

lrgpr: Low Rank Gaussian Process Regression

Gabriel E. Hoffman¹

lrgpr version 0.1.3 as of July 29, 2014

Abstract

The **lrgpr** package provides an interactive R interface for fitting linear mixed models, also known as Gaussian Process Regression, on very large datasets. The package provides user-friendly interfaces for the linear mixed model and a computationally efficient extension termed Low Rank Gaussian Process Regression, as well as standard linear and logistic regression models. **lrgpr** allows fitting millions of regression models on a desktop computer by using an efficient implementation, parallelization and out-of-core computing for datasets that are too large to fit in main memory.

Contents

1	Overview	2
2	Integration with the R environment: Using the <code>big.matrix</code> format	2
2.1	Example: Principal components analysis	3
3	Regression with R's <code>glm</code>	4
4	Fitting many regressions with <code>glmApply</code>	7
5	Linear mixed models with <code>lrgpr</code>	7
5.1	The genetic similarity matrix and its spectral decomposition	8
5.1.1	Full rank linear mixed model	8
5.1.2	Low rank linear mixed model: (low rank gaussian process regression)	10
5.2	Proximal contamination	12
6	Fitting many <code>lrgpr</code> regressions with <code>lrgprApply</code>	12
6.1	Re-estimating variance components (δ)	13
6.2	Proximal contamination	13
6.2.1	Interaction model	13
7	Implementation	13

¹Icahn Institute for Genomics and Multiscale Biology, Department of Genetics and Genomic Sciences, Icahn School of Medicine at Mount Sinai, New York, New York, USA

1 Overview

Genome-wide association studies (GWAS) are a widely used approach to identify genetic variants that are associated with variation in a phenotype of interest. Tests of association usually take the form of statistical hypothesis tests in a regression model that accounts for known confounding variables such as sex. Yet due to the complex nature of GWAS datasets, more sophisticated methods are used in order to maintain power while controlling the false positive rate. The linear mixed model is the state-of-the-art method to account for the confounding effects of kinship and population structure in GWAS analysis, and much recent work has focused on this approach (Hoffman, 2013; Listgarten, et al., 2012; Lippert, et al., 2011; Kang, et al., 2010; Zhou and Stephens, 2012).

To date, most software for GWAS data has been designed for a “one-size-fits-all” analysis. Yet as GWAS datasets have become increasingly complex and heterogeneous, there is growing potential for interactive, exploratory data analysis. The **lrgpr** package provides a user-friendly, interactive and computationally efficient analysis framework that facilitates custom, exploratory analysis of large GWAS datasets.

The **lrgpr** package provides:

- Seamless, interactive R interface to arbitrarily large datasets through **bigmemory**’s `big.matrix`
- Scalable linear or logistic regression for millions of hypothesis tests using `glmApply`
- Fitting a full or low rank linear mixed model with `lrgpr`
- Data-adaptive construction of the genetic similarity matrix for the linear mixed model
- Scalable linear mixed model regression for millions of hypothesis tests using `lrgprApply`
- Ability to define arbitrary interaction models and perform composite hypothesis tests with `glmApply`, `lrgpr` and `lrgprApply`

Here I demonstrate some applications of the `lrgpr` package. See the help modules in R like `?lrgpr` and `?lrgprApply` for more detailed documentation.

2 Integration with the R environment: Using the `big.matrix` format

R typically stores all data in memory (i.e. RAM) for fast access. However, GWAS datasets can easily be too large to store in memory. For example, a dataset with 10,000 individuals and 1,000,000 SNPs would require ~ 80 Gb to store in R in its standard format (i.e. as a `double`), and much larger datasets are now very common. Instead, **lrgpr** uses **bigmemory**’s `big.matrix` format to store a very large dataset in binary form directly on the hard drive and avoid using RAM while still allowing interactive access to the data. Thus the `big.matrix` format allows analysis of arbitrarily large datasets on any machine with sufficient hard drive capacity.

Arbitrary data can be converted to `big.matrix` format using **bigmemory**’s `as.big.matrix`. Alternatively, large GWAS data can be converted more efficiently:

Start by loading the package:

```
library(lrgpr)

# Path to Plink files
path <- system.file(package = 'lrgpr')
tped_file <- paste(path, "/extdata/test.tped", sep="")
fam_file <- paste(path, "/extdata/test.tfam", sep="")
map_file <- paste(path, "/extdata/test.map", sep="")

# Convert TPED file to binary format
# Create a binary data file: test.binary
# and a binary file describing this data: test.binary_desc
convertToBinary(tped_file, "./test.binary", "TPED")
```

Supported GWAS data formats include TPED, GEN and DOSAGE.

