

An Automatic Finite-Sample Robustness Metric: Can Dropping a Little Data Make a Big Difference?

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Dropping data: Mexico Microcredit

Example: Angelucci et al. [2015], a randomized controlled trial study of the efficacy of microcredit in Mexico based on 16,560 data points. The variable “Beta” estimates the effect of microcredit in US dollars.

	Beta (SE)
Original result	-4.55 (5.88)

The original conclusion: No evidence that microcredit is effective...

⇒ Standard errors can be inadequate summaries of data sensitivity!

Cannot find influential subsets by brute force! $\binom{16,560}{15} \approx 1.5 \cdot 10^{51}$

We provide a fast, automatic tool to approximately identify the most influential set of points.

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...can be reversed by dropping less than 0.1% of the data.

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We provide a fast, automatic tool to approximately identify the most influential set of points.

- Why and when might you care about sensitivity to data dropping?
- How does our approximation work, and when is it accurate?
(A formalization of the problem and the class of estimators we study.)
- Examine real-life examples of analyses: some sensitive, some not.
(The results may defy your intuition.)
- What kinds of analyses are sensitive to data dropping?
(Including comparison to standard errors and gross-error robustness.)

Dropping data: Motivation

Would you be concerned if you could **reverse your conclusion** by removing a **small proportion** of your data?

Not always! But sometimes, surely yes, especially when you want to **generalize to unseen, systematically different populations**.

Suppose you have a farm, and want to know whether your average yield is > 170 bushels per acre. At harvest, you measure 200 bushels per acre.

- Scenario one: > 170 bushels per acre means you make a profit.
 - Don't care about sensitivity to small subsets.
- Scenario two: Want to recommend methods to a distant friend.
 - Might care about sensitivity to small subsets!

Specifically, often in statistical applications:

- Policy population is different from analyzed population,
- Small fractions of data are missing not-at-random,
- We report a convenient summary (e.g. mean) of a complex effect.

Formalizing the question.

Ordinary least squares

A data point d_n has regressors x_n and response y_n : $d_n = (x_n, y_n)$.

The estimator $\hat{\theta} \in \mathbb{R}^p$ satisfies:

$$\hat{\theta} := \arg \min_{\theta} \frac{1}{2} \sum_{n=1}^N (y_n - \theta^T x_n)^2$$

$$\Leftrightarrow \sum_{n=1}^N (y_n - \hat{\theta}^T x_n) x_n = 0.$$

Make a qualitative decision using:

- A particular component: $\hat{\theta}_k$
- The end of a confidence interval: $\hat{\theta}_k + \frac{1.96}{\sqrt{N}} \hat{\sigma}(\hat{\theta})$

Z-estimators

We observe N data points d_1, \dots, d_N (in any domain).

The estimator $\hat{\theta} \in \mathbb{R}^p$ satisfies:

$$\sum_{n=1}^N G(\hat{\theta}, d_n) = 0_p.$$

$G(\cdot, d_n)$ is “nice,” \mathbb{R}^p -valued.
E.g. OLS, MLE, VB, IV &c.

Make a qualitative decision using $\phi(\hat{\theta})$ for a smooth, real-valued ϕ .

(WLOG try to increase $\phi(\hat{\theta})$.)

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- There are $\binom{N}{\lfloor \alpha N \rfloor}$ sets to check. (Huge even for $\alpha \ll 1$.)
- Evaluating $\hat{\theta}$ re-solving the estimating equation.
 - E.g., re-computing the OLS estimator.
 - Other examples are even harder (VB, machine learning)

Idea: Smoothly approximate the effect of leaving out points.

We have N data points d_1, \dots, d_N , a quantity of interest $\phi(\cdot)$, and

$$\sum_{n=1}^N G(\hat{\theta}, d_n) = 0_P \quad .$$

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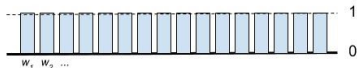
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Original weights: $\vec{1} = (1, \dots, 1)$



Leave points out by setting their elements of w to zero.

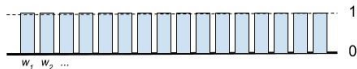


The map $w \mapsto \phi(\hat{\theta}(w))$ is well-defined even for continuous weights.

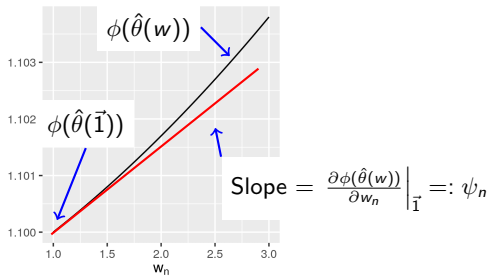
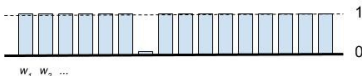
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The values $N\psi_n$ are the **empirical influence function** [Hampel, 1986]. We call ψ_n an “influence scores.”

