Topics in instrumental variable estimation: structural mean models and bounds

Tom Palmer

MRC CAiTE Centre, School of Social and Community Medicine, University of Bristol, UK

February 2012

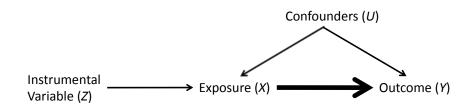




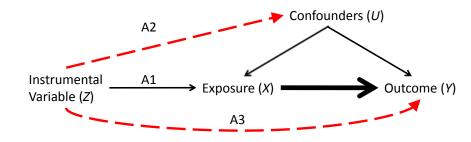
Outline

- 1. Introduction to instrumental variables
 - Assumptions
 - Applications
 - ▶ Mendelian randomization
 - ► Noncompliance in RCTs
 - ► Test of presence of effect
 - ► Estimators: ratio, two-stage least squares
- 2. Structural mean models
 - Potential outcomes and causal parameters
 - Additive SMM: G-estimation example
 - Multiplicative SMM: estimation with multiple instruments
- 3. Nonparametric bounds
 - Extensions
 - Limitations
- 4. Summary

1. Introduction to instrumental variables: Assumptions



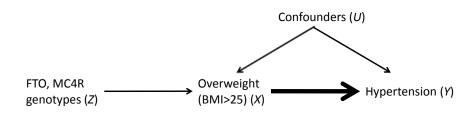
1. Introduction to instrumental variables: Assumptions



- A1. Instrumental variable associated with exposure $(Z \perp \!\!\!\! \perp X)$
- A2. Instrumental variable independent of confounders $(Z \perp \!\!\! \perp U)$
- A3. No direct effect of instrumental variable on outcome exclusion restriction $(Y \perp \!\!\! \perp Z \mid (X, U))$

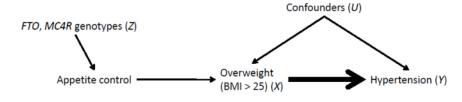
Application: Mendelian randomization

Use of genotypes robustly associated with exposures (from replicated genome-wide association studies, $P < 5 \times 10^{-8}$) as instrumental variables (Davey Smith & Ebrahim, 2003)



Application: Mendelian randomization

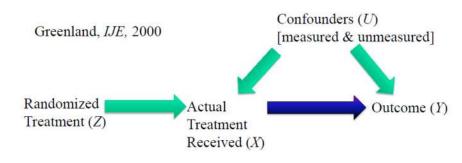
Use of genotypes robustly associated with exposures (from replicated genome-wide association studies, $P < 5 \times 10^{-8}$) as instrumental variables (Davey Smith & Ebrahim, 2003)



IV does not have to be causal for exposure

Application: randomized controlled trials

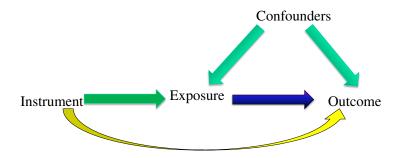
Correcting for noncompliance in randomized controlled trials



Test of presence of causal effect

If the IV conditions hold, then a test of the instrument-outcome association

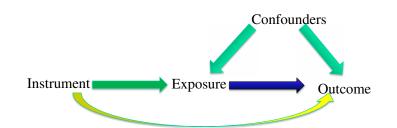
is a test for the presence of a causal effect of exposure on outcome



Ratio estimator

Assumptions: everything linear, binary instrument





Standard error & confidence interval from delta-method/Fieller's theorem (Thomas et al. *Ann Epi*, 2007)

Two-stage least squares estimator

Estimation with multiple instruments

Two-stage least squares

$$X = \beta_0 + \beta_{z1}Z_1 + ... + \beta_{zn}Z_n$$

Stage 1

- Regress Exposure on instrument/s
- Generate predicted values of exposure

$$\hat{X} = \hat{\beta}_0 + \hat{\beta}_Z Z$$

Stage 2

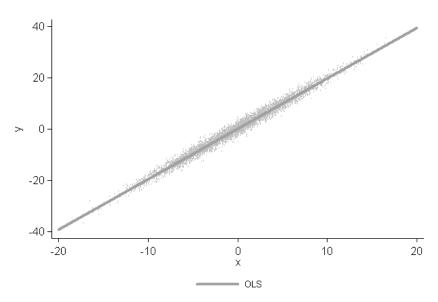
- Regress Outcome on predicted values of exposure
- (adjust standard errors)

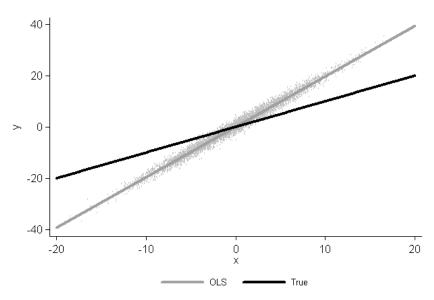
$$Y = \alpha + \beta_{XY} \hat{X}$$

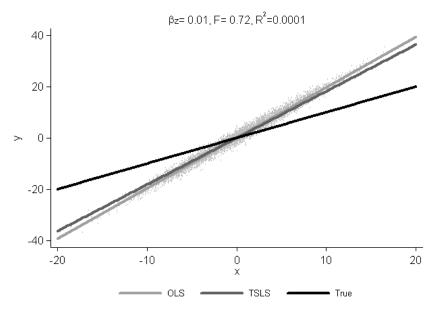
Stata comands: ivregress, ivreg2

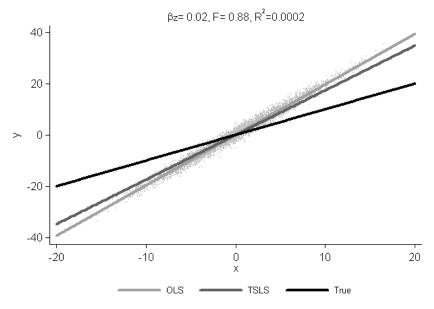
Automatically correct standard errors!!

