

An Automatic Finite-Sample Robustness Metric: Can Dropping a Little Data Make a Big Difference?

Ryan Giordano (rgiordan@mit.edu)¹
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¹With coauthors Rachael Meager (LSE) and Tamara Broderick (MIT)

Dropping data: Motivation

More data & cheaper computation \Rightarrow

Statistical analyses are playing larger roles in decision making.

Decisions are important: We want **trustworthy** conclusions.

Data / models not always perfect: We want **robust** conclusions.

Would you be concerned if you could **reverse your conclusion** by removing a **small proportion** (say, 0.1%) of your data?

Running example: Angelucci et al. [2015], a randomized controlled trial study of the efficacy of microcredit based on 16,560 data points.

We can reverse the studies qualitative conclusions by removing 15 observations ($< 0.1\%$ of the data).

How do we find sets of influential points? Difficult in general!

We provide a **automatic approximation** with finite-sample guarantees.

Studying the approximation reveals the causes of non-robustness.

Dropping data: Mexico Microcredit

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The variable “Beta” estimates the effect of microcredit in US dollars.

	Beta (SE)
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Original conclusion:

There is no evidence that microcredit is effective.

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The culprit is signal to noise ratio.

By the end of the talk, we will see that the sensitivity is due to

- High variability of the outcome (household profit) relative to
- A small signal driving the conclusion (statistical significance)

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Would you be concerned if you could **reverse your conclusion** by removing a **small proportion** (say, 0.1%) of your data?

Not always! But sometimes, surely yes.

Thinking without random noise can be helpful.

Suppose you have a farm, and want to know whether your average yield is greater than 170 bushels per acre. At harvest, you measure 200 bushels per acre.

- Scenario one: If your yield is greater than 170 bushels per acre, you make a profit.
 - Don't care about sensitivity to small subsets
- Scenario two: You want to recommend your farming methods to a friend across the valley.
 - Might care about sensitivity to small subsets

For example, often in economics:

- Small fractions of data are missing not-at-random,
- Policy population is different from analyzed population,
- We report a convenient summary (e.g. mean) of a complex effect,
- Models are stylized proxies of reality.

Which estimators do we study?

Z-estimators. Suppose we have N data points $\vec{d} = d_1, \dots, d_N$. Then:

$$\hat{\theta} := \vec{\theta} \text{ such that } \sum_{n=1}^N G(\vec{\theta}, d_n) = 0_P.$$

Examples: MLE, OLS, VB, &c (all minimizers of smooth empirical loss).

Function of interest. Qualitative decision based on $\phi(\hat{\theta}) \in \mathbb{R}$. E.g.:

- A particular component: $\phi(\theta) = \theta_d$
- The end of a confidence interval: $\phi(\theta) = \theta_d + \frac{1.96}{\sqrt{N}} \hat{\sigma}(\hat{\theta})$

Fix a proportion $0 < \alpha \ll 1$ of points to drop and find a set $\mathcal{S} \subset \{1, \dots, N\}$ with $|\mathcal{S}| \leq \lfloor \alpha N \rfloor$ that extremizes $\phi(\hat{\theta})$ when dropped.

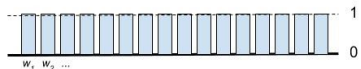
- **Problem:** There are many sets with $|\mathcal{S}| \leq \lfloor \alpha N \rfloor$.
 - E.g., in Angelucci et al. [2015], $\binom{16,560}{15} \approx 1.5 \cdot 10^{51}$
- **Problem:** Evaluating $\phi(\hat{\theta}(\vec{d}_{-\mathcal{S}}))$ requires an estimation problem.
 - E.g., in Angelucci et al. [2015] computing the OLS estimator.
 - Other examples are even harder (VB, machine learning)

An approximation is needed!

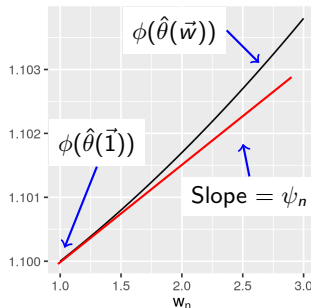
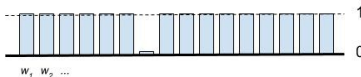
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$$\hat{\theta} := \vec{\theta} \text{ such that } \sum_{n=1}^N G(\vec{\theta}, d_n) = 0_P.$$

Original weights: $\vec{1} = (1, \dots, 1)$



Leave points out by setting their elements of \vec{w} to zero.



