Sampling Distributions and Central Limit Theorem

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 - Sampling Distributions
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- Lately we've been discussing random variables, which are used to represent the numeric outcome of a random process
- ▶ The act of data collection is a random process
 - ▶ We don't know which cases from the population will be sampled
 - We don't know which study participants will be randomized to the treatment/control group
- Further, any descriptive summary of our sample data (ie: means, proportions, correlations, etc.) is the observed outcome of a random variable

The Sample Average as a Random Variable

► Consider a sample of *n* cases from a population, the sample average is calculated:

$$\bar{x} = \frac{x_1 + x_2 + x_3 + \dots + x_n}{n}$$

▶ In addition to usefulness in describing the center of a quantitative variable's distribution, lots of statistical theory has been developed for understanding variability in sample averages

Proportions are Averages

- Now, consider a binary categorical variable
 - ▶ Because binary variables involve only two categories, we can map their outcomes to the numeric values of 0 and 1
 - ► For example, consider a coin flip, we could map the outcome "Heads" to "1" and the outcome "Tails" to "0"

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 - ► For example, consider a coin flip, we could map the outcome "Heads" to "1" and the outcome "Tails" to "0"
- ▶ By coding the outcomes using 1s and 0s, we can see that the sample proportion is also an average:

$$\hat{p} = \frac{1 + 0 + 1 + 1 + 0 + \dots + 1}{n}$$

▶ If we mapped "Heads" to a value of "1", \hat{p} would refer to the proportion of heads in our sample

The Distribution of the Sample Proportion

- ► According to the US Census, 27.5% of the adult population are college graduates
- ▶ Randomly sampling *n* adults represents a *random process*
 - The proportion of college graduates in a sample, \hat{p} , is a random variable

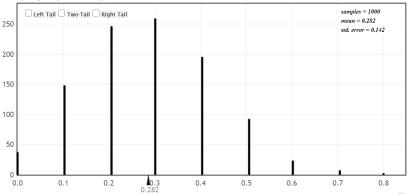
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- ightharpoonup Randomly sampling n adults represents a random process
 - The proportion of college graduates in a sample, \hat{p} , is a random variable
- Let's explore some different outcomes of this random variable for two different sampling protocols: random samples of size n=10, and random samples of size n=100

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 - ▶ Thus, the sample space is: $\{0/10, 1/10, 2/10, ..., 10/10\}$

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 - ▶ Thus, the sample space is: $\{0/10, 1/10, 2/10, ..., 10/10\}$
- Nather than trying to perform probability calculations, we'll instead look at repeatedly drawing different random samples (of size n=10) to judge the likelihood of each of these outcomes

Sampling Dotplot of Proportion



Each dot represents the proportion of college graduates in a different random sample of size n = 10

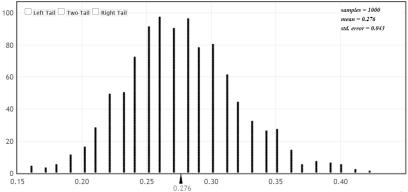


▶ Due to the relatively small number of discrete outcomes, it's reasonable to use a table to convey a probability model for the sample proportion:

Sample Proportion ($n = 10$)	Probability
0/10	40/1000 = 0.04
1/10	150/1000 = 0.15
2/10	250/1000 = 0.25
3/10	270/1000 = 0.27
4/10	190/1000 = 0.19
10/10	0/1000 = 0

- For a random sample of n=100, there are now 101 discrete outcomes that could be observed for the sample proportion $\{0/100, 1/100, 2/100, \dots, 100/100\}$
 - At this point, it's impractical to write-out probabilities using a table, instead it makes more sense to treat the sample proportion as a *continuous random variable*





Notice this distribution is roughly bell-shaped, it's centered at the population proportion (approximately), and has a spread described by the standard error



A Normal Model?

- ➤ You might be thinking that we can apply a Normal model here, but getting the proper Normal distribution requires us to get the center and spread correct
 - StatKey will report these values (which it finds via simulation), but we'll discuss their origin at length in the next couple of videos

Sampling Distributions

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- For example, different random samples of size n=100 (of US adults) yield sample proportions (of college graduates) that are on average 0.043 off from their expected value of 0.275
 - Different random samples of size n = 10 yield sample proportions that are on average 0.142 off from their expected value of 0.275

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- \triangleright For example, different random samples of size n=100 (of US adults) yield sample proportions (of college graduates) that are on average 0.043 off from their expected value of 0.275
 - ▶ Different random samples of size n = 10 yield sample proportions that are on average 0.142 off from their expected value of 0.275
 - This should make sense, larger samples contain more information about the population and therefore provide estimates that are more reliable (ie: tend to have less sampling error)

Closing Remarks

- So far, we've seen that descriptive statistics calculated using sample data can be viewed as a realized outcome of a random variable
- ➤ To understand the role of random chance in observed sample data, we can explore the sampling distributions of these random variables
 - If we can come up with an accurate probability model for the sampling distribution, we can use it to evaluate random chance as a viable explanation for trends that occur in our sample data

John Kerrich

- ▶ John Kerrich, a South African mathematician, was visiting Copenhagen in 1940
- When Germany invaded Denmark he was sent to an internment camp, where he spend the next five years
- ➤ To pass time, Kerrich conducted experiments exploring sampling and probability theory
 - One of these experiments involved flipping a coin 10,000 times

Kerrich's Experiment and Probability

- We know that a fair coin shows "Heads" with a probability of 50%
- ➤ So, in a random sample of *n* coin flips, we'd expect roughly even numbers of "Heads" and "Tails"
 - ► We'll explore the results of Kerrich's experiment to see why the sample average is so special

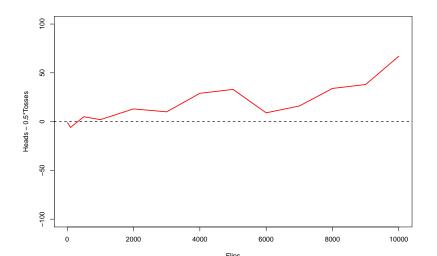
Kerrich's Results

Number of Tosses (n)	Number of Heads	Heads - 0.5*Tosses
10	4	-1
100	44	-6
500	255	5
1,000	502	2
2,000	1,013	13
3,000	1,510	10
4,000	2,029	29
5,000	2,533	33
6,000	3,009	9
7,000	3,516	16
8,000	4,034	34
9,000	4,538	38
10,000	5,067	67



Kerrich's Results

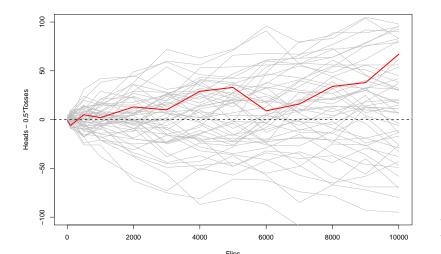
It seems like the number of heads and tails are actually getting further apart... could this be a fluke?





Kerrich's Experiment (repeated 50 times)

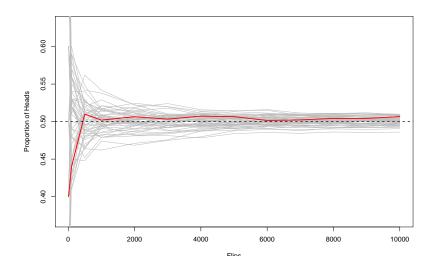
No, the phenomenon occurs systematically when repeating Kerrich's experiment





Kerrich's Experiment (sample proportions)

The *sample proportion* of heads behaves exactly as we'd expect, but why?





Central Limit Theorem

- Suppose $X_1, X_2, ..., X_n$ are independent random variables with a common expected value E(X) and variance Var(X) (see previous notes for definitions of these two terms)
- Let \bar{X} denote the average of all n random variables, **Central** Limit Theorem (CLT) states:

$$\sqrt{n}\Big(rac{ar{X}-E(X)}{\sqrt{Var(X)}}\Big)
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$$\sqrt{n}\left(\frac{\bar{X}-E(X)}{\sqrt{Var(X)}}\right) \to N(0,1)$$

Often it is more useful to think of CLT in the following way (which abuses notation):

$$\bar{X} \sim N\bigg(E(X), rac{SD(X)}{\sqrt{n}}\bigg)$$



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- ► The sample proportion is comprised of *n* different binary variables (taking on values of 1 and 0)
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 - E(X) = p * 1 + 0 * (1 p) = p
 - Var(X) = $p * (1-p)^2 + (1-p) * (0-p)^2 = p * (1-p)$
- ▶ Thus, the *sampling distribution* of sample proportions is:

$$\hat{p} \sim N(p, \sqrt{p(1-p)/n})$$

The Power of CLT

- Central Limit Theorem is one of the most important theoretical results in all of statistics
- ▶ In real-world applications, it is nearly impossible to know the probability distribution of something that is only observed once (remember that real researchers can only afford to collect a single sample)

The Power of CLT

- Central Limit Theorem is one of the most important theoretical results in all of statistics
- In real-world applications, it is nearly impossible to know the probability distribution of something that is only observed once (remember that real researchers can only afford to collect a single sample)
- But by focusing on the sample average this isn't an issue, as CLT provides us the distribution of sample averages
 - That is, we are able to use CLT to understand the sampling variability of our study, despite only getting to see a single sample!

Example

- Let's consider a random sample of n = 100 coin flips
 - ▶ What proportion of heads might we expect? It'll likely be close to 50%, but we know there's sampling variability, the question is how much...

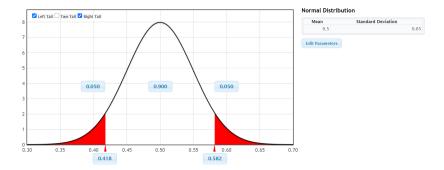
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- Let's consider a random sample of n = 100 coin flips
 - ▶ What proportion of heads might we expect? It'll likely be close to 50%, but we know there's sampling variability, the question is how much...
- Each coin flip is a random variable an expected value of 0.5, so Central Limit Theorem tells us that proportion of heads in random samples of n = 100 coin flips follows a Normal distribution:

$$\hat{p} \sim N(0.5, \sqrt{0.5(1-0.5)/100})$$

 \triangleright To understand the sampling variability of n = 100 coin flips, we might look at the *interval* that defines what we'd expect to see 90% of the time

Example



▶ We'd expect 90% of different random samples to result in sample proportions between 0.418 and 0.582



Assumptions

Using the Central Limit theorem to determine the distribution of sample averages is only appropriate when the following conditions are met:

- 1) Independence the cases in the sample (ie: the individual contributions to the sample average) are not related to each other
- 2) Large population less that 10% of the population is being sampled (otherwise removing the already sampled individuals has too much of an impact on the probability of selection)
- 3) Large sample $n\hat{p} \ge 10$ and $n(1-\hat{p}) \ge 10$

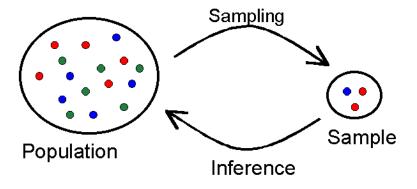
Most of the time, it's only the third condition that is problematic

Closing Remarks

- Central Limit theorem provides a theoretical basis for focusing on sample averages when attempting to characterize a population
 - Put differently, CLT allows us to understand the sampling variability of the sample average without needing to actually take multiple different samples!

Sampling and Statistical Inference

- ► A fundamental goal of statisticians is to use information from a sample to make reliable statements about a population
 - ► This idea is called **statistical inference**



 $Image\ credit:\ http://testofhypothesis.blogspot.com/2014/09/the-sample.html$



Statistical Inference - Notation

Statisticians use different notation to distinguish *population* parameters (things we want to know) from *estimates* (things derived from a sample). For a few common measures, this notation is summarized below:

	Population Parameter	Estimate (from sample)
Mean	μ	\bar{x}
Standard Deviation	σ	s
Proportion	р	p p
Correlation	ho	r
Regression	$eta_{f 0},eta_{f 1}$	b_0, b_1

For example, μ is the mean of the target population, while \bar{x} is the mean of the cases that ended up in the sample



Point Estimation

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- ▶ If a sampling protocol is *unbiased*, the sample average is a sensible estimate of the population mean
 - This is called a **point estimate**, referring to the fact that it is a single value
- From our study of sampling distributions, we know that the existence of sampling variability means a point estimate is almost certainly wrong (at least to some degree)
 - This suggests that we can more appropriately describe what we think is true of the population by reporting an interval estimate that accounts for sampling variability

Point vs. Interval Estimation

To summarize:

- ▶ Point estimation uses sample data to produce a *single "most likely" estimate* of a population characteristic, which will almost always miss the target (at least by some degree)
- ▶ **Interval estimation** uses sample data to produce a *range of plausible estimates* of a population characteristic, an approach that has a much better chance at capturing the truth

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An analogy:

Using only a point estimate is like fishing in a murky lake with a spear. We can throw a spear where we saw a fish, but we will probably miss. On the other hand, if we toss a net in that area, we have a good chance of catching the fish.

Margin of Error

Most interval estimates have the form:

Point Estimate \pm Margin of Error

We often report these intervals using only their endpoints:

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- We'd like the margin of error to be constructed in way that carries a quantifiable claim of precision
 - ie: 80% of the time an interval with this type of margin of error will contain the population characteristic
 - Without an accompanying claim regarding precision, reporting a margin of error is not particularly useful



So, what can we say about a population proportion, p, based upon an observed sample proportion, \hat{p} ? Consider a representative sample of 100 infants used to estimate the proportion of all babies who are born prematurely

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- True or false? "Although we don't know p, if we attach a large margin error to our point estimate, the interval estimate $14\% \pm 10\% = (4\%, 24\%)$ probably contains p"
 - False we don't know how reliable this margin of error is, perhaps an MOE of 10% is not wide enough



Conclusion

- ▶ This presentation introduces the idea of interval estimation
 - ► The key concept is that point estimates are almost always off, but by attaching a margin of error we can more reliably describe the population of interest
- ▶ In class this week, we'll further explore this concept and learn how to use sampling distributions to come up with interval estimates that have *meaningful margins of error*