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Directed Polymers in Random Environment

Chapter 3

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Abstract

These notes are based on the book "Directed Polymers in Random Environment" [2], and are intended for self-study and understand better these topics.

1 Martingale Approach and the L^2 region.

In this section we will study the partition function Z_n . Recall that $\lambda(\beta) = \log \mathbb{P}(e^{\beta\omega})$, so $\mathbb{P}(e^{\beta\omega}) = e^{\lambda(\beta)}$. Therefore,

 $\mathbb{P}(Z_n) = \mathbb{P}(P[e^{\beta H_n}]) = P(\mathbb{P}(exp\{\sum_{i=1}^n \omega(i, S_i)\})) = P(e^{n\lambda(\beta)}) = e^{n\lambda(\beta)}$, where we used the fact that the environment is i.i.d under \mathbb{P} . Because this calculation, we will consider the normalized partition function

$$W_n := \frac{Z_n}{e^{n\lambda(\beta)}}, n \ge 1 \tag{1.0.1}$$

Now, fix a path x, and consider $\overline{\xi_n} = \overline{\xi_n}(x) := e^{\beta H_n(x) - n\lambda(\beta)}$. Then, for that path, $H_n(x)$ is a random walk. In particular, ξ_n is a martingale on $(\Omega, \mathcal{F}, \mathbb{P})$ respect to the filtration $(\mathcal{F}_n)_{n \in \mathbb{N}}$ given by $\mathcal{F}_n := \sigma(\omega(j, x) : 1 \le j \le n, x \in \mathbb{Z}^d)$. Then, if x_1, \dots, x_k are paths and a_1, \dots, a_k are scalars, $\sum_{i=1}^k a_k \overline{\xi_n}(x_i)$ is also a martingale. Because $P(\overline{\xi_n}) = \sum_{x \in \mathbb{Z}^d} P(x) \overline{\xi_n}(x)$ then $W_n = P(\overline{\xi_n})$ also is a positive martingale.

Recall now the martingale convergence theorem, that basically says that if a martingale has L^1 norm uniformly bounded, then the limit exists \mathbb{P} almost surely. In our case, the limit $\lim_{n\to\infty} W_n$ will be called W_{∞} . A natural question is ask if $W_{\infty} = 0$ or not. In fact, note first that by

$$Z_{n+m} = Z_n \times P_n^{\beta,\omega}(Z_m \circ \theta_{n,S_n})$$
(1.0.2)

we can write

$$W_{n+m} = P(\overline{\xi_n} \times W_m \circ \theta_{n,S_n})$$

Taking $m \to \infty$ we deduce

$$\begin{split} W_{\infty} &= \lim_{m \to \infty} W_m = P(\overline{\xi_n} \times \lim_{m \to \infty} W_m \circ \theta_{n,S_n}) = P(\overline{\xi_n} \times W_{\infty} \circ \theta_{n,S_n}) \\ &= \sum_x P(S_n = x) \overline{\xi_n}(x) \times W_{\infty} \circ \theta_{n,x} = \sum_x P(S_n = x) e^{\beta H_n(x) - n\lambda(\beta)} \times W_{\infty} \circ \theta_{n,x} \\ &= Z_n e^{-n\lambda(\beta)} \sum_x \frac{1}{Z_n} (S_n = x) e^{\beta H_n(x)} \times W_{\infty} \circ \theta_{n,x} = W_n \times \sum_x P_n^{\beta,\omega}(S_n = x) \times W_{\infty} \circ \theta_{n,x} \end{split}$$

Because $W_n > 0$ for all n,we have

$$\{W_{\infty} = 0\} = \bigcap_{x \in \mathbb{Z}^d : P(S_n = x)} \{W_{\infty} \circ \theta_{n,x} = 0\} \in \mathcal{F}_n$$

Thus, the event $\{W_{\infty} = 0\}$ belongs to the tail σ algebra, and we can use the Kolmogorov 0-1 law to deduce the following theorem:

Theorem 1.1. The limit

$$W_{\infty} = \lim_{n \to \infty} W_n$$

exists $\mathbb{P}-$ a.s. Moreover, the limit is strictly positive, or is zero \mathbb{P} almost surely.

It's natural introduce some terminology that distinguish this dichotomy:

Definition 1.2. We say that the polymer is in the weak disorder when $W_{\infty} > 0$ – a.s and in the strong disorder if $W_{\infty} = 0$ – a.s

This definition is an analogous of the high/low temperature defined in the last section. In fact, there will be a similar statement related to a critical value.

