

An Automatic Finite-Sample Robustness Metric: Can Dropping a Little Data Make a Big Difference?

Ryan Giordano (rgiordan@mit.edu)¹
January 2022

¹With coauthors Rachael Meager (LSE) and Tamara Broderick (MIT)

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The conclusions of one's statistical analysis may depend on only a **small fraction of the data**, even for **highly significant results in correctly specified models**.

We provide a **generally applicable tool** to detect such sensitivity. Our methods are **efficiently and automatically computable**, and come with **finite-sample accuracy guarantees** and **clear intuition**.

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Dropping data: Mexico Microcredit

Example: Angelucci et al. [2015], a randomized controlled trial study of the efficacy of microcredit in Mexico based on 16,560 data points. The variable “Beta” estimates the effect of microcredit in US dollars.

	Beta (SE)
Original result	-4.55 (5.88)

The original conclusion: No evidence that microcredit is effective...

⇒ Standard errors can be inadequate summaries of data sensitivity!

Cannot find influential subsets by brute force! $\binom{16,560}{15} \approx 1.5 \cdot 10^{51}$

We provide a fast, automatic tool to approximately identify the most influential set of points.

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...can be reversed by dropping less than 0.1% of the data.

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We provide a fast, automatic tool to approximately identify the most influential set of points.

- Why and when might you care about sensitivity to data dropping?
- How does our approximation work, and when is it accurate?
(A formalization of the problem and the class of estimators we study.)
- Examine real-life examples of analyses: some sensitive, some not.
(The results may defy your intuition.)
- What kinds of analyses are sensitive to data dropping?
(Including comparison to standard errors and gross-error robustness.)

Dropping data: Motivation

Would you be concerned if you could **reverse your conclusion** by removing a **small proportion** of your data?

Not always! But sometimes, surely yes.

Thinking without random noise can be helpful.

Suppose you have a farm, and want to know whether your average yield is > 170 bushels per acre. At harvest, you measure 200 bushels per acre.

- Scenario one: If your yield is greater than 170 bushels per acre, you make a profit.
 - Don't care about sensitivity to small subsets
- Scenario two: You want to recommend your farming methods to a friend across the valley.
 - Might care about sensitivity to small subsets

For example, often in economics:

- Policy population is different from analyzed population,
- Small fractions of data are missing not-at-random,
- We report a convenient summary (e.g. mean) of a complex effect.

Formalizing the question.

Ordinary least squares

A data point d_n has regressors x_n and response y_n : $d_n = (x_n, y_n)$.

The estimator $\hat{\theta} \in \mathbb{R}^p$ satisfies:

$$\hat{\theta} := \arg \min_{\theta} \frac{1}{2} \sum_{n=1}^N (y_n - \theta^T x_n)^2$$

$$\Leftrightarrow \sum_{n=1}^N (y_n - \hat{\theta}^T x_n) x_n = 0.$$

Make a qualitative decision using:

- A particular component: θ_k
- The end of a confidence interval: $\theta_k + \frac{1.96}{\sqrt{N}} \hat{\sigma}(\hat{\theta})$

Z-estimators

We observe N data points d_1, \dots, d_N (in any domain).

The estimator $\hat{\theta} \in \mathbb{R}^p$ satisfies:

$$\sum_{n=1}^N G(\hat{\theta}, d_n) = 0_p.$$

$G(\cdot, d_n)$ is “nice,” \mathbb{R}^p -valued.
E.g. OLS, MLE, VB, IV &c.

Make a qualitative decision using $\phi(\hat{\theta})$ for a smooth, real-valued ϕ .

(WLOG try to increase $\phi(\hat{\theta})$.)

Question: Can we make a big change in $\phi(\hat{\theta})$ by dropping $\lfloor \alpha N \rfloor$ datapoints, for some small proportion α ?

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- There are $\binom{N}{\lfloor \alpha N \rfloor}$ sets to check. (Huge even for $\alpha \ll 1$.)
- Evaluating $\hat{\theta}$ re-solving the estimating equation.
 - E.g., re-computing the OLS estimator.
 - Other examples are even harder (VB, machine learning)

Idea: Smoothly approximate the effect of leaving out points.

We have N data points d_1, \dots, d_N , a quantity of interest $\phi(\cdot)$, and

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Original weights: $\vec{1} = (1, \dots, 1)$



Leave points out by setting their elements of \vec{w} to zero.

