

Math of Big Data, Summer 2018

Prof: Gu

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HW #: 1

Day: Mon. Tue. Wed. Thu. Fri.

Date: 05/15/2018

No.	Points	Acknowledgments
1		Tim Player, Jacky Lee
2		
Total		

This Assignment is (check one):



On Time



Late, without deduction



Late, with deduction

Comments: Feel free to work with other students, but make sure you write up the homework and code on your own (no copying homework *or* code; no pair programming). Feel free to ask students or instructors for help debugging code or whatever else, though.

The starter code for problem 2 part c and d can be found under the Resource tab on course website.

Note: You need to create a Github account for submission of the coding part of the homework. Please create a repository on Github to hold all your code and include your Github account username as part of the answer to problem 2.

Problem 1. (Linear Transformation)

Let $\mathbf{y} = A\mathbf{x} + \mathbf{b}$ be a random vector. Show that expectation is linear:

$$\mathbb{E}[\mathbf{y}] = \mathbb{E}[A\mathbf{x} + \mathbf{b}] = A\mathbb{E}[\mathbf{x}] + \mathbf{b}.$$

Also show that

$$\text{cov}[\mathbf{y}] = \text{cov}[A\mathbf{x} + \mathbf{b}] = A \text{cov}[\mathbf{x}] A^\top = A\Sigma A^\top.$$

Solution:

(a) First, we prove some small lemmas.

Lemma 1.1. *Let X be a continuous random variable, and a be a scalar. Then $\mathbb{E}[aX] = a\mathbb{E}[X]$.*

Proof. Let X admit a density function $f(x)$. Then

$$\begin{aligned} \mathbb{E}[aX] &= \int_{-\infty}^{\infty} axf(x) \, dx \\ &= a \int_{-\infty}^{\infty} xf(x) \, dx \\ &= a\mathbb{E}[X] \end{aligned}$$

Lemma 1.2. *Let X be a random variable, and let a be a scalar. Then $\mathbb{E}[X + a] = \mathbb{E}[X] + \mathbb{E}[a] = \mathbb{E}[X] + a$.* ■

Proof. Let X admit a density function $f(x)$. Then over \mathbb{R} , $f(x)$ must integrate to 1. Hence

$$\begin{aligned} \int_{-\infty}^{\infty} (x + a)f(x) \, dx &= \int_{-\infty}^{\infty} xf(x) + af(x) \, dx \\ &= \int_{-\infty}^{\infty} xf(x) \, dx + \int_{-\infty}^{\infty} af(x) \, dx \\ &= \int_{-\infty}^{\infty} xf(x) \, dx + a \int_{-\infty}^{\infty} f(x) \, dx \\ &= \mathbb{E}[X] + a \cdot 1 \\ &= \mathbb{E}[X] + a \end{aligned}$$

Lemma 1.3. *Let X and Y be continuous random variables. Then $\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y]$.* ■

Proof. Let $X + Y$ have a density function $f(x, y)$. Then

$$\begin{aligned} \mathbb{E}[X + Y] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x + y)f(x, y) \, dx \, dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xf(x, y) \, dx \, dy + \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} yf(x, y) \, dx \, dy \end{aligned}$$

$$\begin{aligned} \mathbb{E}[\mathbf{y}] &= \mathbb{E}[A\mathbf{x} + \mathbf{b}] \\ &= \mathbb{E}[A\mathbf{x}] + \mathbb{E}[\mathbf{b}] \end{aligned}$$

$$= \mathbb{E}[A\mathbf{x}] + \mathbf{b}$$

(we got from the second line to the third by the fact that $\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y]$ for random variables X and Y , and from the second to the third by the fact that the expected value of a constant is the constant itself). It remains to show that $\mathbb{E}[A\mathbf{x}] = A\mathbb{E}[\mathbf{x}]$. Suppose \mathbf{x} is an n -dimensional random vector in a space X :

$$\mathbf{x} = \begin{bmatrix} x_0 \\ x_1 \\ \vdots \\ x_n \end{bmatrix}$$

where each of the $x_i \in \mathbb{R}$. By definition, the expectation value of \mathbf{x} is given by

$$\mathbb{E}[\mathbf{x}] = \begin{bmatrix} \mathbb{E}[x_0] \\ \mathbb{E}[x_1] \\ \vdots \\ \mathbb{E}[x_n] \end{bmatrix}$$

and for any continuous random variable, we have

$$\mathbb{E}[Y] = \int_{-\infty}^{\infty} y f(y) \, dy$$

where $f(y)$ is some probability density function. For each of the x_i , define $f_i(x_i)$ to be the corresponding probability density function. Then

$$\mathbb{E}[\mathbf{x}] = \begin{bmatrix} \int_{-\infty}^{\infty} x_0 f_0(x_0) \, dx_0 \\ \int_{-\infty}^{\infty} x_1 f_1(x_1) \, dx_1 \\ \vdots \\ \int_{-\infty}^{\infty} x_n f_n(x_n) \, dx_n \end{bmatrix}$$

