# Variational Methods for Latent Variable Problems (part 2)

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#### **Outline**

#### Outline for today:

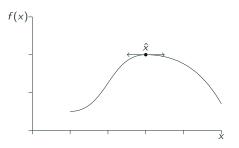
- What counts as variational inference?
- Kullback-Leibler (KL) divergence and "standard" variational inference
- The classical EM algorithm as a special case of variational inference
- Variational inference as a generalization of the EM algorithm
- A quick and incomplete sketch of further topics in variational inference

#### What counts as variational inference?

Lots of very different procedures go by the name "variational inference." I propose an (idosyncratic) enompassing definition based on the use cases and the name:

#### Variational inference is inference using optimization.

Think "calculus of variations:" an optimum  $\hat{x} = \operatorname*{argmax} f(x)$  is characterized by  $df/dx|_{\hat{x}} = 0$ , i.e. where small variations in  $\hat{x}$  result in no changes to the value of  $f(\hat{x})$ .



#### By this definition,

- The maximum likelihood estimator (MLE) is VI.
- The Laplace approximation to a Bayesian posterior is VI.
- Markov chain Monte Carlo (MCMC) is not VI.

#### What counts as variational inference?

A more common definition of VI is the following.

Suppose we have a random variable  $\xi$  and a distribution  $\mathfrak{p}(\xi)$  that we want to know.

Let y denote data and  $\theta$  a parameter. Examples:

- The variable is  $\theta$ , and we wish to know the posterior  $\mathfrak{p}(\theta|y)$  (Bayes)
- The variable is y, and we wish to know  $\mathfrak{p}(y)$  (MLE)
- The variable is y, and we wish to know the map  $\theta \mapsto \mathfrak{p}(y|\theta) = \int p(y,z|\theta)dz$  (marginal MLE)

Let  $\mathcal{Q}$  be some class of distributions which may or may not contain  $\mathfrak{p}(\xi)$ .

Variational inference finds the distribution in  $\mathcal Q$  closest to  $\mathfrak p$  according to some measure of "divergence" between distributions:

$$\mathfrak{q}^* = \operatorname*{argmin}_{\mathfrak{q} \in \mathcal{Q}} D(\mathfrak{q}, \mathfrak{p}).$$

The most common choice of "divergence" is the **Kullback-Leibler** (KL) divergence, though other choices are possible (e.g. Li and Turner [2016], Liu and Wang [2016], Ambrogioni et al. [2018]).

#### KL divergence

The KL divergence is defined as:

$$\mathrm{KL}\left(\mathfrak{q}||\mathfrak{p}
ight) := \underset{\mathfrak{q}(\xi)}{\mathbb{E}}\left[\log\mathfrak{q}(\xi)\right] - \underset{\mathfrak{q}(\xi)}{\mathbb{E}}\left[\log\mathfrak{p}(\xi)\right]$$

Some points to be aware of:

- $KL(\mathfrak{q}||\mathfrak{p}) > 0$
- $KL(\mathfrak{g}||\mathfrak{p}) = 0 \Rightarrow \mathfrak{p} = \mathfrak{q}$
- $KL(\mathfrak{q}||\mathfrak{p}) \neq KL(\mathfrak{p}||\mathfrak{q})$
- $\mathrm{KL}(\mathfrak{q}||\mathfrak{p})$  is a "strict" measure of closeness [Gibbs and Su, 2002]

Why use KL divergence?

**Phony answer:** The KL divergence has an information theoretic interpretation [Kullback and Leibler, 1951].

Real answer: Mathematical convenience (normalizing constants pop out).

