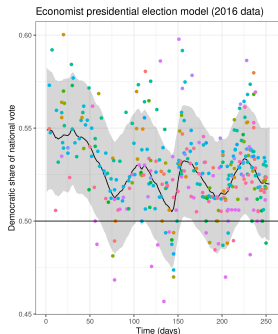


Approximate data deletion and replication with the Bayesian influence function

Ryan Giordano (rgiordano@berkeley.edu, UC Berkeley), Tamara Broderick (MIT)

MIT Robustness and Influence Functions Workshop

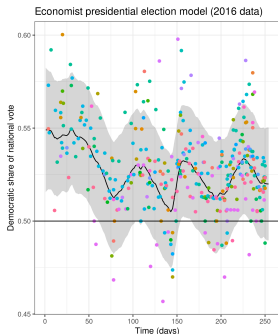
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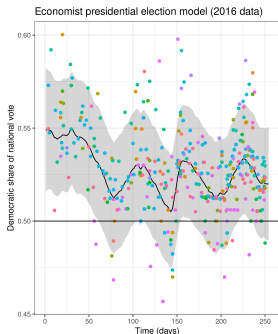
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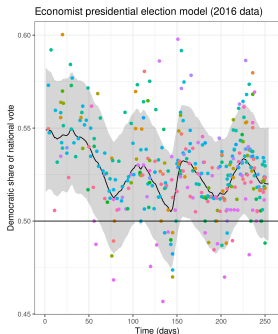
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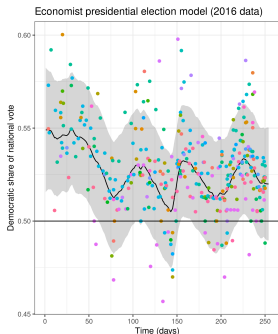
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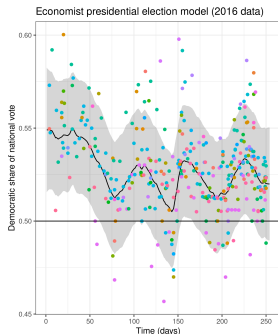
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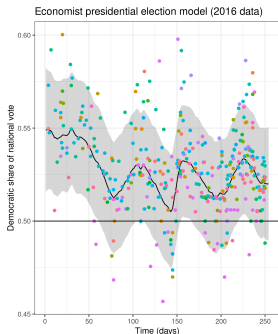
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We want to know $\mathbb{E}_{p(\theta|X)}[f(\theta)]$.

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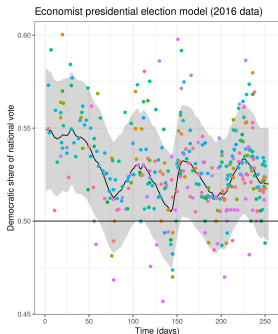
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- We are interested re-sampling for *this* election, not a hypothetical future election

Idea: Re-fit with bootstrap samples of data [Huggins and Miller, 2023]

Problem: Each MCMC run takes about 10 hours (Stan, six cores).

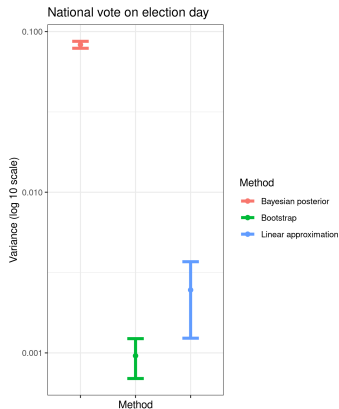
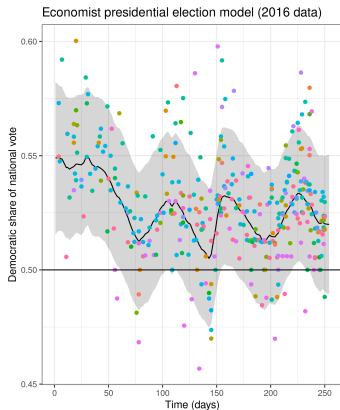
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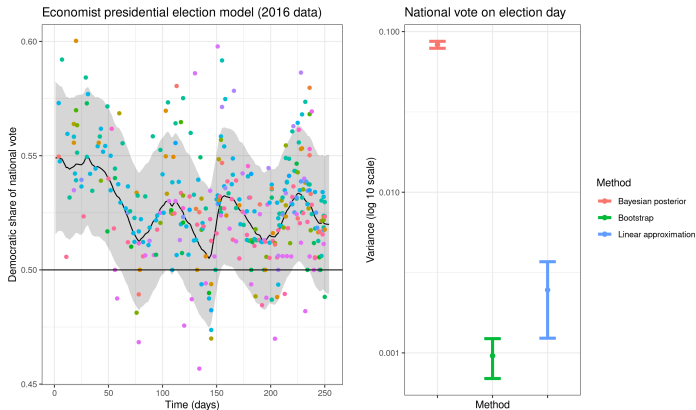


