

Self-employment and Labor Market Risks

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Abstract

I study the labor market risks associated with being self-employed. I document that the self-employed are subject to larger labor income fluctuations than employees and that they frequently transition into unemployment. Given the self-employed are not eligible to unemployment insurance, I analyze the provision of benefits targeted at these risks using a calibrated search model with (i) precautionary savings, (ii) work opportunities in paid- and self-employment, (iii) skill heterogeneity. This exercise suggests that extending the current US unemployment insurance scheme to the self-employed does not improve the welfare of the low-skill self-employed while making their unemployment spells longer. At the calibrated parameters, the tax burden of financing these benefits limits the welfare gains of the low-skill self-employed.

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1 Introduction

Job loss ranks amongst the most significant risks that workers face over the course of their career. Many countries target sizable transfers to the unemployed in the form of unemployment insurance (UI), and a substantial literature aims at characterizing the optimal UI contract.¹

By contrast, while the self-employed account on average for fifteen percent of employment across OECD countries, there is little evidence on the labor market risks associated with self-employment.² And, because the majority of the self-employed are not eligible to UI in these countries, traditional social insurance programs are not well-designed to alleviate these risks.³

This paper studies the labor market risks associated with self-employment and the provision of benefits targeted at these risks. I first provide new empirical evidence on the earnings risks faced by the self-employed using US monthly survey data. The use of data at monthly frequency is important to accurately measure the drivers of these risks for the self-employed. It allows to separate earnings fluctuations associated with a given self-employment spell from transitions to unemployment or wage work. I show that: (i) earnings are substantially more volatile during self-employment spells than during paid-employment spells; (ii) there are frequent direct transitions from self-employment to unemployment.

My second contribution is to build a framework to assess several benefit schemes

¹Following Jacobson et al. (1993), many studies have confirmed that job loss has long-term negative effects on workers' earnings. Examples of work on the optimal UI contract include Acemoglu and Shimer (1999), Chetty (2008), and Kolsrud et al. (2018).

²Throughout the paper, I use the terms "self-employed" and "self-employment" to designate all workers who get most of their labor income from a business. This definition is further clarified when I introduce the data in Section 2.

³Among these countries, a handful offer some form of public unemployment insurance for some narrow group of self-employed workers, such as artists and writers in Germany. These schemes are reviewed in details in OECD (2018). In the US, some owners of incorporated businesses (S- and C-corps) can become eligible under very specific conditions. Sole proprietors, partnership-owned business, Limited Liability Corporations, and independent contractors do not qualify. This restriction was abolished as part of the US government response to Covid-19 (CARES Act).

targeted at these risks. I develop and calibrate a search model with precautionary savings that is consistent with the patterns of labor market risks found in the data. I use this framework to quantify the welfare changes and distortions in labor supply resulting from the introduction of these schemes.

A key pre-requisite to my calibration is to allow for substantial worker heterogeneity in the model. This heterogeneity is important because the data show that self-employed workers are over-represented in the tails of the earnings distribution, so the underlying skills are likely to differ widely across workers. To discipline this feature of the model, I follow Bonhomme et al. (2019) and use a k-means algorithm to partition workers based on their observed labor income and likelihood to become unemployed. These two measures are strongly correlated in the resulting clusters of workers. For instance, relative to the high earnings group in self-employment, the low earners are close to five times as likely to make a transition from self-employment to unemployment. Building heterogeneity in such a way into the model allows to study how self-employed workers at different skill levels respond to the introduction of additional benefit entitlements.

The model is calibrated to replicate the empirical evidence on the exposure of the self-employed to labor market risks. Paid-employment and self-employment each come with specific labor income and unemployment shocks. In line with the data, the model captures the substantial direct flows between paid-employment, self-employment, and unemployment, thus allowing for moves between paid- and self-employment to represent a response to unemployment shocks or bad realizations of earnings. In addition, workers can also partially insure against labor market risks by borrowing and drawing down their savings, and by relying on household-level income (spousal income and welfare transfers). My calibration matches the substantial fraction of households with low wealth holdings in the data. Within the model, many self-employed households therefore have limited means to self-insure.

I use the calibrated model to study several policies designed to alleviate the labor

market risks associated with self-employment. I center on two broad classes of policies: a UI scheme for the self-employed and a basic income scheme. This exercise suggests that effectively improving the welfare of the self-employed at the bottom of the earnings distribution requires some degree of progressivity in financing these benefits. My results indicate that extending the UI system currently in place in the US, including the cap on income used as a basis to compute unemployment tax liabilities, has barely any impact on the welfare of low earners. This result stems from the relatively large tax burden falling on the lowest earners, though the overall policy is still actuarially fair for them. At the calibrated parameters, the benchmark system is regressive: the lowest-earners face the largest average contribution rate, limiting the welfare gains achieved by this policy.

I further study a clearly redistributive policy to illustrate the trade-offs in providing insurance to the lowest earners in self-employment. A modest basic income scheme is clearly welfare-enhancing for the low-earners group, and it reduces the time they spend unemployed. At the top of the earnings distribution, however, this type of unconditional transfer results in substantial welfare losses for high earners, who bear the burden of financing the policy.

More generally, the analysis developed in this paper is relevant for the design of policy in several dimensions. Like regular paid employees, first, a fraction of the self-employed are not insulated from labor market risks. Second, the rise of alternative work arrangements with the emergence of labor platforms (firms that match workers to customers without being bound to them by an employment contract) may further increase the number of self-employed with low earnings in the future. My framework highlights some of the key trade-offs in providing additional benefits to this group of workers.

Related literature. This work is related to the large literature studying the risks faced by wage workers in the labor market, most notably unemployment, and the associated optimal provision of unemployment insurance (Chetty, 2008; Kolsrud et al., 2018). My

approach is closest to the series of papers that study UI benefits within the context of a fully specified structural model (Hansen and Imrohoroglu, 1992; Acemoglu and Shimer, 1999; Lentz, 2009; Krusell et al., 2010). I depart from these studies by focusing on the labor market risks specifically associated with being self-employed.

This work also relates to the extensive literature estimating income processes (e.g., Meghir and Pistaferri, 2004; Guvenen, 2009) and the degree of insurance households can achieve in response to the resulting income shocks (e.g., Blundell et al., 2008; Kaplan and Violante, 2010). While the majority of papers in this literature drop the self-employed, whose earnings are found to be substantially more volatile, I specifically center on this category of workers.⁴ Using data at monthly frequency, I can further unpack the drivers of earnings fluctuations by separately studying transitions between employment states and earnings shocks within paid-employment and self-employment spells.

This paper also contributes to the growing body of work on self-employment. This literature primarily studies the decision to become self-employed (Hamilton, 2000; Hurst and Lusardi, 2004; Levine and Rubinstein, 2017; Humphries, 2017; Catherine, 2019; Jones and Pratap, 2020). This line of research does not directly consider the insurance dimension of self-employment at the household level. Perhaps the closest paper to mine in that regard is Catherine (2019), who quantifies the value of paid-employment as a back up option for workers deciding whether to start a business. This last paper uses data on yearly earnings and does not consider the risk of becoming unemployed.

A branch of the self-employment literature also studies self-employment as a path out of unemployment by analyzing the outcomes of the ensuing businesses (Hombert et al., 2020; Camarero Garcia and Murmann, 2020). While this channel is present in my model, I instead stress the risks to earnings conditional on workers being self-employed. I can then evaluate several policies targeted at these risks within my framework.

⁴In the income process section of their review paper on earnings and consumption, Meghir and Pistaferri (2011) write: “the focus is mainly on employed workers and self-employed workers are typically also dropped.”

Outline. The next section presents the data. Section 3 introduces the model. Section 4 discusses the calibration procedure. Section 5 describes the policy experiments, and Section 6 concludes.

2 Data

This section provides empirical evidence on the earnings risks associated with being self-employed. An accurate description of these risks in the data requires: (i) frequent records of each individual’s labor market history to capture transitions in and out of unemployment, (ii) information on household income and wealth, which represent key self-insurance channels, and (iii) a large enough sample, since the self-employed represent a small fraction of total employment. The Survey of Income and Program Participation meets these requirements (Census Bureau, 2014, SIPP thereafter). I pool together the four panels spanning 1996-2013.

I stress two key dimensions of earnings risks in these data: shocks to labor income while employed (an intensive margin risk) and unemployment (an extensive margin risk). I start by showing how individuals with at least some experience of self-employment over the sample period differ from individuals always working as paid employees. I then document the drivers of earnings risks for the self-employed.

2.1 The self-employed in the SIPP

The sample is restricted to the working-age individuals with the largest earnings in each household over the duration of the survey. This restriction is made both to center on the individuals with the strongest ties to the labor market and to account for household-level insurance channels, since shared assets, earnings from other household members, and welfare programs represent additional sources of insurance against earnings risks. I incorporate these additional sources of insurance in the model introduced in Section 3.

Workers are assigned to one of three labor market states in each month: paid-employed (P), self-employed (S), or unemployed (U). When in employment, they are categorized as paid- or self-employed based on their primary source of earnings in a month. A person can therefore be paid- or self-employed in different months over the duration of the survey. I choose to focus on earnings as hours appear to be imprecisely measured in the SIPP. Earnings also more directly relate to the resources workers can actually set aside in anticipation of labor market shocks. For an overwhelming majority of workers, this definition clearly singles out one labor form (paid- or self-employment) as an individual's main source of earnings (see Figure C.2). I use a standard definition of unemployment, based on whether workers actively search for work over the duration of a non-employment spell. Appendix A provides a complete description of the data and the definition of each labor market state.

Defining the self-employed as workers deriving most of their earnings from their business encompasses individuals working in very different ways, both in terms of scale and industry. It groups freelancers, contractors, doctors and lawyers, as well as owner-managers with potentially many employees. However, all these workers have in common that they are not the employee of a distinct person (in legal terms, some of them may pay themselves a wage through their business). The model introduced in subsequent sections allows for permanent worker heterogeneity in skills to capture the range of situations covered by this definition.

This definition of self-employment is broader than the typical plant or establishment level dataset, which are restricted to production units with paid employees.⁵ In the data, the majority of businesses from which self-employment income is derived are relatively limited operations.⁶ Table 1 reports some key business characteristics. In the self-employed sample, the reported median business assets for workers in self-employment are slightly

⁵See the body of work in the tradition of Davis et al. (1998).

⁶This is consistent with the evidence reported in Hurst and Pugsley (2011), who find that most new business owners do not intend to grow or bring new products to the market.

| Business characteristic | Value |
|---|-------|
| Reported business wealth (Median, \$2009) | |
| Assets | 6,355 |
| Debt | 0 |
| Equity | 4,770 |
| Share incorporated | 0.320 |
| Share with more than 25 employees | 0.050 |

Table 1: Business characteristics in the SIPP. Note: Business wealth is obtained by multiplying the owner’s share and summing the data for all businesses, if several businesses are reported. The incorporation and employment variables take value one if any of the reported businesses satisfies the condition.

above \$6,000. Only 32% of the corresponding businesses are incorporated, and only 5% of the self-employed declare that their business has ever had more than 25 employees since it was established.

