

# Confidence Intervals for Numerical Data

Ryan Miller

# Confidence Intervals

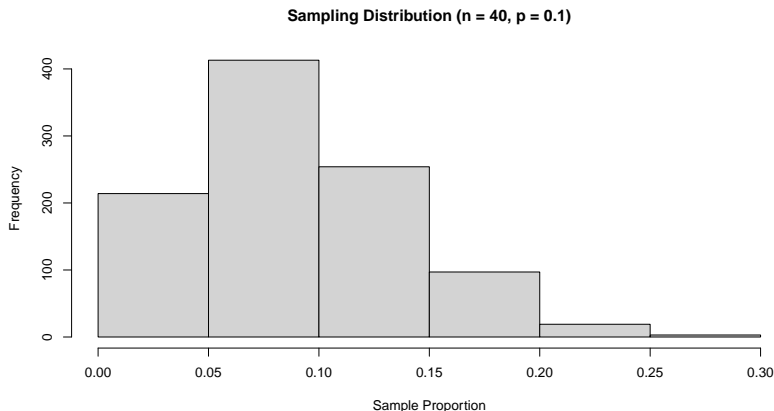
- ▶ Previously, we covered a couple different methods for constructing confidence interval estimates for categorical data (proportions)
- ▶ These approaches relied upon finding a reasonable *probability model* for the **sampling distribution** (of the sample proportion)
  - ▶ Binomial distribution (Clopper-Pearson intervals)
  - ▶ Normal distribution (CLT intervals)

# CLT Confidence Intervals for a Proportion

- ▶ Recall that the CLT approach was only valid when the conditions  $n\hat{p} \geq 10$  and  $n(1 - \hat{p}) \geq 10$  were both met
  - ▶ Thus, there are two scenarios where the CLT approach fails
- 1) The sample proportion is too close a boundary of either 0 and 1 (relative to the sample size)
- 2) The sample size is too small (and the proportion is away from the boundaries)

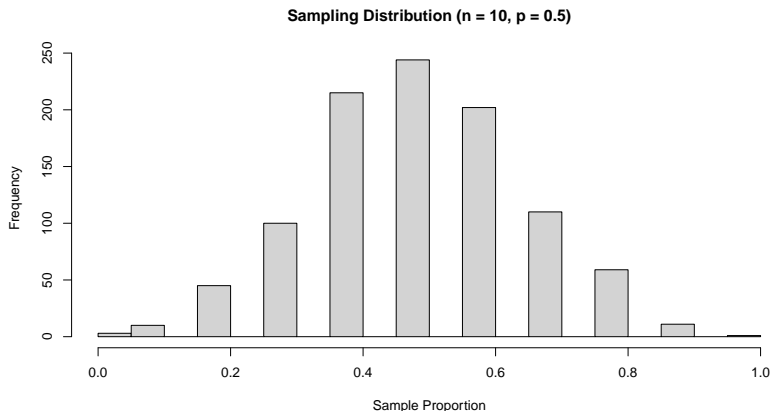
# CLT Confidence Intervals for a Proportion

- ▶ In this scenario ( $n = 40$  and  $p = 0.1$ ), notice  $n\hat{p} \approx 4$  and the sampling distribution is *skewed right*



# CLT Confidence Intervals for a Proportion

- In this scenario ( $n = 10$  and  $p = 0.5$ ), notice  $n * \hat{p} \approx 5$  and the sampling distribution is symmetric, but it is too discrete to be properly represented by a Normal curve



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  - ▶ In the second scenario ( $n = 10$  and  $p = 0.5$ ), 871 of 1000 “95% confidence” intervals contain  $p$

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- ▶ Fortunately, we could use the binomial distribution to calculate Clopper-Pearson (Exact Binomial) intervals in situations like these



# How About Numerical Data?

- ▶ Now let's consider a numeric random variable,  $X$  (ie: an individual's height, income, cholesterol, etc.)
  - ▶ Central Limit Theorem suggests the following sampling distribution for  $\bar{x}$ , the sample average of  $X$

$$\bar{x} \sim N(E(X) = \mu, \sqrt{\text{Var}(X)/n} = \frac{\sigma}{\sqrt{n}})$$

- ▶ This result suggests confidence intervals of the form:

$$\bar{x} \pm z^* \frac{\sigma}{\sqrt{n}}$$

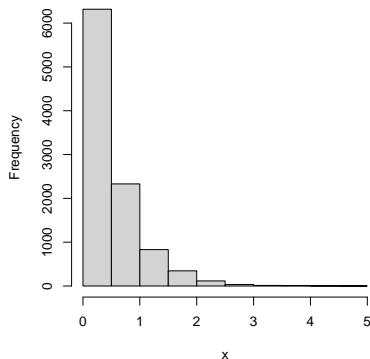
# CLT Confidence Intervals for a Mean

- ▶ As we saw with categorical data, these CLT-based intervals will only be valid when the *sampling distribution* is approximately Normal
  - ▶ For numerical data, there are two ways this could occur:
    - 1) The sample size is relatively large ( $n \geq 30$ ), regardless of how the population is distributed
    - 2) The sample size is small, but the *population distribution* is Normal

