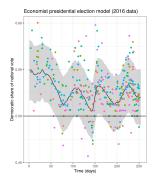
Approximate data deletion and replication with the Bayesian influence function

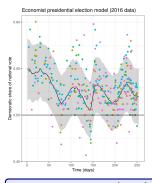
Ryan Giordano (rgiordano berkeley.edu, UC Berkeley), Tamara Broderick (MIT)

MIT Robustness and Influence Functions Workshop July 2024



A time series model to predict the 2016 US presidential election outcome from polling data.

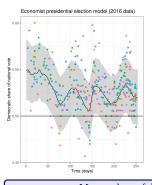
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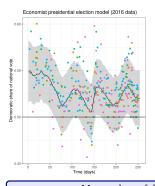


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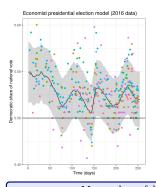
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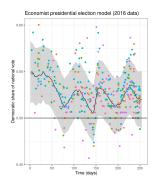
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Practically: We compute Markov chain Monte Carlo (MCMC) draws from $p(\theta|X)$.

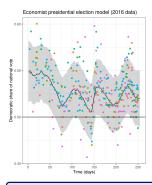


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- $X = x_1, ..., x_N =$ Polling data (N = 361).
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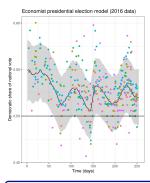
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- · We are interested re-sampling for this election, not a hypothetical future election

,

Results

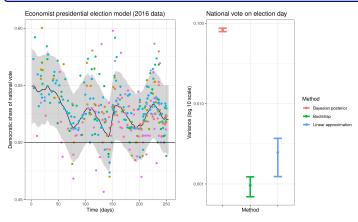
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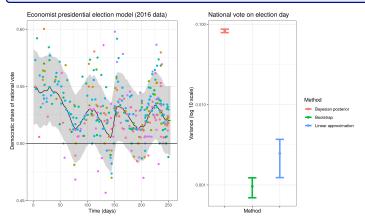


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Compute time for 100 bootstraps:

Compute time for the linear approximation: (But note the approximation has some error)

Seconds

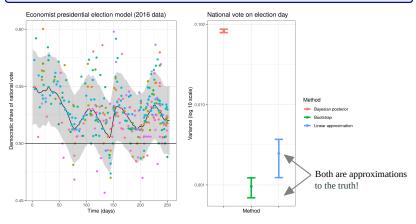
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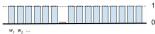
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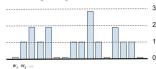
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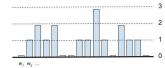
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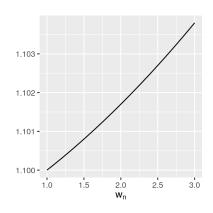


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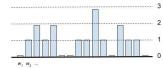
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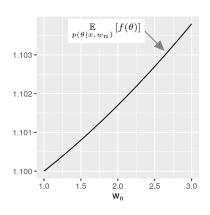


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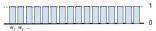




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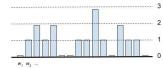
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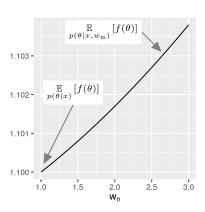


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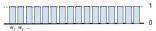




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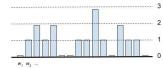
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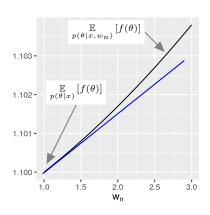


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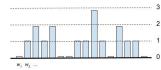
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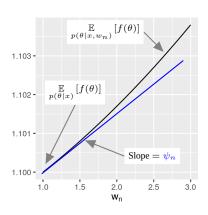


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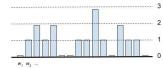
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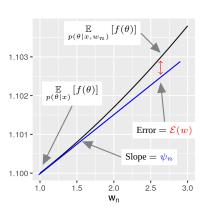


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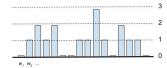
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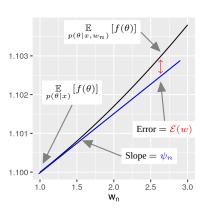


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The re-scaled slope $N\psi_n$ is known as the "influence function" at data point x_n .

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$$\approx \underset{p(w)}{\operatorname{Var}} \left(\underset{p(\theta|x)}{\mathbb{E}} [f(\theta)] + \psi_n(w_n - 1) \right) \quad \text{(assuming the error is small)}$$

$$= \sum_{n=1}^N \left(\psi_n - \overline{\psi} \right)^2.$$

The final line is also known as the "infinitesimal jackknife" variance approximation.

