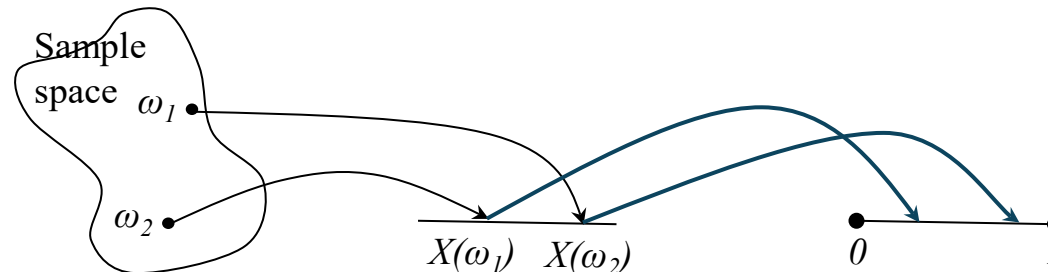


Random Variable

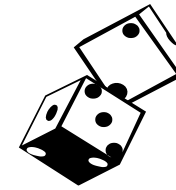
Random Variable Definition

- It frequently occurs that in performing an experiment we are mainly interested in some functions of the outcome as opposed to the outcome itself
 - Rotery, casino, game, etc.
- These quantities of interest, or more formally, these real-valued functions defined on the sample space, are known as **random variables (RVs)**.
- More mathematical expression of RVs
 - Given a probability space (S, F, P) , a random variable is a measurable function (mapping) from S to the real line $X: S \rightarrow R$
 - $X(\{H\}) = 100, X(\{H\}) = -50$
 - $X(\{1,2\}) = 100, X(\{3,4,5,6\}) = -50$
- Since the value of a random variable is defined by the outcome of the experiment, we may assign probabilities to the possible values of the random variable

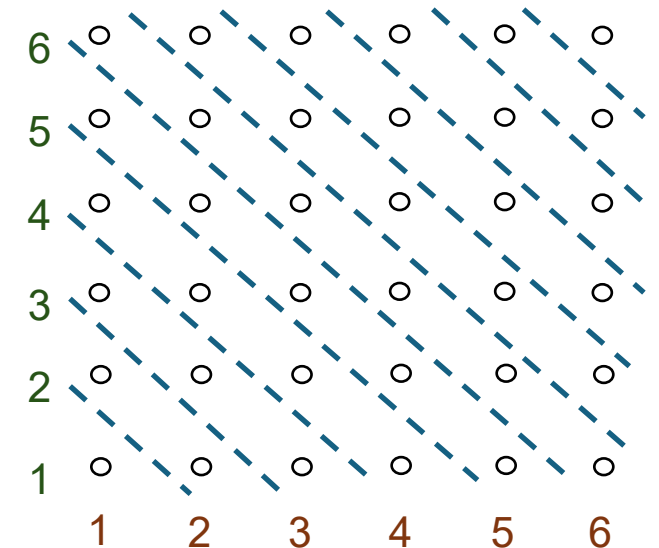


Random Variable: Examples

- Ex 2.1] Letting X denote the random variable that is defined as the sum of two fair dice



- The outcomes of two dice = (ω_1, ω_2)
- RV $X(\omega_1, \omega_2) = \omega_1 + \omega_2$
- Possible values of $X = \{2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12\}$
- Probabilities to the possible values of the random variable
 - $P(X = 2) = P\{(1,1)\} = 1/36$
 - $P(X = 3) = P\{(1,2), (2,1)\} = 2/36$
 - ...
 - $P(X = 11) = P\{(5,6), (6,5)\} = 2/36$
 - $P(X = 12) = P\{(6,6)\} = 1/36$
 - $1 = P\{\cup_{n=2}^{12}\{X = n\}\} = \sum_{n=2}^{12} P(X = n)$



