

# Lecture 3: Normal Linear Models — Types and Matrix Form

## MATH3823 Generalised Linear Models

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# Reading

**Course notes:** Chapter 2, Sections 2.3–2.4

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# Types of Normal Linear Models

$p$	Explanatory Variables	Model Name
1	Quantitative	Simple linear regression
$> 1$	Quantitative	Multiple linear regression
1	Dichotomous (2 levels)	Two-sample $t$ -test
1	Polytomous ( $k$ levels)	One-way ANOVA
$> 1$	Qualitative	Multi-way ANOVA
$> 1$	Mixed (quant. + qual.)	ANCOVA

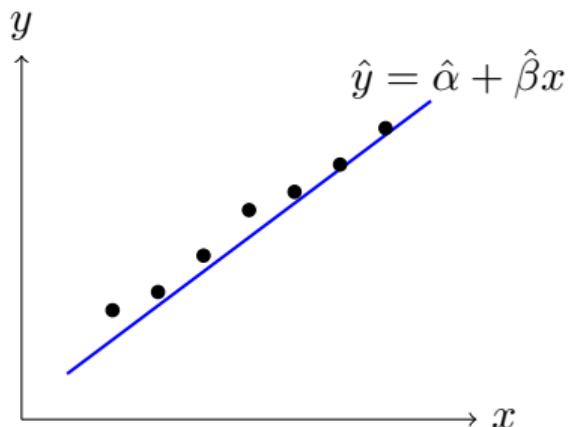
**Key insight:** All of these are *special cases* of the general linear model.

# Simple Linear Regression

**Model:**

$$y_i = \alpha + \beta x_i + \epsilon_i, \quad \epsilon_i \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma^2)$$

**Example:** Height vs. weight, dose vs. response



# Multiple Linear Regression

**Model:**

$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \cdots + \beta_p x_{ip} + \epsilon_i$$

**Example:** House price depending on size, age, location, etc.

**Interpretation:**

- $\beta_j$  = change in  $\mathbb{E}[Y]$  per unit increase in  $x_j$ , *holding other variables constant*
- This is a **partial** or **adjusted** effect

**Warning:** In observational studies, “holding constant” is conceptual, not causal.

## Two-Sample $t$ -Test as a Linear Model

**Setting:** Compare means of two groups

**Traditional formulation:**

$$Y_{1j} \sim \mathcal{N}(\mu_1, \sigma^2), \quad j = 1, \dots, n_1$$

$$Y_{2j} \sim \mathcal{N}(\mu_2, \sigma^2), \quad j = 1, \dots, n_2$$

**As a linear model:**

$$y_i = \alpha + \gamma \cdot \text{Group}_i + \epsilon_i$$

where  $\text{Group}_i = 0$  for group 1,  $\text{Group}_i = 1$  for group 2.

**Parameters:**

- $\alpha = \mu_1$  (mean of reference group)
- $\gamma = \mu_2 - \mu_1$  (difference in means)

# One-Way ANOVA as a Linear Model

**Setting:** Compare means across  $k$  groups

**Model:**

$$y_{ij} = \mu + \alpha_i + \epsilon_{ij}, \quad i = 1, \dots, k, \quad j = 1, \dots, n_i$$

**Problem:** The model is **overparameterized**.

- We have  $k + 1$  parameters  $(\mu, \alpha_1, \dots, \alpha_k)$
- But only  $k$  group means to estimate

**Solution:** Add a constraint (identifiability condition)

## Parameter Constraints

### Option 1: Corner constraint (R default)

$$\alpha_1 = 0$$

- Group 1 is the reference category
- $\mu$  = mean of group 1
- $\alpha_i$  = difference between group  $i$  and group 1

### Option 2: Sum-to-zero constraint

$$\sum_{i=1}^k \alpha_i = 0$$

- $\mu$  = grand mean (average of group means)
- $\alpha_i$  = deviation of group  $i$  from grand mean

# Matrix Formulation: The General Linear Model

For  $n$  observations:

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$$

where:

- $\mathbf{Y} = (Y_1, \dots, Y_n)'$  is the  $n \times 1$  response vector
- $\mathbf{X}$  is the  $n \times p$  **design matrix**
- $\boldsymbol{\beta} = (\beta_1, \dots, \beta_p)'$  is the  $p \times 1$  parameter vector
- $\boldsymbol{\epsilon} = (\epsilon_1, \dots, \epsilon_n)'$  is the  $n \times 1$  error vector

