

# Markov Chain Monte Carlo simulation components

## bivariate normal distribution

A bivariate normal distribution. Could also simulate this using the MASS library (mvnrmf)

```
mu1 <- 0 # expected value of x
mu2 <- 0.5 # expected value of y
sig1 <- 0.5 # variance of x
sig2 <- 2 # variance of y
rho <- 0.5 # corr(x, y)

### Some additional variables for x-axis and y-axis
xm <- -3
xp <- 3
ym <- -3
yp <- 3

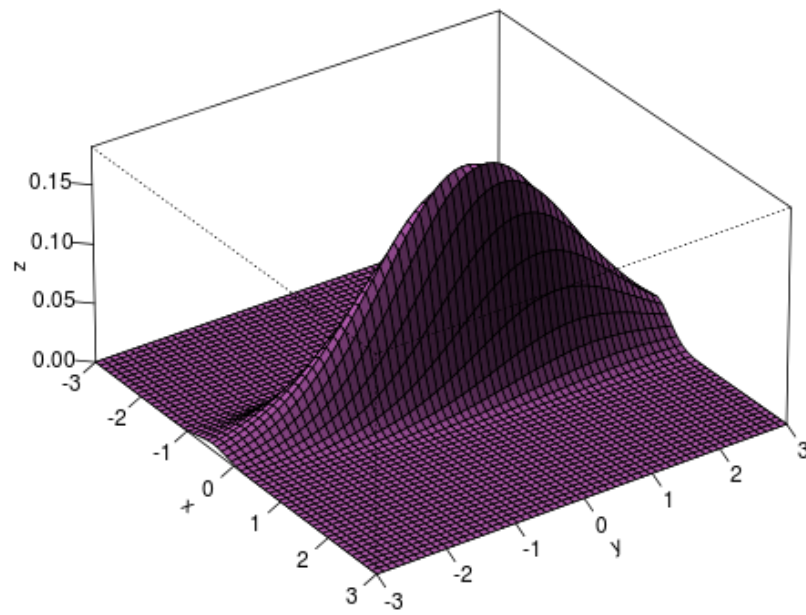
x <- seq(xm, xp, length = as.integer((xp + abs(xm)) * 10)) # vector series x
y <- seq(ym, yp, length = as.integer((yp + abs(ym)) * 10)) # vector series y

### Core function
bivariate <- function(x, y) {
  term1 <- 1/(2 * pi * sig1 * sig2 * sqrt(1 - rho^2))
  term2 <- (x - mu1)^2/sig1^2
  term3 <- -(2 * rho * (x - mu1) * (y - mu2))/(sig1 * sig2)
  term4 <- (y - mu2)^2/sig2^2
  z <- term2 + term3 + term4
  term5 <- exp((-z/(2 * (1 - rho^2))))
  return(term1 * exp(-z * (2 * (1 - rho^2))))
}

### Computes the density values
z <- outer(x, y, bivariate)

### Plot
persp(x, y, z, main = "Bivariate Normal Distribution", sub = bquote(bold(mu[1]) ==
  .(mu1) ~ ", " ~ sigma[1] == .(sig1) ~ ", " ~ mu[2] == .(mu2) ~ ", " ~ sigma[2] ==
  .(sig2) ~ ", " ~ rho == .(rho)), col = "orchid2", theta = 55, phi = 30,
  r = 40, d = 0.1, expand = 0.5, ltheta = 90, lphi = 180, shade = 0.4, ticktype = "detail",
  nticks = 5)
```

### Bivariate Normal Distribution



$$\mu_1 = 0, \sigma_1 = 0.5, \mu_2 = 0.5, \sigma_2 = 2, \rho = 0.5$$

Figure 1: plot of chunk bivar\_norm

## Curved and folded surfaces - not normal

```
require(ggplot2)
require(reshape2)
require(gridExtra)
n = 100
x = y = (scat <- sort(rnorm(n) + rchisq(n, df = 4)))
fun <- function(x, y) {
  r <- sqrt(x^2 + y^2)
  10 * sin(r)/r
}
z <- outer(x, y, fun)
scat.df <- data.frame(x = x, y = x, z = z)

p = persp(x, y, z, theta = 30, phi = 30, expand = 0.5, shade = 0.75, ticktype = "detailed",
  xlab = "X", ylab = "Y")
```

## Probability density functions

```
library(ggplot2)
library(reshape2)
nval = 10^4
normal <- rnorm(nval, 0.5, 0.5)
pois <- rpois(nval, 0.1)
beta <- rbeta(nval, 6, 4)
dist <- data.frame(normal = normal, poisson = pois, beta = beta)
dist_m <- melt(dist)

## Using as id variables

names(dist_m) <- c("distribution", "value")
ggplot(dist_m, aes(x = value, fill = distribution)) + geom_density(alpha = 0.6) +
  ggtitle("Probability density functions")
```

## Random sampling in R

This is meant to show that there is very little correlation between random numbers generated in R.

```
x = runif(Nsim)
```