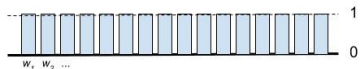
The slopes $\psi_n := \left. \frac{\partial \phi(\hat{\theta}(\vec{w}))}{\partial w_n} \right|_{\vec{1}}$ are values of the **empirical influence function** [Hampel, 1986]. We call them “influence scores.”

Second-order derivatives control the error of the linear approximation.

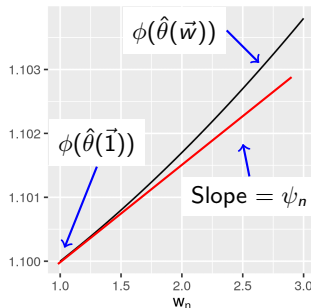
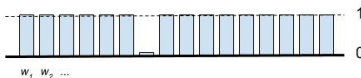
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Taylor series approximation.

Problem: How large can you make $\phi(\hat{\theta}(\vec{w}))$ leaving out no more than $\lfloor \alpha N \rfloor$ points? **Combinatorially hard!**

To simplify the search over \vec{w} , we form the Taylor series approximation:

$$\phi(\hat{\theta}(\vec{w})) \approx \phi^{\text{lin}}(\vec{w}) := \phi(\hat{\theta}(\vec{1})) + \sum_{n=1}^N \psi_n(\vec{w}_n - 1)$$

Approximate solution: How large can you make $\phi^{\text{lin}}(\vec{w})$ leaving out no more than $\lfloor \alpha N \rfloor$ points? **Easy!**

The most influential points for $\phi^{\text{lin}}(\vec{w})$ have the most negative ψ_n .

We provide **finite-sample theory** showing that

$$\left| \phi(\hat{\theta}(\vec{w})) - \phi^{\text{lin}}(\vec{w}) \right| = O \left(\left\| \frac{1}{N}(\vec{w} - \vec{1}) \right\|_2^2 \right) = O(\alpha) \text{ as } \alpha \rightarrow 0.$$

How to compute the influence scores ψ_n ?

By the chain rule, $\psi_n = \left. \frac{\partial \phi(\hat{\theta}(\vec{w}))}{\partial \vec{w}_n} \right|_{\vec{1}} = \left. \frac{d\phi(\theta)}{d\theta^T} \right|_{\hat{\theta}} \left. \frac{\partial \hat{\theta}(\vec{w})}{\partial \vec{w}_n} \right|_{\vec{1}}.$

Recall that $\hat{\theta}(\vec{w}) := \vec{\theta}$ such that $\sum_{n=1}^N \vec{w}_n G(\vec{\theta}, d_n) = 0_P.$

The **implicit function theorem** expresses $\left. \frac{\partial \hat{\theta}(\vec{w})}{\partial \vec{w}_n} \right|_{\vec{1}}$ as a linear system.

Computation of ψ_n is fully automatable from a software implementation of $G(\cdot, \cdot)$ and $\phi(\cdot)$ with **automatic differentiation** [Baydin et al., 2017].

We have an R package, `rgiordan/zaminfluence`, for OLS and IV.

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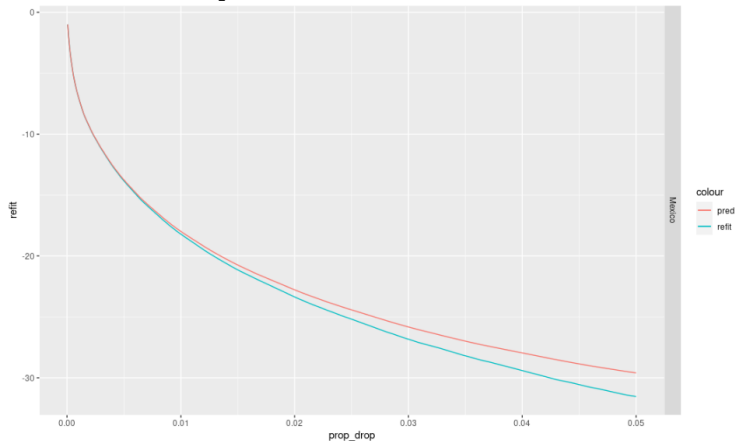
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- 6 **Optional:** Compute $\hat{\theta}(\vec{w}^*)$, and verify that $\phi(\hat{\theta}(\vec{w}^*)) - \phi(\hat{\theta}) \geq \Delta$.