2.2 Self-employed and paid employees

I split the sample of main earners into two sub-samples: The “self-employed” sample, which is made of workers who are self-employed at some point during the survey, and the “paid employees” sample, which is made of workers always working as employees over the duration of the survey. The self-employed sample represents about fifteen percent of the entire sample of main earners.⁷

Table 2 gives some basic demographic statistics on these two groups of workers. The self-employed are more likely to be married older men. They are less likely to belong to a minority. These two groups are otherwise similar in terms of education level, with the self-employed slightly more likely to hold a post-graduate qualification.

Table 3 compares the paid employees and self-employed samples in terms of income

⁷Fifteen percent is higher than the share of self-employment in total employment implied by the data from the Bureau of Labor Statistics. The sample is, however, selected on labor market attachment. Figure C.1a confirms that the self-employment rate derived from the SIPP is very close to the BLS series when taking all working age individuals into account.

| | Paid employees | Self-employed |
|-------------------------|----------------|---------------|
| Age | 39.864 | 42.575 |
| Gender (woman=1) | 0.415 | 0.263 |
| Married (married=1) | 0.539 | 0.616 |
| Race (non-white=1) | 0.196 | 0.138 |
| HS Graduate | 0.902 | 0.894 |
| College Graduate | 0.454 | 0.474 |
| Post-graduate | 0.099 | 0.130 |
| Number of individuals | 105,178 | 18,079 |
| Average months in panel | 35.663 | 36.960 |

Table 2: Demographic characteristics by subsample.

and wealth. All monetary values are given in 2009 real dollars throughout. To account for differences in monthly fluctuations in income across samples, all income measures are reported as twelve-month averages (including any potential zeros). In line with prior work, the self-employed sample is characterized by more heterogeneity along both the income and wealth margins (Hamilton, 2000; Åstebro et al., 2011). For instance, the tenth percentile of total household income is \$1,186 (\$1,891) and the ninetieth percentile is \$13,204 (\$10,772), respectively for workers with (without) self-employment.

In terms of wealth, a substantial fraction of the self-employed have not accumulated more wealth than paid employees. The median net liquid wealth (net worth excluding business, home, and vehicle equity, aggregated at the household level) is approximately the same for the self-employed and employees (around \$3,000). Access to credit also appears similar: the distribution of unsecured debt is very close in both samples. Taken together, these statistics suggest that, in terms of household finance, a substantial fraction of the self-employed is no less insulated from earnings risks than paid employees.

Figure 1 shows the distribution of yearly income growth year-on-year in each sample. Large changes in income are markedly more common in the self-employed sample. This pattern is found both for the labor income of the main earner (Figure 1a) and total household income (Figure 1b). Across these different measures, the self-employed appear

| | | Percentile | | | | |
|--------------------|----------------|------------|--------|-------|--------|---------|
| | Sample | p10 | p25 | p50 | p75 | p90 |
| Income (\$2009) | | | | | | |
| (12-month average) | | | | | | |
| Main earner | Paid employees | 997 | 1,861 | 3,102 | 4,828 | 7,113 |
| | Self-employed | 357 | 1,157 | 2,582 | 4,951 | 8,952 |
| Other earners | Paid employees | 0 | 0 | 297 | 2,027 | 3,587 |
| | Self-employed | 0 | 0 | 58 | 1,671 | 3,436 |
| Welfare | Paid employees | 0 | 0 | 0 | 320 | 1,319 |
| | Self-employed | 0 | 0 | 0 | 256 | 1,293 |
| Household | Paid employees | 1,891 | 3,082 | 4,953 | 7,545 | 10,772 |
| | Self-employed | 1,186 | 2,409 | 4,542 | 7,957 | 13,204 |
| Wealth (\$2009) | | | | | | |
| Net liquid wealth | Paid employees | -12,572 | -885 | 3,296 | 49,859 | 185,564 |
| | Self-employed | -16,254 | -1,323 | 2,992 | 67,000 | 256,079 |
| Unsecured debt | Paid employees | 0 | 0 | 1,472 | 8,887 | 23,983 |
| | Self-employed | 0 | 0 | 1,408 | 9,834 | 27,693 |

Table 3: Earnings and wealth: self-employed vs employee sample.

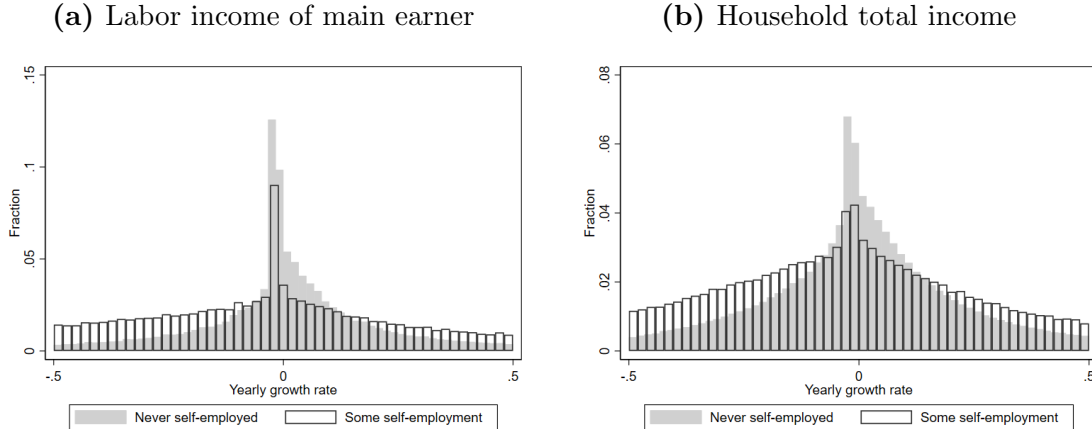


Figure 1: Income growth year-on-year: self-employed vs paid employee sample.

to be exposed to more substantial earnings risks than paid employees.

2.3 The self-employed’s exposure to labor market risks

I describe the drivers of earnings changes in the self-employed sample. These empirical regularities represent some of the key moment targets used to calibrate the model introduced in Section 3.

Table 4 gives information on the monthly flows between paid-employment (P), self-employment (S), and unemployment (U). In the self-employed sample, workers are about equally likely to exit unemployment in paid-employment (UP rate of 14%) or self-employment (US rate of 9.5%). The implied aggregate transition rate from unemployment (23.5%) is of the same magnitude as in the paid employee samples (21.1%).⁸

In terms of transitions to unemployment, the self-employed do make direct transitions from self-employment to unemployment (SU rate of 0.5%). While their chance to become unemployed after a spell as an employee is larger (PU rate of 2.4%), workers in the self-employed sample are on average markedly more likely to experience a PU transition than workers in the employee sample: the PU rate in the employees sample is 0.9%. Using this figure as a benchmark, workers in self-employment then make a SU transition at half the

⁸By definition, workers in the “paid employees” sample do not become self-employed.

| | Self-employed | Paid employees |
|--|---------------|----------------|
| Transition rates between employment states (monthly, origin to destination) | | |
| <i>UP</i> | 0.141 | 0.211 |
| <i>US</i> | 0.094 | — |
| <i>PU</i> | 0.024 | 0.009 |
| <i>SU</i> | 0.005 | — |
| <i>SP</i> | 0.008 | — |
| <i>PS</i> | 0.022 | — |
| Share by employment state | | |
| <i>U</i> | 0.046 | 0.044 |
| <i>P</i> | 0.269 | 0.956 |
| <i>S</i> | 0.685 | — |

Table 4: Employment state: flows and shares in each sample

rate of regular wage workers. Transitions to unemployment then represent an important driver of earnings risks in the self-employed sample.

There are also large direct flows (without an unemployment spell in between) between paid- and self-employment. In a typical month, for instance, there is a 2.2% chance that a paid-employed worker makes a direct transition to self-employment (*PS*).

Table 5 reports summary statistics on labor income and wealth for the self-employed sample. Within this sample, earnings also show more dispersion in self-employment than in paid-employment. The 10th percentile of labor income is almost \$200 lower in self-employment. It appears that workers start their own activity or maintain them with lower associated labor income than as an employee.

In addition, conditional on being continuously employed in the same labor form, earnings are substantially more volatile during a self-employment spell than during a paid-employment spell. Over a six-month period, for instance, there is a 25% chance that one's earnings drop by more than 20% in self-employment. The corresponding drop in earnings for workers in paid-employment is 4.5%.

These findings are robust to taking averages of income within an employment type

| | Percentile | | | | |
|---|------------|--------|--------|--------|---------|
| | p10 | p25 | p50 | p75 | p90 |
| Labor income (\$2009, monthly) | | | | | |
| Paid-employed (P) | 974 | 1,674 | 2,907 | 4,792 | 7,722 |
| Self-employed (S) | 790 | 1,605 | 3,144 | 5,889 | 10,081 |
| Labor income (\$2009, 6-month average) | | | | | |
| Paid-employed (P) | 1,137 | 1,814 | 3,073 | 4,939 | 7,846 |
| Self-employed (S) | 822 | 1,690 | 3,227 | 5,914 | 9,864 |
| Labor income growth within emp. type (6-month growth) | | | | | |
| Paid-employed (P) | -0.315 | -0.045 | -0.008 | 0.114 | 0.541 |
| Self-employed (S) | -0.561 | -0.219 | -0.008 | 0.320 | 1.341 |
| Labor income growth within emp. type (6-month growth, 6-month average) | | | | | |
| Paid-employed (P) | -0.223 | -0.050 | -0.006 | 0.107 | 0.381 |
| Self-employed (S) | -0.475 | -0.189 | -0.006 | 0.306 | 1.102 |
| Net liquid wealth (\$2009) | | | | | |
| Paid-employed (P) | -17,208 | -2,518 | 753 | 40,276 | 193,287 |
| Self-employed (S) | -15,459 | -764 | 5,818 | 83,850 | 287,921 |

Table 5: Earnings and wealth by labor force status in the self-employed sample.

over several periods. Table 5 reports the distribution of labor income and labor income changes using wages and business earnings averaged over six months. The larger dispersion of earnings in self-employment, both for income in levels and income growth, still holds in this case.

In terms of liquid wealth holdings, finally, a substantial fraction of these workers do not have access to substantial liquid savings to self-insure against unemployment risk and fluctuations in income. Workers in self-employment, who are not eligible to UI benefits, more than 25% have negative liquid wealth.

Summary. Three main points emerge from this descriptive analysis with regard to the self-employed’s exposure to labor market risks. First, while a fraction of the self-employed rank amongst the highest earners, they are also over-represented at the lower end of the earnings distribution. Second, there are some clear labor market risks associated with self-employment. When self-employed, workers become unemployed at a non-trivial rate, and they experience higher labor income volatility than in paid-employment. Third, many of the self-employed have limited wealth reserves that they can draw on in the face of these events. In the next sections, I present a quantitative model that can account for these empirical regularities.