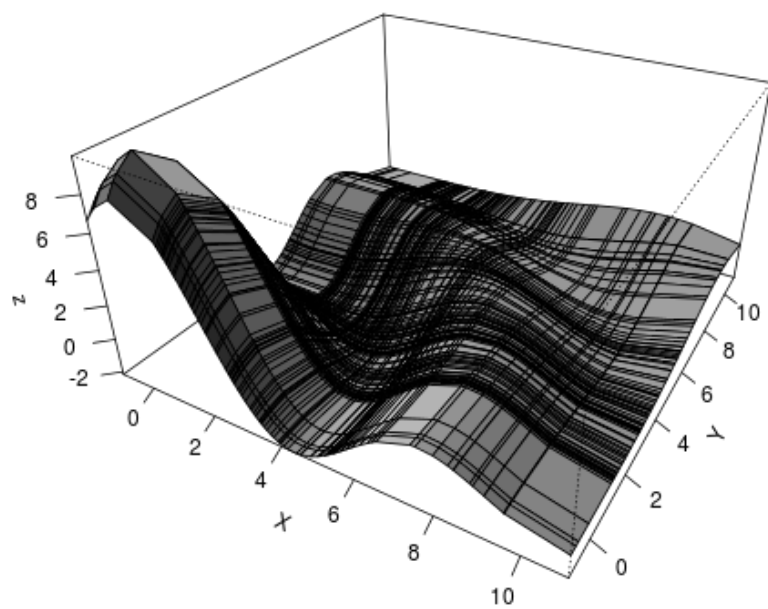


Figure 2: plot of chunk curve\_fold

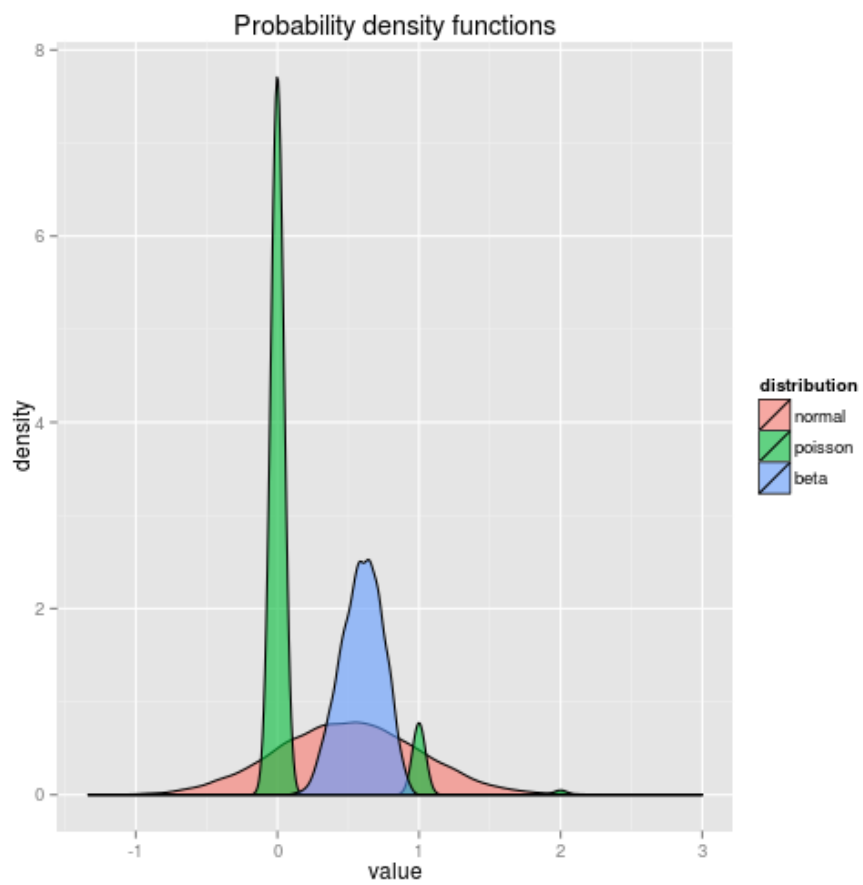


Figure 3: plot of chunk pdfs

```
## Error: object 'Nsim' not found

x2 = x[-1]
par(mfrow = c(1, 3))
hist(x)
plot(x1, x2)

## Error: error in evaluating the argument 'x' in selecting a method for
## function 'plot': Error: object 'x1' not found

acf(x)
```

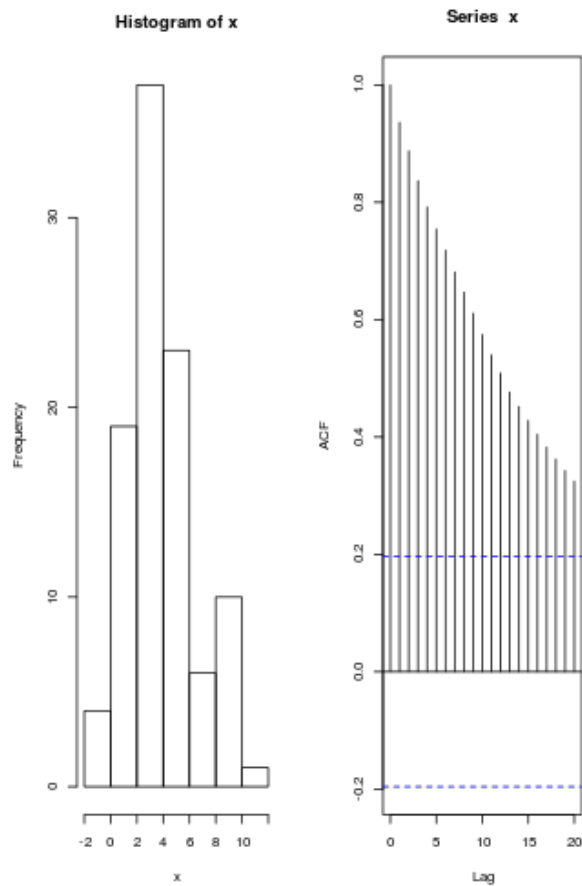


Figure 4: plot of chunk random\_numbers

## Using uniform distribution to generate a beta distribution

This comes from [suess\_introduction.2010, 32].

A real function of a random variable is another random variable. Random variables with a wide variety of distributions can be obtained by transforming a standard uniform random variable  $U \approx UNIF(0,1)$ . Let  $U \approx UNIF(0,1)$ . ... We seek the distribution of  $X = U^2$  [suess\_introduction.2010, 32].

The example uses a uniform random variable to create a beta distribution.

```
set.seed(1234)
m = 10000
u = runif(m)
x = u^2
xx = seq(0, 1, by = 0.001)
cut.u = (0:10)/10
cut.x = cut.u^2
par(mfrow = c(1, 2))
hist(u, breaks = cut.u, prob = T, ylim = c(0, 10))
lines(xx, dunif(xx), col = "blue")
hist(x, breaks = cut.x, prob = T, ylim = c(0, 10))
lines(xx, 0.5 * xx^-0.5, col = "blue")

par(mfrow = c(1, 1))
```

## Use inverse transform to generate distribution

A uniform random variable can be used to generate an arbitrary distribution using the the inverse probability transform.

$$F(x) = \int_{-\infty}^x f(t) dt$$

where  $F$  is the cumulative distribution function (cdf), For instance, to generate an exponential random variable

```
Nsim = 10^5
U = runif(Nsim)
X = -log(U) #
Y = rexp(Nsim)
par(mfrow = c(1, 2))
hist(X, freq = F, main = "Exp from uniform", breaks = 20)
hist(Y, freq = F, main = "Exp from R", breaks = 20)
```