Remark 1.3. We can prove that $W_{\infty} > 0 \Rightarrow p(\beta) = \lambda(\beta)$. Indeed, we have a.s

$$p(\beta) = \lim_{n \to \infty} \frac{1}{n} \log(Z_n(\beta, \omega)) = \lambda(\beta) + \lim_{n \to \infty} \frac{1}{n} \log(W_n) = \lambda(\beta)$$

because $\log(W_n) \to \log(W_\infty) < \infty$

Now we state the analogous of Theorem 1.19, chapter 2:

Proposition 1.4. There exists $\overline{\beta}_c = \overline{\beta}_c(\mathbb{P}, d) \in [0, \infty]$ such that

$$\begin{cases} W_{\infty} > 0 & \beta \in (0, \overline{\beta}_c) \\ W_{\infty} = 0 & \beta > \overline{\beta}_c \end{cases}$$

Proof. Let $\delta \in (0,1)$ arbitrary. The idea is prove the result for $(W_n^{\delta})_n$ for every δ . Note that since $\mathbb{P}(W_n) = 1$, then $(W_n^{\delta})_n$ is uniformly integrable (e.g. Hölder). We have then the a.s convergence of $W_n^{\delta} \to W_{\infty}^{\delta}$ and $\mathbb{P}(W_n^{\delta}) \to \mathbb{P}(W_{\infty}^{\delta})$, which is 0 in the strong disorder case, and strictly positive in the weak one. Is enough to prove that the map $\beta \to \mathbb{P}(W_n^{\delta})$ in non-increasing on \mathbb{R}_+ . This implies that the map $\beta \to \mathbb{P}(W_{\infty}^{\delta})$ also is non-increasing on \mathbb{R}_+ . To finish, define $\overline{\beta}_c := \inf\{\beta \geq 0 : \mathbb{P}(W_n^{\delta}) = 0\} \in [0, \infty]$. We compute

$$\frac{d}{d\beta}\mathbb{P}(W_n^{\delta}) = \mathbb{P}\left(\frac{d}{d\beta}W_n^{\delta}\right) = \delta\mathbb{P}\left(W_n^{\delta-1}\frac{d}{d\beta}Z_ne^{-n\lambda}\right) = \delta\mathbb{P}\left(W_n^{\delta-1}\left[e^{-n\lambda}\frac{d}{d\beta}Z_n - n\lambda'e^{-n\lambda}Z_n\right]\right)$$

$$= \delta\mathbb{P}\left(W_n^{\delta-1}\left[e^{-n\lambda}P(H_n\overline{\xi_n}) - n\lambda'P(\overline{\xi_n})\right]\right) = \delta\mathbb{P}(W_n^{\delta-1}P[\{H_n - n\lambda'\}\overline{\xi_n}])$$

$$= \delta P(\mathbb{P}[\overline{\xi_n}W_n^{\delta-1}\{H_n - n\lambda'\}])$$
(1.0.3)

Now, we apply the FGK-Harris inequality as follows: we consider the measure $\overline{\xi}_n d\mathbb{P}$ for fixed path x, the decreasing function (on ω) $W_n^{\delta-1}$ and the non-decreasing one $H_n - n\lambda'$. Then we obtain

$$\frac{d}{d\beta}\mathbb{P}(W_n^{\delta}) = \delta P(\mathbb{P}[\overline{\xi}_n W_n^{\delta-1}\{H_n - n\lambda'\}]) \le \delta P(\mathbb{P}[\overline{\xi}_n W_n^{\delta-1}]\mathbb{P}[\overline{\xi}_n\{H_n - n\lambda'\}]) = 0$$

The last equality is by

$$\mathbb{P}[\omega e^{\beta \omega}] = \lambda'(\beta) e^{\lambda(\beta)} \tag{1.0.4}$$

Because the last theorem and Theorem 1.19,chapter 2,it's expected that $\beta_c = \overline{\beta}_c$, but at the moment this question remains open. However, if $d \in \{1, 2\}$, the equality is satisfied. Note that always $\overline{\beta}_c \leq \beta_c$. In effect, if $\beta = \beta_c + \epsilon$ with $\epsilon > 0$, then $\beta \geq \overline{\beta}_c$, otherwise $W_{\infty}(\beta) > 0 \Rightarrow p(\beta) = \lambda(\beta)$ (by Remark 1.3) $\Rightarrow \beta \leq \beta_c$, a contradiction.

Question 1. $\beta_c = \overline{\beta}_c$?

Question 2. What happens at $\overline{\beta}_c$?

