# An Automatic Finite-Sample Robustness Metric: Can Dropping a Little Data Make a Big Difference?

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### Dropping data: Mexico Microcredit

**Example:** Angelucci et al. [2015], a randomized controlled trial study of the efficacy of microcredit in Mexico based on N=16,560 data points. A regression was run to estimate the average effect of microcredit.

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Cannot find influential subsets by brute force!

We provide a fast, automatic tool to approximately identify the most influential set of points.

#### Outline

- Why and when might you care about sensitivity to data dropping?
- How do we identify influential sets? When is our method accurate?
   (A formalization of the problem and the class of estimators we study.)
- Examine real-life examples of analyses: some sensitive, some not. (The results may defy your intuition.)
- What kinds of analyses are sensitive to data dropping?
   (Comparison to standard errors, gross errors, and how to mitigate.)

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Specifically, often in statistical applications:

- Policy population is different from analyzed population,
- Small fractions of data are missing not-at-random,
- We report a convenient summary (e.g. mean) of a complex effect.

**Example: Least squares** 

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Make a qualitative decision using  $\phi(\hat{\theta})$  for a smooth, real-valued  $\phi$ .

(WLOG try to increase  $\phi(\hat{\theta})$ .)

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Original weights:  $\vec{1} = (1, \dots, 1)$ 



Leave points out by setting their elements of w to zero.



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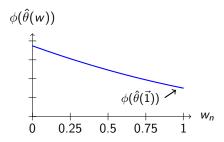
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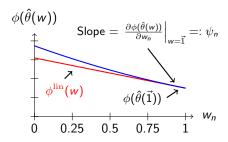


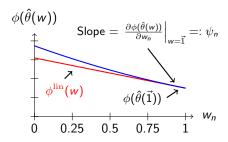
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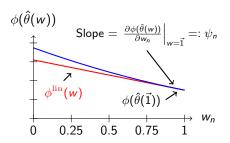
The map  $w \mapsto \phi(\hat{\theta}(w))$  is well-defined even for continuous weights.







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We can use  $\psi_n$  to form a Taylor series approximation:

$$\phi(\hat{\theta}(w)) \approx \phi^{\text{lin}}(w) := \phi(\hat{\theta}(\vec{1})) + \sum_{n=1}^{N} \psi_n(w_n - 1)$$

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Our procedure: (see rgiordan/zaminfluence on github)

- Compute your original estimator  $\hat{\theta}(\vec{1})$ .
- **②** Compute and sort the influence scores  $\psi_{(1)}, \ldots, \psi_{(N)}$ .
- **③** Check if  $-\sum_{n=1}^{\lfloor \alpha N \rfloor} \psi_{(n)}$  is large enough to change your conclusions.

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How to compute the  $\psi_n$ 's? And how accurate is the approximation?

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Recall that  $\sum_{n=1}^{N} w_n G(\hat{\theta}(w), d_n) = 0_P$  for all w near  $\vec{1}$ .

 $\Rightarrow$  By the **implicit function theorem**, we can write  $\frac{\partial \hat{\theta}(w)}{\partial w_n}\Big|_{w=\vec{1}}$  as a linear system involving  $G(\cdot, \cdot)$  and its derivatives.

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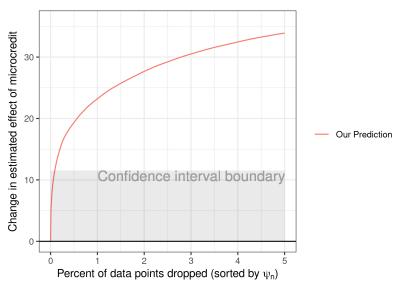
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- $\Rightarrow$  The  $\psi_n$  are automatically computable from  $\hat{\theta}(\vec{1})$  and software implementations of  $G(\cdot, \cdot)$  and  $\phi(\cdot)$  using **automatic differentiation**.