Random Variable: Examples

- Ex.2.3] Suppose that we toss a coin having a probability p of coming up heads, until the first head appears. Letting N denote the number of flips required, then assuming that the outcome of successive flips are independent, N is a random variable taking on one of the values 1, 2, 3, . . . , with respective probabilities

$$\begin{aligned}P\{N = 1\} &= P\{H\} = p, \\P\{N = 2\} &= P\{(T, H)\} = (1 - p)p, \\P\{N = 3\} &= P\{(T, T, H)\} = (1 - p)^2p,\end{aligned}$$

$$P\{N = n\} = P\{(T, T, \dots, T, H)\} = (1 - p)^{n-1}p, n \geq 1$$

- Indicator random variable

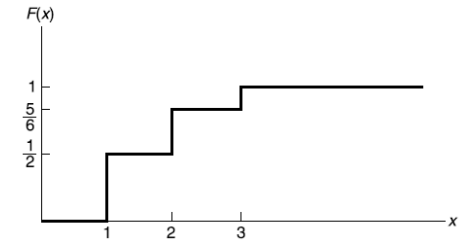
$$I_E(x) = 1 \text{ if } x \in E, 0 \text{ otherwise.}$$

Cumulative Distribution Function (CDF)

- If a random variable takes on either a finite or a countable number of possible values, the RV is called **discrete**.
- If a continuum of possible values, **continuous**.
- The **cumulative distribution function (CDF)** $F(\cdot)$ of the random variable X is defined for any real number b , $-\infty < b < \infty$ by $F(b) = P\{X \leq b\}$.
- Properties of CDF F
 1. $F(b)$ is a **nondecreasing** function of b
 2. $\lim_{b \rightarrow \infty} F(b) = F(\infty) = 1$
 3. $\lim_{b \rightarrow -\infty} F(b) = F(-\infty) = 0$
- Ex. $P\{a < X \leq b\} = F(b) - F(a)$ for all $a < b$

Discrete RVs

- A random variable that can take on at most countable number of possible values
- Probability Mass Function (pmf): for a discrete random variable X , pmf $p(a)$ of X is defined as $p(a) = P\{X = a\}$.
- Properties of pmf
 1. $p(x_i) > 0, i = 1, 2, \dots$
 2. $p(x) = 0$ all other values of x
 3. $\sum_{i=1, \dots, \infty} p(x_i) = 1$
 4. cdf vs. pmf: $F(a) = \sum_{x_i \leq a} p(x_i)$
- Ex. $p(1) = 1/2, p(2) = 1/3, p(3) = 1/6$
 $F(a) = 0 \quad a < 1, F(a) = \frac{1}{2} \quad 1 \leq a < 2, F(a) = \frac{5}{6} \quad 2 \leq a < 3, F(a) = 1 \quad 3 \leq a$
- Discrete random variables are often classified according to their pmf
 - Ex: Bernoulli RV, Binomial RV, Geometric RV, Poisson RV



Example of Discrete RVs: The Bernoulli RV

- A random variable X is said to be a **Bernoulli** RV for given sample space $S = \{A, A^c\}$ and some $p \in (0,1)$ if its pmf is given by
$$P\{A\} = p, P\{A^c\} = 1 - p$$
- For example, suppose that a trial, or an experiment, whose outcome can be classified as either a “success” or as a “failure” is performed. If we let X equal 1 if the outcome is a success and 0 if it is a failure, then the probability mass function of X is given by
$$\begin{aligned} p(0) &= P\{X = 0\} = 1 - p, \\ p(1) &= P\{X = 1\} = p \end{aligned} \tag{2.2}$$
where $p, 0 \leq p \leq 1$, is the probability that the trial is a “success.”

