

Conceptual and methodological problems with bias
detection and avoidance in natural language
processing

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Chapter 1

Introduction

Placeholder

Chapter 2

Cosine similarity and bias detection

Placeholder

2.1 Word embeddings

2.2 Cosine similarity and distance

2.3 Cosine distance in a one-class bias detection

2.4 Cosine distance in a multi-class bias detection

2.5 Limitations of the approach

Chapter 3

Walkthrough with the religion dataset

3.1 Loading and understanding the dataset

We start with loading the libraries needed for the analysis.

```
library(ggplot2)
library(ggthemes)
library(rethinking)
library(tidyverse)
library(ggpubr)
library(kableExtra)
library(dplyr)
library(ggExtra)
library(cowplot)
```

We will use the choice of protected words and stereotypical predicates used in REF. This is a decent point of departure, not only we want to compare our method to that of REF, but also because this data format is fairly general (as contrasted, say, with a set

up for binary stereotypes). Note also that the method we develop here can fairly easily be run for different stereotypization patterns. Let's start with explaining the method and its deployment using a dataset obtained for the religion-related protected words.

Let's load, clean a bit and inspect the head of the religion dataset we prepared. In order to obtain this dataset, we calculated the cosine distance between each protected word and each word from both the bias-related attribute groups, which were used in the original study, and to neutral and human control attributes which we added as control groups. For instance, for religion, the bias-related predicates (coming from the original study in REF) include muslim bias attributes, jew bias attributes, christian bias attributes (see a list in the APPENDIX).

We decided to add control groups in the form of two classes — neutral words and human-related words. Without a proper control group it is quite hard to compare the resulting cosine distances and decide on their significance in bias detection. We prepared approximately 230 neutral words to double-check the prima-facie neutral hypothesis that their cosine similarity to the protected words will oscillate around 0 (that is, the distances will be around 1). This provides us with a more reliable point of reference. Moreover, we added human attributes that are associated with people in general to investigate whether the smaller cosine distance between protected words and stereotypes can result simply from the fact that the stereotype predicates are associated with humans. For two control groups, we have randomly drawn 230 words that do not express any property usually attributed to humans, and human related attributes.

```
religion <- read.csv("../datasets/religionReddit.csv")[-1]
colnames(religion) <- c("protectedWord", "wordToCompare", "wordClass",
                        "cosineDistance", "cosineSimilarity", "connection")
#head(religion)
#library(plyr)
religion$wordClass <- as.factor(religion$wordClass)
```

Table 3.1: Head of the religion dataset.

protectedWord	wordToCompare	wordClass	cosineDistance	cosineSimilarity	connection
judaism	violent	muslim	0.7141939	0.2858061	different
judaism	terrorist	muslim	0.7461333	0.2538667	different
judaism	dirty	muslim	1.2002599	-0.2002599	different
judaism	uneducated	muslim	0.7885469	0.2114531	different
judaism	greedy	jewish	1.0026172	-0.0026172	associated
judaism	cheap	jewish	1.2323229	-0.2323229	associated

```
#levels(religion$wordClass)
#religion$wordClass <- mapvalues(religion$wordClass,
#                                from = levels(religion$wordClass),
#                                to = c("christian", "human", "jewish", "muslim", "neutral"))
levels(religion$wordClass) <- c("christian", "human", "jewish", "muslim", "neutral")
head(religion) %>% kable(format = "latex", booktabs=T,
                        linesep = "", escape = FALSE,
                        caption = "Head of the religion dataset.") %>% kable_styling(latex_opti
```

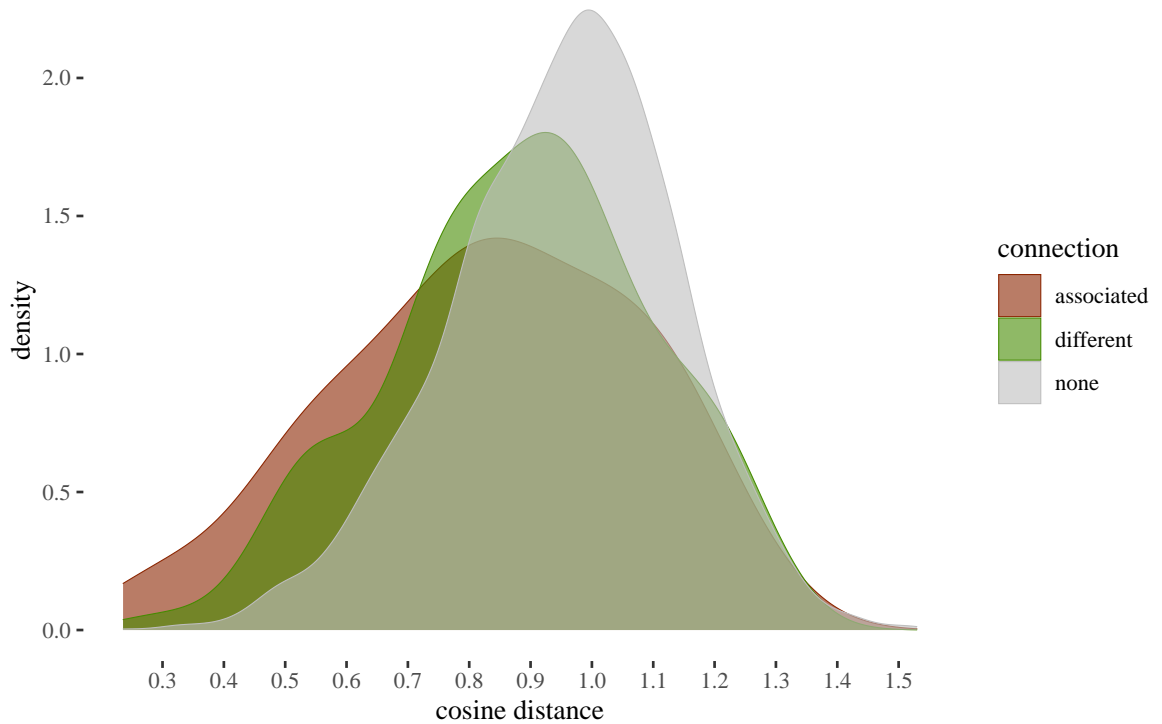
The `protectedWord` column contains words from a protected class that (in a perfect world according to the assumptions of the original study) should not be associated with harmful stereotypes. `wordToCompare` contains attributes, including stereotypes and control group words. For each row we compute the cosine distances between a given protected word and a given attribute word. `wordClass` tells us which class an attribute is supposed to be stereotypically associated with, that is, whether the word from `wordToCompare` is associated stereotypically with jews, christians or muslims, or whether it belongs to a control group. `cosineDistance` is simply a calculation of the cosine distance between protected word and attribute. `cosineSimilarity` contains the result of subtracting cosine distance from 1. `connection` contains information about the relation type between a protected word and an attribute. If the attribute is e.g. a harmful jewish stereotype and the protected word is also from the judaism group, the

connection has value **associated**. If the attribute is still stereotypically jewish, but the protected word comes from another religion, the connection is labelled as **different**. If the attribute belongs to a neutral group then the connection is labelled as **none** and if an attribute belongs to the **human** class, then the connection is labelled as **human**.

3.2 First look at the empirical distributions

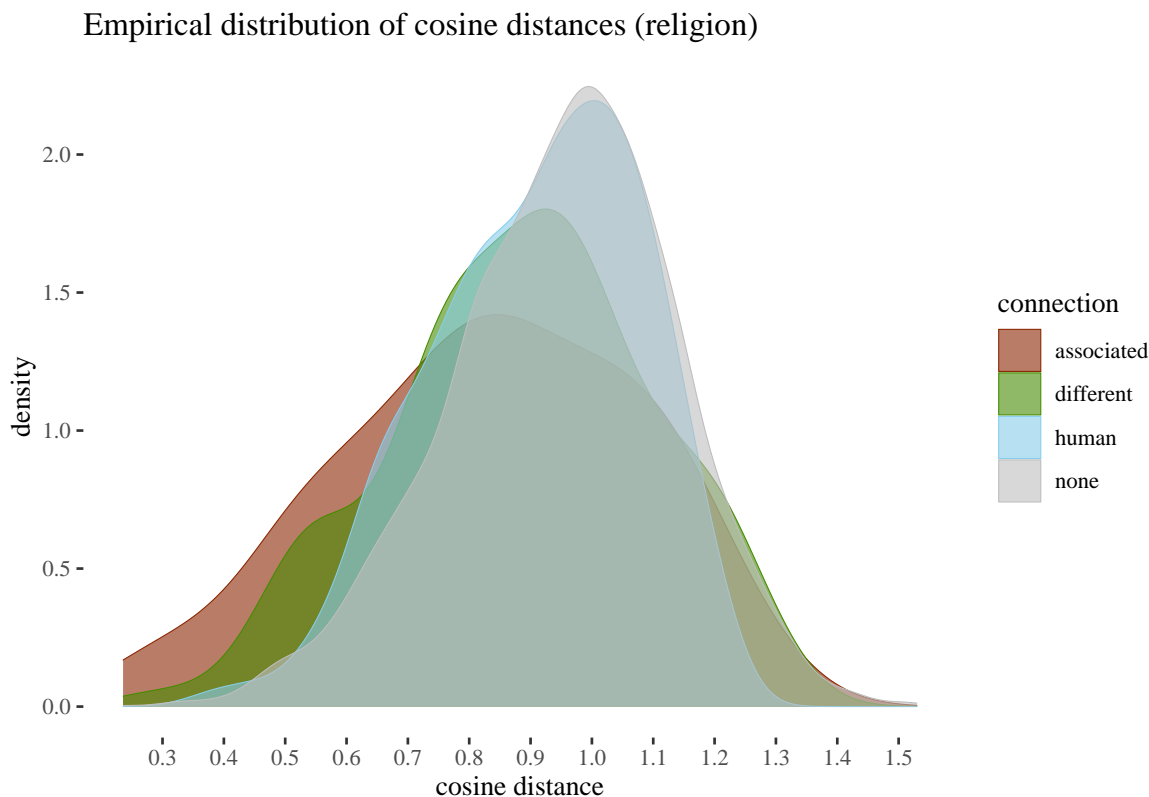
First let's take a look at the empirical distribution of distances by the connection type, initially ignoring the human control class for now.