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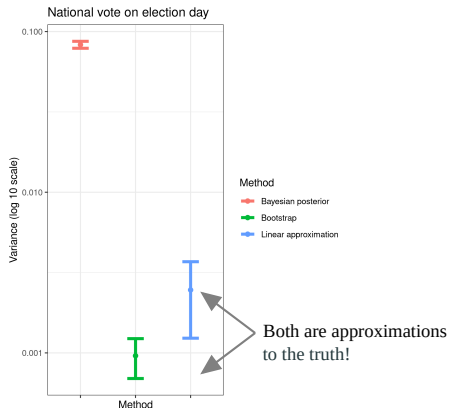
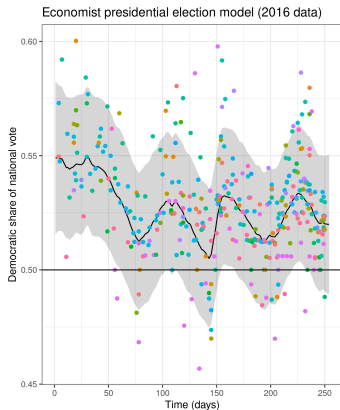
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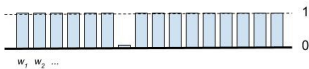
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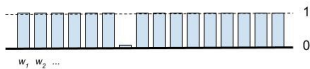
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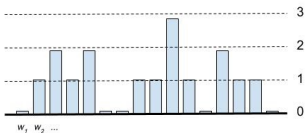
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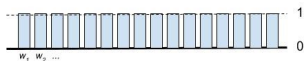


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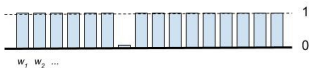
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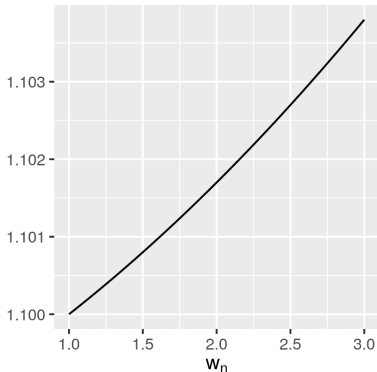
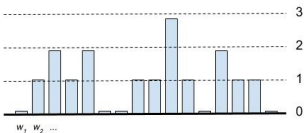
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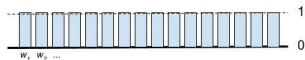


