

Reading: Chapter 4 of *Probability, Statistics, and Random Processes* by A. Leon-Garcia

1. Find the characteristic function of a normal distribution with mean m and variance σ^2 .

Hint: Let $k = m + j\omega\sigma^2$, then $j\omega x - \frac{(x-m)^2}{2\sigma^2} = \frac{-(x-k)^2 + 2mj\omega\sigma^2 - \omega^2\sigma^4}{2\sigma^2}$, $\int_{-\infty}^{\infty} e^{-x^2} dx = \sqrt{\pi}$.

Solution:

The characteristic function of a random variable X is defined as following:

$$\Phi_X(\omega) = E[e^{j\omega X}] = \int_{-\infty}^{\infty} f_X(x) e^{j\omega x} dx$$

Thus, the characteristic function of a normal distribution with mean m and variance σ^2 is:

$$\begin{aligned} \Phi_X(\omega) &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-m)^2}{2\sigma^2}} e^{j\omega x} dx \\ &= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^{j\omega x - \frac{(x-m)^2}{2\sigma^2}} dx \\ &= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^{\frac{-(x-k)^2 + 2mj\omega\sigma^2 - \omega^2\sigma^4}{2\sigma^2}} dx \quad (\text{from hint}) \\ &= e^{\frac{2mj\omega\sigma^2 - \omega^2\sigma^4}{2\sigma^2}} \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^{-\frac{(x-k)^2}{2\sigma^2}} dx \\ &= e^{mj\omega - \frac{1}{2}\omega^2\sigma^2} \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^{-\left(\frac{x-k}{\sqrt{2\sigma^2}}\right)^2} dx \\ &= e^{mj\omega - \frac{1}{2}\omega^2\sigma^2} \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^{-z^2} \sqrt{2\sigma^2} dz \\ &= e^{mj\omega - \frac{1}{2}\omega^2\sigma^2} \frac{1}{\sqrt{2\pi\sigma^2}} \sqrt{2\sigma^2} \int_{-\infty}^{\infty} e^{-z^2} dz \\ &= e^{mj\omega - \frac{1}{2}\omega^2\sigma^2} \frac{1}{\sqrt{2\pi\sigma^2}} \sqrt{2\pi\sigma^2} \quad (\text{from hint}) \\ &= e^{mj\omega - \frac{1}{2}\omega^2\sigma^2} \end{aligned}$$

2. Find the characteristic function of $Y = aX + b$ where X is a Gaussian random variable.

Solution:

The characteristic function of $Y = aX + b$ is given by

$$\begin{aligned} \Phi_Y(\omega) &= E[e^{j\omega Y}] \\ &= E[e^{j\omega(aX+b)}] \\ &= e^{j\omega b} E[e^{ja\omega X}] \\ &= e^{j\omega b} e^{jam\omega - a^2\sigma^2\omega^2/2} \end{aligned}$$

which is the characteristic function of Gaussian random variable with mean of $am + b$ and variance of $a^2\sigma^2$

3. Suppose Y is a Laplacian random variable, which is defined by the PDF:

$$f_Y(y) = he^{-4\alpha|y|}, \quad -\infty < y < \infty, \alpha > 0.$$

- (a) Determine the value of h as a function of α .

Solution:

$$f_Y(y) = he^{-4\alpha|y|} \begin{cases} he^{4\alpha y}, & -\infty < y < 0 \\ he^{-4\alpha y}, & 0 \leq y < \infty \end{cases}$$

$$\begin{aligned} \int_{-\infty}^{\infty} f_Y(y) dy &= 1 \\ \int_{-\infty}^0 he^{4\alpha y} dy + \int_0^{\infty} he^{-4\alpha y} dy &= 1 \\ 2 \int_0^{\infty} he^{-4\alpha y} dy &= 1 \\ 2h \frac{e^{-4\alpha y}}{-4\alpha} \Big|_0^{\infty} &= 1 \\ \frac{h}{2\alpha} &= 1 \end{aligned}$$

Thus, $h = 2\alpha$

- (b) Determine the CDF of the random variable Y .

Solution:

Region 1, $-\infty < y < 0$

$$F_Y(y) = \int_{-\infty}^y 2\alpha e^{4\alpha y} dy = 2\alpha \frac{e^{4\alpha y}}{4\alpha} \Big|_{-\infty}^y = \frac{1}{2} e^{4\alpha y}$$

Region 2, $0 \leq y < \infty$

$$F_Y(y) = \int_{-\infty}^0 2\alpha e^{4\alpha y} dy + \int_0^y 2\alpha e^{-4\alpha y} dy = \frac{1}{2} + 2\alpha \frac{e^{-4\alpha y}}{-4\alpha} \Big|_0^y = \frac{1}{2} + \frac{1}{2}(1 - e^{-4\alpha y}) = 1 - \frac{1}{2} e^{-4\alpha y}$$

4. Let X be a Laplacian random variable, which is defined by the PDF:

$$f_X(x) = \frac{\alpha}{2} e^{-\alpha|x|}, \quad -\infty < x < \infty, \alpha > 0.$$

- (a) Find the characteristic function of the Laplacian random variable.

