

Data Analysis: Statistical Modeling and Computation in Applications

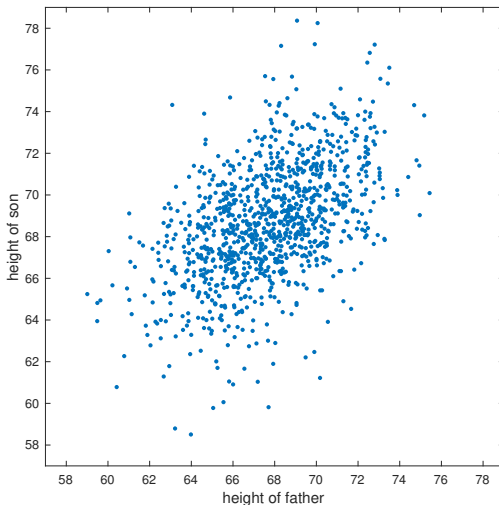
Correlation and Least Squares Regression

Outline

- Correlation
- Regression line
- Evaluation
- Multiple regression
- Computing the estimator
- Variable selection and regularization

Scatter diagram: height of 1078 fathers and their sons

Is there an association?
What kind?

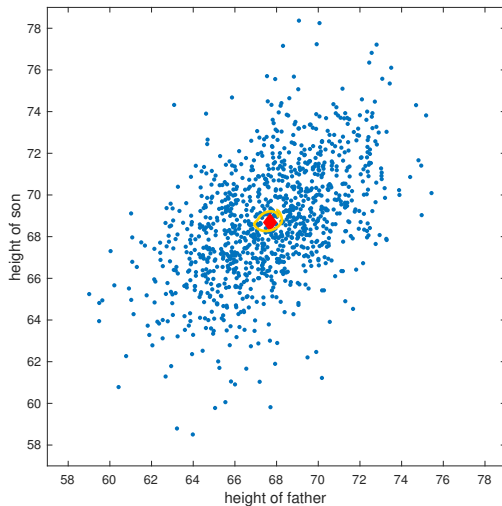


Data: Pearson K and Lee A. (1903). On the laws of inheritance in man. *Biometrika*, 2:357-462.

Downloaded from <https://myweb.uiowa.edu/pbreheny/data/pearson.html>

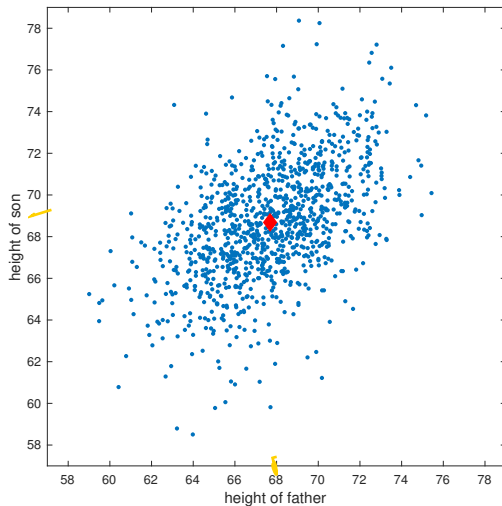
Summarizing the Plot

- average \bar{x} , \bar{y}



Summarizing the Plot

- average \bar{x} , \bar{y}
fathers: $\bar{x} \approx 68$,
sons: $\bar{y} \approx 69$



Summarizing the Plot

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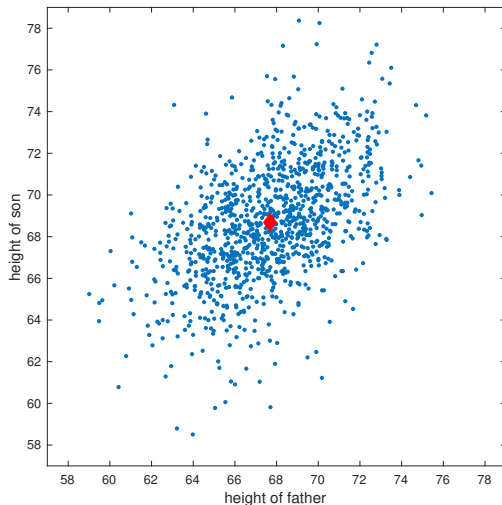
fathers: $\bar{x} \approx 68$,