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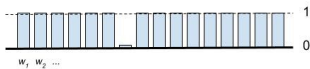
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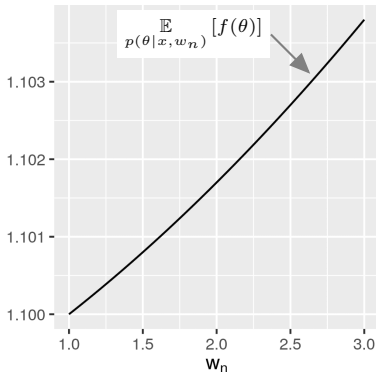
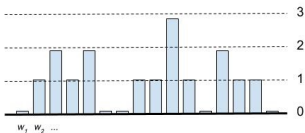
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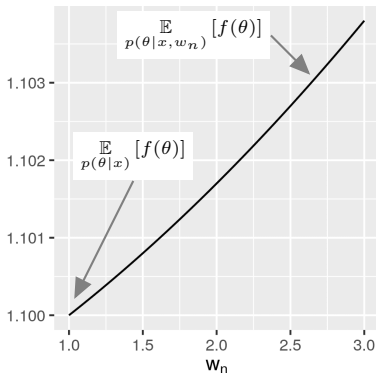
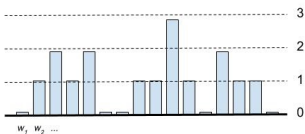
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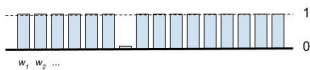
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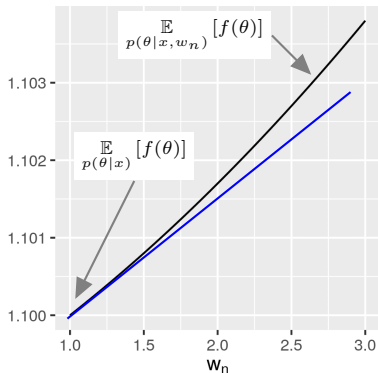
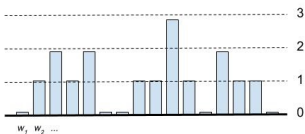
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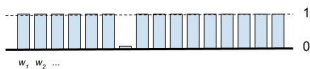
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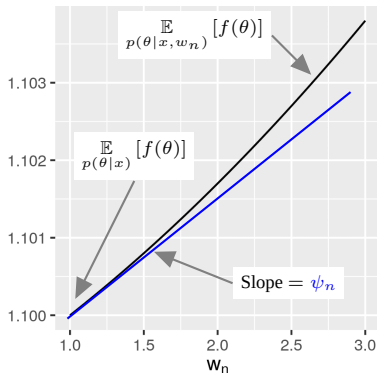
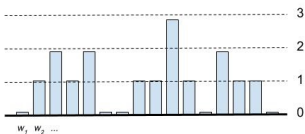
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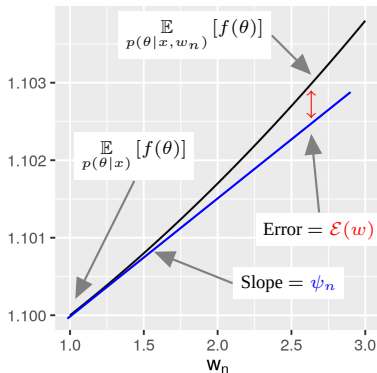
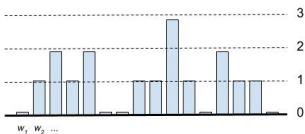
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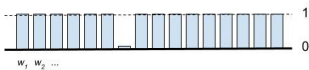
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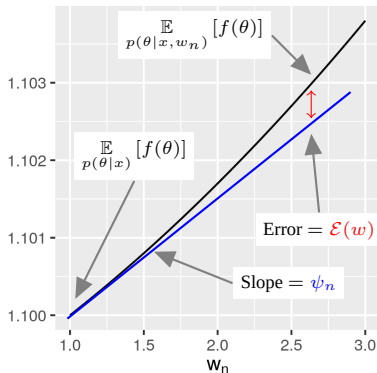
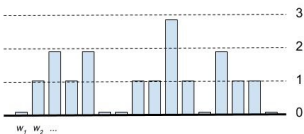
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$$\begin{aligned}\text{Bootstrap variance} &= \text{Var}_{p(w)} \left(\mathbb{E}_{p(\theta|x,w)} [f(\theta)] \right) \\ &\approx \text{Var}_{p(w)} \left(\mathbb{E}_{p(\theta|x)} [f(\theta)] + \psi_n(w_n - 1) \right) \quad (\text{assuming the error is small}) \\ &= \sum_{n=1}^N \left(\psi_n - \bar{\psi} \right)^2.\end{aligned}$$

The final line is also known as the “infinitesimal jackknife” variance approximation.

