coin: A Computational Framework for Conditional Inference

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1 Introduction

The coin package implements a unified approach for conditional inference procedures commonly known as *permutation tests*. The theoretical basis of design and implementation is the unified framework for permutation tests given by Strasser and Weber (1999). For a very flexible formulation of multivariate linear statistics, Strasser and Weber (1999) derived the conditional expectation and covariance of the conditional (permutation) distribution as well as the multivariate limiting distribution.

Conditional counterparts of a large amount of classical (unconditional) test procedures for continuous, categorical and censored data are part of this framework, for example the Cochran-Mantel-Haenszel test for independence in general contingency tables, linear association tests for ordered categorical data, linear rank tests and multivariate permutation tests.

The conceptual framework of permutation tests by Strasser and Weber (1999) for arbitrary problems is available via the generic independence_test. Because

concenience functions for the most prominent problems are available, users will not have to use this extremely flexible procedure. Currently, the conditional variants of the following test procedures are available:

oneway_testtwo- and K-sample permutation testwilcox_testWilcoxon-Mann-Whitney rank sum testnormal_testvan der Waerden normal quantile testmedian_testMedian test

cmh_testCochran-Mantel-Haenszel testlbl_testlinear-by-linear association testsurv_testtwo- and K-sample logrank testmaxstat_testmaximally selected statistics)

spearman_testSpearman's testfriedman_testFriedman test

wilcoxsign_test Wilcoxon-Signed-Rank test marginal homogenity test.

Those convenience functions essentially perform a certain transformation of the data, e.g., a rank transformation, and call <code>independence_test</code> for the computation of linear statistics, expectation and covariance and the test statistic as well as their null-distribution. The exact null-distribution can be approximated either by the asymptotic distribution or via conditional Monte-Carlo for all test procedures, the exact null-distribution is available for special cases. Moreover, all test procedures allow for the specification of blocks for stratification.

2 Permutation Tests

In the following we assume that we are provided with n observations

$$(\mathbf{Y}_i, \mathbf{X}_i, w_i, b_i), \quad i = 1, \dots, n.$$

The variables **Y** and **X** from sample spaces \mathcal{Y} and \mathcal{X} may be measured at arbitrary scales and may be multivariate as well. In addition to those measurements, case weights w and a factor b coding blocks may be available. For the sake of simplicity, we assume $w_i = 1$ and $b_i = 0, i = 1, \ldots, n$ for the moment.

We are interested in testing the null-hypothesis of independence of \mathbf{Y} and \mathbf{X}

$$H_0: D(\mathbf{Y}|\mathbf{X}) = D(\mathbf{Y})$$

against arbitrary alternatives. Strasser and Weber (1999) suggest to derive scalar test statistics for testing H_0 from multivariate linear statistics of the form

$$\mathbf{T} = \operatorname{vec}\left(\sum_{i=1}^{n} w_{i} g(\mathbf{X}_{i}) h(\mathbf{Y}_{i}, (\mathbf{Y}_{1}, \dots, \mathbf{Y}_{n}))^{\top}\right) \in \mathbb{R}^{pq}.$$
 (1)

Here, $g: \mathcal{X} \to \mathbb{R}^p$ is a transformation of the **X** measurements and the *influence* function $h: \mathcal{Y} \times \mathcal{Y}^n \to \mathbb{R}^q$ depends on the responses $(\mathbf{Y}_1, \dots, \mathbf{Y}_n)$ in a permutation symmetric way. We will give specific examples how to choose g and h later on.

The distribution of \mathbf{T} depends on the joint distribution of \mathbf{Y} and \mathbf{X} , which is unknown under almost all practical circumstances. At least under the null hypothesis one can dispose of this dependency by fixing $\mathbf{X}_1, \ldots, \mathbf{X}_n$ and conditioning on all possible permutations S of the responses $\mathbf{Y}_1, \ldots, \mathbf{Y}_n$. This principle leads to test procedures known as permutation tests.

The conditional expectation $\mu \in \mathbb{R}^{pq}$ and covariance $\Sigma \in \mathbb{R}^{pq \times pq}$ of **T** under H_0 given all permutations $\sigma \in S$ of the responses are derived by Strasser and Weber (1999):

$$\mu = \mathbb{E}(\mathbf{T}|S) = \operatorname{vec}\left(\left(\sum_{i=1}^{n} w_{i}g(\mathbf{X}_{i})\right)\mathbb{E}(h|S)^{\top}\right),$$

$$\Sigma = \mathbb{V}(\mathbf{T}|S)$$

$$= \frac{\mathbf{w}.}{\mathbf{w}. - 1}\mathbb{V}(h|S) \otimes \left(\sum_{i} w_{i}g(\mathbf{X}_{i}) \otimes w_{i}g(\mathbf{X}_{i})^{\top}\right)$$