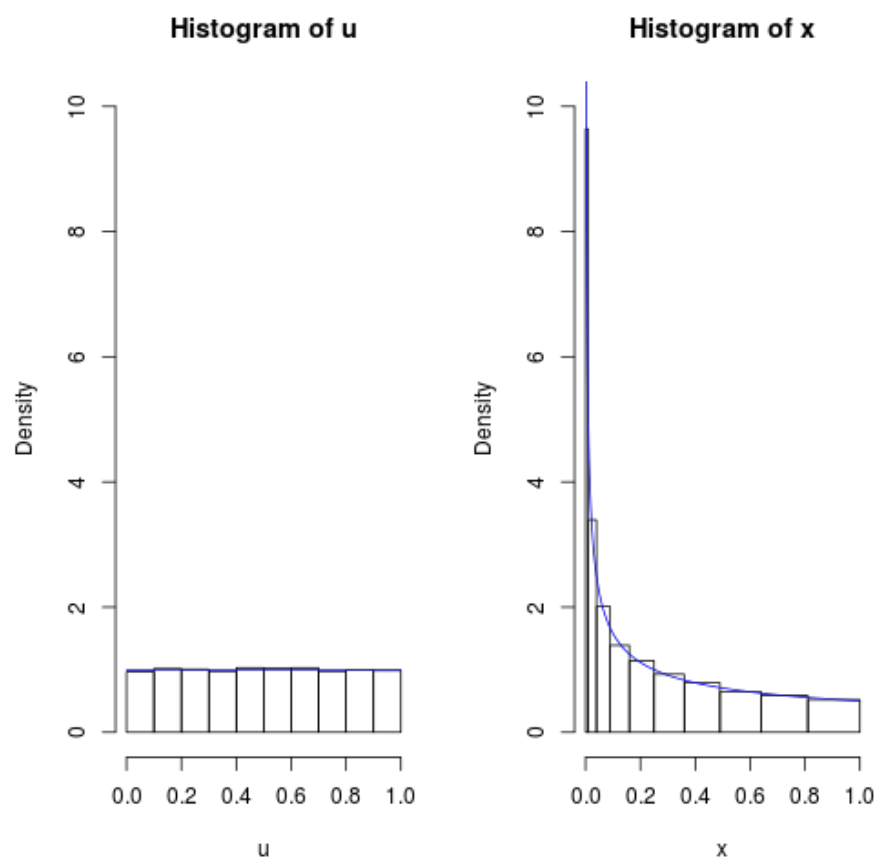


Figure 5: plot of chunk uniform\_beta



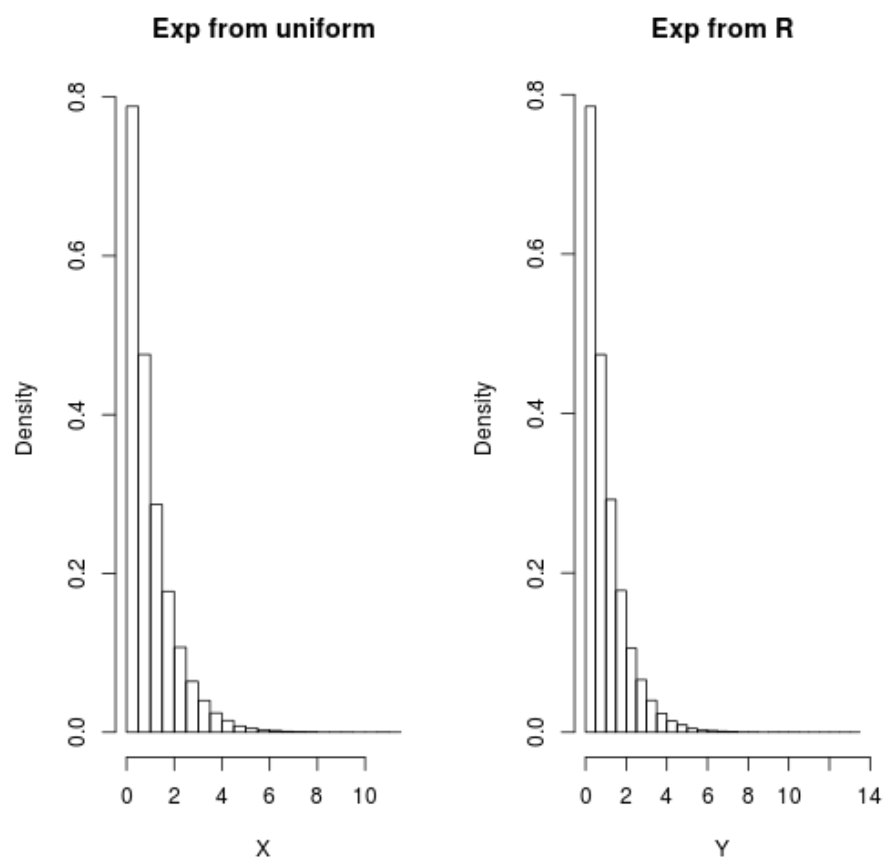


Figure 6: plot of chunk inverse\_transform

## Another demo of use of inverse transform

The point is that a supply of random variables can be used to generate different distribution; [robert\_introducing\_2010,p.44]

```
Nsim = 10^5
mu = 0.4
beta = 0.4
U = runif(Nsim)
X = -mu + beta * (log((1 - U)/U)) #I had to verify this by solving the cdf for X
Y = rlogis(Nsim, location = mu, scale = beta)
par(mfrow = c(1, 2))
hist(X, freq = F, main = "Logistic from uniform")
hist(Y, freq = F, main = "Logistic from R")
```

## Accept-reject method to simulate beta

[robert\_introducing\_2010, 53-54], but code taken from mcms package

## Markov-chain random walk to generate normal distribution

‘Consider the Markov chain defined by  $X(t+1) = \text{sigma}X(t) + \text{epsilon}(t)$  where  $\text{epsilon}(t) \sim \text{Normal}(0,1)$ ’, [robert\_introducing\_2010, p.169]

```
X <- vector(length = 10^4)
X[1] = runif(1)
sigma = 0.9
for (t in 1:10^4) {
  X[t + 1] = sigma * X[t] + runif(1, min = 0, max = 1)
}
Y = rnorm(10^4, 0, 1/(1 - sigma^2))
par(mfrow = c(1, 2))
hist(X, breaks = 200, freq = F, main = "Markov chain generated normal")
hist(Y, breaks = 200, freq = F, main = "Stationary distribution")
```

## Beta distribution generated using Metropolis-Hastings algorithm

To simulate a beta distribution: ‘we can just as well use a Metropolis-Hastings algorithm, where the target density  $f$  is the  $\text{Be}(2.7,6.3)$  density and the candidate  $q$  is uniform over  $[0,1]$ ’ [robert\_introducing\_2010,p.172]

