# Ch 4.4.1: Linear Discriminant Analysis

Lecture 8 - CMSE 381

Michigan State University

::

Dept of Computational Mathematics, Science & Engineering

February 5, 2024

#### Announcements

Last time:

Logistic Regression

#### **Announcements:**

- Third homework due Friday! Covers:
  - ▶ Mon 2/12 Review Midterm 1
  - ▶ Weds 2/14 No class
  - ► Fri 2/16 Midterm 1
- Office hours
  - Mon: 4-6pm; Tue: 12:30-2:30pm;Wed: 7-9pm; Thu: 12:30-2:30pm

2/39

#### Covered in this lecture

- Bayes theorem
- Linear Discriminant Analysis,
- Quadratic Discriminant Analysis

(MSU-CMSE) February 5, 2024

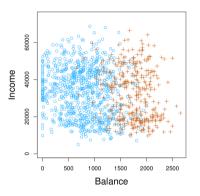
#### Section 1

Logistic Regression Review

(MSU-CMSE) February 5, 2024

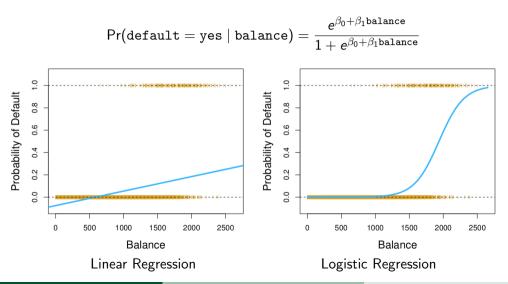
#### What is classification

- Classification: When the response variable is qualitative
- Goal: Model the probability that Y belongs to a particular category
- Example data:
   p(balance) = Pr(default = yes | balance)



5/39

## Logistic Regression



(MSU-CMSE) February 5, 2024

#### Odds

$$\frac{p(x)}{1 - p(x)} = \frac{\Pr(Y = 1 \mid X = x)}{1 - \Pr(Y = 1 \mid X = x)} = \frac{\Pr(Y = 1 \mid X = x)}{\Pr(Y = 0 \mid X = x)}$$

- Logistic function is chosen so that odds are linear
- Can take any value from 0 (low odds) to ∞ (high odds)

- 1 in 4 people with odds of 1/3 will default since p(X) = 0.25 and 0.25/(1-0.25) = 1/3
- 9 in 10 people with odds of 9 will default since p(X) = 0.9 and 0.9/(1-0.9) = 9

7/39

### How to get logistic function

Assume the (natural) log odds (logits) follow a linear model

$$\log\left(\frac{p(x)}{1-p(x)}\right) = \beta_0 + \beta_1 x$$

Solve for p(x):

$$p(x) = \frac{e^{\beta_0 + \beta_1 x}}{1 + e^{\beta_0 + \beta_1 x}}$$

8 / 39

Playing with the logistic function: https://www.desmos.com/calculator/jzsakksqcm

## Interpreting the coefficients

$$p(x) = \frac{e^{\beta_0 + \beta_1 x}}{1 + e^{\beta_0 + \beta_1 x}}$$

$$\log\left(\frac{p(x)}{1-p(x)}\right) = \beta_0 + \beta_1 x$$

 $eta_1$  means increasing x by one unit increases the log odds by 1 unit

9/39

## Estimating Coefficients: Maximum Likelihood Estimation

• **Likelihood:** Probability that data is generated from a model

$$\ell(model) = \Pr[data \mid model]$$

• Find the most likely model

$$\max_{model} \ell(model)$$

Hard to maximize likelihood, instead maximize log

$$\max_{model} \log(\ell(model))$$

 Strictly increasing log function doesn't change maximum

$$egin{aligned} \mathsf{Pr}(\mathit{Y}=1\mid \mathit{X}) &= \mathit{p}(\mathit{X}) \ &= rac{e^{eta_0 + eta_1 \mathit{X}}}{1 + e^{eta_0 + eta_1 \mathit{X}}} \end{aligned}$$

$$\ell(\beta_0, \beta_1) = \prod_{i | y_i = 1} p(x_i) \prod_{i' | y_{i'} = 0} (1 - p(x_{i'}))$$

10 / 39

## Multiple Logistic Regression

#### Multiple features:

$$p(X) = \frac{e^{\beta_0 + \beta_1 X_1 + \dots + \beta_\rho X_\rho}}{1 + e^{\beta_0 + \beta_1 X_1 + \dots + \beta_\rho X_\rho}}$$

#### **Equivalent to:**

$$\log\left(\frac{p(X)}{1-p(X)}\right) = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p$$

## Multinomial Logistic Regression

What if we have a categorical variable with more than two levels (let's say K of them)?

