A General Framework for Clustering and Distribution Matching with Bandit Feedback

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7 Oct. 2024

Group meeting

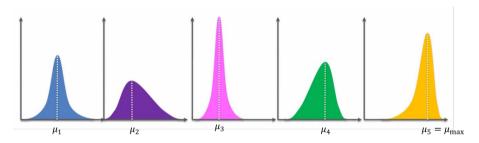
Joint work with Yuqi Huang, Vincent Tan, and Jonathan Scarlett

Outline

- Problem formulation
- Lower Bound
- Algorithm
- Theoretical and experimental results

Multi-armed bandits (MABs)

Motivation: A suitable mathematical model for applications such as drug design, online advertisement, and online recommendation systems



There are two main objectives:

- Regret minimization: Maximize the total reward after T pulls
- **2** Pure exploration: Answer a specific question about K unknown distributions e.g., best arm identification, odd arm identification, ϵ -good arm identification

Clustering and Distribution Matching Problem

Any pure exploration problem can be viewed as a sequential multi-hypothesis testing with bandit feedback [Prabhu et al. 2022].

- We are given K arms.
- Arm distributions are on a finite alphabet \mathcal{X} .
- ullet Each hypothesis σ is denoted by a partition of a subset of [K]

$$\sigma = \{\mathcal{A}_1^{\sigma}, \dots, \mathcal{A}_M^{\sigma}\}$$

where \mathcal{A}_{m}^{σ} 's are disjoint and $\bigcup_{m \in [M]} \mathcal{A}_{m}^{\sigma} \subseteq [K]$.

- For each $m \in [M]$, \mathcal{A}_m^{σ} indicates a cluster with identical distributions \implies arm $i, j \in \mathcal{A}_m^{\sigma}$, then $P_i = P_j$.
- $\mathcal{A}_{M+1}^{\sigma} \triangleq [K] \setminus \bigcup_{m \in [M]} \mathcal{A}_{m}^{\sigma}$ is called the unconstrained group, which doesn't restrict the arm distributions in it.
- We assume $|\mathcal{A}_m^{\sigma}| \geq 2$ for $m \leq M$. Hence, $K \geq 2M$.
- Arms from distinct subsets in $\{A_1^{\sigma}, \dots, A_{M+1}^{\sigma}\}$ follow distinct distributions.

Clustering and Distribution Matching Problem

Let $A_t \in [K]$ be the arm pulled at time t. Let $X_{t,A_t} \in \mathcal{X}$ be the reward at time t from arm A_t .

- We design an online algorithm: A_t may depend only on $(A_1, X_{1,A_1}, A_2, X_{2,A_2}, \dots, A_{t-1}, X_{t-1,A_{t-1}})$.
- Let $P = (P_1, \dots, P_K)$ be the underlying problem instance whose hypothesis is σ_P .
- Fixed confidence: Algorithm stops at a random time τ and outputs a hypothesis $\hat{\sigma}(\tau) \in \mathcal{C}$.

Goal: Design an online algorithm such that

- **①** δ -correct: $\mathbb{P}\left[\hat{\sigma}(\tau) \neq \sigma_P\right] \leq \delta$ and $\mathbb{P}\left[\tau < \infty\right] = 1$

Distinguishability of Hypotheses

A hypothesis σ is said to *dominate* another hypothesis σ' if every subset in the partitioning of σ , $\mathcal{A}^{\sigma}_{[M]}$, is a subset of some subset in the partitioning of σ' , $\mathcal{A}^{\sigma'}_{[M]}$.

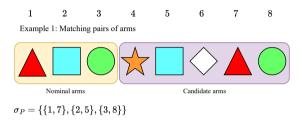
Consider two hypotheses $\sigma_1 = \{\{1,2\},\{4,5\}\}$ and $\sigma_2 = \{\{1,2,3\},\{4,5\}\}$ with M=2 and K=5. The equality relations implied by σ_1 ($P_1=P_2$ and $P_4=P_5$) are contained in the equality relations implied by σ_2 ($P_1=P_2$, $P_1=P_3$, and $P_4=P_5$). Hence, σ_1 dominates σ_2 (σ_1 is less stringent than σ_2).