Example of Discrete RVs: The Binomial RV

- Suppose that n independent trials, each of which results in a “success” with probability p and in a “failure” with probability $1 - p$, are to be performed.
- If X represents the number of successes that occur in the n trials, then X is said to be a **binomial** random variable with parameters (n, p) .
- The pmf of a binomial RV with (n, p)
$$p(i) = \binom{n}{i} p^i (1 - p)^{n-i} \text{ for } i = 0, 1, \dots, n \quad (2.3)$$

where $\binom{n}{i} = \frac{n!}{(n-i)!i!}$.
- $\sum_{i=0}^n p(i) = \sum_{i=0}^n \binom{n}{i} p^i (1 - p)^{n-i} = (p + (1 - p))^n = 1$

Example of Discrete RVs: The Binomial RV

- Ex. 2.7] It is known that any item produced by a certain machine will be defective with probability 0.1, independently of any other item. What is the probability that in a sample of three items, at most one will be defective?
 - If X is the number of defective items in the sample, then X is a binomial random variable with parameters $(3, 0.1)$. Hence, the desired probability is given by
 - $P\{X = 0\} + P\{X = 1\} = \binom{3}{0}(0.1)^0(0.9)^3 + \binom{3}{1}(0.1)^1(0.9)^2 = 0.972$

Example of Discrete RVs: The Geometric RV

- Suppose that independent trials, each having probability p of being a success, are performed until a success occurs.
- If we let X be the number of trials required until the first success, then X is said to be a **geometric** random variable with parameter p .
- Its probability mass function is given by
$$p(n) = P\{X = n\} = (1 - p)^{n-1}p, n = 1, 2, \dots \quad (2.4)$$
- To check that $p(n)$ is a probability mass function, we note that
$$\sum_{n=1}^{\infty} p(n) = p \sum_{n=1}^{\infty} (1 - p)^{n-1} = 1$$

Example of Discrete RVs: The Poisson RV

- A random variable X , taking on one of the values $0, 1, 2, \dots$, is said to be a **Poisson** random variable with parameter λ , if for some $\lambda > 0$,

$$p(i) = P\{X = i\} = e^{-\lambda} \frac{\lambda^i}{i!}, i = 0, 1, 2, 3, \dots \quad (2.5)$$

- To check that $p(n)$ is a probability mass function, we note that

$$\sum_{i=0}^{\infty} p(i) = \sum_{i=0}^{\infty} e^{-\lambda} \frac{\lambda^i}{i!} = e^{-\lambda} \sum_{i=0}^{\infty} \frac{\lambda^i}{i!} = e^{-\lambda} e^{\lambda} = 1$$

- Binomial vs. Poisson RVs

- $n \gg 1, p \ll 1$, let $\lambda = np$

- $P\{X = 1\} = \binom{n}{i} p^i (1-p)^{n-i} = \frac{n!}{(n-i)!i!} \left(\frac{\lambda}{n}\right)^i \left(1 - \frac{\lambda}{n}\right)^{n-i} = \frac{n(n-1)\dots(n-i+1)}{n^i} \frac{\lambda^i}{i!} \frac{\left(1 - \frac{\lambda}{n}\right)^n}{\left(1 - \frac{\lambda}{n}\right)^i}$

- For large n and small p

- $\left(1 - \frac{\lambda}{n}\right)^n \approx e^{-\lambda}, \left(1 - \frac{\lambda}{n}\right)^i \approx 1, \frac{n(n-1)\dots(n-i+1)}{n^i} \approx 1$

- $P\{X = i\} \approx e^{-\lambda} \frac{\lambda^i}{i!}$

Example of Discrete RVs: The Poisson RV

- Ex. 2.10] Suppose that the number of typographical errors on a single page of this book has a Poisson distribution with parameter $\lambda = 1$. Calculate the probability that there is at least one error on this page.
 - $P\{X \geq 1\} = 1 - P\{X = 0\} = 1 - e^{-1} = 0.633$

Discrete probability distribution	Probability mass function, $p(x)$	Moment generating function, $\phi(t)$	Mean	Variance
Binomial with parameters n, p , $0 \leq p \leq 1$	$\binom{n}{x} p^x (1-p)^{n-x}$, $x = 0, 1, \dots, n$	$(pe^t + (1-p))^n$	np	$np(1-p)$
Poisson with parameter $\lambda > 0$	$e^{-\lambda} \frac{\lambda^x}{x!}$, $x = 0, 1, 2, \dots$	$\exp\{\lambda(e^t - 1)\}$	λ	λ
Geometric with parameter $0 \leq p \leq 1$	$p(1-p)^{x-1}$, $x = 1, 2, \dots$	$\frac{pe^t}{1 - (1-p)e^t}$	$\frac{1}{p}$	$\frac{1-p}{p^2}$