**Example:** the MLE minimizes KL divergence. Suppose that  $x_n \stackrel{iid}{\sim} \mathfrak{p}(\cdot)$ , and  $\mathfrak{q}(\cdot|\theta) \in \mathcal{Q}$  is a (possibly misspecified) parameteric family of data distributions. Then

$$\begin{split} \hat{\theta} := & \operatorname*{argmin}_{\theta} \operatorname{KL}\left(\mathfrak{p}||\mathfrak{q}\right) = \operatorname*{argmin}_{\theta} \left( - \underset{\mathfrak{p}(x_1)}{\mathbb{E}} \left[\log \mathfrak{q}(x_1|\theta)\right] + \underset{\mathfrak{p}(x_1)}{\mathbb{E}} \left[\log \mathfrak{p}(x_1)\right] \right) \\ = & \operatorname*{argmax}_{\theta} \underset{\mathfrak{p}(x_1)}{\mathbb{E}} \left[\log \mathfrak{q}(x_1|\theta)\right] \approx \operatorname*{argmax}_{\theta} \frac{1}{N} \sum_{n=1}^{N} \log \mathfrak{q}(x_n|\theta). \end{split}$$

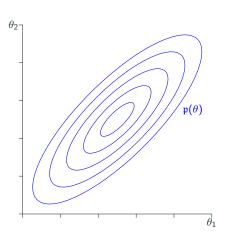
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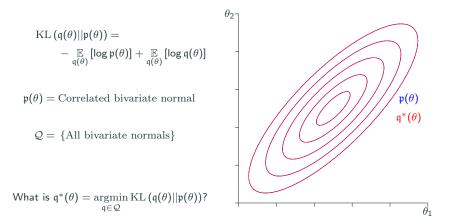
$$\begin{split} \operatorname{KL}\left(\mathfrak{q}(\theta)||\mathfrak{p}(\theta)\right) &= \\ &- \underset{\mathfrak{q}(\theta)}{\mathbb{E}} \left[\log \mathfrak{p}(\theta)\right] + \underset{\mathfrak{q}(\theta)}{\mathbb{E}} \left[\log \mathfrak{q}(\theta)\right] \end{split}$$

 $\mathfrak{p}(\theta) = \text{Correlated bivariate normal}$ 

 $\mathcal{Q} = \, \{ \text{All bivariate normals} \}$ 

What is  $q^*(\theta) = \operatorname*{argmin}_{q \in \mathcal{Q}} \mathrm{KL}\left(q(\theta)||p(\theta)\right)$ ?





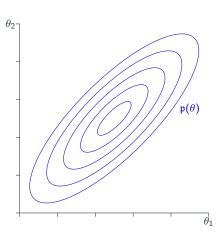
Sufficiently expressive families recover the target distribution.

$$\begin{split} \mathrm{KL}\left(\mathfrak{q}(\theta)||\mathfrak{p}(\theta)\right) &= \\ &- \underset{\mathfrak{q}(\theta)}{\mathbb{E}}\left[\log\mathfrak{p}(\theta)\right] + \underset{\mathfrak{q}(\theta)}{\mathbb{E}}\left[\log\mathfrak{q}(\theta)\right] \end{split}$$

 $\mathfrak{p}(\theta)$  = Correlated bivariate normal

 $Q = \{Independent bivariate normals\}$ 

What is  $q^*(\theta) = \operatorname*{argmin}_{q \in \mathcal{Q}} \mathrm{KL}\left(q(\theta)||\mathfrak{p}(\theta)\right)$ ?

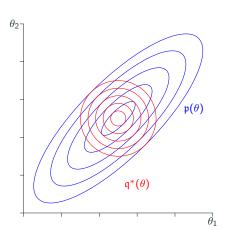


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 $Q = \{Independent bivariate normals\}$ 

What is 
$$q^*(\theta) = \underset{q \in \mathcal{Q}}{\operatorname{argmin}} \operatorname{KL}(q(\theta)||p(\theta))$$
?



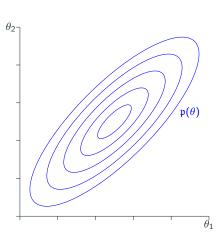
KL minimizers "fit inside" the second argument.

$$\begin{split} \mathrm{KL}\left(\mathfrak{p}(\theta)||\mathfrak{q}(\theta)\right) &= \\ &- \underset{\mathfrak{p}(\theta)}{\mathbb{E}} \left[\log \mathfrak{q}(\theta)\right] + \underset{\mathfrak{p}(\theta)}{\mathbb{E}} \left[\log \mathfrak{p}(\theta)\right] \end{split}$$

 $\mathfrak{p}(\theta)$  = Correlated bivariate normal

 $Q = \{Independent \ bivariate \ normals\}$ 

What is  $q^*(\theta) = \operatorname*{argmin}_{q \in \mathcal{Q}} \mathrm{KL}\left(\mathfrak{p}(\theta)||q(\theta)\right)$ ?