It's expected that $W_{\infty}(\overline{\beta}_c) = 0$

1.1 The second moment method and the L^2 region

We recall some knows facts about the return probability of the simple random walk, denoted by π_d . More precisely,

$$\pi_d := P(S_n = 0 \text{ for some } n \ge 1) \tag{1.1.1}$$

Then $\pi_d = 1$ if $d \in \{1, 2\}$ and < 1 for $d \ge 3$.In fact, $\pi_d < 1$ if $d \ge 3$.For more information,see [5] Now we define the important L2 condition:

Definition 1.5 (L2 condition). Given $d \geq 3$, we say that β satisfies the L2 condition if

$$\lambda_2(\beta) := \lambda(2\beta) - 2\lambda(\beta) < \log(1/\pi_d) \tag{1.1.2}$$

Note that $\lambda_2(0) = 0$ and $\lambda'_2(\beta) = 2(\lambda'(2\beta) - \lambda'(\beta)) > 0$, since λ' is increasing. Therefore, $\lambda_2(\beta)$ is increasing in \mathbb{R}_+ , and if $d \geq 3$, the L2 condition es equivalent to $\beta < \beta_{L^2}$, with

$$\beta_{L^2} := \inf\{\beta \ge 0 : \lambda_2 \le \log(1/\pi_d)\}$$
 (1.1.3)

Theorem 1.6. Suppose that $d \geq 3$ and the L2 condition is satisfied. Then, $W_{\infty} > 0$ a.s. In particular, $p = \lambda$ if $\beta \leq \beta_{L^2}$

Proof. Let's consider the product space $(\Omega^2, \mathcal{F}^{\otimes 2})$ and the probability measure $P^{\otimes 2}(dx, d\tilde{x})$. This is the distribution of two independent walks S, \tilde{S} . Note that

$$W_n^2 = P \frac{Z_n^2}{e^{2n\lambda(\beta)}} = P(e^{\beta H_n(S)}) P(e^{\beta H_n(\tilde{S})}) e^{-2n\lambda(\beta)} = P^{\otimes 2} (e^{\beta [H_n(S) - H_n(\tilde{S})] - 2n\lambda(\beta)})$$

Now we compute $\mathbb{P}(W_n^2)$:

$$\mathbb{P}(W_n^2) = \stackrel{Fubini}{=} P^{\otimes 2} \mathbb{P} \left[\prod_{t=1}^n e^{\beta[\omega(t,S_t) + \omega(t,\tilde{S}_t)] - 2\lambda(\beta)} \right] = P^{\otimes 2} \left[\prod_{t=1}^n \mathbb{P} \left(e^{\beta[\omega(t,S_t) + \omega(t,\tilde{S}_t)] - 2\lambda(\beta)} (1_{S_t = \tilde{S}_t} + 1_{S_t \neq \tilde{S}_t}) \right) \right]$$

$$= P^{\otimes 2} \left[\prod_{t=1}^n e^{\lambda(2\beta)] - 2\lambda(\beta)} (1_{S_t = \tilde{S}_t}) + 1_{S_t \neq \tilde{S}_t} \right]$$

$$= P^{\otimes 2} \left[\prod_{t=1}^n e^{\lambda_2(\beta) 1} S_t = \tilde{S}_t} \right] = P^{\otimes 2} \left[e^{\lambda_2(\beta) N_n} \right]$$

where $N_n = N_n(S, \tilde{S}) := \sum_{t=1}^n 1_{S_t = \tilde{S}_t}$ is the number of intersections of S, \tilde{S} up to time n. Taking $n \to \infty$, we have $N_n \to \infty$, and by the Monotone Convergence Theorem,

$$\mathbb{P}(W_n^2) \to P^{\otimes 2}[e^{\lambda_2(\beta)N_\infty}]$$

Also note that if $X_n = S_n - \tilde{S}_n$, then N_{∞} is the number of visits to zero of the symmetric random walk X_n (this walk is not a nearest-neighbor one). Then N_{∞} is geometrically distributed with success probability π_d . Then, we have

$$P^{\otimes 2}[e^{\lambda_2(\beta)N_{\infty}}] = \sum_{k=0}^{\infty} (1 - \pi_d) \pi_d^k e^{k\lambda_2} = \begin{cases} \frac{1 - \pi_d}{1 - \pi_d e^{\lambda_s}} & \pi_d e^{\lambda_2} < 1\\ +\infty & \pi_d e^{\lambda_2} \ge 1 \end{cases}$$

Therefore, $\sup_n \mathbb{P}(W_n^2) < \infty \Leftrightarrow \lambda_2 + \log(\pi_d) < 0 \Leftrightarrow \lambda_2 < \log(1/\pi_d)$. In that case, the martingale W_n is bounded in L^2 . This implies not only the convergence to W_∞ , but also convergence in L^1 . Thus, $\mathbb{P}(W_\infty) = 1$ and necessarily W_∞ must be strictly positive.

Corollary 1.7. Let $s = ess \sup_{\mathbb{P}} \omega(t, x)$. The function $\lambda_2(\beta)$ is increasing on \mathbb{R}_+ , with $\lambda_2(+\infty) = -\log(\mathbb{P}(\omega(t, x) = s))$. Thus, the L2 condition holds for all $\beta \geq 0$ as soon as $\mathbb{P}(\omega(t, x) = s) > \pi_d$.