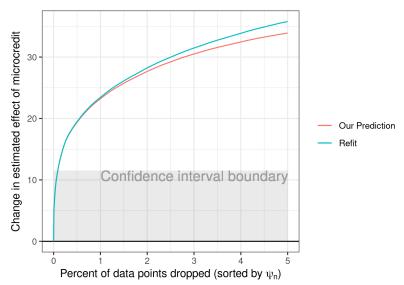
```
> import jax
> import jax.numpy as np
> def phi(theta):
> ... computations using np and theta ...
> return value
>
> # Exact gradient of phi (first term in the chain rule above):
> jax.grad(phi)(theta_opt)
```

See rgiordan/vittles on github.

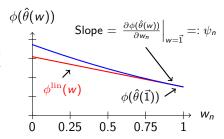
Checking the approximation for Mexico microcredit.



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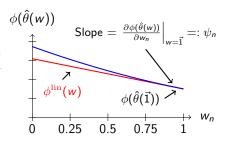
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#### But you don't need to rely on the theory to show non-robustness!

Our method returns which points to drop. **Re-running once** without those points provides an **exact lower bound** on the worst-case sensitivity.

## Selected experimental results.

Original estimate (SE)	Refit estimate (SE)	Observations dropped
-4.549 (5.879)	7.030 (2.549)*	15 = 0.09%

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0.029 (0.005)*	-0.009 (0.004)*	224 = 0.96%

Table: Medicaid profit results (N=23361) [Finkelstein et al., 2012]

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#### Contrast with sampling variability.

A 95% CI is given by  $\phi(\hat{\theta}(\vec{1})) \pm \frac{1.96}{\sqrt{N}} \hat{\sigma}_{\phi}$ . We reject  $\phi(\hat{\theta}(\vec{1})) + \Delta$  when

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*Proof:* P-hacked effect sizes are of the order  $\frac{1.96}{\sqrt{N}}\hat{\sigma}_{\phi}$ .

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Robust to gross errors: ("gross error robustness")

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- Dropping non-robustness is not driven by misspecification. *Proof:* Small  $\Delta$  are dropping non-robust irrespective of specification.
- Gross outliers primarily affect dropping robustness through  $\hat{\sigma}_{\phi}$ .

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Dropping robustness should augment other forms of robustness.

#### How to make an analysis less sensitive?

Robust to data dropping: ("dropping robustness")

SNR := 
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To achieve dropping robustness, reduce  $\hat{\sigma}_{\phi}$  and / or increase  $\Delta$ . *Proof:* Across typical distributions,  $\hat{\mathscr{S}}_{\alpha}$  varies little. (Details in paper.)

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In the Mexico microcredit example,

$$\hat{\sigma}_{\phi} = 757.8$$
  $\phi(\hat{\theta}(\vec{1})) = -4.55$   $N = 16,560$ 

The study overcame a very low signal to noise ratio with a very large N.

This (canonical) response to low signal to noise ratio — to gather more data — produces small SEs, but cannot produce dropping robustness.

#### Conclusion

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- We can quickly and automatically find an approximate influential set which is accurate for small sets.
- Data dropping robustness is principally determined by the signal to noise ratio, and captures sensitivity distinct from sampling and gross error sensitivity.

#### Links and references

Tamara Broderick, Ryan Giordano, Rachael Meager (alphabetical authors) "An Automatic Finite-Sample Robustness Metric: Can Dropping a Little Data Change Conclusions?"

https://arxiv.org/abs/2011.14999

Select blog posts with more details: https://rgiordan.github.io

- Data dropping sensitivity overcomes p-hacking
- Collinearity in OLS after dropping
- Influence functions and sums
- Connections to the bootstrap

#### Related software on github:

- rgiordan/zaminfluence (for R)
- rgiordan/vittles (for Python)

#### Some of my work on other forms of robustness:

- Prior sensitivity in Bayesian nonparametrics [Giordano et al., 2021]
- Approximate cross-validation (and other reweightings) [Giordano et al., 2019b,a]
- Covariances and prior sensitivity for mean field VB [Giordano et al., 2015, 2018]
- Model sensitivity of MCMC output [Giordano et al., 2018]
- Frequentist variances of MCMC posteriors (in progress)

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# Extra slides

#### A simulation

For N = 5,000 data points, compute the OLS estimator from:

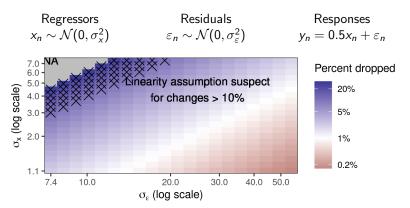


Figure: The approximate perturbation inducing proportion at differing values of  $\sigma_x$  and  $\sigma_\varepsilon$ . Red colors indicate datasets whose sign can is predicted to change when dropping less than 1% of datapoints. The grey areas indicate  $\hat{\Psi}_\alpha = \text{NA}$ , a failure of the linear approximation to locate any way to change the sign.