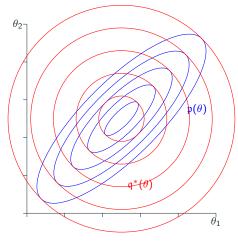


$$\begin{split} \mathrm{KL}\left(\mathfrak{p}(\theta)||\mathfrak{q}(\theta)\right) &= \\ &- \underset{\mathfrak{p}(\theta)}{\mathbb{E}}\left[\log \mathfrak{q}(\theta)\right] + \underset{\mathfrak{p}(\theta)}{\mathbb{E}}\left[\log \mathfrak{p}(\theta)\right] \end{split}$$

 $\mathfrak{p}(\theta)$  = Correlated bivariate normal

 $Q = \{Independent \ bivariate \ normals\}$ 

What is 
$$\mathfrak{q}^*(\theta) = \operatorname*{argmin}_{\mathfrak{q} \in \mathcal{Q}} \mathrm{KL}\left(\mathfrak{p}(\theta)||\mathfrak{q}(\theta)\right)$$
?



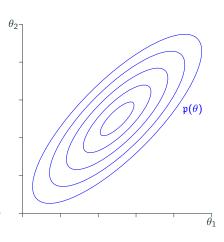
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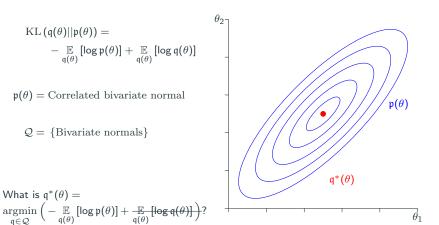
$$\begin{split} \operatorname{KL}\left(\mathfrak{q}(\theta)||\mathfrak{p}(\theta)\right) &= \\ &- \underset{\mathfrak{q}(\theta)}{\mathbb{E}}\left[\log\mathfrak{p}(\theta)\right] + \underset{\mathfrak{q}(\theta)}{\mathbb{E}}\left[\log\mathfrak{q}(\theta)\right] \end{split}$$

 $\mathfrak{p}(\theta) = \text{Correlated bivariate normal}$ 

 $\mathcal{Q} = \, \{ \text{Bivariate normals} \}$ 

What is 
$$q^*(\theta) = \underset{q \in \mathcal{Q}}{\operatorname{argmin}} \left( - \underset{q(\theta)}{\mathbb{E}} [\log \mathfrak{p}(\theta)] + \underset{q(\theta)}{\overline{\mathbb{E}}} [\log \mathfrak{q}(\theta)] \right)?$$





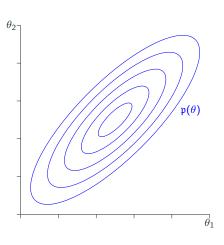
Without the entropy, the KL minimizer concentrates on the maximum of  $\log p(\theta)$ .

$$\begin{split} \operatorname{KL}\left(\mathfrak{q}(\theta)||\mathfrak{p}(\theta)\right) &= \\ &- \underset{\mathfrak{q}(\theta)}{\mathbb{E}}\left[\log\mathfrak{p}(\theta)\right] + \underset{\mathfrak{q}(\theta)}{\mathbb{E}}\left[\log\mathfrak{q}(\theta)\right] \end{split}$$

 $\mathfrak{p}(\theta) = \text{Correlated bivariate normal}$ 

 $\mathcal{Q} = \, \{ \text{Bivariate normals} \}$ 

What is 
$$q^*(\theta) = \underset{q \in \mathcal{Q}}{\operatorname{argmin}} \left( -\frac{\mathbb{E}\left[\log p(\theta)\right]}{q(\theta)} + \underset{q(\theta)}{\mathbb{E}} \left[\log q(\theta)\right] \right)$$
?