Proof. Let q be the law of $\omega(t,x)$. By the Theorem 1.6, it's enough to prove that

$$\lim_{\beta \to \infty} \lambda_2(\beta) = \begin{cases} \infty & s = \infty \\ -\log(q(\{s\})) & s < \infty \end{cases}$$

The claim is clear for $s = \infty$, so we consider the case $s < \infty$. To prove this case, observe that when $q(\{s\}) > 0$, it' enough to prove

$$\lambda(\beta) = \beta s + \log(q(\{s\})) + \epsilon(\beta) \tag{1.1.4}$$

for some $\epsilon(\beta) \stackrel{\beta \to \infty}{\to} 0$. This implies that $\lambda_2(\beta) \stackrel{\beta \to \infty}{\to} -\log(q(\{s\}))$. To prove (1.1.4), we write for h > 0

$$\mathbb{P}(e^{\beta\omega} : \omega = s) := \mathbb{P}(e^{\beta\omega}1_{\omega=s}) \le \mathbb{P}(e^{\beta\omega}) = \mathbb{P}(e^{\beta\omega} : \omega \in [s - h, s]) + \mathbb{P}(e^{\beta\omega} : \omega \le s - h)
= e^{\beta s}(q([s - h, s]) + e^{-\beta h}q((-\infty, s - h])
\le e^{\beta s}q(\{s\})(q([s - h, s]) + e^{-\beta h})/q(\{s\})$$

Taking logarithms, we deduce

$$\beta s + \log(q(\{s\})) \le \lambda(\beta) \le \beta s + \log(q(\{s\})) + \log\left(\frac{q([s-h,s]) + e^{-\beta h}}{q(\{s\})}\right)$$

Finally observe that $\inf \left\{ \log \left(\frac{q([s-h,s]) + e^{-\beta h}}{q(\{s\})} \right) : h > 0 \right\}$ decreases to 1 when $\beta \to +\infty$

Example 1.8. If $\omega \sim N(0,1)$, then $\lambda(\beta) = \beta^2/2$. Therefore, $\lambda_2(\beta) = \beta^2$. Hence, the L2 condition holds if $\beta < \sqrt{\log(1/\pi_d)}$

Definition 1.9 (L^2 region). We call L^2 region to the set of β 's such that the L^2 condition (1.1.2) holds.

1.2 Diffusive Behavior in L^2 Region

In this section we assume $d \geq 3$, β in the L^2 region. The next theorem states that in this region, the environment does not change the transversal fluctuations of the polymer for large d and small β .

Theorem 1.10. Under the assumptions of Theorem 1.6, we have

$$\lim_{n \to \infty} P_n^{\beta,\omega}(|S_n|^2)/n = 1 \mathbb{P} a.s \tag{1.2.1}$$

and for all $f \in C(\mathbb{R}^d)$ with at most polynomial growth at infinity

$$\lim_{n \to \infty} P_n^{\beta,\omega}[f(S_n \sqrt{n})] = (2\pi)^{-d/2} \int_{\mathbb{R}^d} f(x \sqrt{d}) exp(-|x^2|/2) dx \ \mathbb{P} \ a.s \tag{1.2.2}$$

In particular, if $Z \sim N_d(0, d^{-1}I_d)$, we have

$$P_n^{\beta,\omega}(S_n \sqrt{n} \in \cdot) \to P(Z \in \cdot) \mathbb{P} a.s$$

Before giving a proof of this theorem, we will define a family of martingales $(M_n)_{n\geq 1}$ on $(\Omega, \mathcal{F}, \mathbb{P})$ of the form

$$M_n = P(\phi(n, S_n)\overline{\xi}_n) \tag{1.2.3}$$

Where we recall that $\overline{\xi}_n = e^{\beta H_n(x) - n\lambda(\beta)}$ and $\phi : \mathbb{N} \times \mathbb{Z}^d \to \mathbb{R}$ is a function for which we assume the following:

(P1) There are constants $C_i, p \in \mathbb{N}, i = 0, 1, 2$ such that

$$|\phi(n,x)| \le C_0 + C_1 |x|^p + C_2 n^{p/2} \text{ for all } (n,x) \in \mathbb{N} \times \mathbb{Z}^d$$
 (1.2.4)

(P2) $\Phi_n := \phi(n, S_n)$ is a martingale on $(\Omega_{traj}, \mathcal{G}, P)$ with respect to the filtration $\mathcal{G}_n = \sigma(S_j : j \le n)$

Note that M_n is a (\mathcal{F}_n) martingale on $(\Omega, \mathcal{F}, \mathbb{P})$. In effect, if $\mathbb{P}^{\mathcal{F}_n}$, $P^{\mathcal{G}_n}$ denote the conditional expectations, we have

$$\mathbb{P}^{\mathcal{F}_n} M_{n+1} = P[\phi(n+1, S_{n+1})\overline{\xi}_n \mathbb{P}^{\mathcal{F}_n} e^{\beta \omega(n+1, S_{n+1}) - \lambda}] = P[\phi(n+1, S_{n+1})\overline{\xi}_n]$$

$$= P[\overline{\xi}_n P^{\mathcal{G}_n} \phi(n+1, S_{n+1})] = P[\overline{\xi}_n \phi(n, S_n)] = M_n$$
(1.2.5)

We will need the following proposition before proving Theorem 1.10.