We can use ψ_n to form a Taylor series approximation:

$$\phi(\hat{\theta}(w)) \approx \phi^{\text{lin}}(w) := \phi(\hat{\theta}(\vec{1})) + \sum_{n=1}^N \psi_n (w_n - 1)$$

Taylor series approximation.

Problem: How much can you change $\phi(\hat{\theta}(w))$ dropping $\lfloor \alpha N \rfloor$ points?
Combinatorially hard by brute force!

Approximate Problem: How much can you change $\phi^{\text{lin}}(\hat{\theta}(w))$ dropping $\lfloor \alpha N \rfloor$ points? **Easy!**

$$\phi^{\text{lin}}(w) := \phi(\hat{\theta}(\vec{1})) + \sum_{n=1}^N \psi_n(w_n - 1)$$

Dropped points have $w_n - 1 = -1$. Kept points have $w_n - 1 = 0$
 \Rightarrow The most influential points for $\phi^{\text{lin}}(w)$ have the most negative ψ_n .

Procedure: (see rgiordan/zaminfluence on github)

- 1 Compute your original estimator $\hat{\theta}(\vec{1})$.
- 2 Compute and sort the influence scores $\psi_{(1)}, \dots, \psi_{(N)}$.
- 3 Worry if $-\sum_{n=1}^{\lfloor \alpha N \rfloor} \psi_{(n)}$ is large enough to change your conclusions.

How to compute the ψ_n 's? And how accurate is the approximation?

How to compute the influence scores?

How can we compute the influence scores $\psi_n = \left. \frac{\partial \phi(\hat{\theta}(w))}{\partial w_n} \right|_{\vec{1}}$?

By the **chain rule**, $\psi_n = \left. \frac{\partial \phi(\theta)}{\partial \theta} \right|_{\hat{\theta}(\vec{1})} \left. \frac{\partial \hat{\theta}(w)}{\partial w_n} \right|_{\vec{1}}$.

Recall that $\sum_{n=1}^N w_n G(\hat{\theta}(w), d_n) = 0_P$ for all w near $\vec{1}$.

\Rightarrow By the **implicit function theorem**, we can write $\left. \frac{\partial \hat{\theta}(w)}{\partial w_n} \right|_{\vec{1}}$ as a linear system involving $G(\cdot, \cdot)$ and its derivatives.

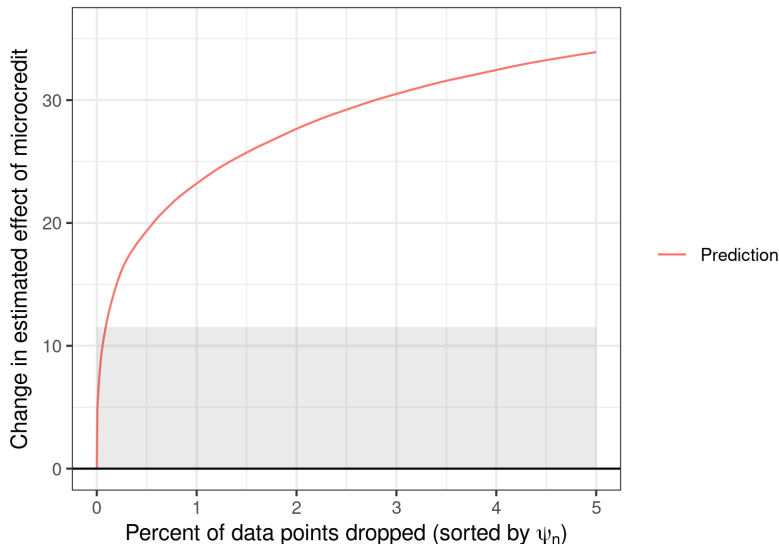
\Rightarrow The ψ_n are automatically computable from $\hat{\theta}(\vec{1})$ and software implementations of $G(\cdot, \cdot)$ and $\phi(\cdot)$ using **automatic differentiation**.

```
import jax
import jax.numpy as np
def phi(theta):
    ... computations using np and theta ...
    return value

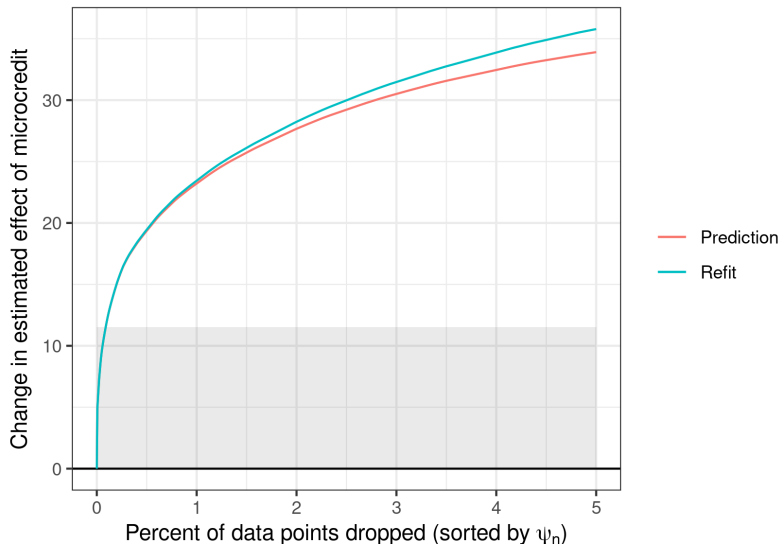
# Exact gradient of phi (1st term in the chain rule):
jax.grad(phi)(theta_opt)
```