Continuous RVs

- RV X is said to be a **continuous RV** if there exists a non-negative function $f(x)$, defined for all real $x \in \{-\infty, \infty\}$ having the property that for any set B of real numbers

$$P\{X \in B\} = \int_B f(x)dx = 1$$

- The function $f(x)$ is called the probability density function (pdf) of the random variable X

- Properties of pdf of X

- $P\{X \in (-\infty, \infty)\} = \int_{-\infty}^{\infty} f(x)dx = 1$
- $P\{a \leq X \leq b\} = \int_a^b f(x)dx$
- $P\{X = a\} = \int_a^a f(x)dx = 0$
- $F(a) = P\{X \in (-\infty, a)\} = \int_{-\infty}^a f(x)dx$
- $\frac{dF(a)}{da} = f(a)$

- Ex. Uniform Rv, Exponential Rv, Gamma RV, Normal RV

Example of Continuous RVs: Uniform RV

- A RV is said to be uniformly distributed over the interval (α, β) if its probability density function is given by

$$f(x) = \begin{cases} \frac{1}{\beta - \alpha}, & \alpha < x < \beta \\ 0, & \text{otherwise} \end{cases}$$

- Ex. 2.14] $X \sim U(0, 10)$

- $P\{X < 3\} = \int_{-\infty}^3 f(x) dx = \int_0^3 \frac{1}{10} dx = \frac{3}{10}$
- $P\{X > 7\} = \int_7^{\infty} f(x) dx = \int_7^{10} \frac{1}{10} dx = \frac{3}{10}$
- $P\{1 \leq X < 6\} = \int_1^6 f(x) dx = \int_1^6 \frac{1}{10} dx = \frac{5}{10} = \frac{1}{2}$

Example of Continuous RVs: Exponential RV

- A continuous RV is said to be an exponential RV with parameter λ if its probability density function is given, for $\lambda > 0$, by

$$f(x) = \begin{cases} \lambda e^{-\lambda x}, & \text{if } x \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

$$F(a) = \int_0^a \lambda e^{-\lambda x} dx = 1 - e^{-\lambda a}$$

Example of Continuous RVs: Gamma RV

- A continuous RV is said to be a gamma RV with parameter α, λ , if its probability density function is given, for $\alpha > 0, \lambda > 0$, by

$$f(x) = \begin{cases} \frac{\lambda e^{-\lambda x} (\lambda x)^{\alpha-1}}{\Gamma(\alpha)}, & \text{if } x \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

where

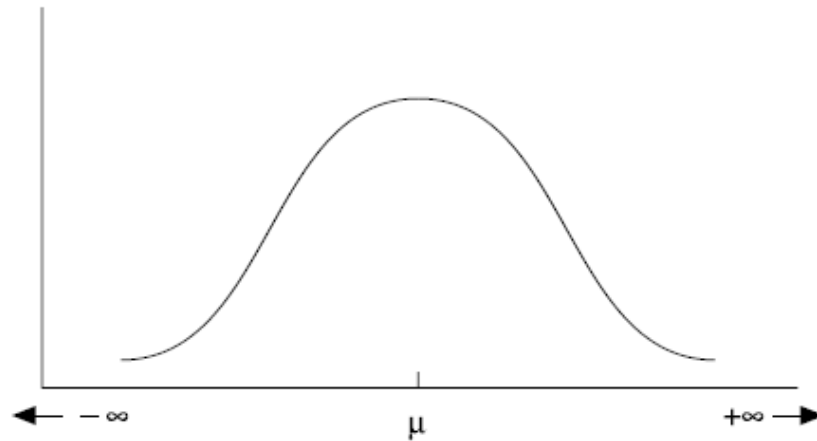
$$\Gamma(\alpha) = \int_0^{\infty} e^{-x} x^{\alpha-1} dx$$

Example of Continuous RVs: Normal RV

- A continuous RV is said to be a normal RV with parameter μ and σ^2 , if its probability density function is given by

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

for $-\infty < x < \infty$.