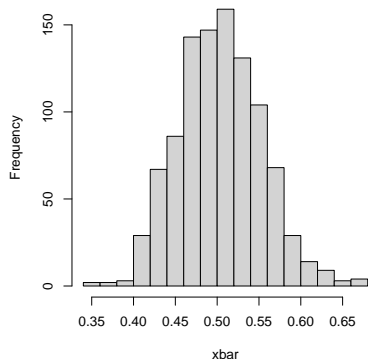
# CLT Confidence Intervals for a Mean

- ▶ The sampling distribution of  $\bar{x}$  is approximately Normal for  $n = 100$ , even when the population is heavily right-skewed

Population Dist

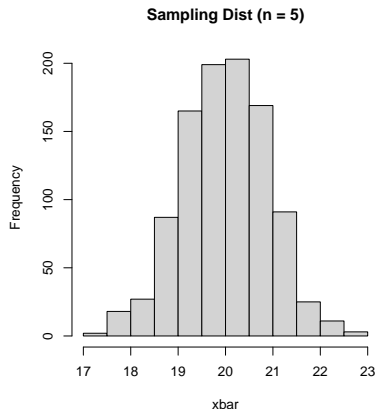
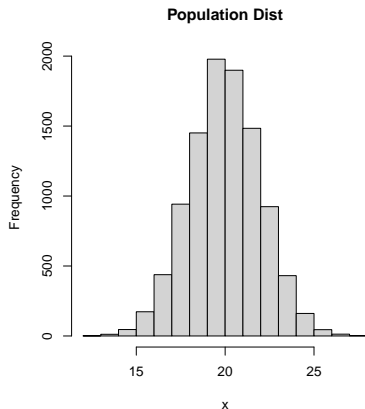


Sampling Dist (n = 100)



# CLT Confidence Intervals for a Mean

- ▶ The sampling distribution of  $\bar{x}$  is approximately Normal for very small samples ( $n = 5$ ) if the population is Normally distributed



- ▶ Now consider using the 1,000 different samples shown in the previous histograms to construct a 95% confidence interval estimate of  $\mu$  (the population mean)
  - ▶ This would look like  $\bar{x} \pm 1.96 \frac{s}{\sqrt{n}}$  for each sample (recognize that we're using the *sample standard deviation*,  $s$ , as an estimate of the population standard deviation,  $\sigma$ )
- ▶ How many of the 1,000 intervals would you expect to contain  $\mu$ ?

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- ▶ How many of the 1,000 intervals would you expect to contain  $\mu$ ?
  - ▶ In the first scenario ( $n = 100$  and right-skewed population), 953 of 1000 these “95% confidence” intervals contain  $\mu$
  - ▶ In the second scenario ( $n = 5$  and Normal population), 924 of 1000 “95% confidence” intervals contain  $\mu$

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- ▶ In 1906, Gosset took a leave of absence to go work with Karl Pearson (creator of the correlation coefficient) on the problem

# Student's $t$ -distribution

- ▶ Gosset discovered the flaw was due to using the sample standard deviation,  $s$ , in place of the population standard deviation,  $\sigma$ 
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- ▶ Simply “plugging in”  $s$  into the CLT result introduces a new source of variability (due to the imperfect estimation of  $\sigma$ )
  - ▶ Not accounting for this additional variability is the flaw in the previously constructed confidence intervals ( $n = 5$ , Normal population scenario)

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- ▶ Usually the person who discovers an important results gets to name it
  - ▶ However, Gosset had to publish his work under the name “Student” because Guinness didn't want competitors knowing it employed statisticians!
  - ▶ Gosset's result, called Student's  $t$ -distribution, is among the most widely-used statistical results of all time

