

Adjoint-aided inference of Gaussian process driven differential equations

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Inference for complex models

> Nature, 2002 Apr 18;416(6882):726-9. doi: 10.1038/416726a.

Using the fossil record to estimate the age of the last common ancestor of extant primates

Simon Tavaré¹, Charles R Marshall, Oliver Will, Christophe Soligo, Robert D Martin

Syst Biol., 2011 Jan; 60(1): 16-31.

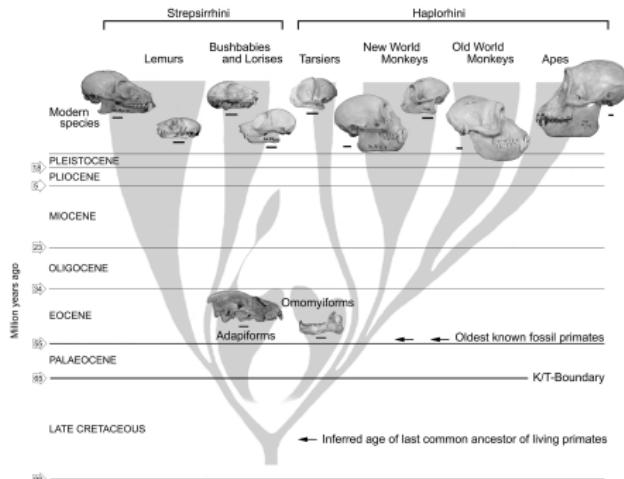
Published online 2010 Nov 4. doi: [10.1093/sysbio/syz054](https://doi.org/10.1093/sysbio/syz054)

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Dating Primate Divergences through an Integrated Analysis of Palaeontological and Molecular Data

Richard D. Wilkinson,^{1,*} Michael F. Steiner,^{2,3,4,5} Christophe Soligo,⁶ Robert D. Martin,⁷ Zheng Yang,⁸ and Simon Tavaré⁹



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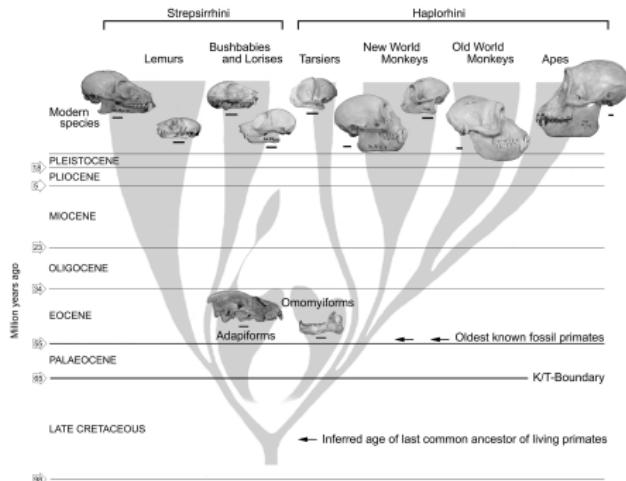
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Genetics, 1997 Feb; 145(2): 505-518.

doi: [10.1093/genetics/145.2.505](https://doi.org/10.1093/genetics/145.2.505)



Inferring Coalescence Times from DNA Sequence Data

S. Tavaré, D. J. Balding, R. C. Griffiths, and P. Donnelly

ABC: given data $D = f(\theta) + e$, find $\pi(\theta|D)$

- Draw θ from $\pi(\theta)$
- Simulate $X \sim f(\theta)$
- Accept θ if $\rho(D, X) \leq \epsilon$

Works for any simulator f – no knowledge required

Project team

Paterne



Engineer



Mike



Mauricio



Chris

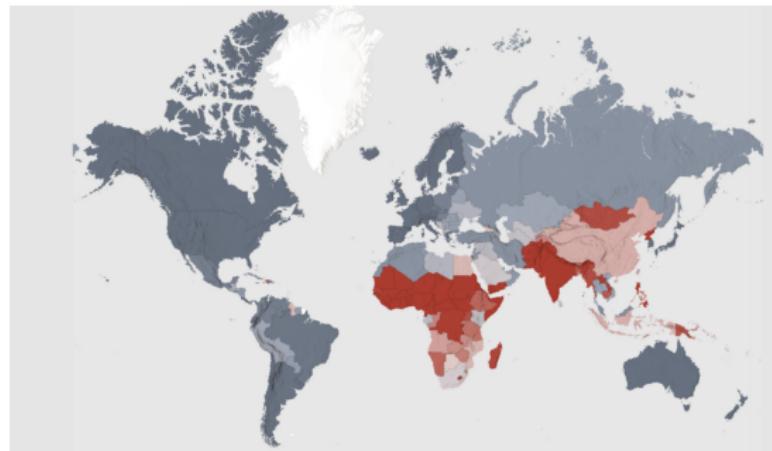


Funders:

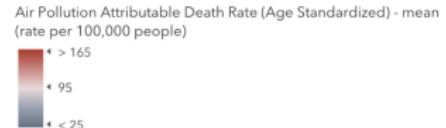


Air pollution

7 million people die every year from exposure to air pollution, the majority in LMICs.