Mexico example:

See `microcredit_profit_sandbox.R`.



Selected experimental results.

Study case	Original estimate (SE)	Target change	Refit estimate	Observations dropped
Mexico	-4.549 (5.879)	Sign change	0.398 (3.194)	1 = 0.01%
		Significance change	-10.962 (5.565)*	14 = 0.08%
		Significant sign change	7.030 (2.549)*	15 = 0.09%

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Table: Medicaid profit results [Finkelstein et al., 2012]

A simulation

For $N = 5,000$ data points, compute the OLS estimator from:

Regressors
 $x_n \sim \mathcal{N}(0, \sigma_x^2)$

Residuals
 $\varepsilon_n \sim \mathcal{N}(0, \sigma_\varepsilon^2)$

Responses
 $y_n = 0.5x_n + \varepsilon_n$

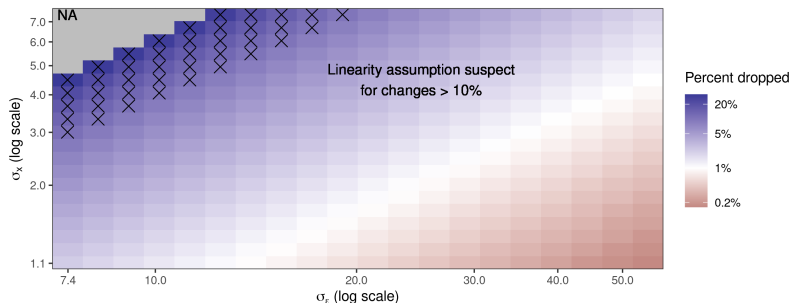


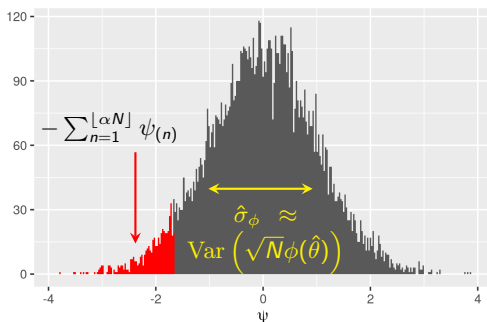
Figure: The approximate perturbation inducing proportion at differing values of σ_x and σ_ε . Red colors indicate datasets whose sign can be predicted to change when dropping less than 1% of datapoints. The grey areas indicate $\hat{\Psi}_\alpha = \text{NA}$, a failure of the linear approximation to locate any way to change the sign.

What makes an estimator non-robust? A tail sum.

We show that $\phi^{\text{lin}}(\vec{w}^*) - \phi(\hat{\theta}) = -\sum_{n=1}^{\lfloor \alpha N \rfloor} \psi_{(n)} =: \hat{\sigma}_{\phi} \hat{\mathcal{T}}_{\alpha}$ where

- The “noise” $\hat{\sigma}_{\phi}^2 \rightarrow \text{Var}(\sqrt{N}\phi)$
 - $\hat{\sigma}_{\phi}^2$ is the robust “sandwich” variance estimator [Hampel, 1986]
- The “shape” $\hat{\mathcal{T}}_{\alpha} \leq \sqrt{\alpha(1-\alpha)}$ determined by ψ_n distribution
 - $\hat{\mathcal{T}}_{\alpha}$ converges to a nonzero constant

Influence score histogram (N = 10000, $\alpha = 0.05$)



Example.

