

ML in Economics and Finance: Where do We Go Now? - Part II

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- 1. What is ML, anyway?
 - 2. Causality in High Dimensions
 - 3. (Seriously) Heterogeneous Partial Effects
 - 4. Solving Large-Scale General Equilibrium Models
- } **Yesterday**
- } **Today**

Heterogeneous Partial Effects

Let Y be an outcome and X, Z be features (covariates). We frequently want to approximate

$$h(x, z) \equiv \mathbb{E}[Y|X = x, Z = z]$$

and the **partial effects**

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- Approach 1: impose a parametric model for h , e.g. linear regression. Pros and cons?
- Approach 2: use fully nonparametric methods. Pros and cons?

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- Approach 2: use fully nonparametric methods. Pros and cons?
- The third way is the charm: a bit of structure, a bit of ML!

Example I - Heterogenous Treatment Effects

- Outcomes Y_i depend on a treatment $X_i \in \mathbb{R}$ and covariates $Z_i \in \mathbb{R}^p$;
- The dose X_i depends on observables Z_i ;
- Potential outcomes $Y_i(x)$ for each dose $x \in \mathbb{R}$;

Example I - Heterogenous Treatment Effects

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- The dose X_i depends on observables Z_i ;
- Potential outcomes $Y_i(x)$ for each dose $x \in \mathbb{R}$;
- The conditional average effect of increasing the dose is $\tau(x, z) \equiv \frac{\partial}{\partial x} \mathbb{E}[Y(x)|Z = z]$;
- If $\mathbb{E}[Y(x)|X = x, Z = z] = \mathbb{E}[Y(x)|Z = z]$ (common assumption in the literature), then

$$\tau(x, z) = \frac{\partial}{\partial x} \mathbb{E}[Y|X = x, Z = z] = \frac{\partial h(x, z)}{\partial x}$$

Example II - Grouped Heterogeneity

Consider the following model:

$$Y_i = \alpha(Z_i) + X_i'\beta + \varepsilon_i, \quad \mathbb{E}[\varepsilon_i|X_i, Z_i] = 0,$$

- X_i affects Y_i homogeneously;
- Intercept $\alpha(Z_i)$ varies with Z_i , maybe in a highly nonlinear way;
- Since Z_i and X_i can be correlated, this can affect inference about β ;
- [Bonhomme and Manresa \(2015\)](#) studied how democracy affects national income using this model;
- In their case: $\alpha(Z_i)$ is constant across groups but Z_i defines membership;
- In our notation: $h(z, x) = \alpha(z) + x'\beta$

How can we balance flexibility and interpretability?

Masini and Medeiros (2025)  proposed a middle ground:

$$h(x, z) = x^\top \beta(z), \quad \frac{\partial h(x, z)}{\partial x} = \beta(z)$$

- The partial effect of X on Y varies with Z through $\beta(z)$;
- $\beta(\cdot)$ is a Lipschitz function that can be estimated with ML methods;
- No need to numerically approximate $\frac{\partial}{\partial x} h(x, z)$;

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But what is a Random Forest, anyway? 🤔

Questions?

Quick Intro to Random Forests

Random Trees

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- If $X \in \mathbb{R}^p$, consider a finite partition $\{S_1, S_2, \dots, S_m\}$ of \mathbb{R}^p ;
- Each S_i is a hyperrectangle defined by recursive binary splits on the covariates;
- On each S_i , f is constant: $f(x) = \mu_i$ for all $x \in S_i$;
- After a tree has been estimated ("grown"), $\hat{f}(x_i) = \mu_i$ if $x_i \in S_i$;

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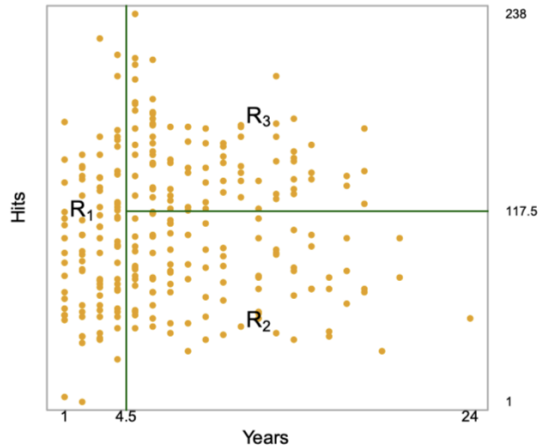
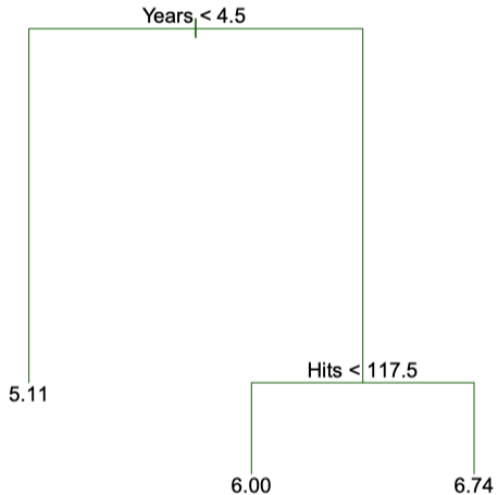
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The really complicated part: there are *so many* partitions... how to pick one?