Solution:

The characteristic function of a random variable X is defined as following:

$$\Phi_X(\omega) = E[e^{j\omega X}] = \int_{-\infty}^{\infty} f_X(x) e^{j\omega x} dx$$

The characteristic function of a Laplacian random variable X is defined as following:

$$\begin{aligned}
\Phi_X(\omega) &= E[e^{j\omega X}] \\
&= \int_{-\infty}^{\infty} \left(\frac{\alpha}{2} e^{-\alpha|x|} \right) e^{j\omega x} dx \\
&= \int_{-\infty}^{\infty} \left(\frac{\alpha}{2} e^{-\alpha|x|} \right) (\cos(\omega x) + j\sin(\omega x)) dx \\
&= \frac{\alpha}{2} \int_{-\infty}^{\infty} e^{-\alpha|x|} \cos(\omega x) dx + j \frac{\alpha}{2} \int_{-\infty}^{\infty} e^{-\alpha|x|} \sin(\omega x) dx
\end{aligned}$$

Since, $\sin(-\omega x) = -\sin(\omega x)$ and, $\cos(-\omega x) = \cos(\omega x)$. Thus,

$$\begin{aligned}
\Phi_X(\omega) &= \frac{\alpha}{2} \int_{-\infty}^{\infty} e^{-\alpha|x|} \cos(\omega x) dx + j \frac{\alpha}{2} \int_{-\infty}^{\infty} e^{-\alpha|x|} \sin(\omega x) dx \\
&= \frac{\alpha}{2} \int_{-\infty}^{\infty} e^{-\alpha|x|} \cos(\omega x) dx \\
&= \frac{\alpha}{2} \left[\int_{-\infty}^0 e^{\alpha x} \cos(\omega x) dx + \int_0^{\infty} e^{-\alpha x} \cos(\omega x) dx \right] \\
&= \frac{\alpha}{2} \left[\int_0^{\infty} e^{-\alpha x} \cos(-\omega x) dx + \int_0^{\infty} e^{-\alpha x} \cos(\omega x) dx \right] \\
&= 2 \frac{\alpha}{2} \int_0^{\infty} e^{-\alpha x} \cos(\omega x) dx \\
&= \alpha \frac{e^{-\alpha x} (\omega \sin(\omega x) - \alpha \cos(\omega x))}{\alpha^2 + \omega^2} \Big|_0^{\infty} \\
&= \frac{\alpha^2}{\alpha^2 + \omega^2}
\end{aligned}$$

(b) Find the mean and variance of X by applying the moment theorem.

Solution:

$$\begin{aligned}
E[X] &= \frac{1}{j} \frac{d}{d\omega} \Phi_X(\omega) \Big|_{\omega=0} \\
&= \frac{1}{j} \frac{d}{d\omega} \left(\frac{\alpha^2}{\alpha^2 + \omega^2} \right) \Big|_{\omega=0} \\
&= \frac{1}{j} \frac{2\alpha^2 \omega}{(\omega^2 + \alpha^2)^2} \Big|_{\omega=0} \\
&= 0
\end{aligned}$$

$$\begin{aligned}
Var[X] &= E[X^2] - E[X]^2 \\
&= E[X^2] - 0 \\
&= \frac{1}{j^2} \frac{d^2}{d\omega^2} \Phi_X(\omega) \Big|_{\omega=0} \\
&= \frac{1}{j^2} \frac{d^2}{d\omega^2} \left(\frac{\alpha^2}{\alpha^2 + \omega^2} \right) \Big|_{\omega=0} \\
&= \frac{1}{j^2} \left(\frac{8\alpha^2\omega^2}{(\omega^2 + \alpha^2)^3} - \frac{2\alpha^2}{(\omega^2 + \alpha^2)^2} \right) \Big|_{\omega=0} \\
&= \frac{2}{\alpha^2}
\end{aligned}$$

5. Compare the Chebyshev inequality and the exact probability for the event $\{|X - m| > c\}$ as a function of c for X is a continuous uniform random variable in the interval $[-b, b]$.

Solution:

Chebyshev inequality states that,

$$P\{|X - m| > c\} \leq \frac{\sigma^2}{c^2}$$

When X is a uniform random variable in the interval $[-b, b]$, $P[X \leq c] = \frac{c-a}{b-a}$ for c in $[-b, b]$. $E[X] = 0$ and $Var[X] = \frac{b^2}{3}$. For $0 \leq c \leq b$:

$$\begin{aligned}
P\{|X - m| > c\} &= P(X - m > c) + P(X - m < -c) \\
&= [1 - P(X - m < c)] + P(X - m < -c) \\
&= [1 - P(X < c + m)] + P(X < -c + m) \\
&= \left[1 - \frac{c+b}{b+b}\right] + \frac{-c+b}{b+b} \\
&= 1 - \frac{c}{b}
\end{aligned}$$

Plug in the respective values to the Chebyshev inequality:

$$P\{|X - m| > c\} \leq \frac{\sigma^2}{c^2}$$

When X is a uniform random variable in the interval $[-b, b]$, $P[X \leq c] = \frac{c-a}{b-a}$ for c in $[-b, b]$. $E[X] = 0$ and $Var[X] = \frac{b^2}{3}$. For $0 \leq c \leq b$:

$$\begin{aligned}
P\{|X - m| > c\} &\leq \frac{\sigma^2}{c^2} \\
1 - \frac{c}{b} &\leq \frac{b^2}{3} \times \frac{1}{c^2} \\
\frac{b-c}{b} &\leq \frac{b^2}{3c^2} \\
0 &\leq 3c^3 + b^3 - 3bc^2
\end{aligned}$$

The inequality always holds for $0 \leq c \leq b$.