#### Plan A

Play the dummy variable game:

Make K the baseline:

$$\Pr(Y = k | X = x) = \frac{e^{\beta_{k0} + \beta_{k1} x_1 + \dots + \beta_{kp} x_p}}{1 + \sum_{l=1}^{K-1} e^{\beta_{l0} + \beta_{l1} x_1 + \dots + \beta_{lp} x_p}}$$

$$\Pr(Y = K | X = x) = \frac{1}{1 + \sum_{l=1}^{K-1} e^{\beta_{l0} + \beta_{l1} x_1 + \dots + \beta_{lp} x_p}}.$$

Calculated so that log odds between two classes is linear:

$$\log \left( \frac{\Pr(Y = k \mid X = x)}{\Pr(Y = K \mid X = x)} \right) = \beta_{k0} + \beta_{k1} x_1 + \dots + \beta_{kp} x_p$$

(MSU-CMSE) February 5, 2024

## Example

Predict  $Y \in \{\text{stroke, overdose, seizure}\}\$ for hospital visits based on Xp

$$\begin{split} \Pr(Y = \texttt{stroke} \mid X = x) &= \frac{\exp(\beta_{\texttt{str},0} + \beta_{\texttt{str},1}x)}{1 + \exp(\beta_{\texttt{str},0} + \beta_{\texttt{str},1}x) + \exp(\beta_{\texttt{0D},0} + \beta_{\texttt{0D},1}x)} \\ \Pr(Y = \texttt{overdose} \mid X = x) &= \frac{\exp(\beta_{\texttt{0D},0} + \beta_{\texttt{0D},1}x)}{1 + \exp(\beta_{\texttt{str},0} + \beta_{\texttt{str},1}x) + \exp(\beta_{\texttt{0D},0} + \beta_{\texttt{0D},1}x)} \\ \Pr(Y = \texttt{seizure} \mid X = x) &= \frac{1}{1 + \exp(\beta_{\texttt{str},0} + \beta_{\texttt{str},1}x) + \exp(\beta_{\texttt{0D},0} + \beta_{\texttt{0D},1}x)} \end{split}$$

MSU-CMSE) February 5, 2024

#### Softmax

#### Plan B: Softmax coding

Treat all levels symmetrically

$$\Pr(Y = k | X = x) = \frac{e^{\beta_{k0} + \beta_{k1} x_1 + \dots + \beta_{kp} x_p}}{\sum_{l=1}^{K} e^{\beta_{l0} + \beta_{l1} x_1 + \dots + \beta_{lp} x_p}}.$$

Calculated so that log odds between two classes is linear

$$\log\left(\frac{\Pr(Y=k|X=x)}{\Pr(Y=k'|X=x)}\right) = (\beta_{k0} - \beta_{k'0}) + (\beta_{k1} - \beta_{k'1})x_1 + \dots + (\beta_{kp} - \beta_{k'p})x_p.$$

(MSU-CMSE) February 5, 2024

### Softmax example

$$\begin{split} & \Pr(Y = \texttt{stroke} \mid X = x) \\ & = \frac{\exp(\beta_{\texttt{str},0} + \beta_{\texttt{str},1}x)}{\exp(\beta_{\texttt{str},0} + \beta_{\texttt{str},1}x) + \exp(\beta_{\texttt{0D},0} + \beta_{\texttt{0D},1}x) + \exp(\beta_{\texttt{seiz},0} + \beta_{\texttt{seiz},1}x)} \\ & \Pr(Y = \texttt{overdose} \mid X = x) \\ & = \frac{\exp(\beta_{\texttt{0D},0} + \beta_{\texttt{0D},1}x)}{\exp(\beta_{\texttt{str},0} + \beta_{\texttt{str},1}x) + \exp(\beta_{\texttt{0D},0} + \beta_{\texttt{0D},1}x) + \exp(\beta_{\texttt{seiz},0} + \beta_{\texttt{seiz},1}x)} \\ & \Pr(Y = \texttt{seizure} \mid X = x) \\ & = \frac{\exp(\beta_{\texttt{seiz},0} + \beta_{\texttt{seiz},1}x)}{\exp(\beta_{\texttt{str},0} + \beta_{\texttt{str},1}x) + \exp(\beta_{\texttt{0D},0} + \beta_{\texttt{0D},1}x) + \exp(\beta_{\texttt{seiz},0} + \beta_{\texttt{seiz},1}x)} \end{split}$$

(MSU-CMSE) February 5, 2024

#### Section 2

Generative Models

MSU-CMSE) February 5, 2024

#### Goal:

Another way to approximate 
$$Pr(Y = k \mid X = x)$$

How? BAYES!!!!