Global Particulate Matter (PM) 2.5 between 1998-2016 - Country

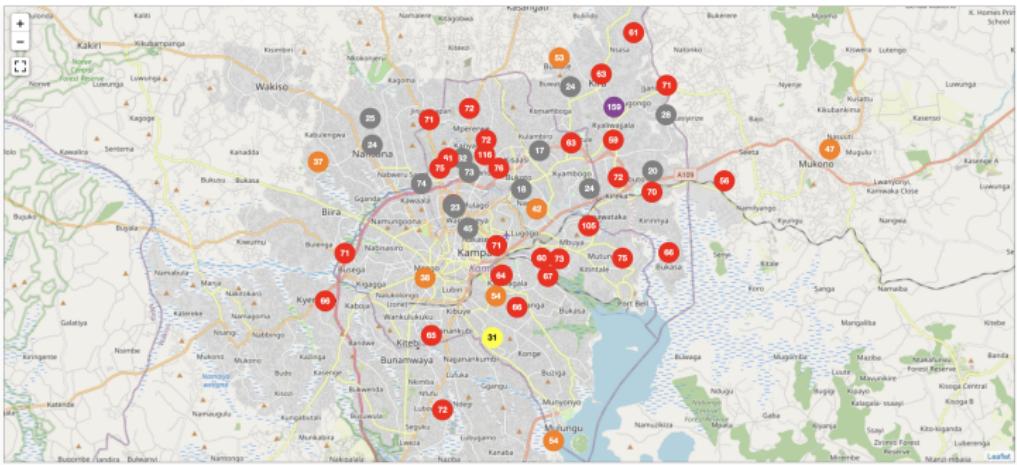


The UK government estimates the annual mortality of human-made air pollution to be 28,000 to 36,000 deaths, and costs UK $\sim \text{£}10^{10}$

Kampala and AirQo



- AirQo, a portable air quality monitor
 - Measures particulate matter
 - Solar powered or other available power sources
 - Cellular data transmission
 - Weather proof for unique African settings



AQI Key



Modelling air pollution

Model pollution concentration $u(x, t)$ at location x at time t .

We want to

- infer air pollution (and predict future pollution levels)
- infer pollution sources

Standard non-parametric models (e.g., Gaussian processes) unable to do this.

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Instead build data models that *know* some physics

$$\frac{\partial u}{\partial t} = \nabla \cdot (\mathbf{p}_1 u) + \nabla \cdot (\mathbf{p}_2 \nabla u) - p_3 u + f$$

Here $f(x, t)$ represents the pollution source.

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Here $f(x, t)$ represents the pollution source.

Given noisy measurements of pollution levels $z_i = h_i(u) + e_i$ can we infer

- the concentration field $u(x, t)$?
- the source $f(x, t)$?
- ...

General linear systems

$$\mathcal{L}u = f$$

Linear systems with unknown parameters

Consider

$$\mathcal{L}u = f$$

where

- \mathcal{L} = linear operator
- f = forcing function.
- u dependent quantity, e.g. pollution concentration.

Finding u given \mathcal{L} and f is the **forward problem**.

Linear systems with unknown parameters

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- \mathcal{L} = linear operator
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Finding u given \mathcal{L} and f is the **forward problem**.

Inverse problem: infer u, f given noisy observations of u

$$z = h(u) + N(0, \Sigma).$$

Note: MCMC likely to be prohibitively expensive: each iteration requires a solution of the forward problem.

Linear systems with unknown parameters

Least squares/maximum-likelihood estimation:

$$\begin{aligned} \min_f \quad & (z - h(u))^\top (z - h(u)) \\ \text{subject to} \quad & \mathcal{L}u = f. \end{aligned}$$

Bayes: find

$$\pi(f|z, \mathcal{L}).$$

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Adjoints can help in both cases

- We can solve both problems with n simulator evaluations, where $n = \text{number of data points.}$

What is an adjoint?

See Estep 2004

Suppose \mathcal{U} and \mathcal{V} are Hilbert spaces

- i.e. vector spaces with an inner product $\langle u, u' \rangle$,
and $\mathcal{L} : \mathcal{U} \mapsto \mathcal{V}$ a linear operator between spaces.

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Thus for all $v \in \mathcal{V}$ we've associated a unique $u \in \mathcal{U}$.

$$\mathcal{L}^* : v \mapsto u.$$

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By definition

$$\langle \mathcal{L}u, v \rangle = \langle u, \mathcal{L}^*v \rangle \text{ the 'bilinear identity'}$$

Example 0

In the finite dimensional case,

$$\mathcal{L}u = Au \text{ for some matrix } A.$$

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Then

$$\mathcal{L}^*v = A^\top v$$

That is

$$\langle Au, v \rangle = \langle u, A^\top v \rangle$$

Efficient inference

$$\mathcal{L}u = f, \quad z_i = h_i(u) + e$$

If the observation operator is linear

$$h_i(u) = \langle h_i, u \rangle$$

we can consider the n adjoint systems

$$\mathcal{L}^* v_i = h_i \text{ for } i = 1, \dots, n.$$

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Then

$$\begin{aligned} h_i(u) &= \langle h_i, u \rangle = \langle \mathcal{L}^*v_i, u \rangle = \langle v_i, \mathcal{L}u \rangle \\ &= \langle v_i, f \rangle, \end{aligned}$$

by the bilinear identity.