Data is loaded into R by pointing a variable to the description file:

```
# attach data by reading the description file
# NOTE: A big.matrix MUST be re-attached in each R session.
#       It CANNOT be restored from a previous R session using save()/load()
X <- attach.big.matrix("./test.binary_descr")
```

Note that this data is not loaded into memory. Instead X points to the location of the data on the hard drive, and the data is loaded into memory only when the user asks for it. The user can treat X as a standard R matrix with the caveat that user must be careful not to load too much data into memory at the same time. For example, the 10th SNP can be accessed seamlessly:

```
# data is only loaded into memory when it is accessed
X[,10]
```

and the user can perform arbitrary operations on each column individually without loading the whole dataset at the same time:

```
# data is only loaded into memory when it is accessed
column_means <- c()

for(j in 1:ncol(X)) {
  column_means[j] <- mean(X[,j])
}
```

Accessing the data in this way creates a standard R matrix that can be passed to any function. Thus the user can process X as if it were stored in memory; the complexity of storing the data on the hard drive is completely hidden from the user.

Processing the entire dataset in R can be costly for large datasets, so **lrgpr** provides functions to report common statistics more efficiently by performing all computations in C/C++:

```
# Report allele frequencies using C/C++ backend
freq <- getAlleleFreq(X)
```

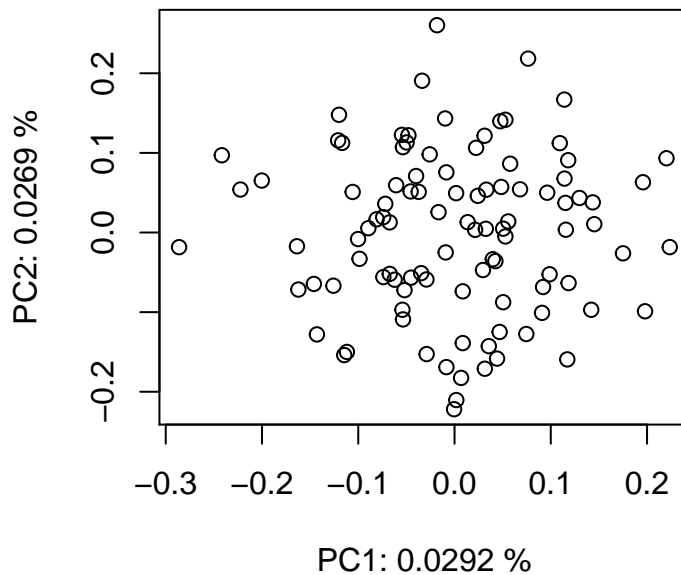
2.1 Example: Principal components analysis

It is simple to apply standard statistical methods to a subset of the data pointed to by X. This example shows an application of principal components analysis common in GWAS analysis (Price, et al., 2006).

```
# Extract 200 markers and perform SVD on scaled markers
# Use round so that indices are integers
j <- round(seq(1, ncol(X), length.out=200))

# Note that a minor allele frequency (maf) of 0 will produce an error
# since scale() divides by the maf
# This operation is only valid on markers that vary in this dataset
# Use getAlleleFreq() (or getAlleleVariance() for dosage data)
# to filter out markers that don't have variation
dcmp <- svd(scale(X[,j]))

# Make plot
percentVar <- dcmp$d^2 / sum(dcmp$d^2)
xlab <- paste("PC1:", format(percentVar[1], digits=3), "%")
ylab <- paste("PC2:", format(percentVar[2], digits=3), "%")
plot(dcmp$u[,1:2], xlab=xlab, ylab=ylab)
```



3 Regression with R's glm

R fits generalized linear models using the `glm` function that provides flexibility in defining the regression model. I describe `glm` very briefly here so that I can later describe how **lrgpr** mirrors this functionality. For example, we can read Plink's FAM file and fit a simple model:

```
# Read FAM file
FAM <- read.fam(fam_file)
y <- FAM$phenotype
sex <- FAM$sex
```

```
# Fit a linear model
fit <- glm( y ~ sex )
```