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Furthermore, by the mean value theorem, for some \tilde{w} ,

$$\mathcal{E}(w) = \frac{1}{2} \sum_{n=1}^N \sum_{n'=1}^N \mathcal{E}_{nn'}(w) (w_n - 1)(w_{n'} - 1) \quad \text{where}$$
$$\mathcal{E}_{nn'}(w) := \underbrace{\mathbb{E}_{p(\theta|X,\tilde{w})} [\bar{f}(\theta) \bar{\ell}_n(\theta) \bar{\ell}_{n'}(\theta)]}_{\substack{\text{Cannot compute directly!} \\ \text{(we don't know the intermediate value theorem's } \tilde{w} \text{).} \\ \text{But we can analyze it.}}}$$

Here, an overbar denotes “posterior–mean zero.” For example, $\bar{f}(\theta) := f(\theta) - \mathbb{E}_{p(\theta|X)} [f(\theta)]$.

How good is the linear approximation (IJ covariance) as an approximation of the limiting variance of $\sqrt{N} \mathbb{E}_{p(\theta|X)} [f(\theta)]$?

Theoretical results

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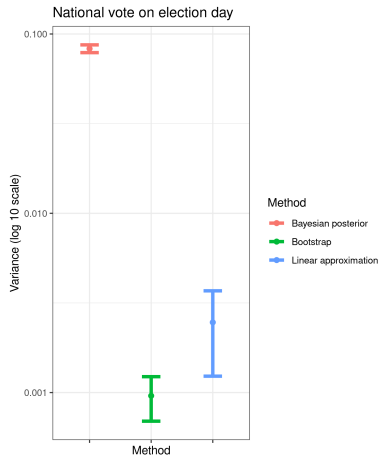
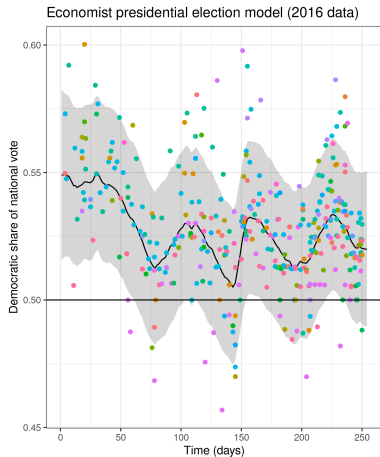
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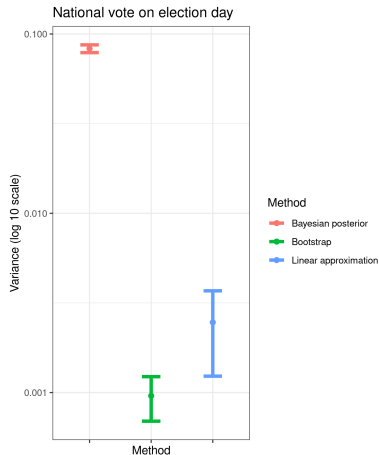
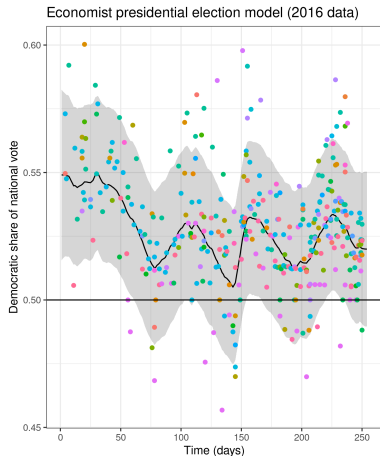
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\Rightarrow High–dimensional Bayesian models are an extremely common class of problems for which the influence function may not provide a good approximation.