17 / 39

## Bayes Theorem

$$P(A \mid B) = \frac{P(A) \cdot P(B \mid A)}{P(B)}$$

- P(A | B) (Posterior): probability of A being true given B
- P(A) (Prior): probability of A being true
- P(B) (Marginalization): probability of B being true
- P(B | A) (Likelihood): probability of B true given that A is true

18 / 39

## Example: Favorite language by year

#### Example:

- dangerous fires are rare (1%)
- but smoke is fairly common (10%) due to barbecues,
- and 90% of dangerous fires make smoke

We can then discover the **probability of dangerous Fire when there is Smoke**:

$$P(Fire|Smoke) = \frac{P(Fire) P(Smoke|Fire)}{P(Smoke)}$$
$$= \frac{1\% \times 90\%}{10\%}$$
$$= 9\%$$

(MSU-CMSE) February 5, 2024

## Example: Favorite language by year

|        | Fresh | Soph | Junior | Senior |    |
|--------|-------|------|--------|--------|----|
| Python | 9     | 14   | 13     | 17     | 53 |
| R      | 14    | 15   | 10     | 8      | 47 |
|        | 23    | 29   | 23     | 25     | ,  |

$$P(Y = py \mid X = jr) = \frac{P(Y = py) \cdot P(X = jr \mid Y = py)}{P(X = jr)}$$

(MSU-CMSE) February 5, 2024

## An equivalent formula

$$P(A \mid B) = \frac{P(A) \cdot P(B \mid A)}{P(B)}$$
  

$$\Leftrightarrow P(A \mid B) = \frac{P(A) \cdot P(B \mid A)}{P(A)P(B|A) + P(A^c)P(B|A^c)}$$

(MSU-CMSE) February 5, 2024

## Following book notation

- Classify an observation into one of K > 2 classes
- $\pi_k$  is the *prior* probability that a randomly chosen observation comes from the kth class, P(Y = k)

- $f_k(X) = \Pr(X \mid Y = k)$  is the density function of X for an observation from the kth class
  - Large  $f_k(x)$  if there is high probability that observation in the kth class has  $X \approx x$
  - ► Small if unlikely that an observation in the kth class has  $X \approx x$

22 / 39

## Bayes to the rescue!

Posterior probability that an observation X = x belongs to the kth class:

$$p_k(x) = \Pr(Y = k \mid X = x) = \frac{\pi_k f_k(x)}{\sum_{\ell=1}^K \pi_\ell f_\ell(x)}$$

- Second equation is Bayes,  $p_k(x)$  is our notation for it
- Plug in estimates for  $\pi_k$  and  $f_k(x)$  to get an estimate
- Estimate for  $\pi_k$  is easy with a random sample, just take proportion that are kth class

23 / 39

• Estimating desnity  $f_k$  is hard(er).....

#### Section 3

Linear Discriminant Analysis for p = 1

MSU-CMSE) February 5, 2024

## Assumptions

Assume  $f_k(x)$  is normal/Gaussian:

$$f_k(x) = \frac{1}{\sqrt{2\pi}} \exp\left(\frac{-1}{2\sigma_k^2}(x - \mu_k)^2\right)$$

- $\mu_k = \text{mean of } k \text{th class}$
- $\sigma_k^2$  = variance of kth class
- Assume  $\sigma_1^2 = \cdots = \sigma_K^2$

$$p_k(x) = \frac{\pi_k \frac{1}{\sqrt{2\pi}} \exp\left(\frac{-1}{2\sigma^2} (x - \mu_k)^2\right)}{\sum_{\ell=1}^K \pi_\ell \frac{1}{\sqrt{2\pi}} \exp\left(\frac{-1}{2\sigma^2} (x - \mu_\ell)^2\right)}$$