We can see the model coefficients and statistics:

```
# Simple view
fit

##
## Call:  glm(formula = y ~ sex)
##
## Coefficients:
## (Intercept)      sex2
##    -0.1169      0.0737
##
## Degrees of Freedom: 99 Total (i.e. Null);  98 Residual
## Null Deviance:      92.8
## Residual Deviance: 92.7  AIC: 282
```

```

# Expanded view
summary(fit)

##
## Call:
## glm(formula = y ~ sex)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.2615  -0.7680   0.0955   0.7277   2.1681
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  -0.1169     0.1336  -0.87    0.38
## sex2          0.0737     0.1948   0.38    0.71
##
## (Dispersion parameter for gaussian family taken to be 0.9454)
##
##      Null deviance: 92.786  on 99  degrees of freedom
## Residual deviance: 92.650  on 98  degrees of freedom
## AIC: 282.2
##
## Number of Fisher Scoring iterations: 2

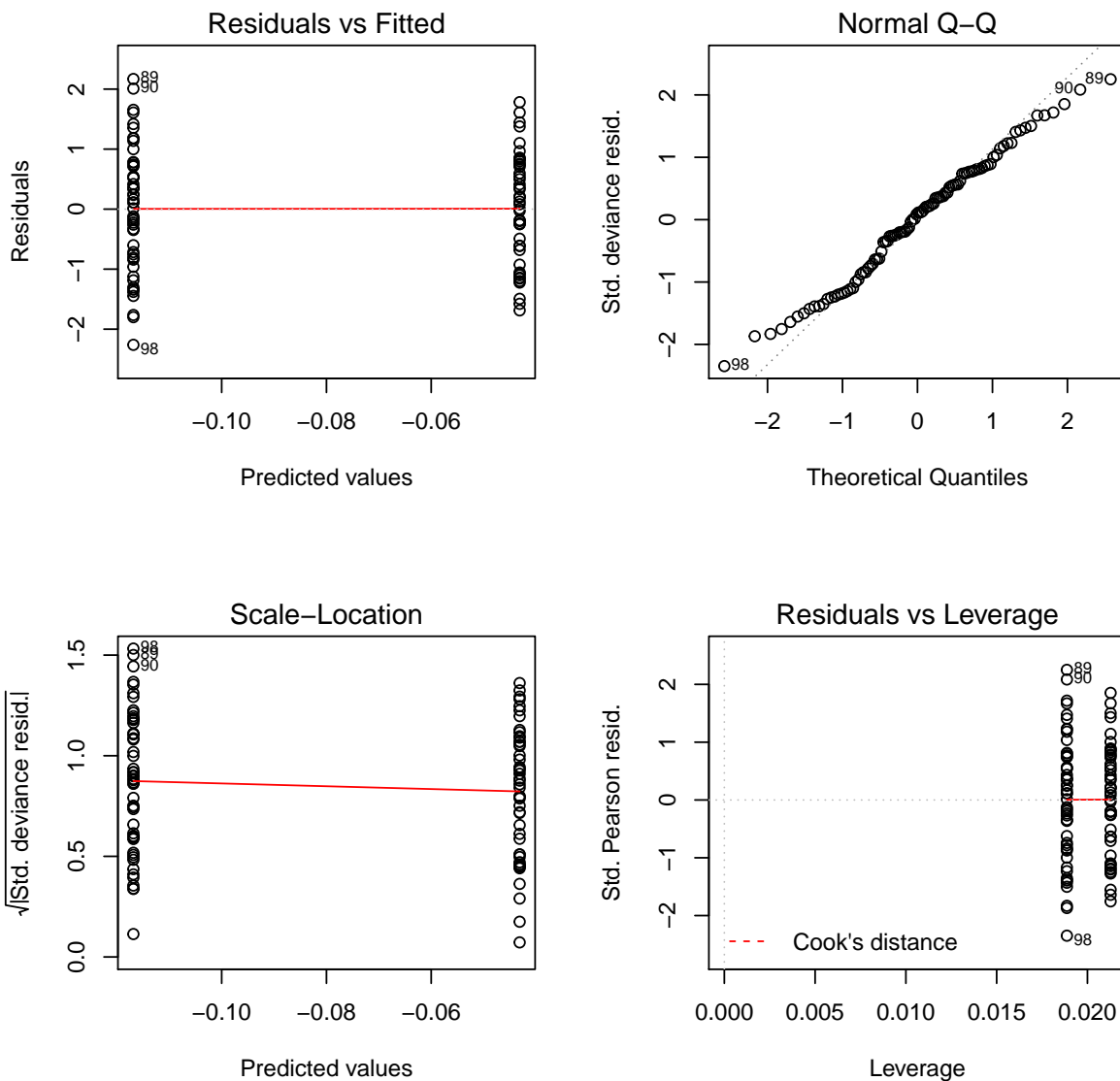
```

and plot model diagnostics:

```

# Plot diagnostics
par(mfrow=c(2,2))
plot(fit)

```



The user can also specify and perform hypothesis tests on more complex models involving interactions:

```
# Fit a linear model
fit <- glm( y ~ sex + X[,10]*X[,100] )
```

```
library(aod)
# Composite hypothesis test of additive and interaction terms
# for the two markers
# Note that wald.test is from the aod package
wald.test(vcov(fit), coef(fit), Terms=3:5)

## Wald test:
## -----
##
## Chi-squared test:
## X2 = 1.2, df = 3, P(> X2) = 0.76
```

4 Fitting many regressions with glmApply

Analysis of GWAS data involves fitting a regression model for each SNP individually. The user could write a simple for loop that calls the glm function, but the computation can be accelerated by doing the calculation in the C/C++ backend and reporting the results.