Observations and consequences



Observations and consequences



Preprint: Giordano and Broderick [2023] (arXiv:2305.06466)

- Detailed proofs
- Simple analytical examples
- Simulated and real-world experiments

- A. Gelman and M. Heidemanns. The Economist: Forecasting the US elections., 2020. URL <https://projects.economist.com/us-2020-forecast/president>. Data and model accessed Oct., 2020.
- R. Giordano and T. Broderick. The Bayesian infinitesimal jackknife for variance. *arXiv preprint arXiv:2305.06466*, 2023.
- J. Huggins and J. Miller. Reproducible model selection using bagged posteriors. *Bayesian Analysis*, 18(1):79–104, 2023.

How can we use the approximation?

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Cross validation. Let $w_{(-n)}$ leave out point n , and loss $f(\theta) = -\ell(x_n|\theta)$.

$$\text{LOO CV loss at point } n = \mathbb{E}_{p(\theta|x, w_{(-n)})} [f(\theta)] \approx \mathbb{E}_{p(\theta|x)} [f(\theta)] - \psi_n$$

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Example: Approximate bootstrap.

Draw bootstrap weights $w \sim p(w) = \text{Multinomial}(N, N^{-1})$.

$$\begin{aligned} \text{Bootstrap variance} &= \text{Var}_{p(w)} \left(\mathbb{E}_{p(\theta|x, w)} [f(\theta)] \right) \\ &\approx \text{Var}_{p(w)} \left(\mathbb{E}_{p(\theta|x)} [f(\theta)] + \psi_n(w_n - 1) \right) \\ &= \sum_{n=1}^N \left(\psi_n - \bar{\psi} \right)^2. \end{aligned}$$

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Influential subsets: Approximate maximum influence perturbation (AMIP).

Let $W_{(-K)}$ denote weights leaving out K points.

$$\max_{w \in W_{(-K)}} \left(\mathbb{E}_{p(\theta|x, w)} [f(\theta)] - \mathbb{E}_{p(\theta|x)} [f(\theta)] \right) \approx - \sum_{n=1}^K \psi_{(n)}.$$

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Consider $p(X|\gamma) = \prod_{n=1}^N \text{NegativeBinomial}(x_n|\gamma)$. Here, $\theta = \gamma$ is a scalar.

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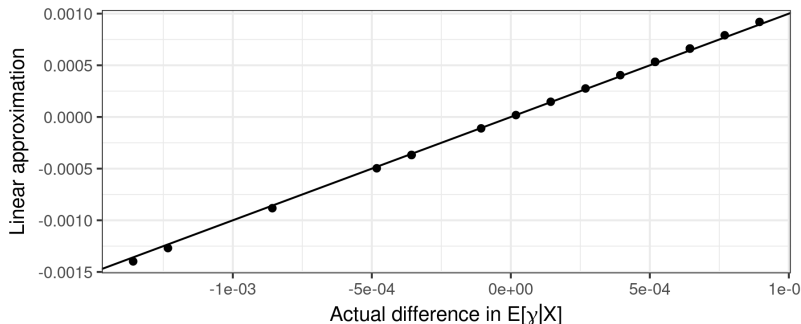
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Negative Binomial model
leaving out single datapoints with $N = 800$



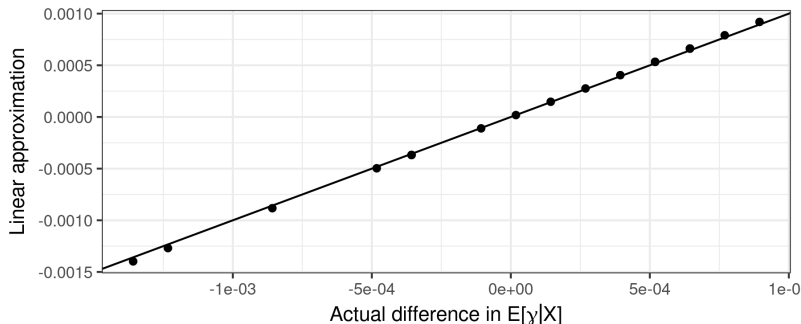
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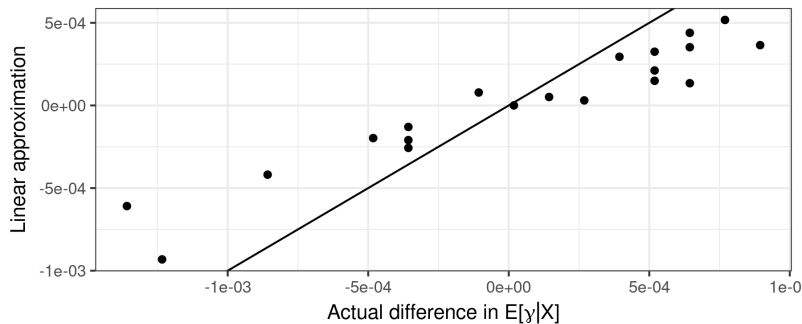
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Problem: Most computationally hard Bayesian problems don't concentrate.