$$z_i = h_i(u) + e_i = \langle v_i, f \rangle + e_i$$

$$\text{where } \mathcal{L}^* v_i = h_i$$

Suppose f is a parametric model with a linear dependence upon some unknown parameters q :

$$f(\cdot) = \sum_{m=1}^M q_m \phi_m(\cdot) \quad (1)$$

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$$\text{then } h_i(u) = \langle v_i, \sum_{m=1}^M q_m \phi_m \rangle = \sum_{m=1}^M q_m \langle v_i, \phi_m \rangle.$$

A linear model!

The complete observation vector z can then be written as

$$\begin{aligned} z &= \begin{pmatrix} \langle v_1, \phi_1 \rangle & \dots & \langle v_1, \phi_M \rangle \\ \vdots & & \vdots \\ \langle v_n, \phi_1 \rangle & \dots & \langle v_n, \phi_M \rangle \end{pmatrix} \begin{pmatrix} q_1 \\ q_M \end{pmatrix} + e \\ &= \Phi q + e \end{aligned} \quad (2)$$

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Thus

$$\begin{aligned} \min_f \quad S(f) &= (z - h(u))^\top (z - h(u)) \\ \text{subject to} \quad \mathcal{L}u &= f \end{aligned}$$

is equivalent to

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The solution is

$$\hat{q} = (\Phi^\top \Phi)^{-1} \Phi^\top z$$

with $\text{Var}(\hat{q}) = \sigma^2(\Phi^\top \Phi)^{-1}$ when e_i are uncorrelated and homoscedastic with variance σ^2 .

In a Bayesian setting, if we assume *a priori* that $q \sim \mathcal{N}_M(\mu_0, \Sigma_0)$, then the posterior for q given z (and other parameters) is

$$q | z \sim \mathcal{N}_M(\mu_n, \Sigma_n) \quad (3)$$

where

$$\mu_n = \Sigma_n \left(\frac{1}{\sigma^2} \Phi^\top z + \Sigma_0^{-1} \mu_0 \right), \quad \Sigma_n = \left(\frac{1}{\sigma^2} \Phi^\top \Phi + \Sigma_0^{-1} \right)^{-1}. \quad (4)$$

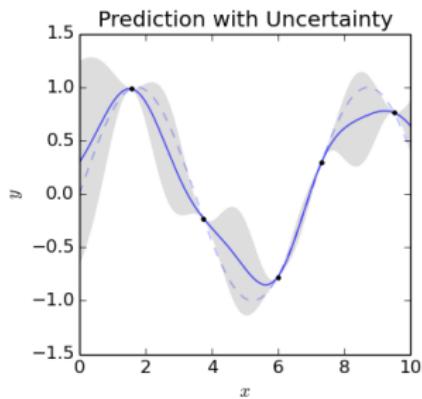
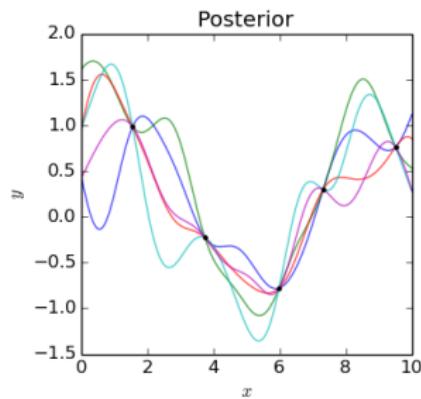
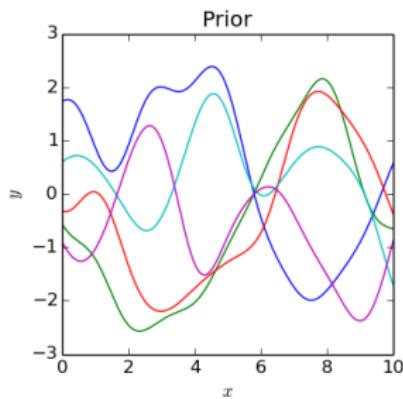
Gaussian Processes

Suppose we model unknown function $f = \{f(x) : x \in \mathcal{X}\}$ as a Gaussian process (GP)

$$f \sim GP(m, k)$$

where we need to specify the prior mean and covariance functions

$$\mathbb{E}f(x) = m(x), \quad \text{Cov}(f(x), f(x')) = k(x, x').$$



Parameterizing GPs

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How can we use GPs within the adjoint framework developed earlier?

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- Let \mathcal{F} be the RKHS (function space) associated with kernel k , i.e., $f \in \mathcal{F}$
- Consider $\{\phi_1(x), \phi_2(x), \dots\}$ an orthonormal basis for \mathcal{F} .

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- Consider $\{\phi_1(x), \phi_2(x), \dots\}$ an orthonormal basis for \mathcal{F} .