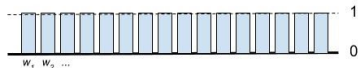


The map $\vec{w} \mapsto \phi(\hat{\theta}(\vec{w}))$ is well-defined even for continuous weights.

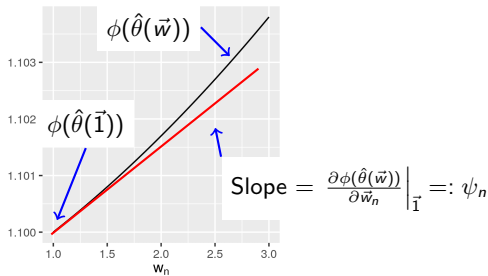
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The values $N\psi_n$ are the **empirical influence function** [Hampel, 1986]. We call ψ_n an “influence scores.”

We can use ψ_n to form a Taylor series approximation:

$$\phi(\hat{\theta}(\vec{w})) \approx \phi^{\text{lin}}(\vec{w}) := \phi(\hat{\theta}(\vec{1})) + \sum_{n=1}^N \psi_n (\vec{w}_n - 1)$$

Taylor series approximation.

Problem: How much can you change $\phi(\hat{\theta}(\vec{w}))$ dropping $\lfloor \alpha N \rfloor$ points?
Combinatorially hard by brute force!

Approximate Problem: How much can you change $\phi^{\text{lin}}(\hat{\theta}(\vec{w}))$ dropping $\lfloor \alpha N \rfloor$ points? **Easy!**

$$\phi^{\text{lin}}(\vec{w}) := \phi(\hat{\theta}(\vec{1})) + \sum_{n=1}^N \psi_n(\vec{w}_n - 1)$$

Dropped points have $\vec{w}_n - 1 = -1$. Kept points have $\vec{w}_n - 1 = 0$
 \Rightarrow The most influential points for $\phi^{\text{lin}}(\vec{w})$ have the most negative ψ_n .

Procedure: (see rgiordan/zaminfluence on github)

- 1 Compute your original estimator $\hat{\theta}(\vec{1})$.
- 2 Compute and sort the influence scores $\psi_{(1)}, \dots, \psi_{(N)}$.
- 3 Worry if $-\sum_{n=1}^{\lfloor \alpha N \rfloor} \psi_{(n)}$ is large enough to change your conclusions.

How to compute the ψ_n 's? And how accurate is the approximation?

How to compute the influence scores?

How can we compute the influence scores $\psi_n = \left. \frac{\partial \phi(\hat{\theta}(\vec{w}))}{\partial \vec{w}_n} \right|_{\vec{1}}$?

By the **chain rule**, $\psi_n = \left. \frac{\partial \phi(\theta)}{\partial \theta} \right|_{\hat{\theta}(\vec{1})} \left. \frac{\partial \hat{\theta}(\vec{w})}{\partial \vec{w}_n} \right|_{\vec{1}}$.

Recall that $\sum_{n=1}^N \vec{w}_n G(\hat{\theta}(\vec{w}), d_n) = 0_P$ for all \vec{w} near $\vec{1}$.

\Rightarrow By the **implicit function theorem**, we can write $\left. \frac{\partial \hat{\theta}(\vec{w})}{\partial \vec{w}_n} \right|_{\vec{1}}$ as a linear system involving $G(\cdot, \cdot)$ and its derivatives.

\Rightarrow The ψ_n are automatically computable from $\hat{\theta}(\vec{1})$ and software implementations of $G(\cdot, \cdot)$ and $\phi(\cdot)$ using **automatic differentiation**.

```
import jax
import jax.numpy as np
def phi(theta):
    ... computations using np and theta ...
    return value

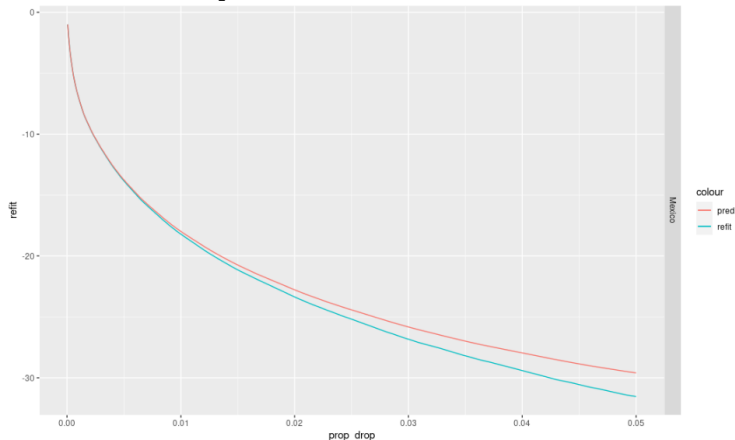
# Exact gradient of phi (1st term in the chain rule):
jax.grad(phi)(theta_opt)
```