See [rgiordan/vittles](#) on github.

How accurate is the approximation?

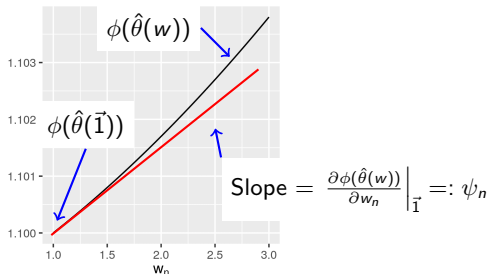


How accurate is the approximation?



How accurate is the approximation?

By controlling the curvature, we can control the error in the linear approximation.



We provide **finite-sample theory** [Giordano et al., 2019] showing that

$$\left| \phi(\hat{\theta}(w)) - \phi^{\text{lin}}(w) \right| = O \left(\left\| \frac{1}{N}(w - \vec{1}) \right\|_2^2 \right) = O(\alpha) \text{ as } \alpha \rightarrow 0.$$

But you don't need to rely on the theory!

Our method returns which points to drop. **Re-running once** without those points provides an **exact lower bound** on the worst-case sensitivity.

Selected experimental results.

Original estimate (SE)	Refit estimate (SE)	Observations dropped
-4.549 (5.879)	7.030 (2.549)*	15 = 0.09%

Table: Microcredit Mexico results [Angelucci et al., 2015].

A * indicates statistical significance at the 95% level.

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Original estimate (SE)	Refit estimate (SE)	Observations dropped
0.029 (0.005)*	-0.009 (0.004)*	224 = 0.96%

Table: Medicaid profit results [Finkelstein et al., 2012]

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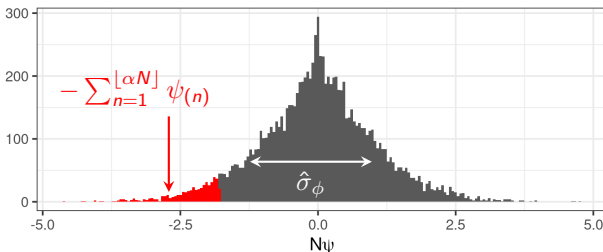
What makes an analysis sensitive?

We are “sensitive to data dropping” if, for some Δ large enough to change conclusions, $\exists w^*$ dropping $\lfloor \alpha N \rfloor$ points such that

$$\text{“Signal”} := \Delta < \phi^{\text{lin}}(w^*) - \phi(\hat{\theta}(\vec{1})) = - \sum_{n=1}^{\lfloor \alpha N \rfloor} \psi_{(n)} =: \hat{\sigma}_{\phi} \hat{\mathcal{T}}_{\alpha}$$

- The “noise” $\hat{\sigma}_{\phi}^2 \rightarrow \text{Var}(\sqrt{N}\phi)$ (“sandwich” variance estimator)
- The “shape” $\hat{\mathcal{T}}_{\alpha} := \frac{-\sum_{n=1}^{\lfloor \alpha N \rfloor} \psi_{(n)}}{\hat{\sigma}_{\phi}} \rightarrow \text{nonzero constant} \leq \sqrt{\alpha(1-\alpha)}$

Influence score histogram (N = 10000, $\alpha = 0.05$)



Example.

α := Proportion of points to drop

Δ := Signal (difference large enough to change conclusions)

$\hat{\sigma}_\phi$:= Noise (consistent estimator of $\text{Var}(\sqrt{N}\phi)$)

$\hat{\mathcal{T}}_\alpha$:= Shape (bounded by $\sqrt{\alpha(1-\alpha)}$ and given by $N\psi_n$ tail shape)

Sensitive to data dropping if:

$$\phi^{\text{lin}}(w^*) - \phi(\hat{\theta}(\vec{1})) = \hat{\sigma}_\phi \hat{\mathcal{T}}_\alpha \geq \Delta \quad \Leftrightarrow \quad \frac{\Delta}{\hat{\sigma}_\phi} \leq \hat{\mathcal{T}}_\alpha.$$

The **signal to noise ratio** $\frac{\Delta}{\hat{\sigma}_\phi}$ determines sensitivity to data dropping.