Assumption: For a given clustering problem C,

- there exists no hypothesis pair $(\sigma, \sigma') \in \mathcal{C}^2$ for which $\sigma \neq \sigma'$ and σ dominates σ' ,
- ② for each problem instance $P \in \Lambda$, there exists a unique hypothesis $\sigma \in \mathcal{C}$ such that $P \in \Lambda_{\sigma}$.

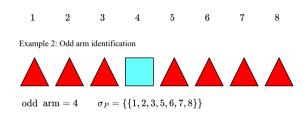
 Λ_{σ} denotes all instances assoc. with σ and Λ denotes all instances included in a given problem.

Example 1: Matching pairs



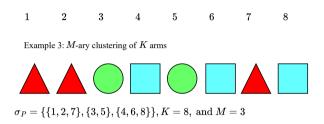
- Each nominal arm has exactly 1 match in the set of candidate arms.
- The designer knows which arms are nominal arms (here, $\{1,2,3\}$).
- Offline version of this problem is studied by [Zhou et al. 2024]. The offline is referred to as sequence matching.

Example 2: Odd arm identification



- Odd arm follows a different distribution than the rest.
- Studied by [Vaidhiyan and Sundaresan, 2018] and [Karthik and Sundaresan, 2020].

Example 3: M-ary clustering of K arms



- K arms are partitioned into M groups whose size can be as small as 1.
- Highest number of hypotheses for fixed K and M (appr. $M^K/M!$).
- Studied by [Yang et al., 2024] for *d*-dimensional Gaussian arms. Their algorithm utilizes *K*-means algorithm, which relies on the fact that the arms are Gaussian distributed.

Preliminary Definitions

Generalization of (Generalized Jensen-Shannon (GJS) divergence)

$$G(P_{\mathcal{A}}, w_{\mathcal{A}}) \triangleq egin{cases} 0 & ext{if } w_i = 0, \ orall i \in \mathcal{A} \ \sum_{i \in \mathcal{A}} w_i D(P_i \| W) & ext{otherwise} \end{cases}$$

where
$$W \triangleq \frac{\sum_{i \in \mathcal{A}} w_i P_i}{\sum_{i \in \mathcal{A}} w_i} \in \mathcal{P}(\mathcal{X}).$$

Score functions

$$g_{P}^{\sigma}(w) \triangleq \sum_{m=1}^{M} G(P_{\mathcal{A}_{m}^{\sigma}}, w_{\mathcal{A}_{m}^{\sigma}}) = \sum_{m=1}^{M} \inf_{Q \in \mathcal{P}(\mathcal{X})} \sum_{i \in \mathcal{A}_{m}^{\sigma}} w_{i} D(P_{i} \| Q)$$

$$G_{P}^{\sigma}(w) \triangleq \min_{\sigma' \in \mathcal{C} \setminus \{\sigma\}} g_{P}^{\sigma'}(w)$$

$$T(P, \sigma_{P}) \triangleq \max_{w \in \Sigma_{K}} G_{P}^{\sigma_{P}}(w) = \sup_{w \in \Sigma_{K}} \inf_{P' \in \text{Alt}(P)} \sum_{i \in [K]} w_{i} D(P_{i} \| P'_{i})$$

Intuition on G

Let $(X_i^{n_i})_{i \in [B]}$ be $B \ge 2$ collection of sequences of lengths n_1, \ldots, n_B from a finite alphabet \mathcal{X} . Let $N = \sum_{i=1}^B n_i$, $X^N = (X_1^{n_1}, \ldots, X_M^{n_B})$, and

$$H_0: X^N \sim P^N$$
 for some $P \in \mathcal{P}(\mathcal{X})$
 $H_1: X_i^{n_i} \sim P_i^{n_i}, i \in [B]$ for some $P_{[B]} \in \mathcal{P}^B(\mathcal{X})$