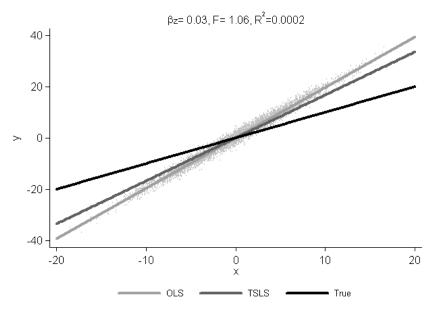
Binary outcome: parameter = risk difference

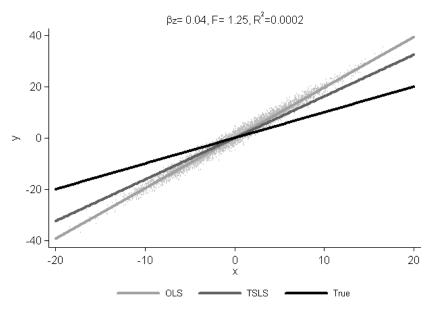


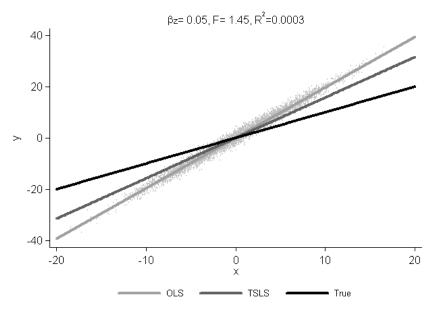


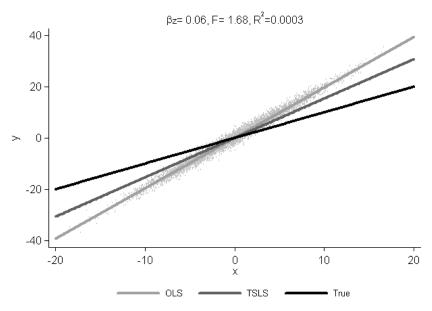


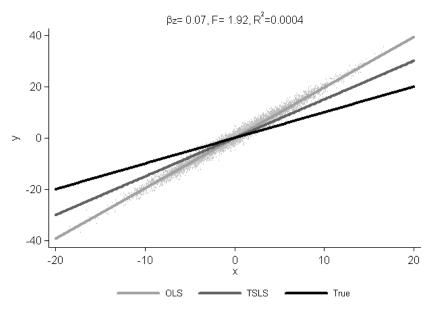


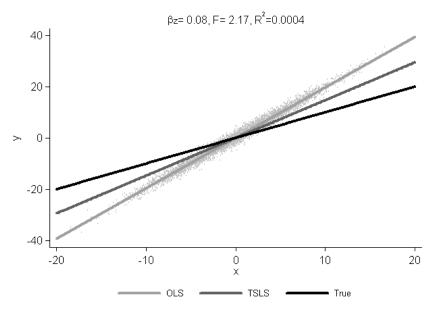


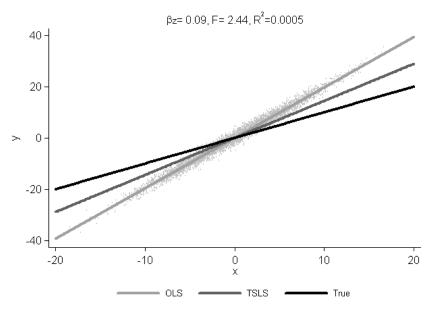


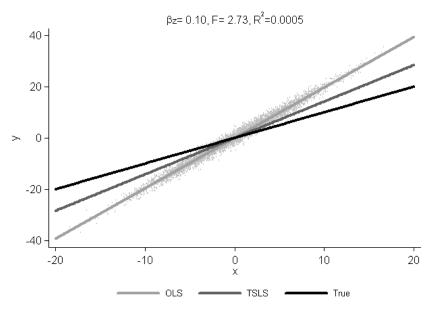


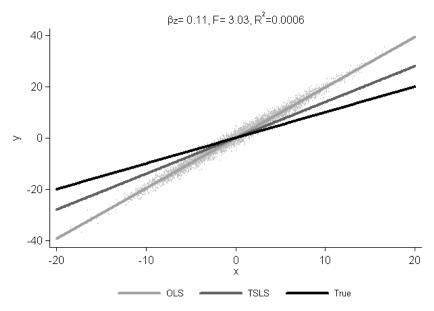


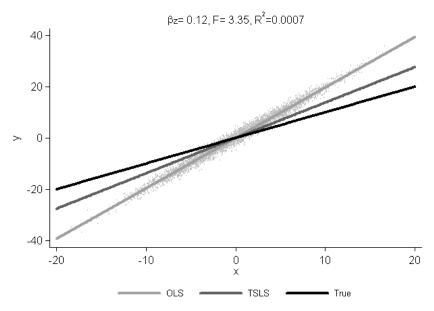


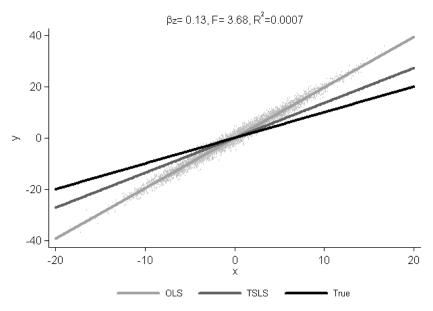


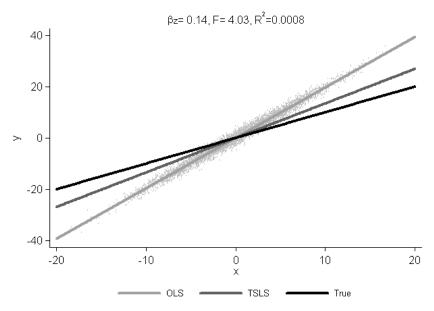


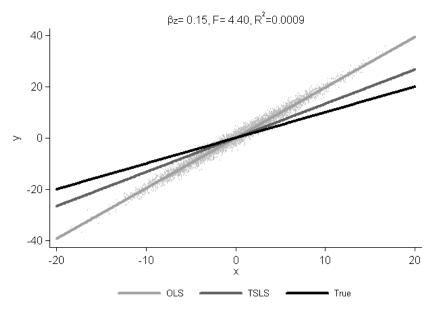


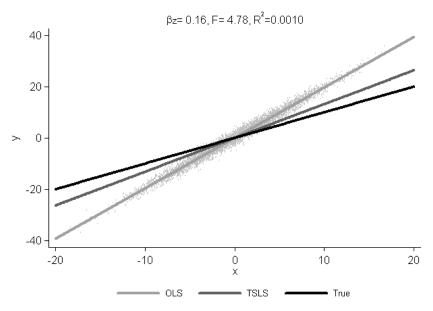


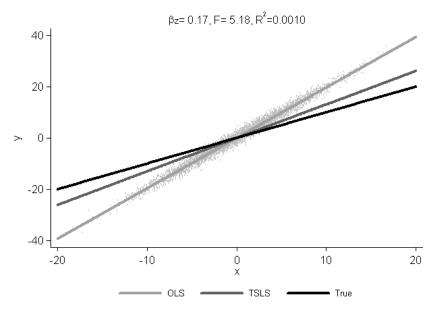


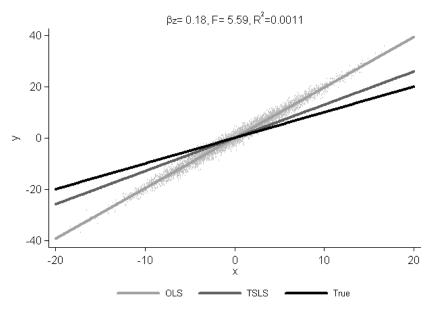


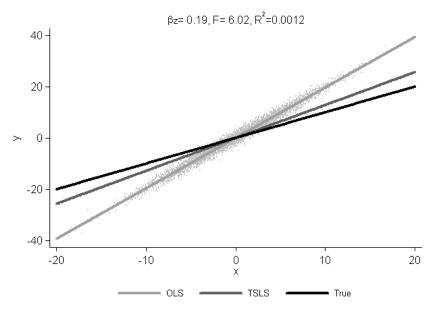


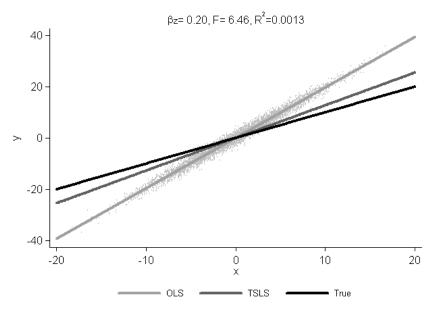


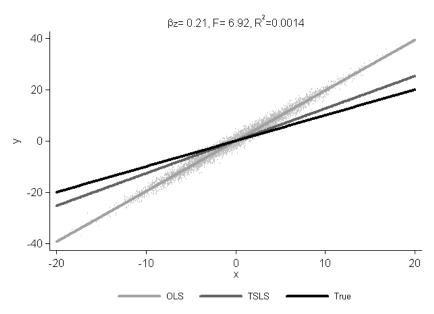


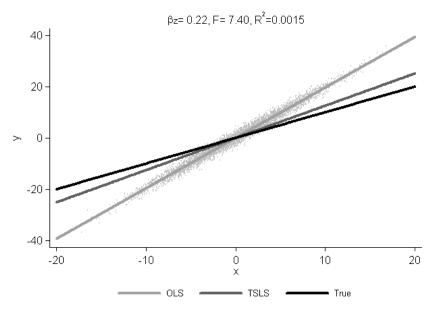


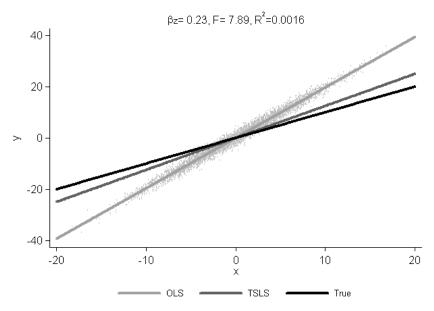


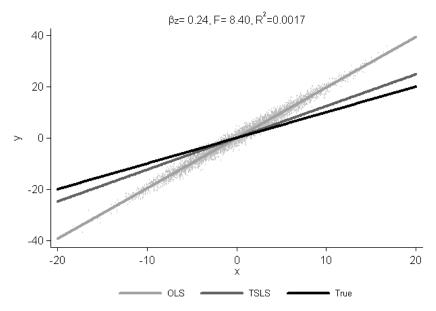


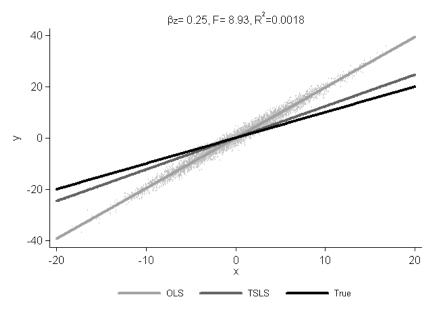


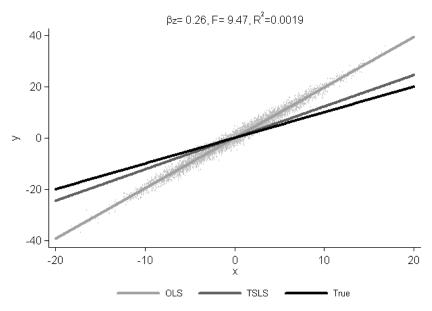


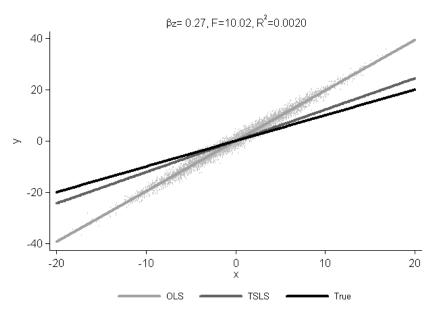


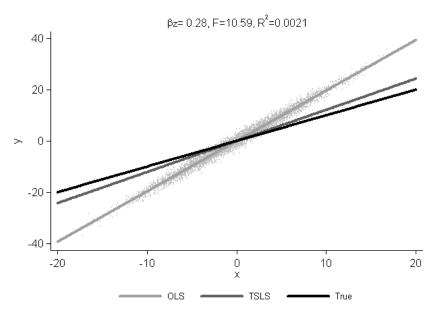


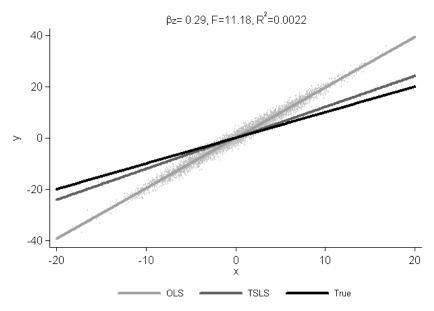


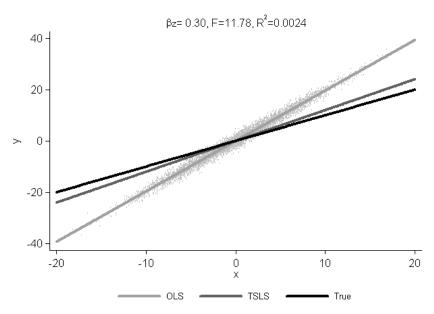






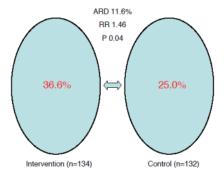






Additive SMM example I

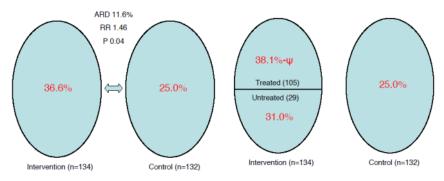
- ► Ten Have, Elliott, Joffe, Zanutto, & Datto, 2004 266 African American adults with high cholesterol and/or hypertension
- ► Control: usual care (conventional nutritional info)
- ▶ Intervention: usual care + home-based audio tapes
- ▶ Outcome: +ve response: beneficial change in cholesterol,
 - -ve response: otherwise



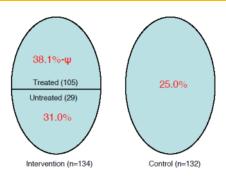
Additive SMM example I

- ► Ten Have et al., 2004 266 African American adults with high cholesterol and/or hypertension
- ► Control: usual care (conventional nutritional info)
- ▶ Intervention: usual care + home-based audio tapes
- ▶ Outcome: +ve response: beneficial change in cholesterol,

-ve response: otherwise



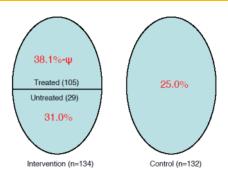
Additive SMM example II



Ratio estimate:
$$\psi = \frac{E(Y|Z=1) - E(Y|Z=0)}{E(X|Z=1) - E(X|Z=0)} = \frac{36.6 - 25.0}{105/134 - 0}$$

= 11.6/78.4 = 14.8% (95%CI 0.8%, 28.7%; $P = 0.04$)

Additive SMM example II



Ratio estimate:
$$\psi = \frac{E(Y|Z=1) - E(Y|Z=0)}{E(X|Z=1) - E(X|Z=0)} = \frac{36.6 - 25.0}{105/134 - 0}$$

= 11.6/78.4 = 14.8% (95%CI 0.8%, 28.7%; $P = 0.04$)

G-estimation: what would have happened if no-one was treated ASMM estimate: $(38.1-\psi)(105/134)+31.0(29/134)=25.0$ $\psi=(38.1\times105+31.0\times29-25.0\times134)/105=14.8\%$

Multiplicative SMM

Notation: X exposure/treatment, Y outcome, Z instrument, Y(X=0) exposure/treatment free potential outcome

Robins, Rotnitzky, & Scharfstein, 1999; Hernán & Robins, 2006

$$\begin{split} \log(E[Y|X,Z]) - \log(E[Y(0)|X,Z]) &= \psi X \\ \frac{E[Y|X,Z]}{E[Y(0)|X,Z]} &= \exp(\psi X) \\ \psi : & \text{ log causal risk ratio} \\ \text{Rearrange: } Y(0) &= Y \exp(-\psi X) \end{split}$$

Under the instrumental variable assumptions (Robins, 1989):

$$Y(0) \perp \!\!\! \perp Z$$

 $Y \exp(-\psi X) \perp \!\!\! \perp Z$

Under the instrumental variable assumptions (Robins, 1989):

$$Y(0) \perp \!\!\! \perp Z$$
 $Y \exp(-\psi X) \perp \!\!\! \perp Z$ trick: $Y \exp(-\psi X) - Y(0) \perp \!\!\! \perp Z$

Under the instrumental variable assumptions (Robins, 1989):

$$Y(0) \perp \!\!\! \perp Z$$
 $Y \exp(-\psi X) \perp \!\!\! \perp Z$ trick: $Y \exp(-\psi X) - Y(0) \perp \!\!\! \perp Z$

Moment conditions

$$Z = 0,1$$

$$E[(Y \exp(-\psi X) - Y(0))1] = 0$$

$$E[(Y \exp(-\psi X) - Y(0))Z_1] = 0$$

Under the instrumental variable assumptions (Robins, 1989):

$$Y(0) \perp \!\!\! \perp Z$$

$$Y \exp(-\psi X) \perp \!\!\! \perp Z$$
 trick: $Y \exp(-\psi X) - Y(0) \perp \!\!\! \perp Z$

Moment conditions

$$Z=0,1,2,3$$

Over-identified

$$E[(Y \exp(-\psi X) - Y(0))1] = 0$$

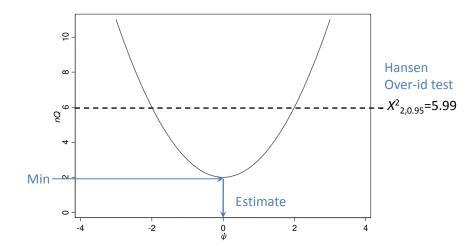
$$E[(Y \exp(-\psi X) - Y(0))Z_1] = 0$$

$$E[(Y \exp(-\psi X) - Y(0))Z_2] = 0$$

$$E[(Y \exp(-\psi X) - Y(0))Z_3] = 0$$

Generalised Method of Moments

Minimises quadratic form: $Q = m'W^{-1}m$



Implementation in Stata & R

Stata: gmm command

```
gmm (y*exp(-1*x*{psi}) - {ey0}), instruments(z1 z2 z3)
```

Implementation in Stata & R

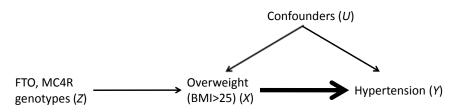
Stata: gmm command

```
gmm (y*exp(-1*x*{psi}) - {ey0}), instruments(z1 z2 z3)
```

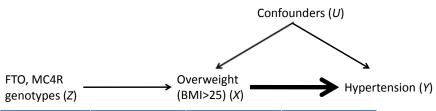
R: gmm package (Chaussé, 2010)

```
library(gmm)
msmmMoments <- function(theta,x){
    # extract variables from x
    Y <- x[,1]; X <- x[,2]; Z1 <- x[,3]; Z2 <- x[,4]; Z3 <- x[,5]
    # moments
    m1 <- (Y*exp(- X*theta[2]) - theta[1])
    m2 <- (Y*exp(- X*theta[2]) - theta[1])*Z1
    m3 <- (Y*exp(- X*theta[2]) - theta[1])*Z2
    m4 <- (Y*exp(- X*theta[2]) - theta[1])*Z3
    return(cbind(m1,m2,m3,m4))
}
fit <- gmm(msmmMoments, data, t0=c(0,0))</pre>
```

Copenhagen example descriptive statistics 1



Copenhagen example descriptive statistics 1



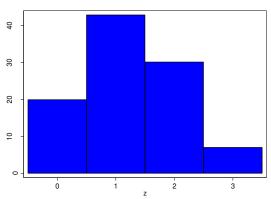
	No Hypertension	Hypertension	Total
Not	10,066	13,909	23,975
Overweight	42%	58%	
Overweight	6,906 22%	24,642 78%	31,548
Total	16,972	38,551	55,523
	31%	69%	χ² P<0.001

Risk ratio for hypertension 1.35 (1.32, 1.37)

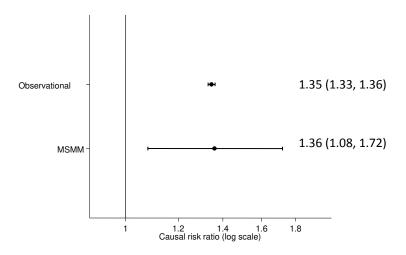
Copenhagen example descriptive statistics 2

Distribution of instrument (Z)

FTO	MC4R	Ζ	Freq
0	0	0	0.20
0	1	1	0.15
1	0	1	0.27
1	1	2	0.21
2	0	2	0.09
2	1	3	0.07

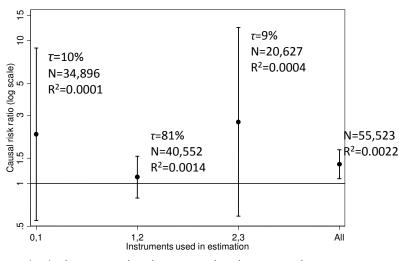


Copenhagen example Multiplicative SMM estimates



MSMM: Hansen over-identification test P = 0.31 E[Y(0)] = 0.58 (0.50, 0.65)

Copenhagen example local risk ratios



Check: $(0.10 \times 2.21) + (0.81 \times 1.11) + (0.09 \times 2.69) = 1.36$

3. Nonparametric bounds I

X,Y,Z all binary: IV assumptions—instrumental inequality (Bonet, 2001)

Denote
$$p_{yx.z} = P(Y = y, X = x \mid Z = z)$$

$$p_{00.0} + p_{10.1} \leq 1$$

$$p_{10.0} + p_{00.1} \leq 1$$

$$p_{11.0} + p_{01.1} \leq 1$$

$$p_{01.0} + p_{11.1} \leq 1$$

- Not a statistical test
- ▶ If fail then the IV assumptions must be violated
- ► But IV assumptions can be violated without failing the inequalities

Nonparametric bounds II

Bounds on intervention probabilities (Balke & Pearl, 1997):

$$P(Y(X = 0)), P(Y(X = 1))$$

Bounds on the Average Causal Effect (ACE):

$$ACE = P(Y(X = 1)) - P(Y(X = 0))$$

Nonparametric bounds II

Bounds on intervention probabilities (Balke & Pearl, 1997):

$$P(Y(X = 0)), P(Y(X = 1))$$

Bounds on the Average Causal Effect (ACE):

ACE =
$$P(Y(X = 1)) - P(Y(X = 0))$$

$$\max \left\{ \begin{cases} p_{00.0} + p_{11.1} - 1 \\ p_{00.1} + p_{11.1} - 1 \\ p_{11.0} + p_{00.1} - 1 \\ p_{00.0} + p_{11.0} - 1 \\ p_{00.0} + p_{11.0} - 1 \\ 2p_{00.0} + p_{11.0} + p_{10.0} + p_{11.1} - 2 \\ p_{00.0} + 2p_{11.0} + p_{00.1} + p_{01.1} - 2 \\ p_{10.0} + p_{11.0} + 2p_{00.1} + p_{11.1} - 2 \\ p_{00.0} + p_{01.0} + p_{00.1} + p_{01.1} - 2 \\ p_{00.0} + p_{01.0} + p_{00.1} + p_{01.1} - 2 \\ p_{00.0} + p_{01.0} - p_{00.1} + p_{01.1} - 2 \\ p_{00.0} + p_{01.0} - p_{00.1} + 2p_{11.1} - 2 \\ \end{cases} \right\} \leq \mathsf{ACE} \leq \min \left\{ \begin{cases} 1 - p_{10.0} - p_{01.1} \\ 1 - p_{01.0} - p_{10.0} \\ 1 - p_{01.0} - p_{10.1} \\ 2 - p_{01.0} - p_{10.0} - p_{10.1} - p_{11.1} \\ 2 - p_{10.0} - p_{11.0} - 2p_{01.1} - p_{10.1} \\ 2 - p_{00.0} - p_{01.0} - p_{01.1} - 2p_{10.1} \\ 2 - p_{00.0} - p_{01.0} - p_{01.1} - 2p_{10.1} \\ \end{cases} \right\}$$

Nonparametric bounds: Example

	Z = 0		Z = 1	
	Y = 0	Y = 1	Y = 0	Y = 1
X = 0	74	11514	34	2385
X = 1	0	0	12	9663

Table: Vitamin A supplementation data (Balke & Pearl, 1997).

Nonparametric bounds: Example

	Z = 0		Z = 1	
	Y = 0	Y = 1	Y = 0	Y = 1
X = 0	74	11514	34	2385
X = 1	0	0	12	9663

Table: Vitamin A supplementation data (Balke & Pearl, 1997).

	Bounds
Instrumental inequality	satisfied
P(Y(X=0))	(0.9936, 0.9936)
P(Y(X=1))	(0.7990, 0.9990)
Average Causal Effect (risk difference)	(-0.1946, 0.0054)
Causal risk ratio	(0.8042, 1.0054)

TSLS: ACE 0.0032 (95%CI 0.0010, 0.0055)

Nonparametric bounds: Interpretation

- Not same as a confidence interval
- ▶ Bounds of [-0.1946, 0.0054]: there exists some distribution involving *U* that yields a true ACE as small as -19.46%, and another that gives a true ACE as large as 0.54%, with both distributions satisfying the IV assumptions and having the same observed marginal frequencies on (*X*, *Y*, *Z*)

Nonparametric bounds: Extensions

- Monotonicity assumption: for all values u of U $P(X = 1 \mid Z = 1, U = u) \ge P(X = 1 \mid Z = 0, U = u)$ gives (slightly) tighter inequalities and bounds
- Extensions for a 3rd instrument category
- ▶ Data structures: (X, Z) & (Y, Z) in different samples

Nonparametric bounds: Limitation

Simulate data (N=10,000), two outcomes; both do not fulfil IV assumptions

$$Z \sim Bern(0.5), \quad U \sim Bern(0.5)$$

 $p_X = 0.05 + 0.1Z + 0.1U, \quad X \sim Bern(p_X)$
 $p_1 = 0.1 + 0.2Z + 0.05X + 0.1U, \quad Y_1 \sim Bern(p_1)$
 $p_2 = 0.1 + 0.05Z + 0.05X + 0.1U, \quad Y_2 \sim Bern(p_2)$

For Y_1 IV instrumental inequality is not satisfied

Nonparametric bounds: Limitation

Simulate data (N=10,000), two outcomes; both do not fulfil IV assumptions

$$Z \sim Bern(0.5), \quad U \sim Bern(0.5)$$

 $p_X = 0.05 + 0.1Z + 0.1U, \quad X \sim Bern(p_X)$
 $p_1 = 0.1 + 0.2Z + 0.05X + 0.1U, \quad Y_1 \sim Bern(p_1)$
 $p_2 = 0.1 + 0.05Z + 0.05X + 0.1U, \quad Y_2 \sim Bern(p_2)$

For Y_1 IV instrumental inequality is not satisfied For Y_2 we get

	Bounds
Instrumental inequality	satisfied
P(Y(X=0))	(0.1542, 0.2352)
P(Y(X=1))	(0.0585, 0.8464)
Average Causal Effect (risk difference)	(-0.1767, 0.6922)

4. Summary

- Instrumental variable assumptions
 - ► Not fully testable from observational data (as for all causal inf.)
- Applications
 - Mendelian randomization (IV: genotypes)
 - ► Correct for noncompliance (IV: randomized treatment)
- ► Test for presence of effect
- Estimators: ratio and two-stage least squares
- ▶ Structural Mean Models: G-estimation $Y(0) \perp \!\!\! \perp Z$
- Additive SMM
- Multiplicative SMM
- ► Review of methods Palmer, Sterne, et al., 2011; Palmer, Lawlor, et al., 2011
- ► Nonparametric bounds
- Stata command: bpbounds Palmer, Ramsahai, Didelez, & Sheehan,
 2011

Acknowledgements

- ► MRC Collaborative grant G0601625
- ► MRC CAiTE Centre grant G0600705
- ► ESRC grant RES-060-23-0011
- With thanks to Frank Windmeijer, Paul Clarke, Nuala Sheehan, Vanessa Didelez, Debbie Lawlor, Jonathan Sterne, Stijn Vansteelandt, Roland Ramsahai, George Davey Smith, Sha Meng, Neil Davies, Roger Harbord, Nic Timpson, Borge Nordestgaard.

References I

- Balke, A., & Pearl, J. (1997). Bounds on treatment effects from studies with imperfect compliance. *Journal of the American Statistical Association*, 92(439), 1172–1176.
- Bonet, B. (2001). Instrumentality tests revisited. In J. Breese & D. Koller (Eds.), Proc. 17th conf. on uncertainty in artificial intelligence (pp. 48–55). Seattle, WA: Morgan Kaufmann.
- Chaussé, P. (2010). Computing generalized method of moments and generalized empirical likelihood with R. *Journal of Statistical Software*, *34*(11), 1–35. Available from http://www.jstatsoft.org/v34/i11/
- Davey Smith, G., & Ebrahim, S. (2003). 'Mendelian randomization': can genetic epidemiology contribute to understanding environmental determinants of disease. *International Journal of Epidemiology*, 32, 1–22.
- Hernán, M. A., & Robins, J. M. (2006). Instruments for Causal Inference. An Epidemiologist's Dream? *Epidemiology*, 17, 360–372.
- Imbens, G. W., & Angrist, J. D. (1994). Identification and Estimation of Local Average Treatment Effects. *Econometrica*, 62, 467–467.
- Palmer, T. M., Lawlor, D. A., Harbord, R. M., Sheehan, N. A., Tobias, J. H., Timpson, N. J., et al. (2011). Using multiple genetic variants as instrumental variables for modifiable risk factors. Statistical Methods in Medical Research. (Published online before print January 7)

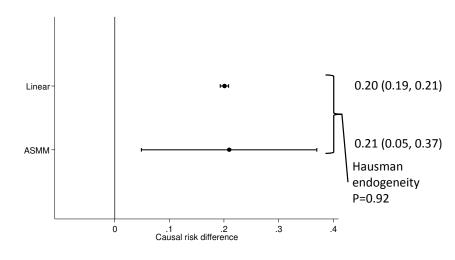
References II

- Palmer, T. M., Ramsahai, R. R., Didelez, V., & Sheehan, N. A. (2011). Nonparametric bounds for the causal effect in a binary instrumental-variable model. Stata Journal, 11(3), 345-367(23).
- Palmer, T. M., Sterne, J. A. C., Harbord, R. M., Lawlor, D. A., Sheehan, N. A., Meng, S., et al. (2011). Instrumental variable estimation of causal risk ratios and causal odds ratios in mendelian randomization analyses. *American Journal* of Epidemiology, 173, 1392–1403.
- Robins, J. M. (1989). Health services research methodology: A focus on aids. In L. Sechrest, H. Freeman, & A. Mulley (Eds.), (chap. The analysis of randomized and non-randomized AIDS treatment trials using a new approach to causal inference in longitudinal studies). Washington DC, US: US Public Health Service.
- Robins, J. M., Rotnitzky, A., & Scharfstein, D. O. (1999). Statistical models in epidemiology: The environment and clinical trials. In M. E. Halloran & D. Berry (Eds.), (pp. 1–92). New York, US: Springer.
- Ten Have, T. R., Elliott, M. R., Joffe, M., Zanutto, E., & Datto, C. (2004). Causal models for randomized physician encouragement trials in treating primary care depression. *Journal of the American Statistical Association*, 99(465), 16-25.