See [rgiordan/vittles](#) on github.

How accurate is the approximation?

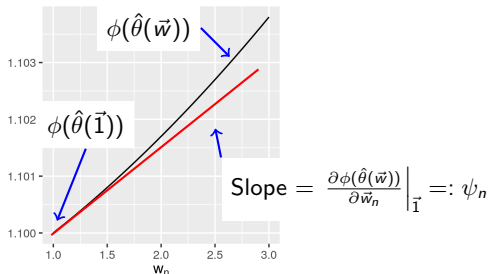
Mexico example:

See `microcredit_profit_sandbox.R`.



How accurate is the approximation?

By controlling the curvature, we can control the error in the linear approximation.



We provide **finite-sample theory** [Giordano et al., 2019] showing that

$$\left| \phi(\hat{\theta}(\vec{w})) - \phi^{\text{lin}}(\vec{w}) \right| = O \left(\left\| \frac{1}{N}(\vec{w} - \vec{1}) \right\|_2^2 \right) = O(\alpha) \text{ as } \alpha \rightarrow 0.$$

But you don't need to rely on the theory!

Our method returns which points to drop. **Re-running once** without those points provides an **exact lower bound** on the worst-case sensitivity.

Selected experimental results.

Original estimate (SE)	Refit estimate (SE)	Observations dropped
-4.549 (5.879)	7.030 (2.549)*	15 = 0.09%

Table: Microcredit Mexico results [Angelucci et al., 2015].

A * indicates statistical significance at the 95% level.

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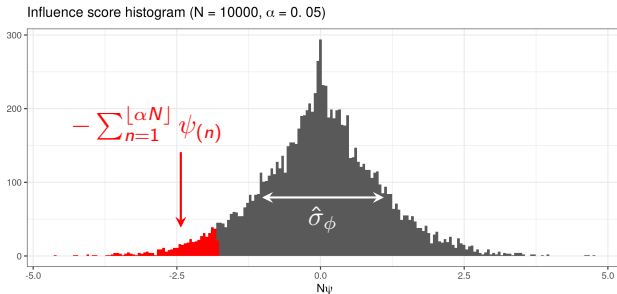
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What makes an analysis sensitive?

Suppose we are willing to drop $\lfloor \alpha N \rfloor$ points. We are “sensitive to data dropping” if, for some Δ large enough to change our conclusions,

$$\text{“Signal”} := \Delta < \phi^{\text{lin}}(\vec{w}^*) - \phi(\hat{\theta}) = - \sum_{n=1}^{\lfloor \alpha N \rfloor} \psi_{(n)} =: \hat{\sigma}_{\phi} \hat{\mathcal{J}}_{\alpha}$$

- The “noise” $\hat{\sigma}_{\phi}^2 \rightarrow \text{Var}(\sqrt{N}\phi)$ (“sandwich” variance estimator)
- The “shape” $\hat{\mathcal{J}}_{\alpha} := \frac{-\sum_{n=1}^{\lfloor \alpha N \rfloor} \psi_{(n)}}{\hat{\sigma}_{\phi}} \rightarrow \text{nonzero constant} \leq \sqrt{\alpha(1-\alpha)}$



Example.

α := Proportion of points to drop

Δ := Signal (difference large enough to change conclusions)

$\hat{\sigma}_\phi$:= Noise (estimator of $\text{Var}(\sqrt{N}\phi)$)

$\hat{\mathcal{J}}_\alpha$:= Shape (some constant determined by α and $N\psi_n$ tail shape)

Sensitive to data dropping if:

$$\phi^{\text{lin}}(\vec{w}^*) - \phi(\hat{\theta}) = \hat{\sigma}_\phi \hat{\mathcal{J}}_\alpha \geq \Delta \quad \Leftrightarrow \quad \frac{\Delta}{\hat{\sigma}_\phi} \leq \hat{\mathcal{J}}_\alpha.$$

The **signal to noise ratio** $\frac{\Delta}{\hat{\sigma}_\phi}$ determines sensitivity to data dropping.

