An Automatic Finite-Sample Robustness Metric: Can Dropping a Little Data Make a Big Difference?

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Dropping data: Motivation

More data & cheaper computation \Rightarrow Statistical analyses are playing larger roles in decision making.

Decisions are important: We want **trustworthy** conclusions. Data / models not always perfect: We want **robust** conclusions.

Would you be concerned if you could **reverse your conclusion** by removing a **small proportion** (say, 0.1%) of your data?

Running example: Angelucci et al. [2015], a randomized controlled trial study of the efficacy of microcredit based on 16,560 data points. We can reverse the studies qualitative conclusions by removing 15 observations (< 0.1% of the data).

How do we find sets of influential points? Difficult in general!

We provide a **automatic approximation** with finite-sample guarantees.

Studying the approximation reveals the causes of non-robustness.

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The culprit is signal to noise ratio.

By the end of the talk, we will see that the sensitivity is due to

- High variability of the outcome (hosehold profit) relative to
- A small signal driving the conclusion (statistical significance)

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Would you be concerned if you could **reverse your conclusion** by removing a **small proportion** (say, 0.1%) of your data? Not always! But sometimes, surely yes.

Thinking without random noise can be helpful.

Suppose you have a farm, and want to know whether your average yield is greater than 170 bushels per acre. At harvest, you measure 200 bushels per acre.

- Scenario one: If your yield is greater than 170 bushels per acre, you
 make a profit.
 - Don't care about sensitivity to small subsets
- Scenario two: You want to recommend your farming methods to a friend across the valley.
 - Might care about sensitivity to small subsets

For example, often in economics:

- Small fractions of data are missing not-at-random,
- Policy population is different from analyzed population,
- We report a convenient summary (e.g. mean) of a complex effect,
- Models are stylized proxies of reality.

Which estimators do we study?

Z-estimators. Suppose we have N data points $\vec{d} = d_1, \dots, d_N$. Then:

$$\hat{\theta} := \vec{\theta}$$
 such that $\sum_{n=1}^{N} G(\vec{\theta}, d_n) = 0_P$.

Examples: MLE, OLS, VB, &c (all minimizers of smooth empirical loss).

Function of interest. Qualitative decision based on $\phi(\hat{\theta}) \in \mathbb{R}$. E.g.:

- A particular component: $\phi(\theta) = \theta_d$
- The end of a confidence interval: $\phi(\theta) = \theta_d + \frac{1.96}{\sqrt{N}}\hat{\sigma}(\hat{\theta})$

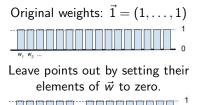
Fix a proportion $0 < \alpha \ll 1$ of points to drop and find a set $\mathcal{S} \subset \{1, \dots N\}$ with $|\mathcal{S}| \leq \lfloor \alpha N \rfloor$ that extremizes $\phi(\hat{\theta})$ when dropped.

- **Problem:** There are many sets with $|\mathcal{S}| \leq \lfloor \alpha N \rfloor$. • E.g., in Angelucci et al. [2015], $\binom{16,560}{15} \approx 1.5 \cdot 10^{51}$
- ullet Problem: Evaluating $\phi(\hat{ heta}(ec{d}_{-\mathcal{S}}))$ requires an estimation problem.
 - E.g., in Angelucci et al. [2015] computing the OLS estimator.
 - Other examples are even harder (VB, machine learning)

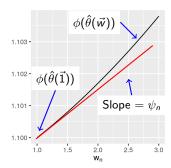
An approximation is needed!

Which estimators do we study?

$$\hat{\theta} := \vec{\theta} \text{ such that } \sum_{n=1}^{N} G(\vec{\theta}, d_n) = 0_P.$$



W, W2 ...

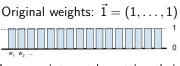


The slopes $\psi_n := \frac{\partial \phi(\hat{\theta}(\vec{w}))}{\partial \vec{w}_n}\Big|_{\vec{1}}$ are values of the **empirical influence** function [Hampel, 1986]. We call them "influence scores."

Second-order derivatives control the error of the linear approximation.

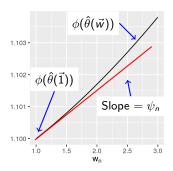
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$$\hat{\theta}(\vec{w}) := \vec{\theta}$$
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Leave points out by setting their elements of \vec{w} to zero.





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Problem: How large can you make $\phi(\hat{\theta}(\vec{w}))$ leaving out no more than $\lfloor \alpha N \rfloor$ points? **Combinatorially hard!**

To simplify the search over \vec{w} , we form the Taylor series approximation:

$$\phi(\hat{ heta}(\vec{w})) pprox \phi^{ ext{lin}}(\vec{w}) := \phi(\hat{ heta}(\vec{1})) + \sum_{n=1}^{N} \psi_n(\vec{w}_n - 1)$$

Approximate solution: How large can you make $\phi^{\text{lin}}(\vec{w})$ leaving out no more than $|\alpha N|$ points? **Easy!**

The most influential points for $\phi^{\text{lin}}(\vec{w})$ have the most negative ψ_n .

We provide finite-sample theory showing that

$$\left|\phi(\hat{ heta}(ec{w})) - \phi^{\mathrm{lin}}(ec{w})
ight| = O\left(\left\|rac{1}{N}(ec{w} - ec{1})
ight\|_2^2\right) = O\left(lpha
ight) ext{ as } lpha o 0.$$

How to compute the influence scores ψ_n ?