3 A search model with self-employment

I build a search model in which risk-averse workers can save and borrow to jointly capture (i) workers’ transitions between paid-employment, self-employment, and unemployment; (ii) fluctuations in labor income; (iii) a self-insurance motive. This framework relates to several studies that describe a frictional labor market where workers can move across several forms of employment, such as formal and informal employment (Meghir et al., 2015) or public sector and private sector employment (Bradley et al., 2017). I depart from these prior studies by making agents risk-averse in an incomplete market framework.⁹ Agents can self-insure against labor market risks by borrowing or drawing down their savings, as well as by relying on additional sources of income at the household level, such as spousal earnings and welfare transfers.

3.1 Environment

Time is discrete. The labor force is represented by a continuum of working age individuals with measure one. These workers are the model counterpart to the “main earners” defined

⁹See also Lise (2013) for a search model with savings in which workers climb a single job ladder.

in Section 2. They are risk-averse and discount the future at rate $\beta < 1$. Their per-period utility of consumption $c > 0$ is given by the utility function u .¹⁰ Workers are allowed to borrow and save using a risk-free asset, a , with rate of return r and borrowing limit $\underline{a} \leq 0$ exogenously given.

Workers can be in one of four labor market states s : paid-employment (P), self-employment (S), unemployed on UI benefits (B), or unemployed not eligible to benefits (U). They can search for work opportunities either as paid- or self-employed when unemployed. While employed, they can only search for opportunities in the alternative employment form.¹¹ In the baseline model, workers in unemployment are only eligible to UI benefits if they were previously employees. Self-employed workers terminating their business are not eligible to such transfers.

Workers differ in skills, which translate in permanent differences in earnings potential. I index this heterogeneity by $k = 1, \dots, K$. A worker's skill level conditions the distributions of business and job opportunities from which they draw. These distributions are denoted by F_k^P and F_k^S , respectively for workers of type k drawing a wage or "self-employment income." There is no recall of past jobs or business opportunities. In addition, workers experience fluctuations in labor income while working at a job or at their business. The income processes governing these shocks, conditional on current labor income, also depend on k and are given by $Q_k^P(\cdot|w)$ and $Q_k^S(\cdot|y)$.

A worker's skill level also conditions the destruction rate of employment opportunities. Jobs and businesses disappear, respectively, with exogenous probability δ_k^P and δ_k^S . These destruction rates are indexed on ability k to allow for differential exposure to unemployment risk by worker type. In addition, workers are always free to leave their current job or business, in which case they become unemployed without access to unemployment benefits (they go straight to labor market state $s = U$).

¹⁰ $u : \mathbb{R}_+^* \rightarrow \mathbb{R}$ is assumed to satisfy $u' > 0$, $u'' < 0$, and $\lim_{c \rightarrow 0} u'(c) = \infty$.

¹¹ SS and PP transitions are captured by earnings shocks.

Unemployment insurance (UI) has two dimensions in the model: $b(\cdot)$ and T . $b(\cdot)$ is a benefit function, mapping workers' wage in their last job to some benefit level. $T \geq 1$ controls the potential duration of benefits payments, with the eligibility to benefits expiring at rate $1/T$.¹² These assumptions restrict UI to belong to the “constant benefit, finite duration” class of UI policies, in line with the US system.

There are additional income sources that accrue to the household beyond the main earner's labor income and the (potentially negative) returns to wealth ra . There may be additional earners in the household. On top of unemployment insurance, there may also be other welfare programs to which the household is eligible. Total household income is modeled as a tuple of functions, $\{Y_k^s, s \in \{U, P, S\}\}$, of the main earner's labor income in the corresponding labor market state s and worker type k . These functions give a reduced-form representation of the phasing-out of welfare programs as earnings increase and of added worker effects.

3.2 Timing

Each period t unfolds as follows:

1. Earning realization. Workers in paid- and self-employment get a new labor income draw, respectively from the conditional distributions $Q_k^P(\cdot|w)$ and $Q_k^S(\cdot|y)$.
2. Quits and separations. Conditional on earnings in the present period, workers can decide to quit, in which case they become unemployed with no access to UI benefits (U). If not, they are hit by an exogenous destruction shock with probability δ_k^s , $s \in \{P, S\}$.
3. Search. Workers not separated, as well as those previously in unemployment sample job offers and self-employment opportunities, respectively, with probabilities

¹²This simplification is introduced to avoid making the eligibility period to unemployment benefits a state variable in state B . Because agents are risk-averse in the model, this simplification underestimates the value of unemployment benefits relative to a model with a deterministic eligibility period.

λ^{sP} , $s \in \{U, S\}$ and λ^{sS} , $s \in \{U, P\}$. These probabilities are assumed to be mutually exclusive, so that unemployed workers, who can search for work opportunities both as paid- or self-employed, get at most one labor income draw in each period, either from F_k^P or F_k^S . If they choose to pursue this job or business opportunity, they immediately switch to this new labor form and earn the associated wage or business income.

4. Consumption and savings. Household income accrues to all workers. UI benefits are paid out to the eligible fraction of unemployed workers. Agents then choose consumption c and next period's net wealth a' .

3.3 Worker's problem

Notations. Let $R_k^s(a, y)$ be the present value of being in state s with net wealth holdings a and labor income y for a worker of type k at the start of the quits and separations stage. Let $V_k^s(a, y)$ stand for the worker's present value at the start of the consumption and savings stage. I denote the net gains from getting a job or business opportunity (a draw from $F_k^{s'}$) as

$$\phi_k^{ss'}(a, y) := \int \max \left\{ V_k^{s'}(a, \tilde{y}) - V_k^s(a, y), 0 \right\} dF_k^{s'}(\tilde{y}),$$

where s is the worker's current state, and $s' \in \{P, S\}$, which by assumption is only defined for $s \neq s'$.

The notation $\phi_k^{ss'}(a, y)$ implicitly defines a reservation income given a worker's state variables. For instance, a worker with current assets a and currently in unemployment (labor force status $s = U$) who gets a business idea accepts any idea generating income at least greater than $\underline{y}_k^{US}(a)$: $V_k^S(a, \underline{y}_k^{US}(a)) = V_k^U(a)$. The net gains from getting such a

business opportunity can then be written

$$\phi_k^{US}(a) = \int_{\underline{y}_k^{US}(a)}^{\bar{y}} V_k^S(a, \tilde{y}) - V_k^U(a) dF_k^S(\tilde{y}).$$

Value functions: paid-employed. The value of holding a job with current wage w and wealth a at the beginning of the period is given by

$$R_k^P(a, w) = \max \left\{ V_k^U(a), \delta_k^P V_k^B(a, w) + (1 - \delta_k^P) \left[V_k^P(a, w) + \lambda^{PS} \phi_k^{PS}(a, w) \right] \right\}. \quad (1)$$

The max operator corresponds to the worker's choice to remain in paid-employment given the current realization of wages. The worker's outside option is to become unemployed and start searching next period. Because only layoffs (δ_k^P -shocks) entitle workers to unemployment benefits, voluntary quits result in unemployment without benefits ($V_k^U(a)$).

The value of being in paid-employment with wage w and wealth a at the consumption and savings stage writes

$$\begin{aligned} V_k^P(a, w) &= \max_{c, \tilde{a}} \left\{ u(c) + \beta \int R_k^P(\tilde{a}, \tilde{w}) dQ_k^P(\tilde{w}|w) \right\} \\ \text{s.t. } c + \frac{\tilde{a}}{1+r} &= Y_k^P(w) + a; \\ \tilde{a} &\geq \underline{a}. \end{aligned} \quad (2)$$

Value functions: self-employed. The value of a business at the beginning of the period mirrors that of job holders in equation (1), with the difference that unemployment transitions are all to the no-benefits state (U):

$$R_k^S(a, y) = \max \left\{ V_k^U(a), \delta_k^S V_k^U(a) + (1 - \delta_k^S) \left[V_k^S(a, y) + \lambda^{SP} \phi_k^{SP}(a, y) \right] \right\}. \quad (3)$$

Similarly to the case of paid-employment (P), workers can decide to terminate their business given the current realization of income and become unemployed (first term in the max operator). Otherwise, they continue with their current business and are subject to exogenous destruction shocks with probability $\delta_k^S < 1$.

The self-employed's savings problem is given by

$$\begin{aligned} V_k^S(a, y) &= \max_{c, \tilde{a}} \left\{ u(\kappa \cdot c) + \beta \int R_k^S(\tilde{a}, \tilde{y}) dQ_k^S(\tilde{y}|y) \right\} \\ \text{s.t. } c + \frac{\tilde{a}}{1+r} &= Y_k^S(y) + a; \\ \tilde{a} &\geq \underline{a}. \end{aligned} \tag{4}$$

The key difference with the savings problem of the paid-employed is that self-employment comes with non-pecuniary benefits, a factor $\kappa \geq 1$, scaling consumption directly in the utility function. These non-pecuniary benefits capture a preference for being self-employed, which could stem from the intrinsic value of being one's own boss, from unreported earnings, or from other tax advantages.¹³ Empirically, this feature is motivated in the data by the non-trivial growth in earnings observed following a direct SP transition, which is interpreted as the compensating differential of self-employment. Non-pecuniary benefits can potentially lead workers to turn down a job offering a better pay than what they get from their current business.¹⁴

Value functions: unemployed. The probabilities of getting a job opportunity (λ_{UP}) or business opportunity (λ_{US}) are assumed to be mutually exclusive. The values of searching for employment opportunities are then given by

$$R_k^B(a, w) = V_k^B(a, w) + \lambda^{UP} \phi_k^{BP}(a, w) + \lambda^{US} \phi_k^{BS}(a, w), \tag{5}$$

¹³Private consumption through one's business is one example of such a tax advantage.

¹⁴Most structural models of self-employment allow for some degree of preference for business ownership (Humphries, 2017; Catherine, 2019; Jones and Pratap, 2020).

and

$$R_k^U(a) = V_k^U(a) + \lambda^{UP} \phi_k^{UP}(a) + \lambda^{US} \phi_k^{US}(a), \quad (6)$$

respectively for workers eligible (state $s = B$) and not eligible (state $s = U$) to unemployment benefits. In equation (5), w denotes the individual's last wage before becoming unemployed, which is part of the state-space as it determines their UI benefit income.

Equations (5) and (6) imply an indirect comparison between work opportunities in paid- and self-employment. Consider a currently unemployed worker with a potential job w in hand (a draw from F_k^P which occurs with chance λ^{UP}) decides whether to wait for better options either working for others (P) or for themselves (S). This decision is captured by the functions $\phi_k^{BP}(a, w)$ (or $\phi_k^{UP}(a)$ if the worker is no longer eligible to UI benefits), which denote the net gains from a draw from F_k^P .