Assumptions:

$$\boldsymbol{\epsilon} \sim \mathcal{N}_n(\mathbf{0}, \sigma^2 \mathbf{I}_n)$$

## Design Matrix: Simple Linear Regression

**Model:**  $y_i = \alpha + \beta x_i + \epsilon_i$

$$\underbrace{\begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix}}_{\mathbf{Y}} = \underbrace{\begin{pmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_n \end{pmatrix}}_{\mathbf{X}} \underbrace{\begin{pmatrix} \alpha \\ \beta \end{pmatrix}}_{\boldsymbol{\beta}} + \underbrace{\begin{pmatrix} \epsilon_1 \\ \epsilon_2 \\ \vdots \\ \epsilon_n \end{pmatrix}}_{\boldsymbol{\epsilon}}$$

- First column of  $\mathbf{X}$  is all 1s (for intercept)
- Second column contains the  $x$  values

## Design Matrix: One-Way ANOVA

**Model:**  $y_{ij} = \mu + \alpha_i + \epsilon_{ij}$  with  $k = 3$  groups

**Full (overparameterized) design matrix:**

$$\mathbf{X}_{\text{full}} = \begin{pmatrix} 1 & 1 & 0 & 0 \\ 1 & 1 & 0 & 0 \\ 1 & 0 & 1 & 0 \\ 1 & 0 & 1 & 0 \\ 1 & 0 & 0 & 1 \\ 1 & 0 & 0 & 1 \end{pmatrix} \quad (\text{not full rank})$$

**With corner constraint ( $\alpha_1 = 0$ ):**

$$\mathbf{X} = \begin{pmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 1 & 0 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \end{pmatrix} \quad \boldsymbol{\beta} = \begin{pmatrix} \mu \\ \alpha_2 \\ \alpha_3 \end{pmatrix}$$

# Constructing the Design Matrix

## Recipe:

- ① Start with a column of 1s (intercept)
- ② For each **quantitative** variable: add one column of values
- ③ For each **qualitative** variable with  $k$  levels:
  - Create  $k$  dummy (indicator) columns
  - Remove one column to avoid singularity
- ④ For interactions: multiply corresponding columns element-wise

**Result:** A full-rank  $n \times p$  matrix where  $p = \text{number of free parameters}$ .

# Least Squares Solution in Matrix Form

Residual sum of squares:

$$\text{RSS}(\boldsymbol{\beta}) = (\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})'(\mathbf{Y} - \mathbf{X}\boldsymbol{\beta})$$

Normal equations:

$$\mathbf{X}'\mathbf{X}\hat{\boldsymbol{\beta}} = \mathbf{X}'\mathbf{Y}$$

Least squares estimator:

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$$

Properties:

- $\mathbb{E}[\hat{\boldsymbol{\beta}}] = \boldsymbol{\beta}$  (unbiased)
- $\text{Var}[\hat{\boldsymbol{\beta}}] = \sigma^2(\mathbf{X}'\mathbf{X})^{-1}$

## Fitted Values and Residuals

Fitted values:

$$\hat{\mathbf{Y}} = \mathbf{X}\hat{\boldsymbol{\beta}} = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} = \mathbf{H}\mathbf{Y}$$

where  $\mathbf{H} = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'$  is the **hat matrix**.

Residuals:

$$\mathbf{r} = \mathbf{Y} - \hat{\mathbf{Y}} = (\mathbf{I} - \mathbf{H})\mathbf{Y}$$

Error variance estimate:

$$\hat{\sigma}^2 = \frac{\mathbf{r}'\mathbf{r}}{n-p} = \frac{\text{RSS}}{n-p}$$

# Viewing the Design Matrix in R

```
# Create a factor variable  
group <- factor(c("A", "A", "B", "B", "C", "C"))  
  
# See the design matrix  
model.matrix(~ group)
```

## Output:

	(Intercept)	groupB	groupC
1	1	0	0
2	1	0	0
3	1	1	0
4	1	1	0
5	1	0	1
6	1	0	1

Note: Group A is the reference (no column for it).

# Summary

## Key points:

- Many statistical tests are special cases of linear models
- The matrix formulation  $\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \epsilon$  unifies all cases
- Design matrix  $\mathbf{X}$  encodes the model structure
- Qualitative variables require dummy coding
- Constraints are needed to avoid overparameterization
- Least squares solution:  $\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$

**Next lecture:** Model notation, R formula syntax, and fitting in practice.