Empirical distribution of cosine distances (religion), no human attributes.



The first impression is that while there is a shift for associated words towards smaller cosine distances as compared to the neutral words, slightly surprisingly a slightly weaker shift in the same direction is visible for attributes associated with different stereotypes.

Moreover, the empirical distributions overlap to a large extent and the means grouped by connection type do not seem too far from each other. In fact, as there is a lot of variety in the cosine distances (as we will soon see), we need to gauge the uncertainty involved, and to look more carefully at individual protected words to get a better idea of how the cosine distance distribution changes for different attribute groups and different protected classes. Now, let's add the human attributes to the picture:



Notice that the distribution for **human** (even though we did our best not to include in it any stereotype-related attributes) is left-skewed, with much overlap with **associated** and **different**, which illustrates the need to take being associated with humans as an important predictor.

Our focus lies in **connection** as a predictor. Moreover, later on we'll be interested in looking at the protected words separately, and at protected words split by connection.

For technical reasons it is useful to represent these factors as integer vectors.

```
religion$con <- as.integer(religion$connection)

## Warning: NAs introduced by coercion

religion$pw <- as.integer(religion$protectedWord)

## Warning: NAs introduced by coercion

religion$pwFactor <- factor(paste0(religion$protectedWord, religion$connection))
religion$pwIndex <- as.integer(religion$pwFactor)
```

A short script, `cleanDataset` to make this faster, so equivalently:

```
source("../functions/cleanDataset.R")
religion <- read.csv("../datasets/religionReddit.csv")[-1]
religion <- cleanDataset(religion, c("christian", "human", "jewish", "muslim", "neutral"))

## Warning in cleanDataset(religion, c("christian", "human", "jewish", "muslim", :
## NAs introduced by coercion

## Warning in cleanDataset(religion, c("christian", "human", "jewish", "muslim", :
## NAs introduced by coercion
```

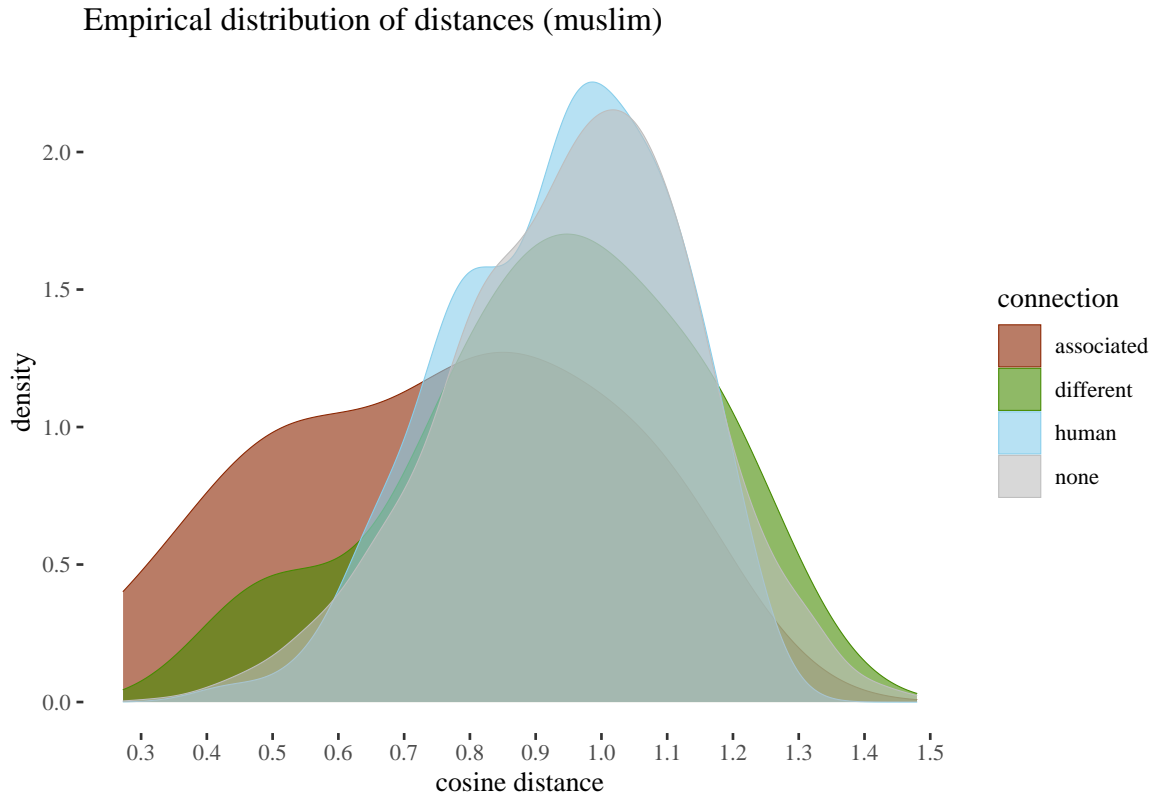
3.3 Looking at the islam-related words

For now, let's focus on five protected words related to islam ("imam," "islam," "mosque," "muslim," and "quran"). The word list associates with islam four stereotypical attributes ("violent," "terrorist," "uneducated" and "dirty"). First, we select and plot the empirical distributions for these protected words.

```
library(tidyverse)

muslimWords <- c("imam", "islam", "mosque", "muslim", "quran")
muslim <- religion %>% filter(protectedWord %in% muslimWords)
ggplot(muslim, aes(x = cosineDistance, fill = connection, color = connection))+
```

```
geom_density(alpha=0.6,size = .2)+
scale_fill_manual(values = c("orangered4","chartreuse4", "skyblue", "gray"))+
scale_x_continuous(breaks = seq(0.3,1.5, by = 0.1))+xlab("cosine distance")+
scale_color_manual(values = c("orangered4","chartreuse4","skyblue","gray"))+
theme_tufte()+ggtitle("Empirical distribution of distances (muslim)")
```



Once we focus on words related to islam, the associated bias seems to be stronger than in the whole dataset. This is a step towards illustrating that the distribution of bias is uneven.

Now, say we want to look at a single protected word. Since the dataset also contains comparison multiple control neutral and human attributes, we randomly select only 5 from none and 5 from human control groups of those for the visualisation purposes.

```
library(tidyverse)
```

```
## -- Attaching packages ----- tidyverse 1.3.1 --
```

```
## v ggplot2 3.3.3      v purrr  0.3.4
## v tibble  3.1.2      v dplyr  1.0.6
## v tidyr   1.1.3      v stringr 1.4.0
## v readr   1.4.0      v forcats 0.5.1
```

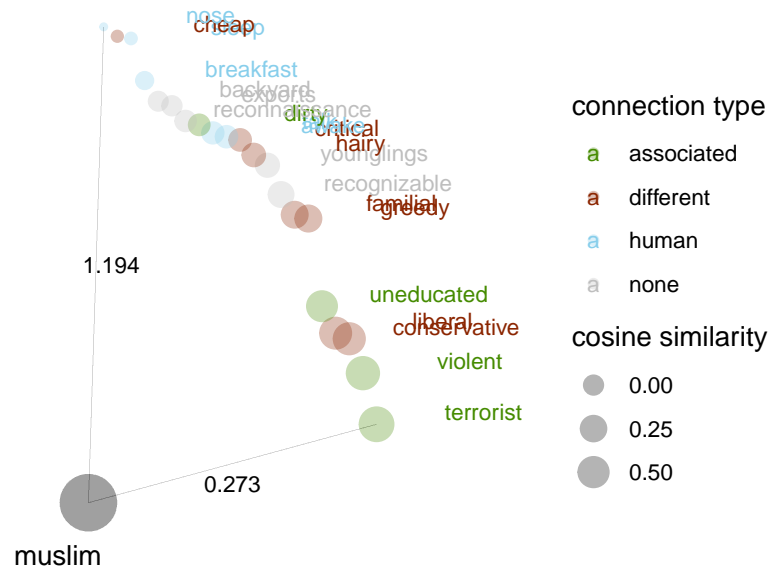
```
## -- Conflicts ----- tidyverse_conflicts() --
```

```
## x dplyr::filter() masks stats::filter()
## x dplyr::lag()     masks stats::lag()
```

```
muslimClass <- muslim %>% filter(protectedWord == "muslim")
neutralSample <- sample_n(filter(muslimClass,connection == "none"), 5)
humanSample <- sample_n(filter(muslimClass,connection == "human"), 5)
muslimVis <- muslimClass %>% filter(connection != "none" & connection != "human")
muslimVis <- rbind(muslimVis,neutralSample,humanSample)
```

```
## Warning in `[<-factor`(`*tmp*`, ri, value = structure(c("muslim_stereotype", :
## invalid factor level, NA generated
```

```
#we plug in our visualisation script
source("../functions/visualisationTools.R")
#two arguments: dataset and protected word
visualiseProtected(muslimVis,"muslim")
```

