



Two-Factor Hull-White Model for Interest Rate Derivative Products in Bloomberg

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Abstract

This document describes the detailed implementation of the Two-Factor Hull-White model for path-dependent interest rate derivatives

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1 Two-factor Hull-White model

1.1 Model dynamics

1.1.1 Risk neutral measure

Under the risk neutral measure Q , numeraire is given by the money market account, i.e.

$$\mathcal{N}_{rn}(t, \mathcal{F}_t) := \exp \left[\int_0^t r(s) ds \right]. \quad (1)$$

The two-factor Hull-White model for short rate $r(t)$ is

$$r(t) = \theta(t) + \mathbf{1}^\top \mathbf{X}(t), \quad (2)$$

where $\theta(t)$ is a deterministic function to match the initial yield curve. $\mathbf{X}(t)$ is a two dimensional Ornstein-Uhlenbeck process,

$$d\mathbf{X}(t) = -\boldsymbol{\kappa}\mathbf{X}(t)dt + \boldsymbol{\sigma}(t)d\mathbf{W}^Q(t), \quad \mathbf{X}(0) = \mathbf{0}, \quad (3)$$

$$\left(d\mathbf{W}^Q(t) \right) \left(d\mathbf{W}^Q(t) \right)^\top = \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix} dt := \boldsymbol{\rho} dt. \quad (4)$$

with parameters $\boldsymbol{\kappa}$, $\boldsymbol{\sigma}(t)$ and $\boldsymbol{\rho}$.

$$\boldsymbol{\kappa} = \begin{bmatrix} \kappa_1 & 0 \\ 0 & \kappa_2 \end{bmatrix}, \quad (5)$$

$$\boldsymbol{\sigma}(t) = \begin{bmatrix} \sigma_1(t) & 0 \\ 0 & \sigma_2(t) \end{bmatrix}. \quad (6)$$

Constants κ_1 and κ_2 are speeds of mean-reversion for the first and second factor, respectively. We assume that the first factor captures long term dynamics, while the second factor captures short term dynamics. Therefore, $0 \leq \kappa_1 < \kappa_2$. $\sigma_u(t) \geq 0$ ($u = 1, 2$) are volatility related parameters obtained by calibrating to the market volatilities. ρ is the correlation between the two factors.

The solution to the SDE specified in [Eqs. 3](#) and [Eqs. 4](#) is

$$\mathbf{X}(t) = \mathbf{h}(t - t_0)\mathbf{x} + \int_{t_0}^t \mathbf{h}(t - s)\boldsymbol{\sigma}(s)d\mathbf{W}^Q(s), \quad \text{for } \mathbf{X}(t_0) = \mathbf{x} \quad (7)$$

$$= \int_0^t \mathbf{h}(t - s)\boldsymbol{\sigma}(s)d\mathbf{W}^Q(s) \quad \text{for } \mathbf{X}(0) = \mathbf{0}, \quad (8)$$

where

$$\mathbf{h}(t) := e^{-\boldsymbol{\kappa}t} = \begin{bmatrix} h_1(t) & 0 \\ 0 & h_2(t) \end{bmatrix}, \quad (9)$$

$$h_u(t) := h(t; \kappa_u) := e^{-\kappa_u t}.$$

From Eqs. 7,

$$\left[\mathbf{X}(t) | \mathbf{X}(t_0) = \mathbf{x} \right] \sim \mathcal{N} \left(\boldsymbol{\mu}^Q(t_0, t; \mathbf{x}), \boldsymbol{\Sigma}(t_0, t) \right),$$

where

$$\boldsymbol{\mu}^Q(t_0, t; \mathbf{x}) = \mathbf{h}(t - t_0) \mathbf{x} \quad (10)$$

$$\boldsymbol{\Sigma}(t_0, t) = \boldsymbol{\nu}(t) - \mathbf{h}(t - t_0) \boldsymbol{\nu}(t_0) \mathbf{h}(t - t_0). \quad (11)$$

where, $\boldsymbol{\nu}(t)$ is the co-variance of $\mathbf{X}(t)$, i.e.

$$\boldsymbol{\nu}(t) := \text{Var}[\mathbf{X}(t) \mathbf{X}^\top(t)] = \int_0^t \mathbf{h}(t-s) \boldsymbol{\sigma}(s) \boldsymbol{\rho} \boldsymbol{\sigma}(s) \mathbf{h}(t-s) ds = \begin{bmatrix} \nu_{11}(t) & \nu_{12}(t) \\ \nu_{21}(t) & \nu_{22}(t) \end{bmatrix}, \quad (12)$$

$$\nu_{uv}(t) = \rho_{uv} \int_0^t \sigma_u(s) \sigma_v(s) e^{-(\kappa_u + \kappa_v)(t-s)} ds. \quad (13)$$

1.1.2 Terminal forward measure

Under terminal forward measure Q_T , numeraire is zero coupon bond with maturity T , i.e.

$$\mathcal{N}_{Q_T}(t, \mathcal{F}_t) := \mathbb{P}(t, T). \quad (14)$$

Therefore, Eqs. 3 becomes

$$d\mathbf{X}(t) = \left(-\boldsymbol{\kappa} \mathbf{X}(t) - \boldsymbol{\sigma}(t) \boldsymbol{\rho} \boldsymbol{\sigma}(t) \mathbf{H}(T-t) \mathbf{1} \right) dt + \boldsymbol{\sigma}(t) d\mathbf{W}^{Q_T}(t), \quad (15)$$

where,

$$\mathbf{H}(t) := \int_0^t \mathbf{h}(s) ds = \begin{bmatrix} H_1(t) & 0 \\ 0 & H_2(t) \end{bmatrix}. \quad (16)$$