Continuous probability distribution	Probability density function, $f(x)$	Moment generating function, $\phi(t)$	Mean	Variance
Uniform over (a, b)	$f(x) = \begin{cases} \frac{1}{b-a}, & a < x < b \\ 0, & \text{otherwise} \end{cases}$	$\frac{e^{tb} - e^{ta}}{t(b-a)}$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$
Exponential with parameter $\lambda > 0$	$f(x) = \begin{cases} \lambda e^{-\lambda x}, & x > 0 \\ 0, & x < 0 \end{cases}$	$\frac{\lambda}{\lambda - t}$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$
Gamma with parameters $(n, \lambda), \lambda > 0$	$f(x) = \begin{cases} \frac{\lambda e^{-\lambda x} (\lambda x)^{n-1}}{(n-1)!}, & x \geq 0 \\ 0, & x < 0 \end{cases}$	$\left(\frac{\lambda}{\lambda - t} \right)^n$	$\frac{n}{\lambda}$	$\frac{n}{\lambda^2}$
Normal with parameters (μ, σ^2)	$f(x) = \frac{1}{\sqrt{2\pi}\sigma} \times \exp\{-(x - \mu)^2/2\sigma^2\},$ $-\infty < x < \infty$	$\exp \left\{ \mu t + \frac{\sigma^2 t^2}{2} \right\}$	μ	σ^2

Discrete probability distribution	Probability mass function, $p(x)$	Moment generating function, $\phi(t)$	Mean	Variance
Binomial with parameters n, p , $0 \leq p \leq 1$	$\binom{n}{x} p^x (1-p)^{n-x}$, $x = 0, 1, \dots, n$	$(pe^t + (1-p))^n$	np	$np(1-p)$
Poisson with parameter $\lambda > 0$	$e^{-\lambda} \frac{\lambda^x}{x!}$, $x = 0, 1, 2, \dots$	$\exp\{\lambda(e^t - 1)\}$	λ	λ
Geometric with parameter $0 \leq p \leq 1$	$p(1-p)^{x-1}$, $x = 1, 2, \dots$	$\frac{pe^t}{1 - (1-p)e^t}$	$\frac{1}{p}$	$\frac{1-p}{p^2}$
Continuous probability distribution	Probability density function, $f(x)$	Moment generating function, $\phi(t)$	Mean	Variance
Uniform over (a, b)	$f(x) = \begin{cases} \frac{1}{b-a}, & a < x < b \\ 0, & \text{otherwise} \end{cases}$	$\frac{e^{tb} - e^{ta}}{t(b-a)}$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$
Exponential with parameter $\lambda > 0$	$f(x) = \begin{cases} \lambda e^{-\lambda x}, & x > 0 \\ 0, & x < 0 \end{cases}$	$\frac{\lambda}{\lambda - t}$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$
Gamma with parameters (n, λ) , $\lambda > 0$	$f(x) = \begin{cases} \frac{\lambda e^{-\lambda x} (\lambda x)^{n-1}}{(n-1)!}, & x \geq 0 \\ 0, & x < 0 \end{cases}$	$\left(\frac{\lambda}{\lambda - t}\right)^n$	$\frac{n}{\lambda}$	$\frac{n}{\lambda^2}$
Normal with parameters (μ, σ^2)	$f(x) = \frac{1}{\sqrt{2\pi}\sigma} \times \exp\{-(x-\mu)^2/2\sigma^2\}$, $-\infty < x < \infty$	$\exp\left\{\mu t + \frac{\sigma^2 t^2}{2}\right\}$	μ	σ^2

Expectation of a function of a RV

- The expected value of X , $E[X]$, is a **weighted average** of the possible values that X can take on, each value being weighted by the probability that X assumes that value.