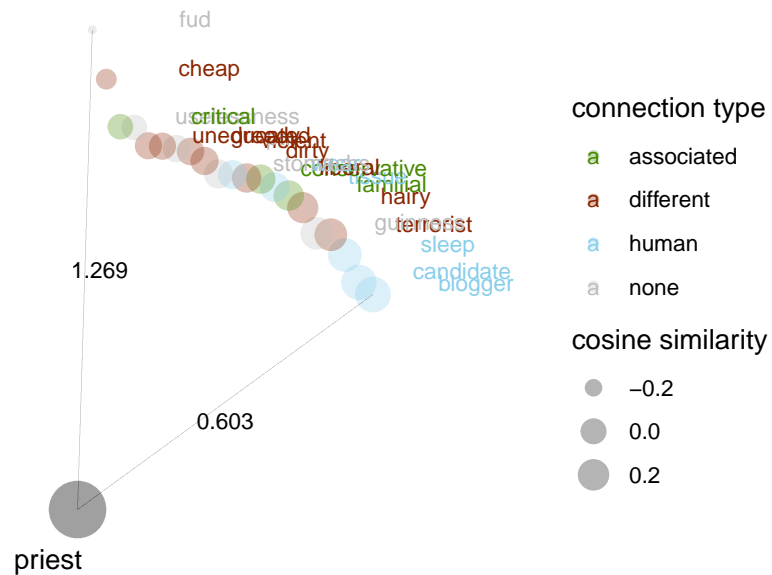
Note that the distance between the grey point and the other points is proportional to cosine distance, the non-grey point size is proportional to cosine similarity to the protected word, and color groups by the connection type. So for *muslim* it seems that the stereotypes coming from the word list are fairly well visible. To give you some taste of how uneven the dataset is, compare this to what happens with *priest*.

```
library(tidyverse)
priestClass <- religion %>% filter(protectedWord == "priest")
neutralSample <- sample_n(filter(priestClass,connection == "none"), 5)
humanSample <- sample_n(filter(priestClass,connection == "human"), 5)
priestVis <- priestClass %>% filter(connection != "none" & connection != "human")
priestVis <- rbind(priestVis,neutralSample,humanSample)
```

```
## Warning in `[<-.factor`(`*tmp*`, ri, value = structure(c("muslim_stereotype", :
## invalid factor level, NA generated
```

```
#we plug in our visualisation script
source("../functions/visualisationTools.R")

#two arguments: dataset and protected word
visualiseProtected(priestVis,"priest")
```



Here you can see that some human attributes are closer than stereotype attributes, and that there is no clear reason to claim that **associated** attributes are closer than **different** or **human** attributes. This, again, illustrates the need of case-by-case analysis with control groups.

The general idea now is that the word lists provided in different pieces of research are just samples of attributes associates with various stereotypes and should be treated as such: the uncertainty involved and the sample sizes should have clear impact on our estimates.

3.4 Bayesian model structure and assumptions

We will now think of cosine distance as the output variable, and will build a few Bayesian models to compare. First, we just build a baseline model which estimates cosine distance to the attributes separately for each protected word. The underlying idea is that different protected words might in general have different relations to all the attributes and these relations should be our point of departure.

Here is the intuition behind the mathematical Bayesian model involved. Our outcome variable is `cosine difference`, which we take to be normally distributed around the predicted mean for a given protected word (that is, we assume the residuals are normally distributed). The simplest model specification is:

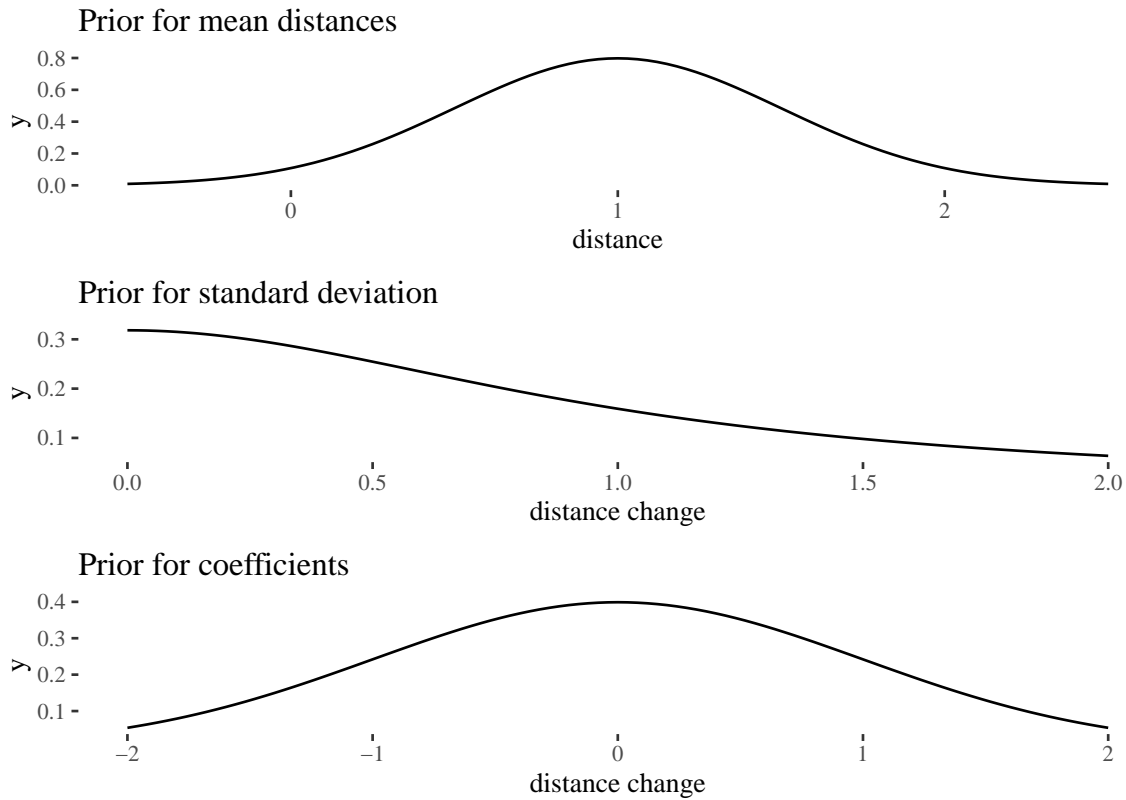
$$\text{cosineDistance}_i \sim \text{dnorm}(\mu_i, \sigma) \quad (3.1)$$

$$\mu_i = m_{pw} \quad (3.2)$$

$$m_{pw} \sim \text{dnorm}(1, .5) \quad (3.3)$$

$$\sigma \sim \text{dcauchy}(0, 1) \quad (3.4)$$

That is, we assume the estimated means might be different for different protected words and our prior for the mean and the overall standard deviation are normal with mean 1 and sd=.5 and half-cauchy with parameters 0,1. Further on we'll also estimate additional impact the connection type may have. For this impact we take a slightly skeptical prior centered around 0 distributed normally with sd = 1. These are fairly weak and slightly skeptical regularizing priors, which can be illustrated as follows:



3.5 Choosing predictors

Now we can define and compile the baseline model. Its parameters will have a posterior distribution obtained using either Hamiltonian Monte Carlo methods (STAN) available through the `rethinking` package.

```
library(rethinking)
options(buildtools.check = function(action) TRUE )
religionBaseline <- ulam(
  alist(
    cosineDistance ~ dnorm(mu,sigma),
    mu <- m[pw],
    m[pw] ~ dnorm(1,.5),
    sigma ~ dcauchy(0,1)
```

```

),
data = religion,
chains=2 , iter=4000 , warmup=1000,
start= list(mu = 1, co = 0, sigma= .3),
log_lik = TRUE, cores=4
)
#saving
#saveRDS(religionBaseline,
#file = "cosineAnalysis/models/religionBaseline.rds")

```

The only reason we need it is the evaluation of connection as a predictor. Does including it in o the model improve the situation? To investigate this, let's now build a model according to the following specification:

$$\text{cosineDistance}_i \sim \text{dnorm}(\mu_i, \sigma) \quad (3.5)$$

$$\mu_i = m_{pw} + co_{con} \quad (3.6)$$

$$m_{pw} \sim \text{dnorm}(1, .5) \quad (3.7)$$

$$co_{con} \sim \text{dnorm}(0, 1) \quad (3.8)$$

$$\sigma \sim \text{dcauchy}(0, 1) \quad (3.9)$$

The idea now is that each connection type comes with its own coefficient *co* that has impact on mean distances for protected words taken separately.

```

library(rethinking)
options(buildtools.check = function(action) TRUE )
religionCoefs <- ulam(
  alist(
    cosineDistance ~ dnorm(mu,sigma),
    mu <- m[pw] + co[con],
    m[pw] ~ dnorm(1,.5),
    co[con] ~dnorm(0,.5),

```

```

    sigma ~ dcauchy(0,1)
  ),
  data = religion,
  chains=2 , iter=8000 , warmup=1000,
  log_lik = TRUE
)

```

First, let's see if this model is really better in terms of the Widely Acceptable Information Criterion (WAIC):

```

##                WAIC SE dWAIC dSE pWAIC weight
## religionCoefs   -2328 93      0  NA    20      1
## religionBaseline -2283 95     45  17    16      0