sons: $\bar{y} \approx 69$

- standard deviation

$$s_x = \frac{1}{N} \sqrt{\sum_{i=1}^N (x_i - \bar{x})^2}$$

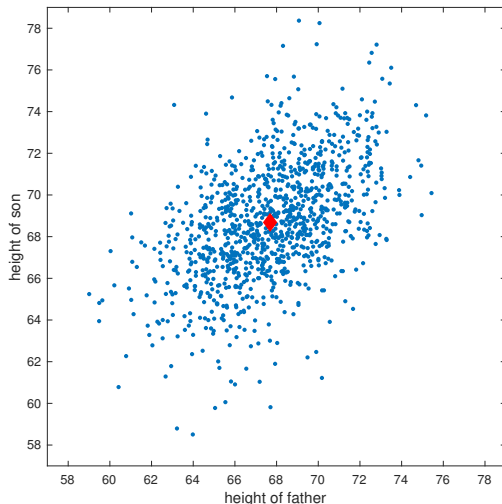
here: $s_x \approx s_y \approx 2.7$



Summarizing the Plot

- average \bar{x} , \bar{y}
fathers: $\bar{x} \approx 68$,
sons: $\bar{y} \approx 69$
- standard deviation
$$s_x = \frac{1}{N} \sqrt{\sum_{i=1}^N (x_i - \bar{x})^2}$$

here: $s_x \approx s_y \approx 2.7$
- correlation coefficient
 $r \approx 0.5$



Correlation Coefficient

$$r = \frac{1}{N} \sum_{i=1}^N \left(\frac{x_i - \bar{x}}{s_x} \right) \left(\frac{y_i - \bar{y}}{s_y} \right) = \frac{\text{cov}(x, y)}{s_x s_y}$$

(convert to standard units and take average product)

Correlation Coefficient

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(convert to standard units and take average product)

- 1 symmetric

Correlation Coefficient

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(convert to standard units and take average product)

- 1 symmetric
- 2 Why standard units?

Correlation Coefficient

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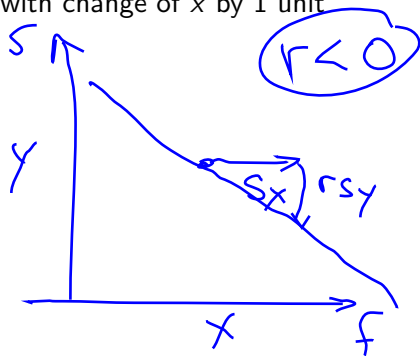
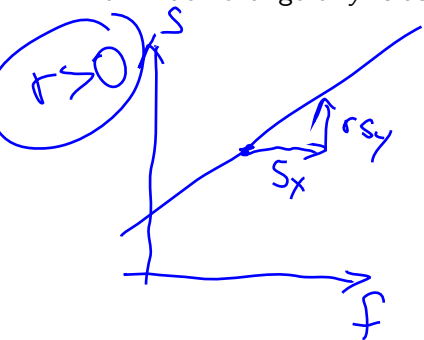
(convert to standard units and take average product)

- ❶ symmetric
- ❷ Why standard units?
adding or multiplying constants to all x_i or y_i does not change r
- ❸ What does $r \approx 0.5$ mean?

What does the Correlation coefficient mean? (1)

$$r = \frac{1}{N} \sum_{i=1}^N \left(\frac{x_i - \bar{x}}{s_x} \right) \left(\frac{y_i - \bar{y}}{s_y} \right)$$

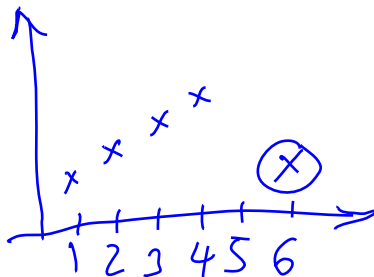
- measures *linear* association between variables:
how much change of y is associated with change of x by 1 unit