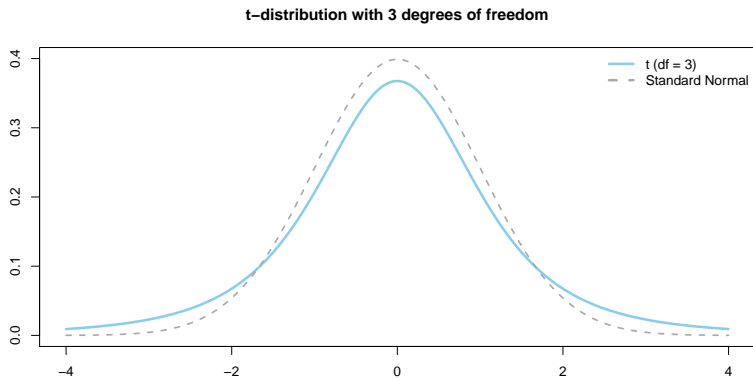
# The $t$ -distribution

- ▶ The  $t$ -curve looks much like a standard Normal curve, but it has *thicker tails* to properly account for additional variability that results from estimating  $\sigma$  using  $s$ 
  - ▶ The amount of additional variability is linked to the sample size through a parameter known as *degrees of freedom* (df)

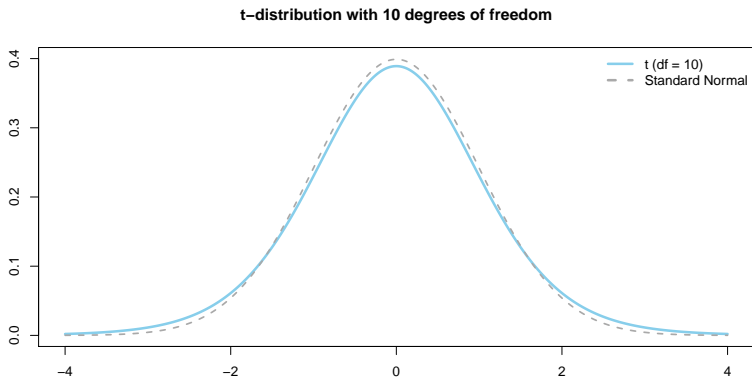
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  - ▶ The amount of additional variability is linked to the sample size through a parameter known as *degrees of freedom* ( $df$ )
- ▶ Similar to Chi-square testing, this is a reference to the number of unique pieces of information available to estimate  $\sigma$ 
  - ▶ If  $\bar{x}$  is assumed known, the sum of deviations  $\sum_{i=1}^n (x_i - \bar{x})$  must add up to zero, thus only  $n - 1$  deviations are necessary to know everything
- ▶ So, when applying the  $t$ -distribution to the mean of a single numeric variable,  $df = n - 1$

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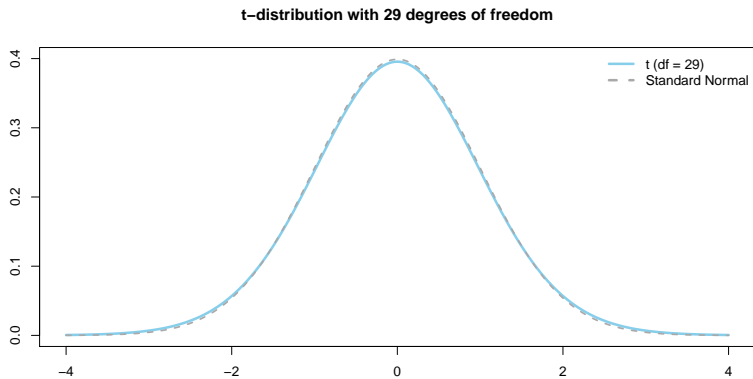


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# How to use the $t$ -distribution

- ▶ For a single mean, we construct a  $P\%$  confidence interval via:

$$\bar{x} \pm t_{n-1}^* \frac{s}{\sqrt{n}}$$

- ▶  $t_{n-1}^*$  is the percentile defining the middle  $P\%$  of the  $t$ -distribution with  $n - 1$  degrees of freedom

```
qt(.975, df = 10)
```

```
## [1] 2.228139
```

```
qt(.975, df = 100)
```

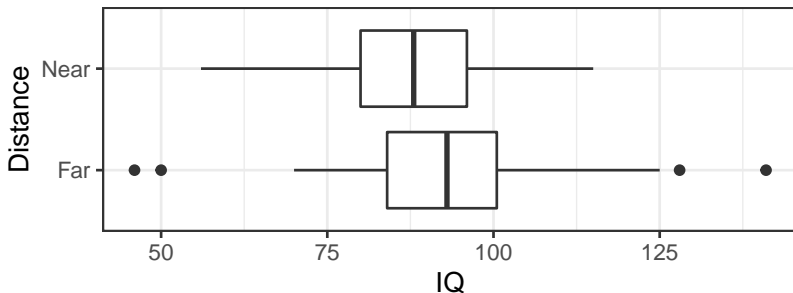
```
## [1] 1.983972
```

```
qt(.975, df = 1000)
```

```
## [1] 1.962339
```