```

Clearly, it should be given weight 1 as compared to the baseline model. So far, we've learned that the connection type actually has predictive value. Let's take a look at the coefficient estimates:

```

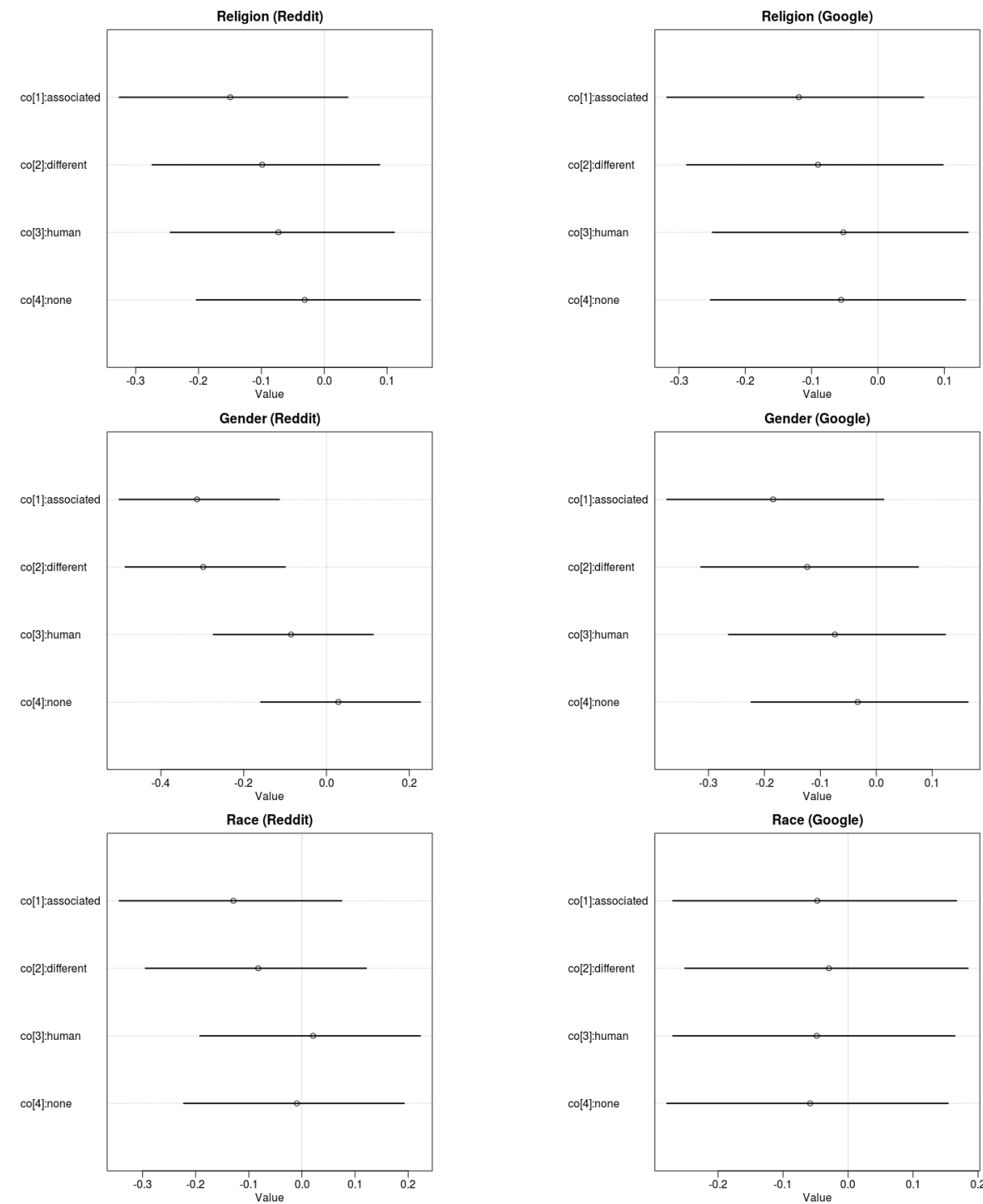
##           mean          sd      5.5%      94.5%    n_eff    Rhat4
## co[1] -0.14956420 0.1151675 -0.3261650 0.03741930 294.2033 1.001449
## co[2] -0.09880543 0.1145024 -0.2736985 0.08813271 291.5044 1.001564
## co[3] -0.07282752 0.1133894 -0.2447778 0.11158986 287.7820 1.001627
## co[4] -0.03103179 0.1131420 -0.2034442 0.15268770 286.8283 1.001606

```

3.6 Dataset-level coefficients

Let's plot them together with their highest posterior density intervals, for the three topic groups.

What should strike us is that while the mean estimates of the coefficients indeed do differ a bit, usually the highest posterior density intervals all include zero, and so we



do not have strong reasons to say that, say, as far as the whole religion dataset is involved, being associated indeed is connected with lower cosine distance. A second striking observation is that the estimated impact for associated stereotypes is quite often not too different from the estimated impact of attributes associated with different stereotypes, and both are sometimes not too far from the estimated impact for simply human attributes. In general, once the uncertainty involved is taken seriously by using control groups and statistical uncertainty estimation that does not dispose of pointwise data, the picture which focuses only on differences between means of means is too simplistic.

But this doesn't mean important differences for some protected words are not there. For one thing, if you start with a word list that is very uneven, the actually not so bad status of some of the protected words might mask a pretty bad situation in which some other protected words are. For comparison, let's see what a model focused on words related to islam tells us.

```
#this is how we build the model
religion <- read.csv("cosineAnalysis/datasets/religionReddit.csv")[-1]
colnames(religion) <- c("protectedWord", "wordToCompare", "wordClass",
                        "cosineDistance", "cosineSimilarity", "connection")
levels(religion$wordClass) <- c("christian", "human", "jewish", "muslim", "neutral")
muslimWords <- c("imam", "islam", "mosque", "muslim", "quran")
muslim <- religion %>% filter(protectedWord %in% muslimWords)
muslim$protectedWord <- droplevels(muslim$protectedWord)
muslim$pw <- as.integer(muslim$protectedWord)
muslim$con <- as.integer(muslim$connection)
muslim$pwFactor <- factor(paste0(muslim$protectedWord, muslim$con))
muslim$pwIndex <- as.integer(muslim$pwFactor)

islamCoefs <- ulam(
  alist(
```



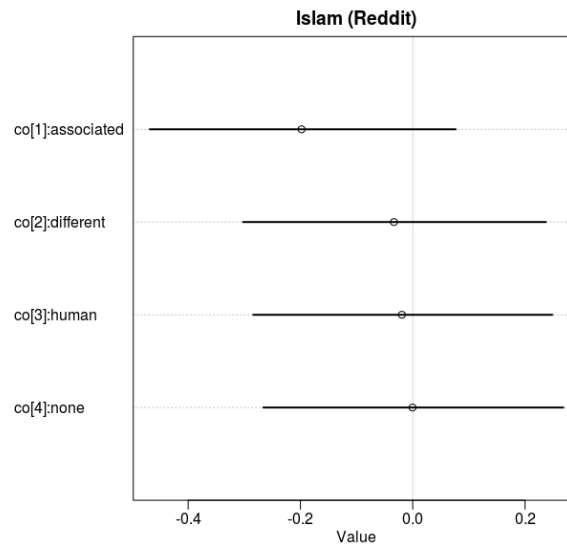
```

    cosineDistance ~ dnorm(mu,sigma),
    mu <- m[pw] + co[con],
    m[pw] ~ dnorm(1,.5),
    co[con] ~ dnorm(0,.5),
    sigma ~ dcauchy(0,1)
  ),
  data = muslim,
  chains=2 , iter=10000 , warmup=1000, cores = 4,
  log_lik = TRUE
)

```

Let's take a look at the coefficients:

##		mean	sd	5.5%	94.5%	n_eff	Rhat4
##	co[1]	-0.1979930035	0.1708785	-0.4682696	0.07634977	1789.894	1.003445
##	co[2]	-0.0334215769	0.1687587	-0.3021954	0.23720938	1738.720	1.003575
##	co[3]	-0.0192492753	0.1675754	-0.2840596	0.24860974	1732.907	1.003755
##	co[4]	-0.0003911363	0.1670610	-0.2661047	0.26815172	1723.758	1.003837



While muslim words were unusual in the sense that the disparity between associated attributes and others is stronger, the evidence is still not conclusive. This is because the

variation even within islam-related words is large enough (and sample sizes sufficiently small) for all the highest posterior density intervals to still include zeros.

So, it seems, taking a closer look does seem to make a difference. The question is, what happens if we do take a close look at the level of protected words?

Chapter 4

Protected-word level analysis

4.1 Model structure and assumptions

Let's turn then to data analysis that takes a look at protected words separately. This time for each combination of a protected word and a connection status we will have a separate mean cosine distance estimate, each coming with its own highest posterior density interval. This means we will use indices that are result from all such combinations (and then we will split them up in the model precis to build visualisation, feel free to look at the `visualiseStats.R` script for details).

```
options(buildtools.check = function(action) TRUE ) #removes install pop-up request
religion$pwFactor <- factor(paste0(religion$protectedWord, "-", religion$connection))
religion$pwIndex <- as.integer(religion$pwFactor)

religionSeparate <- ulam(
  alist(
    cosineDistance ~ dnorm(mu,sigma),
    mu <- c[pwIndex],
    c[pwIndex] ~ dnorm(1,.5),
```

```

    sigma ~ dcauchy(0,1)
  ),
  data = religion,
  chains=2 , iter=10000 , warmup=1000,
  start=list(no = 1, a = 0, d = 0, sigma= .3), log_lik = TRUE
)

```

Let's see if the individualized model does better than the previous models in light of WAIC which does add penalty for the number of parameters.

```

compareBaselineCoefsSeparate<- readRDS("../datasets/compareBaselineCoefsSeparate.rds")
compareBaselineCoefsSeparate