$$- \frac{1}{\mathbf{w}. - 1}\mathbb{V}(h|S) \otimes \left(\sum_{i} w_{i}g(\mathbf{X}_{i})\right) \otimes \left(\sum_{i} w_{i}g(\mathbf{X}_{i})\right)^{\top}$$

$$(2)$$

where $\mathbf{w}_{\cdot} = \sum_{i=1}^{n} w_i$ denotes the sum of the case weights, and \otimes is the Kronecker product. The conditional expectation of the influence function is

$$\mathbb{E}(h|S) = \mathbf{w}_{\cdot}^{-1} \sum_{i} w_{i} h(\mathbf{Y}_{i}, (\mathbf{Y}_{1}, \dots, \mathbf{Y}_{n})) \in \mathbb{R}^{q}$$

with corresponding $q \times q$ covariance matrix

$$\mathbb{V}(h|S) = \mathbf{w}_{\cdot}^{-1} \sum_{i} w_{i} \left(h(\mathbf{Y}_{i}, (\mathbf{Y}_{1}, \dots, \mathbf{Y}_{n})) - \mathbb{E}(h|S) \right)$$
$$\left(h(\mathbf{Y}_{i}, (\mathbf{Y}_{1}, \dots, \mathbf{Y}_{n})) - \mathbb{E}(h|S) \right)^{\top}.$$

Having the conditional expectation and covariance at hand we are able to standardize a linear statistic $\mathbf{T} \in \mathbb{R}^{pq}$ of the form (1). Univariate test statistics c mapping an observed linear statistic $\mathbf{t} \in \mathbb{R}^{pq}$ into the real line can be of arbitrary form. An obvious choice is the maximum of the absolute values of the standardized linear statistic

$$c_{\max}(\mathbf{t}, \mu, \Sigma) = \max \left| \frac{\mathbf{t} - \mu}{\operatorname{diag}(\Sigma)^{1/2}} \right|$$

utilizing the conditional expectation μ and covariance matrix Σ . The application of a quadratic form $c_{\rm quad}(\mathbf{t}, \mu, \Sigma) = (\mathbf{t} - \mu)\Sigma^+(\mathbf{t} - \mu)^\top$ is one alternative, although computationally more expensive because the Moore-Penrose inverse Σ^+ of Σ is involved.

The conditional distribution and thus the P-value of the statistics $c(\mathbf{t}, \mu, \Sigma)$ can be computed in several different ways. For some special forms of the linear statistic, the exact distribution of the test statistic is trackable. For two-sample problems, the shift-algorithm by Streitberg and Röhmel (1986) and Streitberg and Röhmel (1987) and the split-up algorithm by van de Wiel (2001) are implemented as part of the package. Conditional Monte-Carlo procedures can be used to approximate the exact distribution. Strasser and Weber (1999) proved (Theorem 2.3) that the conditional distribution of linear statistics T with conditional expectation μ and covariance Σ tends to a multivariate normal distribution with parameters μ and Σ as $n, s \to \infty$. Thus, the asymptotic conditional distribution of test statistics of the form c_{max} is normal and can be computed directly in the univariate case (pq = 1) or approximated by means of quasi-randomized Monte-Carlo procedures in the multivariate setting (Genz, 1992). For quadratic forms c_{quad} which follow a χ^2 distribution with degrees of freedom given by the rank of Σ (Theorem 6.20, Rasch, 1995), exact probabilities can be computed efficiently.

3 Illustrations and Applications

The main workhorse independence_test essentially allows for the specification of \mathbf{Y}, \mathbf{X} and b through a formula interface of the form $\mathbf{y} \sim \mathbf{x} \mid \mathbf{b}$, weights can be defined by a formula with one variable on the right hand side only. Four additional arguments are available for the specification of the transformation g (xtrans), the influence function h (ytrans), the form of the test statistic c (teststat) and the null-distribution (distribution).

Independent K-Sample Problems. When we want to compare the distribution of an univariate qualitative response Y in K groups given by a factor X at K levels, the transformation g is the dummy matrix coding the groups and h is either the identity transformation or a some form of rank transformation.