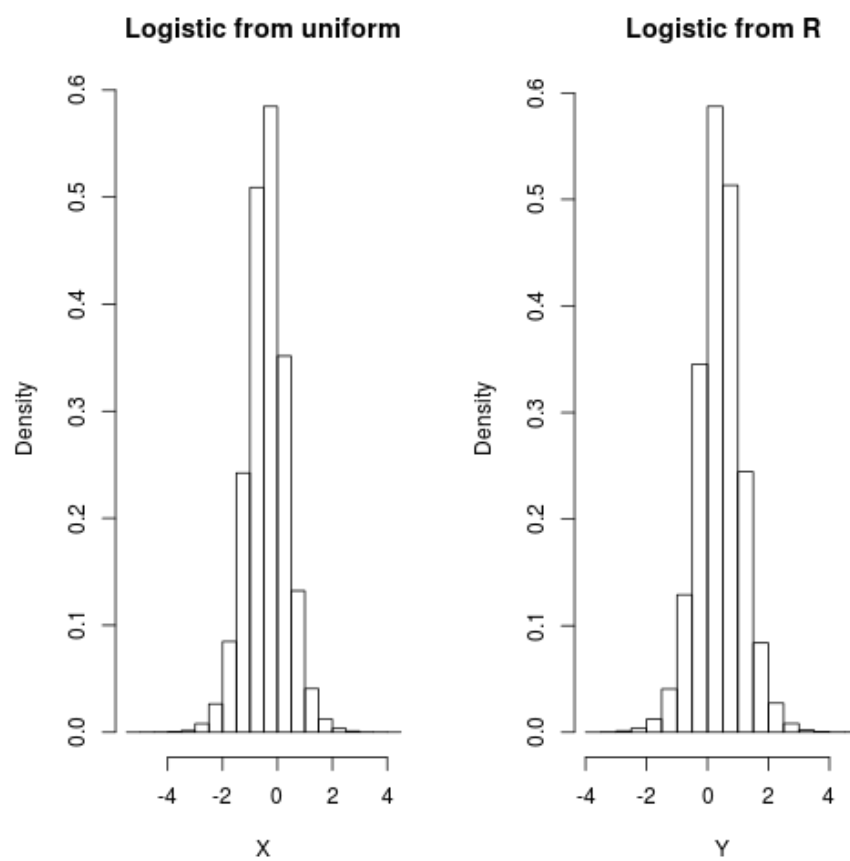


Figure 7: plot of chunk diff\_dist

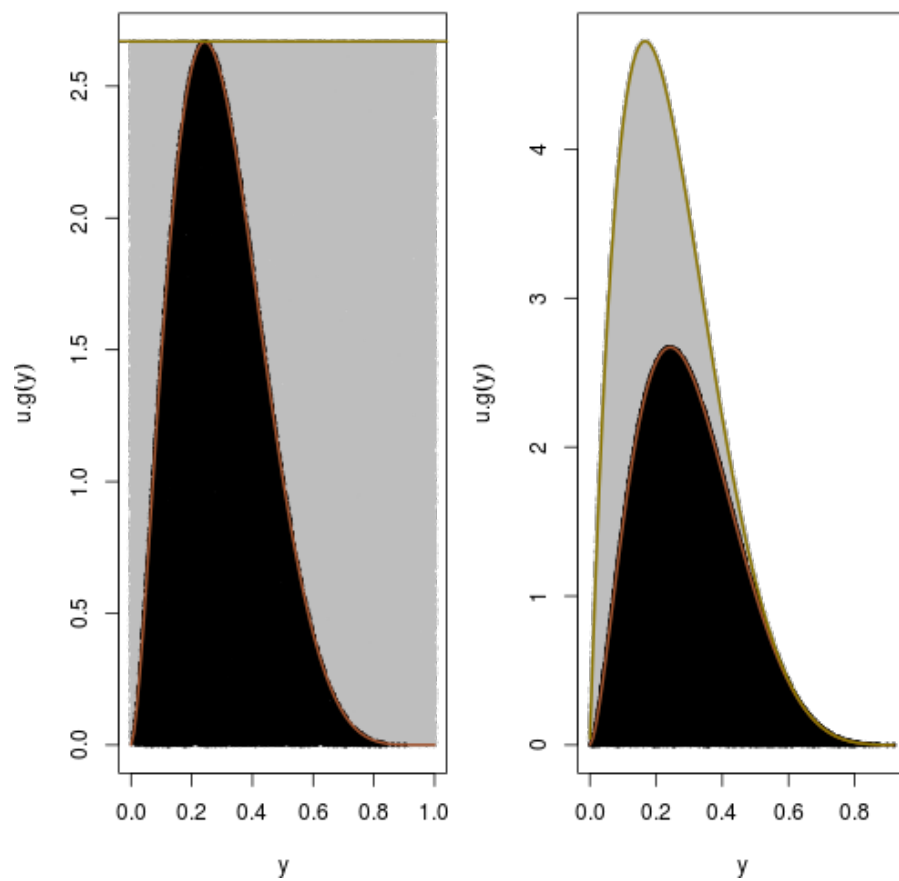


Figure 8: plot of chunk accept\_reject\_beta

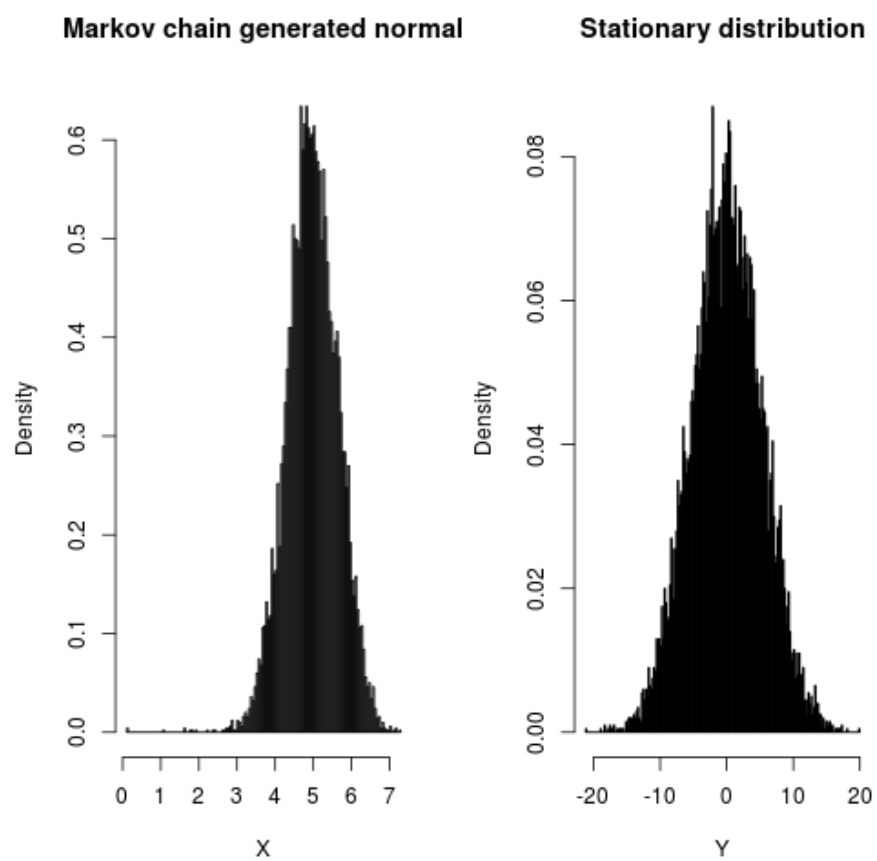


Figure 9: plot of chunk markov\_chain

```

a = 2.7
b = 6.3
c = 2.669
Nsim = 5000
X = rep(runif(1), Nsim)
accept <- vector(mode = "logical", length = Nsim)
for (i in 2:Nsim) {
  Y = runif(1)
  rho = dbeta(Y, a, b)/dbeta(X[i - 1], a, b)
  accept[i] = runif(1) < rho
  X[i] = X[i - 1] + (Y - X[i - 1]) * (accept[i])
}

Z = rbeta(5000, a, b)
par(mfrow = c(1, 2))
hist(X, freq = F, breaks = 200, main = "Sample generated by Metropolis-Hastings")
hist(Z, freq = F, breaks = 200, main = "Sample generated by exact iid")

print(ks.test(jitter(X), rbeta(5000, a, b)))

##
## Two-sample Kolmogorov-Smirnov test
##
## data: jitter(X) and rbeta(5000, a, b)
## D = 0.0198, p-value = 0.2809
## alternative hypothesis: two-sided

X.df <- data.frame(X = X, Z = Z)

ggplot(X.df, aes(x = X)) + geom_density()

ggplot(X.df, aes(x = Z)) + geom_density()