We can then approximate f using a truncated basis expansion

$$\begin{aligned} f(x) \approx f_q(x) &= \sum_{j=1}^M q_j \phi_j(x) \text{ where } a \text{ priori } q_i \sim N(0, \lambda_i^2) \\ &= \Phi \mathbf{q} + e \end{aligned}$$

We've approximated the GP by a finite dimensional linear model.

Choice of basis

$$f(x) = \sum_{j=1}^{\infty} q_j \phi_j(x)$$

- **Random Fourier features:** $\phi_i(x) = \cos(w_i x + b_i)$ where $w_i, b_i \sim p(\cdot)$

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- **Random Fourier features:** $\phi_i(x) = \cos(w_i x + b_i)$ where $w_i, b_i \sim p(\cdot)$
- **Mercer basis:** $\phi_i(x) = \lambda_i \psi(x)$ where $\lambda_i, \phi_i(\cdot)$ are eigenpairs of

$$T_k(f)(\cdot) = \int_{\mathcal{X}} k(x, \cdot) f(x) dx.$$

- **Laplacian basis:** useful for non-Euclidean domains...

Example 1: Ordinary differential equation

Consider the ordinary differential equation

$$-D\ddot{u} + \nu\dot{u} + u = f(t) \quad \text{with } u(0) = \dot{u}(0) = 0.$$

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$$\begin{aligned}\langle \mathcal{L}u, v \rangle &= \int_0^T \mathcal{L}u(t)v(t)dt = \int_0^T (-D\ddot{u} + \nu\dot{u} + u)v dt \\ &= [-D\dot{u}v]_0^T + \int_0^T D\dot{u}\dot{v} dt + [\nu uv]_0^T - \int_0^T \nu u\dot{v} dt + \int_0^T u v dt\end{aligned}$$

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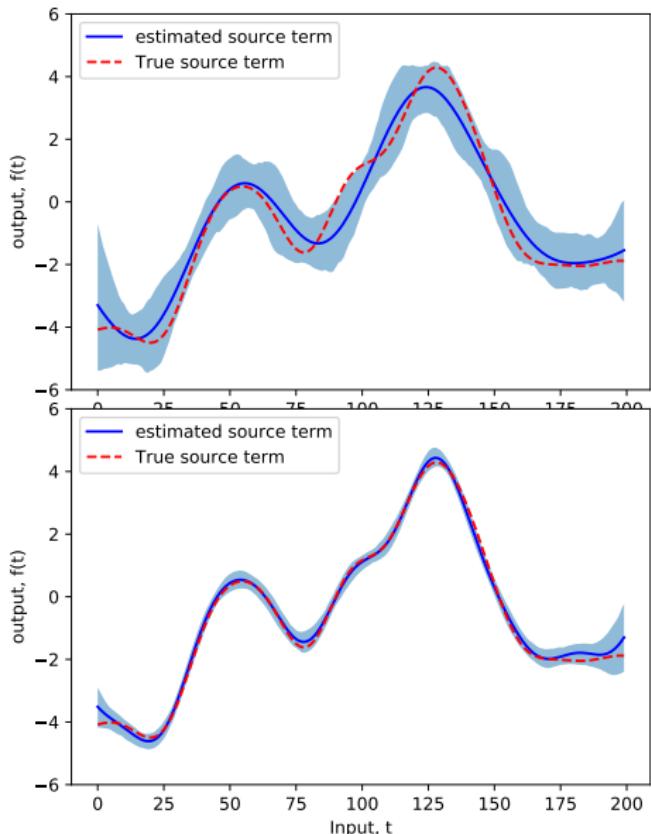
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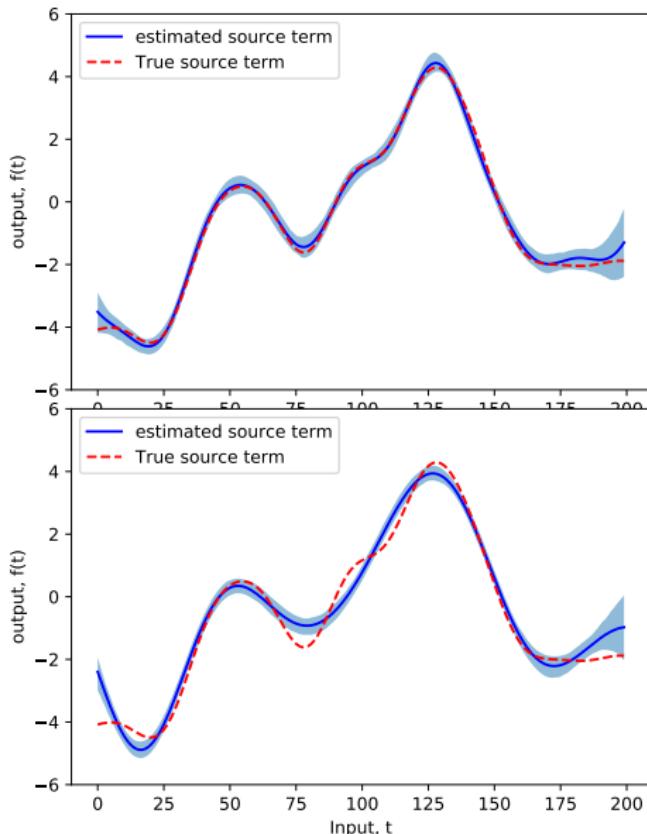
Example 1: Posterior mean and 95% CI (blue), true (red)