For example, the Kruskal-Wallis test may be computed as follows (example taken from Hollander and Wolfe, 1999, Table 6.3, page 200):

```
R> library(coin)
```

```
Loading required package: survival
Loading required package: splines
Loading required package: mvtnorm

R> YOY <- data.frame(length = c(46, 28, 46, 37, 32,
+ 41, 42, 45, 38, 44, 42, 60, 32, 42, 45, 58,
+ 27, 51, 42, 52, 38, 33, 26, 25, 28, 28, 26,
+ 27, 27, 27, 31, 30, 27, 29, 30, 25, 25, 24,
+ 27, 30), site = factor(c(rep("I", 10), rep("II",
+ 10), rep("III", 10), rep("IV", 10))))
R> it <- independence_test(length ~ site, data = YOY,
```

```
ytrafo = function(data) trafo(data, numeric_trafo = rank),
      teststat = "quadtype")
R> it
         Asymptotical General Independence Test
data: length by groups I, II, III, IV
T = 22.8524, df = 3, p-value = 4.335e-05
The linear statistic T is the sum of the ranks in each group and can be extracted
R> statistic(it, "linear")
    [,1]
Ι
     278
     307
II
III 119
ΙV
     116
The conditional expectation and covariance are available from
R> expectation(it)
    [,1]
Ι
     205
     205
II
III 205
     205
ΙV
R> covariance(it)
           [,1]
                      [,2]
                                 [,3]
                                            [,4]
[1,] 1019.0385 -339.6795 -339.6795 -339.6795
[2,] -339.6795 1019.0385 -339.6795 -339.6795
[3,] -339.6795 -339.6795 1019.0385 -339.6795
[4,] -339.6795 -339.6795 -339.6795 1019.0385
and the standardized linear statistic (\mathbf{T} - \mu) \operatorname{diag}(\Sigma)^{1/2} is
R> statistic(it, "standardized")
          [,1]
Ι
     2.286797
```

Since a quadratic form of the test statistic was requested via teststat = "quadtype", the test statistic is

3.195250

III -2.694035 IV -2.788013

ΙΙ

```
R> statistic(it)
```

[1] 22.85242

By default, the asymptotic distribution of the test statistic is computed, the p-value is

```
R> pvalue(it)
```

[1] 4.334659e-05

Life is much simpler with convenience functions very similar to those available in package stats for a long time. The exact null-distribution of the Kruskal-Wallis test can be approximated by 9999 Monte-Carlo replications via

Approximative Kruskal-Wallis Test

```
data: length by groups I, II, III, IV T = 22.8524, p-value < 2.2e-16
```

with p-value (and 99% confidence interval) of

R> pvalue(kw)

[1] 0

99 percent confidence interval:

0.000000000 0.0005297444

Of course it is possible to choose a c_{\max} type test statistic instead of a quadratic form.

Independence in Contingency Tables. Independence in general two- or three-dimensional contingency tables can be tested by the Cochran-Mantel-Haenszel test. Here, both g and h are dummy matrices (example data from Agresti, 2002, Table 7.8, page 288):

```
R> data(jobsatisfaction, package = "coin")
R> it <- cmh_test(jobsatisfaction)
R> it
```

Asymptotical Generalised Cochran-Mantel-Haenszel Test

The standardized contingency table allowing for an inspection of the deviation from the null-hypothesis of independence of income and jobsatisfaction (stratified by gender) is

R> statistic(it, "standardized")

	Very Dissatisfied A I	Little Dissatisfied
<5000	1.3112789	0.69201053
5000-15000	0.6481783	0.83462550
15000-25000	-1.0958361	-1.50130926
>25000	-1.0377629	-0.08983052
	Moderately Satisfied	Very Satisfied
<5000	-0.2478705	-0.9293458
5000-15000	0.5175755	-1.6257547
15000-25000	0.2361231	1.4614123
>25000	-0.5946119	1.2031648

Ordered Alternatives. Of course, both job satisfaction and income are ordered variables. When **Y** is measured at J levels and **X** at K levels, **Y** and **X** are associated with score vectors $\xi \in \mathbb{R}^J$ and $\gamma \in \mathbb{R}^K$, respectively. The linear statistic is now a linear combination of the linear statistic **T** of the form

$$\mathbf{MT} = \operatorname{vec}\left(\sum_{i=1}^{n} w_{i} \gamma^{\top} g(\mathbf{X}_{i}) \left(\xi^{\top} h(\mathbf{Y}_{i}, (\mathbf{Y}_{1}, \dots, \mathbf{Y}_{n}))^{\top}\right).\right)$$

By default, scores are $\xi = 1, \dots, J$ and $\gamma = 1, \dots, K$.

R> lbl_test(jobsatisfaction)

Asymptotical Linear-by-Linear Association Test

```
data: Job.Satisfaction (ordered) by
          groups <5000 < 5000-15000 < 15000-25000 < >25000
          stratified by Gender
T = 6.6235, df = 1, p-value = 0.01006
```

The scores ξ and γ can be specified to the linear-by-linear association test via a list those names correspond to the variable names

Asymptotical Linear-by-Linear Association Test

4 Quality Assurance

The test procedures implemented in package coin are continuously checked against results obtained by the corresponding implementations in package stats (where available). In addition, the test statistics and exact, approximativ and asymptotic p-values for data examples given in the StatXact-6 user manual Mehta and Patel (2003) are compared with the results reported in the StatXact-6 manual. For details on the test procedures we refer to the R-transcript files in directory coin/tests.

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