```

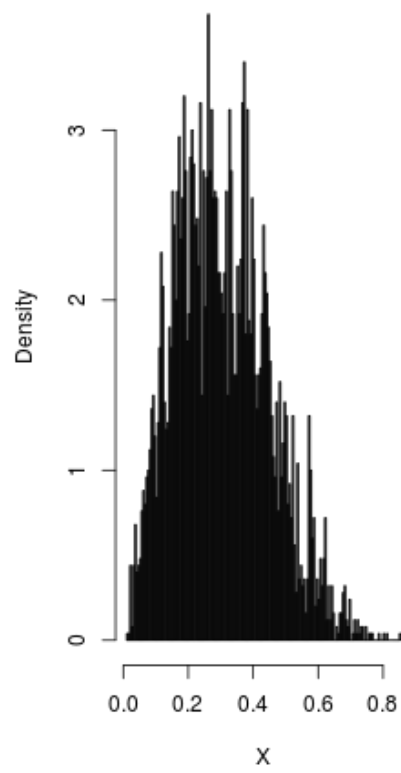
Note the acceptance rate on this: **45.72%**.

## Two-stage Gibbs sampling

Two-stage gibbs sampler

Gibbs\_pair\_of\_distributions\_example7.2 from [robert\_introducing\_2010]