- top: $n = 10$ data points, $M = 100$ basis vectors
- bottom: $n = 100$ and $M = 100$

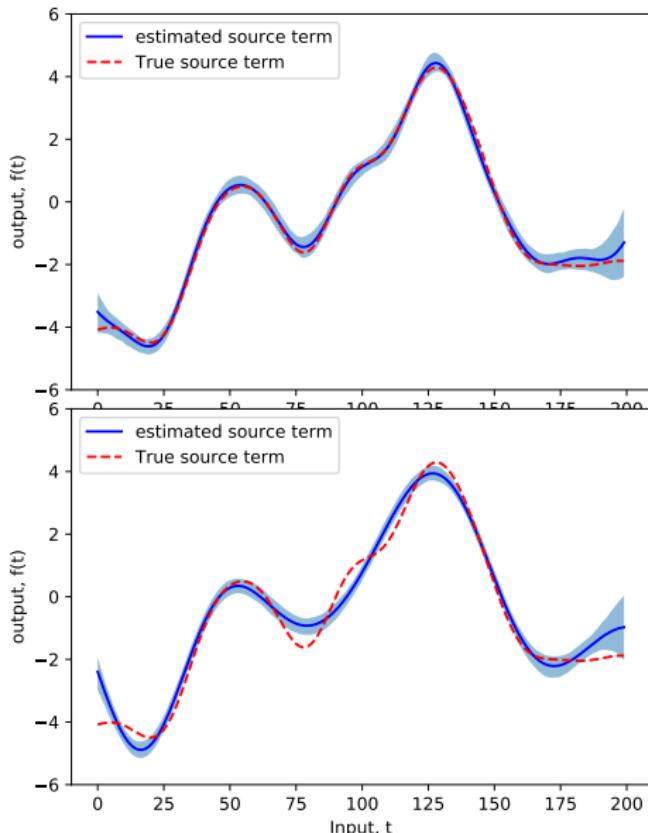
Results required 10 and 100 ODE solves respectively.

Example 1: Too few features



- $n = 100$ data points
- top: $M = 100$ basis vectors
 - bottom: $M = 10$
- NB: overconfident and wrong when $M = 10$ - misspecified model!

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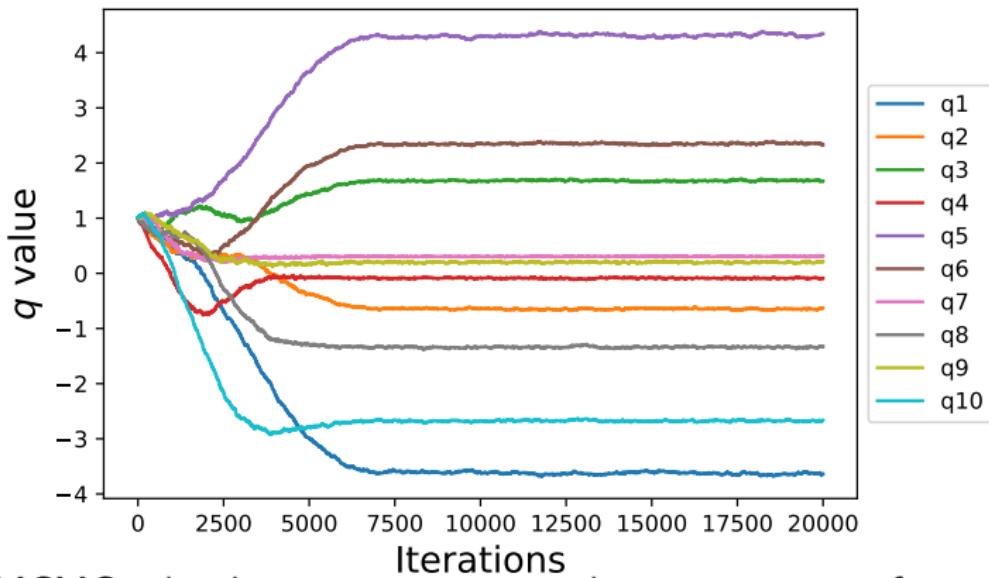
NB: overconfident and wrong when $M = 10$ - misspecified model!

We need to include enough features to have sufficient modelling flexibility.

Using additional features doesn't require additional ODE solves.

MCMC is fine as long as you have a small number of features.

But even with only 10 features, we need ~ 1000 s of ODE solves vs 10 ODE solves for the adjoint method.



MCMC takes longer to converge when we use more features.

Example 2: PDE

Advection-diffusion-reaction is a linear operator:

$$\mathcal{L}u = \frac{\partial u}{\partial t} - \nabla \cdot (\mathbf{p}_1 u) - \nabla \cdot (p_2 \nabla u) + p_3 u$$

Forward problem: solve (for some initial and boundary conditions)

$$\mathcal{L}u = f \text{ on } \mathcal{X} \times [0, T].$$

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and estimate q given $z_i = \langle h_i, u \rangle + N(0, \sigma)$.