A standard analysis can be performed with **lrgpr**'s glmApply:

```
# Standard GWAS regression analysis
pValues <- glmApply( y ~ SNP + sex, features=X, terms=2)$pValues
```

where the formula follows standard glm syntax, features specifies the big.matrix or standard R matrix with columns as genetic markers, and terms specifies which variables from the formula are in the hypothesis test. Note that SNP is a placeholder for each successive column in features, and if terms is omitted the hypothesis test includes all terms corresponding to SNP:

```
# Standard GWAS regression analysis
pValues <- glmApply( y ~ SNP + sex, features=X)$pValues
```

This code is the same as:

```
pValuesR <- c()

for(j in 1:ncol(X)) {
  # fit model
  fit <- glm(y ~ X[,j] + sex)

  # perform hypothesis test
  pValuesR[j] <- wald.test(vcov(fit), coef(fit), Terms=2)$result$chi2[3]
}
```

but calling glmApply is simpler and much faster for large datasets.

Note that *markers* can easily be excluded from the analysis using the `cincl` or `cexcl` arguments to glmApply. See ?glmApply for details. However, there is currently no simple way to exclude *samples* from an analysis with glmApply (or lrgprApply). This can be accomplished by regenerating the binary file with `convertToBinary`, or by using the R interface to the data matrix to process `X[i, j]` and using glm or lrgpr. But note that you must select a small enough subset of the data with `X[i, j]` for it to fit in memory.

Note that glmApply can also efficiently fit multivariate regression models with multiple response variables. See ?glmApply for details.

5 Linear mixed models with lrgpr

The lrgpr function fits full and low rank linear mixed models and follows the syntax and behavior of glm.

The linear mixed model has the form:

$$\mathbf{y} \sim \mathcal{N}(\mathbf{X}\boldsymbol{\beta}, \mathbf{K}\sigma_a^2 + \mathbf{I}\sigma_e^2),$$

where \mathbf{y} ($n \times 1$) is the vector of phenotype values, n is the sample size, \mathbf{X} ($n \times c$) is the design matrix of c fixed effects, $\boldsymbol{\beta}$ ($c \times 1$) is the vector of coefficients, \mathbf{K} ($n \times n$) is the genetic similarity matrix (GSM), \mathbf{I} ($n \times n$) is the identity matrix, σ_a^2 is the magnitude of the genetic variance, and σ_e^2 is the magnitude of the residual variance. The parameters are estimated by maximizing the log-loglikelihood using the algorithms of Lippert, et al. (2011) and Listgarten, et al. (2012).

When \mathbf{K} is full rank, the coefficient estimates and their sample variance are, respectively,

$$\begin{aligned}\hat{\boldsymbol{\beta}} &= (\mathbf{X}^T \boldsymbol{\Omega} \mathbf{X})^{-1} \mathbf{X}^T \boldsymbol{\Omega} \mathbf{y} \\ \hat{\boldsymbol{\Sigma}} &= (\mathbf{X}^T \boldsymbol{\Omega} \mathbf{X})^{-1} \sigma_a^2\end{aligned}$$

where $\Omega = (\mathbf{K} + \mathbf{I} \frac{\sigma_e^2}{\sigma_a^2})^{-1}$ and the values of σ_a^2 and σ_e^2 are already estimated by maximum likelihood. Following standard likelihood theory, the Wald test is

$$\hat{\beta}_h^T (\hat{\Sigma}_h)^{-1} \hat{\beta}_h \sim \chi_{|h|}^2$$

where h specifies the fixed effects being tested and $|h|$ is the number of entries. When \mathbf{K} is not full rank the estimates and hypothesis tests are analogous and follow directly from Listgarten, et al. (2012).

5.1 The genetic similarity matrix and its spectral decomposition

Following the algorithms of Lippert, et al. (2011) and Listgarten, et al. (2012), the genetic similarity matrix enters the linear mixed model only through its spectral decomposition. There are multiple ways to compute \mathbf{K} or its spectral decomposition.