```

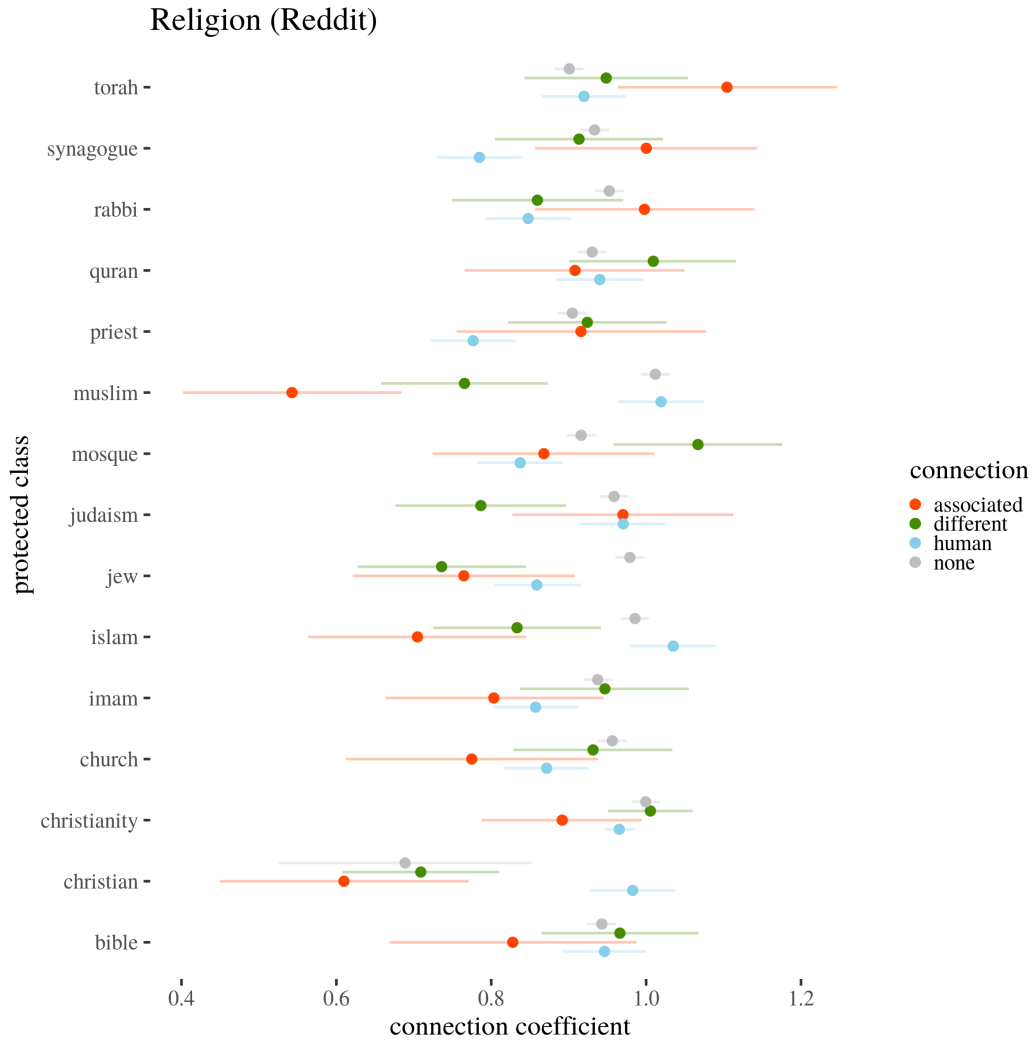
```

##                WAIC SE dWAIC dSE pWAIC weight
## religionSeparate -2400 93      0  NA    60      1
## religionCoefs    -2328 93     72  29    20      0
## religionBaseline -2283 95    117  37    16      0

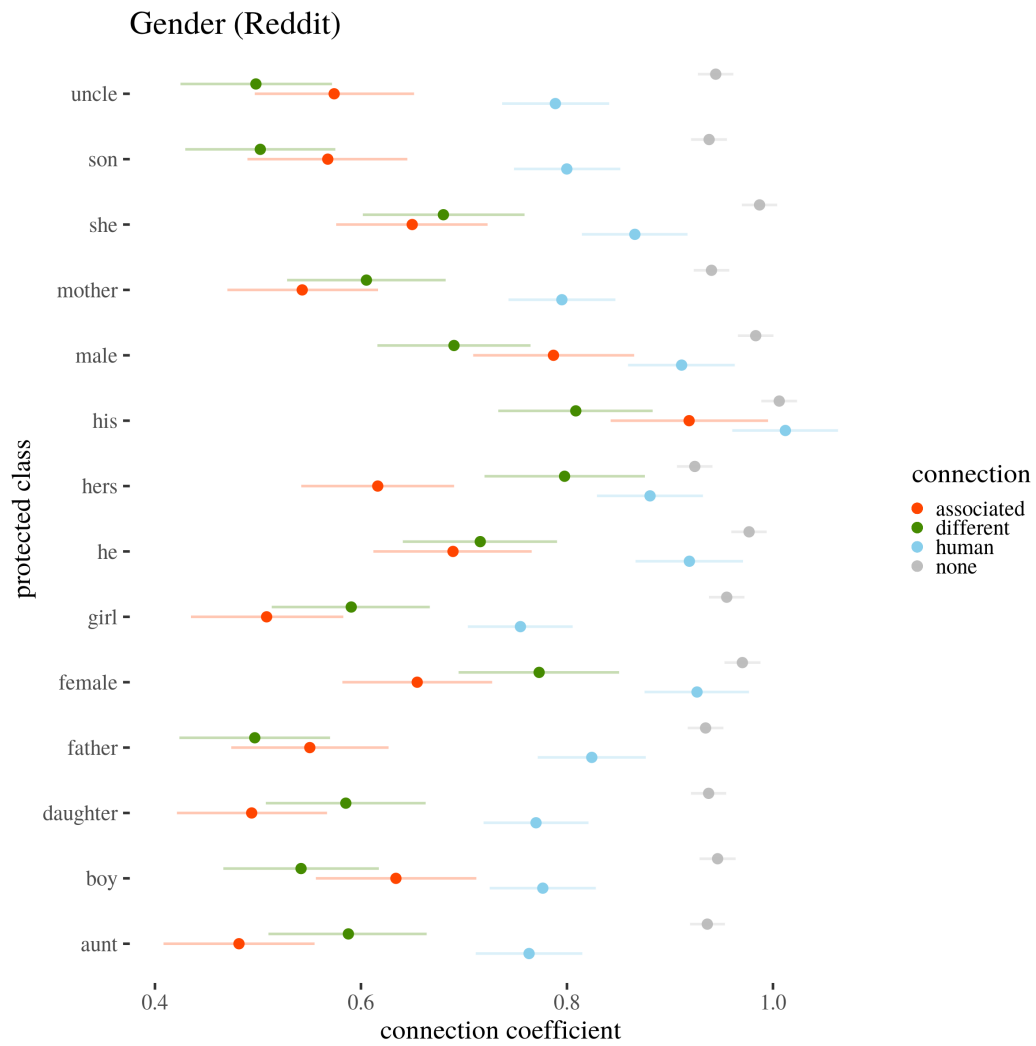
```

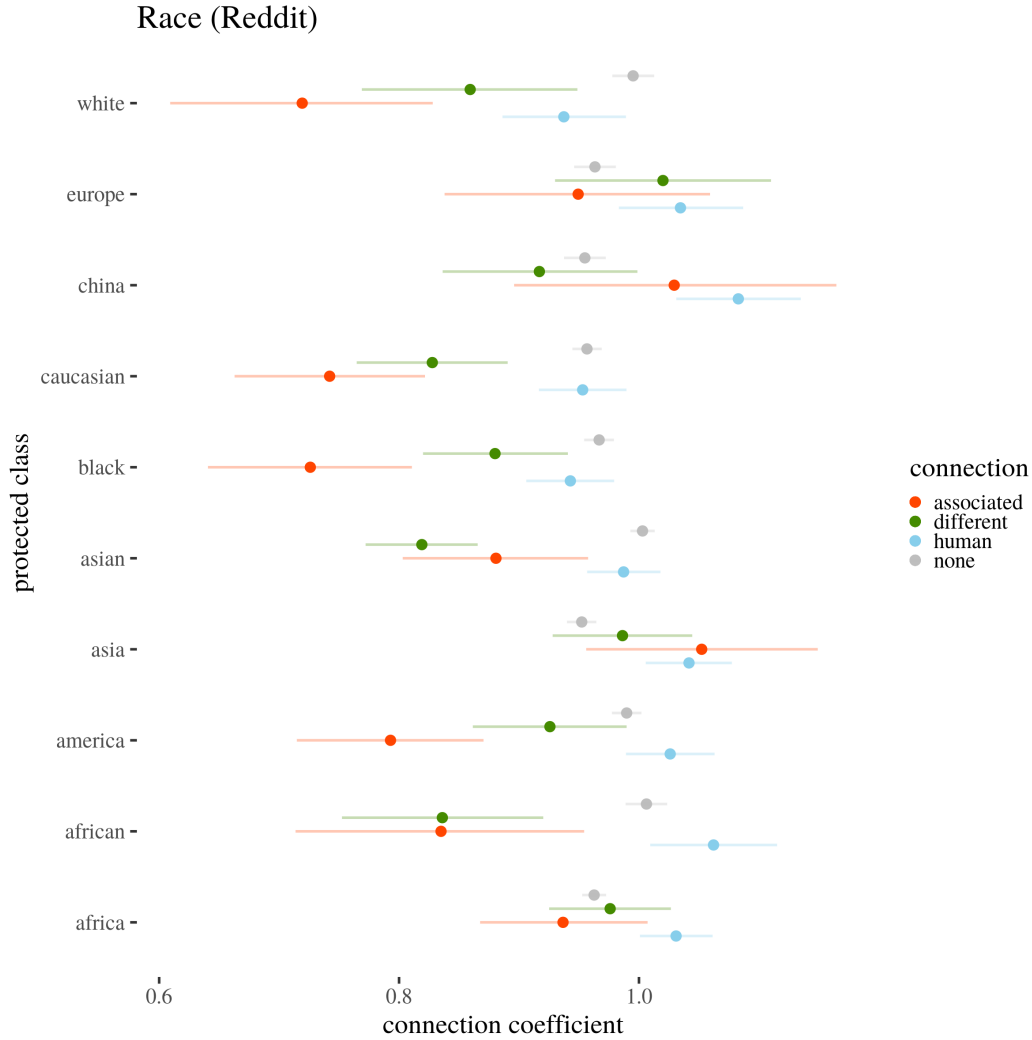
4.2 Protected classes in Reddit and Google embeddings

It seems that we do want to prefer this model, despite its relative complication. Now, what does it tell us about the protected words? Let's visualise the predicted means together with 89% highest posterior density intervals.