Typically h_i will be a sensor function that might average the pollution at a specific location over a short window

$$\langle h_i, u \rangle = \frac{1}{|\mathcal{T}|} \int_{\mathcal{T}} u(x_i, t) dt$$

Example 2: PDE adjoint

The adjoint system is again derived by integrating by parts twice:

$$\mathcal{L}^*v = -\frac{\partial v}{\partial t} - \mathbf{p}_1 \cdot \nabla v - \nabla \cdot (p_2 \nabla v) + p_3 u.$$

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For n observations we need n adjoint equations!

$$\mathcal{L}^*v_i = h_i \text{ in } \mathcal{X} \times [0, T] \text{ for } i = 1, \dots, n.$$

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$$\mathcal{L}^* v = -\frac{\partial v}{\partial t} - \mathbf{p}_1 \cdot \nabla v - \nabla \cdot (p_2 \nabla v) + p_3 u.$$

For n observations we need n adjoint equations!

$$\mathcal{L}^* v_i = h_i \text{ in } \mathcal{X} \times [0, T] \text{ for } i = 1, \dots, n.$$

If we use initial and boundary conditions

$$u(x, 0) = 0 \text{ for } x \in \mathcal{X} \text{ and } \nabla_n u = 0 \text{ for } x \in \partial \mathcal{X}$$

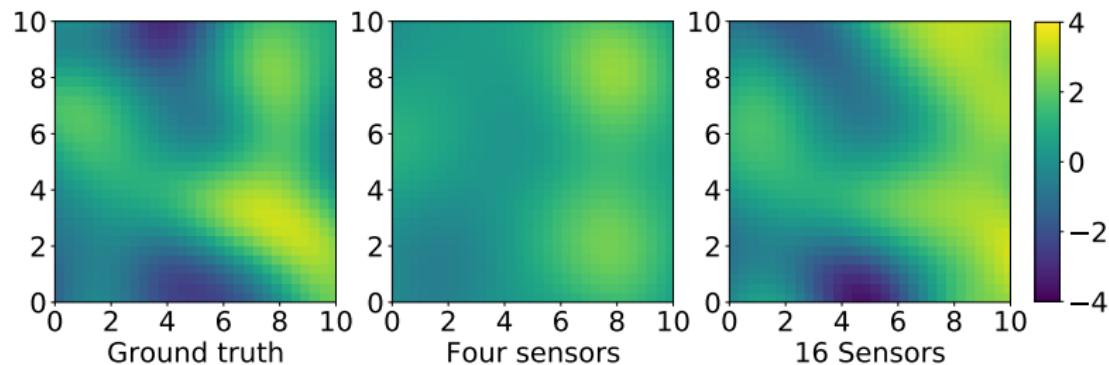
then the final and boundary conditions on the adjoint system are

$$v_i(x, T) = 0 \text{ for } x \in \mathcal{X}$$

$$\mathbf{p}_1 v_i(x, t) + p_2 \nabla v_i(x, t) = 0 \text{ for } x \in \partial \Omega \text{ and } t \in [0, T].$$

Results: $n = 20$ (4 sensors) and $n = 80$ (16), noise = 10%

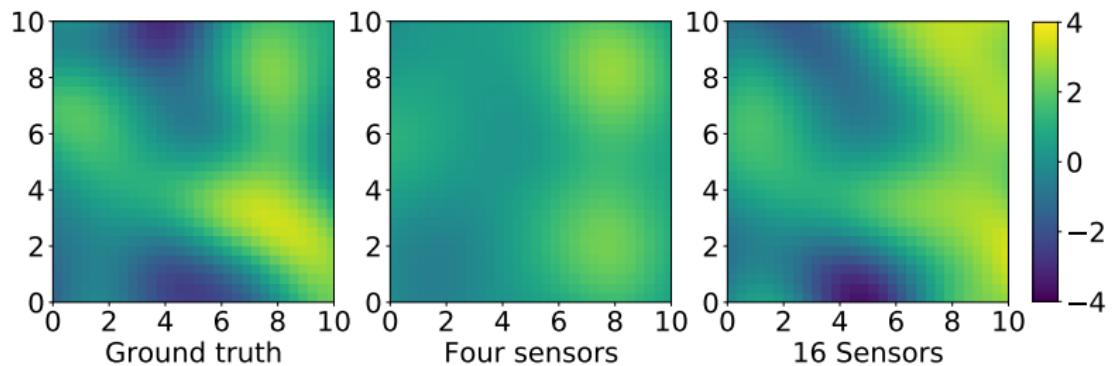
Posterior mean of time slice $u(x, 5)$ - more sensors, improved estimates!