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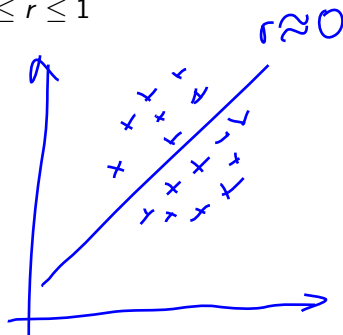
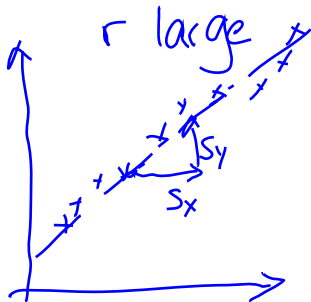
- measures *linear* association between variables:
how much change of y is associated with change of x by 1 unit



What does the Correlation coefficient mean? (2)

$$r = \frac{1}{N} \sum_{i=1}^N \left(\frac{x_i - \bar{x}}{s_x} \right) \left(\frac{y_i - \bar{y}}{s_y} \right)$$

- measures *clusteredness* along a line: $-1 \leq r \leq 1$
sign?



Examples

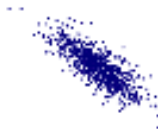
1.



2.



3.



4.



5.



6.



Examples

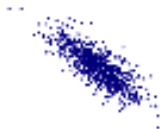
1. $r = 1$



2.



3.



4.



5.



6.



Examples

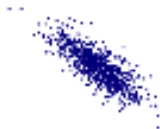
1. $r = 1$



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Examples

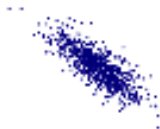
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Examples

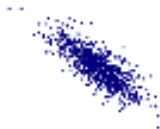
1. $r = 1$



2. $r = -1$



3. $r = -0.8$



4. $r = 0$



5.



6.



Examples

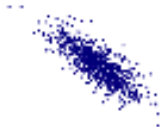
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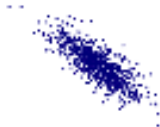


Examples

1. $r = 1$



3. $r = -0.8$



5. $r = 0$



2. $r = -1$



4. $r = 0$



6. $r = 0$



Careful with nonlinearities and outliers!

Correlation coefficient: summary

$$r = \frac{1}{N} \sum_{i=1}^N \left(\frac{x_i - \bar{x}}{s_x} \right) \left(\frac{y_i - \bar{y}}{s_y} \right)$$

- measures *linear* association between variables:
- measures clusteredness along a line
- symmetric (swapping x and y)
- between -1 and 1 , and invariant to
 - adding a constant to all x_i or all y_i
 - multiplying to all x_i (all y_i) by a positive constant

Data Analysis: Statistical Modeling and Computation in Applications

Correlation and Least Squares Regression Part 4

Outline

- Correlation
- Regression line
- Evaluation
- Multiple regression
- Computing the estimator
- Variable selection and regularization

Multiple regression

- Model: $y_i = \beta_0 + x_{i1}\beta_1 + x_{i2}\beta_2 + \epsilon_i$

ozone	radiation	temp
41	190	67
36	118	72
12	149	74
18	313	62

y_i

x_{i1}

x_{i2}

Multiple regression

ozone	radiation	temp
41	190	67
36	118	72
12	149	74
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- Model: $y_i = \beta_0 + x_{i1}\beta_1 + x_{i2}\beta_2 + \epsilon_i$
- vector form: $y_i = \mathbf{x}_i\boldsymbol{\beta} + \epsilon_i$

Multiple regression

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41	190	67
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- Model: $y_i = \beta_0 + x_{i1}\beta_1 + x_{i2}\beta_2 + \epsilon_i$ ←
- vector form: $y_i = \mathbf{x}_i\boldsymbol{\beta} + \epsilon_i$
- Matrix-vector form: $\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$

$$41 = \beta_0 + 190\beta_1 + 67\beta_2 + \epsilon_1$$

$$\begin{pmatrix} 41 \\ 36 \\ 12 \\ 18 \end{pmatrix} = \begin{pmatrix} 1 & 190 & 67 \\ 1 & 118 & 72 \\ 1 & 149 & 74 \\ 1 & 313 & 62 \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{pmatrix} + \begin{pmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \epsilon_4 \end{pmatrix}$$