Let $A \in M_{m \times n}(\mathbb{R})$, with rows $\mathbf{a}_0, \mathbf{a}_1, \dots, \mathbf{a}_m$. Then $A\mathbb{E}[\mathbf{x}]$ is given by

$$\begin{aligned} A\mathbb{E}[\mathbf{x}] &= \begin{bmatrix} a_{0,0} & a_{0,1} & \cdots & a_{0,n} \\ a_{1,0} & a_{1,1} & \cdots & a_{1,n} \\ \vdots & \vdots & \ddots & \vdots \\ a_{m,0} & a_{m,1} & \cdots & a_{m,n} \end{bmatrix} \begin{bmatrix} \mathbb{E}[x_0] \\ \mathbb{E}[x_1] \\ \vdots \\ \mathbb{E}[x_n] \end{bmatrix} \\ &= \begin{bmatrix} \sum_{i=0}^n a_{0,i} \mathbb{E}[x_i] \\ \sum_{i=0}^n a_{1,i} \mathbb{E}[x_i] \\ \vdots \\ \sum_{i=0}^n a_{m,i} \mathbb{E}[x_i] \end{bmatrix} \end{aligned}$$

Now, we manipulate this expression to obtain $\mathbb{E}[A\mathbf{x}]$. Because the expected value of a sum is the sum of expected values, we have

$$\begin{bmatrix} \sum_{i=0}^n a_{0,i} \mathbb{E}[x_i] \\ \sum_{i=0}^n a_{1,i} \mathbb{E}[x_i] \\ \vdots \\ \sum_{i=0}^n a_{n,i} \mathbb{E}[x_i] \end{bmatrix} = \begin{bmatrix} \mathbb{E} \left[\sum_{i=0}^n a_{0,i} x_i \right] \\ \mathbb{E} \left[\sum_{i=0}^n a_{1,i} x_i \right] \\ \vdots \\ \mathbb{E} \left[\sum_{i=0}^n a_{n,i} x_i \right] \end{bmatrix}$$

and now because the expected value of a vector is a vector of expected values,

$$\begin{bmatrix} \mathbb{E} \left[\sum_{i=0}^n a_{0,i} x_i \right] \\ \mathbb{E} \left[\sum_{i=0}^n a_{1,i} x_i \right] \\ \vdots \\ \mathbb{E} \left[\sum_{i=0}^n a_{n,i} x_i \right] \end{bmatrix} = \mathbb{E} \begin{bmatrix} \sum_{i=0}^n a_{0,i} x_i \\ \sum_{i=0}^n a_{1,i} x_i \\ \vdots \\ \sum_{i=0}^n a_{n,i} x_i \end{bmatrix}$$

(b) The covariance matrix $\text{cov}[\mathbf{y}]$ is defined by

$$\begin{aligned} \text{cov}[\mathbf{y}] &= \begin{bmatrix} \mathbb{E}[(y_0 - \mathbb{E}[y_0])(y_0 - \mathbb{E}[y_0])] & \mathbb{E}[(y_0 - \mathbb{E}[y_0])(y_1 - \mathbb{E}[y_1])] & \cdots & \mathbb{E}[(y_0 - \mathbb{E}[y_0])(y_n - \mathbb{E}[y_n])] \\ \mathbb{E}[(y_1 - \mathbb{E}[y_1])(y_0 - \mathbb{E}[y_0])] & \mathbb{E}[(y_1 - \mathbb{E}[y_1])(y_1 - \mathbb{E}[y_1])] & \cdots & \mathbb{E}[(y_1 - \mathbb{E}[y_1])(y_n - \mathbb{E}[y_n])] \\ \vdots & \vdots & \ddots & \vdots \\ \mathbb{E}[(y_n - \mathbb{E}[y_n])(y_0 - \mathbb{E}[y_0])] & \mathbb{E}[(y_n - \mathbb{E}[y_n])(y_1 - \mathbb{E}[y_1])] & \cdots & \mathbb{E}[(y_n - \mathbb{E}[y_n])(y_n - \mathbb{E}[y_n])] \end{bmatrix} \\ &= \begin{bmatrix} \text{cov}(y_0, y_0) & \text{cov}(y_0, y_1) & \cdots & \text{cov}(y_0, y_n) \\ \text{cov}(y_1, y_0) & \text{cov}(y_1, y_1) & \cdots & \text{cov}(y_1, y_n) \\ \vdots & \vdots & \ddots & \vdots \\ \text{cov}(y_n, y_0) & \text{cov}(y_n, y_1) & \cdots & \text{cov}(y_n, y_n) \end{bmatrix} \end{aligned}$$

it is a property of covariance that for all random variables X and Y , with scalars a and b , $\text{cov}(X + a, Y + b) = \text{cov}(X, Y)$. Hence, since each $y_i = (A\mathbf{x})_i + b_i$, then $\forall i, j \in \{0, 1, \dots, n\}$,

$$\begin{aligned} \text{cov}[y_i, y_j] &= \text{cov}[(A\mathbf{x})_i + b_i, (A\mathbf{x})_j + b_j] \\ &= \text{cov}[(A\mathbf{x})_i, (A\mathbf{x})_j] \end{aligned}$$

thus, examining the matrix above, we see $\text{cov}[A\mathbf{x} + \mathbf{b}] = \text{cov}[A\mathbf{x}]$. It remains to show $\text{cov}[A\mathbf{x}] = A \text{cov}[\mathbf{x}] A^\top$

Problem 2.

Given the dataset $\mathcal{D} = \{(x, y)\} = \{(0, 1), (2, 3), (3, 6), (4, 8)\}$

- (a) Find the least squares estimate $y = \boldsymbol{\theta}^\top \mathbf{x}$ by hand using Cramer's Rule.
 - (b) Use the normal equations to find the same solution and verify it is the same as part (a).
 - (c) Plot the data and the optimal linear fit you found.
 - (d) Find randomly generate 100 points near the line with white Gaussian noise and then compute the least squares estimate (using a computer). Verify that this new line is close to the original and plot the new dataset, the old line, and the new line.
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Solution: