

MATH3823 Generalized Linear Models

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1/12/23

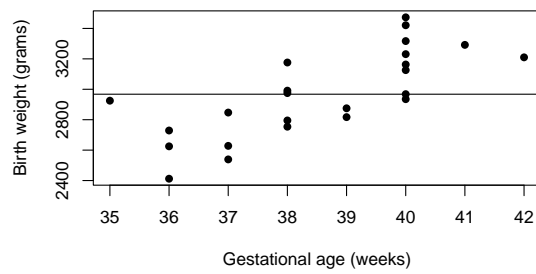
Table of contents

Appendix C: Background to Analysis of Variance	3
C.1 Analysis of Variance	3
C.2 Distributions derived from the Gaussian distribution	5
The Gaussian (normal) distribution	5
The Chi-squared distribution	5
The t- and F-distributions	6

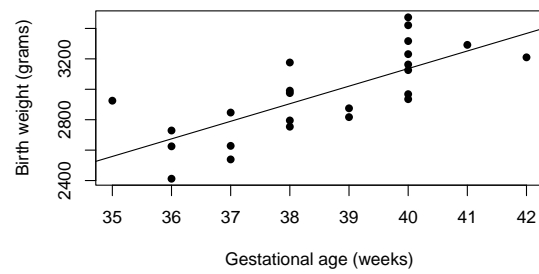
Appendix C: Background to Analysis of Variance

C.1 Analysis of Variance

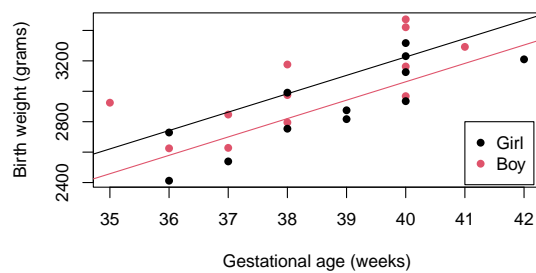
Consider the four models fitted to the birth weight data. Figure 1 shows the data set along with the corresponding fitted model as a single line, for the models which do not take **Sex** into account, and two lines, for the models which include **Sex**.



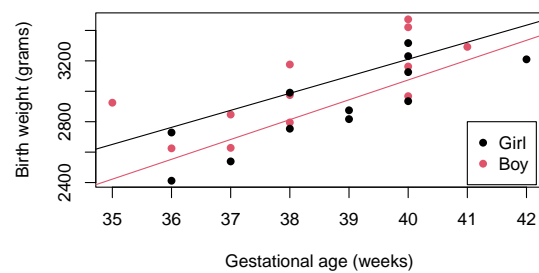
(a) Model 0



(b) Model 1



(c) Model 2



(d) Model 3

Figure 1: Birthweight and gestational age data with superimposed fitted regression lines from various competing models.

The *residual sum of squares* (RSS) takes into account the vertical distance between the fitted model and the data values. Let R_k denote the residual sum of squares for Model k : $R_k = \sum_{i=1}^n (y_i - \hat{\mu}_{ki})^2$, where $\hat{\mu}_{ki}$ is the fitted value for individual i under Model k and let r_k denote the corresponding *residual degrees of freedom* for Model k (the number of observations minus the number of model parameters). The table below shows these values for the four models fitted to the data.

Table 1: Summary of the residual sums of squares

Model k	R_k	r_k	$R_k - R_{k-1}$	$r_k - r_{k-1}$
0	1829873	23		
1	816074	22	1013799	1
2	658771	21	157304	1
3	652425	20	6346	1

The table also shows the change in residual sums of squares, $R_k - R_{k-1}$, which measures the improvement in the fit due to the extra parameters used in Model k compared to Model $k - 1$. The RSS and changes in RSS values are also shown in Figure 2. It is clear that there is a substantial reduction in RSS moving from Model 0 to Model 1, but small reductions as further parameters are added to the model. We might guess that Model 1 will be the “best” model, but as it is not acceptable to base a choice on our personal subjective opinion but instead a sequence of hypothesis tests will be used.

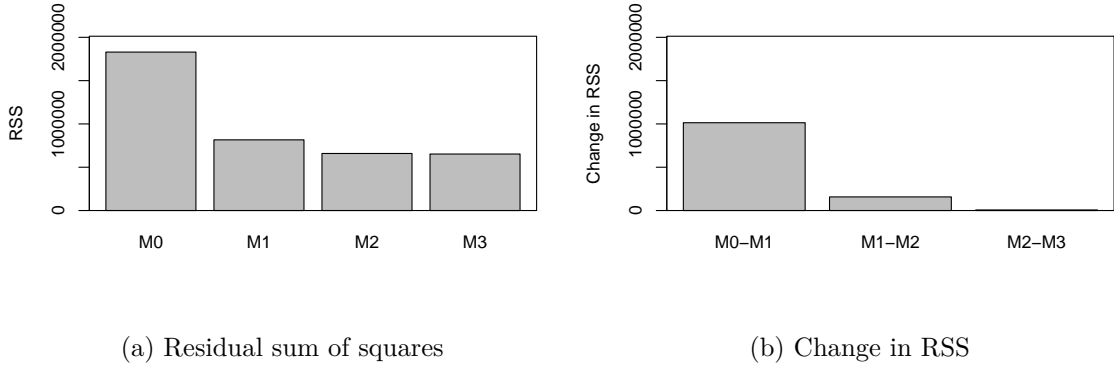


Figure 2: Birthweight and gestational age data with superimposed fitted regression lines from various competing models.

Here, a sequence of three hypothesis tests is considered: Starting with

Test 1 H_0 : Model 0 is true; H_1 : Model 1 is true.