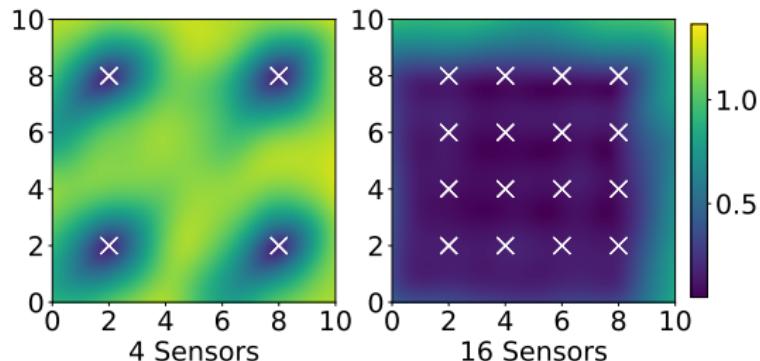
Variance of $u(x, 5)$: Wind from the south west.

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Conclusions

Adjoints of linear systems

- an intrusive method; development does require some work but can be automated
- Requires n adjoint solves to infer the posterior
 - ▶ essentially insensitive to the number of basis functions used
 - ▶ In contrast, MCMC requires a typically an *a priori* unknown number of simulations (but is largely independent of n).
- Gives numerically stable derivatives of the cost function with respect to other parameters, $\frac{dS}{dp}$ etc.
- Opportunities for additional efficiencies...
 - ▶ Efficient use of adjoint simulations
 - ▶ Multi-level approaches
 - ▶ Gradient based optimization
 - ▶ Sequential data

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Thank you for listening!

Link to Green's function approach

Consider the linear system

$$\mathcal{L}u = f \quad \text{for } x \in \Omega$$

The Green's function for this system, $G_y(x)$, satisfies

$$\mathcal{L}^* G_y(x) = \delta_y(x) \quad \text{for } x \in \Omega$$

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If $f \sim GP(0, k)$, then u is also distributed as a Gaussian process,

$$u \sim GP(0, k_u)$$

with covariance function

$$k_u(y, y') = \int G_y(x) \int G_{y'}(x') k(x, x') dx' dx.$$

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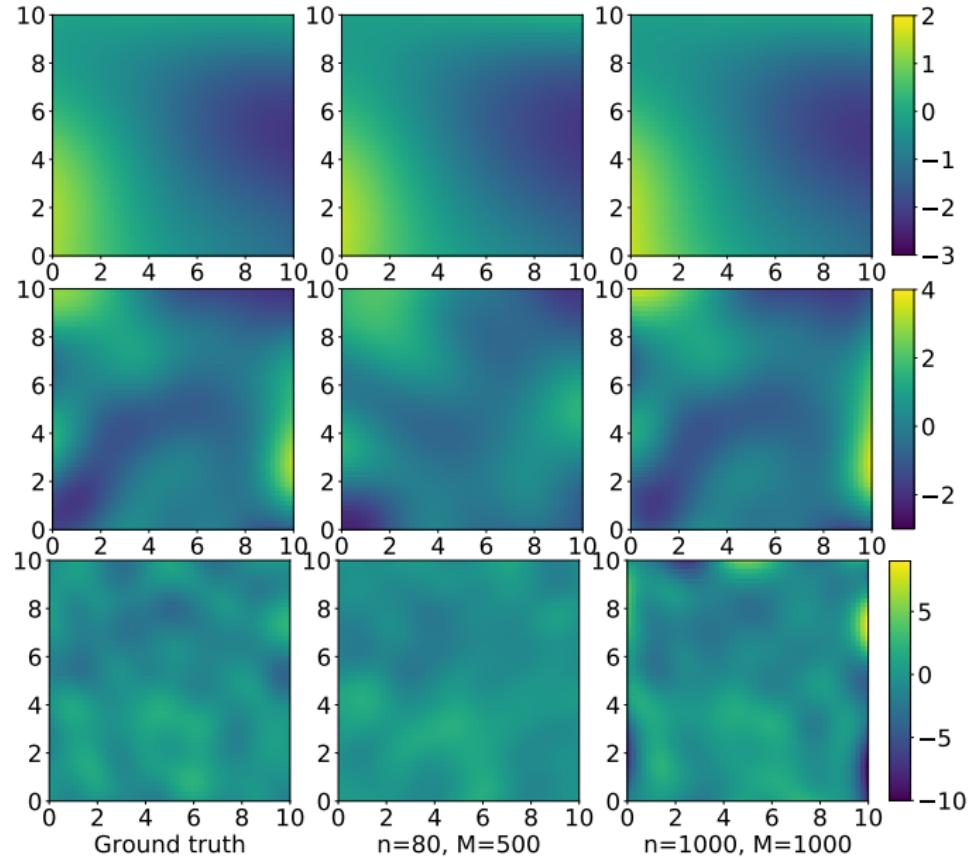
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In contrast, our approach relies on

- existence of the adjoint operator \mathcal{L}^*
- ability to solve adjoint systems numerically - deploy modern finite element solvers (efficient, stable, and offer good error-control).

Recommendation: Use Green's function approach only when G known and covariance integral tractable.

Effect of length scale, $\lambda = 5, 2, 1$



Example 2: Results

Mean square error vs number of features and sensors

Median MSE as a function of number of sensors and RFFs.