The consumption-savings problem for unemployed workers on benefits is given by

$$\begin{aligned} V_k^B(a, w) &= \max_{c, \tilde{a}} \left\{ u(c) + \beta [\delta^B R_k^U(\tilde{a}) + (1 - \delta^B) R_k^B(\tilde{a}, w)] \right\} \\ \text{s.t. } c + \frac{\tilde{a}}{1+r} &= Y_k^U + a + b(w); \\ \tilde{a} &\geq \underline{a}, \end{aligned} \quad (7)$$

where the term in squared brackets after the discount factor comes from benefits expiring at rate $\delta^B := 1/T$. Similarly to equation (5), w is part of the state-space because benefits are indexed on the last wage received by the worker.

The consumption-savings problem for unemployed workers not eligible to benefits

(state $s = U$) follows directly from adapting equation (7) and is given by

$$\begin{aligned}
V_k^U(a) &= \max_{c, \tilde{a}} \left\{ u(c) + \beta R_k^U(\tilde{a}) \right\} \\
\text{s.t. } \quad c + \frac{\tilde{a}}{1+r} &= Y_k^U + a; \\
\tilde{a} &\geq \underline{a}.
\end{aligned} \tag{8}$$

3.4 Stationary equilibrium

Taken together, the optimal choices of consumption and savings, as well as the reservation strategies implied by workers' decisions to quit and to follow through on alternative employment opportunities imply a stationary distribution over employment states $\{U, B, S, P\}$, labor income (y, w) , and net wealth (a) for each skill level k . These distributions are denoted Γ_k in what follows. A formal statement of this definition is given in Appendix B.

3.5 Discussion

I conclude the exposition of the model by briefly discussing some of my modeling choices. First, I do not model investment and the associated accumulation of business capital. Hurst and Lusardi (2004) argue that liquidity constraints are unlikely to be binding for a majority of entrepreneurs in the US, which would imply that rental markets for capital function well. This is also consistent with the fact that many businesses operate with a low capital stock. Median business assets in the SIPP are less than \$6,500, suggesting that many self-employed have limited capital requirements. For the subset of workers in self-employment operating on a larger scale, an interesting extension would consider how these workers allocate their portfolio between liquid assets, their business, and other illiquid assets, given the labor market risks associated with this employment form.

In addition, I abstract from general equilibrium effects. The self-employed account

for less than 15 percent of workers in the sample. Their borrowing and saving behavior can therefore be expected to have a very limited impact on the aggregate interest rate. Besides, while my definition of self-employment covers the majority of firm owners, most of the businesses associated with self-employment have very few, if any, employees. The data suggest that only 5 percent of these businesses have ever had more than 25 employees. Large firms account for the majority of US employment (in data from the Census, 57% of US employment is at firms with more than 100 workers), and the managers at these larger firms are unlikely to be affected by the introduction of additional benefits targeted at the self-employed. It therefore seems reasonable that the labor market parameters (arrival rates of opportunities and distributions of income in paid- and self-employment) remain constant given the type of policies considered.

Lastly, I model worker skills k as uni-dimensional. As such, I implicitly assume that these skills are transferable across employment forms. By focusing on data from the self-employed sample in the calibration, I ensure that workers are skilled enough in self-employment that it becomes their main source of labor income at some point. Because many self-employed in the data are not observed working as employees, recovering a joint distribution of skills in both labor forms is not feasible. The life cycle income profiles reported in Humphries (2017) for various groups of self-employed suggest that experience as a wage worker is, to a large extent, transferable to self-employment, conditional on making the transition.

4 Calibration

My calibration strategy proceeds in two steps. Worker heterogeneity is first introduced by clustering workers with similar labor market outcomes. In a second step, I target key moments from the data to calibrate the remaining parameters, conditional on the heterogeneity uncovered in the first step. Data from the self-employed sample (workers

with at least some self-employment over the duration of the survey) are used throughout.

4.1 Worker heterogeneity

To discipline permanent worker heterogeneity in the model (the index k), I rely on a clustering tool from machine learning to partition workers into distinct groups.¹⁵ This classification aims at capturing the sizable heterogeneity in earnings found in the self-employed sample (see Section 2).

Clustering algorithm. Following Bonhomme et al. (2019), I use a k-means algorithm to group workers in earnings classes in the SIPP data. This procedure finds the best partition of the data according to the following objective function

$$\arg \min_{\tilde{h}, k_1, \dots, k_N} \sum_{i=1}^N \left\| \hat{h}_i - \tilde{h}(k_i) \right\|^2, \quad (9)$$

where i in $\{1, \dots, N\}$ denotes individuals, $k_i \in \{1, \dots, K\}$ indicates the group to which individual i is assigned, with $1 < K \leq N$, and h is a vector of characteristics, with h_i and $h(k_i)$ denoting the vectors for individual i and group k_i respectively. Each element in $\tilde{h}(k_i)$ is computed by averaging over the members of the group. The solution to the minimization problem in (9) then assigns a cluster to each i such that the squared Euclidean distance between i 's vector of characteristics and the average of these characteristics in i 's group is minimized.¹⁶

The logic behind this clustering step is in essence the same as that described in Bonhomme et al. (2019). They first cluster firms based on their empirical distribution of earnings amongst employed workers, before using the resulting classes in a series of mix-

¹⁵Thereafter, I use the terms “class”, “group”, and “cluster” interchangeably to designate the outcome of the clustering procedure.

¹⁶Standard algorithms to efficiently solve this global minimization problem are readily available in standard packages. I use the implementation of the “Hartigan-Wong” algorithm in the base R function “kmeans”.

ture models. The estimated partition then captures both observed and unobserved firm heterogeneity.

In my framework, I rely on a similar partition of the data to discipline worker heterogeneity in earnings. In the model, this partition then translates into cluster-specific distributions of employment opportunities $\{F_k^P, F_k^S\}$ and unemployment shocks $\{\delta_k^P, \delta_k^S\}$.

Implementation. To implement this procedure in practice, one needs to choose the number of clusters, K , and the vector of features \hat{h}_i . I set $K = 4$, which represents a good trade-off between capturing a reasonable degree of heterogeneity in worker ability and computation time.¹⁷

The vector of features \hat{h}_i includes two key outcomes: a measure of labor income when employed and a measure of exposure to unemployment. The former is the empirical cdf of earnings

$$\widehat{\text{ECDF}}_i(y) := \frac{1}{\sum_t \mathbb{1}\{\text{Employment}_{it} = 1\}} \sum_t \mathbb{1}\{y_{it} \leq y\},$$

with y_{it} denoting labor income for i in period t and Employment_{it} the corresponding indicator for employment. This distribution is based on all earnings, irrespective of whether they come from paid- or self-employment at a point in time.¹⁸ It is computed for each decile of the empirical distribution of labor income in the sample.

Exposure to unemployment is the fraction of time workers spend in non-employment over the duration of the survey. Letting T_i be the number of months the individual is observed in the panel (typically slightly under three years in the 1996-2008 SIPP panels),

¹⁷Figure C.3 shows how the total within sum of square changes with K .

¹⁸I do not further distinguish between labor income from paid- versus self-employment since many workers are never observed in paid-employment in my sample.

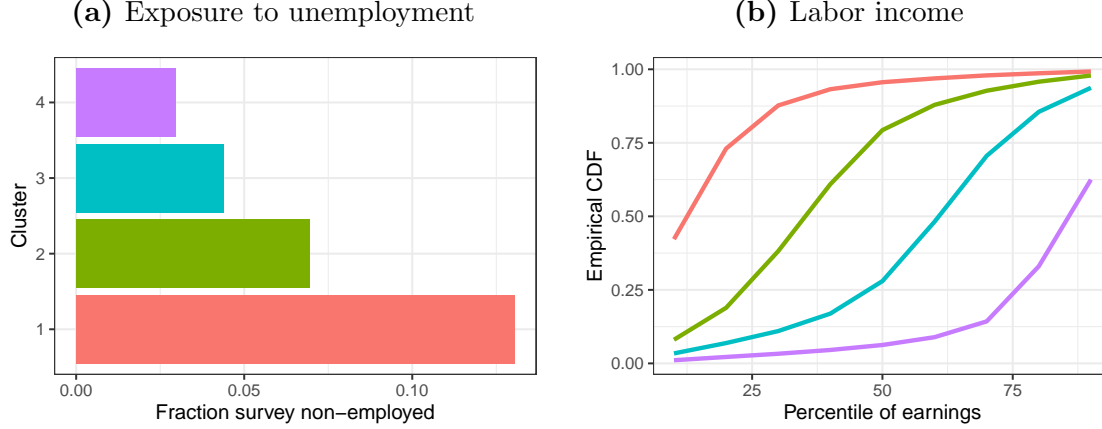


Figure 2: Worker clusters obtained from clustering algorithm. The Figure depicts the average feature in each cluster with $K = 4$. The clusters are ordered based on the value of the empirical cdf of labor income at the median.

it is given by

$$\frac{1}{T_i} \sum_t \mathbb{1}\{\text{Employment}_{it} = 0\}.$$

Results. The resulting average features for each worker class, $\tilde{h}(k_i)$, are depicted in Figure 2. The Figure shows a clear relationship between the two measures used to build the clusters: the group with the largest labor income is also least likely to be unemployed. This negative relationship between earnings when employed and unemployment risk holds for all worker groups.

I subsequently label the four clusters of workers based on the median earnings in each group. As shown in Table 6, the “Low”, “Medium-Low”, “Medium-High”, and “High” groups correspond to clusters with a median monthly labor income of, respectively, \$1,049, \$2,230, \$3,830, and \$7,494.

Table 6 further shows how worker classes differ along several dimensions not directly used in the clustering procedure. In terms of demographic composition, workers with lower earnings are more likely to be women and belong to a minority, and are less likely to be married and have a college education. (Recall that the sample is restricted to the

| | Worker earnings group | | | |
|---------------------------------|-----------------------|------------|-------------|--------|
| | Low | Medium-Low | Medium-High | High |
| Median earnings (\$2009) | 1,049 | 2,230 | 3,830 | 7,494 |
| Demographic characteristics | | | | |
| Age | 42.44 | 41.45 | 42.63 | 44.20 |
| Gender (woman=1) | 0.47 | 0.28 | 0.19 | 0.15 |
| Married (married=1) | 0.42 | 0.57 | 0.68 | 0.76 |
| Race (non-white=1) | 0.19 | 0.15 | 0.12 | 0.10 |
| College Graduate | 0.45 | 0.47 | 0.47 | 0.51 |
| Post-graduate | 0.07 | 0.09 | 0.13 | 0.24 |
| Business characteristics | | | | |
| Median business equity (\$2009) | 476 | 2,458 | 8,819 | 22,658 |
| Incorporated | 0.17 | 0.27 | 0.36 | 0.46 |
| More than 25 employees | 0.02 | 0.03 | 0.05 | 0.10 |
| Number of workers (N_k) | 3,515 | 5,366 | 5,373 | 3,825 |

Table 6: Characteristics of worker clusters.

main earner in each household.) Reassuringly, age is roughly the same in all groups on average, so the classification is not driven by workers' accumulation of human capital. The resulting clusters of workers also clearly relate to the large variation in the scale of the corresponding businesses in terms of equity, employment, and legal status. For example, the group of highest earners are close to three times as likely to be incorporated (46%) as the group of lowest earners (17%).¹⁹ The worker classes resulting from the algorithm therefore capture meaningful observed characteristics, as well as potentially unobserved ability traits determining earnings.