5.1.1 Full rank linear mixed model

When the genetic similarity matrix, \mathbf{K} , is based on a genome-wide set of markers, the linear mixed model is full rank. However, only the spectral decomposition of \mathbf{K} is needed, and this can be computed directly from the genotype matrix in R:

```
# Keep only SNPs where there is variation in this dataset
# With SNP data coded as 0,1,2 screening by the allele
# frequency would be sufficient
# With continuous dosage data, we need to consider the variance of the SNP
v <- getAlleleVariance(X)

# get markers that pass filter
passFilter <- which(v > .01)

# Prune to every 100th marker that passes filter
index <- passFilter[seq(1, length(passFilter), by=100)]

# Perform spectral decomposition on the set of SNPs
# after replacing missing values with the per-SNP mean
# Each SNP is centered and scaled
dcmp <- svd(scale(set_missing_to_mean(X[,index])))
```

We can then fit the linear mixed model using the same syntax as for `glm`:

```
fit <- lrgpr(y ~ sex, dcmp)
```

Alternatively, the genetic similarity matrix \mathbf{K} from another program (Yang, et al., 2011; Kang, et al., 2010; Zhou and Stephens, 2012) can easily be imported into R:

```
# Read in GSM and perform spectral decomposition
K <- read.table(paste(path, "/extdata/K.txt", sep=""))
dcmp <- eigen(K, symmetric=TRUE)
```

Features of lrgpr

Since `lrgpr` behaves like `glm` and we can see the model coefficients and statistics:

```
# Simple view
fit

##
## Call:
## lrgpr(formula = y ~ sex, decomp = dcmp)
##
```



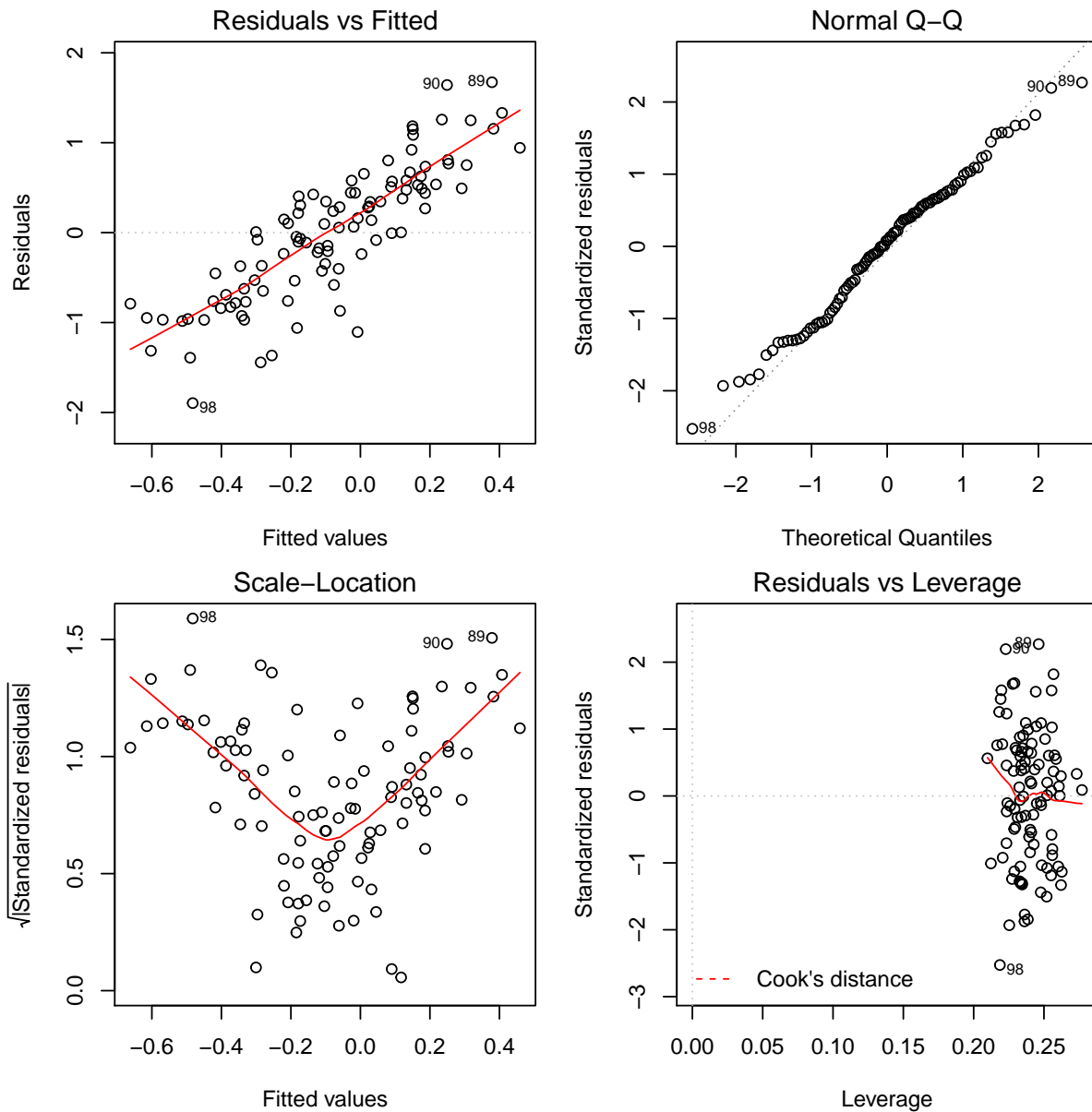
```
## Coefficients:
## (Intercept)      sex2
##      -0.12136      0.08332
##
## Variance components:
## sigSq_a sigSq_e delta
##  0.01489  0.91178 61.23545
```

```
# Expanded view
summary(fit)

##
## Call:
## lrgpr(formula = y ~ sex, decomp = dcmp)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -2.13773 -0.71825  0.08448  0.67686  2.05327
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.12136      0.1312  -0.9249   0.3550
## sex2         0.08332      0.1915   0.4352   0.6634
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.9646 on 97.38 degrees of freedom
## Multiple R-squared:  0.02667 , Adjusted R-squared:  0.0005208
##
##
## Variance components:
## sigSq_a sigSq_e delta
##  0.01489  0.91178 61.23545
```