\mathbf{y} \mathbf{X} $\boldsymbol{\beta}$ $\boldsymbol{\epsilon}$

Multiple regression

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- vector form: $y_i = \mathbf{x}_i\boldsymbol{\beta} + \epsilon_i$
- Matrix-vector form: $\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$
 - \mathbf{y} dependent / response variable: $N \times 1$

$$\underset{\text{Y}}{\mathbf{N}} \begin{pmatrix} 41 \\ 36 \\ 12 \\ 18 \end{pmatrix} = \begin{pmatrix} 1 & 190 & 67 \\ 1 & 118 & 72 \\ 1 & 149 & 74 \\ 1 & 313 & 62 \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{pmatrix} + \begin{pmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \epsilon_4 \end{pmatrix}$$

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 - \mathbf{y} dependent / response variable: $N \times 1$
 - \mathbf{X} design matrix: $N \times p$

$$\begin{pmatrix} 41 \\ 36 \\ 12 \\ 18 \end{pmatrix} \stackrel{N}{=} \begin{pmatrix} 1 & 190 & 67 \\ 1 & 118 & 72 \\ 1 & 149 & 74 \\ 1 & 313 & 62 \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{pmatrix} + \begin{pmatrix} \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \epsilon_4 \end{pmatrix}$$

Note: A blue bracket labeled 'P' is drawn above the design matrix X, indicating its dimensions are N x p.

Multiple regression

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 - $\boldsymbol{\beta}$ parameters: $p \times 1$
 - $\boldsymbol{\epsilon}$: random error / disturbances
 ϵ_i are iid, $\mathbb{E}[\epsilon_i] = 0$, $\text{Var}(\epsilon_i) = \sigma^2$

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Examples of multiple regression

- **Simple linear regression:**

$$p = 2, X = \begin{pmatrix} 1 & x_1 \\ \vdots & \vdots \\ 1 & x_N \end{pmatrix}, \beta = \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix}, y_i = \beta_0 + \beta_1 x_1$$

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$$p = 2, X = \begin{pmatrix} 1 & x_1 \\ \vdots & \vdots \\ 1 & x_N \end{pmatrix}, \beta = \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix}, y_i = \beta_0 + \beta_1 x_1$$

- **Quadratic (polynomial) regression:**

$$p = 3, X = \begin{pmatrix} 1 & x_1 & x_1^2 \\ \vdots & \vdots & \vdots \\ 1 & x_N & x_N^2 \end{pmatrix}, \beta = \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{pmatrix}, y_i = \beta_0 + \beta_1 x_{i1} + \underline{\beta_2 x_{i1}^2}$$

Examples of multiple regression

- **Effect on groups.** Consider an example where we have data obtained on different days. The effect of the days can be modeled as

$$y_i = \underbrace{\beta_0}_{\text{day 1}} + \underbrace{\beta_1}_{\text{day 2}} + \underbrace{\beta_2}_{\text{day 3}} + \epsilon_i$$

Examples of multiple regression

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$$p = 3, \quad X = \begin{pmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ \vdots & \vdots & \vdots \\ 0 & 1 & 0 \\ \vdots & \vdots & \vdots \\ 0 & 0 & 1 \\ \vdots & \vdots & \vdots \end{pmatrix}, \quad \beta = \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{pmatrix}$$

Multiple regression

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$\mathbf{y} = \mathbf{X} \boldsymbol{\beta} + \boldsymbol{\epsilon}$

Ordinary Least Squares estimator (OLS)

- model: $y_i = \beta_0 + x_{i1}\beta_1 + x_{i2}\beta_2 + \epsilon_i$
- fitted values: $\hat{y}_i = \hat{\beta}_0 + x_{i1}\hat{\beta}_1 + x_{i2}\hat{\beta}_2$
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- least squares:

$$\hat{\beta} = \arg \min_{\beta} \sum_{i=1}^N (y_i - \mathbf{x}_i\beta)^2 = \arg \min_{\beta} \underbrace{\|\mathbf{y} - \mathbf{X}\beta\|^2}$$

$$\begin{aligned} \|a\|^2 &= a^T a \\ &= \sum_{j=1}^n a_j^2 \end{aligned}$$

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$$\hat{\boldsymbol{\beta}} = \arg \min_{\boldsymbol{\beta}} \sum_{i=1}^N (y_i - \mathbf{x}_i\boldsymbol{\beta})^2 = \arg \min_{\boldsymbol{\beta}} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2$$