Report non-robustness if:

$$\phi^{\text{lin}}(\vec{w}^*) - \phi(\hat{\theta}) = \hat{\sigma}_{\phi} \hat{\mathcal{T}}_{\alpha} \geq \Delta \quad \Leftrightarrow \quad \frac{\Delta}{\hat{\sigma}_{\phi}} \leq \hat{\mathcal{T}}_{\alpha}.$$

The **signal to noise ratio** $\frac{\Delta}{\hat{\sigma}_{\phi}}$ determines sensitivity to data dropping.

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Let's analyze with $\alpha = 0.01 = 1\%$.

$$\begin{array}{llll}
 \phi(\hat{\theta}) = & -0.029 & (\text{Increase QOI by defn}) & \Delta = 0.029 \\
 \hat{\sigma}_{\phi} = & 0.766 & (\text{Noise}) & \frac{1}{\sqrt{N}} \hat{\sigma}_{\phi} = 0.005 \quad (\text{SE}) \\
 \hat{\mathcal{T}}_{\alpha} = & 0.046 & (\text{Shape}) & \frac{1.96}{\sqrt{N}} = 0.0128 \rightarrow 0 \text{ as } N \rightarrow \infty \\
 \hat{\mathcal{T}}_{\alpha} \hat{\sigma}_{\phi} = & 0.035 & (\text{Data dropping sensitivity}) & \frac{1.96}{\sqrt{N}} \hat{\sigma}_{\phi} = 0.010 \quad (\text{SE sensitivity})
 \end{array}$$

The noise is much larger than the signal \Rightarrow Sensitive to data dropping.

Corollaries.

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Both $\hat{\mathcal{J}}_{\alpha}$ and $\hat{\sigma}_{\phi}$ typically converge to nonzero constants.

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Corollary: Gross outliers primarily affect robustness through $\hat{\sigma}_{\phi}$.

See paper for more details.

Other forms of robustness

We proceeded as follows:

- ① Took presence of datapoints as a model input,
- ② Formed an automatically-computable differential approximation,
- ③ Provided theory by analyzing higher-order derivatives,
- ④ Studied its effectiveness in problems with open-access data.

Presence of datapoints is only one model input of many!

- Prior sensitivity in Bayesian nonparametrics [Giordano et al., 2021]
- Model sensitivity of MCMC output [Giordano et al., 2018]
- Cross-validation [Giordano et al., 2019]
- Frequentist variances of MCMC posteriors (in progress)

Tamara Broderick, Ryan Giordano, Rachael Meager (alphabetical authors)
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<https://arxiv.org/abs/2011.14999>

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- M. Angelucci and G. De Giorgi. Indirect effects of an aid program: How do cash transfers affect ineligibles' consumption? *American Economic Review*, 99(1):486–508, 2009.
- M. Angelucci, D. Karlan, and J. Zinman. Microcredit impacts: Evidence from a randomized microcredit program placement experiment by Compartamos Banco. *American Economic Journal: Applied Economics*, 7(1):151–82, 2015.
- A. Baydin, B. Pearlmutter, A. Radul, and J. Siskind. Automatic differentiation in machine learning: A survey. *The Journal of Machine Learning Research*, 18(1):5595–5637, 2017.
- A. Finkelstein, S. Taubman, B. Wright, M. Bernstein, J. Gruber, J. Newhouse, H. Allen, K. Baicker, and Oregon Health Study Group. The Oregon health insurance experiment: Evidence from the first year. *The Quarterly Journal of Economics*, 127(3):1057–1106, 2012.
- R. Giordano, T. Broderick, and M. I. Jordan. Covariances, robustness and variational Bayes. *The Journal of Machine Learning Research*, 19(1):1981–2029, 2018.
- R. Giordano, W. Stephenson, R. Liu, M. I. Jordan, and T. Broderick. A swiss army infinitesimal jackknife. In *The 22nd International Conference on Artificial Intelligence and Statistics*, pages 1139–1147. PMLR, 2019.
- R. Giordano, R. Liu, M. I. Jordan, and T. Broderick. Evaluating sensitivity to the stick-breaking prior in Bayesian nonparametrics. 2021.
- F. Hampel. *Robust statistics: The approach based on influence functions*, volume 196. Wiley-Interscience, 1986.