Proposition 1.11. Consider the martingale $(M_n)_{n\geq 1}$ defined in 1.2.3. Suppose that $d\geq 3$ the L2 condition (1.1.2), (P1) and (P2) are satisfied. Then, there exists $\kappa \in [0, p/2)$ such that

$$\max_{0 \le j \le n} |M_j| = O(n^{\kappa}) \text{ as } n \to \infty$$
 (1.2.6)

If in addition, $p < \frac{1}{2}d - 1$, then

$$\lim_{n \to \infty} M_n \text{ exists } \mathbb{P} \text{ a.s and in } L^2(\mathbb{P})$$
 (1.2.7)

Sketch of the proof: As is shown in [4], we have

$$\mathbb{P}(M_n^2) = O(b_n), \quad b_n = \sum_{j=1}^n j^{p-d/2}$$
(1.2.8)

Let $M_n^* := \max_{0 \le j \le n} |M_j|$. Note that $\sqrt{b_n} = O(n^{p/2-d/4+1/2})$. If $d \ge 3$, then $\frac{1}{2} - \frac{d}{4} \le -\frac{1}{4}$. Thus, $b_n = O(n^{p/2-1/4})$. Therefore, to prove the first part of the proposition, is enough to show that if $\delta > 0$ is small enough, then

$$M_n^* = O(n^{\delta} \sqrt{b_n}) \tag{1.2.9}$$

when $n \to \infty$, \mathbb{P} a.s. Because the monotonicity of M_n^* and the polynomial growth of $n^{\delta}\sqrt{b_n}$, it's enough to prove (1.2.9) along a subsequence $\{n^k : n \in \mathbb{N}\}$ with $k \ge 2$. If $k > 1/\delta$, using the Doob's inequality for martingales, we have

$$\mathbb{P}(M_{n^k}^* > n^{k\delta} \sqrt{b_{n^k}}) \le \mathbb{P}(M_{n^k}^* > n \sqrt{b_{n^k}}) \le \mathbb{P}[(M_{n^k}^*)^2] / (n^2 b_{n^k}) \le 4\mathbb{P}[(M_{n^k})^2] / (n^2 b_{n^k}) \le Cn^{-2}$$

Where the last inequality comes from (1.2.8). Using Borel-Cantelli, we have that

$$\mathbb{P}(M_{n^k}^* \leq n^{k\delta} \sqrt{b_{n^k}} \text{ for large enough n's}) = 1$$

That concludes the first part of the proposition. The second part comes from the Martingale Convergence Theorem.

Proof of Theorem 1.10. We the first part. Take $\phi(n,x) := |x|^2 - n$. In that case, p = 2. By Proposition 1.11, there exists $\kappa \in [0,1)$ such that

$$P_n^{\beta,\omega}[|S_n|^2] - n = \frac{1}{Z_n} \sum_x |x|^2 e^{\beta H_n(x)} P(S_n = x) - n = \frac{1}{W_n} \sum_x |x|^2 \overline{\xi}_n(x) P(S_n = x) - n$$
$$= P[|S_n|^2 \overline{\xi}_n] / W_n - n = P[\phi(S_n, n) \overline{\xi}_n] / W_n = O(n^{\kappa})$$

Dividing by n and taking $n \to \infty$, we conclude the first part.

Now we sketch the second part. We will use the standard multi-index notation, that is, if $a=(a_j)_{j=1}^d$, then $|a|_1=a_1+\cdots a_d, x^a=x_1^{a_1}\cdots x_d^{a_d}, (\frac{\partial}{\partial x})^a=(\frac{\partial}{\partial x_1})^{a_1}\cdots (\frac{\partial}{\partial x_d})^{a_d}$ for $x\in\mathbb{R}^d$. It's enough to prove the second part if $f(x)=x^a$. We prove this by induction in $|a|_1$. Let define

$$\phi(n,x) := \left(\frac{\partial}{\partial \theta}\right)^a exp(\theta \cdot x - n\rho(\theta))|_{\theta=0}$$
$$\psi(n,x) := \left(\frac{\partial}{\partial \theta}\right)^a exp(\theta \cdot x - n\frac{|\theta|^2}{2d})|_{\theta=0}$$

where $\rho(\theta) := \log(\frac{1}{d} \sum_{j=1}^{d} \cosh(\theta_j))$.