4.2 Model parametrization

UI benefits. The actual unemployment insurance system in the US is determined at the state level. I follow Chetty (2008) and Saporta-Eksten (2014) and approximate these

¹⁹I also report the distribution of worker groups by industry in Figure C.4.

policies as a fifty percent replacement rate for a maximum duration of six months. I also cap the maximum monthly benefit payments at \$2,000.²⁰ In the notation of the model, the UI system is then given by the replacement rate function $b(w) = \max\{2,000; 0.5 \cdot w\}$ and $T = 6$. The relevant model wage w is assumed to be the last wage realization before the worker becomes unemployed.

Household income functions. I obtain a mapping between an individual’s current labor income and household income by estimating the following regressions

$$\ln Y_{it}^s = \alpha_k^s + \beta_k^s \cdot \ln y_{it}^s + \epsilon_{it}, \quad k = 1, \dots, K, \quad s \in \{P, S, U\}, \quad (10)$$

where Y_{it}^s is total household income, including other earners and welfare transfers, but net of UI payments, $\ln y_{it}^s$ is the main earner’s labor income, and ϵ_{it} is an error term. The index k refers to the worker’s assigned cluster, and s denotes their current state in the labor market. (The slope β_k^s is omitted for $s = U$, since labor income is zero in this case by definition.) The estimated coefficients $(\hat{\alpha}_k^s, \hat{\beta}_k^s)$ give the corresponding household income function by cluster and labor force status. These regressions are reported in Table C.1.

This functional form assumption allows to capture some degree of household heterogeneity in total household income by worker cluster without introducing additional state variables in the model. For example, by running the regressions in (10) separately by k , this specification allows for assortative mating, as spousal income is allowed to depend on the skill level of the household’s main earner. This heterogeneity in income at the household level represents another difference in the scope for insurance.

²⁰This further approximates the caps on UI benefit payments enacted by most states. See the “Significant Provisions Of State Unemployment Insurance Laws” tables from the Department of Labor for the corresponding weekly ceilings by state.

Utility and labor income distributions. Utility from consumption is given by a standard CRRA specification $u(c) = (1 - \gamma)^{-1} c^{1-\gamma}$. The distribution of labor income draws, F_k^s , is assumed to follow a log-normal distribution with mean μ_k^s and standard deviation σ_k^s for each $k \in \{1, \dots, K\}$ and $s \in \{P, S\}$. Shocks to labor income are modeled as a chance $1 - \rho^s$ to redraw from the corresponding F_k^s distribution in each period.²¹

4.3 Calibration procedure

I set the returns on net liquid wealth r to zero and the coefficient of risk-aversion γ to $3/2$. The lower bound on net liquid wealth is allowed to vary by worker group according to $\underline{a}_k = \underline{a} \cdot Y_k^{\text{p50}}$, where Y_k^{p50} denotes median household income in a cluster and the factor \underline{a} is taken from Kaplan and Violante (2014).²² The borrowing constraint is therefore tighter for workers with lower earnings potential.

The remaining parameters are calibrated by matching a set of moments from the data to their counterpart in the simulated model. Starting from the invariant distribution of workers in each group, Γ_k , I simulate a panel of N_k workers (the number of observations in each cluster k in the data) for 48 months. I stack all cluster-specific panels and obtain the model simulated moments by using the exact same definitions as those used to construct the moments derived from the data.

Though all parameters are jointly calibrated, I motivate my choice of moment targets by mapping each set of calibrated parameters to specific moments. The job and business arrival rates $(\lambda^{US}, \lambda^{UP}, \lambda^{PS}, \lambda^{SP})$ are chosen to replicate the corresponding monthly transition rates. The job and business destruction rates (δ_k^P, δ_k^S) directly relate to the k -specific transition rates into unemployment, respectively from paid- and self-employment. The parameters governing the distributions of labor income draws (μ_k^s, σ_k^s) are set to

²¹This income process better fits the distribution of labor income growth within labor form, which exhibits a clear peak at zero, than a standard AR(1) process.

²²These authors set this factor to the median ratio of households' "total credit limit" to their quarterly income, which is equal to 74% in the Survey of Consumer Finances. I then set $\underline{a} = -.74 \times 3$.

match the observed distributions of income by worker group k and employment state s . The rates of persistence of income (ρ^P, ρ^S) are disciplined by the six-month growth of income for workers continuously employed in paid- and self-employment, respectively. The discount factor (β) is set to match the median net liquid wealth in the data. Finally, the non-pecuniary benefits of self-employment (κ) is controlled by the gap in median labor income growth following a direct PS and SP transition, where labor income growth is computed only for transitions with at least six months of continuous employment history on either side of the transition.

The classification of workers is implemented by relating class-specific parameters to class-specific targets. In line with my choice of clustering variables (unemployment risk and labor income) I therefore make the job and business destruction parameters and labor income parameters class-specific.²³ I stress that, while there is a tight link between these parameters and the corresponding moments, they still need to be calibrated jointly due to the selection across P , S , and U , that naturally arises in the model. As an example, the model simulated distribution of labor income in self-employment is shaped by workers' reservation income strategy with respect to paid-employment and unemployment given their current wealth level.

4.4 Results

I report the model fit to the targeted moments in tables 7 and 8. Table 7 first shows the fit to the moments that are specific to each worker class.²⁴ Both for PU and SU transitions, there is a clear decreasing pattern where higher earners, defined by median earnings in the class, are much less likely to transition to unemployment. There is an almost one percent chance for the lowest earners to make a SU transition in a given month, while the

²³The UP and US transition rates, which also relate to unemployment, do not exhibit a clear pattern by class in the data, so the corresponding arrival rates are common to all workers.

²⁴The corresponding class-specific parameters are given in Table C.2.

highest earners face a .2 percent chance of a similar event. The class-specific destruction parameters allow the model to very closely replicate this heterogeneity in the risk to transition into unemployment.

The labor income distributions also show that, even within a worker class, labor income tends to be markedly more dispersed in self-employment than in paid-employment. While median earnings are very similar for both types of employment in most worker classes, self-employment income tends to have a longer left tails for the lowest earners. Overall, the model replicates these distribution well, though for some worker classes it does not fully get the tails right given the log-normal assumption.

Table 8 shows the model fit to the moment computed on the whole sample, and Table 9 lists the parameters common to all workers. The model replicates well the large flows across unemployment, paid-, and self-employment observed in the data. In terms of the arrival rates of job and business opportunities for employed workers (λ^{SP} and λ^{PS}), the model sees transitions to paid-employment as relatively straightforward for self-employed workers (a six percent chance in every month), though many of these opportunities are turned down (the actual transition rate is slightly under one percent). The model also emulates the difference in earnings volatility observed within employment type found in the data, with workers in self-employment experiencing markedly larger changes in earnings over a six-month period of continuous employment.

The median growth rate of earnings following an employment-to-employment transition differs starkly between *SP* and *PS* transitions. The calibrated non-pecuniary benefits of self-employment, κ , implies that consumption is 2.5% more valuable in self-employment for these workers. The model can replicate a large fraction of the gap in median labor income growth between the two types of transitions found in the data.

Regarding net liquid wealth, finally, the model implies median net liquid wealth holdings in line with the data, though, in spite of having access to UI benefits in the event they become unemployed, workers in paid-employment still tend to accumulate slightly more

| Moment | Worker earnings group | | | | | | | |
|-------------------------|-----------------------|--------|------------|--------|-------------|--------|-------|--------|
| | Low | | Medium-Low | | Medium-High | | High | |
| | Model | Actual | Model | Actual | Model | Actual | Model | Actual |
| Transition rates to U | | | | | | | | |
| PU | 0.041 | 0.044 | 0.029 | 0.028 | 0.018 | 0.018 | 0.016 | 0.015 |
| SU | 0.012 | 0.009 | 0.006 | 0.006 | 0.004 | 0.004 | 0.002 | 0.002 |
| Earnings in P (log) | | | | | | | | |
| p10 | 6.743 | 6.170 | 7.161 | 7.078 | 7.813 | 7.718 | 8.441 | 8.305 |
| p25 | 6.882 | 6.609 | 7.452 | 7.410 | 8.010 | 7.988 | 8.676 | 8.644 |
| p50 | 7.022 | 7.041 | 7.744 | 7.721 | 8.245 | 8.262 | 8.958 | 8.941 |
| p75 | 7.189 | 7.366 | 7.987 | 8.035 | 8.480 | 8.541 | 9.287 | 9.231 |
| p90 | 7.328 | 7.663 | 8.278 | 8.277 | 8.716 | 8.807 | 9.522 | 9.523 |
| Earnings in S (log) | | | | | | | | |
| p10 | 5.993 | 5.924 | 6.967 | 6.769 | 7.545 | 7.353 | 8.340 | 8.136 |
| p25 | 6.429 | 6.352 | 7.312 | 7.253 | 7.883 | 7.845 | 8.638 | 8.534 |
| p50 | 6.952 | 6.894 | 7.656 | 7.695 | 8.221 | 8.246 | 8.936 | 8.939 |
| p75 | 7.388 | 7.316 | 8.069 | 8.048 | 8.627 | 8.614 | 9.294 | 9.333 |
| p90 | 7.824 | 7.803 | 8.414 | 8.461 | 8.965 | 8.981 | 9.592 | 9.592 |

Table 7: Model fit—moments specific to worker class.

| Moment | Model | Actual |
|--|--------|--------|
| Transition rates | | |
| <i>UP</i> | 0.164 | 0.141 |
| <i>US</i> | 0.106 | 0.094 |
| <i>SP</i> | 0.007 | 0.008 |
| <i>PS</i> | 0.022 | 0.022 |
| Earnings growth in paid-employment (6 months) | | |
| p10 | -0.385 | -0.315 |
| p25 | -0.136 | -0.045 |
| p50 | 0.000 | -0.008 |
| p75 | 0.087 | 0.114 |
| p90 | 0.518 | 0.541 |
| Earnings growth in self-employment (6 months) | | |
| p10 | -0.582 | -0.561 |
| p25 | -0.287 | -0.219 |
| p50 | 0.000 | -0.008 |
| p75 | 0.317 | 0.320 |
| p90 | 1.250 | 1.341 |
| Median earnings growth after direct transition | | |
| After <i>PS</i> | 0.058 | -0.028 |
| After <i>SP</i> | 0.243 | 0.218 |
| Fraction below actual median wealth (\$2009) | | |
| <i>P</i> (p50 in data: \$ 753) | 0.334 | 0.500 |
| <i>S</i> (p50 in data: \$5818) | 0.461 | 0.500 |

Table 8: Model fit—moments computed on whole sample.

| Parameter | Description | Value |
|-----------------|--|-------|
| Set exogenously | | |
| r | Rate of return on liquid wealth | 0.0 |
| γ | Coefficient of rel. risk aversion | 1.5 |
| Calibrated | | |
| β | Discount factor (monthly) | 0.992 |
| κ | $u(\kappa \cdot c)$ in self-employment | 1.025 |
| λ^{UP} | Arrival rate P opportunity from U | 0.168 |
| λ^{US} | Arrival rate S opportunity from U | 0.111 |
| λ^{SP} | Arrival rate P opportunity from S | 0.067 |
| λ^{PS} | Arrival rate S opportunity from P | 0.030 |
| ρ^P | Persistence income in P | 0.818 |
| ρ^S | Persistence income in S | 0.775 |

Table 9: Parameters—Common across worker classes.

wealth than in the data. I also stress that despite having access to some insurance through additional welfare programs and the income of other household members in unemployment (captured by Y_k^U) and despite setting the coefficient of relative risk-aversion at a relatively low value ($\gamma = 3/2$), the yearly discount rate required to match the fraction of workers at low wealth levels is high (approximately 0.10). The gap between the discount rate and the interest rate is implicitly the cost of using savings as insurance, and the gap required to match median net liquid wealth holdings in the data is large. Lentz (2009) notes that workers tend to value unemployment benefits more when this gap is larger in his structural model estimated on Danish data. Through the lens of the model, additional social insurance is therefore potentially quite valuable for some groups of self-employed.

