^2$  in the  $L^2$  region.

Now we state the Local Limit Theorem:

**Theorem 1.12** (Local Limit Theorem). *Assume (1.1.2). Then for all  $A < \infty$  and all sequence of integers  $l_n \rightarrow \infty$  with  $l_n = o(n^\alpha)$  for some  $\alpha < 1/2$ , we have*

$$P[\bar{\xi}_{1,n}|S_n = y] = P(\bar{\xi}_{1,l_n})P_y(\bar{\xi}_{1,l_n}^{\leftarrow}) + \delta_n^y \quad (1.3.3)$$

with

$$\lim_{n \rightarrow \infty} \sup\{\mathbb{P}(|\delta_n^y|^2) : |y| \leq An^{1/2}, n \leftrightarrow y\} = 0$$

Moreover,

$$W_n(y) = W_\infty \times (W_\infty \circ \theta_{n,y}^{\leftarrow}) + \epsilon_n^y \quad (1.3.4)$$

Where the error term  $\epsilon_n^y \rightarrow 0$  uniformly on  $\{y : |y| \leq An^{1/2}, n \leftrightarrow y\}$

**Remark 1.13.**

1. Note that (1.3.3) can be written as  $W_n(y) = W_{l_n} \times P_y(\bar{\xi}_{1,l_n}^{\leftarrow}) + \delta_n^y$
2. The result says that  $P_n^{\beta,\omega}(S_n = y) \simeq W_\infty \circ \theta_{n,y}^{\leftarrow} \times P(S_n = y)$ . The interpretation is that the polymer measure is close to a Gaussian measure, up to a factor that depend of the endpoint.
3. Intuition: The polymer only is affected by the environment in the endpoints, and behaves like a Gaussian in the middle

*Sketch of the proof of Theorem 1.12.* The details are found in [6]

**Step 1:** With  $l \leq n/2$ , we approximate  $W_n(y)$  with  $P[\bar{\xi}_{1,l_n} \bar{\xi}_{n-l_n,n}]$  in  $L^2$ :

$$\lim_{n \rightarrow \infty} \sup_{|y| \leq An^{1/2}} \|W_n(y) - P[\bar{\xi}_{1,l_n} \bar{\xi}_{n-l_n,n}]\|_2^2 = 0$$

**Step 2:** Using the standard local limit theorem for random walks, we deduce that

$$\lim_{n \rightarrow \infty} \|P[\bar{\xi}_{1,l_n} \bar{\xi}_{n-l_n,n}] - P_y(\bar{\xi}_{1,l_n}^{\leftarrow})\|_1 = 0 \quad (1.3.5)$$

□

## 1.4 Rate of Martingale Convergence

First we define two modes of convergence in distribution. Let  $(Y_n)$  be a sequence of random variables defined in a common probability space  $(\Omega, \mathcal{F}, P)$ , such that  $Y_n$  converges in distribution to  $Y$ .

1. The convergence is stable if for all  $B \in \mathcal{F}$  with  $P(B) > 0$ , the conditional law of  $Y_n$  given  $B$  converges to some probability distribution depending on  $B$ .
2. The convergence is called mixing if it's stable, and the limit of conditional laws does not depend on  $B$  (and therefore, is the law of  $Y$ ).

The stable convergence says that for a random variable  $Z$  defined on  $(\Omega, \mathcal{F}, P)$ ,  $(Y_n, Z)$  converges in law to some coupling of  $Y, Z$  in an extended space. The mixing convergence says that  $Y_n$  is asymptotically independent of all event  $A \in \mathcal{F}$  (Recall the mixing in Ergodic Theory). more precisely, if  $G_n$  is the left hand of (1.4.2), then  $P(\{G_n \leq y\} \cap A) \rightarrow P(G \leq y)P(A)$  for  $A \in \mathcal{F}$

We now state the theorem, found in [3]:

**Theorem 1.14.** *For  $d \geq 3$ , there exists some  $\beta_0 > 0$  such that for  $|\beta| < \beta_0$ :*

$$n^{\frac{d-2}{4}}(W_\infty - W_n) \Rightarrow \sigma_1 W_\infty G \text{ in distribution} \quad (1.4.1)$$

and

$$n^{\frac{d-2}{4}} \frac{(W_\infty - W_n)}{W_n} \Rightarrow \sigma_1 G \text{ in distribution} \quad (1.4.2)$$

where

$$\sigma_1^2 := \frac{d^{d/2}(1 - \pi_d)}{2^{d/2}(d - 2)\pi^{d/2}\pi_d} \times \text{Var}(W_\infty) \quad (1.4.3)$$

where  $G \sim N(0, 1)$ , which is independent of  $W_\infty$ . Moreover, the convergence in (1.4.1) is stable, and the convergence in (1.4.2) is mixing

The theorem is based on a Central Limit Theorem for infinite martingale arrays:

**Theorem 1.15.** [3] For  $n \geq 1$ , let  $\{(S_{n,i}, \mathcal{F}_{n,i}) : i \geq 0\}$  be a martingale defined on a probability space  $(\Omega, \mathcal{F}, P)$ , with  $S_{n,0} = 0$  and

$$\sup_{n,i \geq 1} P(S_{n,i}^2) < \infty$$

Let  $X_{n,i} = S_{n,i} - S_{n,i-1}$ ,  $i \geq 1$  be the martingale differences, and  $S_{n,\infty} = \lim_{i \rightarrow \infty} S_{n,i}$  be the a.s limit of  $(S_{n,i} : i \geq 0)$ . Suppose that:

1. The conditional variance converges in probability: for a real random variable  $V \in [0, \infty)$ ,

$$V_{n,\infty}^2 := \sum_{i=1}^{\infty} P(X_{n,i}^2 | \mathcal{F}_{n,i-1}) \rightarrow V^2 \text{ in probability.}$$

2. The conditional Lindeberg condition holds:

$$\forall \epsilon > 0, \sum_{i=1}^{\infty} P(X_{n,i}^2 1_{|X_{n,i}| > \epsilon} | \mathcal{F}_{n,i-1}) \rightarrow 0 \text{ in probability.}$$

3.  $\mathcal{F}_{n,i} \subset \mathcal{F}_{n,i+1}$  for all  $n, i \geq 1$ .

Then

$$S_{n,\infty} \rightarrow VG \text{ in distribution} \quad (1.4.4)$$

where  $G \sim N(0, 1)$  and independent of  $V$ . If additionally  $V \neq 0$  a.s, then

$$\frac{S_{n,\infty}}{V_{n,\infty}} \rightarrow G \text{ in distribution} \quad (1.4.5)$$

Moreover, the convergence in (1.4.4) is stable, and the convergence in (1.4.5) is mixing.

*Sketch of the proof of Theorem 1.14.* We write  $n^{\frac{d-2}{4}}(W_\infty - W_n) = n^{\frac{d-2}{4}} \sum_{k=n}^{\infty} D_{k+1}$ , where  $D_{k+1} := W_{k+1} - W_k$ ,  $k \geq n$  forms a sequence of martingale differences with respect to the sequence of filtrations  $\mathcal{F}^{(n)} = (\mathcal{F}_i^{(n)})_{i \geq 0}$ , where  $\mathcal{F}_i^{(n)} = \sigma(W_j : j \leq i + n)$ . Using the last theorem with  $X_{n,i} = X_{n+i} \mathcal{F}_{n,i} = \mathcal{F}_{n+i}$ , is enough to show that

1.

$$s_n^2 := n^{\frac{d-1}{1}} \sum_{k \geq n} \mathbb{P}[D_{k+1}^2 | \mathcal{F}_k] \rightarrow \sigma_1^2 W_\infty^2 \quad (1.4.6)$$

in probability

2. The following Lindeberg condition holds:

$$\forall \epsilon > 0, n^{\frac{d-1}{2}} \sum_{k \geq n} \mathbb{P}[D_{k+1}^2 1_{n^{\frac{d-1}{4}} |D_{k+1}| > \epsilon} | \mathcal{F}_k] \rightarrow 0 \quad (1.4.7)$$

in probability

So here,  $V_{n,\infty} = \sigma_1 W_n, V = \sigma_1 W_\infty \neq 0$  because we are in the  $L^2$  region. Also, note that  $W_n/W_\infty \rightarrow 1$  in probability, so we can change  $W_\infty$  by  $W_{n,\infty}$  in (1.4.3).

To prove (1.4.6), it's shown in [3] that there exists  $\beta_0 > 0$  such that for  $|\beta| < |\beta_0|$  we have, as  $n \rightarrow \infty$ :

$$\mathbb{P}[(W_n - W_\infty)^4] \rightarrow 0 \quad (1.4.8)$$

$$\mathbb{P}(s_n^4) - \sigma_1^4 \mathbb{P}(W_n^4) \rightarrow 0 \quad (1.4.9)$$

$$\mathbb{P}(s_n^2 W_n^2) - \sigma_1^2 \mathbb{P}(W_n^4) \rightarrow 0 \quad (1.4.10)$$

Using the three equations we deduce that

$$\mathbb{P}[(s_n^2 - \sigma_1^2 W_n^2)^2] = \mathbb{P}(s_n^4) - 2\sigma_1^2 \mathbb{P}(s_n^2 W_n^2) + \sigma_1^4 \mathbb{P}(W_n^4) \rightarrow 0 \quad (1.4.11)$$

Thus,  $s_n^2 - \sigma_1^2 W_n^2 \rightarrow 0$  in  $L^2$ . As  $W_n^2 \rightarrow W_\infty^2$  in  $L^2$ , it follows that  $s_n^2 \rightarrow \sigma_1^2 W_\infty^2$  in  $L^2$ .

To check (1.4.7), is verified for  $q > 1$ , when  $|\beta| > 0$  is small enough, we have

$$\mathbb{P}(D_{k+1}^4) = O(k^{-d/q}), \quad k \geq 1$$

In that case, when  $\beta$  is small,  $n^{d-2} \sum_{k \geq n} \mathbb{P}(D_{k+1}^4) \rightarrow 0$ . This implies that  $n^{d-2} \sum_{k \geq n} \mathbb{P}(D_{k+1}^4 | \mathcal{F}_k) \rightarrow 0$ , and at the same, implies (1.4.7) by Cauchy Schwartz.  $\square$

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