Sample generated by Metropolis-Hastings



Sample generated by exact iid

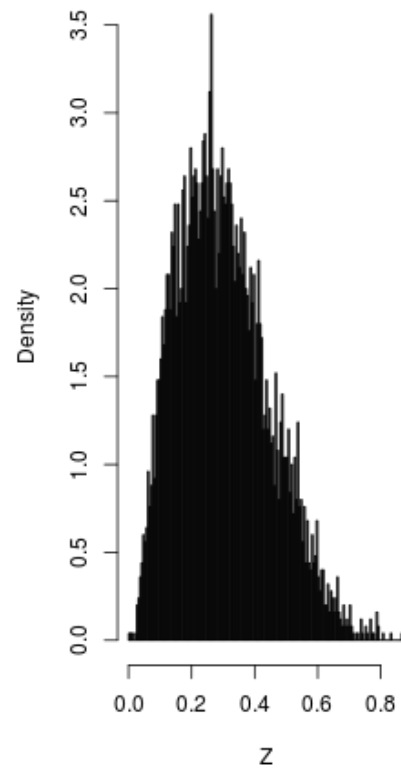


Figure 10: plot of chunk beta\_metro\_hast

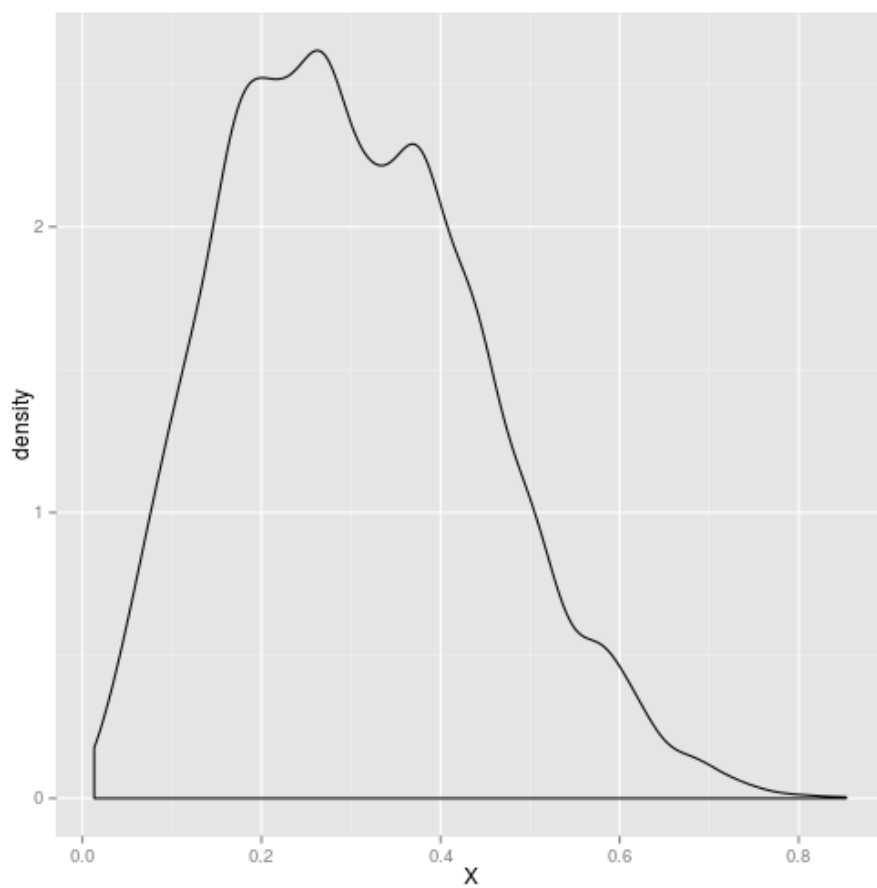


Figure 11: plot of chunk beta\_metro\_hast



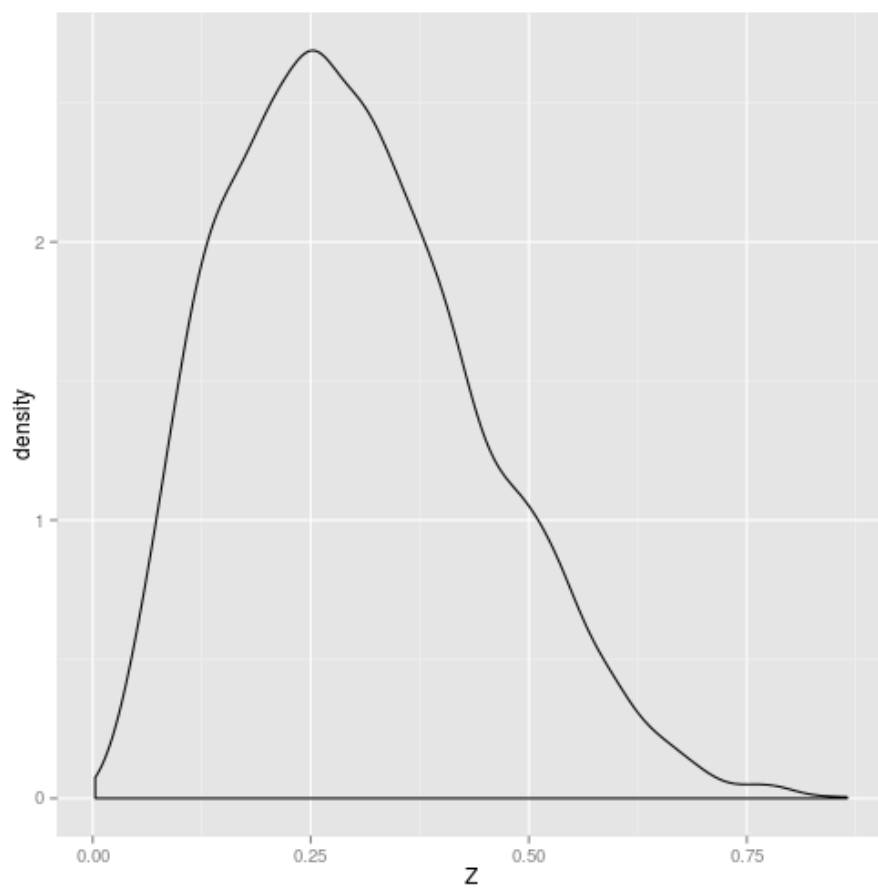


Figure 12: plot of chunk beta\_metro\_hast

```

Nsim = 5000
n = 15
a = 3
b = 7
X = T = array(0, dim = c(Nsim, 1))
T[1] = rbeta(1, a, b)
X[1] = rbinom(1, n, T[1])
for (i in 2:Nsim) {
  X[i] = rbinom(1, n, T[i - 1])
  T[i] = rbeta(1, a + X[i], n - X[i] + b)
}
par(mfrow = c(1, 2))
hist(T, freq = F, main = "theta")
hist(X, freq = F, main = "X")

```

## 6.4 A Simple Gibbs Sampler

This sampler implements a screening test. It comes from [suess\_introduction.2010]

Variable|Value ———|———— D|the proportion of infected population. T|the proportion of population that tests positive. eta |sensitivity of the test (correctly identified positives) theta|specificity of the test (correctly identified negatives) gamma|predictive value of a positive test delta|predictive value of a negative test

```

m = 80000
eta = 0.99
theta = 0.97
gamma = 0.4024
delta = 0.9998
d = numeric(m)
d[1] = 0
t = numeric(m)

for (n in 2:m) {
  if (d[n - 1] == 1)
    t[n - 1] = rbinom(1, 1, eta) else t[n - 1] = rbinom(1, 1, 1 - theta)

  if (t[n - 1] == 1)
    d[n] = rbinom(1, 1, gamma) else d[n] = rbinom(1, 1, 1 - delta)
}

runprop = cumsum(d)/1:m
mean(d[m/2 + 1]:m)

```

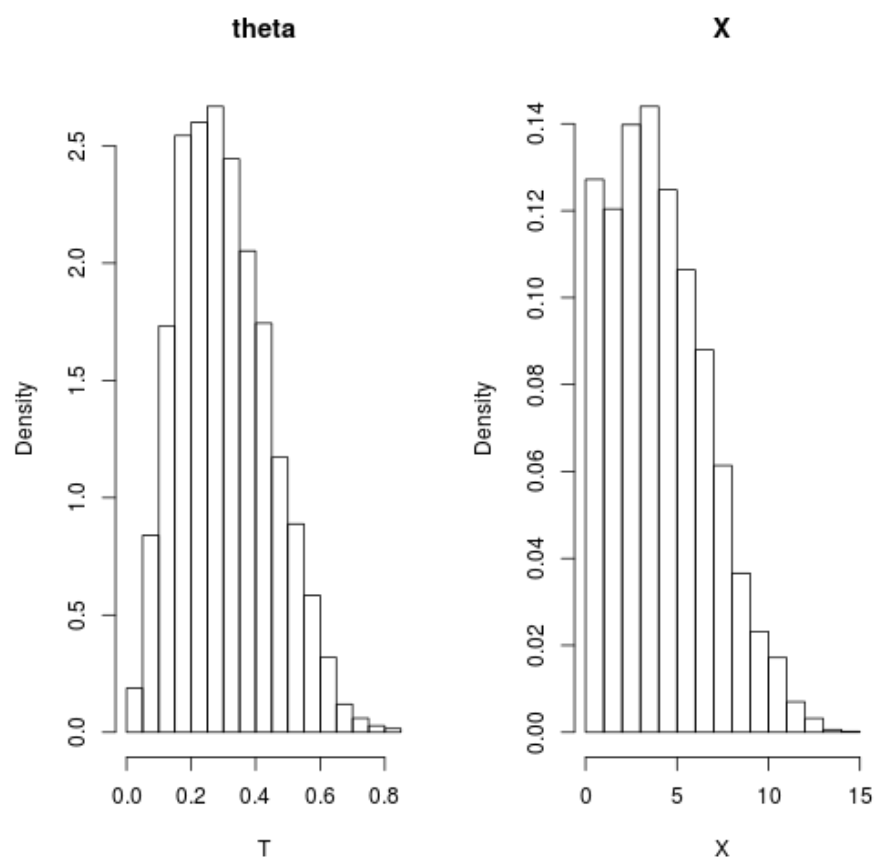


Figure 13: plot of chunk gibbs\_pairs

```
## [1] 40000
```

```
par(mfrow = c(1, 2))  
plot(runprop, type = "l", ylim = c(0, 0.05), xlab = "Step", ylab = "Running Proportion Infected")  
acf(d, ylim = c(-0.1, 0.4), xlim = c(1, 10))
```