Contrast with standard errors. A 95% CI is given by $\phi(\hat{\theta}(\vec{1})) \pm \frac{1.96}{\sqrt{N}} \hat{\sigma}_\phi$.

We fail to reject the value $\phi(\hat{\theta}(\vec{1})) + \Delta$ when

$$\phi(\hat{\theta}(\vec{1})) + \Delta \leq \phi(\hat{\theta}(\vec{1})) + \frac{1.96}{\sqrt{N}} \hat{\sigma}_\phi \quad \Leftrightarrow \quad \frac{\Delta}{\hat{\sigma}_\phi} \leq \frac{1.96}{\sqrt{N}}.$$

Robust to data dropping:
("dropping robustness")

$$\text{SNR} = \frac{\Delta}{\hat{\sigma}_\phi} > \hat{\mathcal{J}}_\alpha$$

Robust to sampling variation:
("sampling robustness")

$$\text{SNR} = \frac{\Delta}{\hat{\sigma}_\phi} > \frac{1.96}{\sqrt{N}} \hat{\sigma}_\phi$$

-
- **Dropping robustness \neq sampling robustness in general.**

Proof: $\hat{\mathcal{J}}_\alpha \neq \frac{1.96}{\sqrt{N}} \hat{\sigma}_\phi$.

- **When the SNR is small, sufficiently large N produces sampling robustness, but not necessarily dropping robustness.**

Proof: $\frac{1.96}{\sqrt{N}} \hat{\sigma}_\phi \rightarrow 0$, but $\hat{\mathcal{J}}_\alpha \rightarrow$ a nonzero constant.

- **Statistical insignificance is dropping non-robust for large N .**

Proof: Insignificance means $|\phi(\hat{\theta}(\vec{1}))| \leq \frac{1.96}{\sqrt{N}} \hat{\sigma}_\phi$.

\Rightarrow A result can be made significant by a change of no more than $\frac{1.96}{\sqrt{N}} \hat{\sigma}_\phi$.

\Rightarrow The SNR for a conclusion of "insignificance" is $\frac{\Delta}{\hat{\sigma}_\phi} \leq \frac{1.96}{\sqrt{N}} \rightarrow 0 \leq \hat{\mathcal{J}}_\alpha$.

Corollaries.

Robust to data dropping:
("dropping robustness")

$$\text{SNR} = \frac{\Delta}{\hat{\sigma}_\phi} > \hat{\mathcal{I}}_\alpha$$

Robust to gross errors:
("gross error robustness")

Gross outliers cannot produce
arbitrarily large changes to ϕ .

- **Dropping non-robustness is not driven by misspecification.**

Proof: Small Δ are dropping non-robust irrespective of specification.

- **Gross outliers primarily affect dropping robustness through $\hat{\sigma}_\phi$.**

Proof: For a fixed $\hat{\sigma}_\phi$, outliers decrease $\hat{\mathcal{I}}_\alpha$. (Details in paper.)

- **To achieve dropping robustness, reduce $\hat{\sigma}_\phi$ and / or increase Δ .**

Proof: Across typical distributions, $\hat{\mathcal{I}}_\alpha$ varies little. (Details in paper.)

- You may be concerned if you could reverse your conclusion by removing a small proportion of your data.

Conclusion

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- We can quickly and automatically find an approximate influential set which is accurate for small sets.

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- We can quickly and automatically find an approximate influential set which is accurate for small sets.
- Data dropping robustness is principally determined by the signal to noise ratio, and captures sensitivity distinct from sampling and gross error sensitivity.

Links and references

Tamara Broderick, Ryan Giordano, Rachael Meager (alphabetical authors)
“An Automatic Finite-Sample Robustness Metric: Can Dropping a Little Data Change Conclusions?”

<https://arxiv.org/abs/2011.14999>

Blog posts with more details:

- Colinearity in OLS after dropping
 - Connections to the bootstrap
 - Data dropping sensitivity overcomes p-hacking
 - When a norm is the quantity of interest
-

Related software on github:

- [rgiordan/zaminfluence](#) (for R)
 - [rgiordan/vittles](#) (for Python)
-

Some of my work on other forms of robustness:

- Prior sensitivity in Bayesian nonparametrics [Giordano et al., 2021]
- Model sensitivity of MCMC output [Giordano et al., 2018]
- Cross-validation [Giordano et al., 2019]
- Frequentist variances of MCMC posteriors (in progress)

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