and plot model diagnostics:

```
# Evaluate model fit diagnostics
fitDiag <- lrgpr(y ~ sex, dcmp, diagnostic=TRUE)
# Plot diagnostics
# Note that plot() only works when lrgpr() is run with diagnostic=TRUE
# Producing the diagnostic statistics can be computationally expensive for
# large sample sizes, so it is set to FALSE by default
par(mfrow=c(2,2), mar=c(4,4.3,2,1));
plot(fitDiag)
```



The user can also specify and perform hypothesis tests on more complex models involving interactions:

```
# Fit a linear model
fit <- lrgpr( y ~ sex + X[,10]*X[,100], dcmp)
```

```
# Composite hypothesis test of additive and interaction terms
# for the two markers
wald(fit, terms=3:5)

##      chisq df p.value
## 0.6289  3  0.8898
```

This functionality can be used to perform custom, exploratory analysis on arbitrarily large datasets.

5.1.2 Low rank linear mixed model: (low rank gaussian process regression)

The full rank linear mixed model can be very computationally expensive for large datasets and a low rank version was proposed to address this issue (Lippert, et al., 2011; Listgarten, et al., 2012). A low rank GSM, or more precisely its spectral decomposition, can be constructed by using fewer SNPs than the sample size. Instead of

capturing the genome-wide similarity, the GSM can be constructed based on the k SNP's that are most correlated with the phenotype. Following Lippert, et al. (2011) and Listgarten, et al. (2012), we construct the GSM using the most significant markers from an uncorrected association test:

```
# Uncorrected single SNP test
pValues <- glmApply( y ~ SNP + sex, features=X)$pValues
```

```
# sort markers based on p-value
ord <- order(pValues, decreasing=FALSE)
```

Now construct the spectral decomposition from the top $k = 100$ SNPs and fit the low rank linear mixed model:

```
k <- 100
dcmp <- svd(scale(set_missing_to_mean(X[,ord[1:k]])))
```

```
# Fit the low rank model
fit <- lrgpr(y ~ sex, dcmp)
```

Learning the optimal rank of the low rank linear mixed model

We arbitrarily selected the top $k = 100$ SNPs above, but following Listgarten, et al. (2012) we can use cross-validation to select k based on the data. Here the cross-validation error is calculated for multiple values of k :