5 Policy analysis

What are the labor market and welfare effects of providing benefits targeted at the labor market risks associated with self-employment? This section assesses the impact of several counterfactual policies through the lens of the calibrated model. I emphasize how

heterogeneous groups of self-employed workers respond to these policies.

5.1 Alternative policies

UI for the self-employed. A first potential policy is to extend UI benefits to the self-employed. I rely on a similar formulation as the one used in the baseline model. The UI policy for the self-employed is given by $b(y) = \max\{0.5 \cdot y; 2,000\}$: benefits are paid at a fifty percent replacement rate of workers' last income capped at \$2,000. Benefits expire at rate $1/T$ with $T = 6$. These benefits are financed by taxing the labor income of workers in self-employment according to the tax schedule τ^S . Introducing some notation, τ^S must satisfy

$$\sum_k \omega_k \left[\int \tau^S(y) d\Gamma_k^S(a, y) - \int b(y) d\Gamma_k^C(a, y) \right] = 0, \quad (11)$$

where state $s = C$ denotes previously self-employed workers currently on benefits (the counterpart to state $s = B$), Γ_k^s denotes the measure of workers in labor force state s and worker class k , and $\omega_k := N_k/N$ is the share of workers in skill group k . Equation (11) imposes that the budget of the policy is balanced overall, but it does not imply that it is actuarially fair within each group k .

When writing down the self-employed's problem, I assume that, contrary to the paid-employed, the self-employed can choose to terminate their activity and still be eligible to unemployment benefits. The distinction between layoffs and voluntary quits, which conditions eligibility in many UI systems, is irrelevant for the self-employed. Formally, the value of self-employment at the beginning of the search stage (equation (3) in the baseline model) now reads

$$R_k^S(a, y) = \max \left\{ V_k^C(a, y), \delta_k^S V_k^C(a, y) + (1 - \delta_k^S) \left[V_k^S(a, y) + \lambda^{SP} \phi_k^{SP}(a, y) \right] \right\},$$

where $V_k^C(a, y)$ is the value of becoming unemployed when benefits are paid out. There are then two ways for the self-employed to cash UI benefits: either following a δ_k^S -shock (involuntary) or by “choosing” to become unemployed (voluntary).

This formulation abstracts from some of the challenges associated with designing a UI program specifically for the self-employed. Because many self-employed run fairly informal businesses (the incorporation rate is 32% in the self-employed sample), such a program would require many self-employed to make their business more formal in legal terms, which would form the basis to both collect UI contributions (a role taken up by the employer for traditional wage workers) and record business closures (the condition to claim benefits).²⁵ Keeping track more formally of workers’ employment history in self-employment would further allow to introduce specific experience ratings and search requirements for this group.

Basic income policy. Given the potential challenges in administering a UI program for the self-employed, I also consider a basic income supplement policy giving workers a “top up” on existing transfers, irrespective of their labor force status. By decoupling transfers and labor force status, such a policy design sidesteps the practical difficulties in keeping track of workers’ self-employment histories, but is also more directly targeted at the low earners.

This policy is implemented through the following tax function $\mathcal{T}(Y) := -\tilde{b} + \tilde{\tau}Y$, where \tilde{b} is some unconditional cash transfer and $\tilde{\tau}$ is a constant tax rate of household income Y . The parameters $(\tilde{b}, \tilde{\tau})$ are assumed to be such that the government’s budget satisfies

$$\sum_{s,k} \omega_k \int \mathcal{T}(Y_k^s(y)) d\Gamma_k^s(a, y) = 0. \quad (12)$$

²⁵Requiring businesses operating on a limited scale to file accounting information seems possible given transactions are increasingly digitalized.

In implementing this policy scenario, it is assumed that the UI program for the paid-employed is discontinued and the corresponding resources added to the overall budget.

Policy scenarios. I consider two alternative UI scenarios and one basic income scenario. The first UI scenario (UI-A) simply is identical to the way UI contributions are collected in the actual US system. Specifically, the tax schedule to finance UI is set to $\tau^S(y) = \min\{\tau \cdot y; 7,000/12\}$. In practice, all States cap the annual earnings subject to UI contributions.²⁶ Because UI contributions are infra-marginal for all but the lowest earners in such a system, I also consider a second, more progressive, schedule (UI-P) where UI contributions are paid out on all self-employment income at a constant rate $\tau^S(y) = \tau \cdot y$, without a cap. The benefit entitlements are kept identical to the first UI scenario. I finally study a basic income policy (BI), an unconditional cash transfer $b = 100$ financed by levying taxes at a constant rate $\tilde{\tau}$ of household income. The value of b is chosen to make the total size of the program (total government income) broadly consistent with that implied by the two UI policies.

These policies are implemented by finding the tax rates satisfying the budget balance conditions in equations (11) and (12). In addition, I allow these tax rates to have an effect along the intensive margin of labor supply. I simply input an elasticity of labor income to the marginal tax rate of 0.5.²⁷ The policies and corresponding tax rates are summarized in Table 10.

5.2 Results

Labor market. Table 11 shows how employment choices are distorted in response to the introduction of each policy. The Table gives the percentage change in each transition rate

²⁶\$7,000 corresponds to the States with the lowest threshold. Detailed tables on UI contribution rates by State can be found at https://oui.doleta.gov/unemploy/avg_employ.asp.

²⁷This value is an upper bound on the steady-state intensive margin labor supply elasticity in Chetty et al. (2011), who report 0.33 as their central estimate.

| Policy | Benefits | Contributions | Tax rate |
|-----------------------------|--------------------------------|-------------------------------------|----------|
| Unemployment insurance (UI) | | | |
| UI-Align | $\max\{0.5 \cdot y^S; 2,000\}$ | $\tau \cdot \min\{y^S; 7,000/12\}$ | 0.025 |
| UI-Progressive | $\max\{0.5 \cdot y^S; 2,000\}$ | $\tau \cdot y^S$ | 0.003 |
| Basic income (BI) | $\tilde{b} = 100$ | $\tilde{\tau} \cdot Y^s, \forall s$ | 0.014 |

Table 10: Summary of policy scenarios.

with respect to its baseline value. Though the arrival rates of employment opportunities are exogenous, the reservation income at which workers decide to stop searching responds to policy changes. The endogenous choice to become unemployed after bad realizations of income (as described in (1) and (3)) is similarly affected in each policy scenario.

I contrast how labor supply decisions along the extensive margin differ for workers at the top (“High”, median monthly earnings of \$7,494) and bottom (“Low” median monthly earnings of \$1,049) of the skill distribution. UI benefits for the self-employed has a clear moral hazard component for the “Low” group. Their SU rate increases by close to twenty percent; their US rate drops by about two percent as well. For this group, capping UI contributions to finance UI payments implies that they shoulder more of the tax burden. As such, while self-employment becomes more attractive for workers currently in employment (decrease in SP , increase in PS), these changes are small. This is in contrast to the more progressive policy (UI-P), which makes self-employment a lot more attractive in the “Low” group, since they get the same insurance level with a lower tax burden. By comparison, workers in the “High” group do not respond strongly along the worker flow margin, since they are further from the borrowing limit. Their choice of employment state is therefore not impacted by the availability of UI benefits in self-employment.

The basic income policy affects labor supply decisions in a very different way. Workers in both the “Low” and “High” earnings groups respond similarly, though the precise magnitude differs. The combination of removing UI benefits for the paid-employed and

| | Worker earnings group | | | | | |
|---|-----------------------|-------|--------|-------|-------|--------|
| | Low | | | High | | |
| | UI-A | UI-P | BI | UI-A | UI-P | BI |
| Transition rates (% change wrt baseline) | | | | | | |
| <i>UP</i> | -0.05 | -0.70 | -0.18 | 0.88 | -0.23 | 1.08 |
| <i>US</i> | -2.57 | 0.45 | 6.15 | 0.12 | -0.10 | 0.84 |
| <i>PS</i> | 0.41 | 5.30 | 23.16 | 0.24 | -0.39 | 2.22 |
| <i>SP</i> | -1.35 | -4.81 | -18.67 | -2.00 | 0.72 | -18.38 |
| <i>PU</i> | -0.23 | -0.13 | -0.17 | 2.03 | 2.52 | 1.14 |
| <i>SU</i> | 23.77 | 17.16 | -7.61 | -0.08 | 0.15 | 0.17 |
| Share of labor force (% change wrt baseline) | | | | | | |
| <i>U</i> | 5.31 | 2.72 | -10.40 | -0.12 | 1.89 | -6.38 |
| <i>P</i> | 1.10 | -1.67 | -12.80 | -0.77 | 1.15 | -14.89 |
| <i>S</i> | -2.74 | 1.57 | 19.64 | 0.09 | -0.15 | 1.74 |

Table 11: Change in transition rates in each policy scenario.

introducing an unconditional cash transfer makes self-employment more attractive, especially with regard to the choices of workers already in employment. This shift is reflected in both *SP* and *PS* transitions. As an example, *SP* transition rate drops by 20% in both skill groups. A key difference of this policy scenario with respect to the UI policies is the absence of moral hazard effects, since the payment associated with the BI policy is not conditional on a worker’s current labor force status. By boosting earnings in employment and providing them with additional insurance in case they become unemployed, this policy strongly favors employment in the “Low” group, both along the *SU* margin (decrease by 13%) and *US* margin (increase by 7%).