Hyafil and Rivest (1976): this is harder than you think! The problem is NP-complete!

Example: Predicting Baseball Player Salaries



How to pick a split point? Use some greed!

Let's say you want to split on feature X_j at point δ :

$$S_1 \equiv \{x \in \mathbb{R}^p : x_j \leq \delta\}, \quad S_2 \equiv \{x \in \mathbb{R}^p : x_j > \delta\}$$

$$\mu_1 \equiv \sum_{i: x_i \in S_1} \frac{Y_i}{n_1}, \quad \mu_2 \equiv \sum_{i: x_i \in S_2} \frac{Y_i}{n_2}$$

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- Define $SSR(\delta) \equiv \sum_{i: x_i \in S_1} (Y_i - \mu_1)^2 + \sum_{i: x_i \in S_2} (Y_i - \mu_2)^2$
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- Example: minimum number of observations per leaf;

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This is the so-called the **CART** algorithm due to [Breiman et al. \(1984\)](#).

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Key insight:

- Each tree is noisy and biased, but averaging them reduces variance dramatically;
- Randomness *decorrelates* the trees, making averaging powerful;

Questions?

Back to Partial Effects

The Main Insight

We have a random sample $\{(Y_i, X_i, Z_i)\}_{i=1}^n$ from

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Suppose you want to split at $Z_j \leq \delta$ as before. Then:

$$\begin{aligned} S_1 &\equiv \{z \in \mathbb{R}^p : z_j \leq \delta\}, & S_2 &\equiv \{z \in \mathbb{R}^p : z_j > \delta\} \\ (\hat{\beta}_1) &\equiv \arg \min_{\beta} \sum_{i: Z_{i,j} \in S_1} (Y_i - X_i^\top \beta)^2, & (\hat{\beta}_2) &\equiv \arg \min_{\beta} \sum_{i: Z_{i,j} \in S_2} (Y_i - X_i^\top \beta)^2 \\ SSR(\delta) &\equiv \sum_{i: Z_{i,j} \in S_1} (Y_i - X_i^\top \hat{\beta}_1)^2 + \sum_{i: Z_{i,j} \in S_2} (Y_i - X_i^\top \hat{\beta}_2)^2 \end{aligned}$$

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Pick δ to minimize $SSR(\delta)$!

The Algorithm

Pick a number of trees B and a minimum leaf size k . For $b = 1, \dots, B$:

1. Draw a bootstrap sample of size $s \leq n$ from the data;
2. Divide the data into two halves \mathcal{A} and \mathcal{B} ;
3. Using \mathcal{B} , keep splitting at random dimensions j using the previous criterion;
4. Stop when all leaves have less than $2k - 1$ and more than k observations;
5. Using \mathcal{A} , estimate $\beta(z)$ using only observations in the leaf where z falls;

The final estimate is

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This algorithm uses **honest trees**! Similar intuition to cross-fitting.

Cool Properties and Limitations

Cool properties:

- Highly interpretable and relatively mild assumptions on $\beta(z)$;
- Easy confidence intervals for $\beta(z)$ at any point z :

$$\Omega^{-1/2}(z) \left(\hat{\beta}(z) - \beta(z) \right) \xrightarrow{d} \mathcal{N}(0, I_q)$$

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Limitations:

- The dimension of X_i should be small relative to n ;
- The dimension of Z_i cannot be *that* large relative to n ;
- Pointwise inference only;
- It cannot be readily applied to time series and panel data;
- It can be computationally demanding in large datasets;

Questions?

Solving Large-Scale General Equilibrium Models

Thank you!
See you tomorrow, stay tuned!

Appendix and References

References



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