- If RV X is a discrete RV having a pmf $p(x)$, then $E[X]$ is defined by

$$E[X] = \sum_{x:p(x)>0} xp(x)$$

- Ex. 2.16~2.19] find the followings
 - Expectation of a Bernoulli Random Variable
 - Expectation of a Binomial Random Variable
 - Expectation of a Geometric Random Variable
 - Expectation of a Poisson Random Variable

- If RV X is a continuous RV having a pdf $f(x)$, then $E[X]$ is defined by

$$E[X] = \int_{-\infty}^{\infty} xf(x) dx$$

- Ex. 2.20~2.22] find the followings
 - Expectation of a Uniform Random Variable
 - Expectation of a Exponential Random Variable
 - Expectation of a Normal Random Variable

Discrete probability distribution	Probability mass function, $p(x)$	Moment generating function, $\phi(t)$	Mean	Variance
Binomial with parameters n, p , $0 \leq p \leq 1$	$\binom{n}{x} p^x (1-p)^{n-x}$, $x = 0, 1, \dots, n$	$(pe^t + (1-p))^n$	np	$np(1-p)$
Poisson with parameter $\lambda > 0$	$e^{-\lambda} \frac{\lambda^x}{x!}$, $x = 0, 1, 2, \dots$	$\exp\{\lambda(e^t - 1)\}$	λ	λ
Geometric with parameter $0 \leq p \leq 1$	$p(1-p)^{x-1}$, $x = 1, 2, \dots$	$\frac{pe^t}{1 - (1-p)e^t}$	$\frac{1}{p}$	$\frac{1-p}{p^2}$

Continuous probability distribution	Probability density function, $f(x)$	Moment generating function, $\phi(t)$	Mean	Variance
Uniform over (a, b)	$f(x) = \begin{cases} \frac{1}{b-a}, & a < x < b \\ 0, & \text{otherwise} \end{cases}$	$\frac{e^{tb} - e^{ta}}{t(b-a)}$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$
Exponential with parameter $\lambda > 0$	$f(x) = \begin{cases} \lambda e^{-\lambda x}, & x > 0 \\ 0, & x < 0 \end{cases}$	$\frac{\lambda}{\lambda - t}$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$
Gamma with parameters (n, λ) , $\lambda > 0$	$f(x) = \begin{cases} \frac{\lambda e^{-\lambda x} (\lambda x)^{n-1}}{(n-1)!}, & x \geq 0 \\ 0, & x < 0 \end{cases}$	$\left(\frac{\lambda}{\lambda - t} \right)^n$	$\frac{n}{\lambda}$	$\frac{n}{\lambda^2}$
Normal with parameters (μ, σ^2)	$f(x) = \frac{1}{\sqrt{2\pi}\sigma} \times \exp\{-(x - \mu)^2/2\sigma^2\},$ $-\infty < x < \infty$	$\exp\left\{\mu t + \frac{\sigma^2 t^2}{2}\right\}$	μ	σ^2

Expectation of a function of a RV

- Given a RV X and its probability distribution, what is the expectation of a function of X ?
 1. Since $g(X)$ is itself a random variable, it must have a probability distribution, which should be computable from a knowledge of the distribution of X . Once we have obtained **the distribution of $g(X)$** , we can then compute $E[g(X)]$ by the definition of the expectation.
 2. Another way is to compute the expectation of a function of X from **a knowledge of the distribution of X** . See Proposition 2.1.

Expectation of a function of a RV

- Ex. 2.23] Suppose X has the following probability mass function:
 $p(0) = 0.2, p(1) = 0.5, p(2) = 0.3$. Calculate $E[X^2]$.
 - Letting $Y = X^2$, we have that Y is a random variable that can take on one of the values $0^2, 1^2, 2^2$ with respective probabilities
$$\begin{aligned}p_Y(0) &= P\{Y = 0^2\} = 0.2, \\p_Y(1) &= P\{Y = 1^2\} = 0.5, \\p_Y(4) &= P\{Y = 2^2\} = 0.3\end{aligned}$$
 - Hence, $E[X^2] = E[Y] = 0(0.2) + 1(0.5) + 4(0.3) = 1.7$
 - Note that $E[X^2] \neq E[X]^2$

Expectation of a function of a RV

- Proposition 2.1

1. If X is a discrete random variable with probability mass function $p(x)$, then for any real-valued function g ,

$$E[g(X)] = \sum_{x:p(x)>0} g(x)p(x)$$

2. If X is a continuous random variable with probability density function $f(x)$, then for any real-valued function g ,

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)f(x) dx$$

- Ex 2.26] Let X be uniformly distributed over $(0,1)$, $E[X^3]$?