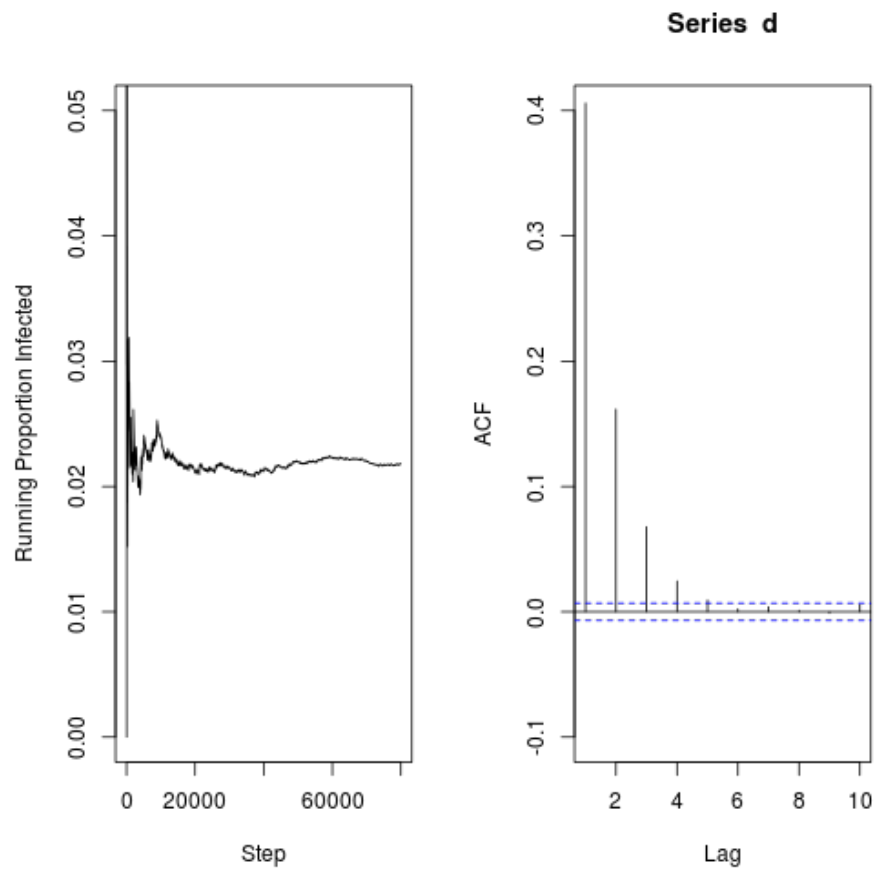


Figure 14: plot of chunk simple\_gibbs

```
par(mfrow = c(1, 1))  
acf(d, plot = F)
```

```
##  
## Autocorrelations of series 'd', by lag
```

```
##
##      0      1      2      3      4      5      6      7      8      9
## 1.000 0.406 0.162 0.068 0.025 0.009 0.002 0.004 0.001 -0.001
##    10    11    12    13    14    15    16    17    18    19
## 0.006 0.007 0.005 0.007 0.003 0.005 0.002 -0.001 0.000 0.001
##    20    21    22    23    24    25    26    27    28    29
## 0.004 -0.001 -0.006 -0.008 -0.007 -0.001 0.001 0.005 0.003 0.003
##    30    31    32    33    34    35    36    37    38    39
## 0.001 0.001 -0.001 0.001 0.002 0.001 -0.002 -0.004 -0.005 -0.007
##    40    41    42    43    44    45    46    47    48    49
## -0.007 -0.006 -0.006 -0.002 0.005 0.009 0.007 0.002 -0.001 0.002
```

## Simulation by sampling

Imagine trying to work out how likely we are to get 90 good chips in batch of 100 from a production line where we know that 5 in every 100 chips are faulty. There are two ways of doing it. The first is analytical.

```
choose(90, 5)/choose(100, 5)
```

```
## [1] 0.5838
```

The second way is to simulate a sample of the chips, and calculate the probability of getting 90% good ones:

```
set.seed(1237)
m = 1e+05
good = numeric(m)

for (i in 1:m) {
  pick = sample(1:100, 5)
  good[i] = sum(pick <= 90)
}

mean(good == 5)
```

```
## [1] 0.5829
```

```
good.df <- data.frame(x = good)
ggplot(good.df, aes(x = x)) + geom_bar()
```

```
## Error: Don't know how to add geom_bar() to a plot
```

## Bivariate normal distribution

[@suess\_introduction\_2010, 177-178]

```
set.seed(1234)
m = 40000
rho = 0.8
sgm = sqrt(1 - rho^2)
xc = yc = numeric(m)
# vectors of state components
xc[1] = -3
yc[1] = 3
# arbitrary starting values
jl = 1
jr = 1
# l and r limits of proposed jumps
for (i in 2:m) {
  xc[i] = xc[i - 1]
  yc[i] = yc[i - 1]
  # if jump rejected
  xp = runif(1, xc[i - 1] - jl, xc[i - 1] + jr) # proposed x coord
  yp = runif(1, yc[i - 1] - jl, yc[i - 1] + jr) # proposed y coord
  nmtr = dnorm(xp) * dnorm(yp, rho * xp, sgm)
  dntr = dnorm(xc[i - 1]) * dnorm(yc[i - 1], rho * xc[i - 1], sgm)
  r = nmtr/dntr
  # density ratio
  acc = (min(r, 1) > runif(1))
  # jump if acc == T
  if (acc) {
    xc[i] = xp
    yc[i] = yp
  }
}
x = xc[(m/2 + 1):m]
y = yc[(m/2 + 1):m]
# states after burn-in
round(c(mean(x), mean(y), sd(x), sd(y), cor(x, y)), 4)
```

```
## [1] -0.0348 -0.0354  0.9966  0.9992  0.7994

mean(diff(x) == 0)

## [1] 0.4316

# proportion of proposals rejected
mean(pmax(x, y) >= 1.25)

## [1] 0.1472

# prop. of subj. getting certificates
par(mfrow = c(1, 2), pty = "s")
plot(xc[1:100], yc[1:100], xlim = c(-4, 4), ylim = c(-4, 4), type = "l")
plot(x, y, xlim = c(-4, 4), ylim = c(-4, 4), pch = ".")

par(mfrow = c(1, 1), pty = "m")
round(c(mean(x), mean(y), sd(x), sd(y), cor(x, y)), 4)

## [1] -0.0348 -0.0354  0.9966  0.9992  0.7994
```

## Examples of chains with larger state spaces

[@suess.introduction.2010, 180]

```
set.seed(1235)
m = 20000
rho = 0.8
sgm = sqrt(1 - rho^2)
xc = yc = numeric(m)
# vectors of state components
xc[1] = -3
yc[1] = 3
# arbitrary starting values
for (i in 2:m) {
  xc[i] = rnorm(1, rho * yc[i - 1], sgm)
  yc[i] = rnorm(1, rho * xc[i], sgm)
}
x = xc[(m/2 + 1):m]
y = yc[(m/2 + 1):m]
# states after burn-in
round(c(mean(x), mean(y), sd(x), sd(y), cor(x, y)), 4)
```

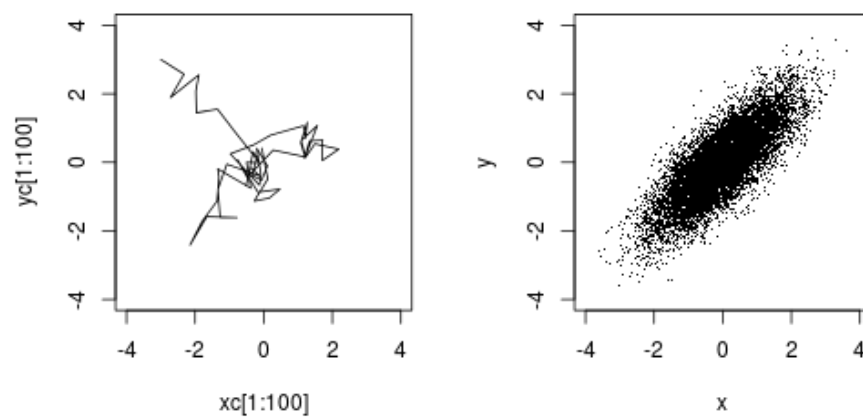


Figure 15: plot of chunk normal\_bivariate\_metropolis



```
## [1] 0.0083 0.0077 1.0046 1.0073 0.8044

best = pmax(x, y)
mean(best >= 1.25)

## [1] 0.1527

# prop. getting certif.
summary(best)

##      Min. 1st Qu.  Median    Mean 3rd Qu.    Max.
## -3.600  -0.391   0.263   0.259   0.909   3.390

par(mfrow = c(1, 2), pty = "s")
hist(best, prob = T, col = "wheat", main = "")

## Warning: the condition has length > 1 and only the first element will be
## used

## Warning: the condition has length > 1 and only the first element will be
## used

abline(v = 1.25, lwd = 2, col = "red")
plot(x, y, xlim = c(-4, 4), ylim = c(-4, 4), pch = ".")
lines(c(-5, 1.25, 1.25), c(1.25, 1.25, -5), lwd = 2, col = "red")

par(mfrow = c(1, 1), pty = "m")

round(c(mean(x), mean(y), sd(x), sd(y), cor(x, y)), 4)

## [1] 0.0083 0.0077 1.0046 1.0073 0.8044
```

## Gibbs sampler on the normal mixture model

In chapter 7 [robert\_introducing\_2010] apply Gibbs sampling to the normal mixture model.