Welfare. My main measure of welfare changes is the compensating cash grant, $CG_k^s(a, y)$, which I define as

$$V_k^s(a + CG_k^s(a, y), y) = \tilde{V}_k^s(a, y).$$

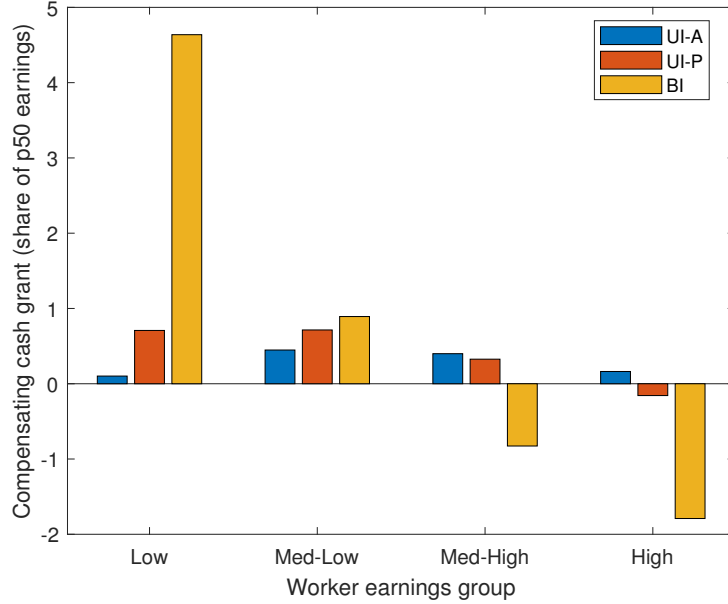


Figure 3: Change in welfare in each policy scenario.

In this last expression, $CG_k^s(a, y)$ denotes the cash transfer making the agent indifferent between the baseline economy (value function V_k^s) and the economy under an alternative policy (value function \tilde{V}_k^s). I report the average of these measures by worker group.

Figure 3 reports the welfare effects of each counterfactual policy, broken down by worker group. The measure of welfare changes, CG_k , is scaled by median earnings in each cluster to make it comparable across groups. Detailed results broken down by worker class k and labor market state s are given in Table C.3.²⁸

The results suggest that introducing a UI policy close to the current UI scheme for the paid-employed has a small, but mostly positive welfare impact. Across policies and groups, workers do not value the compensating cash grant at more than one month of median earnings overall. This result is striking given that, with a relatively high discount rate, precautionary savings are costly for these workers.

The “align” policy scenario (UI-A) is the only policy considered that translates in positive welfare gains for all worker classes and employment states. Given the relatively

²⁸I also report results for the equivalent variation in consumption used in Krusell et al. (2010) with very similar results.

low limit on taxable earnings, the largest welfare gains are obtained for the “Medium-Low” and “Medium-High” earnings groups (median monthly earnings \$2,227 and \$3,840, respectively). For the “Low” group, the welfare gains from having access to an extra source of insurance are in large part offset by the tax burden implied by the cap on taxable earnings. This comes on top of the increase in unemployment documented in Table 11 for this group of workers .

Introducing UI in a more progressive fashion by removing the limit on taxable labor income changes the distribution of welfare gains. The UI-P policy implies that workers in the “Low” and “Medium-low” earnings groups witness the largest welfare gains. These welfare gains have a very similar magnitude in both groups. However, in this policy scenario, workers in the “High” group experience a loss in welfare, since their exposure to unemployment is limited, they hold more wealth to self-insure, and the tax burden is larger for them once the cap on UI contributions is lifted.

The basic income policy translates into much larger welfare changes for the various groups of workers. For instance, the compensating cash grant implied by the BI policy is more than four times the median monthly labor income in the “Low” group. By comparison, the compensating cash grant implied by the UI-P policy for this group is equal to one month of median labor income. However, these sizable welfare gains for the “Low” group come with large losses for workers in the “Medium-high” and “High” earnings groups, who bear the burden of financing the policy. Using the same compensating cash grant metrics, the loss in welfare represents about two months of median earnings in the “High” group.

Summary. Targeting the risks faced by the self-employed in the labor market involves some stark policy trade-offs. Unconditional cash transfers are a potentially useful tool to alleviate these risks for the low earners, and they do not discourage labor supply, but this policy comes at the cost of sizable welfare losses for high earners.

UI policies tend to have much more balanced welfare effects, even positive for all workers in the “align” scenario. They are also difficult to target to low earners. Specifically, the counterfactuals suggest that, without some degree of progressivity in the financing of UI contributions, self-employed workers in the “Low” group barely witness any welfare gains. In addition, UI policies have clear moral hazard effects as they lead workers in the “Low” group to spend substantially more time unemployed.

6 Conclusion

In this paper, I study the labor market risks associated with self-employment. My first set of results are empirical. I show that the large earnings fluctuations experienced by the self-employed are driven both by labor income fluctuations during self-employment spells and transitions to unemployment. In addition, many self-employed do not earn more or have more liquid wealth than employees, thus potentially limiting their resources to self-insure. These findings are for a sample of individuals who account for the majority of earnings within their household.

I then calibrate a search model with precautionary savings that can replicate these empirical regularities. I use the model to assess how different groups of self-employed value a range of policies providing them with extra insurance against labor market risks. My second set of results point to the difficulty of improving the welfare of the self-employed with low earnings without some degree of progressivity in the design of such policies. Specifically, my analysis suggests that extending to the self-employed the UI system currently in place in the US results in (i) an increase in unemployment, (ii) very small welfare gains for the self-employed with low earnings. This type of framework can inform the current policy debate on the rise of “gig economy” work and the degree to which this category of workers can potentially be reached by conventional social insurance programs.

The model in this paper offers a unified framework to study the labor market risks faced

by all non-employees. In doing so, it abstracts from explicitly modeling some dimensions of running a business that are potentially relevant for a subset of the self-employed. An important direction for future work would be to understand how business owners allocate their portfolio between liquid assets, illiquid assets, and their business. Given the inherently risky nature of owning a business, such an extension would provide additional insights into the self-insurance dimension of entering and exiting self-employment.

Appendix

A Data

The main data source is the Survey of Income and Program Participation (SIPP), a survey with detailed information on households' use of welfare programs (Census Bureau, 2014). I use data from the 1996, 2001, 2004, and 2008 panels. This Appendix expands on the definition of paid-employment and self-employment and describes the construction of the “main earners” and “self-employed” samples.

Labor force status. The SIPP contains two key pieces of information to classify individuals as unemployed, paid-, or self-employed: a week-by-week account of their employment state (employed, on layoff, unemployed, non-participating) and information on up to two jobs and two businesses (job/business identifier, wages, profits, incorporation status, and so on).

Employed individuals are classified as paid- or self-employed on the basis of the jobs or businesses for which they report most of their earnings. I start by cleaning the job and business identifiers provided in the survey, so that they are consistent with the start and end date reported for each job or business. For a large majority of workers, only one type of employment is reported: wage work or self-employment. In the case where workers report both working as an employee and having a business, they are assigned the status in which they report the largest earnings during the overlapping spell. In the event workers report more than one job or business, I simply add the corresponding wages/profits to get a measure of earnings in paid- or self-employment. I proceed similarly to define the business characteristics reported in the main text. For categorical variables (incorporation status and number of employees) I use the maximum across the two businesses.

The SIPP reports three types of non-employment week-by-week: on layoff, no job

looking for work, no job not looking for work. To address potential differences in reporting between previously paid- and self-employed workers, I define as unemployment spells all non-employment spells of at most fifty weeks if the individual declares to be looking for a job at some point during the spell or if the spell ends with a transition to employment. The first part of the definition is standard (Chetty, 2008, Appendix B). The second part accounts for the fact former business owners might associate unemployment with the receipt of UI benefits.

Finally, I follow the convention in the Current Population Survey and build a monthly panel of employment state for each individual based on the second week (the first full week) in each month. Keeping only working-age individuals, I check the validity of my definitions by plotting the implied unemployment rate (Figure C.1b) and self-employment rate (Figure C.1a) against data from the Bureau of Labor Statistics (BLS).²⁹ The SIPP data are aggregated using cross-sectional weights. The figures show that my definitions yield sensible unemployment and self-employment rates.

Sample: Main earners. I retain individuals aged at least 25 or at most 65 across the survey. In each household, I only keep the main earner, defined as the individuals with the largest labor earnings across the survey, irrespective of their gender. I also exclude individuals who are mostly out of the labor force over the survey period. The aim of these restrictions is to reduce the sample to the individuals most strongly attached to the labor market. Taken together, these restrictions yield a dataset of 122,662 individuals which are followed on average for 36 months.

Sample: Self-employed. The key analysis and estimation sample is made of individuals who are part of the main earner sample and spending at least one month in

²⁹The corresponding BLS series are LNU02048984 (“Incorporated Self-Employment”), LNU02027714 (“Non-incorporated Self-Employment”), LNU03000000 (“Unemployment”), and LNU02000000 (“Employment”).

self-employment over the duration of the panel. This additional restriction leaves me with a sample of 17,884 workers. All the statistics targeted in the calibration procedure are drawn from the self-employed sample.

B Equilibrium definition

I introduce some additional notation for workers' reservation strategies to formally describe the equilibrium. Let d_k^s be an indicator denoting workers' decision to quit their job or shut down their business given the realization of income. For instance, for a self-employed worker deciding to terminate their activity,

$$d_k^S(a, y) := \mathbb{1} \left\{ V_k^U(a) > \delta_k^S V_k^U(a) + (1 - \delta_k^S) \left[V_k^S(a, y) + \lambda^{SP} \phi_k^{SP}(a, y) \right] \right\}.$$

Let $\rho_k^{ss'}(a, y^s)$ be the reservation income draw from $F_k^{s'}$ making workers indifferent between their current labor force status s and employment in s' . This reservation income is implicitly given by $V_k^s(a, y^s) = V_k^{s'}(a, \rho_k^{ss'}(a, y^s))$.

For each $k \in \{1, \dots, K\}$ and $s \in \{U, B, P, S\}$, a stationary equilibrium is a set of value functions V_k^s and R_k^s , decision rules d_k^S , $\rho_k^{ss'}$, \hat{a}_k^s and \hat{c}_k^s , and a distribution Γ_k across labor force states, wealth and income, such that

1. The R_k^s functions are defined by equations (1), (3), (5), and (6);
2. The choice to terminate employment, $d_k^S(a, y)$, and the reservation income functions $\rho_k^{ss'}(a, y^s)$ solve (1), (3), (5), and (6);
3. The V_k^s functions are defined by equations (2), (4), (7), (8);
4. The asset and consumption choice functions, \hat{a}_k^s and \hat{c}_k^s , solve equations (2), (4), (7), (8);

5. Finally define \mathbf{Q}_k the operator mapping the current distribution of workers to that in the next period. This operator arises from workers' optimal consumption and savings decisions, as well as their reservation strategies. The associated stationary distribution of workers Γ_k solves $\Gamma_k = \mathbf{Q}_k \circ \Gamma_k$.