- $E[X^3] = \int_0^1 X^3 dx = \frac{1}{4}$

- Corollary 2.2

- If a and b are constants, then $E[aX + b] = aE[X] + b$

Expectation of a function of a RV

- The expected value of a RV X , $E[X]$, is also referred to as **the mean or the first moment of X** .
- The quantity $E[X^n]$, $n \geq 1$, is called the n_{th} moment of X
- The variance of a RV X , denoted by $Var(X)$, is defined by $Var(X) = E[(X - E[X])^2]$, deviation of X from the mean.
- $$\begin{aligned} Var(X) &= E[(X - E[X])^2] \\ &= E[X^2 - 2E[X]X + E[X]^2] \\ &= E[X^2] - E[2E[X]X] + E[E[X]^2] \\ &= E[X^2] - 2E[X]^2 + E[X]^2 \\ &= E[X^2] - E[X]^2 \end{aligned}$$

Expectation of a function of a RV

- Ex 2.27] $Var(X)$ of the normal RV with μ and σ .

- $Var(X) = E[(X - \mu)^2]$
$$= \int_{-\infty}^{\infty} (x - \mu)^2 \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx = \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} (x - \mu)^2 e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$
$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} y^2 e^{-\frac{y^2}{2}} dy$$
$$= \frac{\sigma^2}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{y^2}{2}} dy$$
$$= \sigma^2$$

Discrete probability distribution	Probability mass function, $p(x)$	Moment generating function, $\phi(t)$	Mean	Variance
Binomial with parameters n, p , $0 \leq p \leq 1$	$\binom{n}{x} p^x (1-p)^{n-x}$, $x = 0, 1, \dots, n$	$(pe^t + (1-p))^n$	np	$np(1-p)$
Poisson with parameter $\lambda > 0$	$e^{-\lambda} \frac{\lambda^x}{x!}$, $x = 0, 1, 2, \dots$	$\exp\{\lambda(e^t - 1)\}$	λ	λ
Geometric with parameter $0 \leq p \leq 1$	$p(1-p)^{x-1}$, $x = 1, 2, \dots$	$\frac{pe^t}{1 - (1-p)e^t}$	$\frac{1}{p}$	$\frac{1-p}{p^2}$

Continuous probability distribution	Probability density function, $f(x)$	Moment generating function, $\phi(t)$	Mean	Variance
Uniform over (a, b)	$f(x) = \begin{cases} \frac{1}{b-a}, & a < x < b \\ 0, & \text{otherwise} \end{cases}$	$\frac{e^{tb} - e^{ta}}{t(b-a)}$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$
Exponential with parameter $\lambda > 0$	$f(x) = \begin{cases} \lambda e^{-\lambda x}, & x > 0 \\ 0, & x < 0 \end{cases}$	$\frac{\lambda}{\lambda - t}$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$
Gamma with parameters (n, λ) , $\lambda > 0$	$f(x) = \begin{cases} \frac{\lambda e^{-\lambda x} (\lambda x)^{n-1}}{(n-1)!}, & x \geq 0 \\ 0, & x < 0 \end{cases}$	$\left(\frac{\lambda}{\lambda - t} \right)^n$	$\frac{n}{\lambda}$	$\frac{n}{\lambda^2}$
Normal with parameters (μ, σ^2)	$f(x) = \frac{1}{\sqrt{2\pi}\sigma} \times \exp\{-(x - \mu)^2/2\sigma^2\},$ $-\infty < x < \infty$	$\exp \left\{ \mu t + \frac{\sigma^2 t^2}{2} \right\}$	μ	σ^2