By the chain rule,
$$\psi_n = \frac{\partial \phi(\hat{\theta}(\vec{w}))}{\partial \vec{w}_n} \Big|_{\vec{1}} = \frac{\mathrm{d}\phi(\theta)}{\mathrm{d}\theta^T} \Big|_{\hat{\theta}} \frac{\partial \hat{\theta}(\vec{w})}{\partial \vec{w}_n} \Big|_{\vec{1}}.$$

Recall that $\hat{\theta}(\vec{w}) := \vec{\theta}$ such that $\sum_{n=1}^{N} \vec{w}_n G(\vec{\theta}, d_n) = 0_P$.

The implicit function theorem expresses $\frac{\partial \hat{\theta}(\vec{w})}{\partial \vec{w}_n}\Big|_{\vec{1}}$ as a linear system.

Computation of ψ_n is fully automatible from a software implemenation of $G(\cdot, \cdot)$ and $\phi(\cdot)$ with **automatic differentiation** [Baydin et al., 2017].

We have an R package, rgiordan/zaminfluence, for OLS and IV.

Procedure:

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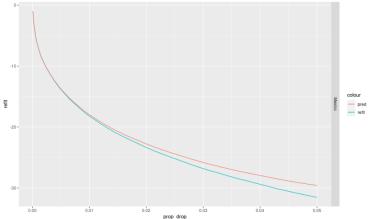
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- **3** Report non-robustness if $\phi^{\text{lin}}(\vec{w}^*) \phi(\hat{\theta}) = -\sum_{n=1}^{\lfloor \alpha N \rfloor} \psi_{(n)} \geq \Delta$.

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- **Optional:** Compute $\hat{\theta}(\vec{w}^*)$, and verify that $\phi(\hat{\theta}(\vec{w}^*)) \phi(\hat{\theta}) \geq \Delta$.

Mexico example:

See ${\tt microcredit_profit_sandbox.R.}$



Selected experimental results.

- The "Refit estimate" column shows the result of re-fitting the model removing the points with the largest influence scores.
- Stars indicate significance at the 5% level.
- Refits that achieved the desired change are bolded.

Study case	Original estimate	Target change	Refit estimate	Observations dropped
Mexico	-4.549 (5.879)	Sign change Significance change Significant sign change	0.398 (3.194) -10.962 (5.565)* 7.030 (2.549)*	1 = 0.01% 14 = 0.08% 15 = 0.09%

Table: Microcredit Mexico results [Angelucci et al., 2015].

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Poor, period 10	33.861 (4.468)*	Sign change Significance change Significant sign change	-2.559 (3.541) 4.806 (3.684) -9.416 (3.296)*	697 = 6.63% 435 = 4.14% 986 = 9.37%
Non-poor, period 10	21.493 (9.405)*	Sign change Significance change Significant sign change	-0.573 (6.750) 16.262 (8.927) -10.845 (6.467)	30 = 0.70% 3 = 0.07% 92 = 2.16%

Table: Cash transfers results. [Angelucci and De Giorgi, 2009]

A simulation

For N = 5,000 data points, compute the OLS estimator from:

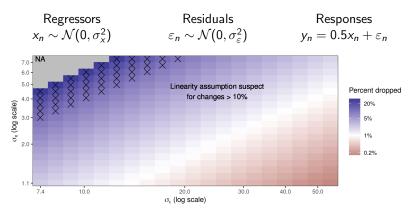


Figure: The approximate perturbation inducing proportion at differing values of σ_x and σ_ε . Red colors indicate datasets whose sign can is predicted to change when dropping less than 1% of datapoints. The grey areas indicate $\hat{\Psi}_\alpha = \text{NA}$, a failure of the linear approximation to locate any way to change the sign.

What makes an estimator non-robust? A tail sum.

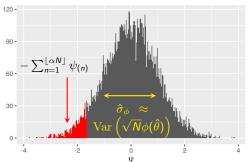
Report non-robustness if:

$$\Delta \leq \phi^{\text{lin}}(\vec{w}^*) - \phi(\hat{\theta}) = -\sum_{n=1}^{\lfloor \alpha N \rfloor} \psi_{(n)} =: \hat{\sigma}_{\phi} \hat{\mathcal{J}}_{\alpha}$$

We will show that:

- The "noise" $\hat{\sigma}_{\phi}^2 o \mathrm{Var}(\sqrt{\textit{N}}\phi)$ [Hampel, 1986]
- ullet The "shape" $\hat{\mathscr{T}}_{lpha} \leq \sqrt{lpha(1-lpha)}$ and converges to a nonzero constant

Influence score histogram (N = 10000, α = 0.05)



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We call $\frac{\Delta}{\hat{\sigma}_{\phi}}$ the "signal to noise ratio."

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Corollary: Gross outliers primarily affect robustness through $\hat{\sigma}_{\phi}$. Cauchy-Schwartz is tight when all the influence scores are the same.

Conclusion: Related work and future directions

Links and references

Tamara Broderick, Ryan Giordano, Rachael Meager (alphabetical authors)

"An Automatic Finite-Sample Robustness Metric: Can Dropping a Little Data Change Conclusions?"

https://arxiv.org/abs/2011.14999

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