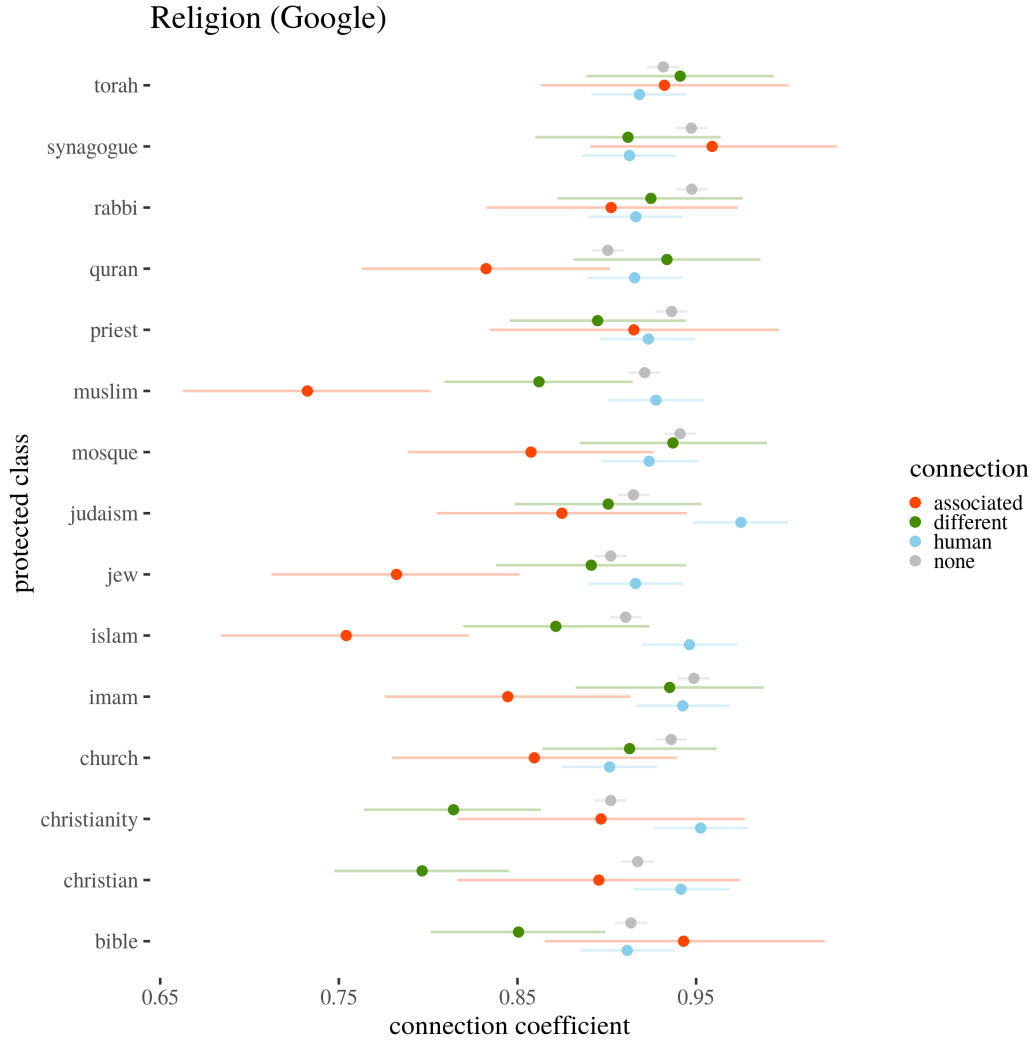
Before we move on, let's perform analogous analyses for the remaining types of supposed bias: gender and race (the model building is analogous).