Sensors

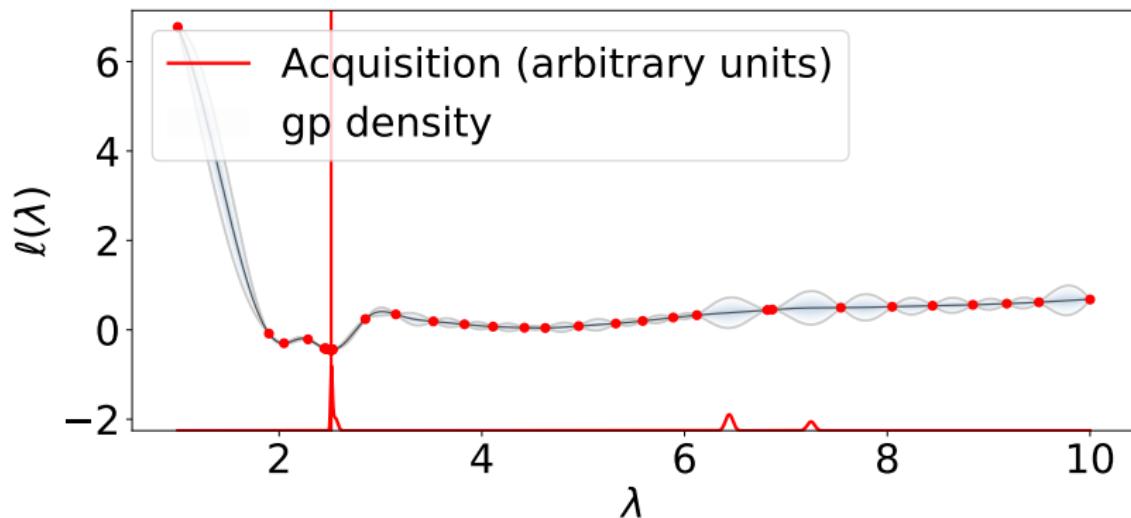
Features

	10	50	100	200	300
1	3.42 (2.82,4.39)	3.27 (3.13,3.38)	3.24 (3.10,3.37)	3.27 (3.17,3.44)	3.24
4	7.12 (1.57,28.81)	2.39 (2.06,2.62)	2.41 (2.13,2.60)	2.45 (2.32,2.57)	2.50
9	2.38 (1.41,4.40)	2.12 (1.48,3.98)	1.70 (1.49,2.07)	1.48 (1.40,1.72)	1.47
16	1.73 (1.23,3.28)	3.99 (2.32,10.90)	2.18 (1.72,3.54)	1.3 (1.02,1.68)	1.12
25	1.35 (1.19,3.09)	8.93 (4.92,39.86)	4.36 (2.53,8.20)	1.86 (1.43,2.75)	1.35
25 (MH)	3.27 (1.73,6.12)	-	-	-	-

MH algorithm did not converge after 20,000 iterations for 50 or more RFFs.

Non-linear parameter estimation

A naive way to estimate the non-linear parameters is via Bayesian optimization iteration



Preprint showing how to use the adjoint sensitivity soon....

Example 1: Matrix system

Suppose $X = Y = \mathbb{R}^d$. A linear operator $\mathcal{L}_p : X \rightarrow Y$ can be written as

$$\mathcal{L}_p x = A_p x \text{ where } A_p \in \mathbb{R}^d$$

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The **adjoint operator** is

$$\mathcal{L}_p^* y = A_p^\top y$$

as we can see that

$$\begin{aligned}\langle A_p x, y \rangle &= (A_p x)^\top y \\ &= x^\top (A^\top y) \\ &= \langle x, A_p^\top y \rangle\end{aligned}$$

Sensitivity

Consider the quantity of interest (QoI)

$$h(x) \equiv \langle g, x \rangle = g^\top x$$

for some $g \in \mathbb{R}^d$, where x is the solution to $h(x, p) := f - Ax = 0$.

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Define Lagrangian the

$$L = g^\top x + y^\top h(x, p)$$

Think of $y \in \mathbb{R}^d$ as Lagrange multipliers.

$$L = g^\top x + y^\top h(x, p)$$

Differentiating with respect to p gives

$$\frac{dL}{dp} = g^\top \frac{dx}{dp} + y^\top \left(\frac{dh}{dx} \frac{dx}{dp} + \frac{dh}{dp} \right)$$

This is true for all y , so if we set $g^\top + y^\top \frac{dh}{dx} = 0$ then we get

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- Autodiff software (eg TensorFlow, JAX etc) will give us this, but can be unreliable for differential equations with long iterative loops

Non-identifiable linear model

Let

$$A_p = \begin{pmatrix} 2 + p_2^2 & -1 \\ 1 & 1 + p_1^2 \end{pmatrix} \text{ and } f_q = \begin{pmatrix} q_1 \\ q_2 \end{pmatrix} = q_1 \begin{pmatrix} 1 \\ 0 \end{pmatrix} + q_2 \begin{pmatrix} 0 \\ 1 \end{pmatrix}$$

and suppose we're given 4 observations with

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Consider the solution to the unconstrained optimization problem.

$$x^* = \arg \min_x (z - G^\top x)^\top (z - G^\top x)$$

The basis functions used for f form a complete basis for \mathbb{R}^2 , and we can always find a q so that $A_p x^* = f_q$ (for all p as A_p is invertible).