- setting derivative to zero gives *normal equations*

$$(\mathbf{X}^\top \mathbf{X})^{-1} \underbrace{\mathbf{X}^\top \mathbf{X}} \hat{\boldsymbol{\beta}} = \mathbf{X}^\top \mathbf{y}$$

Ordinary Least Squares estimator (OLS)

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$$\hat{\boldsymbol{\beta}} = \arg \min_{\boldsymbol{\beta}} \sum_{i=1}^N (y_i - \mathbf{x}_i\boldsymbol{\beta})^2 = \arg \min_{\boldsymbol{\beta}} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2$$

- setting derivative to zero gives *normal equations*

$$\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} = \mathbf{X}^T \mathbf{y}$$

- if $\mathbf{X}^T \mathbf{X}$ is invertible, then $\hat{\boldsymbol{\beta}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$

$$\begin{aligned} \hat{\mathbf{y}} &= \mathbf{X} \hat{\boldsymbol{\beta}} \\ &= \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} \end{aligned}$$

hat matrix

Ordinary Least Squares estimator (OLS)

- model: $y_i = \beta_0 + x_{i1}\beta_1 + x_{i2}\beta_2 + \epsilon_i$
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$$\hat{\boldsymbol{\beta}} = \arg \min_{\boldsymbol{\beta}} \sum_{i=1}^N (y_i - \mathbf{x}_i\boldsymbol{\beta})^2 = \arg \min_{\boldsymbol{\beta}} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2$$

- setting derivative to zero gives *normal equations*

$$\mathbf{X}^\top \mathbf{X} \hat{\boldsymbol{\beta}} = \mathbf{X}^\top \mathbf{y}$$

- if $\mathbf{X}^\top \mathbf{X}$ is invertible, then $\hat{\boldsymbol{\beta}} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y}$
- fitted values: $\hat{\mathbf{y}} = \mathbf{X}\hat{\boldsymbol{\beta}} = \underbrace{\mathbf{X}(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top}_{\text{"hat matrix"}} \mathbf{y}$

Deriving the normal equations

- least squares objective:

$$f(\boldsymbol{\beta}) = \sum_{i=1}^N (y_i - \mathbf{x}_i \boldsymbol{\beta})^2 = (\mathbf{y} - \mathbf{X}\boldsymbol{\beta})^\top (\mathbf{y} - \mathbf{X}\boldsymbol{\beta})$$

Deriving the normal equations


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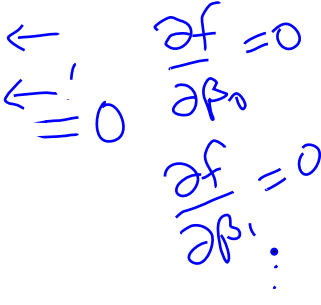
- set gradient to zero. *Gradient* is the vector of partial derivatives:

Deriving the normal equations

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- set gradient to zero. **Gradient** is the vector of partial derivatives:

$$\nabla_{\beta} f(\beta) = \begin{pmatrix} \frac{\partial f}{\partial \beta_0} \\ \frac{\partial f}{\partial \beta_1} \\ \vdots \\ \frac{\partial f}{\partial \beta_{p-1}} \end{pmatrix}$$


If β is $p \times 1$, then $\nabla_{\beta} f(\beta)$ is $p \times 1$.

Partial derivative

- example: 1 data point, $p = 2$:

$$f(\beta) = (y_1 - x_{11}\beta_1 - \beta_0)^2$$

Handwritten blue annotations: A circle around β_1 in the equation, with an arrow pointing down to \hat{y}_1 . To the right, the handwritten equation $\frac{\partial f}{\partial \beta_1} = 0$.