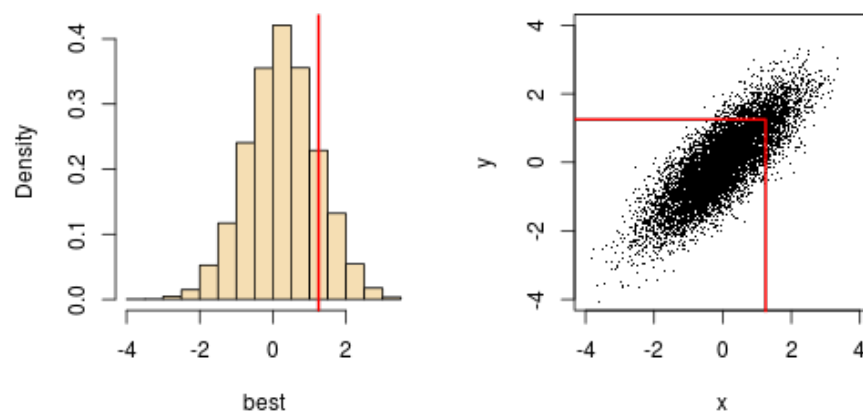


Figure 16: plot of chunk normal\_bivariate\_gibbs

```

Niter = 10^4
v = 1
da = sample(c(rnorm(10^2), 2.5 + rnorm(4 * 10^2)))

like = function(mu) {
  sum(log((0.2 * dnorm(da - mu[1]) + 0.8 * dnorm(da - mu[2]))))
}

mu1 = mu2 = seq(-2, 5, le = 250)
lli = matrix(0, ncol = 250, nrow = 250)
for (i in 1:250) for (j in 1:250) lli[i, j] = like(c(mu1[i], mu2[j]))

x = prop = runif(2, -2, 5)
the = matrix(x, ncol = 2)
curlike = hval = like(x)
for (i in 2:Niter) {
  pp = 1/(1 + ((0.8 * dnorm(da, mean = the[i - 1, 2]))/(0.2 * dnorm(da, mean = the[i - 1, 1]))))
  z = 2 - (runif(length(da)) < pp)
  prop[1] = (v * sum(da[z == 1]))/(sum(z == 1) * v + 1) + rnorm(1) * sqrt(v/(1 + sum(z == 1) * v))
  prop[2] = (v * sum(da[z == 2]))/(sum(z == 2) * v + 1) + rnorm(1) * sqrt(v/(1 + sum(z == 2) * v))
  curlike = like(prop)
  hval = c(hval, curlike)
  the = rbind(the, prop)
}

par(mar = c(4, 4, 1, 1))
image(mu1, mu2, lli, xlab = expression(mu[1]), ylab = expression(mu[2]))
contour(mu1, mu2, -lli, nle = 100, add = T)

## Warning: the condition has length > 1 and only the first element will be
## used

## Warning: the condition has length > 1 and only the first element will be
## used

points(the[, 1], the[, 2], cex = 0.6, pch = 19)
lines(the[, 1], the[, 2], cex = 0.6, pch = 19)

```

hastings 1970 paper - code from mcsim

```
## Error: object 'nsim' not found
```

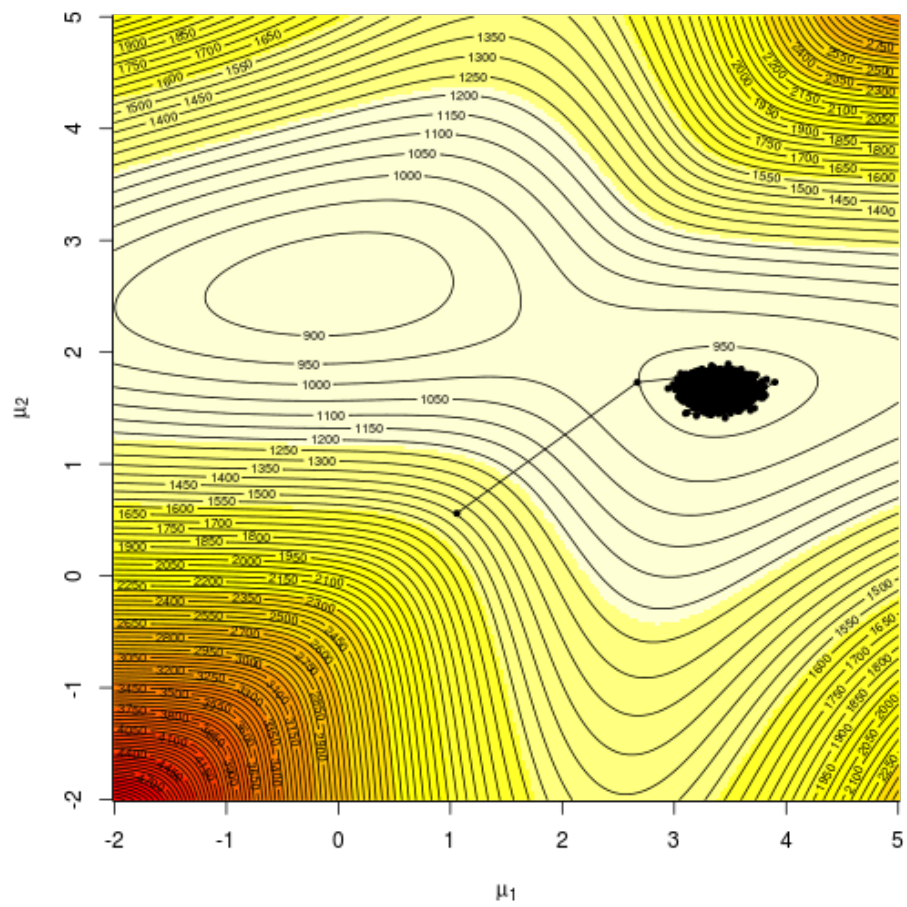


Figure 17: plot of chunk normal\_mixture\_gibbs

```
## Error: object 'nsim' not found
```

```
## Error: error in evaluating the argument 'x' in selecting a method for  
## function 'plot': Error: object 'nsim' not found
```

```
## Error: incorrect number of dimensions
```

```
## Error: incorrect number of dimensions
```

## References