## Example - Lead Exposure and IQ

Researchers in El Paso, TX measured the IQ scores (age-adjusted) of 57 children who lived within 1 mile of a lead smelter and 67 children who lived at least 1 mile away



1. Do these data appear to be normally distributed?
2. Could there be an association between Distance and IQ?

## Example - Lead Exposure and IQ

- ▶ Load these data into R using the following code
- ▶ Then use the  $t$ -distribution to come up with separate 90% confidence intervals for the population IQ in each group (Near and Far)

```
data <- read.csv("https://remiller1450.github.io/data/LeadIQ.csv")
near <- data$IQ[data$Distance == "Near"]
far <- data$IQ[data$Distance == "Far"]
```

## Example (solution, “Near” group)

```
xbar_near = mean(near)
sd_near = sd(near)
n_near = length(near)
ts_near = qt(.95, df = n_near - 1)

lower_end = xbar_near - ts_near*sd_near/sqrt(n_near)
upper_end = xbar_near + ts_near*sd_near/sqrt(n_near)
c(lower_end, upper_end)

## [1] 86.49585 91.89012
```

## Example (solution - “Far” group)

```
xbar_far = mean(far)
sd_far = sd(far)
n_far = length(far)
ts_far = qt(.95, df = n_far - 1)

lower_end = xbar_far - ts_far*sd_far/sqrt(n_far)
upper_end = xbar_far + ts_far*sd_far/sqrt(n_far)
c(lower_end, upper_end)

## [1] 89.43076 95.94238
```

# Lead Exposure and IQ - Revisited

The previous approach was sub-optimal, it is better to look at the difference in means (rather than each mean separately). To understand why this is, we'll need a new CLT result:

$$\bar{x}_1 - \bar{x}_2 \sim N\left(\mu_1 - \mu_2, \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}\right)$$

Notice how the standard error of a difference in means is *always less than* sum of the standard errors of each mean separately:

$$\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}} < \sqrt{\frac{\sigma_1^2}{n_1}} + \sqrt{\frac{\sigma_2^2}{n_2}}$$

# Lead Exposure and IQ - Degrees of Freedom (difference in means)

- ▶ This result requires estimates of both  $\sigma_1$  and  $\sigma_2$ , so you might be wondering how to determine the correct degrees of freedom. The answer is quite messy. . .

$$df = \frac{\left(\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}\right)^2}{\frac{s_1^2/n_1}{n_1-1} + \frac{s_2^2/n_2}{n_2-1}}$$

- ▶ Don't ever calculate this by hand, use software!



# Lead Exposure and IQ - Degrees of Freedom

The `t.test` function provides a proper 90% CI estimate for the difference in means (far - near)

```
t.test(x = far, y = near, conf.level = .90)$conf.int
```

```
## [1] -0.702856  7.690025
```

```
## attr("conf.level")
```

```
## [1] 0.9
```

# Comments on the $t$ -distribution

- ▶ The  $t$ -distribution should *always* be used when using a sample of numerical data to estimate a population mean (or difference in means)
  - ▶ That said, when  $n \geq 30$  there isn't much of a difference relative to using the Normal distribution

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  - ▶ That said, when  $n \geq 30$  there isn't much of a difference relative to using the Normal distribution
- ▶ When the sample is small *and* the population is Normally distributed, the  $t$ -distribution is necessary for valid statistical inference
  - ▶ When the sample is small *and* the population is skewed, we're out of luck (at least for now)

# Next Steps

- ▶ Next we'll use the  $t$ -distribution as the basis for hypothesis testing of population means (this is called the  $t$ -test)
  - ▶ After that, we'll revisit the situation where the sample is small and the population is skewed and explore a creative approach to achieving valid statistical inference

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  - ▶ After that, we'll revisit the situation where the sample is small and the population is skewed and explore a creative approach to achieving valid statistical inference
- ▶ For now, the main takeaway from this lecture is the  $t$ -distribution
  - ▶ You should know when to use it (numerical data)
  - ▶ You should know why it's important (extra uncertainty using  $s$  as an estimate of  $\sigma$ )
  - ▶ You should know how to use it (the `qt` function, degrees of freedom, etc.)