$$\begin{aligned} \operatorname{KL}\left(\mathfrak{q}(\theta)||\mathfrak{p}(\theta)\right) &= \\ &- \underset{\mathfrak{q}(\theta)}{\mathbb{E}}\left[\log \mathfrak{p}(\theta)\right] + \underset{\mathfrak{q}(\theta)}{\mathbb{E}}\left[\log \mathfrak{q}(\theta)\right] \\ \mathfrak{p}(\theta) &= \operatorname{Correlated bivariate normal} \end{aligned}$$
 
$$\mathcal{Q} = \left\{ \begin{aligned} \operatorname{Bivariate normals} \right\} \end{aligned}$$
 What is  $\mathfrak{q}^*(\theta) = \underset{\mathfrak{q} \in \mathcal{Q}}{\operatorname{argmin}}\left( - \underset{\mathfrak{q}(\theta)}{\underbrace{\mathbb{E}}\left[\log \mathfrak{p}(\theta)\right]} + \underset{\mathfrak{q}(\theta)}{\mathbb{E}}\left[\log \mathfrak{q}(\theta)\right] \right) \end{aligned}$ 

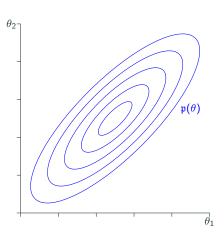
Without  $\log \mathfrak{p}(\theta)$ , the KL minimizer is infinitely dispersed.

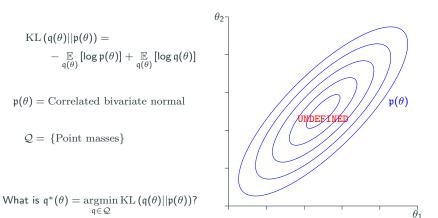
$$\begin{split} \mathrm{KL}\left(\mathfrak{q}(\theta)||\mathfrak{p}(\theta)\right) &= \\ &- \underset{\mathfrak{q}(\theta)}{\mathbb{E}}\left[\log\mathfrak{p}(\theta)\right] + \underset{\mathfrak{q}(\theta)}{\mathbb{E}}\left[\log\mathfrak{q}(\theta)\right] \end{split}$$

 $\mathfrak{p}(\theta)$  = Correlated bivariate normal

$$\mathcal{Q} = \{ \text{Point masses} \}$$

What is 
$$q^*(\theta) = \operatorname*{argmin}_{q \in \mathcal{Q}} \mathrm{KL}\left(q(\theta)||p(\theta)\right)$$
?





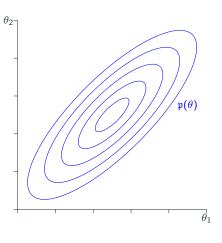
Without a common dominating measure, the KL divergence is undefined.

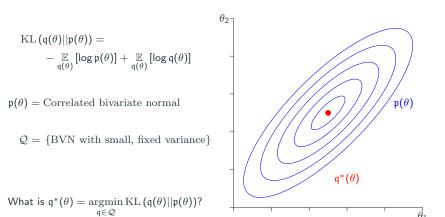
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 $\mathfrak{p}(\theta) = \text{Correlated bivariate normal}$ 

 $\mathcal{Q} = \, \{ \text{BVN with small, fixed variance} \}$ 

What is 
$$q^*(\theta) = \operatorname*{argmin}_{q \in \mathcal{Q}} \mathrm{KL}\left(q(\theta)||p(\theta)\right)$$
?





Sufficently concentrated distributions with constant entropy act like a point mass at the maximum of  $\log p(\theta)$ .

#### **Conclusions**

- Luca Ambrogioni, Umut Güçlü, Yağmur Güçlütürk, Max Hinne, Eric Maris, and Marcel AJ van Gerven. Wasserstein variational inference. arXiv preprint arXiv:1805.11284, 2018.
- Alison L Gibbs and Francis Edward Su. On choosing and bounding probability metrics. *International statistical review*, 70(3):419–435, 2002.
- Solomon Kullback and Richard A Leibler. On information and sufficiency. The annals of mathematical statistics, 22 (1):79–86, 1951.
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- Qiang Liu and Dilin Wang. Stein variational gradient descent: A general purpose bayesian inference algorithm. arXiv preprint arXiv:1608.04471, 2016.