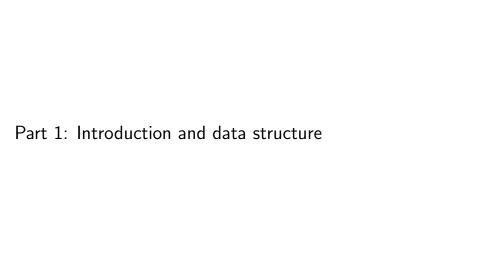


# The 'synbreed' R package

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November 8, 2012



# Summary - synbreed R package

- Add-on for the open source environment for statistical computing
- Three example data sets available in synbreedData package
- Hosted on CRAN: http: //cran.r-project.org/web/packages/synbreed/index.html
   R> install.packages("synbreed")
- Latest development version on R-Forge:
   https://r-forge.r-project.org/R/?group\_id=710
   R> install.packages("synbreed",repos="http://r-forge.r-project.")
- Recent R version required R  $\geq 2.15.1$
- All operating systems
- Once installed, load the package using
   R> library(synbreed)





#### Available documentation

- Publication in Bioinformatics (Wimmer et al. 2012)
- Package vignette

```
R> vignette("IntroSyn")
```

- Website on R-Forge http://synbreed.r-forge.r-project.org/
- Manual and help sites:

```
R> help(package="synbreed")
```

Code demonstrations

```
R> demo(package="synbreed")
```





## Citation Please cite the synbreed package in your work, whenever you use it

# R> citation(package="synbreed")

To cite package 'symbreed' in publications use:

Recommended citation

Wimmer V, Albrecht T, Auinger HJ and Schoen CC (2012) synbreed: a framework for the analysis of genomic prediction data using R. Bioinformatics, 28: 2086-2087

A BibTeX entry for LaTeX users is

@Article{. title = {synbreed: a framework for the analysis of genomic pred.

author = {Valentin Wimmer and Theresa Albrecht and Hans-Juergen journal = {Bioinformatics},

 $year = \{2012\},\$ 

 $volume = \{28\},$ 

number =  $\{15\}$ .  $names = \{2086 - 2087\}$ 

# Design objectives

- User-friendly interface to analyze genomic prediction data
- Analysis framework defined by a single, unified data object
- One (open-source) software package
- Flexible implementation (plant and animal breeding)
- Gateway to other software and R packages
- Teaching tool





# Analysis pipeline

- Data management and storage
- Oata processing: recoding, marker selection and imputing
- Pedigree and marker-based coefficients of relatedness
- Fit BLUP and Bayesian models by a unified interface
- Model validation using cross-validation
- Open Prediction of unphenotyped individuals
- Data visualization





#### Overview

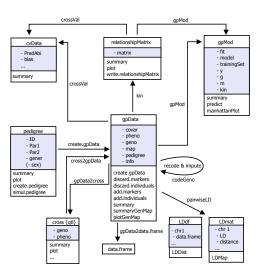


Figure: Classes, functions and methods of the synbreed R package

#### Data structure

- All data for genomic selection are combined in a single object
- Hence, easy data sharing, summary statistics, reduced storage requirements

#### class gpData

- pheno: array (3 dimensions) with phenotypes
- geno: matrix with genotypes (SNP markers)
- map : data.frame with marker map (chr + position)
- pedigree : class "pedigree"
- covar : data.frame with additional covariate information

R> gp <- create.gpData(pheno,geno,map,pedigree,covar,map.unit: "")



# Pedigree

ID	Par1	Par2	gener	sex
Α	-	-	0	
В	-	-	0	
C	Α	В	1	
D	Α	C	2	
Е	D	В	3	

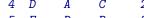
- first generation = 0
- Create pedigree object

$$R > par1 < -c(0,0,"A","A","D")$$

#### ID Par1 Par2 gener

1	. <i>A</i>	0	0	0
2	? B	0	0	0
-		4	D	4







#### Read-in of own data

- Simulated data from XII QTL-MAS Workshop 2008, Uppsala
- Available from http://www.computationalgenetics.se/ QTLMASO8/QTLMAS/DATA.html

#### QTLMAS data

- 50 simulated QTLs (explained variance 0 5 %)
- 5865 individuals (2778 males, 3087 females)
- 6000 markers on 6 chromosomes (each of length 100cM)





### Create object of class gpData - 1

```
R> # Create object of class 'pedigree'
R> ped <- with(dat,create.pedigree(ID=id,Par1=sire,Par2=dam,gener=geR) # Phenotypic data
R> pheno <- data.frame(trait=dat$Phenotype,row.names=dat$id)
R> # covar = tbv
R> covar <- data.frame(tbv=dat$GeneticValue,row.names=dat$id)
R> # genotypic data
```

R> geno <- read.table("genotype\_cor.txt",header=FALSE,stringsAsFactor

R> dat <- read.table("TrueEBV.txt", header=TRUE, stringsAsFactors=FAL

R> # Read file TrueEBV.txt with pedigree, trait, and tbv

# Create object of class gpData - 2

```
R> # gametes to genotypes
R> geno2 <- matrix(data=NA,nrow=nrow(geno),ncol=(ncol(geno)-1)/2)
R> for (j in 1:ncol(geno2)){
     # combine phased data to a genotype
+ geno2[,j] <- paste(as.character(geno[,2*j]),as.character(geno[</pre>
R> # create map
R> # 6 chromosomes with 1000 markers
R> # dist between adjacent markers = 0.1cM
R > chr <- rep(1:6, each=1000)
R > pos \leftarrow rep(seq(from=0, to=99.9, by=.1), times=6)
R> map <- data.frame(chr=chr,pos=pos)</pre>
R> # create gpData object
R> qt1MASdata <- create.gpData(pheno=pheno,geno=geno2,map=map,pedig
R> # save data as object of class gpData in Rdata-format
R> save("qtlMASdata",file="qtlMASdata.Rdata")
R> # for loading data, function load() and ls() might be useful
```

### Example data sets

#### Maize data

- Simulated maize breeding program using DH technology
- 1250 DH lines phenotyped for one quantitative trait and 1117 SNPs

### Mice data (Valdar et al. 2006)

- Heterogeneous stock mice population, publicly available from http://gscan.well.ox.ac.uk
- 2527 individuals with 2 traits (weight [g] at 6 weeks age and growth slope between 6 and 10 weeks age [g/day])
- 1940 individuals genotyped with 12545 SNP markers

#### Cattle data

• 50 individuals genotyped by 7250 SNP markers





#### The simulated maize data

#### **Parameters**

- 10 chromosomes of length 160 cM
- 500 segregating biallelic QTL with equal, additive effects
- Doubled-haploid (DH) lines
- 1250 individuals with genotypes (1117 SNPs) and phenotypes
- One quantitative trait evaluated in a testcross in 3 environments
- $h^2 = 0.46$
- Population structure: 25 biparental families of size 50





# The maize data

R> library(synbreed)

```
R> data(maize)
                                   No. of markers 1117
R> summary(maize)
                                   genotypes 0 1
                                   frequencies 0.339995 0.660005
                                   NA's 0.000 %
object of class 'gpData'
covar
                                  map
 No. of individuals 1610
                                   No. of mapped markers 1117
                                   No. of chromosomes
                                                           10
         phenotyped 1250
          genotyped 1250
pheno
                                   markers per chromosome
 No. of traits:
                                       96 99 122 85 106 154 130 12
     Trait
 Min. :120.7
 1st Qu.:142.8
                                  pedigree
 Median :148.9
                                  Number of
 Mean :148.9
                                   individuals 1610
 3rd Qu.:154.9
                                                219
                                   Par 1
 Max.
       :181.8
                                                221
                                   Par 2
```

15

generations

geno

### Extract parts of the data

An object of class gpData is a list, see

```
R> str(maize)
```

Look a the phenotypic data

```
R> head(maize$pheno[,1,])
11360 11361 11362 11363 11364 11365
148.30 145.35 129.44 158.32 150.27 148.75
```

Look a the genotypic data (individuals 10 to 13, markers 20 to 25)

```
R> maize$geno[10:13,20:25]
```

```
    M20
    M21
    M22
    M23
    M24
    M25

    11369
    1
    1
    1
    0
    1
    1

    11370
    0
    1
    1
    0
    1
    1

    11371
    0
    1
    1
    0
    1
    1

    11372
    1
    1
    1
    0
    1
    1
```





#### The covar element

Generated within create.gpData, a data.frame

R> head(maize\$covar,n=4)

```
id phenotyped genotyped DH tbv family
1 10910
            FALSE
                      FALSE.
                                NA
                                      NA
2 10918
            FALSE
                      FALSE.
                                NA
                                      NA
3 10921
            FALSE
                     FALSE 0
                                NA
                                      NA
4 10924
            FALSE
                      FALSE 0
                                NA
                                      NA
```

- Column id: All names of individuals that either appear in geno, pheno or pedigree
- Column genotyped: Has the individual observations in geno?
- Column phenotyped: Has the individual observations in pheno?

Example: Extract all phenotyped individuals

R> maize\$covar\$id[maize\$covar\$phenotyped]





### Remove and add markers/individuals

- discard.individuals
- discard.markers
- add.individuals
- add.markers

Example: Remove all markers from chromosome 6 to 10

R> maizeChr1to5 <- discard.markers(maize,rownames(maize\$map)[maize\$n

```
R> summary(maizeChr1to5$map)
```

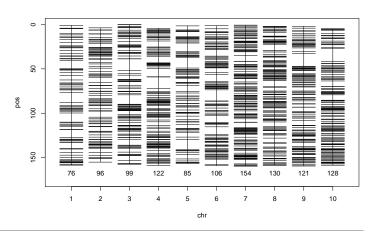
```
chr pos
Min. :1.000 Min. : 0.05
1st Qu.:2.000 1st Qu.: 35.35
Median :3.000 Median : 86.27
Mean :3.092 Mean : 80.74
3rd Qu.:4.000 3rd Qu.:121.44
Max. :5.000 Max. :158.70
```





# Visualization of marker map

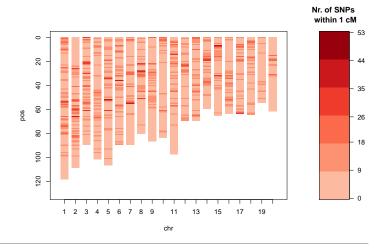
#### R> plotGenMap(maize)





# Visualization of marker map

R> plotGenMap(mice,dense=TRUE,nMarker = FALSE, bw=1)





# Summary of marker map

#### R> summaryGenMap(maize)

```
noM
            length avDist maxDist minDist
                             11.08
        76
            157.52 2.100267
                                     0.10
        96
            151.38 1.593474
                             6.81
                                     0.03
3
                                     0.02
        99
           157.44 1.606531 13.11
       122
            154.34 1.275537 13.11
                                     0.04
5
        85
           155.13 1.846786 11.67
                                     0.01
6
       106
           157.70 1.501905 12.46
                                     0.02
       154
           158.98 1.039085 6.48
                                     0.02
8
       130
           156.62 1.214109
                             7.03
                                     0.05
9
       121
            157.27 1.310583
                             14.21
                                     0.06
10
       128
            153.92 1.211969
                             15.19
                                     0.08
   10 1117 1560.30 1.410027
                             15.19
                                     0.01
```





# Problems 1 - 1 (Corn borer example)

Please read: http://www.rise.gs.tum.de/fileadmin/w00bjb/www/Risk\_book\_Chapters/SchoenWimmer\_revised.pdf

Table: Pedigree, phenotypic values, and marker genotypes for eight simulated maize individuals

Cycle	Individual	Pedigree	Tunnel length	SNP			
-		_	[cm]	1	2	3	4
				(0)*	(1)	(-4)	(4)
1	l1	$P1 \times P2$	13	2	2	0	1
1	12	$P3 \times P4$	17	0	0	0	1
1	l3	-	1	0	1	2	0
2	14	$11 \times 12$	17	1	1	0	2
2	<b>I</b> 5	$11 \times 12$	11	1	1	0	1
2	16	$12 \times 13$	6	0	1	1	0
2	17	$11 \times 12$	-	1	1	0	1
2	18	$I1 \times I2$	-	1	1	0	0

# Problems 1 - 2 (Corn borer example)

- Transfer the pedigree structure of the 8 individuals into an object of class pedigree and plot it.
- Combine all data of the corn borer example in an object of class gpData called cbData. Include pedigree, phenotypes and genotypes (SNPs 1 to 4) and add the names of Table 1 for markers and individuals for all objects.
- Use the summary method for this object. Are all details correct?
- Compute a new object called cbData2 excluding all individuals without phenotypes.
- Use this data to compute a single marker regression for each SNP. Which markers are significant at the 5% error rate.
- Set up a multiple marker regression model using (1) all SNPs and (2) only SNPs 3 and 4. Compare the results and discusss which model you would choose?

Part 2: Processing of marker data

### Processing of marker data

#### Function codeGeno

- Preselection of markers
- Recode marker genotypes
- Impute missing values

```
R> maizeC <- codeGeno(maize,maf=0.05,nmiss=0.1,
+ verbose=TRUE)

step 1 : 0 marker(s) removed with > 10 % missing values
step 2 : Recoding alleles
step 2.1: No duplicated markers discarded
step 5 : 122 marker(s) removed with maf < 0.05
step 6 : No duplicated markers discarded
End : 995 marker(s) remain after the check
Compute pairwise LD measured as r² on chr 1
R> maizeLD <- pairwiseLD(maizeC,chr=1,type="data.frame")</pre>
```



## Algorithm of codeGeno

- Discard markers with fraction > nmiss of missing values
- Recode alleles as number of the minor alleles, i.e. 0, 1 and 2
- Replace missing values by replace.value or impute missing values according to impute.type
- Recode of alleles after imputation, if necessary due to changes in allele frequencies by imputed alleles
- $\odot$  Discard markers with a minor allele frequency of  $\leq$  maf
- O Discard duplicated markers if keep.identical=FALSE
- Restore original data format (gpData, matrix or data.frame)



## Imputing algorithms

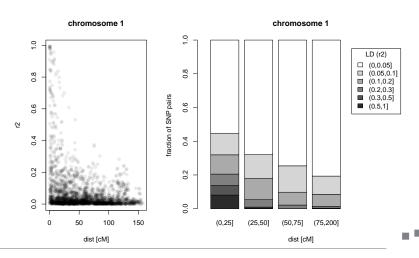
Gaps in the marker matrix can be filled according to

- Beagle (Browning and Browning 2009) (impute.type = "beagle")
  - Imputation within families (only for homozygous inbred lines according to Albrecht et al. (2011), impute.type = "family")
  - Beagle after family (impute.type = "beagleAfterFamily")
  - Random imputation according to the marginal allele distribution (impute.type = "random")
  - A fixed value (impute.type = "fix")



# Visualization of LD decay

R> plot(maizeLD); plot(maizeLD,type="bars")





#### Problems 2 - 1

- Transfer your own data in to class gpData or use the Arabidopsis MAGIC lines population (Kover et al. 2009).
- How many individuals are genotyped/phenotyped?
- Make a new object an retain only those individuals which are phenotyped and genotyped.
- From this object, remove all markers without a map position.
- Make a visualization of the marker map. What is the largest gap between two markers?
- Run and retrace the examples of the function codeGeno.
- What are the observed alleles in your genotypic data? Recode your data into the number of copies of the minor allele. In the same step, remove all marker with more than 10% missing values or a MAF < 0.05. How many markers were removed according to these criteria? Use the argument print.report=TRUE in codeGeno and check the result.</p>

#### Problems 2 - 2

- If there are missing values, try to impute them using Beagle, if not possible, replace them according to the marginal allele distribution.
- Make a histogram of the MAF. What is the median and mean of the MAF?
- ② Compute the LD as  $r^2$  using the gateway to PLINK (only for the first chromosome).
- **•** What is the minimum/mean/maximum LD between two markers? What is the proportion of markers with  $r^2 > 0.20$ ?
- Visualize the LD decay using a scatterplot and stacked histograms?
- Try to add the nonlinear curve according to Hill and Weir (1988) to the scatterplot.
- Make a LD heatmap for the first chromosome.

Part 3: Prediction and validation

#### Prediction models

#### Pedigree-based BLUP PBLUP

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{a} + \mathbf{e}$$
  
 $\mathbf{a} \sim \mathsf{N}(\mathbf{0}, \mathbf{A}\sigma_a^2)$ 

#### Marker-based BLUP GBLUP

$$\begin{array}{lcl} \mathbf{y} & = & \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{u} + \mathbf{e} \\ \mathbf{u} & \sim & \mathsf{N}(\mathbf{0}, \mathbf{U}\sigma_u^2) \\ \mathbf{U} & = & \dfrac{(\mathbf{W} - \mathbf{P})(\mathbf{W} - \mathbf{P})^\top}{2\sum_{j=1}^p p_j(1-p_j)} \end{array}$$

with

**y** Vector of phenotypic records

W Marker matrix

 ${f P}$  Matrix with the allele frequencies  $p_j$ 

 $\mathbf{e} \sim \mathsf{N}(\mathbf{0}, \mathsf{I}\sigma^2)$  vector of residuals





#### Estimation of relatedness

- Pedigree based (expected) and realized kinship coefficients: function kin
  - Additive numerator relationship matrix A (default)
    R> kin(gpData, ret="add")
  - ► Dominance relationship matrix **D** 
    - R> kin(gpData,ret="dom")
  - Kinship matrix  $\mathbf{K} = \frac{1}{2}\mathbf{A}$ 
    - R> kin(gpData,ret="kin")
  - Gametic relationship matrix (dimension  $2n \times 2n$ )

Kinship for the 1250 DH lines

R> A <- kin(maizeC,ret="kin",DH=maize\$covar\$DH)</pre>





## Special case

The phenotypes in the maize data origin from a testcross of DH lines, hence (Albrecht *et al.* 2011)

- The additive relationship matrix must be replaced by the kinship
- The variance of the marker genotypes is

$$4\sum_{j=1}^{p}p_{j}(1-p_{j})$$

Thus

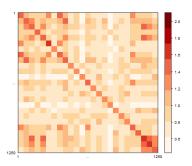
$$\mathbf{U} = \frac{(\mathbf{W} - \mathbf{P})(\mathbf{W} - \mathbf{P})}{4\sum_{j=1}^{p} p_j (1 - p_j)}$$



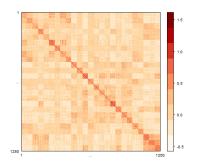
### Example

R> U <- kin(maizeC,ret="realized")/2</pre>

R> plot(A[maize\$covar\$genotyped,maize\$covar\$genotyped]); plot(U)



(a) Pedigree-based relationship



(b) Marker-based relationship



### Equation

$$y = X\beta + Zu + e$$

where

- y is a vectors of phenotypes
- **X** is a design matrix allocating phenotypes to fixed effects
- $\beta$  is a vector of fixed effects
- **Z** is a design matrix allocating phenotypes to random effects
- **u** is a vector of random effects, with  $\mathbf{u} \sim N(0, \mathbf{G}\sigma^2)$
- ${f e}$  is a vector of residuals, with  ${f e} \sim N(0,{f R}\sigma^2)$





$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{u} + \mathbf{e}$$

$$\mathbf{E} \begin{bmatrix} \mathbf{u} \\ \mathbf{e} \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}$$

$$\mathbf{Var} \begin{bmatrix} \mathbf{u} \\ \mathbf{e} \end{bmatrix} = \begin{bmatrix} \mathbf{G} & 0 \\ 0 & \mathbf{R} \end{bmatrix} \sigma^2$$

$$\mathbf{E}(\mathbf{y}) = \mathbf{X}\boldsymbol{\beta}$$

$$\mathbf{Var}(\mathbf{y}) = \mathbf{V} = (\mathbf{Z}\mathbf{G}\mathbf{Z}' + \mathbf{R})\sigma^2$$



$$\left[\begin{array}{c} \boldsymbol{\hat{\beta}} \\ \boldsymbol{\hat{u}} \end{array}\right] = \left[\begin{array}{ccc} \mathbf{X}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{X}'\mathbf{R}^{-1}\mathbf{Z} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{Z}'\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1} \end{array}\right]^{-} \left[\begin{array}{c} \mathbf{X}'\mathbf{R}^{-1}\mathbf{y} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{y} \end{array}\right]$$

- $\hat{\mathbf{u}} = \mathbf{G}\mathbf{Z}'\mathbf{V}^{-1}(\mathbf{y} \mathbf{X}\hat{\boldsymbol{\beta}})$ where  $\hat{\boldsymbol{\beta}}$  is a generalized least square solution  $\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-}\mathbf{X}'\mathbf{V}^{-1}\mathbf{y}$
- Difference to the least square estimate of a LM  $(\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-}\mathbf{X}'\mathbf{y})$  is the decomposition of  $\mathbf{V}$
- V has to be estimated



$$\left[\begin{array}{c} \boldsymbol{\hat{\beta}} \\ \boldsymbol{\hat{u}} \end{array}\right] = \left[\begin{array}{ccc} \mathbf{X}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{X}'\mathbf{R}^{-1}\mathbf{Z} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{Z}'\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1} \end{array}\right]^{-} \left[\begin{array}{c} \mathbf{X}'\mathbf{R}^{-1}\mathbf{y} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{y} \end{array}\right]$$

- $\hat{\mathbf{u}} = \mathbf{G}\mathbf{Z}'\mathbf{V}^{-1}(\mathbf{y} \mathbf{X}\hat{\boldsymbol{\beta}})$ where  $\hat{\boldsymbol{\beta}}$  is a generalized least square solution  $\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-}\mathbf{X}'\mathbf{V}^{-1}\mathbf{y}$
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$$\left[\begin{array}{c} \boldsymbol{\hat{\beta}} \\ \boldsymbol{\hat{u}} \end{array}\right] = \left[\begin{array}{ccc} \mathbf{X}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{X}'\mathbf{R}^{-1}\mathbf{Z} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{Z}'\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1} \end{array}\right]^{-} \left[\begin{array}{c} \mathbf{X}'\mathbf{R}^{-1}\mathbf{y} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{y} \end{array}\right]$$

- $\hat{\mathbf{u}} = \mathbf{G}\mathbf{Z}'\mathbf{V}^{-1}(\mathbf{y} \mathbf{X}\hat{\boldsymbol{\beta}})$ where  $\hat{\boldsymbol{\beta}}$  is a generalized least square solution  $\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-}\mathbf{X}'\mathbf{V}^{-1}\mathbf{y}$
- Difference to the least square estimate of a LM  $(\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-}\mathbf{X}'\mathbf{y})$  is the decomposition of  $\mathbf{V}$
- V has to be estimated



$$\left[\begin{array}{c} \boldsymbol{\hat{\beta}} \\ \boldsymbol{\hat{u}} \end{array}\right] = \left[\begin{array}{ccc} \mathbf{X}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{X}'\mathbf{R}^{-1}\mathbf{Z} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{X} & \mathbf{Z}'\mathbf{R}^{-1}\mathbf{Z} + \mathbf{G}^{-1} \end{array}\right]^{-} \left[\begin{array}{c} \mathbf{X}'\mathbf{R}^{-1}\mathbf{y} \\ \mathbf{Z}'\mathbf{R}^{-1}\mathbf{y} \end{array}\right]$$

- $\hat{\mathbf{u}} = \mathbf{G}\mathbf{Z}'\mathbf{V}^{-1}(\mathbf{y} \mathbf{X}\hat{\boldsymbol{\beta}})$ where  $\hat{\boldsymbol{\beta}}$  is a generalized least square solution  $\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-}\mathbf{X}'\mathbf{V}^{-1}\mathbf{y}$
- Difference to the least square estimate of a LM  $(\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-}\mathbf{X}'\mathbf{y})$  is the decomposition of  $\mathbf{V}$
- V has to be estimated



# Example from Henderson (1977)

У	time	animal
132	1	1
147	2	2
156	1	3
172	2	4

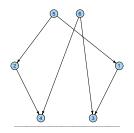


Figure: Pedigree





#### Equation

$$y_{ij} = \beta_i + u_j + e_{ij}$$
  
 $\mathbf{y} = \mathbf{X}\beta + \mathbf{Z}\mathbf{u} + \mathbf{e}$ 

observations

with 
$$\mathbf{y}' = (y_{11}, y_{22}, y_{13}, y_{24})$$
  
 $\beta' = (\beta_1, \beta_2)$ 

$$eta' = (eta_1, eta_2)$$
 time effects (fix)

$$\mathbf{u}' = (u_1, u_2, u_3, u_4)$$
 additive genetic merit (random) with  $\mathbf{u} \sim \mathsf{N}(0, \mathbf{A}\sigma_u^2)$ 

$$\mathbf{u}' = (u_1, u_2, u_3, u_4)$$
 additive genetic merit (random) with  $\mathbf{e}' = (e_{11}, e_{22}, e_{13}, e_{24})$  residuals (random) with  $\mathbf{e} \sim \mathsf{N}(0, \mathbf{I}\sigma^2)$ 

X, Zdesign matrices

$$\begin{bmatrix} 132 \\ 147 \\ 156 \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \\ 1 & 0 \end{bmatrix}$$

$$\begin{bmatrix} 132 \\ 147 \\ 156 \\ 172 \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \\ 1 & 0 \\ 0 & 1 \end{bmatrix} \cdot \begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix} + \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} \cdot \begin{bmatrix} u_1 \\ u_2 \\ u_3 \\ u_4 \end{bmatrix} + \begin{bmatrix} e_{11} \\ e_{22} \\ e_{13} \\ e_{24} \end{bmatrix}$$



#### Equation

$$y_{ij} = \beta_i + u_j + e_{ij}$$
  
 $\mathbf{y} = \mathbf{X}\beta + \mathbf{Z}\mathbf{u} + \mathbf{e}$ 

$$\mathbf{y}' = (y_{11}, y_{22}, y_{13}, y_{24})$$
 observations  $\beta' = (\beta_1, \beta_2)$  time effects (fix)

 $\mathbf{u}' = (u_1, u_2, u_3, u_4)$ 

residuals (random) with  $\mathbf{e} \sim \mathsf{N}(0, \mathbf{I}\sigma^2)$  $\mathbf{e}' = (e_{11}, e_{22}, e_{13}, e_{24})$ 

X, Z

additive genetic merit (random) with 
$$\mathbf{u} \sim N(0, \mathbf{A}\sigma_u^2)$$
  
residuals (random) with  $\mathbf{e} \sim N(0, \mathbf{I}\sigma^2)$ 

design matrices

$$\begin{bmatrix} 0 \\ 1 \\ 0 \\ 1 \end{bmatrix}$$

 $\begin{bmatrix} 132 \\ 147 \\ 156 \\ 172 \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \\ 1 & 0 \\ 0 & 1 \end{bmatrix} \cdot \begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix} + \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} \cdot \begin{bmatrix} u_1 \\ u_2 \\ u_3 \\ u_4 \end{bmatrix} + \begin{bmatrix} e_{11} \\ e_{22} \\ e_{13} \\ e_{24} \end{bmatrix}$ 



$$\begin{split} \mathsf{E} \left[ \begin{array}{c} \mathbf{u} \\ \mathbf{e} \end{array} \right] &= \left[ \begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array} \right] \\ \mathsf{Var} \left[ \begin{array}{c} \mathbf{u} \\ \mathbf{e} \end{array} \right] &= \left[ \begin{array}{cc} \mathbf{G} & \mathbf{0} \\ \mathbf{0} & \mathbf{R} \end{array} \right] \sigma^2 = \left[ \begin{array}{cc} \mathbf{A} \frac{\sigma_u^2}{\sigma^2} & \mathbf{0} \\ \mathbf{0} & \mathbf{I} \end{array} \right] \sigma^2 \end{split}$$

#### Assumption

$$h^{2} = \frac{\sigma_{u}^{2}}{\sigma_{u}^{2} + \sigma^{2}} = 0.25$$

$$\Rightarrow e.g. \ \sigma_{u}^{2} = 0.25 \text{ and } \sigma^{2} = 0.75$$

$$\Rightarrow \mathbf{G}^{-1} = \mathbf{A}^{-1} \frac{\sigma^{2}}{\sigma_{u}^{2}} = 3\mathbf{A}^{-1}$$

Numerator relationship matrix

$$\mathbf{A} = \left[ \begin{array}{ccccc} 1 & 0.25 & 0.5 & 0.125 \\ & 1 & 0.125 & 0.5 \\ & & 1 & 0.3125 \\ & & & 1 \end{array} \right]; \ 3\mathbf{A}^{-1} = \left[ \begin{array}{cccccc} 4.325 & -1.175 & -2.25 & 0.75 \\ -1.175 & 4.325 & 0.75 & -2.25 \\ -2.250 & 0.750 & 4.50 & -1.50 \\ 0.750 & -2.250 & -1.50 & 4.50 \end{array} \right]$$

$$\begin{split} & \mathsf{E}\left[\begin{array}{c} \textbf{u} \\ \textbf{e} \end{array}\right] &= \left[\begin{array}{c} 0 \\ 0 \end{array}\right] \\ & \mathsf{Var}\left[\begin{array}{c} \textbf{u} \\ \textbf{e} \end{array}\right] &= \left[\begin{array}{cc} \textbf{G} & 0 \\ 0 & \textbf{R} \end{array}\right] \sigma^2 = \left[\begin{array}{cc} \textbf{A} \frac{\sigma_u^2}{\sigma^2} & 0 \\ 0 & \textbf{I} \end{array}\right] \sigma^2 \end{split}$$

Assumption

$$h^{2} = \frac{\sigma_{u}^{2}}{\sigma_{u}^{2} + \sigma^{2}} = 0.25$$

$$\Rightarrow e.g. \ \sigma_{u}^{2} = 0.25 \text{ and } \sigma^{2} = 0.75$$

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Numerator relationship matrix

$$\mathbf{A} = \begin{bmatrix} 1 & 0.25 & 0.5 & 0.125 \\ & 1 & 0.125 & 0.5 \\ & & 1 & 0.3125 \\ & & & 1 \end{bmatrix}; \ 3\mathbf{A}^{-1} = \begin{bmatrix} 4.325 & -1.175 & -2.25 & 0.75 \\ -1.175 & 4.325 & 0.75 & -2.25 \\ -2.250 & 0.750 & 4.50 & -1.50 \\ 0.750 & -2.250 & -1.50 & 4.50 \end{bmatrix}$$

### Assumption

$$h^{2} = \frac{\sigma_{u}^{2}}{\sigma_{u}^{2} + \sigma^{2}} = 0.25$$

$$\Rightarrow e.g. \ \sigma_{u}^{2} = 0.25 \text{ and } \sigma^{2} = 0.75$$

$$\Rightarrow \mathbf{G}^{-1} = \mathbf{A}^{-1} \frac{\sigma^{2}}{\sigma_{u}^{2}} = 3\mathbf{A}^{-1}$$

Numerator relationship matrix

$$\mathbf{A} = \left[ \begin{array}{cccc} 1 & 0.25 & 0.5 & 0.125 \\ & 1 & 0.125 & 0.5 \\ & & 1 & 0.3125 \\ & & & 1 \end{array} \right]; \ 3\mathbf{A}^{-1} = \left[ \begin{array}{ccccc} 4.325 & -1.175 & -2.25 & 0.75 \\ -1.175 & 4.325 & 0.75 & -2.25 \\ -2.250 & 0.750 & 4.50 & -1.50 \\ 0.750 & -2.250 & -1.50 & 4.50 \end{array} \right]$$

$$\left[ \begin{array}{c} \boldsymbol{\hat{\beta}} \\ \boldsymbol{\hat{u}} \end{array} \right] = \left[ \begin{array}{cc} \mathbf{X}'\mathbf{X} & \mathbf{X}'\mathbf{Z} \\ \mathbf{Z}'\mathbf{X} & \mathbf{Z}'\mathbf{Z} + \mathbf{A}^{-1}\frac{\sigma^2}{\sigma_u^2} \end{array} \right]^{-} \left[ \begin{array}{c} \mathbf{X}'\mathbf{y} \\ \mathbf{Z}'\mathbf{y} \end{array} \right]$$

Results

$$\hat{\beta} = \begin{bmatrix} 143.89 \\ 159.40 \end{bmatrix}$$
 and  $\hat{\mathbf{u}} = \begin{bmatrix} -2.07 \\ -2.12 \\ 2.28 \\ 2.33 \end{bmatrix}$ 





$$\left[ \begin{array}{c} \boldsymbol{\hat{\beta}} \\ \boldsymbol{\hat{u}} \end{array} \right] = \left[ \begin{array}{cc} \mathbf{X}'\mathbf{X} & \mathbf{X}'\mathbf{Z} \\ \mathbf{Z}'\mathbf{X} & \mathbf{Z}'\mathbf{Z} + \mathbf{A}^{-1}\frac{\sigma^2}{\sigma_u^2} \end{array} \right]^{-} \left[ \begin{array}{c} \mathbf{X}'\mathbf{y} \\ \mathbf{Z}'\mathbf{y} \end{array} \right]$$

$$\begin{bmatrix} \hat{\boldsymbol{\beta}} \\ \hat{\mathbf{u}} \end{bmatrix} = \begin{bmatrix} 2 & 0 & 1 & 0 & 1 & 0 \\ 0 & 2 & 0 & 1 & 0 & 1 \\ 1 & 0 & 5.325 & -1.175 & -2.250 & 0.750 \\ 0 & 1 & -1.175 & 5.325 & 0.750 & -2.250 \\ 1 & 0 & -2.250 & 0.750 & 5.500 & -1.500 \\ 0 & 1 & 0.750 & -2.250 & -1.500 & 5.500 \end{bmatrix}^{-} \begin{bmatrix} 288 \\ 319 \\ 132 \\ 147 \\ 156 \\ 172 \end{bmatrix}^{-}$$

Results

$$\hat{\beta} = \begin{bmatrix} 143.89 \\ 159.40 \end{bmatrix}$$
 and  $\hat{\mathbf{u}} = \begin{bmatrix} -2.07 \\ -2.12 \\ 2.28 \\ 2.33 \end{bmatrix}$ 





#### Function MME

```
R > ped < -create.pedigree(ID=c(6,5,1,2,3,4),Par1=c(0,0,5,5,1,6),Par1=c(0,0,5,5,1,6),Par1=c(0,0,5,5,1,6)
R> gp <- create.gpData(pheno=dat,pedigree=ped)</pre>
R> A <- kin(gp,ret="add")</pre>
R > (X \leftarrow matrix(c(1,0,1,0,0,1,0,1),ncol=2))
     [,1] [,2]
[1,] 1 0
[2,] 0 1
[3,] 1
[4,] 0
R > (Z \leftarrow diag(6)[-c(1,2),])
     [,1] [,2] [,3] [,4] [,5] [,6]
[1,] 0 0 1 0
[2,] 0 0 0 1
[3,] 0 0 0 0 1
[4,] 0
```

R > dat < -data.frame(y=c(132,147,156,172),time=c(1,2,1,2),animal=c(1,2,1,2),anima



#### Function MME

```
R > (AT <- solve(A))
5 1.6666667 0.0 -0.6666667 -0.6666667 0 0
  0.0000000 \ 2.0 \ 0.5000000 \ 0.5000000 \ -1 \ -1
1 -0.6666667 0.5 1.8333333 0.0000000 -1 0
2 -0.6666667 0.5 0.0000000 1.8333333 0 -1
3 0.0000000 -1.0 -1.0000000 0.0000000 2 0
4 0.0000000 -1.0 0.0000000 -1.0000000 0 2
R > RI < - diag(4)
R > res <- MME(X,Z,AI*3,RI,dat$y)
R> res$b: res$u
Γ17 143.8930 159.3976
[1] -1.675214 3.350427 -2.065980 -2.122054 2.280054
[6] 2.326783
```



W.

### Example

Fit models

```
R> modA <- gpMod(maizeC,model="BLUP",kin=A)
R> modU <- gpMod(maizeC,model="BLUP",kin=U)</pre>
```

Predicted genetic values

```
R> gA <- predict(modA)
R> gU <- predict(modU)</pre>
```

Extract true breeding values

```
R> tbv <- maizeC$covar$tbv[maizeC$covar$phenotyped]</pre>
```

Evaluate correlations with tbv

```
R> cor(gA,tbv)
0.587
R> cor(gU,tbv)
0.856
```





```
R> summary(modU)
Object of class 'gpMod'
Model used: BLUP
Nr. observations 1250
Genetic performances:
 Min. 1st Qu. Median Mean 3rd Qu. Max
-19.4200 -3.4210 -0.2841 0.0000 3.2830 15.3000
Model fit
Likelihood kernel: K = (Intercept)
Maximized log likelihood with kernel K is -3223.837
```

Linear Coefficients:
Estimate Std. Error

(Intercept) 148.921 0.197

Estimate Std. Error kinTS 53.055 7.359 In 48.577 2.287

Variance Coefficients:

### Prediction of unphenotyped individuals

Discard last 50 individuals from the data set

```
R> last50 <- rownames(maizeC$pheno)[1201:1250]
```

R> maizeC2 <- discard.individuals(maizeC,last50)</pre>

Fit modU using the variance-covariance structure from the whole data set

```
R> modU24 <- gpMod(maizeC2,model="BLUP",kin=U)</pre>
```

Prediction for the last 50 individuals

R> g <- predict(modU24,rownames(maizeC\$pheno)[1201:1250])</pre>





#### Model cross-validation

- Prospects for GS derived by out-of-sample performance
- Cross-validation as assumption-free method
- Divide data set in *k* mutually exclusive subsets
- k-1 form the estimation set (ES), kth subset is used as independent test set (TS)
- Model validation by
  - Predictive ability r(ĝ<sub>TS</sub>, y<sub>TS</sub>)
  - Prediction bias
- Sampling schemes: random, within family, across family (Albrecht et al. 2011)





### Example

```
R> cv.maize <- crossVal(maizeC,cov.matrix=list(U),k=5,Rep=2,Seed=1
```

#### R> summary(cv.maize)

Object of class 'cvData'

```
5 -fold cross validation with 2 replications
Sampling: random
```