Partial derivative

- example: 1 data point, $p = 2$:

$$f(\beta) = (y_1 - \underline{x_{11}}\beta_1 - \beta_0)^2$$

- derivative:

$$\frac{\partial f}{\partial \beta_1} = -2\underline{x_{11}}(y_1 - \underline{x_{11}}\beta_1 - \beta_0) \stackrel{!}{=} 0$$

Partial derivative

- example: 1 data point, $p = 2$:

$$f(\beta) = (y_1 - x_{11}\beta_1 - \beta_0)^2$$

- derivative:

$$\frac{\partial f}{\partial \beta_1} = -2x_{11}(y_1 - x_{11}\beta_1 - \beta_0)$$

- similarly:

$$\nabla_{\beta} f(\beta) = -2\mathbf{X}^{\top}(\mathbf{y} - \mathbf{X}\beta) \stackrel{!}{=} 0$$

$\mathbf{X}\beta = \mathbf{X}^{\top}\mathbf{X}\mathbf{y}$

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Correlation and Least Squares Regression Part 5

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or $\hat{y}_i = \mathbf{x}_i\hat{\boldsymbol{\beta}}$
- least squares:

$$\hat{\boldsymbol{\beta}} = \arg \min_{\boldsymbol{\beta}} \sum_{i=1}^N (y_i - \mathbf{x}_i\boldsymbol{\beta})^2 = \arg \min_{\boldsymbol{\beta}} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2$$

- setting derivative to zero gives *normal equations*

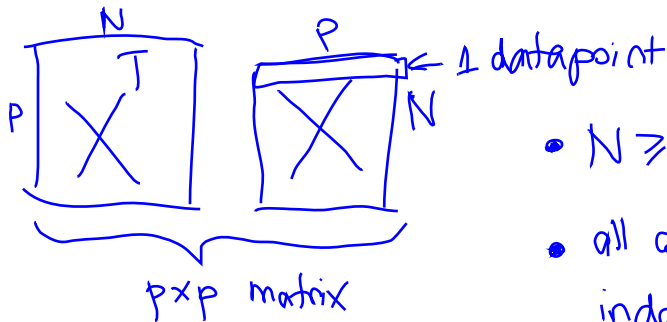
$$\mathbf{X}^\top \mathbf{X} \hat{\boldsymbol{\beta}} = \mathbf{X}^\top \mathbf{y}$$

- if $\mathbf{X}^\top \mathbf{X}$ is invertible, then $\hat{\boldsymbol{\beta}} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y}$



When is $\mathbf{X}^\top \mathbf{X}$ invertible?

- if $\mathbf{X}^\top \mathbf{X}$ has *full rank*:



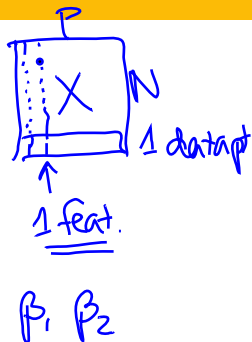
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$$\beta_0 + 2\beta_1 = 5$$

$$N=1$$
$$p=2$$



$$\beta_1 \quad \beta_2$$

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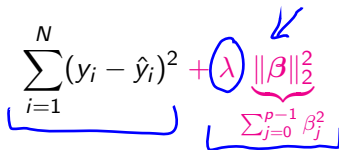
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penalizes large values of β_j
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- ℓ_1 **penalty (Lasso)**: minimize

$$\underbrace{\sum_{i=1}^N (y_i - \hat{y}_i)^2}_{\text{blue}} + \lambda \underbrace{\|\beta\|_1}_{\sum_{j=0}^{p-1} |\beta_j|}$$

prefers *sparse* β (few nonzero coordinates)

Model selection: Which variables to include in the model?

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- **t-test to test $\beta_j = 0$ vs. $\beta_j \neq 0$:** estimate σ^2 as $\hat{\sigma}^2 = \frac{1}{N-p-1} \sum_{i=1}^N (y_i - \hat{y}_i)^2$, then $(N-p-1)\hat{\sigma}^2 \sim \sigma^2 \chi_{N-p}^2$.

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- Use the t-test to determine variables that are not significant. Of those, remove the one with the largest p -value. Re-fit and repeat until all variables have significant p -values.

References

- D. Freedman, R. Pisani, R. Purves. *Statistics*. 2007.
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- D. Freedman. *Statistical Models – Theory and Practice*. 2009.
Chapters 2–4.