We have the following related to these functions:

1.
$$\phi$$
 satisfies $(P1) - (P2)$ with $p = |a|_1$

2.

$$(2\pi)^{-d/2} \int_{\mathbb{R}^d} \psi(1, x / \sqrt{d}) e^{-|x^2|/2} dx = 0$$
 (1.2.10)

Following [1], he proves that

$$\phi(n,x) = x^{a} + \phi_{0}(n,x)$$

$$\psi(n,x) = x^{a} + \psi_{0}(n,x)$$
(1.2.11)

where

$$\phi_0(n,x) = \sum_{\substack{|b|_1 + 2j \le |a|_1 \\ j \ge 1}} A_a(b,j) x^b n^j$$

$$\psi_0(n,x) = \sum_{\substack{|b|_1 + 2j = |a|_1 \\ j > 1}} A_a(b,j) x^b n^j$$

for some $A_a(b,j) \in \mathbb{R}$. In particular, ϕ_0, ψ_0 have the same coefficients for $x^b n^j$ with $|b|_1 + 2j = |a|_1$. Thus, we can write

$$(x\sqrt{n})^a = \phi(n,x)n^{-|a_1|/2} - \psi_0(1,x\sqrt{n}) + [\psi_0(n,x) - \psi_0(n,x)]n^{-|a_1|/2}$$

because $(x\sqrt{n})^a = x^a n^{-|a|_1/2}$ and $\psi_0(1, x\sqrt{n}) = \psi_0(n, x) n^{-|a|_1/2}$. Therefore, we have

$$P_n^{\beta,\omega}[(S_n\sqrt{n})^a] = \frac{1}{W_n}P[\phi(n,S_n)\overline{\xi}_n]n^{-|a_1|/2} - \frac{1}{W_n}P[\psi_0(1,S_n\sqrt{n})\overline{\xi}_n] + \frac{1}{W_n}[\psi_0(n,x) - \psi_0(n,x)]n^{-|a_1|/2}$$

The second term converges to $(2\pi)^{-d/2} \int_{\mathbb{R}^d} (x/\sqrt{d})^a e^{-|x|^2/2} dx$ by (1.2.10),(1.2.11) and the induction hypothesis. The first term converges to zero by Theorem 1.6 and Proposition 1.11. A similar argument applies to the third term.

1.3 Local Limit Theorem in the L^2 region.

In this section we will consider the point to point partition function

$$W_n(y) := P(e^{\beta H_n(S) - n\lambda(\beta)} | S_n = y) \quad , y \in \mathbb{Z}^d$$

$$\tag{1.3.1}$$

That is, we fix the last point to (n, y). To make a distinction, we usually call point to level partition function to W_n .

Let be $n \in \mathbb{N} \setminus \{0\}, x \in \mathbb{Z}^d$ such that $P(S_n = x) > 0$, so $|x|_1 \le n$ and $|x|_1 \equiv n \pmod 2$. We write

$$x \leftrightarrow n \Leftrightarrow P(S_n = x) > 0 \tag{1.3.2}$$

In that case,

$$\mathbb{P}(W_n(x)) = P[\mathbb{P}(e^{\beta H_n(S) - n\lambda(\beta)} | S_n = y)] = P[\mathbb{P}(e^{\beta H_n(S) - n\lambda(\beta)})] = 1$$

and we have

$$W_n = \sum_{y} W_n(y) P(S_n = y)$$

Define the reflection operator $\theta_{n,x}^{\leftarrow}(\omega) = \omega'$, with $\omega'(u,y) := \omega(n-u,x+y)$ In this section we will write for $0 \le k \le n$

$$\overline{\xi}_{k,n}(x) = e^{\beta \sum_{i=k}^{n} \omega(i,x_i) - (n-k+1)\lambda(\beta)}$$

Note that $\overline{\xi}_{1,n} = \overline{\xi}_n$. We also write

$$\overline{\xi}_{k,n}^{\leftarrow} = e^{\beta \sum_{i=k-1}^{n-1} \omega(n-i,x_i) - (n-k+1)\lambda(\beta)}$$

Note that, because the $\omega's$ are i.i.d then $P_y(\overline{\xi}_{1,n}^{\leftarrow}) \stackrel{law}{=} W_n$. Also, the random variable $W_{\infty} \circ \theta_{n,x}^{\leftarrow}$ is well defined for all n, x, thus we have

$$P_y(\overline{\xi}_{1,n}^{\leftarrow}) - W_{\infty} \circ \theta_{n,x}^{\leftarrow} \to 0$$

in probability, and in L^2 in the L^2 region.