C Additional figures and tables

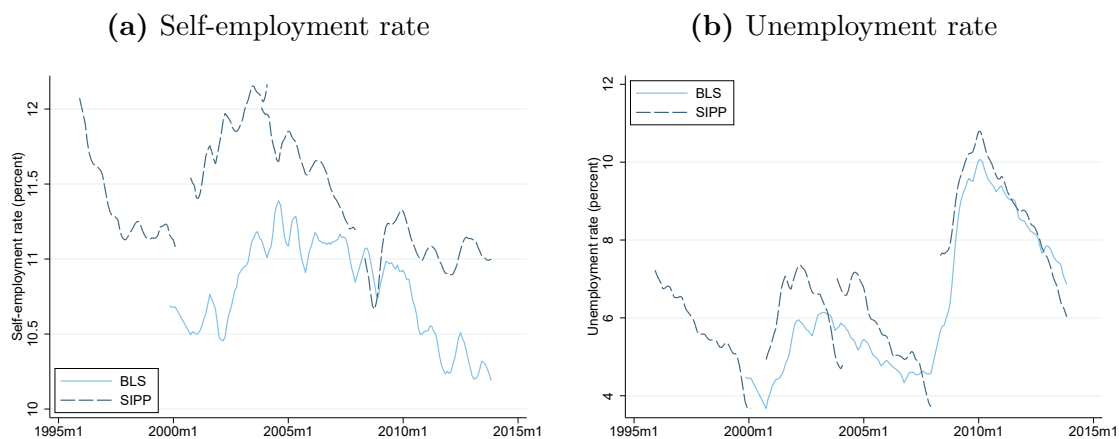


Figure C.1: Benchmark SIPP vs Bureau of Labor Statistics Series (BLS). The self-employment and unemployment rate in the SIPP are based on the definitions of labor force status given in Section 2 and Appendix A. The sample used to build the series include all working age individuals, not just the main earner in each household. Observations are aggregated with cross-sectional weights.

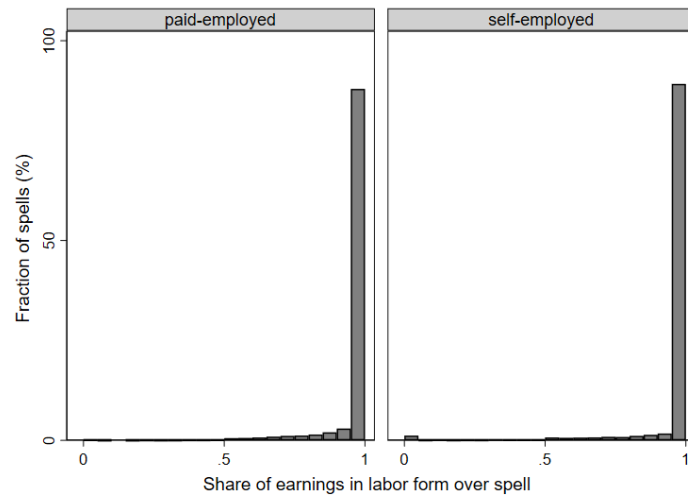


Figure C.2: Share of labor earnings over employment spell, conditional on ascribed labor form.

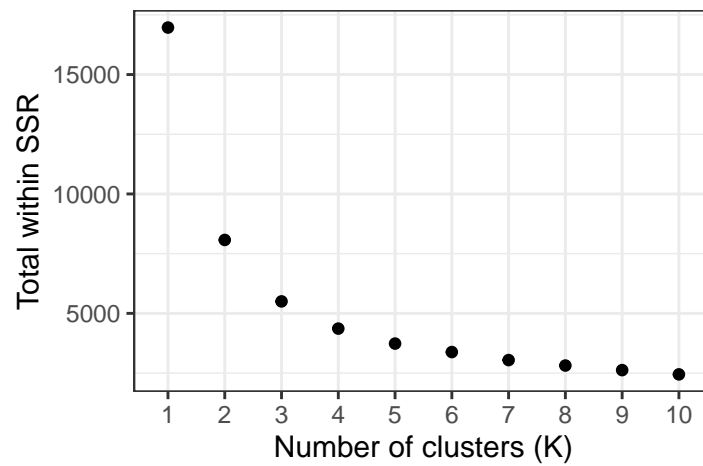


Figure C.3: Change in kmeans fit with number of clusters. The measure of fit is the total within sum of square residuals (SSR).

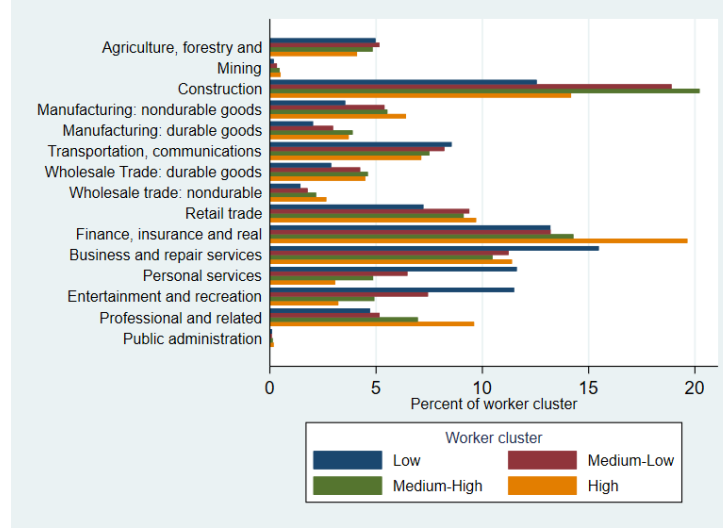


Figure C.4: Industry composition by worker class.

| $\ln Y_{it}^s = \alpha_k^s + \beta_k^s \cdot \ln y_{it}^s + \epsilon_{it}$ | Worker class k (median earnings) | | | |
|--|------------------------------------|---------|---------|---------|
| | \$1,032 | \$2,227 | \$3,840 | \$7,626 |
| Unemployment ($s = U$) | | | | |
| α_k^U | 6.720 | 6.849 | 6.892 | 7.209 |
| Obs. | 9,500 | 8,975 | 5,866 | 2,577 |
| Paid-employment ($s = P$) | | | | |
| α_k^P | 3.460 | 2.978 | 2.894 | 2.437 |
| β_k^P | 0.601 | 0.677 | 0.699 | 0.762 |
| R^2 -adj | 0.232 | 0.343 | 0.435 | 0.603 |
| Obs. | 31,295 | 58,986 | 53,044 | 27,033 |
| Self-employment ($s = S$) | | | | |
| α_k^S | 3.277 | 2.906 | 2.858 | 2.678 |
| β_k^S | 0.629 | 0.688 | 0.705 | 0.736 |
| R^2 -adj | 0.307 | 0.464 | 0.563 | 0.671 |
| Obs. | 54,760 | 99,128 | 113,263 | 82,927 |

Table C.1: Household income functions.

| Parameter | Description | Worker earnings group | | | |
|--------------|--|-----------------------|------------|-------------|-------|
| | | Low | Medium-Low | Medium-High | High |
| δ_k^P | Destruction rate in P | 0.041 | 0.028 | 0.018 | 0.015 |
| δ_k^S | Destruction rate in S | 0.009 | 0.006 | 0.004 | 0.002 |
| μ_k^P | Income draw in P : | 7.022 | 7.647 | 8.206 | 8.864 |
| σ_k^P | $\ln \mathcal{N}(\mu_k^P, \sigma_k^P)$ | 0.232 | 0.405 | 0.327 | 0.392 |
| μ_k^S | Income draw in S : | 6.778 | 7.656 | 8.221 | 8.936 |
| σ_k^S | $\ln \mathcal{N}(\mu_k^S, \sigma_k^S)$ | 0.726 | 0.574 | 0.563 | 0.497 |

Table C.2: Parameters—Specific to worker classes.

| | Worker earnings group | | | | | | | | | | | |
|--|-----------------------|------|------|------------|------|-------|-------------|------|-------|------|-------|-------|
| | Low | | | Medium-Low | | | Medium-High | | | High | | |
| | UI-A | UI-P | BI | UI-A | UI-P | BI | UI-A | UI-P | BI | UI-A | UI-P | BI |
| Equivalent variation in $c - EV_k^s$ (percent) | | | | | | | | | | | | |
| All s | 0.04 | 0.25 | 1.57 | 0.21 | 0.33 | 0.39 | 0.20 | 0.16 | -0.41 | 0.09 | -0.08 | -1.04 |
| P | 0.03 | 0.22 | 1.75 | 0.17 | 0.27 | 0.45 | 0.16 | 0.13 | -0.37 | 0.07 | -0.07 | -1.02 |
| S | 0.03 | 0.28 | 1.74 | 0.20 | 0.33 | 0.46 | 0.19 | 0.16 | -0.37 | 0.09 | -0.09 | -1.02 |
| B | 0.03 | 0.24 | 0.99 | 0.18 | 0.29 | -0.62 | 0.18 | 0.15 | -1.58 | 0.08 | -0.07 | -1.91 |
| U | 0.03 | 0.24 | 1.68 | 0.18 | 0.29 | 0.38 | 0.16 | 0.13 | -0.42 | 0.06 | -0.08 | -1.00 |
| C | 0.56 | 0.79 | — | 1.22 | 1.33 | — | 1.39 | 1.37 | — | 1.37 | 1.22 | — |
| Compensating grant in $a - CG_k^s$ (ratio to median earnings) | | | | | | | | | | | | |
| All s | 0.10 | 0.71 | 4.64 | 0.45 | 0.71 | 0.89 | 0.40 | 0.33 | -0.83 | 0.16 | -0.16 | -1.79 |
| P | 0.09 | 0.63 | 4.38 | 0.37 | 0.59 | 0.35 | 0.32 | 0.27 | -1.21 | 0.13 | -0.12 | -2.03 |
| S | 0.08 | 0.79 | 5.10 | 0.46 | 0.74 | 1.05 | 0.41 | 0.33 | -0.76 | 0.16 | -0.17 | -1.77 |
| B | 0.08 | 0.57 | 2.52 | 0.32 | 0.51 | -0.84 | 0.25 | 0.21 | -1.74 | 0.10 | -0.08 | -1.91 |
| U | 0.07 | 0.62 | 4.56 | 0.33 | 0.54 | 0.70 | 0.32 | 0.26 | -0.82 | 0.13 | -0.17 | -1.96 |
| C | 1.24 | 1.76 | — | 1.81 | 2.00 | — | 1.50 | 1.46 | — | 0.92 | 0.76 | — |

Table C.3: Welfare changes by worker class in response to each alternative policy. The equivalent variation in consumption, $EV_k^s(a, y)$, is defined as $\mathbb{E}_t \sum_{\tilde{t}=0}^{\infty} \beta^{\tilde{t}} u((1 + EV_k^s(a_t, y_t))c_{t+\tilde{t}}) = \mathbb{E}_t \sum_{\tilde{t}=0}^{\infty} \beta^{\tilde{t}} u(\hat{c}_{t+\tilde{t}})$ where \hat{c}_t denotes consumption choices with one of the policies and c_t is consumption in the baseline economy. The compensating cash grant, $CG_k^s(a, y)$, is defined as $V_k^s(a + CG_k^s(a, y), y) = \hat{V}_k^s(a, y)$, where V_k^s and \hat{V}_k^s denote the worker's value function, respectively, in the baseline economy and in the economy with one of the policies.

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