Example: **Poisson model with random effects (REs) λ and fixed effect γ .**

Poisson random effect model
leaving out single datapoints with $N = 800$



A contradiction?

Negative binomial observations.

Asymptotically linear in w .

Poisson observations with random effects.

Asymptotically non-linear in w .

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With a constant regressor, Gamma REs, and one RE per observation,
these are the same model, with the same $p(\gamma|X)$.

Is $\mathbb{E}_{p(\gamma|X,w)} [\gamma]$ linear in the data weights or not?

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$$\log p(X|\gamma, w^m) = \sum_{n=1}^N w_n^m \log p(x_n|\gamma)$$

Poisson observations with random effects.

Asymptotically non-linear in w .

$$\log p(X|\gamma, \lambda, w^c) = \sum_{n=1}^N w_n^c \log p(x_n|\lambda, \gamma)$$

With a constant regressor, Gamma REs, and one RE per observation, these are the same model, with the same $p(\gamma|X)$.

Is $\mathbb{E}_{p(\gamma|X, w)} [\gamma]$ **linear in the data weights** or not?

Trick question! We weight a log likelihood contribution, not a datapoint.

The two weightings are not equivalent in general.

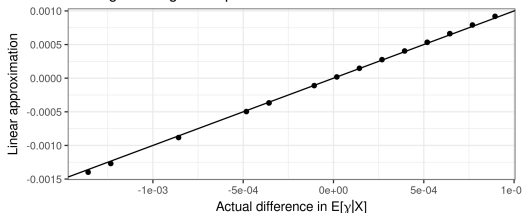
Experimental results

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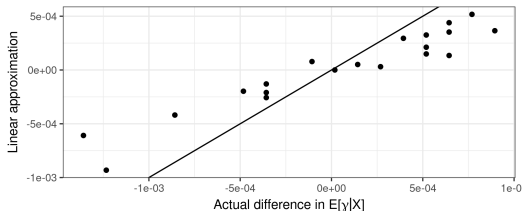
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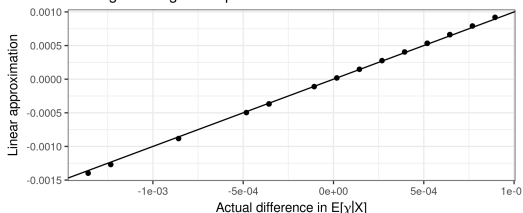
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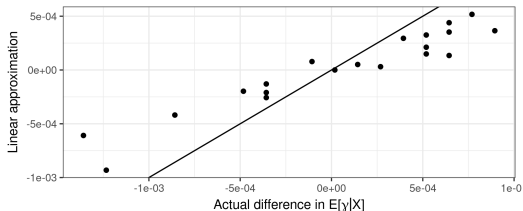
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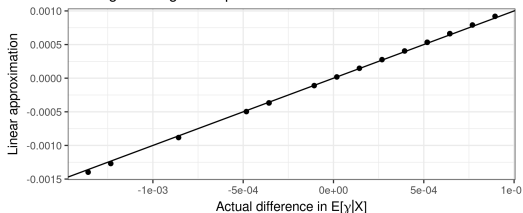
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May still be useful when $p(\lambda | X)$ is *somewhat* concentrated.

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