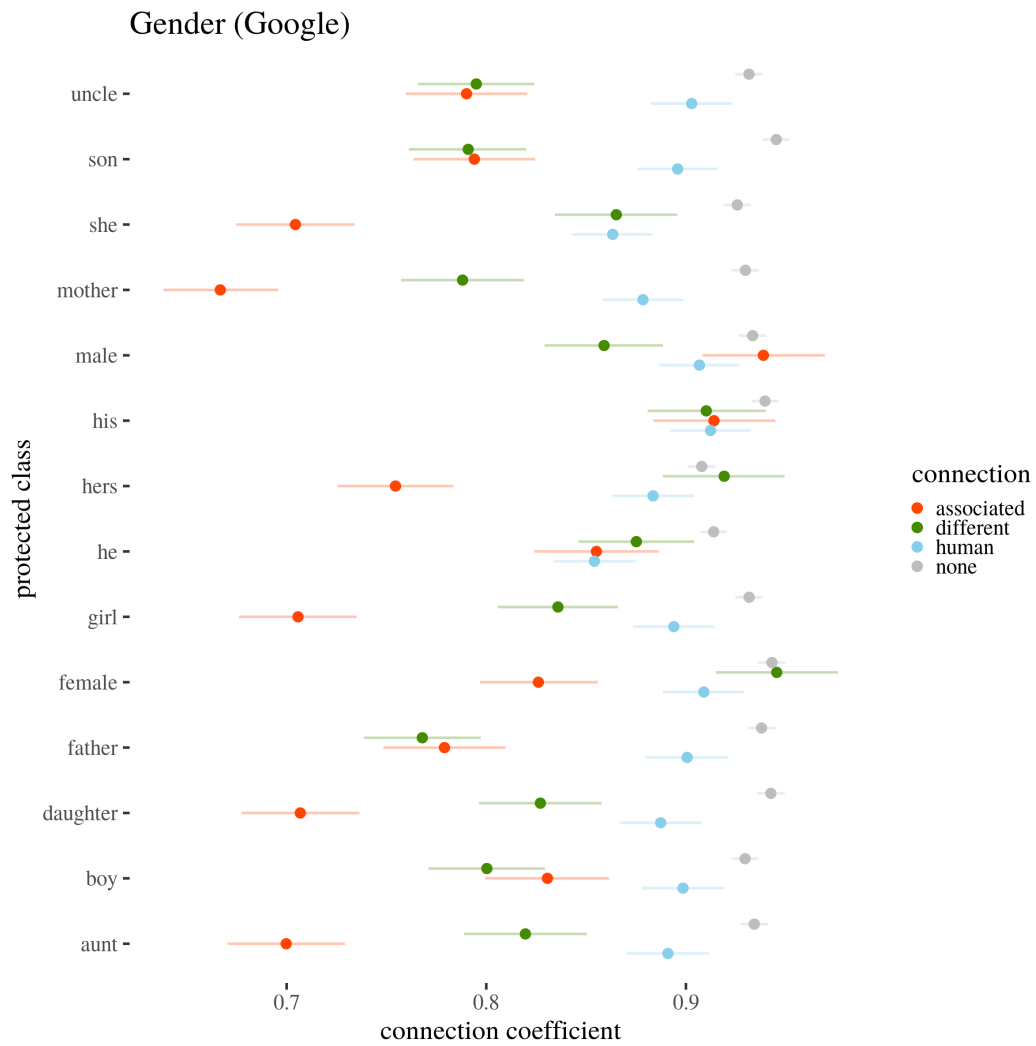


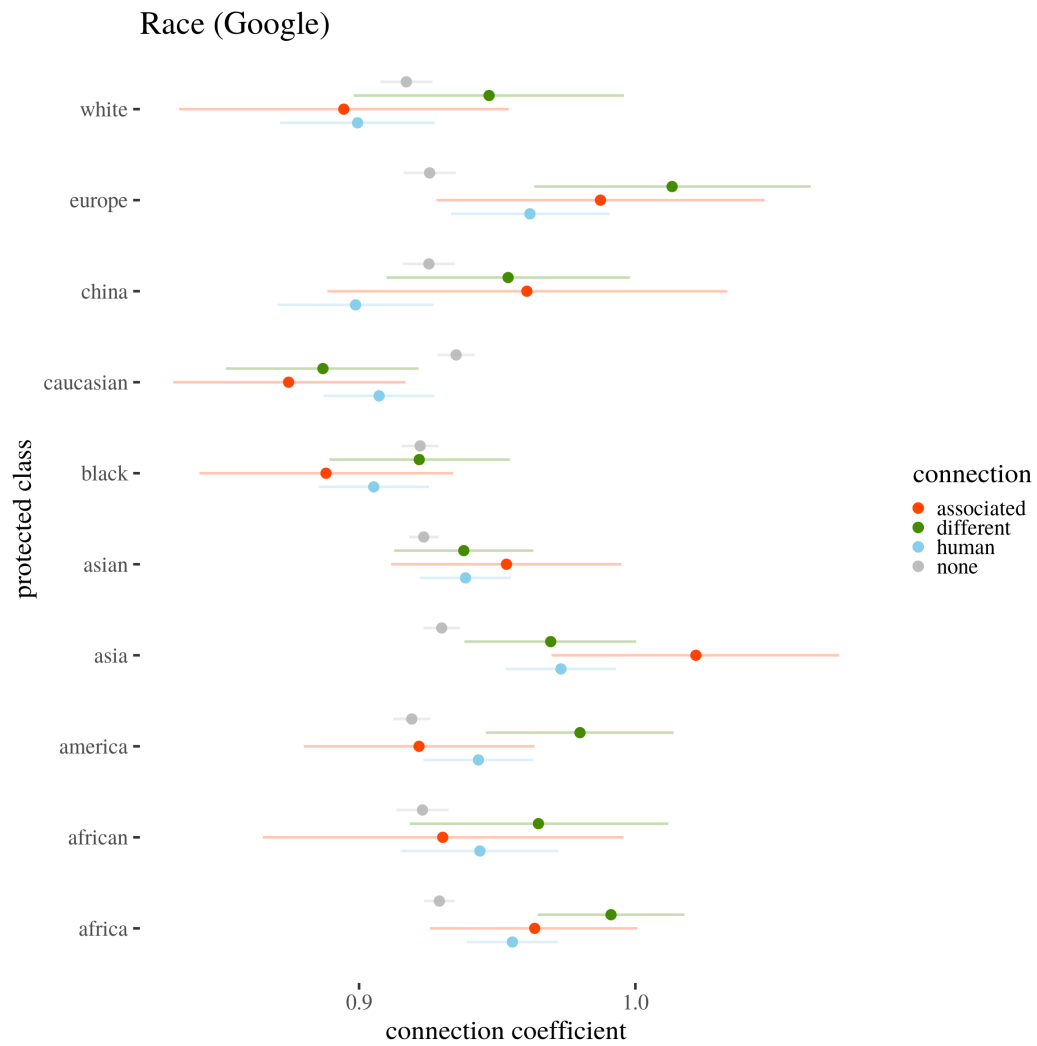


We first encountered the GoogleNews word embedding model in Bolukbasi, Chang, Zou, Saligrama, & Kalai (2016). They use it for their calculations as mentioned in Tolga Code. We decided to compare the results obtained with the use of Reddit L2 model and the ones that we got by GoogleNews model. The details of GoogleNews model can be found here Google News model. One of the main differences between the models is that Reddit word embeddings have 50 dimensions and GoogleNews word embeddings have 300 dimensions. As the dimension is increasing, the vector can capture much more information but this is not always the case. According to some researchers the

amount of dimensions should be chosen by taking into account the type of corpus and its features. In our case both Reddit and GoogleNews models are already trained and ready to use so we do not analyse further the choice of their hyperparameters.







Chapter 5

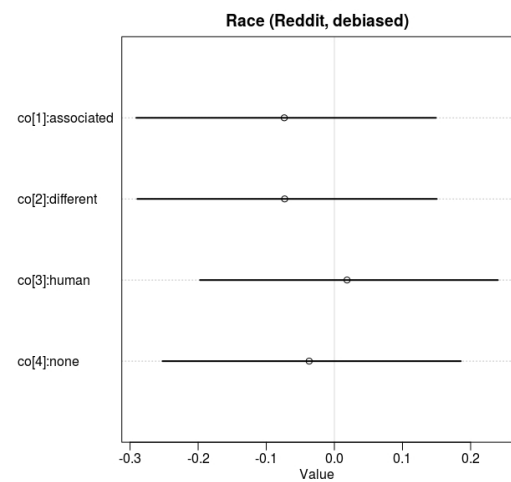
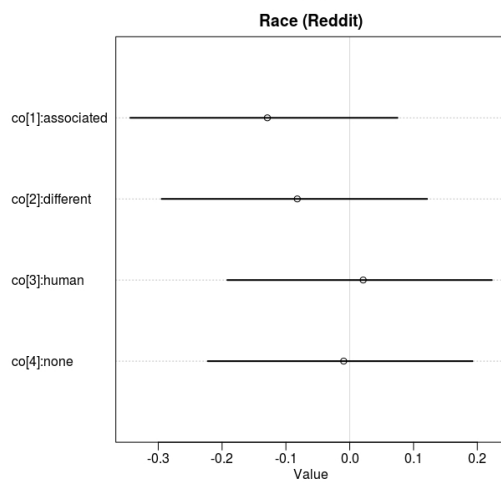
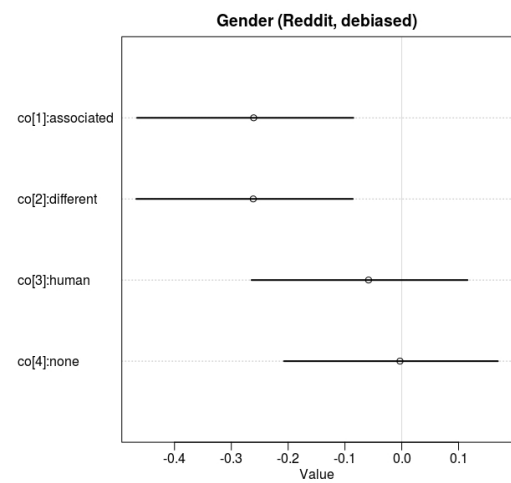
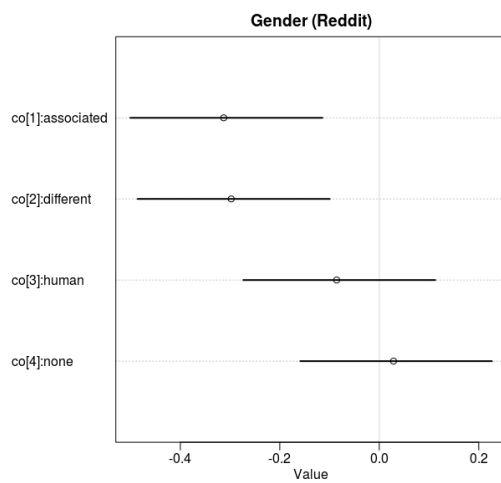
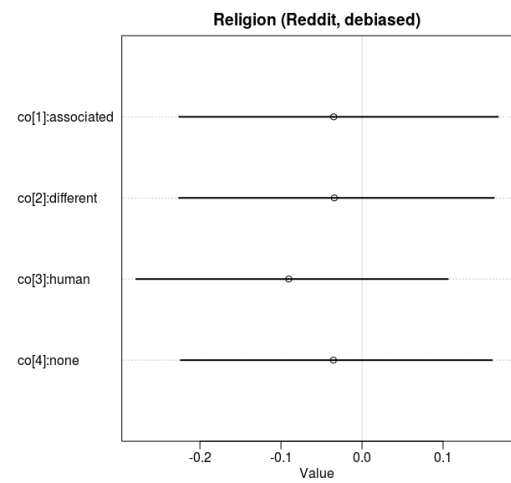
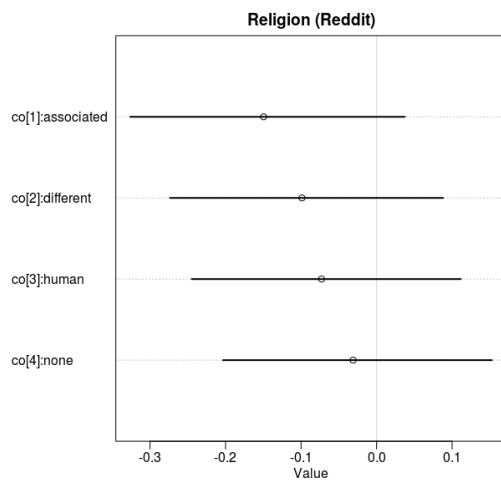
The role of debiasing

We also applied to the word embeddings hard debiasing method found in (manzini2019black?). After debiasing the Reddit word embeddings we calculated the cosine distance again to see if there are any significant changes in terms of the similarity between protected words and the attributes.

5.1 Dataset-level coefficients after debiasing

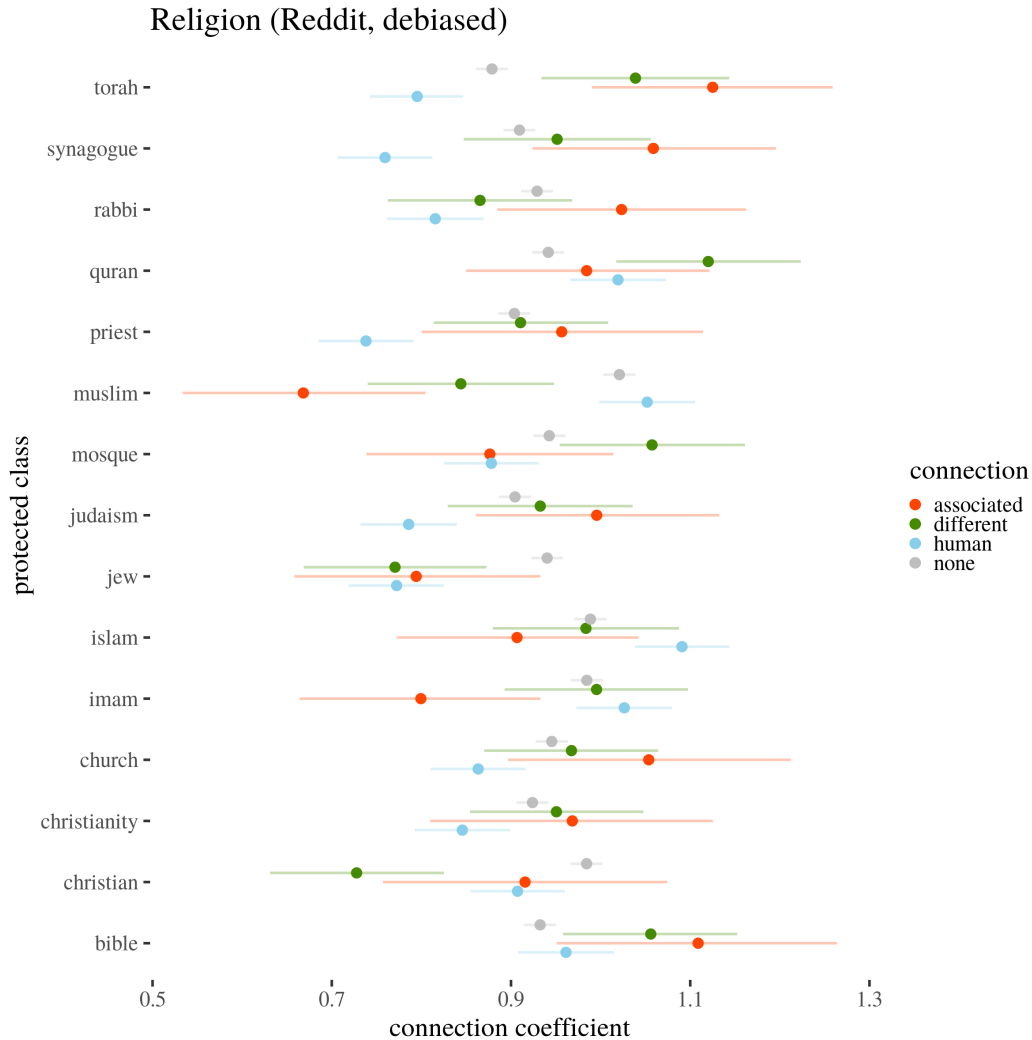
First, let's look at coefficient estimated for the whole datasets, as compared to their estimation prior to debiasing:

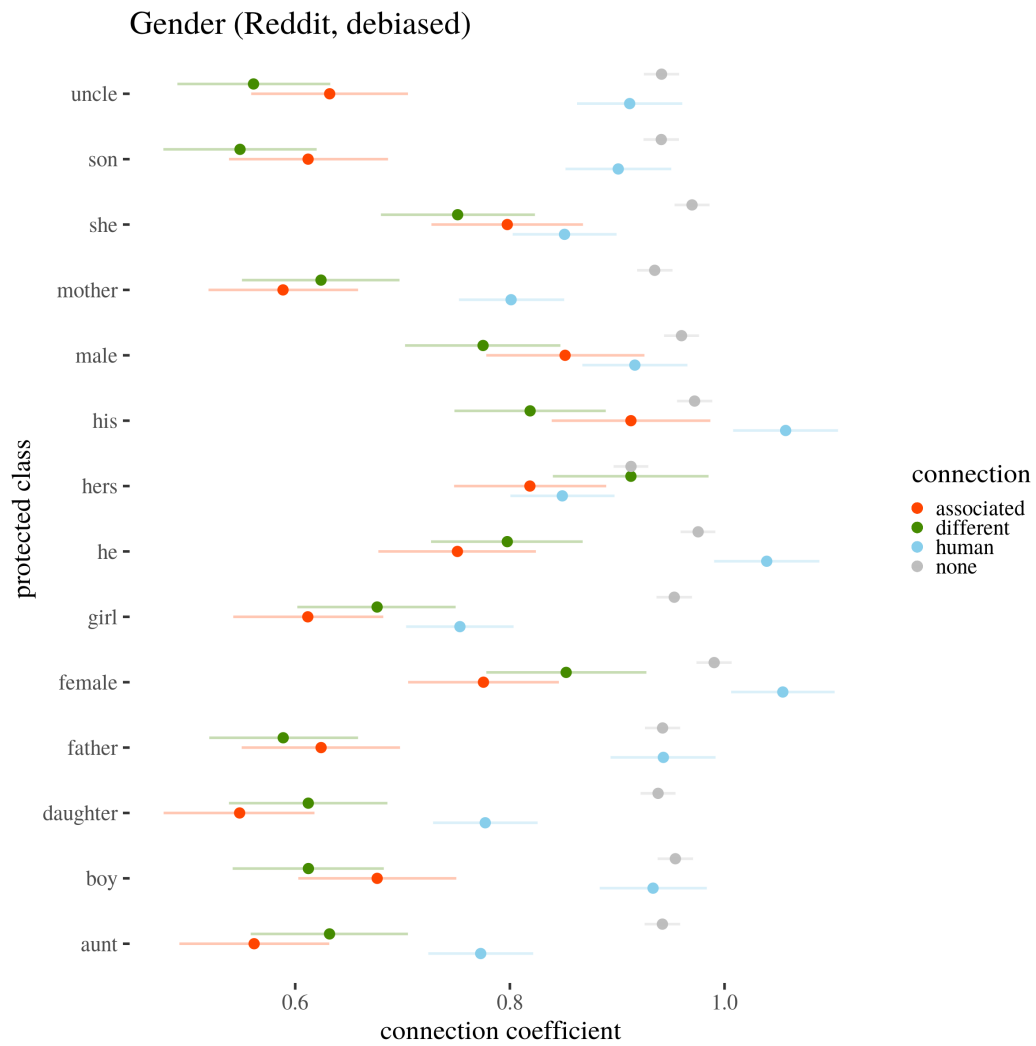
One may observe very slight changes in the estimated coefficients. It seems that in religion dataset the changes are the most significant as the mean moves towards zero which stands for no similarity between words.

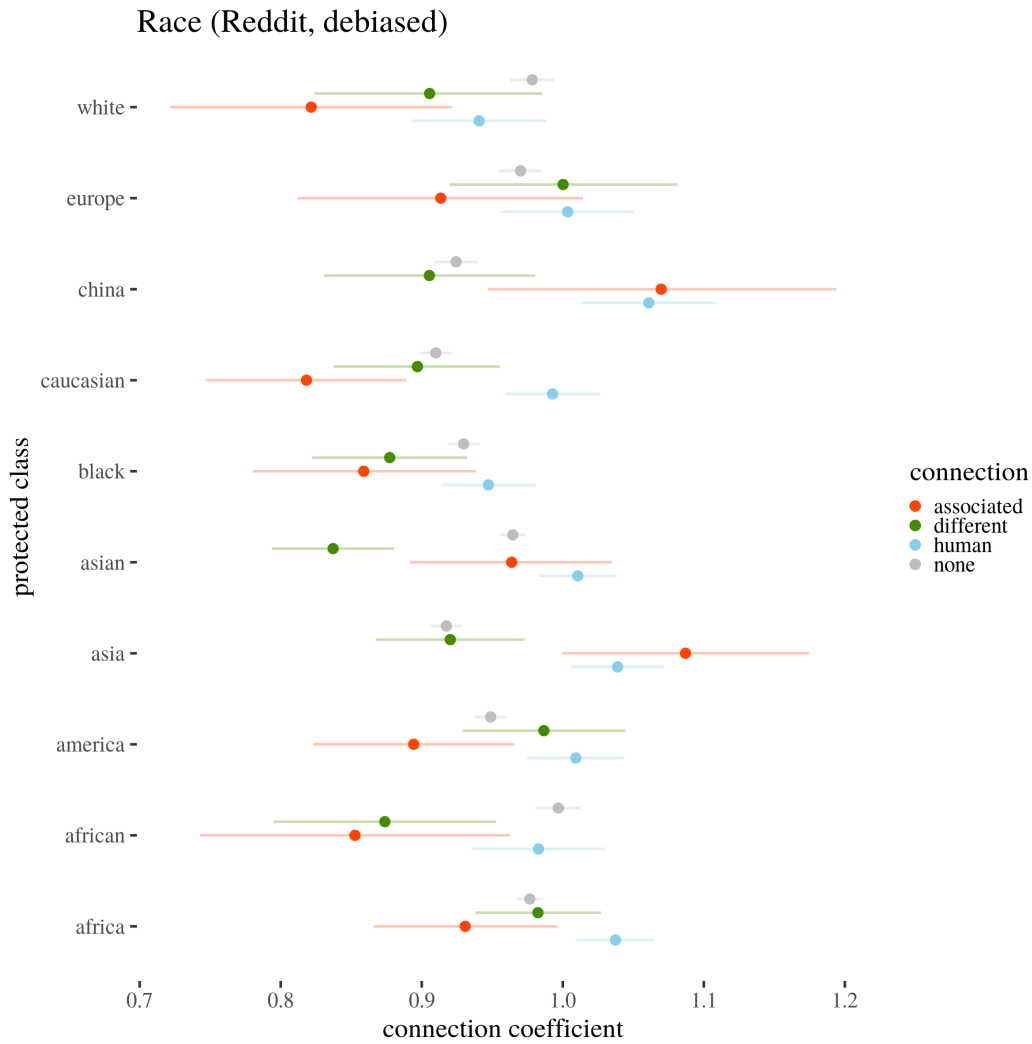


5.2 Protected classes after debiasing

Now, perhaps, the effects of debiasing will be better appreciated if we look at the level of protected words. After all, the hope is, the situation of extremely ill-positioned protected words have improved?







One may notice

- humans in religion got closer (should they?)
- change in Gender is really minor, still zero out of HPDI range, note how different is very similar
- small improvement in race, at the price of moving neutral terms closer

In (manzini2019black?) it is mentioned that debiasing with this method that is based on bias component removal may be insufficient. The reason is that words containing

similar biases may still cluster together even after debiasing.

Chapter 6

Discussion

It is worth diving deeper into some observations that arise while analyzing the results.

- the distributions of cosine distances of whole population and individual words can vary enormously
- including uncertainty shows how in some cases it is not clear why a word is according to this method classified as anomalous
- debiasing with the method provided by (**manzini2019black?**) does not solve the bias issues

Chapter 7

Summary

7.1 Future work:

Bolukbasi, T., Chang, K.-W., Zou, J. Y., Saligrama, V., & Kalai, A. (2016). Man is to computer programmer as woman is to homemaker? Debiasing word embeddings. *CoRR*, *abs/1607.06520*. Retrieved from <http://arxiv.org/abs/1607.06520>