The solution is,

$$\mathbf{X}(t) = -\boldsymbol{\gamma}^{Q_T}(t_0, t) + \mathbf{h}(t - t_0) \mathbf{x} + \int_{t_0}^t \mathbf{h}(t-s) \boldsymbol{\sigma}(s) d\mathbf{W}^{Q_T}(s), \quad \text{for } \mathbf{X}(t_0) = \mathbf{x} \quad (17)$$

$$= -\boldsymbol{\gamma}^{Q_T}(0, t) + \int_0^t \mathbf{h}(t-s) \boldsymbol{\sigma}(s) d\mathbf{W}^{Q_T}(s) \quad \text{for } \mathbf{X}(0) = \mathbf{0} \quad (18)$$

where,

$$\boldsymbol{\gamma}^{Q_T}(t_0, t) := \left[\int_{t_0}^t \mathbf{h}(t-s) \boldsymbol{\sigma}(s) \boldsymbol{\rho} \boldsymbol{\sigma}(s) \mathbf{H}(T-s) ds \right] \mathbf{1}. \quad (19)$$

The drift $\boldsymbol{\gamma}^{Q_T}(t_0, t)$ can be computed as

$$\boldsymbol{\gamma}^{Q_T}(0, t) = \left[\boldsymbol{\nu}^h(t) + \boldsymbol{\nu}(t) \mathbf{H}(T-t) \right] \mathbf{1} \quad \text{for } 0 \leq t \leq T \quad (20)$$

$$\boldsymbol{\gamma}^{Q_T}(t_0, t) = \boldsymbol{\gamma}^{Q_T}(0, t) - \mathbf{h}(t - t_0) \boldsymbol{\gamma}^{Q_T}(0, t_0) \quad \text{for } 0 \leq t_0 \leq t \leq T. \quad (21)$$

From [Eqs. 17](#),

$$\left[\mathbf{X}(t) | \mathbf{X}(t_0) = \mathbf{x} \right] \sim \mathcal{N} \left(\boldsymbol{\mu}^{Q^T}(t_0, t; \mathbf{x}), \boldsymbol{\Sigma}(t_0, t) \right),$$

where

$$\boldsymbol{\mu}^{Q^T}(t_0, t; \mathbf{x}) = -\boldsymbol{\gamma}^{Q^T}(t_0, t) + \mathbf{h}(t - t_0)\mathbf{x} \quad (22)$$

and $\boldsymbol{\Sigma}(t_0, t)$ is the same as that in risk neutral measure in [Eqs. 11](#).

1.1.3 Spot-Libor measure

Under spot-Libor measure Q_β , numeraire is given by the discretely rebalanced bank account (see [\[Brigo & Mercurio\]](#), Section 6.3)

$$\mathcal{N}_\beta(t, \mathcal{F}_t) := \frac{\mathbb{P}(t, T_{\beta(t)})}{\prod_{j=1}^{\beta(t)} \mathbb{P}(T_{j-1}, T_j)} \quad (23)$$

where $0 = T_0 < T_1 < \dots$ are fixed as our discrete tenor structure, $\beta(t)$ is the smallest index such that $t \leq T_{\beta(t)}$. From [Eqs. 14](#) and [Eqs. 23](#), note that the stochastic part of \mathcal{N}_β is the same as that in $\mathcal{N}_{Q_{\beta(t)}}$. Therefore, under the spot-Libor measure, the dynamics changes to

$$d\mathbf{X}(t) = \left(-\kappa \mathbf{X}(t) - \boldsymbol{\sigma}(t) \boldsymbol{\rho} \boldsymbol{\sigma}(t) \mathbf{H}(T_{\beta(t)} - t) \mathbf{1} \right) dt + \boldsymbol{\sigma}(t) d\mathbf{W}^{Q_\beta}(t). \quad (24)$$

Integrating [Eqs. 24](#), we obtain

$$\mathbf{X}(t) = -\boldsymbol{\gamma}^{Q_\beta}(t_0, t) + \mathbf{h}(t - t_0)\mathbf{x} + \int_{t_0}^t \mathbf{h}(t - s) \boldsymbol{\sigma}(s) d\mathbf{W}^{Q_\beta}(s), \quad \text{for } \mathbf{X}(t_0) = \mathbf{x} \quad (25)$$

$$= -\boldsymbol{\gamma}^{Q_\beta}(0, t) + \int_0^t \mathbf{h}(t - s) \boldsymbol{\sigma}(s) d\mathbf{W}^{Q_\beta}(s) \quad \text{for } \mathbf{X}(0) = \mathbf{0} \quad (26)$$

where $\boldsymbol{\gamma}^{Q_\beta}(t_0, t)$ can be calculated stepwise. When $\beta(s) = \beta(t)$ for $t_0 < s < t$, we have $\boldsymbol{\gamma}^{Q_\beta}(t_0, t) = \boldsymbol{\gamma}^{Q_{T_{\beta(t)}}}(t_0, t)$, which is defined in [Eqs. 19](#).

Therefore, the conditional distribution is

$$\left[\mathbf{X}(t) | \mathbf{X}(t_0) = \mathbf{x} \right] \sim \mathcal{N} \left(\boldsymbol{\mu}^{Q_\beta}(t_0, t; \mathbf{x}), \boldsymbol{\Sigma}(t_0, t) \right),$$

where

$$\boldsymbol{\mu}^{Q_\beta}(t_0, t; \mathbf{x}) = -\boldsymbol{\gamma}^{Q_\beta}(t_0, t) + \mathbf{h}(t - t_0)\mathbf{x}. \quad (27)$$

References

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