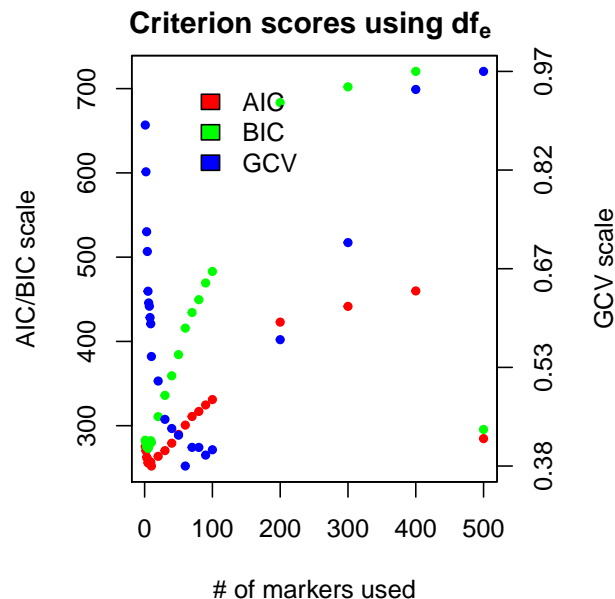
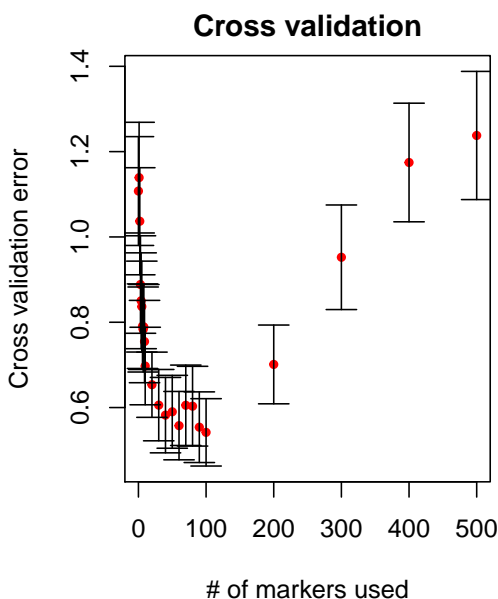
```
# Evaluate cross-validation for multiple rank values
fitcv <- cv.lrgpr( y ~ sex, features=X, order=ord, nfolds=2)
```

Instead of performing computationally expensive cross-validation, we can use model criterion such as Akaike Information Criterion (AIC), Bayesian Information Criterion (BIC) or Generalized Cross-Validation (GCV) to select k based on the degrees of freedom for the linear mixed model described by Hoffman (2013):

```
# Evaluate model criterion
fitcrit <- criterion.lrgpr( y ~ sex, features=X, order=ord )
```

We can plot the cross-validation error and model criterion scores:

```
# Plot the cross-validation and model crition curves
par(mfrow=c(1,2))
plot(fitcv)
plot(fitcrit)
```



and compute to spectral decomposition of the GSM based on the set of SNPs selected by Generalized Cross-validation:

```
k <- fitcrit$best$GCV
dcmp <- svd(scale(set_missing_to_mean(X[,ord[1:k]])))
```

Note that the GCV score is very good proxy for the cross-validation error and most of the difference between these values are driven primarily by the finite sampling in cross-validation.

5.2 Proximal contamination

If the random effect includes the marker to be tested, the correlation between the two can decrease power. More generally, the inclusion of a marker from the same linkage-disequilibrium (LD) block as the marker being tested can decrease power. This is termed *proximal contamination* and Listgarten, et al. (2012) proposed an efficient solution. Intuitively, we would like to use a random effect that excludes any markers from the same LD block as the one being tested. This can be used by dropping a set of markers from the GSM and recomputing the spectral decomposition. However, this can be very computationally expensive and we use the algorithm of Listgarten, et al. (2012) which reuses the same spectral decomposition but which omits the set of proximal contamination markers by modifying the log-loglikelihood calculations of `lrgpr`.

Here, `lrgpr` is evaluated by using markers 2:5 in the random effect and testing marker 1.

```
dcmp <- svd(scale(set_missing_to_mean(X[,2:5])))
fit <- lrgpr(y ~ X[,1], decomp=dcmp)
summary(fit)$coefficients

##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.03198      0.1616 -0.1979  0.8431
## X[, 1]      -0.04524      0.1169 -0.3870  0.6988
```

If the spectral decomposition includes marker 1, then proximal contamination is an issue. We can pass marker 1 into the `W_til` argument to address this.

```
dcmp <- svd(scale(set_missing_to_mean(X[,1:5])))
fit <- lrgpr(y ~ X[,1], decomp=dcmp, W_til=scale(X[,1]))
summary(fit)$coefficients

##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.03198      0.1616 -0.1979  0.8431
## X[, 1]      -0.04524      0.1169 -0.3870  0.6988
```

The results of the two evaluations are exactly the same, with the second having the advantage of avoiding the issue of decreased power due to proximal contamination without having to recompute the spectral decomposition and exclude each marker in turn.