Lemma

Consider X^N and the hypotheses H_0 and H_1 . Let $w_i = \frac{n_i}{N}$ for $i \in [B]$. Denote $\hat{P}_{[B]} = (\hat{P}_{X^{n_i}})_{i \in [B]}$. Then,

$$G(\hat{P}_{[B]}, w_{[B]}) = \frac{1}{N} \log \frac{\max_{P_{[B]} \in \mathcal{P}^{B}(\mathcal{X})} \prod_{i=1}^{B} P_{i}^{n_{i}}(X_{i}^{n_{i}})}{\max_{P \in \mathcal{P}(\mathcal{X})} P^{N}(X^{N})}.$$
 (1)

Lower (Converse) Bound

Theorem

For any δ -correct algorithm π with $\delta \in (0,1)$ and any problem instance $P \in \Lambda$,

$$\mathbb{E}\left[\tau\right] \geq \frac{d(\delta||1-\delta)}{T^*(P)} \geq \frac{1}{T^*(P)} \log \frac{1}{2.4\delta} = \frac{1}{T^*(P)} \log \frac{1}{\delta} + \Theta(1) \tag{2}$$

where
$$T^*(P) = T(P, \sigma_P)$$
 and $d(p||q) = p \log \frac{p}{q} + (1-p) \log \frac{1-p}{1-q}$.

Proof: Apply the standard technique from [Garivier and Kaufmann, 2016]. It uses change of measure and Wald's identity.

Algorithm (TaS-FW)

- We want to design a computationally-efficient Track-and-Stop (TaS) algorithm.
- Ideally, a TaS algorithm computes

$$egin{aligned} w^*(t) &= rg \max_{w \in \Sigma_K} G_{\hat{P}(t-1)}^{\hat{\sigma}(t-1)}(w) \ &= rg \max_{w \in \Sigma_K} \min_{\sigma' \in \mathcal{C} \setminus \{\hat{\sigma}(t-1)\}} g_{\hat{P}(t-1)}^{\sigma'}(w) \end{aligned}$$

at each time t and matches its fraction of arm pulls to the oracle $w^*(t)$.

- But calculating this maximum can be very difficult in the general case (e.g., $|\mathcal{X}| \geq 3$).
- Hence, we linearize the objective function and utilize a modified version of the Frank-Wolfe algorithm (that is tailored to the non-smoothness coming from the minimum of functions).
- The algorithm is inspired by [Wang et al., 2022], which is for general pure exploration problems but does not apply to ours.

Frank-Wolfe Update

Instead of

$$w^*(t) = rg \max_{w \in \Sigma_K} \min_{\sigma' \in \mathcal{C} \setminus \{\hat{\sigma}(t-1)\}} g^{\sigma'}_{\hat{\mathcal{P}}(t-1)}(w)$$

we solve

$$z(t) = \underset{z \in \Sigma_{K}}{\operatorname{arg \, max}} \min_{\substack{h \in H_{G_{\hat{P}(t-1)}^{\hat{\sigma}(t-1)}(x(t-1), r_{t})}} \langle z - x(t-1), h \rangle$$
(3)

$$x(t) = \left(1 - \frac{1}{t}\right)x(t-1) + \frac{1}{t}z(t) = \frac{1}{t}\sum_{s=1}^{t}z(s)$$
 (4)

where the r-subdifferential subspace $H_{G_P^{\sigma}}(w,r)$

$$H_{G_P^{\sigma}}(w,r) \triangleq \operatorname{co}(\nabla g_P^{\sigma'}(w) \colon \sigma' \neq \sigma, g_P^{\sigma'}(w) < G_P^{\sigma}(w) + r)$$

accounts for the non-smoothness in the objective function.

The maximin in (3) is an LP.