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$$\text{LOO CV loss at point } n = \underset{p(\theta|x, w_{(-n)})}{\mathbb{E}} \left[-\ell(x_n|\theta) \right] \underset{p(\theta|x)}{\approx} \underset{p(\theta|x)}{\mathbb{E}} \left[-\ell(x_n|\theta) \right] - \psi_n$$

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How to compute the slopes ψ_n ? How large is the error $\mathcal{E}(w)$?

Expressions for the slope and error

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Furthermore, by the mean value theorem, for some \tilde{w} ,

$$\mathcal{E}(w) = \frac{1}{2} \sum_{n=1}^{N} \sum_{n'=1}^{N} \mathcal{E}_{nn'}(w)(w_n - 1)(w_{n'} - 1) \quad \text{where}$$

$$\mathcal{E}_{nn'}(w) := \underbrace{\mathbb{E}_{p(\theta|X,\bar{w})} \left[\bar{f}(\theta) \bar{\ell}_n(\theta) \bar{\ell}_{n'}(\theta) \right]}_{\text{Cannot compute directly!}}$$

(we don't know the intermediate value theorem's \tilde{w}).

But we can analyze it.

Here, an overbar denotes "posterior–mean zero." For example, $\bar{f}(\theta) := f(\theta) - \underset{p(\theta|X)}{\mathbb{E}}[f(\theta)]$.

Theoretical results

How good is the linear approximation (IJ covariance) as an approximation of the limiting variance of $\sqrt{N}\underset{p(\theta|X)}{\mathbb{E}}[f(\theta)]$?

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Theorem 4 of Giordano and Broderick [2024] (paraphrase & conjecture):

In a flexible class of high-dimensional exponential family models,

even when $p\left(f(\theta)|X\right)$ concentrates marginally (!),

- $\sqrt{N}\mathcal{E}(w)$ does not converge to zero (so the IJ covariance is inconsistent), but...
- $\sqrt{N}\mathcal{E}(w) = \tilde{O}_p(1)$, and proportional to the nuisance parameters' posterior covariance

8

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What if $f(\theta)$ concentrates marginally, but some components don't concentrate?

Theorem 4 of Giordano and Broderick [2024] (paraphrase & conjecture):

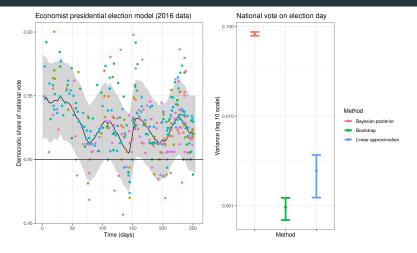
In a flexible class of high-dimensional exponential family models,

even when $p\left(f(\theta)|X\right)$ concentrates marginally (!),

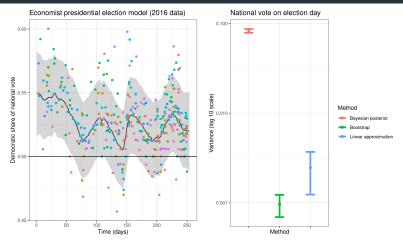
- $\sqrt{N}\mathcal{E}(w)$ does not converge to zero (so the IJ covariance is inconsistent), but...
- $\sqrt{N}\mathcal{E}(w) = \tilde{O}_p(1)$, and proportional to the nuisance parameters' posterior covariance
- \Rightarrow High–dimensional Bayesian models are an extremely common class of problems for which the influence function may not provide a good approximation.

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Observations and consequences



Observations and consequences



Preprint: Giordano and Broderick [2024] (arXiv:2305.06466)

- · Detailed proofs
- · Simple analytical examples
- · Simulated and real-world experiments

9

References

- A. Gelman and M. Heidemanns. The Economist: Forecasting the US elections., 2020. URL https://projects.economist.com/us-2020-forecast/president. Data and model accessed Oct., 2020.
- R. Giordano and T. Broderick. The Bayesian infinitesimal jackknife for variance. arXiv preprint arXiv:2305.06466, 2024.
- J. Huggins and J. Miller. Reproducible model selection using bagged posteriors. Bayesian Analysis, 18(1):79-104, 2023.

Consider $p(X|\gamma) = \prod_{n=1}^N \text{NegativeBinomial}(x_n|\gamma)$. Here, $\theta = \gamma$ is a scalar.

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As $N \to \infty$, $p(\gamma|X)$ concentrates at rate $1/\sqrt{N}$ (Bernstein–von Mises).

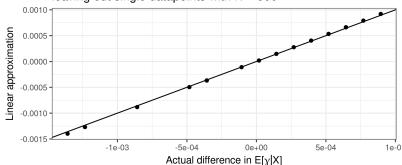
$$\Rightarrow N\left(\underset{p(\gamma|X,w_n)}{\mathbb{E}}[\gamma] - \underset{p(\gamma|X)}{\mathbb{E}}[\gamma]\right) = \psi_n(w_n - 1) + \underset{\boldsymbol{O_p}(N^{-1})}{\boldsymbol{O_p}(N^{-1})}.$$

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Negative Binomial model leaving out single datapoints with N = 800

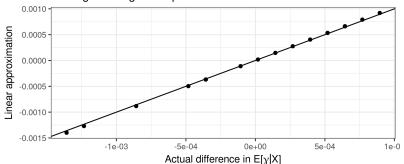


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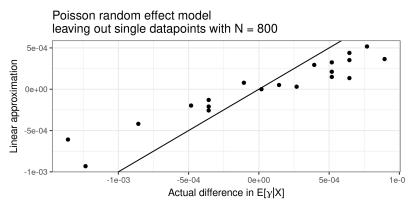
Negative Binomial model leaving out single datapoints with N = 800



Problem: Most computationally hard Bayesian problems don't concentrate.

Experiments

Example: Poisson model with random effects (REs) λ and fixed effect $\gamma.$



A contradiction?

Negative binomial observations.

Asymptotically linear in \boldsymbol{w} .

Poisson observations with random effects.

Asymptotically non-linear in \boldsymbol{w} .

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Asymptotically linear in \boldsymbol{w} .

Asymptotically non-linear in \boldsymbol{w} .

With a constant regressor, Gamma REs, and one RE per observation, these are the same model, with the same $p(\gamma|X)$.

Is $\underset{p(\gamma|X,w)}{\mathbb{E}}[\gamma]$ linear in the data weights or not?

Negative binomial observations.

Poisson observations with random effects.

Asymptotically linear in w.

Asymptotically non-linear in w.

$$\log p(X|\gamma, w^m) = \sum_{n=1}^N w_n^m \log p(x_n|\gamma) \quad \ \log p(X|\gamma, \lambda, w^c) = \sum_{n=1}^N w_n^c \log p(x_n|\lambda, \gamma)$$

With a constant regressor, Gamma REs, and one RE per observation, these are the same model, with the same $p(\gamma|X)$.

Is $\underset{p(\gamma|X,w)}{\mathbb{E}}[\gamma]$ linear in the data weights or not?

Trick question! We weight a log likelihood contribution, not a datapoint.

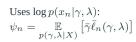
The two weightings are not equivalent in general.

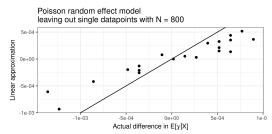
Experimental results

Our results were actually computed on **identical datasets** with G=N and $g_n=n$.

Negative Binomial model

Uses
$$\log p(x_n|\gamma)$$
:
$$\psi_n = \underset{p(\gamma|X)}{\mathbb{E}} \left[\bar{\gamma} \bar{\ell}_n(\gamma) \right]$$





Experimental results

Our results were actually computed on **identical datasets** with G = N and $g_n = n$.

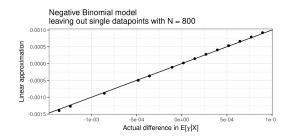
Uses $\log p(x_n|\gamma)$: $\psi_n = \underset{p(\gamma|X)}{\mathbb{E}} \left[\bar{\gamma} \bar{\ell}_n(\gamma) \right]$

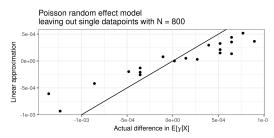
Not computable from $\gamma, \lambda \sim p(\gamma, \lambda|X)$ in general.

Uses $\log p(x_n|\gamma,\lambda)$: $\psi_n = \mathop{\mathbb{E}}_{p(\gamma,\lambda|X)} \left[\bar{\gamma} \bar{\ell}_n(\gamma,\lambda) \right]$

Computable from

$$\gamma, \lambda \sim p(\gamma, \lambda | X).$$





Experimental results

Our results were actually computed on **identical datasets** with G=N and $g_n=n$.

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Computable from $\gamma, \lambda \sim p(\gamma, \lambda | X)$.

May still be useful when $p(\lambda|X)$ is *somewhat* concentrated.