Which can be judged by comparing $R_1 - R_0 = 1013799$ which follows a $\sigma^2\chi^2$ distribution on $r_1 - r_0 = 1$ degrees of freedom ($(R_1 - R_0)/\sigma^2$ follows a χ^2_1 distribution) with $R_1 = 816074$ which follows a $\sigma^2\chi^2$ distribution on $r_1 = 22$ degrees of freedom (R_1/σ^2 follows a

χ_{22}^2 distribution. Fortunately, taking the ratio eliminates σ^2 giving the test statistics

$$F_{01} = \frac{(R_1 - R_0)/(r_1 - r_0)}{R_1/r_1} = \frac{1013799/1}{816074/22} = 27.33$$

If H_0 is true, then we would expect this to be close to 1. The 5%, 1% and 0.1% critical values for the distribution are 4.3, 7.95, 14.38, and the observed F statistics is much larger than all these and hence p-value < 0.001 meaning we reject H_0 in favour of H_1 .

If H_0 had been accepted then the sequence would stop and Model 0 declared the best, whereas H_0 is rejected and the next test is considered

Test 2 H_0 : Model 1 is true; H_1 : Model 2 is true.

If H_0 is accepted here then the sequence stops and Model 1 is declared the best, whereas if H_0 is rejected then the last test is considered

Test 3 H_1 : Model 2 is true; H_1 : Model 3 is true.

If H_0 is accepted here then the sequence stops and Model 2 is declared the best, whereas if H_0 is rejected then Model 3 is declared the best.

C.2 Distributions derived from the Gaussian distribution

The Gaussian (normal) distribution

If $X \sim N(\mu, \sigma^2)$ then

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{1}{2} \frac{(x - \mu)^2}{\sigma^2} \right\}, \quad -\infty < x < \infty.$$

Properties:

1. The parameter $\mu = E[X]$ is a location parameter and $\sigma^2 = \text{Var}[X]$ is a scale parameter.
2. If $X \sim N(\mu, \sigma^2)$ then $aX + b \sim N(a\mu + b, a^2\sigma^2)$.
3. If $X_i \sim N(\mu_i, \sigma_i^2), i = 1, \dots, n$ (independent) then $\sum a_i X_i \sim N(\sum a_i \mu_i, \sum a_i^2 \sigma_i^2)$.
4. A special case is when $\mu = 0$ and $\sigma^2 = 1$ which is called the *standard normal* distribution.

The Chi-squared distribution

If X has a Chi-squared distribution, $X \sim \chi_\nu^2$ then

$$f(x) = \frac{\left(\frac{1}{2}\right)^{\frac{\nu}{2}} x^{\frac{\nu}{2}-1} e^{-\frac{1}{2}x}}{\Gamma\left(\frac{\nu}{2}\right)}, \quad x \geq 0, \nu > 0 \text{ and integer.}$$

with $E[X] = \nu$ and $\text{Var}[X] = 2\nu$.

Properties:

1. The parameter ν is a shape parameter and is called the *degrees of freedom*. The pdf is positive skew, but becomes more symmetric as ν increases.
2. If $Z \sim N(0, 1)$ then $Z^2 \sim \chi_1^2$.
3. If $X_i \sim \chi_{\nu_i}^2, i = 1, \dots, n$ (independent) then $\sum X_i \sim \chi_\nu^2$, where $\nu = \sum \nu_i$.
4. If $Z_i \sim N(0, 1), i = 1, \dots, n$ (independent) then $\sum Z_i^2 \sim \chi_n^2$.
5. This is a special case of the gamma distribution, with $\alpha = \nu/2$ and $\lambda = \frac{1}{2}$, that is $\gamma(\frac{\nu}{2}, \frac{1}{2})$.

The t- and F-distributions

If X has a t-distribution, $X \sim t_\nu$ then

$$f(x) = \frac{1}{\sqrt{\pi\nu}} \frac{\Gamma((\nu+1)/2)}{\Gamma(\nu/2)} \left(1 + \frac{x^2}{\nu}\right)^{-\frac{1}{2}(\nu+1)} \quad -\infty < x < \infty,$$

where $\nu > 0$ and integer.

Properties:

1. The parameter ν is called the *degrees of freedom*.
2. If $X \sim N(0, 1)$ and $Y \sim \chi_\nu^2$ (independent) then

$$\frac{X}{\sqrt{Y/\nu}} \sim t_\nu.$$

3. If $X \sim t_\nu$ then $X^2 \sim F_{1,\nu}$.
4. $t_\nu \rightarrow N(0, 1)$ as $\nu \rightarrow \infty$.

If X has an F-distribution, $X \sim F_{\nu_1, \nu_2}$ then

$$f(x) = \frac{\nu_1^{\frac{\nu_1}{2}} \nu_2^{\frac{\nu_2}{2}} x^{\frac{\nu_1}{2}-1}}{B(\frac{\nu_1}{2}, \frac{\nu_2}{2})(\nu_2 + \nu_1 x)^{\frac{\nu_2+\nu_1}{2}}} \quad x \geq 0.$$

where $\nu_1, \nu_2 > 0$ and integer are known as the *degrees of freedom*.

Properties:

1. The parameters ν_1 and ν_2 are called the degrees of freedom.
2. If $X_1 \sim \chi_{\nu_1}^2$ and $X_2 \sim \chi_{\nu_2}^2$ (independent) then

$$\frac{X_1/\nu_1}{X_2/\nu_2} \sim F_{\nu_1, \nu_2}.$$

3. If $X \sim F_{\nu_1, \nu_2}$ then $1/X \sim F_{\nu_2, \nu_1}$, hence, $Pr(F_{\nu_1, \nu_2} < c) = Pr(F_{\nu_2, \nu_1} > 1/c)$.