Algorithm

Algorithm 1 TaS-FW

```
Input: Target error probability \delta \in (0,1), the collection of hypotheses \mathcal{C}
       Initialization: Sample each arm i \in [K] once, initialize \tilde{x}(K) = \frac{1}{K}\mathbf{1} and N(K) = (1, \dots, 1).
       For t \in \mathbb{N}, set r_t = t^{-4/5}.
       t \leftarrow K
  1: while Z(t) \triangleq t \, G_{\tilde{\rho}(t)}^{\hat{\sigma}(t)}(w(t)) < \beta(t,\delta) \triangleq \beta(t,\delta) = \log \frac{1}{\delta} + (M|\mathcal{X}| + \tilde{K} + 2) \log(t+1) + \log \left(\frac{\pi^2}{6} - 1\right) \, \mathbf{do}
  2: t \leftarrow t + 1
         if t \in \mathcal{I}_f \triangleq \{t \in \mathbb{N} : \lceil \sqrt{t} \log t \rceil = \lceil \sqrt{t+1} \log(t+1) \rceil - 1\} then
         \tilde{z}(t) \leftarrow \frac{1}{K} \mathbf{1} (Forced Exploration)
          else if t \notin \mathcal{I}_{\mathbf{f}} then
              \tilde{z}(t) \leftarrow \operatorname*{arg\,max}_{z \in \Sigma_K} \min_{h \in H_{G^{\frac{\theta}{2}(t-1)}}(\tilde{x}(t-1), r_t)} \langle z - \tilde{x}(t-1), h \rangle
                                                                                                                    (FW Update)
  7:
           end if
          \tilde{x}(t) \leftarrow \left(1 - \frac{1}{4}\right) \tilde{x}(t-1) + \frac{1}{4} \tilde{z}(t)
           Sample the arm A_t \leftarrow \arg\max(t\tilde{x}_i(t) - N_i(t-1))
                                                                                                             (C-tracking rule)
  9:
          Update N(t) \leftarrow N(t-1) + e_A, and the empirical problem instance \hat{P}(t) in (42)
          \hat{\sigma}(t) \leftarrow \operatorname*{arg\,min}_{\sigma \in \mathcal{C}} g^{\sigma}_{\hat{P}(t)}(w(t))
12: end while
Output: \hat{\sigma}(t)
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A Second-order Achievability Bound

Theorem

For any problem instance $P \in \Lambda$, as $\delta \to 0^+$, our algorithm TaS-FW is δ -correct and achieves

$$\mathbb{E}\left[\tau\right] \le \frac{\log\frac{1}{\delta}}{T^*(P)} \left(1 + O\left(\left(\log\frac{1}{\delta}\right)^{-1/4} \sqrt{\log\log\frac{1}{\delta}}\right)\right). \tag{5}$$

- The algorithm is first-order optimal as $\delta \to 0+$.
- We characterize an upper bound on the rate of convergence.
- The tightness of the second-order term is an interesting open problem.

Comparison with Existing Work

Context	Compared paper	Our algorithm	Compared algorithm	Reason
Algorithm based on	Wang et al.	Frank–Wolfe	Frank–Wolfe	computing sup-inf efficiently
Forced exploration	Wang et al., GK, Prabhu et al.	$\sqrt{t} \log t$	\sqrt{t}	$\dim(\Lambda) < \dim(\mathcal{P}^{\kappa}(\mathcal{X}))$
Efficiency	Prabhu et al.	Yes	No in general	Prabhu et al. do not provide an efficient algorithm to compute sup-inf
Approach	Yang et al.	FW + Seq. HT	K-means + simplification in the inner infimum	K-means doesn't work for finite alphabets + simplification isn't always possible
Second-order term	Wang et al., Prabhu et al., Yang et al., GK	Includes	Does not include	Refined analysis

Proof Steps

• C-tracking lemma:

$$\left|N_i(t) - \sum_{s=1}^t \tilde{z}_i(s)\right| \le K - 1 \tag{6}$$

- Establish the Lipschitzness of $g_P^{\sigma}(w)$ and $G_P^{\sigma}(w)$ in w and P.
- FW lemma: Let $\tilde{\Delta}_t \triangleq G_P(w^*) G_P(w(t))$ be the optimality gap. Under the event

$$\max_{z \in \Sigma_K} \min_{h \in H_{G_P}(\tilde{x}(t-1), r_t)} \langle z - \tilde{x}(t-1), h \rangle - \epsilon_t < \min_{h \in H_{G_P}(\tilde{x}(t-1), r_t)} \langle \tilde{z}(t) - \tilde{x}(t-1), h \rangle$$
 (7)

for $t \in \{T_1, \ldots, T_2\} \cap \mathcal{I}_{\mathrm{f}}^{\mathrm{c}}$,

$$\tilde{\Delta}_{T_2} \le \frac{T_1}{T_2} L + 2LT_2^{-1/2} \log T_2 + \frac{1}{T_2} \sum_{t=1}^{T_2} (r_t + \epsilon_t) + 32DKT_2^{-1/2} + \frac{L(K+3)}{T_2}.$$
 (8)

Technical Novelties

• Concentration of $G(\cdot)$:

$$\mathbb{P}\left[N\sum_{m=1}^{M}G(\hat{P}_{\mathcal{A}_{m}},w_{\mathcal{A}_{m}})\geq\beta\right]\leq(N+1)^{M|\mathcal{X}|}\exp\{-\beta\}\tag{9}$$

which is used to prove $\mathbb{P}\left[\hat{\sigma}(\tau) \neq \sigma_P\right] \leq \delta$.

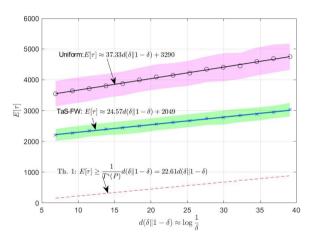
• Establishing a sufficient condition for ϵ_t such that

$$\|\hat{P}(t) - P\|_{\infty} \le \epsilon_t$$
 $\hat{\sigma}(t) = \sigma_P \iff \min_{\sigma' \ne \sigma_P} g_{\hat{P}(t)}^{\sigma'}(w(t)) > g_{\hat{P}(t)}^{\sigma_P}(w(t))$

Because the condition $\hat{\sigma}(t) = \sigma_P$ depends on w(t), we choose a more aggressive forced exploration $(\sqrt{t} \log t \text{ times instead of } \sqrt{t} \text{ times})$

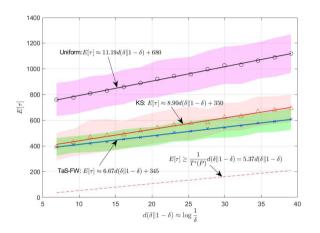
• Carefully choose $\epsilon_t = \Theta(t^{-1/4}\sqrt{\log t})$, T_1 , and T_2 to optimize the second-order term in the upper bound.

Experiment 1



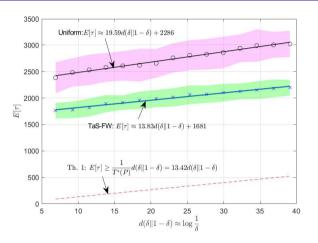
Matching pairs: K = 6, M = 2 with $P_1 = P_3 = (0.1, 0.1, 0.8)$, $P_2 = P_4 = (0.4, 0.4, 0.2)$, $P_5 = (0.5, 0.05, 0.45)$, and $P_6 = (0.1, 0.8, 0.1)$. True hypothesis is $\sigma_P = \{\{1, 3\}, \{2, 4\}\}$.

Experiment 2



Odd arm identification: K = 7, M = 1 with $P_i = (0.1, 0.1, 0.8)$ for $i \in [6]$, and $P_7 = (0.6, 0.2, 0.2)$. True hypothesis is $\sigma_P = \{\{1, \dots, 6\}\}$.

Experiment 3



M-ary clustering of *K* arms: K = 6, M = 3 with $P = (P_1, ..., P_7)$, where $P_1 = P_2 = (0.6, 0.2, 0.2)$, $P_3 = P_4 = (0.25, 0.7, 0.05)$, and $P_5 = P_6 = (0.05, 0.05, 0.90)$. True hypothesis is $\sigma_P = \{\{1, 2\}, \{3, 4\}, \{5, 6\}\}$.

Takeaways

- We develop a generalized framework for PE problems that involve clustering. Our algorithm works for problems such as odd arm id. in [Karthik] and *M*-ary clustering in [Yang et al.] simultaneously.
- ② Our refined algorithm is first-order optimal as δ approaches 0, and the achievability bound includes a second-order term, which is derived by optimizing the design parameters.
- We consider distributions on a finite alphabet size, which are a special case of a vector exponential family. Using existing tools in the literature, the result can be extended to one-parameter exponential families.

Thank you for listening to me!

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