Now we state the Local Limit Theorem:

Theorem 1.12 (Local Limit Theorem). Assume (1.1.2). Then, for all $A < \infty$ and all sequence of integers $l_n \to \infty$ with $l_n = o(n^{\alpha})$ for some $\alpha < 1/2$, we have

$$P[\overline{\xi}_{1,n}|S_n = y] = P(\overline{\xi}_{1,l_n})P_y(\overline{\xi}_{1,l_n}) + \delta_n^y$$
(1.3.3)

with

$$\lim_{n \to \infty} \sup \{ \mathbb{P}(|\delta_n^y|^2) : |y| \le An^{1/2}, n \leftrightarrow y \} = 0$$

Moreover,

$$W_n(y) = W_{\infty} \times (W_{\infty} \circ \theta_{n,y}^{\leftarrow}) + \epsilon_n^y \tag{1.3.4}$$

Where the error term $\epsilon_n^y \to 0$ uniformly on $\{y : |y| \le An^{1/2}, n \leftrightarrow y\}$

Remark 1.13.

- 1. Note that (1.3.3) can be written as $W_n(y) = W_{l_n} \times P_y(\overline{\xi}_{1,l_n}) + \delta_n^y$
- 2. The result says that $P_n^{\beta,\omega}(S_n=y) \simeq W_\infty \circ \theta_{n,y}^\leftarrow \times P(S_n=y)$. The interpretation is that the polymer measure is close to a gaussian measure, up to a factor that depend of the endpoint.
- 3. Intuition: The polymer only is affected by the environment in the endpoints, and behaves like a gaussian in the middle

Sketch of the proof of Theorem 1.12. The details are found in [6]

Step 1: With $l \leq n/2$, we approximate $W_n(y)$ with $P[\overline{\xi}_{1,l_n}\overline{\xi}_{n-l_n,n}]$ in L^2 :

$$\lim_{n \to \infty} \sup_{|y| \le An^{1/2}} ||W_n(y) - P[\overline{\xi}_{1,l_n} \overline{\xi}_{n-l_n,n}]||_2^2 = 0$$

Step 2: Using the standard local limit theorem for random walks, we deduce that

$$\lim_{n \to \infty} ||P[\overline{\xi}_{1,l_n} \overline{\xi}_{n-l_n,n}] - P_y(\overline{\xi}_{1,l_n})||_1 = 0$$
 (1.3.5)

1.4 Rate of Martingale Convergence

First we define two modes of convergence in distribution.Let (Y_n) be a sequence of random variables defined in a common probability space (Ω, \mathcal{F}, P) , such that Y_n converges in distribution to Y.

- 1. The convergence is stable if for all $B \in \mathcal{F}$ with P(B) > 0, the conditional law of Y_n given B converges to some probability distribution depending on B.
- 2. The convergence is called mixing if it's stable, and the limit of conditional laws does not depend on B (and therefore, is the law of Y).

The stable convergence says that for a random variable Z defined on $(\Omega, \mathcal{F}, P), (Y_n, Z)$ converges in law to some coupling of Y, Z in an extended space. The mixing convergence says that Y_n is asymptotically independent of all event $A \in \mathcal{F}(\text{Recall the mixing in Ergodic Theory})$.more precisely, if G_n is the left hand of (1.4.2), then $P(\{G_n \leq y\} \cap A) \to P(G \leq y)P(A)$ for $A \in \mathcal{F}$

We now state the theorem, found in [3]:

Theorem 1.14. For $d \ge 3$, there exists some $\beta_0 > 0$ such that for $|\beta| < \beta_0$:

$$n^{\frac{d-2}{4}}(W_{\infty} - W_n) \Rightarrow \sigma_1 W_{\infty} G$$
 in distribution (1.4.1)

and

$$n^{\frac{d-2}{4}} \frac{(W_{\infty} - W_n)}{W_n} \Rightarrow \sigma_1 G \text{ in distribution}$$
 (1.4.2)

where

$$\sigma_1^2 := \frac{d^{d/2}(1 - \pi_d)}{2^{d/2}(d - 2)\pi^{d/2}\pi_d} \times Var(W_\infty)$$
(1.4.3)

where $G \sim N(0,1)$, which is independent of W_{∞} . Moreover, the convergence in (1.4.1) is stable, and the convergence in (1.4.2) is mixing

The theorem is based on a Central Limit Theorem for infinite martingale arrays:

Theorem 1.15. [3] For $n \geq 1$, let $\{(S_{n,i}, \mathcal{F}_{n,i}) : i \geq 0\}$ be a martingale defined on a probability space (Ω, \mathcal{F}, P) , with $S_{n,0} = 0$ and

$$\sup_{n,i\geq 1} P(S_{n,i}^2) < \infty$$

Let $X_{n,i} = S_{n,i} - S_{n,i-1}, i \ge 1$ be the martingale differences, and $S_{n,\infty} = \lim_{i \to \infty} S_{n,i}$ be the a.s limit of $(S_{n,i} : i \ge 0)$. Suppose that:

1. The conditional variance converges in probability: for a real random variable $V \in [0, \infty)$,

$$V_{n,\infty}^2 := \sum_{i=1}^{\infty} P(X_{n,i}^2 | \mathcal{F}_{n,i-1}) \to V^2$$
 in probability.