25 / 39

#### Bayes Classifier

Same Bayes person, different Bayes definition

#### Bayes classifier

Assign the class k for which  $p_k(x)$  is largest

#### Finding largest k

$$p_k(x) = \frac{\pi_k \frac{1}{\sqrt{2\pi}} \exp\left(\frac{-1}{2\sigma^2} (x - \mu_k)^2\right)}{\sum_{\ell=1}^K \pi_\ell \frac{1}{\sqrt{2\pi}} \exp\left(\frac{-1}{2\sigma^2} (x - \mu_\ell)^2\right)}$$

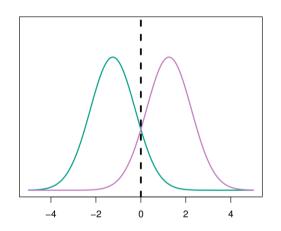
is the same as finding largest k for

$$\delta_k(x) = x \cdot \frac{\mu_k}{\sigma^2} - \frac{\mu_k^2}{2\sigma^2} + \log(\pi_k)$$

## Example when K=2, $\pi_1=\pi_2$

#### **Decision boundary:**

$$x = \frac{\mu_1 + \mu_2}{2}$$

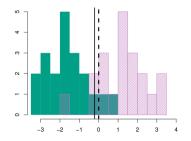


27 / 39

# New plan: Linear Discriminant Analysis (LDA)

#### Estimate

$$\delta_k(x) = x \cdot \frac{\mu_k}{\sigma^2} - \frac{\mu_k^2}{2\sigma^2} + \log(\pi_k)$$



$$\hat{\mu}_k = \frac{1}{n_k} \sum_{i|v_i=k} x_i$$

• 
$$\hat{\sigma}^2 = \frac{1}{n - K} \sum_{k=1}^K \sum_{i|y_i = k} (x_i - \hat{\mu}_k)^2$$

- $\hat{\pi}_k = n_k/n$
- Black solid line: calculated boundary for assignment
- This example,  $n_1 = n_2 = 20$ , so  $\hat{\pi}_1 = \hat{\pi}_2$ , so decision bdry half way between sample means
- Optimal Bayes decision boundary dashed line

## Example 1

Suppose that we wish to predict whether a given stock will issue a dividend this year ("Yes" or "No") based on X, last year's percent proft. We examine a large number of companies and discover that the mean value of X for companies that issued a dividend was  $\bar{X} = 10$ , while the mean for those that didn't was  $\bar{X} = 0$ . In addition. the variance of X for these two sets of companies was  $\hat{\sigma}^2 = 36$ . Finally, 80% of companies issued dividends. Assuming that X follows a normal distribution. predict the probability that a company will issue a dividend this year given that its percentage proft was X = 4 last year.

(MSU-CMSE) February 5, 2024

### Example 2

Assume the probability that a person defaulted on credit card payment is 10%. Based on the given use LDA to predict the default status of a person with a credit card balance of 1800.

| Balance | Prediction |
|---------|------------|
| 0       | No         |
| 500     | No         |
| 1000    | Yes        |
| 1500    | No         |
| 2000    | Yes        |
| 2500    | Yes        |
|         |            |

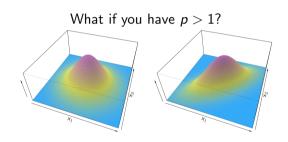
30 / 39

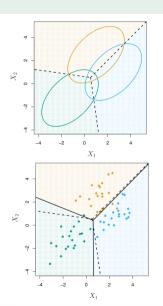
#### LDA review

- Assume observations in each class come from normal
- Class specific means
- Common variance
- Plug in estimates into Bayes classifier

(MSU-CMSE) February 5, 2024

## High dimensional LDA - p>1





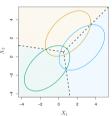
32 / 39

### LDA - p > 1

- Assume observations in the kth drawn from multi-variate normal distribution with the same covariance matrix for all k:  $N(\mu_k, \Sigma)$
- For new date point x, predict the k for which  $p_k(x) = \Pr(Y = k | X = x)$  is largest
- Equivalent to finding k for which  $\delta_k(x)$  is largest

$$\delta_k(x) = x^T \Sigma^{-1} \mu_k - \frac{1}{2} \mu_k^T \Sigma^{-1} \mu_k + \log \pi_k$$

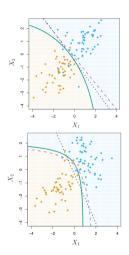
• Approximate  $\mu_i$ 's  $\pi_i$ 's and  $\Sigma$  to find boundary where the returned k switches