Sampling: random
Variance components: committed

Number of random effects: 1

Number of individuals: 1250

Size of the TS: 250 -- 250

#### Results:

	Min	Mean +- pooled SE
Predictive ability:	0.4589	0.5287 +- 0.0079
Bias:	0.8747	1.0061 +- 0.0253

0.5691 1.1179

Max



### Bayesian Lasso

The model (de los Campos et al. 2009)

$$y_i = \mu + \sum_{j=1}^p x_{ij}\beta_j + \varepsilon_i$$

with the prior distributions (Park and Casella 2008)

$$eta_j \sim \mathsf{N}(0,\sigma^2 au_j^2) \ au_j^2 \sim \mathsf{Exp}(\lambda^2) \ \lambda^2 \sim \mathsf{Ga}(lpha,eta) ext{ or } rac{\lambda}{\lambda_{\mathsf{max}}} \sim \mathsf{Beta}(a,b) \ \sigma^2 \sim \chi^{-2}(v,S)$$





## Choice of hyperparameters

According to Pérez et al. (2010):

$$\lambda_{start} = \sqrt{2\sum_{j=1}^p ar{x}_{,j}^2 rac{(1-h^2)}{h^2}}$$

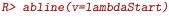
```
R> y <- maize$pheno[,1,]
R> X <- maize$geno
R> sX2 <- sum(X^2)
R> h2 <- 0.5 # priori expectation
R> (lambdaStart <- sqrt(2*sum(X^2)*(1-h2)/h2/nrow(X)))
[1] 38.39858</pre>
```

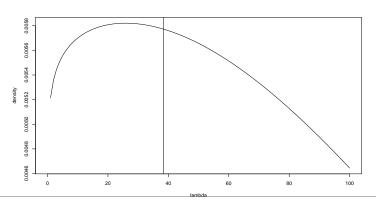




## Choice of hyperparameters

```
R> lambda <- seq(from=0,to=100,by=1)
R> dens <- dgamma(x=lambda^2,shape=.52,rate=3e-5)*lambda*2 # distri
R> plot(dens~lambda,type='l',ylab="density")
R> ablino(y=lambdaStart)
```









### Run Bayesian Lasso

#### Evaluation on the whole data set:

```
R> prior <- list(varE=list(df=3,S=35),lambda = list(shape=0.52,rate=R> modBL <- gpMod(maizeC,model="BL",prior=prior,nIter=6000,burnIn=1000</pre>
```

Use CV to evaluate the predictive ability:

```
R> cv.BL <- crossVal(maizeC,k=5,Rep=2,Seed=123,sampling="random",VC
R> summary(cv.BL)
```





## Gateway from symbreed to package qtl

- Package qt1 for QTL analysis in experimental crosses (Broman et al. 2003)
- Main data class cross
- Conversion from gpData to cross
  - R> gpData2cross(gpDataObj)
- Conversion from cross to gpData
  - R> cross2gpData(cross0bj)





# Problems 3 - 1 (Corn borer example)

- Try to reproduce the results of Illustration 3.1 using function MME.
- 2 Combine the pedigree, the phenotypes and genotypes in an object of class gpData.
- Set up the matrix A for the individuals and plot a heatmap of it.
- Try to reproduce the results of Illustration 4.1 using function MME.
- Set up the matrix U for the individuals and plot a heatmap of it. Discuss the the differences with regard to contents compared to the matrix A.

#### Problems 3 - 2

- Construct a genomic relationship matrix U according to Habier et al. (2007) and fit a GBLUP model. What are the estimated variance components?
- Make a manhattan plot of the estimated marker effects.
- Predict the unphenotyped individuals in your data set using the predict method for the GBLUP model. If all individuals are phenotyped, mask 10% of the phenotypes an predict them.
- Use CV to routinely estimate the predictive ability of the GBLUP model in your data. Commit for each CV model the variance components estimated with the whole data set.
- What is the definition of the bias in the summary of the CV? Try to interpret the values you obtain with your data.

## Problems 3 - 3 (Advanced)

- Try to fit different types of genomic relationship matrices using the function kin. Use them in a linear mixed model as variance-covariance structure (using function gpMod) and compare the variance components you obtain. For further connections between the matrices, see Albrecht et al. (2011). Use CV to estimate the predictive ability of the different models. What do you observe?
- Check the help for function MME. Try to replicate the results from problem 1. First, you need to extract the necessary parts from the gpData object. Next, you need to set up the variance-covariance structure using the U matrix and the estimated variance components from problem 1.
- Use the function gpData2cross to convert your object to an object of class cross for package qtl.
- Use the function scanone of package to scan for QTLs and display the LOD curve you obtain along the genome.

#### Literature

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- Broman, K. W., H. Wu, S. Sen, and G. A. Churchill, 2003 R/qtl: Qtl mapping in experimental crosses. Bioinformatics 7: 889–890. R package version 1.20-15.
- Browning, B. L., and S. R. Browning, 2009 A unified approach to genotype imputation and haplotype-phase inference for large data sets of trios and unrelated individuals. The American Journal of Human Genetics 846: 210–223. Version 3.3.1.
- de los Campos, G., H. Naya, D. Gianola, A. L. José Crossa, E. Manfredi, et al., 2009 Predicting quantitative traits with regression models for dense molecular markers and pedigree. Genetics 182: 375–385.
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- Valdar, W., L. Solberg, D. Gauguier, W. Cookson, J. Rawlins, et al., 2006 Genetic and environmental effects on complex traits in mice. Genetics 174: 959–984.
- Wimmer, V., T. Albrecht, H.-J. Auinger, and C.-C. Schoen, 2012 synbreed: a framework for the analysis of genomic prediction data using r. Bioinformatics 28: 2086–2087.



()