2. The conditional Lindeberg condition holds:

$$\forall \epsilon > 0, \sum_{i=1}^{\infty} P(X_{n,i}^2 1_{|X_{n,i}| > \epsilon} \mathcal{F}_{n,i-1}) \to 0 \text{ in probability.}$$

3. $\mathcal{F}_{n,i} \subset \mathcal{F}_{n,i+1}$ for all $n, i \geq 1$.

Then,

$$S_{n,\infty} \to VG$$
 in distribution (1.4.4)

where $G \sim N(0,1)$ and independent of V.If additionally $V \neq 0$ a.s, then

$$\frac{S_{n,\infty}}{V_{n,\infty}} \to G \text{ in distribution}$$
 (1.4.5)

Moreover, the convergence in (1.4.4) is stable, and the convergence in (1.4.5) is mixing.

Sketch of the proof of Theorem 1.14. We write $n^{\frac{d-2}{4}}(W_{\infty} - W_n) = n^{\frac{d-2}{4}} \sum_{k=n}^{\infty} D_{k+1}$, where $D_{k+1} := W_{k+1} - W_k, k \geq n$ forms a sequence of martingale differences with respect t the sequence of filtrations $\mathcal{F}^{(n)} = (\mathcal{F}_i^{(n)})_{i\geq 0}$, where $\mathcal{F}_i^{(n)}) = \sigma(W_j : j \leq i+n)$. Using the last theorem with $X_{n,i} = X_{n+i}\mathcal{F}_{n,i} = \mathcal{F}_{n+i}$, is enough to show that

1.

$$s_n^2 := n^{\frac{d-1}{1}} \sum_{k \ge n} \mathbb{P}[D_{k+1}^2 | \mathcal{F}_k] \to \sigma_1^2 W_\infty^2$$
 (1.4.6)

in probability

2. The following Lindeberg condition holds:

$$\forall \epsilon > 0, n^{\frac{d-1}{2}} \sum_{k > n} \mathbb{P}[D_{k+1}^2 1_{n^{\frac{d-1}{4}} | D_{k+1}| > \epsilon} | \mathcal{F}_k] \to 0$$
 (1.4.7)

in probability

So here, $V_{n,\infty} = \sigma_1 W_n$, $V = \sigma_1 W_\infty \neq 0$ because we are in the L^2 region. Also, note that $W_n/W_\infty \to 1$ in probability, so we can change W_∞ by $W_{n,\infty}$ in (1.4.3).

To prove (1.4.6), it's shown in [3] that there exists $\beta_0 > 0$ such that for $|\beta| < |\beta_0|$ we have, as $n \to \infty$:

$$\mathbb{P}[(W_n - W_\infty)^4] \to 0 \tag{1.4.8}$$

$$\mathbb{P}(s_n^4) - \sigma_1^4 \mathbb{P}(W_n^4) \to 0 \tag{1.4.9}$$

$$\mathbb{P}(s_n^2 W_n^2) - \sigma_1^2 \mathbb{P}(W_n^4) \to 0 \tag{1.4.10}$$

Using the three equations we deduce that

$$\mathbb{P}[(s_n^2 - \sigma_1^2 W_n^2)^2] = \mathbb{P}(s_n^4) - 2\sigma_1^2 \mathbb{P}(s_n^2 W_n^2) + \sigma_1^4 \mathbb{P}(W_n^4) \to 0 \tag{1.4.11}$$

Thus, $s_n^2 - \sigma_1^2 W_n^2 \to 0$ in L^2 . As $W_n^2 \to W_\infty^2$ in L^2 , it follows that $s_n^2 \to \sigma_1^2 W_\infty^2$ in L^2 .

To check (1.4.7), is verified for q > 1, when $|\beta| > 0$ is small enough, we have

$$\mathbb{P}(D_{k+1}^4) = O(k^{-d/q}), \quad k \ge 1$$

In that case, when β is small, $n^{d-2} \sum_{k \geq n} \mathbb{P}(D_{k+1}^4) \to 0$. This implies that $n^{d-2} \sum_{k \geq n} \mathbb{P}(D_{k+1}^4 | \mathcal{F}_k) \to 0$, and at the same, implies (1.4.7) by Cauchy Schwarz.

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