Copenhagen example comparison estimates

	RR (95% CI)	P over-id
MSMM	1.36 (1.08, 1.72)	0.31
$Y - \exp(\psi X) \perp \!\!\! \perp Z$	1.36 (1.07, 1.75)	0.30
Control function	1.36 (1.08, 1.71)	
	OR (95% CI)	P over-id
LSMM two-stage	1.88 (1.75, 2.02)	
LSMM joint	2.87 (1.25, 6.55)	0.29
$Y - expit(\psi X) \perp \!\!\! \perp Z$	2.69 (1.23, 5.90)	0.30
Control function	2.69 (1.21, 5.97)	

Copenhagen example Additive SMM estimates



MSMM: Hansen over-identification test P = 0.30 E[Y(0)] = 0.58 (0.48, 0.67)

(double) Logistic SMM

$$\mathsf{logit}(p) = \mathsf{log}(p/(1-p)), \mathsf{expit}(x) = e^x/(1+e^x)$$

Goetghebeur, 2010

$$\begin{split} \mathsf{logit}(E[Y|X,Z]) - \mathsf{logit}(E[Y(0)|X,Z]) &= \psi X \\ \psi : \mathsf{log} \ \mathsf{causal} \ \mathsf{odds} \ \mathsf{ratio} \\ \mathsf{Rearrange} \ \mathsf{for} \ Y(0) &= \mathsf{expit}(\mathsf{logit}(Y) - \psi X) \end{split}$$

(double) Logistic SMM

$$\mathsf{logit}(p) = \mathsf{log}(p/(1-p)), \mathsf{expit}(x) = e^x/(1+e^x)$$

Goetghebeur, 2010

$$\begin{split} \mathsf{logit}(E[Y|X,Z]) - \mathsf{logit}(E[Y(0)|X,Z]) &= \psi X \\ \psi : \mathsf{log} \ \mathsf{causal} \ \mathsf{odds} \ \mathsf{ratio} \\ \mathsf{Rearrange} \ \mathsf{for} \ Y(0) &= \mathsf{expit}(\mathsf{logit}(Y) - \psi X) \end{split}$$

- ► Can't be estimated in a single step Robins (1999)
- ► First stage association model Vansteelandt (2003):
 - (i) logistic regression of Y on X & Z & interactions
 - (ii) predict Y, estimate LSMM using predicted Y

(double) Logistic SMM moment conditions

Association model moment conditions

Logistic regression using GMM

$$\begin{split} &E[(Y - \text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z))1] = 0 \\ &E[(Y - \text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z))X] = 0 \\ &E[(Y - \text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z))Z] = 0 \\ &E[(Y - \text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z))XZ] = 0 \end{split}$$

(double) Logistic SMM moment conditions

Association model moment conditions

Logistic regression using GMM

$$\begin{split} &E[(Y - \text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z))1] = 0 \\ &E[(Y - \text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z))X] = 0 \\ &E[(Y - \text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z))Z] = 0 \\ &E[(Y - \text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z))XZ] = 0 \end{split}$$

Causal model moment conditions

$$E[(\operatorname{expit}(\operatorname{logit}(\widehat{p}) - \psi X) - Y(0))1] = 0$$

$$E[(\operatorname{expit}(\operatorname{logit}(\widehat{p}) - \psi X) - Y(0))Z] = 0$$

Problem: SEs incorrect - need association model uncertainty

LSMM joint estimation

Joint estimation = correct SEs Gourieroux (1996)

Vansteelandt & Goetghebeur (2003)

$$\begin{split} E[(Y - \text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z))1] &= 0 \\ E[(Y - \text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z))X] &= 0 \\ E[(Y - \text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z))Z] &= 0 \\ E[(Y - \text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z))XZ] &= 0 \\ E[(\text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z - \psi X) - Y(0))1] &= 0 \\ E[(\text{expit}(\beta_0 + \beta_1 X + \beta_2 Z + \beta_3 X Z - \psi X) - Y(0))Z] &= 0 \end{split}$$

Stata gmm command - allows multiple equations - still 1 line of code

Example: causal model SEs $\times 10$

Including covariates

TSLS: include covariates in both stages

GMM: use covariates as instruments for themselves

Including (pre-exposure) covariates in MSMM

$$Y(0) \perp \!\!\!\perp Z|C$$

 $\log(E[Y|X,Z,C]) - \log(E[Y(0)|X,Z,C]) = \psi X + \psi_c C$

Including covariates

TSLS: include covariates in both stages

GMM: use covariates as instruments for themselves

Including (pre-exposure) covariates in MSMM

$$Y(0) \perp |Z|C$$

$$\log(E[Y|X,Z,C]) - \log(E[Y(0)|X,Z,C]) = \psi X + \psi_c C$$

Copenhagen example estimates

Covariates	RR (95%CI)	Over-id P
	1.36 (1.08, 1.72)	0.31
sex	1.36 (1.07, 1.72)	0.39
sex, age	1.35 (1.07, 1.71)	0.58
sex, age, chol	1.33 (1.05, 1.68)	0.49