The present work is based on the *empirical influence function*. Consider:

- True, unknown distribution function  $F_{\infty}(x) = p(X \le x)$
- Empirical distribution function  $\hat{F}(x) = \frac{1}{N} \sum_{n=1}^{N} \mathbb{I}(x_n \leq x)$
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We estimate with  $T(F_{\infty})$  with  $T(\hat{F})$ . Sample means are an example:

$$T(F) := \int x F(\mathrm{d}x).$$

Z-estimators are, too:

$$T(F) := \theta$$
 such that  $\int G(\theta, x)F(dx) = 0$ .

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Form an (infinite-dimensional) Taylor series expansion at some  $F_0$ :

$$T(F) = T(F_0) + T'(F_0)(F - F_0) + \text{residual}.$$

When the derivative operator takes the form of an integral

$$T'(F_0)\Delta = \int \psi(x; F_0)\Delta(\mathrm{d}x)$$

then  $\psi(x; F_0)$  is known as the *influence function*.

Where to form the expansion? There are at least two reasonable choices:

- The limiting influence function  $\psi(x, F_{\infty})$
- The empirical influence function  $\psi(x, \hat{F})$

- The limiting influence function (LIF)  $\psi(x, F_{\infty})$ 
  - Used in a lot of classical statistics [Mises, 1947, Huber, 1981, Hampel, 1986, Bickel et al., 1993]
  - Unobserved, asymptotic
  - Requires careful functional analysis [Reeds, 1976]
- The empirical influence function (EIF)  $\psi(x, \hat{F})$ 
  - The basis of the present work (also [Giordano et al., 2019b,a])
  - Computable, finite-sample
  - Requires only finite-dimensional calculus

Typically the semantics of the EIF derive from study of the LIF.

Example: 
$$\frac{1}{N} \sum_{n=1}^{N} (N\psi_n)^2 \approx \operatorname{Var}\left(\sqrt{N}\phi(\hat{\theta})\right)$$
.

But the EIF measures what happens when you perturb the data at hand.

Other data perturbations will admit an analysis similar to ours!

#### Local robustness

The present work is an application of *local robustness*. Consider:

- Model parameter  $\lambda$  (e.g., data weights  $\lambda = w$ )
- Set of plausible models  $\mathcal{S}_{\lambda}$  (e.g.  $\mathcal{S}_{\lambda} = W_{\alpha}$ )
- Estimator  $\hat{\theta}(x, \lambda)$  for data x and  $\lambda \in \mathcal{S}_{\lambda}$  (e.g. a Z-estimator)

Global robustness: 
$$\left(\inf_{\lambda \in \mathcal{S}_{\lambda}} \hat{\theta}(x,\lambda), \sup_{\lambda \in \mathcal{S}_{\lambda}} \hat{\theta}(x,\lambda)\right)$$
 (Hard in general!)

Local robustness:  $\left(\inf_{\lambda \in \mathcal{S}_{\lambda}} \hat{\theta}^{lin}(x,\lambda), \sup_{\lambda \in \mathcal{S}_{\lambda}} \hat{\theta}^{lin}(x,\lambda)\right)$ 
...where  $\hat{\theta}^{lin}(x,\lambda) := \hat{\theta}^{lin}(x,\lambda_0) + \left.\frac{\partial \hat{\theta}^{lin}(x,\lambda)}{\partial \lambda}\right|_{\lambda_0} (\lambda - \lambda_0)$ .

#### Many variants are possible!

- Cross-validation [Giordano et al., 2019b]
- Prior sensitivity in Bayesian nonparametrics [Giordano et al., 2021]
- Model sensitivity of MCMC output [Giordano et al., 2018]
- Frequentist variances of MCMC posteriors (in progress)