## Quadratic Discriminant Analysis (QDA)

- Same idea as LDA, but don't assume same covariance matrix
- Assume observations in kth colass drawn from normal distribution  $N(\mu_k, \Sigma_k)$
- Make new predictions based on  $\delta_k(x) = \frac{1}{2}(x \mu_k)^T \Sigma_k^{-1}(x \mu_k) \frac{1}{2} \log |\Sigma_k| + \log \pi_k$
- This setup means decision boundaries are quadratic



(MSU-CMSE)

### Example

Assume the probability that a person defaulted on credit card payment is 10%. Based on the given use LDA to predict the default status of a person with a credit card balance of 1800.

| Balance | Prediction |
|---------|------------|
| 0       | No         |
| 500     | No         |
| 1000    | Yes        |
| 1500    | No         |
| 2000    | Yes        |
| 2500    | Yes        |
|         | 1          |

35 / 39

# Confusion matrix and types of Errors

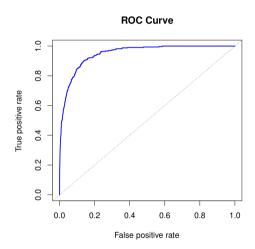
|                   |       | True default status |     |       |
|-------------------|-------|---------------------|-----|-------|
|                   |       | No                  | Yes | Total |
| Predicted         | No    | 9432                | 138 | 9570  |
| $default\ status$ | Yes   | 235                 | 195 | 430   |
|                   | Total | 9667                | 333 | 10000 |

|           |               | True class      |                 |       |
|-----------|---------------|-----------------|-----------------|-------|
|           |               | - or Null       | + or Non-null   | Total |
| Predicted | – or Null     | True Neg. (TN)  | False Neg. (FN) | N*    |
| class     | + or Non-null | False Pos. (FP) | True Pos. (TP)  | P*    |
|           | Total         | N               | P               |       |

| Name             | Definition | Synonyms                                    |
|------------------|------------|---|
| False Pos. rate  | FP/N       | Type I error, 1—Specificity                 |
| True Pos. rate   | TP/P       | 1—Type II error, power, sensitivity, recall |
| Pos. Pred. value | $TP/P^*$   | Precision, 1-false discovery proportion     |
| Neg. Pred. value | $TN/N^*$   |   |

36 / 39

#### **ROC** curve



MSU-CMSE) February 5, 2024

#### Multivariate normal distribution

#### Gaussian (normal) distributions

•  $Z \sim N(0,1)$  means Z follows a standard Gaussian distribution, i.e., has probability density

$$f(z) = \frac{1}{\sqrt{2\pi}}e^{-\frac{x^2}{2}}$$

• If  $Z_1, Z_2, ...., Z_d$  are iid N(0,1) random variables, then say  $\mathbf{Z} := (Z_1, Z_2, ..., Z_d)$  follows a standard multivariate Gaussian distribution on  $\mathbb{R}^d$ , i.e.,  $\mathbf{Z} \sim N(0, I_d)$ .

• Other Gaussian distributions on  $\mathbb{R}^d$  arise by applying (invertible) linear maps and translations to  $\mathbf{Z}$ :

$$\mathsf{z}\mapsto\mathsf{A}\mathsf{z}\mapsto\mathsf{A}\mathsf{z}+\mu.$$

- $f X := f AZ + m \mu \sim \mathcal N(m \mu, f AA^{\sf T}) ext{ has}$   $\mathbb E(f X) = m \mu \quad ext{and } ext{cov}(f X) = f AA^{\sf T}$
- ▶ the (i, j)th entry of cov(X) represents the correlation between X<sub>i</sub> and X<sub>i</sub>.

## Appendix: Review of Multivariate normal distribution

ullet  $oldsymbol{\mathsf{X}} \sim \mathcal{N}(oldsymbol{\mu}, oldsymbol{\Sigma})$  has the probability density function

$$p(\mathbf{x}) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} e^{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \mathbf{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})}$$

 Estimate μ and Σ from data: maximum likelihood estimators

$$\hat{\boldsymbol{\mu}} = \frac{1}{n} \sum_{i=1}^{n} \mathbf{x}_{i}$$

$$\hat{\boldsymbol{\Sigma}} = \frac{1}{n} \sum_{i=1}^{n} (\mathbf{x}_i - \hat{\boldsymbol{\mu}}) (\mathbf{x}_i - \hat{\boldsymbol{\mu}})^T$$

 $ightharpoonup \hat{oldsymbol{\mu}}$  is unbias and  $\hat{oldsymbol{\Sigma}}$  is slightly biased