6 Fitting many `lrgpr` regressions with `lrgprApply`

The functionality of `lrgprApply` is the focus of the `lrgpr` package. Once the spectral decomposition has been computed, we can perform a genome-wide regression analysis with `lrgprApply` which is analogous to `glmApply`:

```
# Linear mixed model analysis
pValuesLMM <- lrgprApply( y ~ SNP + sex, features=X, decomp=dcmp)
```

Moreover, `lrgprApply` can be used to fit a regression model with interactions and the terms can be set to perform the appropriate hypothesis test:

```
# Test SNP x sex interaction
pv1 <- lrgprApply( y ~ SNP*sex, features=X, decomp=dcmp, terms=4)
```

```
# Test SNP x SNP interaction
pv2 <- lrgprApply( y ~ sex + SNP*X[,1], features=X, decomp=dcmp, terms=3:5)
```

The user has a lot of flexibility to design a custom analysis to address the specific question of his or her study.

6.1 Re-estimating variance components (δ)

By default, `lrgprApply` estimates the variance components under the null model and then re-uses these parameter estimates for each marker. This follows Lippert, et al. (2011) and greatly decreases computational time.

It is possible to re-estimate the variance components for each marker with a substantial increase in computational time:

```
# Test SNP x SNP interaction
pv3 <- lrgprApply( y ~ sex + SNP, features=X, decomp=dcmp,
  reEstimateDelta=TRUE)
```

6.2 Proximal contamination

We can also drop proximal markers in `lrgprApply` using either genetic or physical distance to define the genomic window. Using the map file from plink:

```
MAP <- read.table( map_file )
```

we can call perform a genome-wide analysis:

```
pValuesLMMprox <- lrgprApply( y ~ sex + SNP, features=X, decomp=dcmp,
  map=MAP[,c(1,3)], distance=2, dcmp_features=ord[1:k])
```

where entries in `MAP` correspond to the markers in `X`, `map=MAP[,c(1,3)]` indicates the marker names and their position on the genetic map, `distance` determines the size of the window corresponding to the locations in `MAP[,3]`, and `dcmp_features` indicates the indices of `X` used to construct `dcmp`. Following the standard format for plink map files, the physical location can be used by specifying `map=MAP[,c(1,4)]` and setting `distance` appropriately.

6.2.1 Interaction model

When evaluating an interaction model with the formula `y ~ sex + SNP*X[,1]`, the syntax above addresses the proximal contamination due to `SNP`, but not due to `X[,1]`. We can include additional markers in the proximal contamination set by also specifying `W_til`:

```
# Test SNP x SNP interaction
pv2 <- lrgprApply( y ~ sex + SNP*X[,1], features=X, decomp=dcmp,
  map=MAP[,c(1,3)], distance=2, dcmp_features=ord[1:k], W_til=X[,1])
```

7 Implementation

The **lrgpr** package provides an R interface to high performance computational statistics code written in C/C++ that builds off the GNU Scientific Library (GSL). Linear algebra operations are performed by BLAS and LAPACK and the performance is highly dependent on the library installed on the local machine. Data is passed from R to the C++ backend using **Rcpp** and **RcppGSL**. Parallelization is performed by OpenMP. Out-of-core processing is facilitated by **bigmemory**, which stores the full dataset on the hard drive instead of in memory and uses `mmap` from the GNU C Library for random data access.

References

- Hoffman, G.~E. (2013). Correcting for Population Structure and Kinship Using the Linear Mixed Model: Theory and Extensions. *PLoS ONE* **8**, e75707.
- Kang, H.~M., Sul, J.~H., Service, S.~K., Zaitlen, N.~A., Kong, S.-y., Freimer, N.~B., Sabatti, C., and Eskin, E. (2010). Variance component model to account for sample structure in genome-wide association studies. *Nature Genetics* **42**, 348–54.
- Lippert, C., Listgarten, J., Liu, Y., Kadie, C.~M., Davidson, R.~I., and Heckerman, D. (2011). FaST linear mixed models for genome-wide association studies. *Nature Methods* **8**, 833–5.
- Listgarten, J., Lippert, C., Kadie, C.~M., Davidson, R.~I., Eskin, E., and Heckerman, D. (2012). Improved linear mixed models for genome-wide association studies. *Nature Methods* **9**, 525–6.
- Price, A.~L., Patterson, N.~J., Plenge, R.~M., Weinblatt, M.~E., Shadick, N.~A., and Reich, D. (2006). Principal components analysis corrects for stratification in genome-wide association studies. *Nature Genetics* **38**, 904–9.
- Yang, J., Lee, S.~H., Goddard, M.~E., and Visscher, P.~M. (2011). GCTA: a tool for genome-wide complex trait analysis. *American Journal of Human Genetics* **88**, 76–82.
- Zhou, X. and Stephens, M. (2012). Genome-wide efficient mixed-model analysis for association studies. *Nature Genetics* **44**, 821–4.