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Corollaries.

Report non-robustness if:

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Corollary: Leave- $\lfloor \alpha N \rfloor$ -out is different from standard errors.

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Corollary: Leave- $\lfloor \alpha N \rfloor$ -out robustness does not vanish as $N \rightarrow \infty$.

Both $\hat{\mathcal{J}}_{\alpha}$ and $\hat{\sigma}_{\phi}$ typically converge to nonzero constants.

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Corollary: Non-robustness possible even with correct specification.

Report non-robustness if:

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Corollary: To robustify, reduce the noise or increase the signal.

Other forms of robustness

We proceeded as follows:

- 1 Took presence of datapoints as a model input,
- 2 Formed an automatically-computable differential approximation,
- 3 Provided theory by analyzing higher-order derivatives,
- 4 Studied its effectiveness in problems with open-access data.

Presence of datapoints is only one model input of many!

- Prior sensitivity in Bayesian nonparametrics [Giordano et al., 2021]
- Model sensitivity of MCMC output [Gustafson, 2000, Giordano et al., 2018]
- Cross-validation [Giordano et al., 2019, Wilson et al., 2020]
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- Frequentist variances of MCMC posteriors (in progress)

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- We can quickly and automatically find an approximate influential set which is accurate for small sets.
- Robustness to removing small sets is principally determined by the signal to noise ratio.
- In the present work, we studied data dropping. But we provide a framework for studying many other robustness questions, both to data and model perturbations.

Tamara Broderick, Ryan Giordano, Rachael Meager (alphabetical authors)
“An Automatic Finite-Sample Robustness Metric: Can Dropping a Little Data Change Conclusions?”

<https://arxiv.org/abs/2011.14999>

Open-source software:

R package `zaminfluence` <https://github.com/rgiordan/zaminfluence>

Python package `vittles` <https://github.com/rgiordan/vittles>

Some related content can be found on my blog:

<https://rgiordan.github.io/>

- M. Angelucci and G. De Giorgi. Indirect effects of an aid program: How do cash transfers affect ineligibles' consumption? *American Economic Review*, 99(1):486–508, 2009.
- M. Angelucci, D. Karlan, and J. Zinman. Microcredit impacts: Evidence from a randomized microcredit program placement experiment by Compartamos Banco. *American Economic Journal: Applied Economics*, 7(1):151–82, 2015.
- A. Baydin, B. Pearlmutter, A. Radul, and J. Siskind. Automatic differentiation in machine learning: A survey. *The Journal of Machine Learning Research*, 18(1):5595–5637, 2017.
- A. Finkelstein, S. Taubman, B. Wright, M. Bernstein, J. Gruber, J. Newhouse, H. Allen, K. Baicker, and Oregon Health Study Group. The Oregon health insurance experiment: Evidence from the first year. *The Quarterly Journal of Economics*, 127(3):1057–1106, 2012.
- R. Giordano, T. Broderick, and M. I. Jordan. Covariances, robustness and variational Bayes. *The Journal of Machine Learning Research*, 19(1):1981–2029, 2018.
- R. Giordano, W. Stephenson, R. Liu, M. I. Jordan, and T. Broderick. A swiss army infinitesimal jackknife. In *The 22nd International Conference on Artificial Intelligence and Statistics*, pages 1139–1147. PMLR, 2019.
- R. Giordano, R. Liu, M. I. Jordan, and T. Broderick. Evaluating sensitivity to the stick-breaking prior in Bayesian nonparametrics. 2021.
- P. Gustafson. Local robustness in Bayesian analysis. In *Robust Bayesian Analysis*, pages 71–88. Springer, 2000.
- F. Hampel. *Robust statistics: The approach based on influence functions*, volume 196. Wiley-Interscience, 1986.
- A. Wilson, M. Kasy, and L. Mackey. Approximate cross-validation: Guarantees for model assessment and selection. In *International Conference on Artificial Intelligence and Statistics*, pages 4530–4540. PMLR, 2020.